

Solutions to Chapter 3 Exercises

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June 13, 2021

3.05 Let $\hat{\mathbf{e}}$ be the OLS residual from a regression of \mathbf{Y} on \mathbf{X} . Find the OLS coefficient from a regression of $\hat{\mathbf{e}}$ on \mathbf{X} .

Answer:

Step 1: Here, \hat{e} takes the place of Y in the formula for $\hat{\beta}$.

$$\begin{aligned}\tilde{\beta} &= (X'X)^{-1}X'\hat{\mathbf{e}} \\ &= (X'X)^{-1} \left[X'Y - X'X\hat{\beta} \right] \\ &= (X'X)^{-1} \left[X'Y - X'X(X'X)^{-1}X'Y \right] \\ &= (X'X)^{-1} [X'Y - X'Y] \\ &= 0\end{aligned}$$

3.06 Let $\hat{\mathbf{Y}} = \mathbf{X}(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{Y}$. Find the OLS coefficient from a regression of $\hat{\mathbf{Y}}$ on \mathbf{X} .

Answer:

Step 1: Plug into regression formula and expand.

$$\begin{aligned}\tilde{\beta} &= (X'X)^{-1}X'\hat{\mathbf{Y}} \\ &= [(X'X)^{-1}X'X] (X'X)^{-1}X'Y \\ &= I(X'X)^{-1}X'Y \\ &= \hat{\beta}\end{aligned}$$

3.11 Show that when \mathbf{X} contains a constant, $n^{-1} \sum_{i=1}^n \hat{Y}_i = \bar{Y}$.

Answer:

Step 1: Note that $Y_i = \hat{Y}_i + \hat{e}_i$. Also, $\bar{Y} := n^{-1} \sum_{i=1}^n Y_i$.

$$\begin{aligned}\sum_{i=1}^n Y_i &= \sum_{i=1}^n \hat{Y}_i + \sum_{i=1}^n \hat{e}_i \\ \Leftrightarrow n^{-1} \sum_{i=1}^n Y_i &= n^{-1} \sum_{i=1}^n \hat{Y}_i + n^{-1} \sum_{i=1}^n \hat{e}_i \\ \Leftrightarrow \bar{Y} &= n^{-1} \sum_{i=1}^n \hat{Y}_i + n^{-1} \sum_{i=1}^n \hat{e}_i\end{aligned}$$

Step 2: The equation specified in the problem is then only true if $\sum_{i=1}^n \hat{e}_i = 0$.
 X is a matrix of the form:

$$\begin{bmatrix} x_{11} & x_{12} & \cdots & x_{1,k-1} & 1 \\ x_{21} & x_{22} & \cdots & x_{2,k-1} & 1 \\ \vdots & \vdots & & \vdots & \vdots \\ x_{n1} & x_{n2} & \cdots & x_{n,k-1} & 1 \end{bmatrix}$$

Partition X into $[X^* \ O]$ such that X^* is a $n \times k - 1$ matrix of the x observations and O is a $n \times 1$ vector of ones. We want the sum of the residuals to be zero, and recall that we showed (in matrix form) in Exercise 3.05 that $X'\hat{e} = 0$. Also note that the transpose of a partitioned matrix is like the transpose of a regular matrix except the elements are also transposed. So,

$$\begin{aligned} X'\hat{e} &= \begin{bmatrix} X^{*'} \\ O' \end{bmatrix} \hat{e} \\ &= \begin{bmatrix} X^{*'}\hat{e} \\ O'\hat{e} \end{bmatrix} \\ &= 0 \end{aligned}$$

Then $O'\hat{e} = \sum_{i=1}^n \hat{e}_i = 0$.

3.14 Let $\hat{\beta}_n = (\mathbf{X}'_n \mathbf{X}_n)^{-1} \mathbf{X}'_n \mathbf{Y}_n$ denote the OLS estimate where \mathbf{Y}_n is $n \times 1$ and \mathbf{X}_n is $n \times k$. Prove that the OLS estimate computed using an additional observation (Y_{n+1}, X_{n+1}) is

$$\hat{\beta}_{n+1} = \hat{\beta}_n + \frac{1}{1 + X'_{n+1}(X'_n X_n)^{-1} X_{n+1}} (X'_n X_n)^{-1} X_{n+1} (Y_{n+1} - X'_{n+1} \hat{\beta}_n)$$

Answer:

Step 1: Apply the Woodbury Matrix Identity. Let

$$\begin{aligned} A &= X'_n X_n \\ B &= X_{n+1} \\ D &= X'_{n+1} \end{aligned}$$

The choice of D is a bit arbitrary. I think setting the last term equal to C would work, but this way the ‘1’ is already present and the equation is a bit easier to manipulate. Then, since $C = I = 1$ because a single observation is 1 by 1.

$$\begin{aligned}(A + BCD)^{-1} &= A^{-1} - A^{-1}BC(C + CDA^{-1}BC)^{-1}CDA^{-1} \\ &= A^{-1} - A^{-1}B(1 + DA^{-1}B)DA^{-1}\end{aligned}$$

Step 2: Simplify. The terms of the expanded equation are:

$$\begin{aligned}Term1 &= A^{-1}X'_nY_n \\ &= (X'_nX_n)^{-1}X'_nY_n \\ &= \hat{\beta}_n\end{aligned}$$

$$\begin{aligned}Term2 &= A^{-1}X_{n+1}Y_{n+1} \\ &= (X'_nX_n)^{-1}(X_{n+1}Y_{n+1})\end{aligned}$$

$$\begin{aligned}Term3 &= (X'_nX_n)^{-1}X_{n+1} [1 + X'_{n+1}(X'_nX_n)^{-1}X_{n+1}]^{-1} X'_{n+1}(X'_nX_n)^{-1}X'_nY_n \\ &= (X'_nX_n)^{-1}X_{n+1} [1 + X'_{n+1}(X'_nX_n)^{-1}X_{n+1}]^{-1} X'_{n+1}\hat{\beta}_n\end{aligned}$$

$$Term4 = (X'_nX_n)^{-1}X_{n+1} [1 + X'_{n+1}(X'_nX_n)^{-1}X_{n+1}] X'_{n+1}(X'_nX_n)^{-1}X_{n+1}Y_{n+1}$$

Step 3: Combine terms. Note that the term starting with ‘1 + ...’ is a scalar, call it L . Call the term $Z := (X'_nX_n)^{-1}X_{n+1}$. Note also that $L = 1 + X'_{n+1}Z$. The observations are also scalars and may be rearranged. So:

$$\begin{aligned}
\hat{\beta}_{n+1} &= Term1 + Term2 - Term3 - Term4 \\
&= \hat{\beta}_n + ZY_{n+1} - \frac{1}{L}ZX'_{n+1}\hat{\beta}_n - \frac{1}{L}ZX'_{n+1}ZY_{n+1} \\
&= \hat{\beta}_n + ZY_{n+1}\frac{L}{L} - \frac{1}{L}ZZX'_{n+1}Y_{n+1} - -\frac{1}{L}ZX'_{n+1}\hat{\beta}_n \\
&= \hat{\beta}_n + \frac{1}{L}(ZY_{n+1} + ZZX'_{n+1}Y_{n+1} - ZZX'_{n+1}Y_{n+1} - ZX'_{n+1}\hat{\beta}_n) \\
&= \hat{\beta}_n + \frac{1}{L}Z(Y_{n+1} - X'_{n+1}\hat{\beta}_n)
\end{aligned}$$

3.19 For the intercept-only model $Y_i = \beta + e_i$, show that the leave-one-out prediction error is

$$\tilde{e} = \left(\frac{n}{n-1} \right) (Y_i - \bar{Y})$$