Instrumental Variables

UNDERSTANDING IV



Roadmap

Where do (Good) Instruments Come From?

True Lotteries

Natural Experiments

Panel Data

2SLS Mechanics

Just-Identified IV

Overidentification

Weak and Many Instruments

Weak IV

Many IVs

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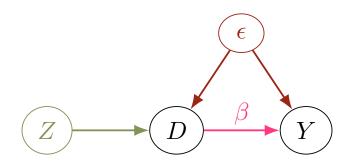
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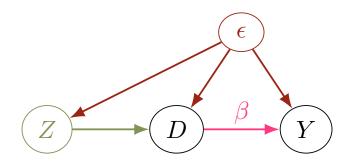
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 More modern IV texts take care to distinguish between these two conceptually distinct requirements...

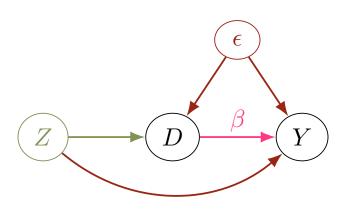
A Valid Instrument



A Violation of As-Good-As-Random Assignment



A Violation of Exclusion



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"Gold standard" IV: a randomized offer to participate in a program, with X_i recording program participation

- Exclusion restriction likely to hold for any Y_i , by construction
- Relevance almost guaranteed (provided people want the program!)

Charter School Lotteries

Abdulkadiroglu et al. (2016) are interested in whether going to a "charter" middle school increases standardized test scores

- Charter students tend to score better, but we worry about selection
- History of doubting educational inputs, since Coleman (1966)

Charter School Lotteries

We leverage an institutional feature of charters: admission lotteries

- When more kids want to enroll than there are seats, admission offers $Z_i \in \{0,1\}$ are effectively drawn from a hat
- Offers plausibly only affect later test scores Y_i by changing charter enrollment $D_i \in \{0,1\}$, so are plausibly valid instruments
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We study a particular charter (UP Academy), which is "takeover"

Two offer IVs: "immediate" (on lottery night) and from a waitlist

Lottery IV Estimates of UP Test Score Effects

TABLE 8—LOTTERY IV ESTIMATES OF UP EFFECTS

				2SLS			
				First stage			
		Comparison group mean (1)	OLS (2)	Immediate offer (3)	Waitlist offer (4)	Enrollment effect (5)	
Panel A. All grades (Sixth through eighth)	Math (N = 2,202)	0.059	0.301 (0.022)	0.760 (0.063)	0.562 (0.067)	0.270 (0.056)	
	ELA $(N = 2,205)$	0.103	0.148 (0.020)	0.759 (0.063)	0.562 (0.067)	0.118 (0.051)	

2. Natural Experiments

Without appealing to literal randomization, we may credibly argue Z_i is as-good-as-randomly assigned conditional on some \mathbf{W}_i

- Such "natural experiments" rely on a selection-on-observables argument (for Z_i , instead D_i)
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Quarter-of-birth

Angrist and Krueger (1991) famously estimate labor market returns to schooling with a creative IV: student quarter-of-birth

- Compulsory schooling requirements prevent students from dropping before the day they turn 16 (used to be more binding)
- Fixed school start dates mean students who drop out at 16 get more or less schooling depending on their birth date

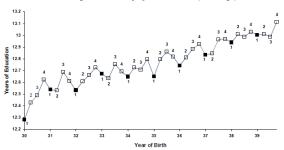
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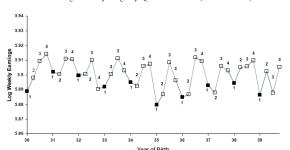
- Compulsory schooling requirements prevent students from dropping before the day they turn 16 (used to be more binding)
- Fixed school start dates mean students who drop out at 16 get more or less schooling depending on their birth date
- Quarter-of-birth seems quasi-randomly assigned is it excludable?
 See Buckles and Hungerman (2013)...

The Quarter-of-Birth Natural Experiment: Visualized

A. Average Education by Quarter of Birth (first stage)



B. Average Weekly Wage by Quarter of Birth (reduced form)



Quarter-of-Birth IV Estimates of Returns to Schooling

Table 4.1.1: 2SLS estimates of the economic returns to schooling

	OLS				2SLS	
	(1)	(2)	(3)	(4)	(5)	(6)
Years of education	0.075 (0.0004)	0.072 (0.0004)	0.103 (0.024)	0.112 (0.021)	0.106 (0.026)	0.108 (0.019)
Covariates:						
9 year of birth dummies 50 state of birth dummies		√			√	√
In struments:			dummy for QOB=1	dummy for QOB=1 or QOB=2	dummy for QOB=1	full set of QOB dummies

3. Panel Data

We might also combine IV + difference-in-differences identification

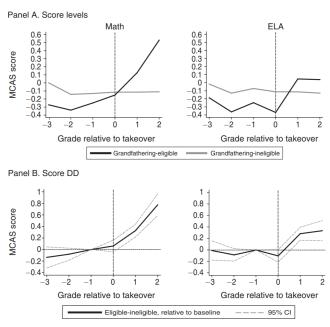
- ullet E.g. instrument with $Z_i imes Post_t$, controlling for Z_i and $Post_t$ FEs
- This requires two parallel trends assumptions, for the RF and FS
- Still need to worry about the exclusion restriction, as always

Example School Lottery

Abdulkadiroglu et al. (2016) complement their lottery analysis of takeover charters with an instrumented diff-in-diff analysis

- Students enrolled in the "legacy" public school were eligible for being "grandfathered" into UP, without having to apply to the charter
- We compare their trends in test scores & enrollment to a matched comparison group of observably-similar students at other schools

Grandfathering IV: Visualized



Grandfathering IV Estimates of UP Test Score Effects

TABLE 7—GRANDFATHERING IV ESTIMATES OF UP EFFECTS

				2SLS		
		Comparison group mean (1)	OLS (2)	First stage (3)	Enrollment effect (4)	
Panel A. All grades						
(Seventh through eighth)	Math $(N = 1,543)$	-0.233	0.400	1.051	0.321	
			(0.032)	(0.040)	(0.039)	
	ELA $(N = 1,539)$	-0.214	0.296	1.040	0.394	
	,		(0.035)	(0.041)	(0.044)	

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$$Y_i = \beta D_i + \mathbf{W}_i' \boldsymbol{\gamma} + \varepsilon_i$$
 (second stage)
 $X_i = \pi Z_i + \mathbf{W}_i' \boldsymbol{\mu} + \eta_i$ (first stage)

where \mathbf{W}_i includes a constant.

The reduced form is:

$$Y_i = \rho Z_i + \mathbf{W}_i' \kappa + \nu_i$$

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• Can use Frisch-Waugh-Lovell to "partial out" \mathbf{W}_i from Y_i , X_i , D_i , and so get back to an IV regression without controls

Sometimes we have more than one instrument $Z_{i\ell}$, for $\ell=1,\ldots,L$. This leads to an "overidentified" IV regression:

$$Y_i = \beta D_i + \mathbf{W}_i' \gamma + \varepsilon_i$$
 (second stage)

$$X_i = \mathbf{Z}_i' \boldsymbol{\pi} + \mathbf{W}_i' \boldsymbol{\mu} + \eta_i$$
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where
$$\mathbf{Z}_i = [Z_{i1}, \dots, Z_{iL}]'$$
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Overidentification can yield tests of IV validity

Intuitively, 2SLS checks whether all the $Z_{i\ell}$ yields the same IV estimate, which is sensible in a constant-effects model...

You'll notice I haven't actually defined 2SLS beyond the simple case

- Before we had $\beta^{IV}=\frac{Cov(Z_i,Y_i)}{Cov(Z_i,D_i)}$ leading to $\widehat{\beta}^{IV}=\frac{Cov(Z_i,Y_i)}{\widehat{Cov}(Z_i,D_i)}$
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A more useful way to define 2SLS is by a two-step procedure:

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The proof of this follows from some (simple) linear algebra

• Intuitively, regressing Y_i on $\widehat{\pi}^{OLS}Z_i$ gives a scaled RF:

$$\widehat{\beta}^{IV} = \frac{\widehat{\rho}^{OLS}}{\widehat{\pi}^{OLS}}$$

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Although easy, you should never do such "manual 2SLS" yourself!

Your point estimates will be right, but your SEs won't be!

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• Specifically the first stage $\ensuremath{\,\overline{ F}}$ -statistic , which tests $\pi^{OLS}=0$

If π^{OLS} is small relative to its standard error, the IV is "weak"

- Typically use the rule-of-thumb of F < 10 (Staiger and Stock 1997)
- In this case the second-stage SEs will be large and the 2SLS estimate will tend to be biased towards the corresponding OLS estimate

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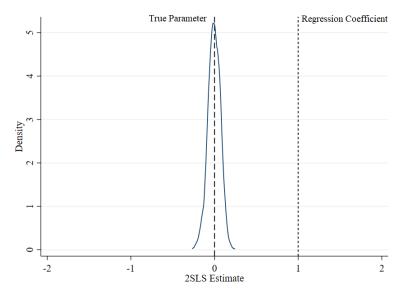
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Much made of this over the years, but Angrist and Kolesár (2022) argue recently that we shouldn't worry too much

- The SE increase tends to be large enough to "cover up" the bias
- Just-id. 2SLS is "approximately median-unbiased" (as it is LIML)

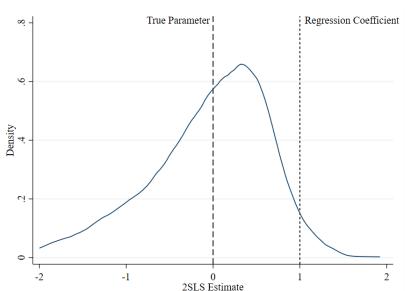
Weak Instruments: Visualized

Monte Carlo: $Y_i = \varepsilon_i$, $D_i = \Pi Z_i + \eta_i$: $\Pi = Var(\varepsilon_i) = Var(\eta_i) = 1$



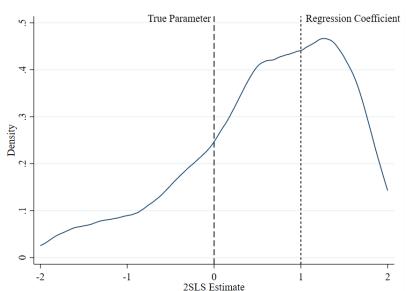
Weak Instruments: Visualized

Monte Carlo: $Y_i = \varepsilon_i$, $D_i = \Pi Z_i + \eta_i$: $\Pi = 0.1$ (Weaker)



Weak Instruments: Visualized

Monte Carlo: $Y_i = \varepsilon_i$, $D_i = \Pi Z_i + \eta_i$: $\Pi = 0.01$ (Very Weak)



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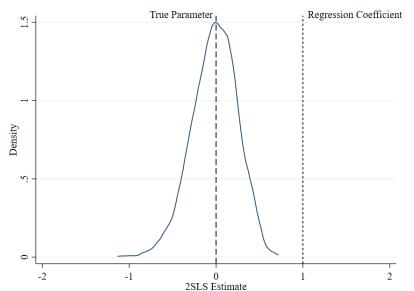
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As we'll see, this bias is especially relevant in judge IV designs

- Potentially many judge assignment indicators as the instrument
- Leave-out corrections (e.g. Angrist et al. 1999) have been adapted to this setting in recent years (e.g. Kolesár 2013)

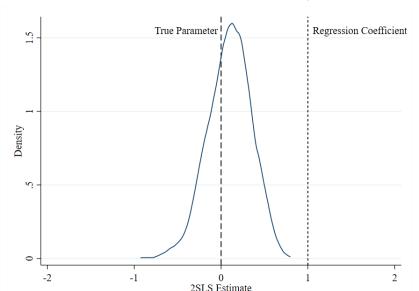
Weak and Many Instruments VIII

Monte Carlo: $Y_i = \varepsilon_i$, $X_i = \Pi Z_{i1} + \eta_i$: IV with one Z_{i1}



Weak and Many Instruments IX

Monte Carlo: $Y_i = \varepsilon_i$, $X_i = \Pi Z_{i1} + \eta_i$: IV with ten Z_{ij}



Weak and Many Instruments X

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