CS 760: Machine Learning

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# Homework 2: Linear Regression

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# **DO NOT POLLUTE!** AVOID PRINTING, OR PRINT 2-SIDED MULTIPAGE.

### Problem 2.1

To minimize the MSE of  $\theta$ , we computer our estimator  $\hat{\theta}$ :

$$\hat{\theta} = \underset{\theta}{\operatorname{arg\,min}} \|\mathbf{y} - \mathbf{X}\boldsymbol{\theta}\|_{\mathbf{2}}^{\mathbf{2}}$$

Expanding that out, we find:

$$||\mathbf{y} - \mathbf{X}\boldsymbol{\theta}||_{\mathbf{2}}^{2} = (\mathbf{y} - \mathbf{X}\boldsymbol{\theta})^{\mathsf{T}}(\mathbf{y} - \mathbf{X}\boldsymbol{\theta})$$
$$= \mathbf{y}^{\mathsf{T}}\mathbf{y} - \mathbf{y}^{\mathsf{T}}\mathbf{X}\boldsymbol{\theta} - \boldsymbol{\theta}^{\mathsf{T}}\mathbf{X}^{\mathsf{T}}\mathbf{y} + \boldsymbol{\theta}^{\mathsf{T}}\mathbf{X}^{\mathsf{T}}\mathbf{X}\boldsymbol{\theta}$$
$$= \mathbf{y}^{\mathsf{T}}\mathbf{y} - 2\boldsymbol{\theta}^{\mathsf{T}}\mathbf{X}^{\mathsf{T}}\mathbf{y} + \boldsymbol{\theta}^{\mathsf{T}}\mathbf{X}^{\mathsf{T}}\mathbf{X}\boldsymbol{\theta}$$

From there, we take the derivative w.r.t.  $\theta$ :

$$\frac{d}{d\theta} = 2\mathbf{X}^\mathsf{T}\mathbf{X}\theta - 2\mathbf{X}^\mathsf{T}\mathsf{y}$$

After setting this derivative to zero and solving for  $\theta$ , we find that our solution is:

$$\hat{\theta} = (\mathbf{X}^\mathsf{T}\mathbf{X})^{-1}\mathbf{X}^\mathsf{T}\mathsf{y}$$

## Problem 2.2

Since know that  $\epsilon$  is distributed normally, we can use the general multivariate normal distribution formula to derive our expression for the MLE of  $\theta^*$ :

$$\mathbb{P}(\mathbf{y},\mathbf{X}|\boldsymbol{\theta},\boldsymbol{\Sigma}^*) = \frac{1}{2\pi^{\frac{n}{2}}|\boldsymbol{\Sigma}^*|^{\frac{1}{2}}}\mathbf{e}^{-\frac{1}{2}(\mathbf{y}-\mathbf{X}\boldsymbol{\theta})^{\mathbf{T}}\boldsymbol{\Sigma}^{*-1}(\mathbf{y}-\mathbf{X}\boldsymbol{\theta})}$$

To find the MLE, we first take the natural log, then we take the derivative and set it to zero. After that, we solve for  $\hat{\theta}$ :

$$\hat{\theta} = (\mathbf{X}^\mathsf{T} \mathbf{\Sigma}^{*-1} \mathbf{X})^{-1} \mathbf{X}^\mathsf{T} \mathbf{\Sigma}^{*-1} \mathsf{y}$$

This is very similar to the solution of Problem 2.1 above, with the exception that it's now being weighted by the inverse covariance matrix  $\Sigma^{*-1}$ 

#### Problem 2.3

The distribution of the MLE of  $\theta^*$  is:

$$\hat{\theta} \sim \mathcal{N}(\theta^*, (X^T \Sigma^{*-1} X)^{-1})$$

#### Problem 2.4

With x being a new sample and  $\hat{\theta}$  being the MLE of  $\theta^*$ , we can say that based on what we derived above, the MLE of  $\hat{y}$  is exactly  $x^T\hat{\theta}$ 

#### Problem 2.5

Borrowing from 2.3 and knowing that  $\hat{y} = x^T \hat{\theta}$ , we say that:

$$\hat{y} \sim \mathcal{N}(x^T \theta^*, x^T (X^T \Sigma^{*-1} X)^{-1} x)$$

### Problem 2.6

We'll take a similiar approach as in Problem 2.2, so first we take the natural log, take the derivative w.r.t.  $\Sigma^*$ , set it to zero, and then solve for  $\Sigma^*$ . Doing so, we get:

$$\Sigma^* = (\mathsf{y} - \mathsf{X}\theta)(\mathsf{y} - \mathsf{X}\theta)^\mathsf{T}$$

### Problem 2.7

- (a) Here we run into an issue, because both  $\Sigma^*$  and  $\theta^*$  are defined in terms of each other. To get around this issue, we use Expectation Maximization to iteratively approach the solution for either terms. Doing this yields an MLE for  $\Sigma^* = 1025.0$  and  $\theta^* = 155$
- (b) Following the same logic as Problem 2.4, the MLE of  $\hat{y}$  is  $x^T\hat{\theta}$

(c) To derive a confidence interval of 95%, we first need to find tau:

$$\tau = \Phi_{\mathcal{N}}^{-1}(\frac{\alpha}{2}|0,\sigma^2\mathbf{x}^\mathsf{T}(\mathbf{X}^\mathsf{T}\mathbf{X})^{-1}\mathbf{x})$$

such that

$$\mathbb{P}(|\hat{y} - \mathsf{y}^*| \le \tau) = 1 - \alpha$$

Here we set  $\alpha=0.05$  to give us a 95% confidence interval  $(\hat{y}-\tau,\hat{y}+\tau)$ 

- (d) No
- (e) Yes