### Resampling Methods for Uncertainty Big Data y Machine Learning para Economía Aplicada

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### Agenda

- 1 Review
  - FWL
- 2 Uncertainty: Motivation
  - Resampling methods
  - Parameter Assessment: the Bootstrap
    - Example: Elasticity of Demand for Gasoline
  - Model Assessment
    - AIC: Akaike Information Criterion
    - SIC/BIC: Schwarz/Bayesian Information Criterion
    - Cross-Validation
- 3 Review



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# Prediction and linear regression

- ▶ We have data  $\{y_i, X_i\}$
- ► Interest on predicting *y*

$$y = f(X) + u \tag{1}$$

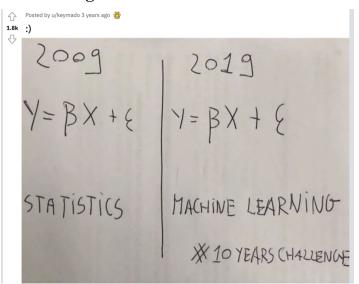
- ▶ When making a prediction we want to minimize the prediction errors
- ▶ A common loss function is the squared loss  $L(e) = e^2$

### Minimizing our losses

▶ The E[L(e)] of using an estimate:  $\hat{f}(x)$ , can be decomposed

$$E(y - \hat{y})^2 = \underbrace{Bias^2(\hat{f}(X)) + V(\hat{f}(X))}_{Reducible} + \underbrace{Var(u)}_{Irreducible}$$
(2)

### Prediction and linear regression



# Prediction and linear regression

► We proposed

$$f(X) = \beta_0 + \beta_1 X_1 + \dots + \beta_p X_p \tag{3}$$

• were estimating f(X) boils down to findig  $\beta$ 

### Linear Regression

• Choose the estimators  $\hat{\beta}$  such that we minimize the E[L(e)] (SSR)

$$\hat{\beta} = \underset{\tilde{\beta}}{\operatorname{argmin}} SSR(\tilde{\beta}) \tag{4}$$

- ightharpoonup Compute  $\beta$ 
  - ▶ QR: Householder transformation, Gram-Schmidt process (similar to FWL)
  - Gradient Descent
- ► Numerical Properties

### Numerical Properties

- ▶ Numerical properties have nothing to do with how the data was generated
- These properties hold for every data set, just because of the way that  $\hat{\beta}$  was calculated
- ► Helps in computing with big data

### Frisch-Waugh-Lovell (FWL) Theorem

- ▶ Lineal Model:  $Y = X\beta + u$
- ► Split it:  $Y = X_1\beta_1 + X_2\beta_2 + u$ 
  - $X = [X_1 X_2], X \text{ is } n \times k, X_1 n \times k_1, X_2 n \times k_2, k = k_1 + k_2$

### Theorem

1 The OLS estimates of  $\beta_2$  from these equations

$$y = X_1 \beta_1 + X_2 \beta_2 + u \tag{5}$$

$$M_{X_1}y = M_{X_1}X_2\beta_2 + residuals \tag{6}$$

are numerically identical

2 the OLS residuals from these regressions are also numerically identical



### Projection

**OLS Residuals:** 

replacing  $\hat{\beta}$ 

$$= y -$$

$$e = y - \hat{y}$$

$$=y-X\hat{\beta}$$

$$e = y - X(X'X)^{-1}X'y$$
  
=  $(I - X(X'X)^{-1}X')y$ 

- ▶ Projection matrix  $P_X = X(X'X)^{-1}X'$
- Annihilator (residual maker) matrix  $M_X = (I P_X)$



(7)

(8)

(9)

(10)

### Projection

- $P_X = X(X'X)^{-1}X'$
- $ightharpoonup M_X = (I P_X)$
- ▶ Both are symmetric
- ▶ Both are idempotent (A'A) = A
- $ightharpoonup P_X X = X$  hence projection matrix
- $ightharpoonup M_X X = 0$  hence annihilator matrix

We can write

$$SSR = e'e = u'M_Xu \tag{11}$$

So we can relate SSR to the true error term u



### **Applications**

- ▶ Why FWL is useful in the context of big volume of data?
- ► An computationally inexpensive way of
  - Removing nuisance parameters
    - ► E.g. the case of multiple fixed effects. The traditional way is either apply the within transformation with respect to the FE with more categories then add one dummy for each category for all the subsequent FE
  - ightharpoonup Computing certain diagnostic statistics: Influential Observations,  $R^2$ , LOOCV.

### Applications: Fixed Effects

► For example: Carneiro, Guimarães, & Portugal (2012) AEJ: Macroeconomics

$$\ln w_{ijft} = x_{it}\beta + \lambda_i + \theta_j + \gamma_f + \delta_t + u_{ijft}$$
(12)

- ▶ Data set 31.6 million observations, with 6.4 million individuals (i), 624 thousand firms (f), and 115 thousand occupations (j), 11 years (t).
- ▶ Storing the required indicator matrices would require 23.4 terabytes of memory
- From their paper

"In our application, we first make use of the Frisch-Waugh-Lovell theorem to remove the influence of the three high-dimensional fixed effects from each individual variable, and, in a second step, implement the final regression using the transformed variables. With a correction to the degrees of freedom, this approach yields the exact least squares solution for the coefficients and standard errors"

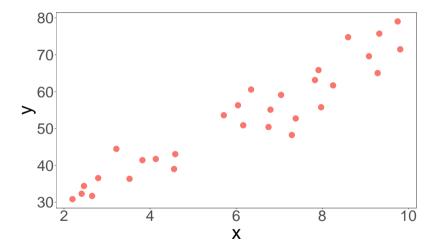
Note the following

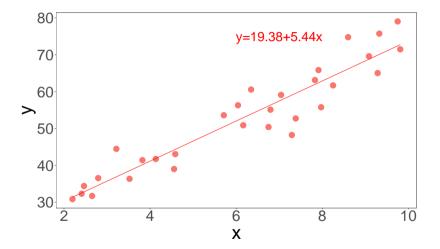
$$\hat{\beta} = (X'X)^{-1}X'y \tag{13}$$

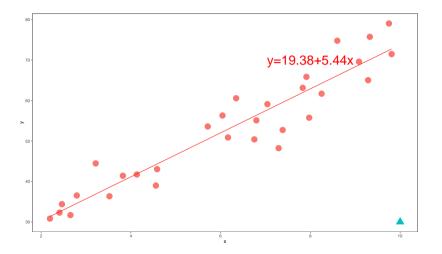
each element of the vector of parameter estimates  $\hat{\beta}$  is simply a weighted average of the elementes of the vector y

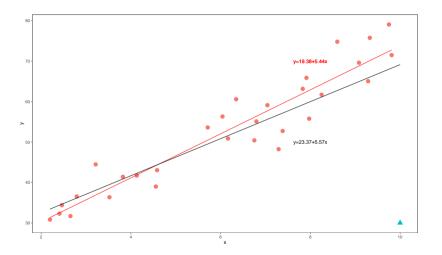
Let's call  $c_j$  the j-th row of the matrix  $(X'X)^{-1}X'$  then

$$\hat{\beta}_j = c_j y \tag{14}$$









Consider a dummy variable  $e_j$  which is an n-vector with element j equal to 1 and the rest is 0. Include it as a regressor

$$y = X\beta + \alpha e_j + u \tag{15}$$

using FWL we can do

$$M_{e_j}y = M_{e_j}X\beta + r \tag{16}$$

- $\triangleright$   $\beta$  and *residuals* from both regressions are identical
- Same estimates as those that would be obtained if we deleted observation j from the sample. We are going to denote this as  $\beta^{(j)}$

#### Note:

$$M_{e_i} = I - e_j (e_i' e_j)^{-1} e_i'$$

$$M_{e_i}y = y - e_j(e'_je_j)^{-1}e'_jy = y - y_je_j$$

 $ightharpoonup M_{e_i} X$  is X with the j row replaced by zeros



### Applications: Influential Observations Let's define a new matrix $Z = [X, e_i]$

$$y = X\beta + \alpha e_j + u$$
$$y = Z\theta + u$$

$$y = P_Z y + M_z y$$
$$= \mathbf{Y} \hat{\mathbf{g}}(j) + \hat{\mathbf{g}}(g)$$

$$= X\hat{\beta}^{(j)} + \hat{\alpha}e_j + M_Z y$$

Pre-multiply by 
$$P_X$$
 (remember  $M_Z P_X = 0$ )

$$P_X y = X \hat{\beta}^{(j)} + \hat{\alpha} P_X e_i$$

$$X\hat{eta} = X\hat{eta}^{(j)} + \hat{lpha}P_Xe_j \ X(\hat{eta} - eta^{(j)}) = \hat{lpha}P_Xe_i$$

$$X(\rho - \rho^{\vee}) = \alpha r \chi e_j$$

$$aP_Xe_j$$
 (22)



(17)

(18)

(19)

(20)

How to calculate  $\alpha$ ? FWL once again

$$M_X y = \hat{\alpha} M_X e_j + res \tag{24}$$

$$\hat{\alpha} = (e_j' M_X e_j)^{-1} e_j' M_X y \tag{25}$$

- $e'_j M_X y$  is the j element of  $M_X y$ , the vector of residuals from the regression including all observations
- $e'_j M_x e_j$  is just a scalar, the diagonal element of  $M_X$ Then

$$\hat{\alpha} = \frac{\hat{u}_j}{1 - h_i} \tag{26}$$

where  $h_i$  is the j diagonal element of  $P_X$ 

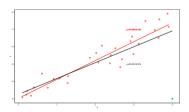


# Applications: Influential Observations Finally we get

$$(\hat{\beta}^{(j)} - \hat{\beta}) = -\frac{1}{1 - h_j} (X'X)^{-1} X_j' \hat{u}_j$$
(27)

Influence depends on two factors

- $\hat{u}_j$  large residual  $\rightarrow$  related to y coordinate
- $\hat{h}_i \rightarrow \text{related to x coordinate}$



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### Motivation

- ► The real world is messy.
- ▶ Recognizing this mess will differentiate a sophisticated and useful analysis from one that is hopelessly naive.
- ► This is especially true for highly complicated models, where it becomes tempting to confuse signal with noise and hence "overfit."
- ▶ The ability to deal with this mess and noise is the most important skill you need.

### What are resampling methods?

- ➤ Tools that involves repeatedly drawing samples from a training set and refitting a model of interest on each sample in order to obtain more information about the fitted model
  - ▶ Parameter Assessment: estimate standard errors
  - ► Model Assessment: finding the best model

### Introduction

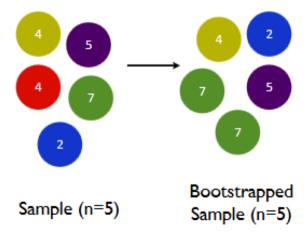
- ► Suppose we have  $y_1, y_2, ..., y_n$  iid  $Y \sim (\mu, \sigma^2)$  (both finite)
- ► We want to estimate

$$Var(\bar{Y})$$
 (28)

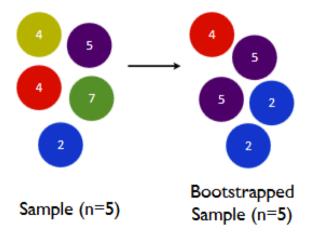
Introduction

- ► Alternative way (no formula!)
  - 1 From the *n* original data points  $y_1, y_2, \ldots, y_n$  take a sample with replacement of size *n*
  - 2 Calculate the sample average of this "pseudo-sample" (Bootstrap sample)
  - 3 Repeat this B times.
  - 4 Compute the variance of the B means

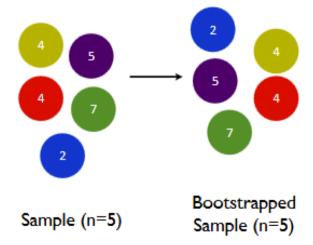
Sampling with replacement



Sampling with replacement



Sampling with replacement



Introduction

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Variance in Linear Regression

- ► Suppose we have  $\{y_i, X_i\}$   $i = \{1, ..., n\}$ iid
- ▶ We want to estimate  $Var(\hat{\beta})$

- ► Sometimes the analytical expression of the variance can be quite complicated.
- ► In these cases bootstrap can be useful
- ► The bootstrap provides a way to perform statistical inference by resampling from the sample.
- ▶ In German the expression *an den eigenen Haaren aus dem Sumpf zu ziehen* nicely captures the idea of the bootstrap "to pull yourself out of the swamp by your own hair."



Two key properties

- ► Two key properties of bootstrapping that make this seemingly crazy idea actually work.
  - 1 Each bootstrap sample must be of the same size (n) as the original sample
  - 2 Each bootstrap sample must be taken with replacement from the original sample.

- ► In general terms:
  - ightharpoonup Sample  $\{y_i, X_i\}$  i = 1, ..., n iid
  - $\triangleright$   $\theta$  is the magnitude of interest
  - 1 Sample of size *n* with replacement (*bootstrap sample*)
  - 2 Compute  $\hat{\theta}_j$   $j = 1, \dots, B$
  - 3 Repeat B times
  - 4 Calculate the magnitude of interest

#### Example: Elasticity of Demand for Gasoline



photo from https://www.dailydot.com/parsec/batman-1966-labels-tumblr-twitter-vine/

#### The Bootstrap

Why it works?

- ► The key is that the distribution of any estimator or statistic is determined by the distribution of the data.
- ▶ While the latter is unknown it can be estimated by the empirical distribution of the data.

#### Prediction

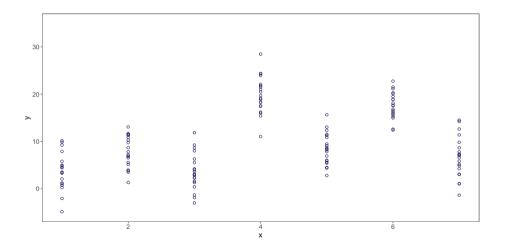
- ightharpoonup Objective predict *y* given *X*.
- ightharpoonup link between y and X:

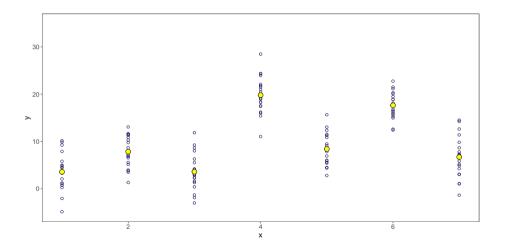
$$y = f(X) + u (29)$$

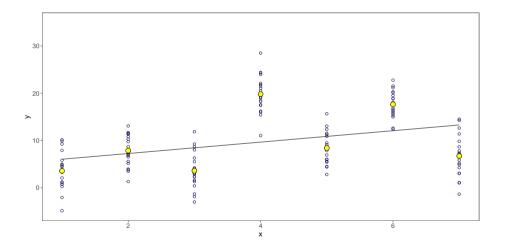
ightharpoonup u rv E(u) = 0 and  $V(u) = \sigma^2$ 

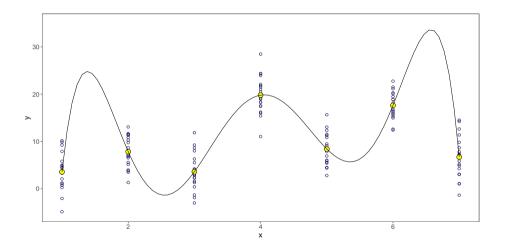
Train and Test Sets. In-Sample and Out-of-Sample Prediction.

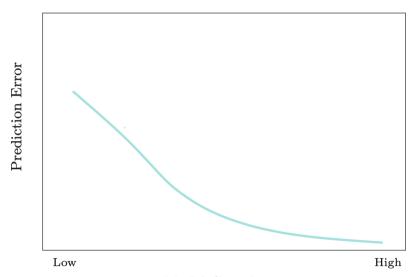
► How do we select a model with the lowest prediction error?



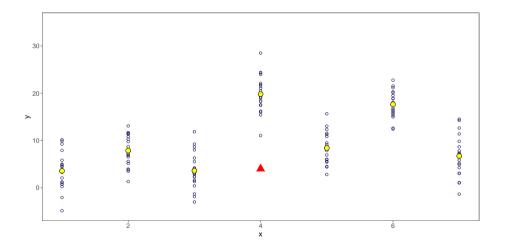


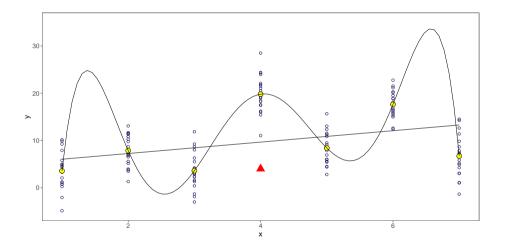


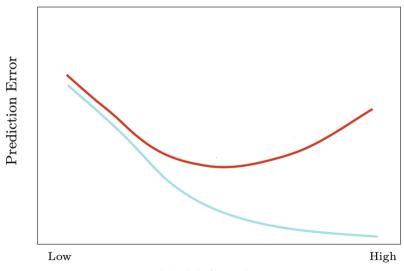




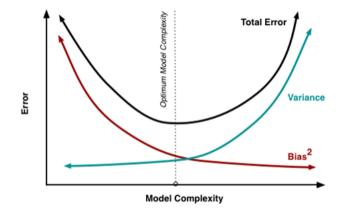
► ML we care about out of sample prediction







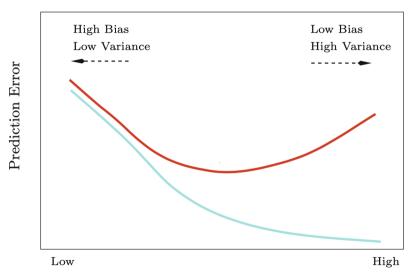
#### Out-of-Sample Error and the Bias Variance Trade-Off



$$E(y - \hat{y})^{2} = \underbrace{Bias^{2}(\hat{f}(X)) + V(\hat{f}(X))}_{Reducible} + \underbrace{Var(u)}_{Irreducible}$$

(30)

#### Out-of-Sample Error and the Bias Variance Trade-Off



- ► ML we care about out of sample prediction
- ► How we estimate the out of sample error?

## In-Sample and Out-of-Sample Errors.

- ► Two important concepts
  - ► *Training error*:

$$Err_{\mathcal{T}rain} = MSE[(y, \hat{y}) | \mathcal{T}rain]$$
 (31)

► *Test Error*:

$$Err_{Test} = MSE[(y, \hat{y}) | Test]$$
 (32)

- ightharpoonup How dowe estimate the  $Err_{Test}$
- ► Two ways
  - ightharpoonup Ex post penalization: AIC, BIC, Adj  $R^2$

- Akaike (1969)
- ► Minimize

$$AIC(j) = log\left(\frac{1}{n}\sum_{i=1}^{n}(y_i - \hat{y})^2\right) - p_j$$
 (33)

## Test Error SIC/BIC

- Schwarz (1978)
- Minimize

$$SIC(j) = log\left(\frac{1}{n}\sum_{i=1}^{n}(y_i - \hat{y})^2\right) - \frac{1}{2}p_j log(n)$$
 (34)

AIC vs BIC

$$AIC(j) = log\left(\frac{1}{n}\sum_{i=1}^{n}(y_i - \hat{y})^2\right) - p_j$$
 (35)

$$SIC(j) = log\left(\frac{1}{n}\sum_{i=1}^{n}(y_i - \hat{y})^2\right) - p_j\frac{1}{2}log(n)$$
 (36)

- ▶ How dowe estimate the  $Err_{Test}$
- ► Two ways
  - ightharpoonup Ex post penalization: AIC, BIC, Adj  $R^2$
  - ► Resampling methods

#### Cross-Validation



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#### Review

- ► Review + FWL
- ► Resampling Methods:
  - ▶ Parameter Assessment: estimate standard errors
  - ► Model Assessment: finding the best model