

Instrumental Variables

UNDERSTANDING IV



Roadmap

Where do (Good) Instruments Come From?

- True Lotteries

- Natural Experiments

- Panel Data

2SLS Mechanics

- Just-Identified IV

- Overidentification

Weak and Many Instruments

- Weak IV

- Many IVs

Subtlties of the Validity Condition

- To apply IV, we need to make a good case for instrument validity (note we can always check relevance!)

Subtleties of the Validity Condition

- To apply IV, we need to make a good case for instrument validity (note we can always check relevance!)
- Consider our simple causal model, $Y_i = \alpha + \beta D_i + \varepsilon_i$. Validity $Cov(Z_i, \varepsilon_i) = 0$ intuitively requires two distinct assumptions:

Subtlties of the Validity Condition

- To apply IV, we need to make a good case for instrument validity (note we can always check relevance!)
- Consider our simple causal model, $Y_i = \alpha + \beta D_i + \varepsilon_i$. Validity $Cov(Z_i, \varepsilon_i) = 0$ intuitively requires two distinct assumptions:
 - *As-good-as-random assignment*: individuals with higher/lower potential earnings face the same distribution of Z_i
 - *Exclusion*: the “assignment” of Z_i only affects Y_i through D_i

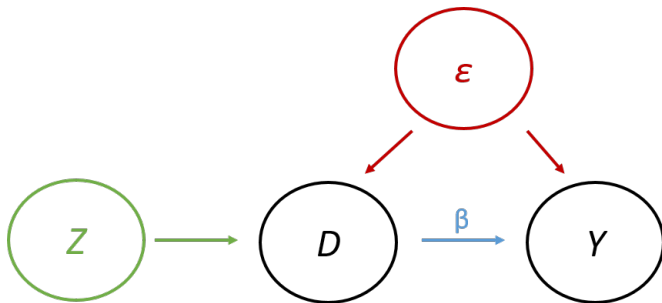
Subtleties of the Validity Condition

- To apply IV, we need to make a good case for instrument validity (note we can always check relevance!)
- Consider our simple causal model, $Y_i = \alpha + \beta D_i + \varepsilon_i$. Validity $Cov(Z_i, \varepsilon_i) = 0$ intuitively requires two distinct assumptions:
 - *As-good-as-random assignment*: individuals with higher/lower potential earnings face the same distribution of Z_i
 - *Exclusion*: the “assignment” of Z_i only affects Y_i through D_i
- Confusingly, old-school econometrics texts sometimes refer to $Cov(Z_i, \varepsilon_i) = 0$ as the “exclusion restriction”

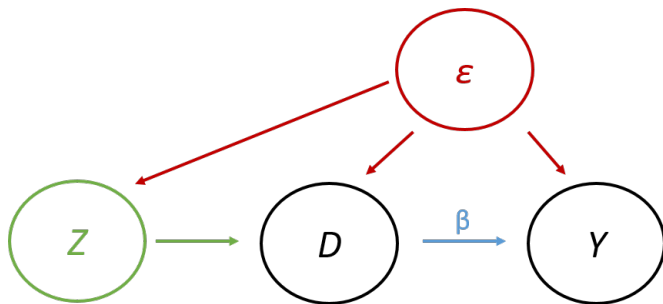
Subtleties of the Validity Condition

- To apply IV, we need to make a good case for instrument validity (note we can always check relevance!)
- Consider our simple causal model, $Y_i = \alpha + \beta D_i + \varepsilon_i$. Validity $Cov(Z_i, \varepsilon_i) = 0$ intuitively requires two distinct assumptions:
 - *As-good-as-random assignment*: individuals with higher/lower potential earnings face the same distribution of Z_i
 - *Exclusion*: the “assignment” of Z_i only affects Y_i through D_i
- Confusingly, old-school econometrics texts sometimes refer to $Cov(Z_i, \varepsilon_i) = 0$ as the “exclusion restriction”
 - More modern IV texts take care to distinguish between these two conceptually distinct requirements...

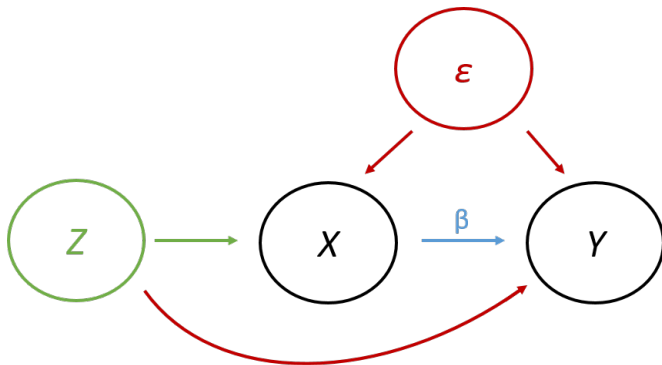
A Valid Instrument



A Violation of As-Good-As-Random Assignment



A Violation of Exclusion



Where do IVs Come From? 1) True Lotteries

- One sure-fire way to ensure that a Z_i is as-good-as-randomly assigned is...

Where do IVs Come From? 1) True Lotteries

- One sure-fire way to ensure that a Z_i is as-good-as-randomly assigned is... to randomly assign it!

Where do IVs Come From? 1) True Lotteries

- One sure-fire way to ensure that a Z_i is as-good-as-randomly assigned is... to randomly assign it!
 - Some of the best IVs come from lotteries, either run by the researcher (e.g. an RCT) or so-called “natural experiments”
 - We still need to worry about violations of the exclusion restriction
 - Relevance holds when Z_i has some effect on X_i

Where do IVs Come From? 1) True Lotteries

- One sure-fire way to ensure that a Z_i is as-good-as-randomly assigned is... to randomly assign it!
 - Some of the best IVs come from lotteries, either run by the researcher (e.g. an RCT) or so-called “natural experiments”
 - We still need to worry about violations of the exclusion restriction
 - Relevance holds when Z_i has some effect on X_i
- “Gold standard” IV: a randomized offer to participate in a program, with X_i recording program participation
 - Exclusion restriction likely to hold for any Y_i , by construction
 - Relevance almost guaranteed (provided people want the program!)

Example: Charter School Lotteries

- Abdulkadiroglu et al. (2016) are interested in whether going to a “charter” middle school increases standardized test scores
 - Charter students tend to score better, but we worry about selection
 - History of doubting educational inputs, since Coleman (1966)

Example: Charter School Lotteries

- Abdulkadiroglu et al. (2016) are interested in whether going to a “charter” middle school increases standardized test scores
 - Charter students tend to score better, but we worry about selection
 - History of doubting educational inputs, since Coleman (1966)
- We leverage an institutional feature of charters: *admission lotteries*
 - When more kids want to enroll than there are seats, admission offers $Z_i \in \{0, 1\}$ are effectively drawn from a hat
 - Offers plausibly only affect later test scores Y_i by changing charter enrollment $D_i \in \{0, 1\}$, so are plausibly valid instruments
 - We need to control for lottery fixed effects (“risk sets”) to make Z_i as-good-as-randomly assigned – more on this soon

Example: Charter School Lotteries

- Abdulkadiroglu et al. (2016) are interested in whether going to a “charter” middle school increases standardized test scores
 - Charter students tend to score better, but we worry about selection
 - History of doubting educational inputs, since Coleman (1966)
- We leverage an institutional feature of charters: *admission lotteries*
 - When more kids want to enroll than there are seats, admission offers $Z_i \in \{0, 1\}$ are effectively drawn from a hat
 - Offers plausibly only affect later test scores Y_i by changing charter enrollment $D_i \in \{0, 1\}$, so are plausibly valid instruments
 - We need to control for lottery fixed effects (“risk sets”) to make Z_i as-good-as-randomly assigned – more on this soon
- We study a particular charter (UP Academy), which is “takeover”
 - Two offer IVs: “immediate” (on lottery night) and from a waitlist

Lottery IV Estimates of UP Test Score Effects

TABLE 8—LOTTERY IV ESTIMATES OF UP EFFECTS

		2SLS				
		Comparison group mean (1)	OLS (2)	First stage		Enrollment effect (5)
				Immediate offer (3)	Waitlist offer (4)	
<i>Panel A. All grades</i>						
(Sixth through eighth)	Math (N = 2,202)	0.059	0.301 (0.022)	0.760 (0.063)	0.562 (0.067)	0.270 (0.056)
	ELA (N = 2,205)	0.103	0.148 (0.020)	0.759 (0.063)	0.562 (0.067)	0.118 (0.051)

Where do IVs Come From? 2) Natural Experiments

- Without appealing to literal randomization, we may credibly argue Z_i is as-good-as-randomly assigned conditional on some \mathbf{W}_i
 - Such “natural experiments” rely on a selection-on-observables argument (for Z_i , instead D_i)
 - Still worry about exclusion: Z_i cannot affect Y_i except through D_i

Where do IVs Come From? 2) Natural Experiments

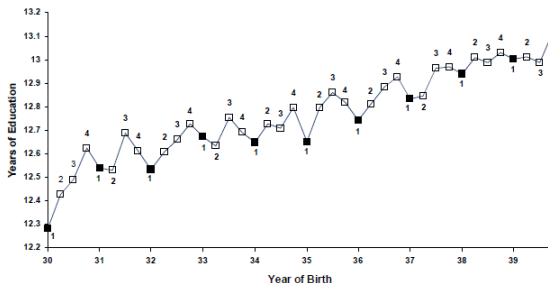
- Without appealing to literal randomization, we may credibly argue Z_i is as-good-as-randomly assigned conditional on some \mathbf{W}_i
 - Such “natural experiments” rely on a selection-on-observables argument (for Z_i , instead D_i)
 - Still worry about exclusion: Z_i cannot affect Y_i except through D_i
- Angrist and Krueger (1991) famously estimate labor market returns to schooling with a creative IV: student quarter-of-birth
 - Compulsory schooling requirements prevent students from dropping before the day they turn 16 (used to be more binding)
 - Fixed school start dates mean students who drop out at 16 get more or less schooling depending on their birth date

Where do IVs Come From? 2) Natural Experiments

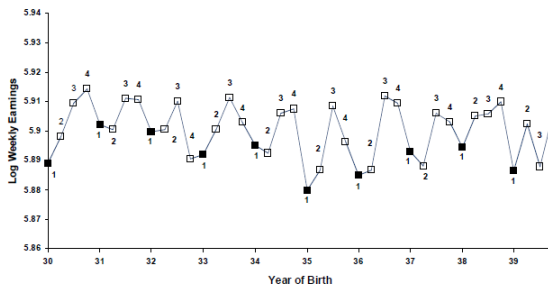
- Without appealing to literal randomization, we may credibly argue Z_i is as-good-as-randomly assigned conditional on some \mathbf{W}_i
 - Such “natural experiments” rely on a selection-on-observables argument (for Z_i , instead D_i)
 - Still worry about exclusion: Z_i cannot affect Y_i except through D_i
- Angrist and Krueger (1991) famously estimate labor market returns to schooling with a creative IV: student quarter-of-birth
 - Compulsory schooling requirements prevent students from dropping before the day they turn 16 (used to be more binding)
 - Fixed school start dates mean students who drop out at 16 get more or less schooling depending on their birth date
 - Quarter-of-birth seems quasi-randomly assigned — is it excludable? See Buckles and Hungerman (2013)...

The Quarter-of-Birth Natural Experiment: Visualized

A. Average Education by Quarter of Birth (first stage)



B. Average Weekly Wage by Quarter of Birth (reduced form)



Quarter-of-Birth IV Estimates of Returns to Schooling

Table 4.1.1: 2SLS estimates of the economic returns to schooling

	OLS		2SLS			
	(1)	(2)	(3)	(4)	(5)	(6)
Years of education	0.075 (0.0004)	0.072 (0.0004)	0.103 (0.024)	0.112 (0.021)	0.106 (0.026)	0.108 (0.019)
<i>Covariates:</i>						
9 year of birth dummies		✓			✓	✓
50 state of birth dummies		✓			✓	✓
<i>Instruments:</i>			dummy for QOB=1	dummy for QOB=1 or QOB=2	dummy for QOB=1	full set of QOB dummies

Where do IVs Come From? 3) Panel Data

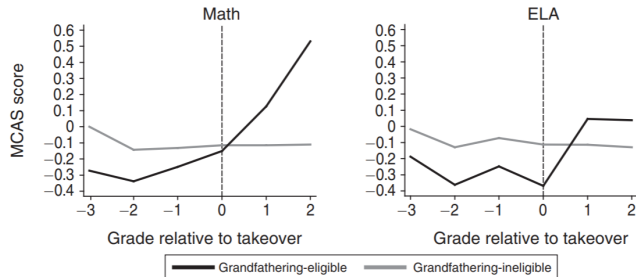
- We might also combine IV + difference-in-difference identification
 - E.g. instrument with $Z_i \times Post_t$, controlling for Z_i and $Post_t$ FEs
 - This requires two parallel trends assumptions, for the RF and FS
 - Still need to worry about the exclusion restriction, as always

Where do IVs Come From? 3) Panel Data

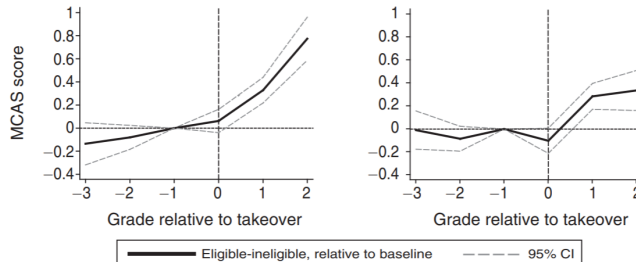
- We might also combine IV + difference-in-difference identification
 - E.g. instrument with $Z_i \times Post_t$, controlling for Z_i and $Post_t$ FEs
 - This requires two parallel trends assumptions, for the RF and FS
 - Still need to worry about the exclusion restriction, as always
- Abdulkadiroglu et al. (2016) complement their lottery analysis of takeover charters with an instrumented diff-in-diff analysis
 - Students enrolled in the “legacy” public school were eligible for being “grandfathered” into UP, without having to apply to the charter
 - We compare their trends in test scores & enrollment to a matched comparison group of observably-similar students at other schools

Grandfathering IV: Visualized

Panel A. Score levels



Panel B. Score DD



Grandfathering IV Estimates of UP Test Score Effects

TABLE 7—GRANDFATHERING IV ESTIMATES OF UP EFFECTS

		Comparison group mean (1)	OLS (2)	2SLS	
				First stage (3)	Enrollment effect (4)
<i>Panel A. All grades</i>					
(Seventh through eighth)	Math (N = 1,543)	−0.233	0.400 (0.032)	1.051 (0.040)	0.321 (0.039)
	ELA (N = 1,539)	−0.214	0.296 (0.035)	1.040 (0.041)	0.394 (0.044)

Roadmap

Where do (Good) Instruments Come From?

- True Lotteries

- Natural Experiments

- Panel Data

2SLS Mechanics

- Just-Identified IV

- Overidentification

Weak and Many Instruments

- Weak IV

- Many IVs

Just-Identified IV

- As you likely know, the general *ivregress* command (or its equivalent in R) allows for controls and multiple treatments / instruments
 - When # treatment = # instruments, we say the IV is “just-identified”:

Just-Identified IV

- As you likely know, the general *ivregress* command (or its equivalent in R) allows for controls and multiple treatments / instruments

→ When # treatment = # instruments, we say the IV is “just-identified”:

$$Y_i = \beta D_i + \mathbf{W}_i' \boldsymbol{\gamma} + \varepsilon_i \quad (\text{second stage})$$

$$X_i = \pi Z_i + \mathbf{W}_i' \boldsymbol{\mu} + \eta_i \quad (\text{first stage})$$

where \mathbf{W}_i includes a constant.

Just-Identified IV

- As you likely know, the general *ivregress* command (or its equivalent in R) allows for controls and multiple treatments / instruments

→ When # treatment = # instruments, we say the IV is “just-identified”:

$$Y_i = \beta D_i + \mathbf{W}_i' \boldsymbol{\gamma} + \varepsilon_i \quad (\text{second stage})$$

$$X_i = \pi Z_i + \mathbf{W}_i' \boldsymbol{\mu} + \eta_i \quad (\text{first stage})$$

where \mathbf{W}_i includes a constant. Reduced form: $Y_i = \rho Z_i + \mathbf{W}_i' \boldsymbol{\kappa} + \nu_i$

Just-Identified IV

- As you likely know, the general *ivregress* command (or its equivalent in R) allows for controls and multiple treatments / instruments

→ When # treatment = # instruments, we say the IV is “just-identified”:

$$Y_i = \beta D_i + \mathbf{W}_i' \boldsymbol{\gamma} + \varepsilon_i \quad (\text{second stage})$$

$$X_i = \pi Z_i + \mathbf{W}_i' \boldsymbol{\mu} + \eta_i \quad (\text{first stage})$$

where \mathbf{W}_i includes a constant. Reduced form: $Y_i = \rho Z_i + \mathbf{W}_i' \boldsymbol{\kappa} + \nu_i$

- Same identification logic as before:

→ Validity: $Cov(Z_i, \varepsilon_i) = 0$, allowing $Cov(Z_i, \mathbf{W}_i) \neq 0$

Just-Identified IV

- As you likely know, the general *ivregress* command (or its equivalent in R) allows for controls and multiple treatments / instruments

→ When # treatment = # instruments, we say the IV is “just-identified”:

$$Y_i = \beta D_i + \mathbf{W}_i' \boldsymbol{\gamma} + \varepsilon_i \quad (\text{second stage})$$

$$X_i = \pi Z_i + \mathbf{W}_i' \boldsymbol{\mu} + \eta_i \quad (\text{first stage})$$

where \mathbf{W}_i includes a constant. Reduced form: $Y_i = \rho Z_i + \mathbf{W}_i' \boldsymbol{\kappa} + \nu_i$

- Same identification logic as before:
 - Validity: $Cov(Z_i, \varepsilon_i) = 0$, allowing $Cov(Z_i, \mathbf{W}_i) \neq 0$
 - Relevance: $\pi \neq 0$, so Z_i and D_i are correlated controlling for \mathbf{W}_i

Just-Identified IV

- As you likely know, the general *ivregress* command (or its equivalent in R) allows for controls and multiple treatments / instruments

→ When # treatment = # instruments, we say the IV is “just-identified”:

$$Y_i = \beta D_i + \mathbf{W}_i' \boldsymbol{\gamma} + \varepsilon_i \quad (\text{second stage})$$

$$X_i = \pi Z_i + \mathbf{W}_i' \boldsymbol{\mu} + \eta_i \quad (\text{first stage})$$

where \mathbf{W}_i includes a constant. Reduced form: $Y_i = \rho Z_i + \mathbf{W}_i' \boldsymbol{\kappa} + \nu_i$

- Same identification logic as before:
 - Validity: $Cov(Z_i, \varepsilon_i) = 0$, allowing $Cov(Z_i, \mathbf{W}_i) \neq 0$
 - Relevance: $\pi \neq 0$, so Z_i and D_i are correlated controlling for \mathbf{W}_i
- IV is still “reduced form over first stage”: $(\beta = \rho/\pi)$

Just-Identified IV

- As you likely know, the general *ivregress* command (or its equivalent in R) allows for controls and multiple treatments / instruments

→ When # treatment = # instruments, we say the IV is “just-identified”:

$$Y_i = \beta D_i + \mathbf{W}_i' \boldsymbol{\gamma} + \varepsilon_i \quad (\text{second stage})$$

$$X_i = \pi Z_i + \mathbf{W}_i' \boldsymbol{\mu} + \eta_i \quad (\text{first stage})$$

where \mathbf{W}_i includes a constant. Reduced form: $Y_i = \rho Z_i + \mathbf{W}_i' \boldsymbol{\kappa} + \nu_i$

- Same identification logic as before:
 - Validity: $Cov(Z_i, \varepsilon_i) = 0$, allowing $Cov(Z_i, \mathbf{W}_i) \neq 0$
 - Relevance: $\pi \neq 0$, so Z_i and D_i are correlated controlling for \mathbf{W}_i
- IV is still “reduced form over first stage”: $(\beta = \rho/\pi)$
 - Can use Frisch-Waugh-Lovell to “partial out” \mathbf{W}_i from Y_i , X_i , D_i , and so get back to an IV regression without controls

Overidentification

- Sometimes we have more than one instrument $Z_{i\ell}$, for $\ell = 1, \dots, L$

Overidentification

- Sometimes we have more than one instrument $Z_{i\ell}$, for $\ell = 1, \dots, L$
→ This leads to an “overidentified” IV regression:

$$Y_i = \beta D_i + \mathbf{W}_i' \boldsymbol{\gamma} + \varepsilon_i \quad (\text{second stage})$$

$$X_i = \mathbf{Z}_i' \boldsymbol{\pi} + \mathbf{W}_i' \boldsymbol{\mu} + \eta_i \quad (\text{first stage})$$

where $\mathbf{Z}_i = [Z_{i1}, \dots, Z_{iL}]'$. Reduced form: $Y_i = \mathbf{Z}_i' \boldsymbol{\rho} + \mathbf{W}_i' \boldsymbol{\kappa} + \nu_i$

Overidentification

- Sometimes we have more than one instrument $Z_{i\ell}$, for $\ell = 1, \dots, L$
 - This leads to an “overidentified” IV regression:

$$Y_i = \beta D_i + \mathbf{W}_i' \boldsymbol{\gamma} + \varepsilon_i \quad (\text{second stage})$$

$$X_i = \mathbf{Z}_i' \boldsymbol{\pi} + \mathbf{W}_i' \boldsymbol{\mu} + \eta_i \quad (\text{first stage})$$

where $\mathbf{Z}_i = [Z_{i1}, \dots, Z_{iL}]'$. Reduced form: $Y_i = \mathbf{Z}_i' \boldsymbol{\rho} + \mathbf{W}_i' \boldsymbol{\kappa} + \nu_i$

- Validity: $Cov(Z_{i\ell}, \varepsilon_i) = 0$ for all ℓ
 - “Overidentified” because we could use any $Z_{i\ell}$ to estimate $\beta = \boldsymbol{\rho}_\ell / \boldsymbol{\pi}_\ell$

Overidentification

- Sometimes we have more than one instrument $Z_{i\ell}$, for $\ell = 1, \dots, L$
 - This leads to an “overidentified” IV regression:

$$Y_i = \beta D_i + \mathbf{W}_i' \boldsymbol{\gamma} + \varepsilon_i \quad (\text{second stage})$$

$$X_i = \mathbf{Z}_i' \boldsymbol{\pi} + \mathbf{W}_i' \boldsymbol{\mu} + \eta_i \quad (\text{first stage})$$

where $\mathbf{Z}_i = [Z_{i1}, \dots, Z_{iL}]'$. Reduced form: $Y_i = \mathbf{Z}_i' \boldsymbol{\rho} + \mathbf{W}_i' \boldsymbol{\kappa} + \nu_i$

- Validity: $Cov(Z_{i\ell}, \varepsilon_i) = 0$ for all ℓ
 - “Overidentified” because we could use any $Z_{i\ell}$ to estimate $\beta = \boldsymbol{\rho}_\ell / \boldsymbol{\pi}_\ell$
 - Relevance: $\boldsymbol{\pi}_\ell \neq 0$ for at least some ℓ

Overidentification

- Sometimes we have more than one instrument $Z_{i\ell}$, for $\ell = 1, \dots, L$
 - This leads to an “overidentified” IV regression:

$$Y_i = \beta D_i + \mathbf{W}_i' \boldsymbol{\gamma} + \varepsilon_i \quad (\text{second stage})$$

$$X_i = \mathbf{Z}_i' \boldsymbol{\pi} + \mathbf{W}_i' \boldsymbol{\mu} + \eta_i \quad (\text{first stage})$$

where $\mathbf{Z}_i = [Z_{i1}, \dots, Z_{iL}]'$. Reduced form: $Y_i = \mathbf{Z}_i' \boldsymbol{\rho} + \mathbf{W}_i' \boldsymbol{\kappa} + \nu_i$

- Validity: $Cov(Z_{i\ell}, \varepsilon_i) = 0$ for all ℓ
 - “Overidentified” because we could use any $Z_{i\ell}$ to estimate $\beta = \boldsymbol{\rho}_\ell / \boldsymbol{\pi}_\ell$
 - Relevance: $\boldsymbol{\pi}_\ell \neq 0$ for at least some ℓ
- Overidentification can yield tests of IV validity
 - Intuitively, 2SLS checks whether all the $Z_{i\ell}$ yields the same IV estimate, which is sensible in a constant-effects model...

Putting the “2S” in “2SLS”

- You'll notice I haven't actually defined 2SLS beyond the simple case
 - Before we had $\beta^{IV} = \frac{Cov(Z_i, Y_i)}{Cov(Z_i, D_i)}$ leading to $\hat{\beta}^{IV} = \frac{\widehat{Cov}(Z_i, Y_i)}{\widehat{Cov}(Z_i, D_i)}$
 - General form follows similarly (sample analog), but is notation-heavy

Putting the “2S” in “2SLS”

- You'll notice I haven't actually defined 2SLS beyond the simple case
 - Before we had $\beta^{IV} = \frac{Cov(Z_i, Y_i)}{Cov(Z_i, D_i)}$ leading to $\hat{\beta}^{IV} = \frac{\widehat{Cov}(Z_i, Y_i)}{\widehat{Cov}(Z_i, D_i)}$
 - General form follows similarly (sample analog), but is notation-heavy
- A more useful way to define 2SLS is by a two-step procedure:
 - First regress D_i on all $Z_{i\ell}$ and W_{ik}
 - Then regress Y_i on the “fitted values” \hat{D}_i and controls W_{ik}

Putting the “2S” in “2SLS”

- You'll notice I haven't actually defined 2SLS beyond the simple case
 - Before we had $\beta^{IV} = \frac{Cov(Z_i, Y_i)}{Cov(Z_i, D_i)}$ leading to $\hat{\beta}^{IV} = \frac{\widehat{Cov}(Z_i, Y_i)}{\widehat{Cov}(Z_i, D_i)}$
 - General form follows similarly (sample analog), but is notation-heavy
- A more useful way to define 2SLS is by a two-step procedure:
 - First regress D_i on all $Z_{i\ell}$ and W_{ik}
 - Then regress Y_i on the “fitted values” \hat{D}_i and controls W_{ik}
- The proof of this follows from some (simple) linear algebra
 - Intuitively, regressing Y_i on $\hat{\pi}^{OLS} Z_i$ gives a scaled RF: $\hat{\beta}^{IV} = \frac{\hat{\rho}^{OLS}}{\hat{\pi}^{OLS}}$

Putting the “2S” in “2SLS”

- You'll notice I haven't actually defined 2SLS beyond the simple case
 - Before we had $\beta^{IV} = \frac{Cov(Z_i, Y_i)}{Cov(Z_i, D_i)}$ leading to $\hat{\beta}^{IV} = \frac{\widehat{Cov}(Z_i, Y_i)}{\widehat{Cov}(Z_i, D_i)}$
 - General form follows similarly (sample analog), but is notation-heavy
- A more useful way to define 2SLS is by a two-step procedure:
 - First regress D_i on all $Z_{i\ell}$ and W_{ik}
 - Then regress Y_i on the “fitted values” \hat{D}_i and controls W_{ik}
- The proof of this follows from some (simple) linear algebra
 - Intuitively, regressing Y_i on $\hat{\pi}^{OLS} Z_i$ gives a scaled RF: $\hat{\beta}^{IV} = \frac{\hat{\rho}^{OLS}}{\hat{\pi}^{OLS}}$
- Although easy, you should never do such “manual 2SLS” yourself!
 - Your point estimates will be right, but your SEs won't be!

Roadmap

Where do (Good) Instruments Come From?

- True Lotteries

- Natural Experiments

- Panel Data

2SLS Mechanics

- Just-Identified IV

- Overidentification

Weak and Many Instruments

- Weak IV

- Many IVs

Weak Instruments

- When running just-identified IV, you should always worry about the “strength” of your instrument
 - Specifically the first stage “F-statistic,” which tests $\pi^{OLS} = 0$

Weak Instruments

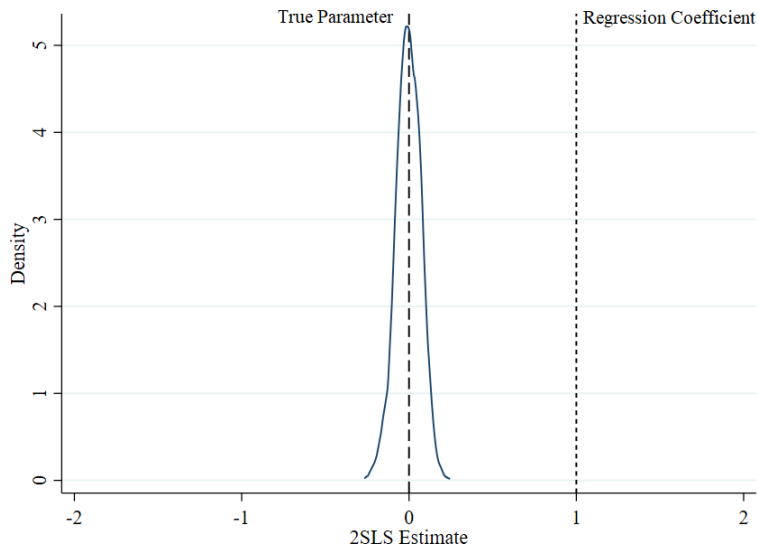
- When running just-identified IV, you should always worry about the “strength” of your instrument
 - Specifically the first stage “F-statistic,” which tests $\pi^{OLS} = 0$
- If π^{OLS} is small relative to its standard error, the IV is “weak”
 - Typically use the rule-of-thumb of $F < 10$ (Staiger and Stock 1997)
 - In this case the second-stage SEs will be large and the 2SLS estimate will tend to be biased towards the corresponding OLS estimate

Weak Instruments

- When running just-identified IV, you should always worry about the “strength” of your instrument
 - Specifically the first stage “F-statistic,” which tests $\pi^{OLS} = 0$
- If π^{OLS} is small relative to its standard error, the IV is “weak”
 - Typically use the rule-of-thumb of $F < 10$ (Staiger and Stock 1997)
 - In this case the second-stage SEs will be large and the 2SLS estimate will tend to be biased towards the corresponding OLS estimate
- Much made of this over the years, but Angrist and Kolesár (2022) argue recently that we shouldn’t worry too much
 - The SE increase tends to be large enough to “cover up” the bias
 - Just-id. 2SLS is “approximately median-unbiased” (as it is LIML)

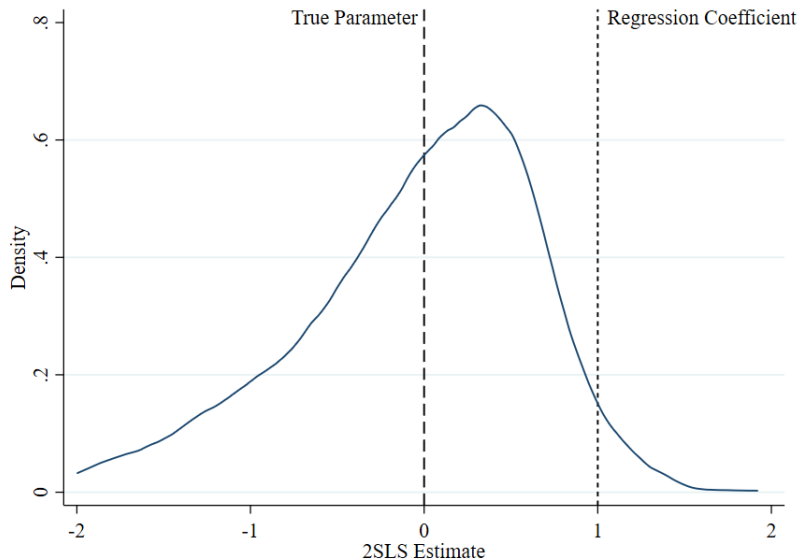
Weak Instruments: Visualized

Monte Carlo: $Y_i = \varepsilon_i$, $D_i = \Pi Z_i + \eta_i$: $\Pi = \text{Var}(\varepsilon_i) = \text{Var}(\eta_i) = 1$



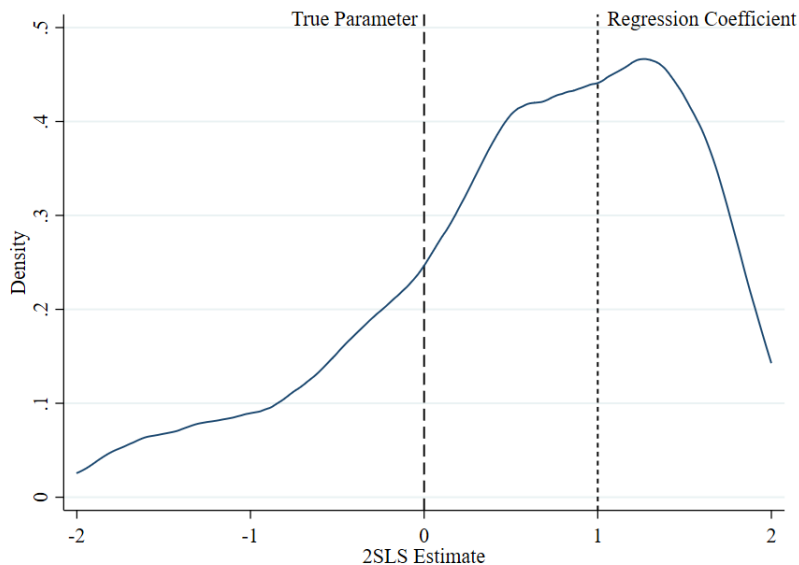
Weak Instruments: Visualized

Monte Carlo: $Y_i = \varepsilon_i$, $D_i = \Pi Z_i + \eta_i$: $\Pi = 0.1$ (Weaker)



Weak Instruments: Visualized

Monte Carlo: $Y_i = \varepsilon_i$, $D_i = \Pi Z_i + \eta_i$: $\Pi = 0.01$ (Very Weak)



Many IVs

- A more pernicious problem is many-instrument bias, when overid
 - Also tends to manifest in low first-stage F 's, so also good to check

Many IVs

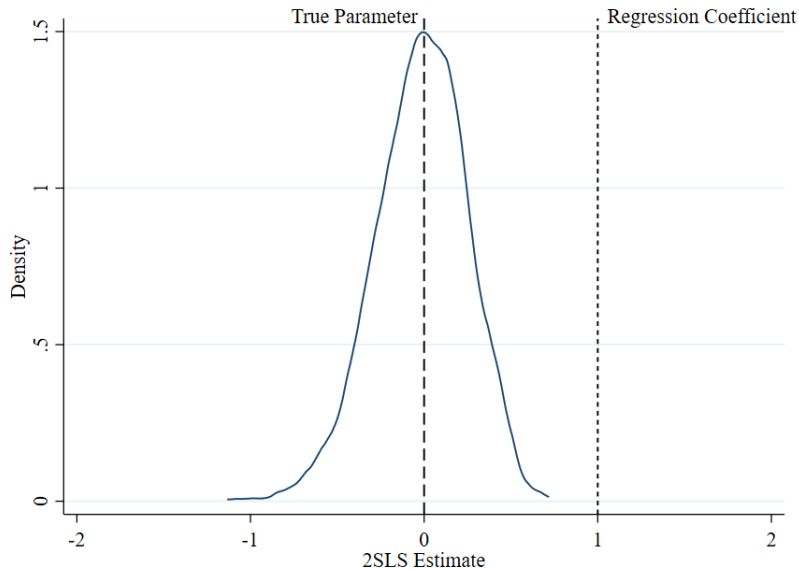
- A more pernicious problem is many-instrument bias, when overid
 - Also tends to manifest in low first-stage F 's, so also good to check
- Many-IV bias is also towards OLS, but unlike before SEs go *down*
 - Intuitively, a more flexible FS tends to fit D_i better → more power
 - But we can have *overfitting* with lots of Z_i → essentially recreate D_i

Many IVs

- A more pernicious problem is many-instrument bias, when overid
 - Also tends to manifest in low first-stage F 's, so also good to check
- Many-IV bias is also towards OLS, but unlike before SEs go *down*
 - Intuitively, a more flexible FS tends to fit D_i better → more power
 - But we can have *overfitting* with lots of Z_i → essentially recreate D_i
- As we'll see, this bias is especially relevant in judge IV designs
 - Potentially many judge assignment indicators as the instrument
 - Leave-out corrections (e.g. Angrist et al. 1999) have been adapted to this setting in recent years (e.g. Kolesár 2013)

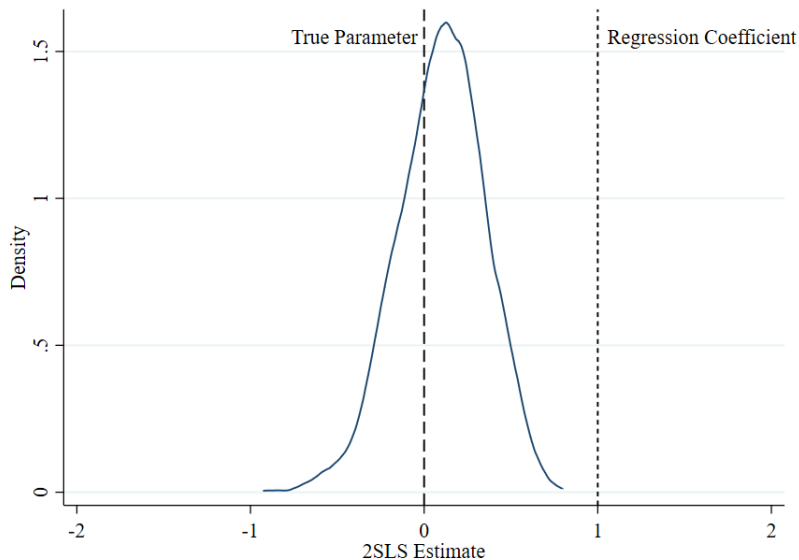
Weak and Many Instruments VIII

Monte Carlo: $Y_i = \varepsilon_i$, $X_i = \Pi Z_{i1} + \eta_i$: IV with one Z_{i1}



Weak and Many Instruments IX

Monte Carlo: $Y_i = \varepsilon_i$, $X_i = \Pi Z_{i1} + \eta_i$: IV with ten Z_{ij}



Weak and Many Instruments X

Monte Carlo: $Y_i = \varepsilon_i$, $X_i = \Pi Z_{i1} + \eta_i$: IV with 100 Z_{ij}

