

HW11

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1. $\{Y_1, Y_2, \dots, Y_6\}$ i.i.d. $\sim \text{Exponential}(\beta)$

$$P(\min\{Y_1, Y_2, \dots, Y_6\} \geq a) = P(Y_1 \geq a) P(Y_2 \geq a) \dots P(Y_6 \geq a) = F_{Y_1}(a)^6 =$$

$$\left(\int_a^\infty \beta^{-1} \exp[-\beta^{-1}y] dy \right)^6 = \exp(-6a/\beta)$$

2. $\{Y_1, Y_2, \dots, Y_n\}$ i.i.d. with pdf $f_Y(y) = e^\theta e^{-y}, y > \theta$

$$f_{Y_{(1)}} = n[1 - F_Y(y)]^{n-1} \cdot f_Y(y) =$$

$$n \left(1 - e^\theta \int_\theta^\infty e^{-y} dy \right) e^\theta e^{-y} = n (1 - e^\theta [-e^{-y}]_\theta^\infty) e^\theta e^{-y} = n(1 - e^{2\theta}) e^\theta e^{-y}$$

3. $\{Y_1, \dots, Y_n\}$ i.i.d. $\sim \text{Beta}(2, 2)$

$$(a) F_{Y_{(n)}}(y) = F_Y(y)^n =$$

$$\left(\frac{\Gamma(2+2)}{\Gamma(2)\Gamma(2)} \int y^{2-1}(1-y)^{2-1} dy \right)^n = \left(\frac{\Gamma(4)}{\Gamma(2)^2} \int y(1-y) dy \right)^n = (3y^2 - 2y^3)^n = y^{2n} \sum_{k=0}^n \binom{n}{k} (3)^k (-2y)^{n-k}$$

$$(b) f_{Y_{(n)}}(y) = \frac{d}{dy} F_{Y_{(n)}}(y) = 6n(y - y^2)(3y^2 - 2y^3)^{n-1}$$

4. $Y = \{Y_1, Y_2, \dots, Y_5\}$ i.i.d. $\sim \text{Uniform}(0, 1)$

$$P(1/4 < Y_{\phi_{0.5}} < 3/4) = P(1/4 < Y_3 < \frac{3}{4}) = F_{Y(3)}(3/4) - F_{Y(3)}(1/4) =$$

$$\begin{aligned} & \frac{5!}{(3-1)!(5-3)!} \int_{1/4}^{3/4} F_Y(y)^{3-1} [1 - F_Y(y)]^{5-3} f_Y(y) dy = 30 \int_{1/4}^{3/4} F_Y(y)^2 [1 - F_Y(y)]^2 f_Y(y) dy \\ & = 30 \int_{1/4}^{3/4} y^2(1-y)^2 \cdot 1 dy = 30 \int_{1/4}^{3/4} y^2 - 2y^3 + y^4 dy = 30 \left[\frac{1}{3}y^3 - \frac{2}{4}y^4 + \frac{1}{5}y^5 \right]_{1/4}^{3/4} = \frac{203}{256} \end{aligned}$$

5.

Theorem. If $\{Y_1, \dots, Y_n\} \sim \text{Uniform}(0, \theta = 1) \implies Y_{(k)} \sim \text{Beta}(\alpha, \beta)$

Proof. We first find the pdf of the k^{th} order statistic. We know that

$$f_Y(y) = 1 \text{ and } F_Y(y) = y$$

then,

$$f_{Y_{(k)}}(y) = \frac{n!}{(k-1)!(n-k)!} [F_Y(y)]^{k-1} [1 - F_Y(y)]^{n-k} = \frac{n!}{(k-1)!(n-k)!} y^{k-1} (1-y)^{n-k}$$

A random variable $Y \sim \text{Beta}(\alpha, \beta)$ iff its pdf is

$$f_Y(y) = \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)} y^{\alpha-1} (1-y)^{\beta-1}$$

Then, we must have that $\alpha = k$ and $\beta = n - k + 1 \implies Y_{(k)} \sim \text{Beta}(k, n - k + 1)$ □

All decimals are rounded to 4 places.

6. $G_k \sim \text{Normal}(14, 2^2)$

(a) Let $G = \{G_1, \dots, G_{100}\}$. We want to know

$$P(\bar{G} > 14.5) = P\left(Z > \frac{14.5 - 14}{2/\sqrt{100}}\right) = .0062$$

(b) $P(g_1 < \bar{G} < g_2) = 0.95$

We have that

$$P(g_1 < \bar{G} < g_2) = P\left(\frac{g_1 - 14}{2/\sqrt{100}} < Z < \frac{g_2 - 14}{2/\sqrt{100}}\right) = 0.95$$

An interval corresponding to 0.95 means that at either side, we have 2.5% is excluded. Then,

$$P(Z > z_2) = 0.025 \implies z_2 = 1.96 \implies P(-1.96 < Z < 1.96) = 0.95$$

So, $\frac{g_1 - 14}{2/\sqrt{100}} = -1.96 \implies g_1 = -1.96 \cdot (2/\sqrt{100}) + 14 = 13.608$, and

$$\frac{g_2 - 14}{2/\sqrt{100}} = 1.96 \implies g_2 = 1.96 \cdot (2/\sqrt{100}) + 14 = 14.392.$$

7. $n = 100, \mu = \mu, \sigma = 2.5 \implies H = \{H_k : k \in [1, 100]\}$ i.i.d. $\sim \text{Normal}(\mu, 2.5^2)$

$$P(|\bar{H} - \mu| < 0.5) =$$

$$P\left(\frac{-0.5}{\sigma/\sqrt{n}} < \frac{\bar{H} - \mu}{\sigma/\sqrt{n}} < \frac{0.5}{\sigma/\sqrt{n}}\right) = P(-2 < Z < 2) = 2(P(Z > 0) - P(Z > 2)) = 0.9544$$

8. $C = \{C_1, \dots, C_{36}\}$ i.i.d. $\sim \text{Normal}(19400, 6000^2)$

$$P(\bar{C} > 20,000) =$$

$$P\left(Z > \frac{20000 - 19400}{6000/\sqrt{36}}\right) = 0.2743$$

9. $C = \{C_1, \dots, C_{2025}\} \sim \text{Normal}(3125, 540^2)$

$$P(c_1 < \bar{C} < c_2) = 0.95 =$$

$$P\left(\frac{c_1 - 3125}{540/\sqrt{2025}} < Z < \frac{c_2 - 3125}{540/\sqrt{2025}}\right) = 0.95 \implies \frac{c_1 - 3125}{540/\sqrt{2025}} = -1.96 \text{ and } \frac{c_2 - 3125}{540/\sqrt{2025}} = 1.96$$

Then, $c_1 = 3148.52, c_2 = 3101.48$.

10. $P(I) = 1/7; P = \sum_{k=1}^{612} P_k \sim \text{Binomial}(612, P(I))$

Let $np = 612 P(I)$, then $P(90 < P < 150) =$

$$P\left(\frac{90 - np}{\sqrt{np(1 - P(I))}} < Z < \frac{150 - np}{\sqrt{np(1 - P(I))}}\right) = P\left(0.2970 - \frac{1}{2} < Z < 7.2281 + \frac{1}{2}\right) = 0.5804$$

11. $n = 160, p = 0.05$

(a) $Y \sim \text{Binomial}(160, 0.05) \implies P(Y \leq 5) = \sum_{y=0}^5 \binom{160}{y} p^y (1-p)^{160-y} = 0.184224295731$

(b) $Y \sim \text{Poisson}(160 \cdot 0.05) \implies P(Y \leq 5) = \sum_{y=0}^5 \frac{\lambda^y e^{-\lambda}}{y!} = 0.19123606208$

(c) $P(Y \leq 5) = P\left(Z \leq \frac{5 - np}{\sqrt{np(1-p)}} + \frac{1}{2}\right) = 0.2810$

$$e = \begin{cases} d_{1h} := D_1(h) + K_1 h + K_2 h^2 \\ d_{2h} := D_1(2h) + 2K_1 h + 4K_2 h^2 \\ d_{3h} := D_1(3h) + 3K_1 h + 9K_2 h^2 \end{cases} \implies e = \alpha d_{1h} + \beta d_{2h} + \gamma d_{3h} \implies \begin{cases} \alpha + \beta + \gamma = 1 \\ \alpha + 2\beta + 3\gamma = 0 \\ \alpha + 4\beta + 9\gamma = 0 \end{cases}$$

$$\therefore \alpha = 3, \beta = -3, \gamma = 1$$