

## Problem 2.1

**a)**

We first compute the likelihood function:

$$\begin{aligned}\mathcal{L}(X|\theta) &= \prod_{i=1}^n f(x_i, \theta) \\ &= \prod_{i=1}^n \frac{x_i}{\theta^2} \exp\left(-\frac{x_i}{\theta}\right) \\ &= \frac{1}{\theta^{2n}} \exp\left(-\frac{1}{\theta} \sum_{i=1}^n x_i\right) \prod_{i=1}^n x_i\end{aligned}$$

Then compute the log likelihood:

$$\ln \mathcal{L}(X|\theta) = -2n \ln(\theta) - \frac{1}{\theta} \sum_{i=1}^n x_i + \sum_{i=1}^n \ln(x_i)$$

Then FOC:

$$\begin{aligned}\ln \mathcal{L}'(X|\theta) &= -\frac{2n}{\theta} + \frac{1}{\theta^2} \sum_{i=1}^n x_i = 0 \\ -2n + \frac{1}{\theta} \sum_{i=1}^n x_i &= 0 \\ \frac{1}{\theta} \sum_{i=1}^n x_i &= 2n \\ \frac{1}{2n} \sum_{i=1}^n x_i &= \theta\end{aligned}$$

**b)**

We already have been given the lower boundary:

$$\frac{1}{n\mathbb{E}\left[\left(\frac{\partial}{\partial\theta}l(X;\theta)\right)^2\right]}$$

So let's first derive the score:

$$\begin{aligned}
 \frac{\partial}{\partial \theta} l(x; \theta) &= \frac{\partial}{\partial \theta} \ln \left( \frac{x}{\theta^2} \exp \left( -\frac{x}{\theta} \right) \right) \\
 &= \frac{\partial}{\partial \theta} \ln x - 2 \ln \theta - \frac{x}{\theta} \\
 &= -\frac{2}{\theta} + \frac{x}{\theta^2}
 \end{aligned}$$

second derivative:

$$\frac{2}{\theta^2} - \frac{2x}{\theta^3}$$

And apply the mean:

$$\begin{aligned}
 -\mathbb{E} \left( \frac{2}{\theta^2} - \frac{2X}{\theta^3} \right) &= -\frac{2}{\theta^2} + \frac{2}{\theta^3} 2\theta \\
 &= -\frac{2}{\theta^2} + \frac{4}{\theta^2} \\
 &= \frac{2}{\theta^2}
 \end{aligned}$$

So we get the lower bound:

$$Var(T) \geq \frac{1}{\frac{2n}{\theta^2}} = \frac{\theta^2}{2n}$$

**c)**

Computing the variance of the maximum likelihood estimator:

$$\begin{aligned}
 Var(\hat{\theta}_{ML}) &= \frac{1}{4n^2} Var \left( \sum_{i=1}^n X_i \right) \\
 &= \frac{1}{4n^2} \sum_{i=1}^n Var(X_i) \\
 &= \frac{1}{4n^2} \sum_{i=1}^n 6\theta^2 - 4\theta^2 \\
 &= \frac{1}{4n^2} n 2\theta^2 \\
 &= \frac{\theta^2}{2n}
 \end{aligned}$$

So the lower bound is reached but is the estimator unbiased?

$$\begin{aligned}\mathbb{E}(\hat{\theta}_{ML}) &= \frac{1}{2n} \sum_{i=1}^n \mathbb{E}(X_i) \\ &= \frac{2n}{2n} \theta = \theta\end{aligned}$$

So yes and it is admissible because of bias variance decomposition.

**d)**

Bias variance decomposition:

$$\begin{aligned}\text{Var}(\hat{\theta}_2) &= \frac{1}{4(n+1)^2} \sum_{i=1}^n \text{Var}(X_i) \\ &= \frac{n6\theta^2}{4(n+1)^2}\end{aligned}$$

And the bias:

$$\begin{aligned}\mathbb{E}(\hat{\theta}_2) &= \frac{1}{2(n+1)} \sum_{i=1}^n \mathbb{E}(X_i) \\ &= \frac{n}{n+1} \theta \\ \text{Bias}(\hat{\theta}_2)^2 &= \frac{\theta^2}{(n+1)^2}\end{aligned}$$

Comming to the risk:

$$\begin{aligned}R(\hat{\theta}_2, \theta) &= \frac{6n\theta^2}{4(n+1)^2} + \frac{\theta^2}{(n+1)^2} \\ &= \frac{(6n-4)\theta^2}{4(n+1)^2} \leq \frac{\theta^2}{2n} \forall n > 0\end{aligned}$$

## Problem 2.2

**a)**

$$\begin{aligned}R(\delta_{a,b}, \mu) &= \mathbb{E}((\mu - a - bX)^2) \\ &= \text{Var}(\delta_{a,b}) + \text{Bias}(\delta_{a,b})^2 \\ &= b^2 + (a + b\mu - \mu)^2 \\ &= b^2 + (a + (b-1)\mu)^2\end{aligned}$$

**b)**

Inserting into the formula from a)

$$R(\delta_{a,b}, \mu) = 1$$

c)

No because the risk increases in  $b$

### Problem 2.3

a)

NP-lemma is applied and we look at the likelihood ratio:

$$\begin{aligned} LQ &= \frac{\mathcal{L}(\vec{X}|p_1)}{\mathcal{L}(\vec{X}|p_0)} > k \quad \forall x \in C \\ \frac{\prod_{i=1}^{10} p_1^{x_i} (1-p_1)^{1-x_i}}{\prod_{i=1}^{10} p_0^{x_i} (1-p_0)^{1-x_i}} &> k \quad \forall x \in C \\ \frac{\left( p_1^{\sum_{i=1}^{10} x_i} (1-p_1)^{10-\sum_{i=1}^{10} x_i} \right)}{\left( p_0^{\sum_{i=1}^{10} x_i} (1-p_0)^{10-\sum_{i=1}^{10} x_i} \right)} &> k \quad \forall x \in C \\ \left( \frac{p_1}{p_0} \right)^{\sum_{i=1}^{10} x_i} \left( \frac{1-p_1}{1-p_0} \right)^{10-\sum_{i=1}^{10} x_i} &> k \quad \forall x \in C \\ \left( \frac{p_1}{p_0} \frac{1-p_0}{1-p_1} \right)^{\sum_{i=1}^{10} x_i} &> k \left( \frac{1-p_0}{1-p_1} \right)^{10} \quad \forall x \in C \\ \sum_{i=1}^{10} x_i &< k^* \quad \forall x \in C \end{aligned}$$

Because  $\log_a(x)$  is a decreasing function for all  $a < 1$ , the inequality shifts in the other direction.

b)

The size is (in this case) the probability of making a type 1 error:

$$\begin{aligned} \alpha &= P(\vec{X} \in C) \\ &= P_{p_0} \left( \sum_{i=1}^{10} x_i \leq c_r - 1 \right) + P_{p_0} \left( \sum_{i=1}^{10} x_i = c_r \right) \gamma \end{aligned}$$

we can choose  $c_r = 6$ , then:

$$P \left( \sum_{i=1}^{10} x_i \leq c_r - 1 \right) = P \left( \sum_{i=1}^{10} x_i \leq 5 \right) = 0.150$$

And choose  $\gamma$  such that:

$$\begin{aligned}
 P\left(\sum_{i=1}^{10} x_i = 6\right) \gamma &= \alpha - P\left(\sum_{i=1}^{10} x_i \leq 5\right) \\
 0.2\gamma &= 0.2 - 0.15 \\
 \frac{1}{5}\gamma &= 5\% \\
 \implies \gamma &= 5 * 5\% = 25\%
 \end{aligned}$$

c)

Assuming 0-1 loss we can say:

$$\begin{aligned}
 R(\phi, p) &= \begin{cases} P_{p_0}(\vec{X} \in C) & p = p_0 \\ P_{p_1}(\vec{X} \notin C) & p = p_1 \end{cases} \\
 &= \begin{cases} P_{p_0}\left(\sum_{i=1}^{10} x_i \leq 6\right) + 0.56 \cdot P_{p_0}\left(\sum_{i=1}^{10} x_i = 7\right) & p = 0.7 \\ P_{p_1}\left(\sum_{i=1}^{10} x_i \geq 8\right) & p = 0.3 \end{cases} \\
 &= \begin{cases} \approx 0.5 & p = 0.7 \\ \approx 0 & p = 0.3 \end{cases}
 \end{aligned}$$

No this is not minimax, because the risk is not invariant of  $p$ .