

# Instrumental Variables

*UNDERSTANDING IV*

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# Roadmap

## Where do (Good) Instruments Come From?

- True Lotteries

- Natural Experiments

- Panel Data

## 2SLS Mechanics

- Just-Identified IV

- Overidentification

## Weak and Many Instruments

- Weak IV

- Many IVs

## Subtlties of the Validity Condition

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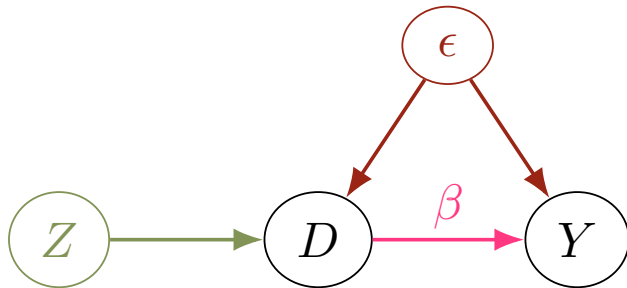
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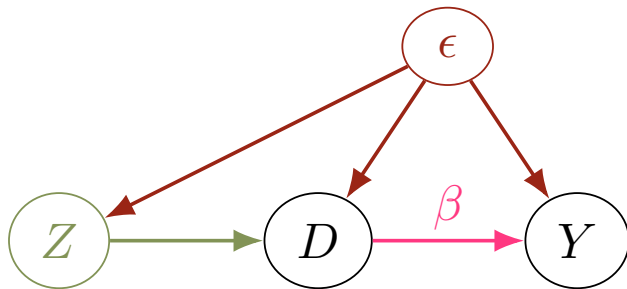
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## A Valid Instrument

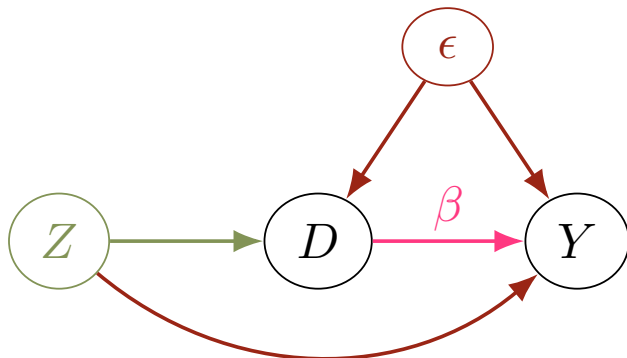




# A Violation of As-Good-As-Random Assignment



## A Violation of Exclusion



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- We still need to worry about violations of the exclusion restriction
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“Gold standard” IV: a randomized offer to participate in a program, with  $X_i$  recording program participation

- Exclusion restriction likely to hold for any  $Y_i$ , by construction
- Relevance almost guaranteed (provided people want the program!)

# Example

## *Charter School Lotteries*

Abdulkadiroglu et al. (2016) are interested in whether going to a “charter” middle school increases standardized test scores

- Charter students tend to score better, but we worry about selection
- History of doubting educational inputs, since Coleman (1966)

# Example

## Charter School Lotteries

We leverage an institutional feature of charters: *admission lotteries*

- When more kids want to enroll than there are seats, admission offers  $Z_i \in \{0, 1\}$  are effectively drawn from a hat
- Offers plausibly only affect later test scores  $Y_i$  by changing charter enrollment  $D_i \in \{0, 1\}$ , so are plausibly valid instruments
- We need to control for lottery fixed effects (“risk sets”) to make  $Z_i$  as-good-as-randomly assigned – more on this soon



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We study a particular charter (UP Academy), which is “takeover”

- Two offer IVs: “immediate” (on lottery night) and from a waitlist

# Lottery IV Estimates of UP Test Score Effects

TABLE 8—LOTTERY IV ESTIMATES OF UP EFFECTS

		2SLS				
		Comparison group mean (1)	OLS (2)	First stage		
				Immediate offer (3)	Waitlist offer (4)	Enrollment effect (5)
<i>Panel A. All grades</i>						
(Sixth through eighth)	Math (N = 2,202)	0.059	0.301 (0.022)	0.760 (0.063)	0.562 (0.067)	0.270 (0.056)
	ELA (N = 2,205)	0.103	0.148 (0.020)	0.759 (0.063)	0.562 (0.067)	0.118 (0.051)

# Where do IVs Come From?

## 2. Natural Experiments

Without appealing to literal randomization, we may credibly argue  $Z_i$  is as-good-as-randomly assigned conditional on some  $\mathbf{W}_i$

- Such “natural experiments” rely on a selection-on-observables argument (for  $Z_i$ , instead  $D_i$ )
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## *Quarter-of-birth*

Angrist and Krueger (1991) famously estimate labor market returns to schooling with a creative IV: student quarter-of-birth

- Compulsory schooling requirements prevent students from dropping before the day they turn 16 (used to be more binding)
- Fixed school start dates mean students who drop out at 16 get more or less schooling depending on their birth date

# Example

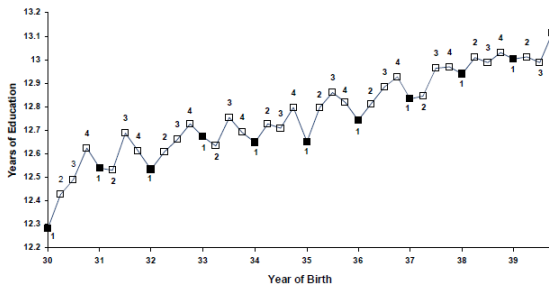
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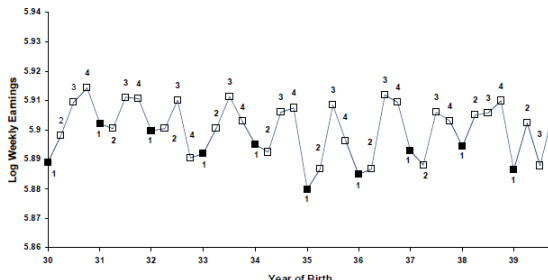
- Compulsory schooling requirements prevent students from dropping before the day they turn 16 (used to be more binding)
- Fixed school start dates mean students who drop out at 16 get more or less schooling depending on their birth date
- Quarter-of-birth seems quasi-randomly assigned — is it excludable? See Buckles and Hungerman (2013)...

# The Quarter-of-Birth Natural Experiment: Visualized

A. Average Education by Quarter of Birth (first stage)



B. Average Weekly Wage by Quarter of Birth (reduced form)



# Quarter-of-Birth IV Estimates of Returns to Schooling

Table 4.1.1: 2SLS estimates of the economic returns to schooling

	OLS		2SLS			
	(1)	(2)	(3)	(4)	(5)	(6)
Years of education	0.075 (0.0004)	0.072 (0.0004)	0.103 (0.024)	0.112 (0.021)	0.106 (0.026)	0.108 (0.019)
<i>Covariates:</i>						
9 year of birth dummies		✓			✓	✓
50 state of birth dummies		✓			✓	✓
<i>Instruments:</i>						
			dummy for QOB=1	dummy for QOB=1 or QOB=2	dummy for QOB=1	full set of QOB dummies



# Where do IVs Come From?

## 3. Panel Data

We might also combine IV + difference-in-differences identification

- E.g. instrument with  $Z_i \times Post_t$ , controlling for  $Z_i$  and  $Post_t$  FEs
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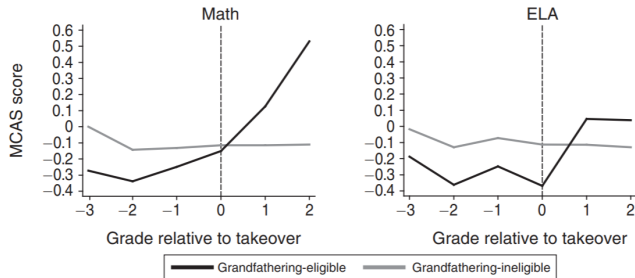
## *School Lottery*

Abdulkadiroglu et al. (2016) complement their lottery analysis of takeover charters with an instrumented diff-in-diff analysis

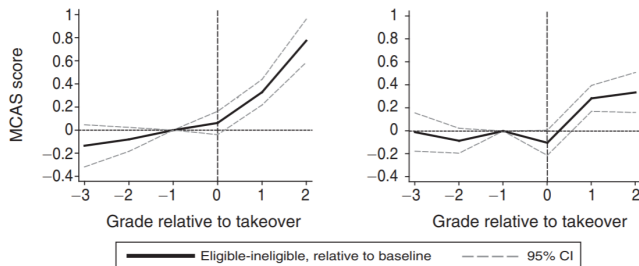
- Students enrolled in the “legacy” public school were eligible for being “grandfathered” into UP, without having to apply to the charter
- We compare their trends in test scores & enrollment to a matched comparison group of observably-similar students at other schools

# Grandfathering IV: Visualized

Panel A. Score levels



Panel B. Score DD



# Grandfathering IV Estimates of UP Test Score Effects

TABLE 7—GRANDFATHERING IV ESTIMATES OF UP EFFECTS

		2SLS			
		Comparison group mean (1)	OLS (2)	First stage (3)	Enrollment effect (4)
<i>Panel A. All grades</i>					
(Seventh through eighth)	Math (N = 1,543)	−0.233	0.400 (0.032)	1.051 (0.040)	0.321 (0.039)
	ELA (N = 1,539)	−0.214	0.296 (0.035)	1.040 (0.041)	0.394 (0.044)

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2SLS Mechanics

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- Overidentification

Weak and Many Instruments

- Weak IV

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# Just-Identified IV

As you likely know, the general `ivregress` command (or its equivalent `fixes::feols` in R) allows for controls and multiple treatments / instruments

- When # treatment = # instruments, we say the IV is “just-identified”:

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$$Y_i = \beta D_i + \mathbf{W}_i' \boldsymbol{\gamma} + \varepsilon_i \quad (\text{second stage})$$

$$X_i = \pi Z_i + \mathbf{W}_i' \boldsymbol{\mu} + \eta_i \quad (\text{first stage})$$

where  $\mathbf{W}_i$  includes a constant.



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- Can use Frisch-Waugh-Lovell to “partial out”  $\mathbf{W}_i$  from  $Y_i$ ,  $X_i$ ,  $D_i$ , and so get back to an IV regression without controls

# Overidentification

Sometimes we have more than one instrument  $Z_{i\ell}$ , for  $\ell = 1, \dots, L$ .

This leads to an “overidentified” IV regression:

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where  $\mathbf{Z}_i = [Z_{i1}, \dots, Z_{iL}]'$ . Reduced form:  $Y_i = \mathbf{Z}_i' \boldsymbol{\rho} + \mathbf{W}_i' \boldsymbol{\kappa} + \nu_i$

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Overidentification can yield tests of IV validity

- Intuitively, 2SLS checks whether all the  $Z_{i\ell}$  yields the same IV estimate, which is sensible in a constant-effects model...

# Putting the “2S” in “2SLS”

You'll notice I haven't actually defined 2SLS beyond the simple case

- Before we had  $\beta^{IV} = \frac{Cov(Z_i, Y_i)}{Cov(Z_i, D_i)}$  leading to  $\hat{\beta}^{IV} = \frac{\widehat{Cov}(Z_i, Y_i)}{\widehat{Cov}(Z_i, D_i)}$
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A more useful way to define 2SLS is by a two-step procedure:

- First regress  $D_i$  on all  $Z_{i\ell}$  and  $W_{ik}$
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The proof of this follows from some (simple) linear algebra

- Intuitively, regressing  $Y_i$  on  $\hat{\pi}^{OLS} Z_i$  gives a scaled RE:

## Putting the “2S” in “2SLS”

$$\hat{\beta}^{IV} = \frac{\hat{\rho}^{OLS}}{\hat{\pi}^{OLS}}$$

Although easy, you should never do such “manual 2SLS” yourself!

- Your point estimates will be right, but your SEs won't be!

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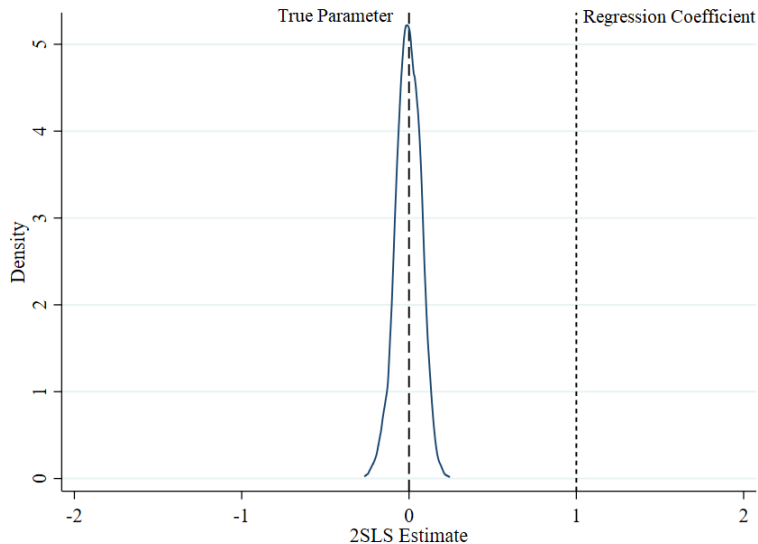
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Much made of this over the years, but Angrist and Kolesár (2022) argue recently that we shouldn't worry too much

- The SE increase tends to be large enough to “cover up” the bias
- Just-id. 2SLS is “approximately median-unbiased” (as it is LIML)

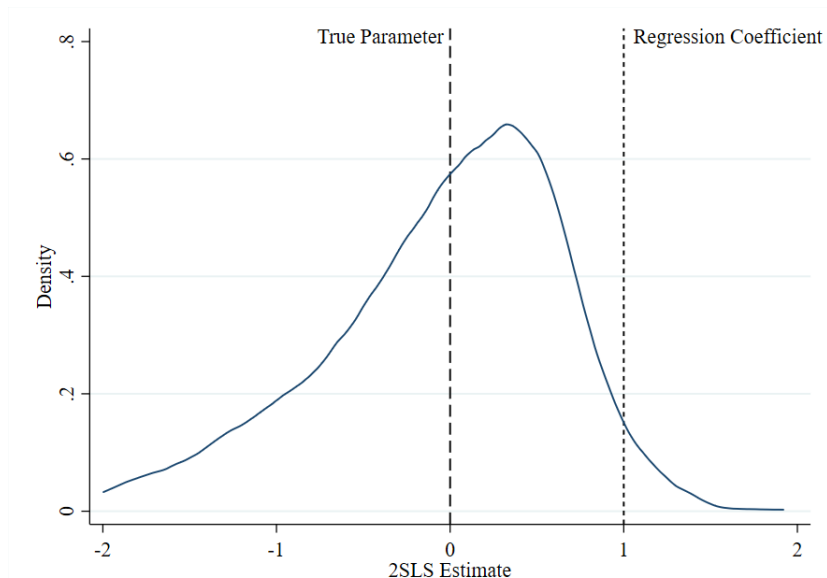
# Weak Instruments: Visualized

Monte Carlo:  $Y_i = \varepsilon_i$ ,  $D_i = \Pi Z_i + \eta_i$ :  $\Pi = \text{Var}(\varepsilon_i) = \text{Var}(\eta_i) = 1$



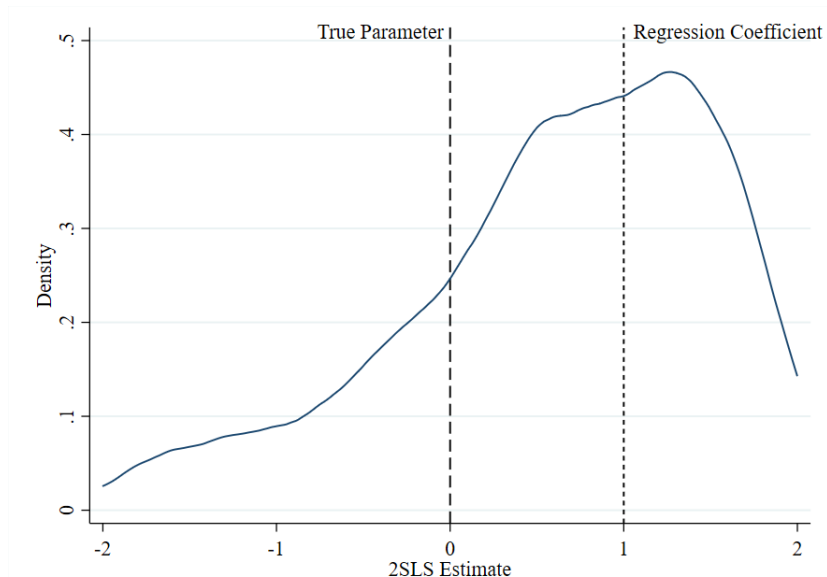
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Monte Carlo:  $Y_i = \varepsilon_i$ ,  $D_i = \Pi Z_i + \eta_i$ :  $\Pi = 0.1$  (Weaker)



# Weak Instruments: Visualized

Monte Carlo:  $Y_i = \varepsilon_i$ ,  $D_i = \Pi Z_i + \eta_i$ ;  $\Pi = 0.01$  (Very Weak)



# Many IVs

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Many-IV bias is also towards OLS, but unlike before SEs go *down*

- Intuitively, a more flexible FS tends to fit  $D_i$  better  $\rightarrow$  more power
- But we can have *overfitting* with lots of  $Z_i \rightarrow$  essentially recreate  $D_i$

# Many IVs

A more pernicious problem is many-instrument bias, when overid

- Also tends to manifest in low first-stage  $F$ 's, so also good to check

Many-IV bias is also towards OLS, but unlike before SEs go *down*

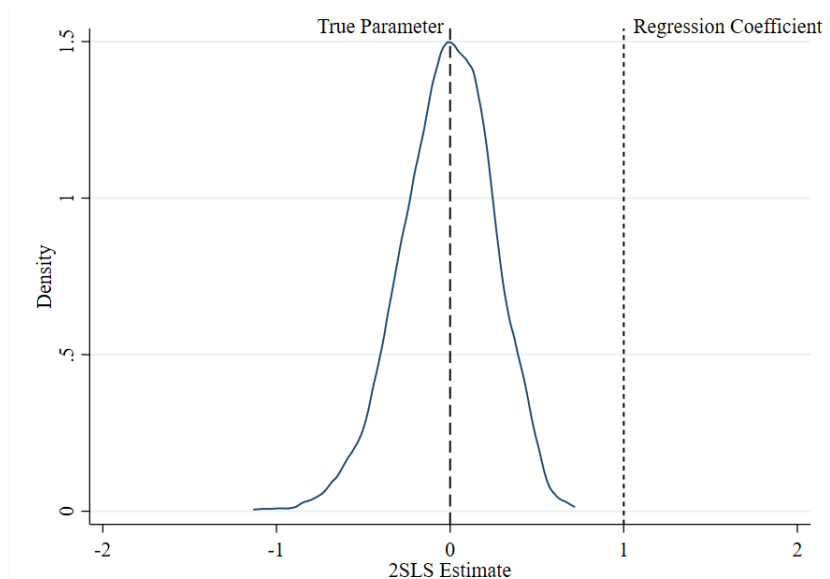
- Intuitively, a more flexible FS tends to fit  $D_i$  better  $\rightarrow$  more power
- But we can have *overfitting* with lots of  $Z_i \rightarrow$  essentially recreate  $D_i$

As we'll see, this bias is especially relevant in judge IV designs

- Potentially many judge assignment indicators as the instrument
- Leave-out corrections (e.g. Angrist et al. 1999) have been adapted to this setting in recent years (e.g. Kolesár 2013)

# Weak and Many Instruments VIII

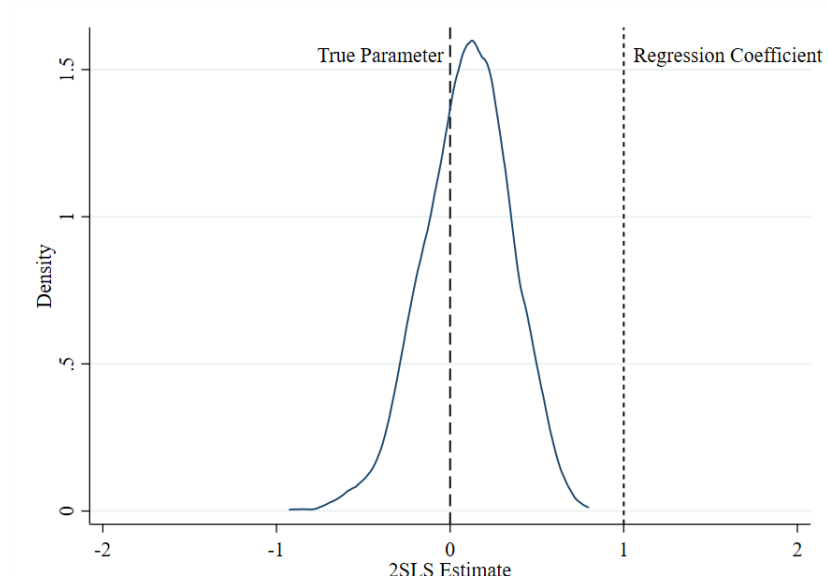
Monte Carlo:  $Y_i = \varepsilon_i$ ,  $X_i = \Pi Z_{i1} + \eta_i$ : IV with one  $Z_{i1}$





# Weak and Many Instruments IX

Monte Carlo:  $Y_i = \varepsilon_i$ ,  $X_i = \Pi Z_{i1} + \eta_i$ : IV with ten  $Z_{ij}$



# Weak and Many Instruments X

Monte Carlo:  $Y_i = \varepsilon_i$ ,  $X_i = \Pi Z_{i1} + \eta_i$ : IV with 100  $Z_{ij}$

