

1 Set Theory

Definition. (*Set Operations*) Given two sets A, B , we have that

- $A - B = A \cap B^c$
- $(A \cap B) \cup (A \cap B^c) = A$
- *Neither* $\implies A^c \cap B^c = (A \cup B)^c$
- *Xor* $\implies (A \cap B^c) \cup (A^c \cap B)$
- *At least one* $\implies A \cup B$

2 Combinatorics

Definition (Counting Tools).

- (a) $\mathbf{m} \times \mathbf{n}$: Number of pairs between m and n items.
- (b) $\mathbf{m}^{\mathbf{n}}$: Number of ways to fill n slots with m objects.
- (c) $\mathbf{P}_{\mathbf{r}}^{\mathbf{n}} = \frac{\mathbf{n}!}{(\mathbf{n}-\mathbf{r})!}$: Number of ways of ordering n distinct objects taken r at a time.
- (d) $\binom{\mathbf{n}}{\mathbf{r}} = \frac{\mathbf{n}!}{\mathbf{r}!(\mathbf{n}-\mathbf{r})!}$: Number of subsets, each of size r , that can be formed from n objects.
- (e) $\binom{\mathbf{n}}{\mathbf{n}_1, \mathbf{n}_2, \dots, \mathbf{n}_k} = \frac{\mathbf{n}!}{\mathbf{n}_1! \mathbf{n}_2! \dots \mathbf{n}_k!}$: Number of ways of partitioning n distinct objects into k distinct groups containing n_1, n_2, \dots, n_k objects.
This has restriction: $\sum_{i=1}^k n_i = n$

3 Conditional Probability

Definition (A given B occurred). $P(A | B) = \frac{P(A \cap B)}{P(B)}$

Definition (Independence). *Two events A, B are independent \iff*

- (i) $P(A | B) = P(A) \iff P(B | A) = P(B)$
- (ii) $P(A \cap B) = P(A)P(B)$

Definition (Probability).

Corollary (Properties).

- (i) $P(A | A \cup B) = \frac{P(A)}{P(A \cup B)}$
- (ii) $P(A \cap B | A \cup B) = \frac{P(A \cap B)}{P(A \cup B)}$
- (iii) $P(A^c | B) = 1 - P(A | B)$
- (iv) $P(A) = P(A \cap B) + P(A \cap B^c) \equiv P(A | B)P(B) + P(A | B^c)P(B^c)$
- (v) If A, B are independent then $\implies (A^c, B), (A, B^c)$ and (A^c, B^c) are all independent.
- (vi) $P(A \cap B) = P(B | A)P(A) = P(A | B)P(B)$
- (vii) $P(A \cap B \cap C) = P(A)P(B | A)P(C | B \cap A)$

3.1 Bayesian

Theorem (Total Law of Probability). *If $\{B_1, B_2, \dots, B_n\}$ is a partition of $S \implies$*

$$P(A) = \sum_{k=1}^n P(A | B_k)P(B_k) = \sum_{k=1}^n P(A \cap B_k)$$

Theorem (Bayes' Theorem).

$$P(B_j | A) = \frac{P(A | B_j)P(B_j)}{\sum_k P(A | B_k)P(B_k)}; P(A | B) = \frac{P(B | A)P(A)}{P(B)}$$

4 Discrete | Continuous Random Variables

Let D and C denote a discrete and continuous random variable respectively.

Definition (Expected value μ). $E[Y] = \begin{cases} D & \implies \sum_y y \cdot P(Y = y) \\ C & \implies \int_{-\infty}^{\infty} y \cdot f_Y(y) dy \end{cases}$

Definition (Variance σ^2). $\text{Var}[Y] = E[Y^2] - E[Y]^2$

Definition (Quantile). ϕ_p : p^{th} quantile of Y . $P(Y \leq \phi_p) = p$ or the probability that Y falls in the ϕ_p quantile.

Corollary (Properties).

- $E[c \in \mathbb{R}] = c$
- $\text{Var}[aY + b] = a^2 \text{Var}[Y]$

Definition (Probability Density Function). *The probability per unit length.*

$$\begin{cases} D & \implies P(Y = y) \\ C & \implies f_Y(y) \end{cases}$$

Definition (Cumulative Distribution Function). *The probability being in a given interval.*

$$\begin{cases} D & \implies P(Y \leq y_*) \sum^{y_*} P(Y = y) \\ C & \implies P(a < Y < b) = F_Y(b) - F_Y(a) = \int_a^b f_Y(y) dy \end{cases}$$

4.0.1 Moment Generating Function

Definition (MGF of Y).

$$M_Y(t) = E[e^{tY}] = \begin{cases} D & \implies \sum_y e^{ty} P(Y = y) \\ C & \implies \int e^{ty} f_Y(y) dy \end{cases}$$

Definition (k^{th} Moment of Y).

$$E[Y^k] = \left. \frac{d^k}{dt^k} M_Y(t) \right|_{t=0} = \begin{cases} k=1 & \implies E[Y] \\ k=2 & \implies E[Y^2] \\ \vdots & \\ k=n & \implies E[Y^n] \end{cases}$$

4.1 Joint Distributions

Definition (Joint CDF). $F(y_1, y_2) = P(Y_1 \leq y_1, Y_2 \leq y_2) =$

$$\int_{-\infty}^{y_2} \int_{-\infty}^{y_1} f(t_1, t_2) dt_1 dt_2$$

Definition (Joint PDF). $f(y_1, y_2)$

Definition (Marginal PDF).

$$(1) f_{y_1}(y_1) = \int f(y_1, y_2) dy_2 \quad (2) f_{y_2}(y_2) = \int f(y_1, y_2) dy_1$$

Definition (Conditional Density).

$$f(y_1 | y_2) = \frac{f(y_1, y_2)}{f_{y_2}(y_2)} = \frac{f(y_1, y_2)}{\int f(y_1, y_2) dy_1}$$

Definition (Conditional Distribution).

$$F(y_1 | y_2) = P(Y_1 \leq y_1 | Y_2 = y_2) = \int_{-\infty}^{y_1} f(t_1 | y_2) dt_1 = \int_{-\infty}^{y_1} \frac{f(t_1, y_2)}{f_{Y_2}(y_2)} dt_1$$

Theorem (Independence). Y_1, Y_2 are independent

- $\iff F(y_1, y_2) = F_{y_1}(y_1)F_{y_2}(y_2)$
- $\iff f(y_1, y_2) = f_{y_1}(y_1)f_{y_2}(y_2)$
- $\iff f(y_1 | y_2) = f_{y_1}(y_1)$
- $\iff E[Y_1, Y_2] = E[Y_1]E[Y_2]$
- $\iff f(y_1, y_2) = g(y_1)h(y_2)$

4.2 Expectation

Definition (Joint Expectation and Variance).

$$(1) E[Y_1 Y_2] = \int \int y_1 y_2 f(y_1, y_2) dy_1 dy_2 \quad (2) \text{Var} [aY_1 \pm bY_2] = a^2 \text{Var} [Y_1] + b^2 \text{Var} [Y_2] \pm 2ab \text{Cov} [Y_1, Y_2]$$

Definition (Covariance).

$$\text{Cov} [Y_1, Y_2] = E[Y_1 Y_2] - E[Y_1] E[Y_2]$$

Corollary (Corollary). If Y_1, Y_2 independent $\implies \text{Cov} [Y_1, Y_2] = 0$

4.2.1 Conditional Expectation

$$E[Y_1 | Y_2 = y_2] = \int y_1 f(y_1 | y_2) dy_1 = \int y_1 \frac{f(y_1, y_2)}{\int f(y_1, y_2) dy_1} dy_1$$

5 Order Statistics

Definition (Order Statistics). Given a set of i.i.d. ordered random variables $Y = \{Y_1 \rightarrow Y_n\}$, denote $\min / \max Y = Y_{(1)}, Y_{(n)}$. Each Y_k has pdf $f_Y(y)$ and cdf $F_Y(y)$.¹

Definition (Sample Mean \bar{Y}). Given $Y = \{Y_1 \rightarrow Y_n\}$ i.i.d., then $\bar{Y} = \frac{1}{n} \sum_{k=1}^n Y_k$

Definition (Sample Median \bar{Y}).

Theorem (PDF and CDF of $Y_{(n)}$).

$$F_{Y_{(n)}}(y) = P(Y_{(n)} \leq y) = P(Y_1 \leq y) P(Y_2 \leq y) \dots P(Y_n \leq y) = [F(y)]^n \quad (1)$$

$$f_{Y_{(n)}}(y) = \frac{d}{dy} [F(y)]^n = n [F_Y(y)]^{n-1} f_Y(y) \quad (2)$$

Theorem (PDF and CDF of $Y_{(1)}$).

$$F_{Y_{(1)}}(y) = 1 - [1 - F_Y(y)]^n \quad (3)$$

$$f_{Y_{(1)}}(y) = n [1 - F_Y(y)]^{n-1} f_Y(y) \quad (4)$$

Theorem (PDF of $Y_{(k)}$).

$$f_{Y_{(k)}} = \frac{n!}{(k-1)!(n-k)!} [F_Y(y)]^{k-1} [1 - F_Y(y)]^{n-k} \quad (5)$$

Theorem (Central Limit Theorem). Given $Y = \{Y_1, Y_2, \dots, Y_n\}$ i.i.d. with any distribution, and $E[Y_k] = \mu$ and $\text{Var} [Y_k] = \sigma^2 \implies$

$$\frac{\bar{Y} - \mu}{\sigma/\sqrt{n}} = \sqrt{\text{Var} [\bar{Y}]} \approx Y \sim \text{Normal}(0, 1)$$

Theorem (Normal Approximation to Binomial). Given $Y \sim \text{Binomial}(n, p)$ and $X_{i \in [1, n]} \sim \text{Bernoulli}(p)$

$$\implies X = \sum_{i=1}^n X_i \sim \text{Binomial}(n, p) \implies \frac{Y}{n} = \bar{X} \implies E[Y/n] = p \text{ and } \text{Var} [Y/n] = \frac{p(1-p)}{n}$$

Theorem (Continuity Correction).

$$\begin{cases} P(Y < y) & \implies P(Y < y - 0.5) \\ P(Y \leq y) & \implies P(Y \leq y + 0.5) \\ P(Y = y) & \implies P(y - 0.5 < Y < y + 0.5) \\ P(Y > y) & \implies P(Y > y + 0.5) \\ P(Y \geq y) & \implies P(Y \geq y - 0.5) \end{cases}$$

¹Note that Y_k is a random variable, while $Y_{(k)}$ is an order statistic.

6 Discrete Distribution Dictionary

- $Y \sim \text{Binomial}(n, p)$: Observing $y \in [0, n]$ successes in fixed n trials.

$$P(Y = y) = \binom{n}{y} p^y (1-p)^{n-y} \implies \mu = n \cdot p \text{ and } \sigma^2 = np(1-p)$$

$$M_Y(t) = [pe^t + (1-p)]^n$$

- $Y \sim \text{NB}(r, p)$: Observing the fixed r^{th} success in $y \in [r, \infty)$ trials.

$$P(Y = y) = \binom{y-1}{r-1} p^r (1-p)^{y-r} \implies \mu = \frac{r}{p} \text{ and } \sigma^2 = \frac{r(1-p)}{p^2}$$

$$M_Y(t) = \left[\frac{pe^t}{1 - (1-p)e^t} \right]^r$$

- $Y \sim \text{Geometric}(p) = \text{NB}(1, p)$: Observing 1 success in $y \in [1, \infty)$ trials.

$$P(Y = y) = (1-p)^{y-1} p \implies \mu = \frac{1}{p} \text{ and } \sigma^2 = \frac{1-p}{p^2}$$

$$M_Y(t) = \frac{pe^t}{1 - (1-p)e^t}$$

- $Y \sim \text{Hypergeometric}(N, r, n)$: Observing $y \in \begin{cases} [0, n] & \text{if } n \leq r, \\ [0, r] & \text{if } n > r \end{cases}$ successes in n draws, without replacement, from a population of N that contains r success states.

$$P(Y = y) = \frac{\binom{r}{y} \binom{N-r}{n-y}}{\binom{N}{n}} \implies \mu = \frac{r \cdot n}{N} \text{ and } \sigma^2 = n \left(\frac{r}{N} \right) \left(\frac{N-r}{N} \right) \left(\frac{N-n}{N-1} \right)$$

- $Y \sim \text{Poisson}(\lambda)$: Observing $y \in [0, \infty)$ independent events that occur with a constant mean rate of λ in a fixed interval of time or area. Note: $Y \sim \text{Poisson}(n \cdot p) = Y \sim \text{Binomial}(n, p)$ if n is very large, or p is small. That is to say, it may be used to approximate the binomial distribution.

$$P(Y = y) = \frac{\lambda^y e^{-\lambda}}{y!} \implies \mu = \lambda \text{ and } \sigma^2 = \lambda$$

$$M_Y(t) = e^{\lambda(e^t - 1)}$$

7 Continuous Distributions Dictionary (ADD CDFs)

- $Y \sim \text{Uniform}(a, b)$:

$$f_Y(y) = \frac{1}{b-a} \text{ and } F_Y(y) = \frac{y}{b-a}; \quad y \in [a, b] \quad (6)$$

$$\mu = \frac{a+b}{2} \text{ and } \sigma^2 = \frac{(b-a)^2}{12} \quad (7)$$

$$M_Y(t) = \frac{e^{tb} - e^{ta}}{t(b-a)} \quad (8)$$

- $Y \sim \text{Normal}(\mu, \sigma^2)$:

$$f_Y(y) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left[-\frac{(y-\mu)^2}{2\sigma^2}\right]; \quad y \in (\pm\infty) \implies \mu = \mu \text{ and } \sigma^2 = \sigma^2 \quad (9)$$

$$M_Y(t) = \exp\left[\mu t + \frac{t^2\sigma^2}{2}\right] \quad (10)$$

- $Y \sim \text{Gamma}(\alpha, \beta)$:

$$f_Y(y) = \frac{y^{\alpha-1} \exp[-y\beta^{-1}]}{\Gamma(\alpha)\beta^\alpha}; \quad y \in (0, \infty) \implies \mu = \alpha\beta \text{ and } \sigma^2 = \alpha\beta^2 \quad (11)$$

$$M_Y(t) = (1 - \beta t)^{-\alpha} \quad (12)$$

- $Y \sim \text{Exponential}(\beta) = Y \sim \text{Gamma}(1, \beta)$: ^{II}

$$f_Y(y) = \beta^{-1} \exp[-y\beta^{-1}] \text{ and } F_Y(y) = \exp(-y/\beta); \quad y \in (0, \infty) \quad (13)$$

$$\implies \mu = \beta \text{ and } \sigma^2 = \beta^2 \quad (14)$$

$$M_Y(t) = (1 - \beta t)^{-1} \quad (15)$$

- $Y \sim \chi^2(\nu) = Y \sim \text{Gamma}(\nu/2, 2)$:

$$f_Y(y) = \frac{y^{\nu/2-1} \exp[-y/2]}{\Gamma(\nu/2)2^{\nu/2}}; \quad y^2 > 0 \implies \mu = \nu \text{ and } \sigma^2 = 2\nu \quad (16)$$

$$M_Y(t) = (1 - 2t)^{-\nu/2} \quad (17)$$

- $Y \sim \text{Beta}(\alpha, \beta)$:

$$f_Y(y) = \frac{\Gamma(\alpha+\beta)}{\Gamma(\alpha)\Gamma(\beta)} y^{\alpha-1} (1-y)^{\beta-1}; \quad y \in (0, 1) \implies \mu = \frac{\alpha}{\alpha+\beta} \text{ and } \sigma^2 = \frac{\alpha\beta}{(\alpha+\beta)^2(\alpha+\beta+1)} \quad (18)$$

7.1 Special Cases

- $Y \sim \chi^2(1) = [Y \sim \text{Normal}(0, 1)]^2$
- $Y \sim \text{Beta}(1, 1) = Y \sim \text{Uniform}(0, 1)$

7.2 Sum of Independent Distributions (Y_1, Y_2)

- $Y_1 + Y_2 \sim \text{Normal}(\mu_1 + \mu_2, \sigma_1^2 + \sigma_2^2)$
- $Y_1 + Y_2 \sim \text{Poisson}(\mu_1 + \mu_2)$
- $Y_1 + Y_2 \sim \text{Binomial}(n_1 + n_2, p)$

7.3 Gamma Function

$$\Gamma(\alpha) = \int_0^\infty y^{\alpha-1} e^{-y} dy \text{ and } \Gamma(k \in \mathbb{N}) = (k-1)!$$

7.4 Standardizing the Normal Distribution

If $Y \sim \text{Normal}(\mu, \sigma^2) \implies$

$$Z = \frac{Y - \mu}{\sigma} \sim \text{Normal}(0, 1)$$

^{II}Memoryless property: $P(Y > y + \alpha \mid Y > \alpha) = P(Y > y)$