

ECON 21020, PSet 6: Solutions

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([] out of 38p) **PART I: Heteroschedasticity and the Linear Probability Model**

([] out of 6p) **Q1: Proof of Statements (a)-(g) in Claim 1**

$$\text{a) NEED OTHER I THINK} \quad (1)$$

$$\text{b) } E[D_i|x] = P(D_i = 1|x) = P(\beta_0 + \beta_1 x \geq u_i|x) = \beta_0 + \beta_1 x \quad (2)$$

$$\text{Since } u_i \sim U[0, 1] \text{ and } \beta_0 + \beta_1 x \in [0, 1] \quad (3)$$

$$\text{c) } \sigma^2(x) = E[D_i|x] - E[D_i|x]^2 \quad (4)$$

$$= (\beta_0 + \beta_1 x) - (\beta_0 + \beta_1 x)^2 \quad (5)$$

$$= \beta_0(1 - \beta_0) + \beta_1(1 - 2\beta_0)x - \beta_1^2 x^2 \quad (6)$$

$$(7)$$

$$\text{d) let } \epsilon_i = D_i - E[D_i|x] \quad (8)$$

$$\text{then } \epsilon_i = D_i - \beta_0 - \beta_1 x_i \quad (9)$$

$$\text{and } D_i = \beta_0 + \beta_1 x_i + \epsilon_i \quad (10)$$

$$(11)$$

$$E[\epsilon_i|x] = E[D_i - E[D_i|x]|x] \quad (12)$$

$$= E[D_i|x] - E[D_i|x] = 0 \quad (13)$$

$$(14)$$

$$Var[\epsilon_i] = E[\epsilon_i^2] - E[\epsilon_i]^2 \quad (15)$$

$$= E[(D_i - E[D_i|x])^2|x] \quad (16)$$

$$\text{By def. this is } = Var[D_i|x] \quad (17)$$

$$(18)$$

$$\text{e) let } E[\hat{\beta}_1] = E\left[\frac{\sum_{i=1}^n (x_i - \bar{x})(D_i - \bar{D})}{\sum_{i=1}^n (x_i - \bar{x})^2}\right] \quad (19)$$

$$\text{Focusing on numerator} = \sum_{i=1}^n (x_i - \bar{x})E[D_i - \bar{D}] \quad (20)$$

$$= \sum_{i=1}^n (x_i - \bar{x})(\beta_0 + \beta_1 x_i - \beta_0 - \beta_1 \bar{x}) \quad (21)$$

$$= \sum_{i=1}^n (x_i - \bar{x})\beta_1(x_i - \bar{x}) \quad (22)$$

$$= \beta_1 \sum_{i=1}^n (x_i - \bar{x})^2 \quad (23)$$

$$\text{Brining back the denominator we get} = \beta_1 \frac{\sum_{i=1}^n (x_i - \bar{x})^2}{\sum_{i=1}^n (x_i - \bar{x})^2} \quad (24)$$

$$= \beta_1 \quad (25)$$

$$(26)$$

f) Since LR.5 fails because $Var[\epsilon_i|x] = \sigma^2(x)$ the OLS estimators are not BLUE, specifically because they don't have the minimum variance (27)

([] out of 5p) **Q2: Deep dive into the Asymptotic Properties of FWLS Estimator**

([] out of 2p) **Q2.a**

$$E[D_i * |x] = w_i E[D_i|x] = \frac{1}{\sqrt{\sigma^2(x)(\beta_0 + \beta_1 x_1 + \beta_2 x_2)}} \text{mm} \quad (28)$$

([] out of 3p) **Q2.b**

The only difference between $\hat{\beta}^{FWLS}$ and $\hat{\beta}^{WLS}$ is the use of the estimator $\hat{\sigma}^2$ instead of the true value of σ^2 . Since the estimator $\hat{\sigma}^2$ converges to the true value of σ^2 as the sample size converges to infinity, therefore it makes sense that $\hat{\beta}^{FWLS}$ would converge to $\hat{\beta}^{WLS}$ and therefore as the sample size converges to the mean of $\hat{\beta}^{FWLS}$ converges to the mean of $\hat{\beta}^{WLS}$ and variance of $\hat{\beta}^{FWLS}$ converges to variance of $\hat{\beta}^{WLS}$.

([] out of 8p) **Q3: Relationship b/w Structural and Reduced-form Parameters**

Reduced-form Parameter	Sign	Reason for Sign
λ_0	+	Reason for the sign of Parameter 1
λ_1	-	Reason for the sign of Parameter 2
λ_2	+	Reason for the sign of Parameter 3
λ_3	-	Reason for the sign of Parameter 4
λ_4	+	Reason for the sign of Parameter 5
λ_5	-	Reason for the sign of Parameter 6
λ_6	+	Reason for the sign of Parameter 7
λ_7	+	Reason for the sign of Parameter 7

Table 1: The Reduced-form Parameters of the Labor Force Participation Decision of a Married Woman

([] out of 4p) Q4: A Linear-in-parameter Model for the LFP decision of a married female

([] out of 2p) Q4.a

$$P(D_i = 1|z) = P(\lambda_0 + \lambda_1 e + \lambda_2 x + \lambda_3 x^2 + \lambda_4 k^{\geq 6} + \lambda_5 k^{<6} + \lambda_6 a + \lambda_7 m > u) = \lambda_0 + \lambda_1 e + \lambda_2 x + \lambda_3 x^2 + \lambda_4 k^{\geq 6} + \lambda_5 k^{<6} + \lambda_6 a + \lambda_7 m \quad (29)$$

([] out of 2p) Q4.b

Since both β_0 and $\alpha_{1,0}$ influence labor force participation but their effects cannot be measured independently of each other, are they not separately identified. This is because β_0 represents the baseline wage for women regardless of the eVars and $\alpha_{1,0}$ represents the baseline enjoyment level of leisure regardless of the eVars. Since there are no data points to find the change in participation from changes in β_0 or $\alpha_{1,0}$, they are not separately identified.

([] out of 2p) Q5: Estimate by OLS a LPM for a Married Female's LFP

Script and Output

Commentary

([] out of 7p) Q6: Estimate by FWLS a LPM for a Married Female's LFP

Script and Output

([] out of 2p) Q7: Verify FWLS Estimate

Script and Output

([] out of 2p) Q8: Compare OLS and FWLS Estimates

Script and Output

Commentary

([] out of 2p) Q9: Complete Statements

Commentary

([] out of 21p) PART II: Three Distributions derived from the Normal Distribution

([] out of 5p) Q10: The Normal Distribution

a) $Y \sim N(a + b\mu, b^2\sigma^2)$ (30)

b) $Y \sim N(\frac{\mu}{c}, \frac{\sigma^2}{c^2})$ (31)

c) $Y \sim N(a + b\mu_1 + c\mu_2, b^2\sigma^2 + c^2\sigma^2 + 2bc\sigma_{1,2})$ (32)

d) $Y \sim N(\mu, \frac{\sigma^2}{n})$ (33)

e) $Y \sim N(0, 1)$ (34)

(35)

([] out of 4p) Q11: Distributions Derived from the Normal Distribution

a) $Y \sim \sigma_2^2\chi_1^2$ (36)

b) $Y \sim \sigma_1^2\chi_1^2 + \sigma_2^2\chi_1^2$ (37)

c) $Y \sim t_q$ d) $Y \sim F_{(1,1)}$ (38)

([] out of 4p) Q12: The Standardized Sample Average for a Normal RS Has a t of Student Distribution

We know $\frac{\bar{X} - \mu}{\sqrt{\frac{S^2}{n}}} \sim N(0, 1)$ and $\frac{S^2}{\sigma^2}$ has chi-squared distribution with $n - 1$ degrees of freedom and since \bar{X} and S^2 are independent, $\frac{\bar{X} - \mu}{\sqrt{\frac{S^2}{n}}} \sim t_{n-1}$ (39)

([] out of 1p) Q13: Support of the Chi-squared Distribution

Script and output

Commentary

([] out of 1p) Q14: Practice with lower.tail argument

Script and output

Commentary

([] out of 4p) Q15: Student's t Distribution Converges to Standard Normal

Script and output

Commentary

([] out of 2p) Q16 Relationship between F and Student t Distributions

Script and output

Commentary

([] out of 41p) PART III: Hypothesis Testing in the MLRM

([] out of 0p) Q17: Load data and estimate log wage model

Script and output

([] out of 12p) Q18: $H_0 : \beta_{\text{exper}} = 0$ versus $H_0 : \beta_{\text{exper}} \neq 0$

([] out of 1p) Q18.a

We test that the percentage change in hourly wage caused by an extra year of labor market experience is zero.

([] out of 3p) Q18.b

([] out of 1p) Q18.c

Script and output

([] out of 1p) Q18.d

Script and output

([] out of 2p) Q18.e Critical Value Approach

Script and Output

([] out of 2p) Q18.f P-value Approach

Script and Output

([] out of 2p) Q19: Definition of critical value and p-value

([] out of 1p) Q19.a

([] out of 1p) Q19.b

([] out of 2p) Q20: Test $H_0 : \beta_{\text{exper}} = 0.007$ versus $H_0 : \beta_{\text{exper}} \neq 0.007$ with PV Approach

([] out of 1p) Q20.a

Script and output

([] out of 1p) Q20.b

Script and output

([] out of 6p) Q21: Use `car::linearHypothesis()` to test simple linear hypotheses

([] out of 4p) Q21.a Test $H_0 : \beta_{\text{exper}} = 0$ versus $H_0 : \beta_{\text{exper}} \neq 0$

([] out of 2p) Q21.a.i Use Syntax A of Function `car::linearHypothesis()`

Script and output

Commentary

([] out of 2p) Q21.a.ii Syntax B of function `car::linearHypothesis()`

Script and output

Commentary

([] out of 2p) Q21.b Test $H_0 : \beta_{\text{exper}} = 0.007$ versus $H_0 : \beta_{\text{exper}} \neq 0.007$

Script and output

([] out of 8p) Q22 Test One-sided Hypothesis $H_0 : \beta_{\text{exper}} \leq 0$ versus $H_1 : \beta_{\text{exper}} > 0$

([] out of 1p) Q22.a

([] out of 2p) Q22.b

([] out of 3p) Q22.c

Script and output

Commentary

([] out of 2p) Q22.d

Script and output

Commentary

([] out of 11p) Q23 Test $H_0 : \beta_{\text{exper}} = \beta_{\text{tenure}}$ versus $H_1 : \beta_{\text{exper}} \neq \beta_{\text{tenure}}$

([] out of 1p) Q23.a Interpretation of the Hypothesis

([] out of 2p) Q23.b Run the test using `car::linearHypothesis()`

Script and output

Commentary

([] out of 4p) Q23.c Run the test using the Reparametrization Approach

Script and output

Commentary

([] out of 4p) Q23.d Run the test using the Model Reformulation Approach

Script and output

Commentary
