# A multi-scale symmetry analysis of daily financial time series using uninterrupted trends returns

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- Dr. Hector Conoronel Brizio, Universidad Veracruzana.

#### Introduction

### Publicly available information

Prices, volume, dividends, beta, etc.

#### Intel Corporation (INTC)

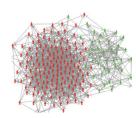
Best Bid/Ask	\$ 29.96 / \$ 29.97
1 Year Target	38
Today's High /Low	\$ 30.50 / \$ 29.45
Share Volume	52,122,067
50 Day Avg. Daily Volume	22,872,515
Previous Close	\$ 32.74
52 Week High/Low	\$ 37.03 / \$ 24.87
Market cap	\$ 141,428,430,000
P/E Ratio	12.81
Forward P/E(1y)	13.76
Earnings Per Share (EPS)	\$ 2.34
Annualized dividend	\$ 0.96
Ex Dividend Date	Nov. 4, 2015
Dividend Payment Date	Dec. 1, 2015
Current Yield	2.93 %
Beta	1.11
NA SDAQ Official Open Price	\$ 29.63
Date of Open Price	Jan. 15, 2016
NA SDAQ Official Close Price	\$ 32.74
Date of Close Price	Jan. 14, 2016
Community Sentiment	<b>▼</b> Bullish



### What is relevant to price

- Interaction of many agents that perform buy/sell operations.
- Every agent has access to the same public information to make an informed decision on every operation. They are rational.
- The direct consequence is that current price already contains all the knowledge of future forecasts.

$$E[P_{t+1}|P_0, P_1, \dots, P_t] = P_t.$$



## Efficient market hypothesis (EMH)

- Bachelier 1900 proposes that since in a voluntary stock exchange between two agents, each of them believes them to be the winning party. If the game is fair (Martingale) it must behave like a random walk. His work was not well received.
- Samuelson 1965 proposes the efficient market hypothesis. If investors were able to incorporate future events in the current price, they would have done so, incorporating their knowledge in the current price.
- Fama 1995 finds empirical evidence that markets behave like a random walk.
- The EMH requires agents to be rational.

### An historical example

- Challenger space shuttle. Four contractors involved: Lockheed, Martin Marietta, Morton Thiokol y Rockwell.
- After 5 months of investigation it was found that the misbehaving component came from Morton Thiokol.
- Markets declared their verdict right away.

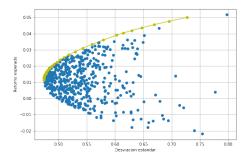


Time from 11:40 a.m. to Close



### If I can't win there is no point

 Markowitz 1952 optimal portfolio theory tell us that we can construct a portfolio such that the relation between expected return and risk is optimal.



- Useful for long term investment.
- EMH implies the reliability of predictive markets.



### There are winners however...

- The existence of agents that can reliably generate profits like Warren Buffet, Peter Lynch and George Soros, implies the possibility that some agents can sometimes predict better than others future outcomes.
- Humans have cognitive biases that systematically makes us averse to lose and have high hopes of winning. We are not always rational.

### Motivation

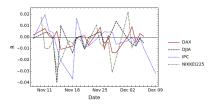
How can we assess the efficiency of a market?

#### **Background**

### **Definitions**

We define the logarithmic returns as

$$R(t, \Delta t) = \log P_{t+\Delta t} - \log P_t.$$



 We can express the efficient market hypothesis in terms of the returns

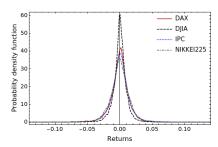
$$E[R_{t+1}|R_0, R_1, \ldots, R_t] = 0.$$

 From this definition of the returns emerge the stylized facts, that are statistical properties common to the returns of most markets.



## Stylized facts I

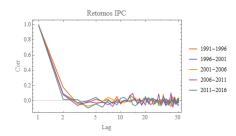
 Win / Loss assymetry. Symmetry is strongly related to the EMH.



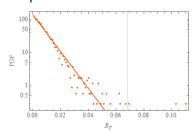
No auto-correlation on the returns



## Stylized facts II

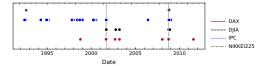


• Returns decay as a power law.



### Stylized facts III

 Some events extremely rare (according to the power law) still happen sometimes. We identify these extreme events as the Dragon Kings described by Sornette. Usually these extreme values signal a big transition in the market like an economic crisis.



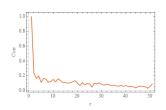
Volatility clustering

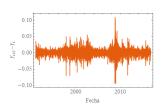
$$C_2(\tau) = \operatorname{Corr}(R(t+\tau,\Delta t)^2, R(t,\Delta t)^2).$$

That is, big changes tend to be followed by big changes of any sign, and small changes tend to be followed also by small changes of any sign.



## Stylized facts IV





# H.F. Coronel-Brizio, et al. (2007) Assessing symmetry of financial returns series. Physica A 383.

#### A test for symmetry

- Based on the  $T_n$  statistic developed by Einmahl and McKeague.
- The null hypothesis is that there is a symmetry point around c, that is  $H_0: F(c-x) = 1 F(x-c)$ , for every x > 0 in the sample  $X_1, \ldots, X_n$ . F is a cumulative distribution function.

# Einmahl, McKeague (2003) Empirical likelihood based hypothesis testing I

Many of the papers cited above consider the case of a known point of symmetry and use a Cramér-von Mises type test statistic. We also assume that the point of symmetry is known, so without loss of generality it is assumed to be zero. Let  $X_1, \ldots, X_n$  be i.i.d. with continuous distribution function F. The null hypothesis of symmetry about zero is

$$H_0: F(-x) = 1 - F(x-),$$
 for all  $x > 0$ .

The local likelihood ratio statistic is defined by

$$R(x) = \frac{\sup\{L(\tilde{F}) : \tilde{F}(-x) = 1 - \tilde{F}(x-)\}}{\sup\{L(\tilde{F})\}}, \qquad x > 0.$$

As in the Introduction, the unrestricted likelihood in the denominator is maximized by setting  $\tilde{F}=F_n$ , the empirical distribution function. The supremum in the numerator can be found by treating  $\tilde{F}$  as a function of  $0 \le p \le 1$ , where  $\tilde{F}$  puts mass p/2 on the interval  $(-\infty, -x]$ , mass p/2 on  $[x, \infty)$  and mass 1-p on (-x, x), with those masses divided equally among

## Einmahl, McKeague (2003) Empirical likelihood based hypothesis testing II

the observations in the respective intervals. That is, the masses on the individual observations in the respective intervals are given by

$$\frac{p/2}{n\hat{p}_1}, \frac{p/2}{n\hat{p}_2}, \frac{1-p}{n(1-\hat{p})},$$

where  $\hat{p} = \hat{p}_1 + \hat{p}_2$ ,  $\hat{p}_1 = F_n(-x)$  and  $\hat{p}_2 = 1 - F_n(x-)$ . The numerator of R(x) is therefore the maximal value of

$$\left(\frac{p/2}{n\hat{p}_1}\right)^{n\hat{p}_1} \left(\frac{p/2}{n\hat{p}_2}\right)^{n\hat{p}_2} \left(\frac{1-p}{n(1-\hat{p})}\right)^{n(1-\hat{p})},$$

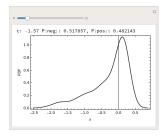
which is easily seen to be attained at  $p = \hat{p}$ . We thus obtain

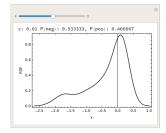
$$\log R(x) = n\hat{p}_1 \log \frac{p}{2\hat{p}_1} + n\hat{p}_2 \log \frac{p}{2\hat{p}_2}$$

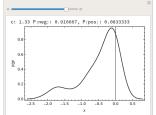
$$= nF_n(-x)\log \frac{F_n(-x) + 1 - F_n(x-)}{2F_n(-x)} + n(1 - F_n(x-))\log \frac{F_n(-x) + 1 - F_n(x-)}{2(1 - F_n(x-))}, \qquad (2.1)$$

### A quick visualization

Simulation of 100 normally distributed numbers. The weight transformation is applied.









### A quick visualization

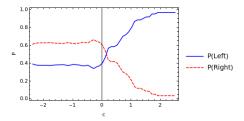


Figura: Left and right cumulative distribution functions.

## H.F. Coronel-Brizio, et al. (2007) Assessing symmetry of financial returns series. Physica A 383.

• The statistic  $T_n$  around a test point c is

$$T_n = -\frac{2}{n} \sum_{i=1}^n \log R(|X_i|).$$

where

$$\log R(x) = nF_n(c-x)\log \frac{F_n(c-x) + 1 - F_n(x-c)}{2F_n(c-x)} + n(1 - F_n(x-c))\log \frac{F_n(c-x) + 1 - F_n(x-c)}{2(1 - F_n(x-c))}.$$

and  $F_n$  is the cumulative empirical distribution function.

# H.F. Coronel-Brizio, et al. (2007) Assessing symmetry of financial returns series. Physica A 383.

• We can test for c in an interval  $[C_{min}, C_{max}]$ , for example

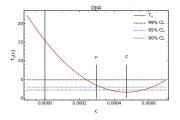


Figura:  $T_n$  symmetry test evaluated on the returns of the DJIA index.

• We can affirm with 95% certainty that the distribution is symmetrical in a interval between 0.0003 and 0.0006.

### The problem of scale

• The time lag in which the returns are defined are particularly relevant when measuring symmetry.

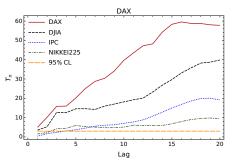


Figura:  $T_n(0)$  with returns of different lags.

Artifacts are introduced with this methodology.

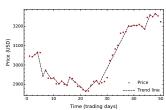


### Definition of trend returns

- Multiscale return definition. Given a financial time series  $P_1, \ldots, P_n$ , we define an "uninterrupted trend" of duration k, a succession of k+1 consecutive values of the given time series where each value is greater than the preceding one.
- We define their corresponding trend return as:

$$R_m^k := \log(P_{m+k}) - \log(P_m) \tag{1}$$

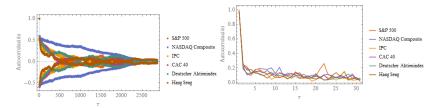
Where m indexes the different trends and k indicates the duration in days of m-Th trend.





### This definition also comes with its stylized facts I

• Volatility clustering is present in trend returns  $R_m^k$ 



- (a) Trend returns autocorrelation. (b) Quadratic trend returns autocorrelation
- So trend returns still can be decomposed as

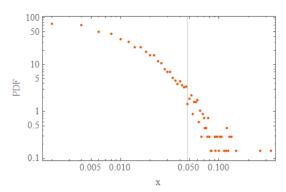
$$R_T(t) = \sigma_R(t)\epsilon(t),$$

where  $\epsilon(t)$  is gaussian noise and  $\sigma(t)$  the volatility.



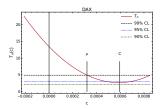
### This definition also comes with its stylized facts II

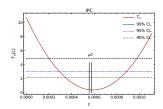
H.F. Coronel-Brizio, A.R. Hernandez-Montoya (2005). On fitting the Pareto-Levy distribution to stock market index data: selecting a suitable cut off value. Physica A 354 437-449. Describes a procedure to select the cutoff point where returns start behaving as a power law.



#### Results

### Symmetry tests: Simple returns





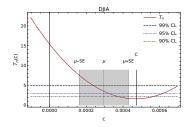
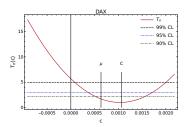
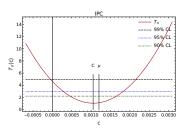


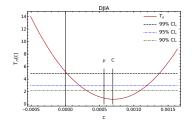
Figura: Symmetry test for 1 day returns

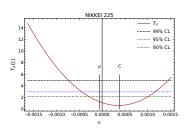


### Symmetry tests: Trend returns

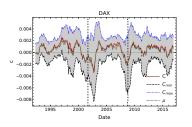


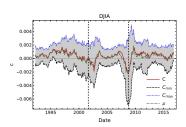


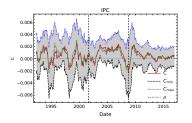


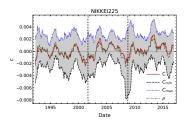


## Time evolution of the most plausible point of symmetry: Simple returns

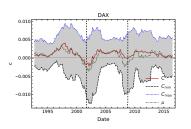


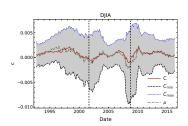


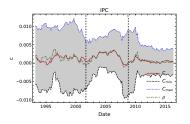


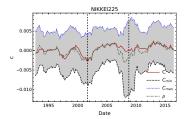


## Time evolution of the most plausible point of symmetry: Trend returns

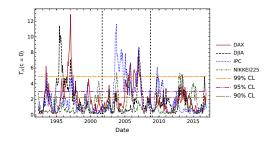


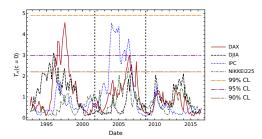






### Symmetry around zero







### Conclusion

- We can measure symmetry of the returns using the T<sub>n</sub> statistic. We also defined the trend returns in order to select a scale independent observable.
- Symmetry is not constant. Particularly symmetry around c=0 tends to be not plausible in dates closer to extreme events.

Thank you all for your attention.

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