

# Sheet 5

## Exercise 16

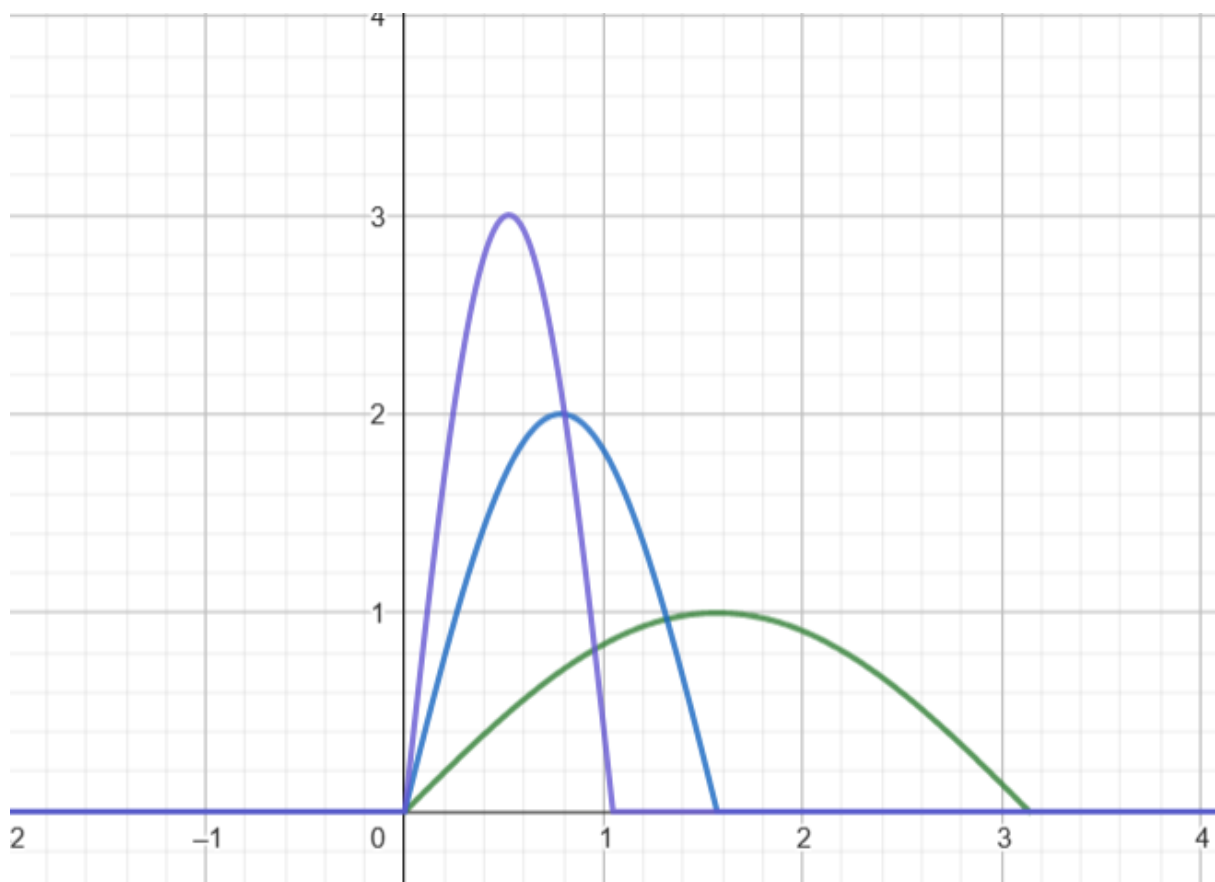
a)

In order to show that  $f_n$  is integrable we need to show two things:

1.  $f_n$  is measurable
2.  $\int |f_n| d\lambda < \infty$

Before starting the proof, a few properties of the sequence need to be established.

In essence the sequence  $f_n$  is compressing the first mode of the sinus function with every iteration. The following plot shows  $f_1$  (green),  $f_2$  (blue),  $f_3$  (purple):



In the first mode, the sinus function reaches a maximum value of 1 at  $\pi/2$ . Therefore we can say:

$$\max_{x \in [0, \pi]} f_n(x) = \max_{x \in [0, \pi/n]} n \sin(nx) = \max_{x \in [0, \pi]} n \sin(x) = n \max_{x \in [0, \pi]} \sin(x) = n(1)$$

Furthermore  $\sin$  is continuous and  $\sin(0) = \sin(\pi) = 0$ . Therefore  $f_n$  is **continuous** on its whole domain. Also  $f_n$  is a positive mapping. We can write:

$$f_n : [0, \pi] \rightarrow \mathbb{R} \Leftrightarrow f_n : [0, \pi] \rightarrow \mathbb{R}_+ \underset{(1)}{\Leftrightarrow} f_n : [0, \pi] \rightarrow [0, n]$$

**s.t.  $f_n$  is  $(\mathcal{B}([0, \pi]) - \mathcal{B}([0, n]))$ measurable**

We have already showed, that  $f_n$  is continuous, therefore  $f_n$  is measurable (Lecture slide 50).

$$\textbf{s.t. } \int |f_n| d\lambda < \infty$$

Using the same strategy from the last sheet, we will assume, that the Riemann and Lebesgue integral are equal. F

$$\int_{[0, \pi]} f_n d\lambda = \int_0^\pi n \sin(nx) 1_{[0, \pi/n]}(x) dx = \int_0^{\pi/n} n \sin(nx) dx$$

Now we can substitute  $t = nx$ :

$$\begin{aligned} \int_0^{\pi/n} n \sin(nx) dx &= \int_0^\pi \sin(t) dt \\ &= (-\cos(t)) \Big|_0^\pi \\ &= (-\cos(\pi) - (-\cos(0))) \\ &= -(-1) - (-1) = 2 < \infty \end{aligned}$$

And because we have established  $f_n$  as a positive mapping it we can argue:

$$\int |f_n| d\lambda = \int f_n d\lambda < \infty$$

**b)**

We will prove a more concrete statement:  $f_n \rightarrow 0$ .

Let any  $x \in (0, \pi]$ . Then we could argue:

$$\exists \hat{n} \in \mathbb{N} : x > \frac{\pi}{n} \quad \forall n \geq \hat{n}$$

But this implies that:

$$f_n(x) = 0 \quad \forall n \geq \hat{n}$$

So we can make the statement:

$$\forall \varepsilon > 0 \exists \hat{n} \in \mathbb{N} : |f_n(x) - 0| = 0 < \varepsilon \quad \forall n \geq \hat{n}$$

Furthermore it holds for all  $n \in \mathbb{N}$ :

$$f_n(0) = n \sin(0) = 0$$

And therefore:

$$f_n(0) \xrightarrow{n \rightarrow \infty} 0$$

Therefore showing that:

$$f_n(x) \xrightarrow{n \rightarrow \infty} 0 \quad \forall x \in [0, \pi]$$

**c)**

The Lemma of Fatou only requires positive mapping, which is given at every element of the sequence  $f_n$ . Therefore we can say that:

$$\int_0^\pi \liminf_{n \rightarrow \infty} f_n dx \leq \liminf_{n \rightarrow \infty} \int_0^\pi f_n dx$$

As we will see in **d)** it does not hold, that  $\mathbb{E}(\lim_{n \rightarrow \infty} f_n) \neq \lim_{n \rightarrow \infty} \mathbb{E}(f_n)$ . So dominated convergence should not hold. We could define:

$$Y(x) = \sup\{f_n(x) | n \in \mathbb{N}\}$$

And this should lead to:

$$\int_{[0, \pi]} Y d\lambda = \infty$$

How exactly this proof works, has been unclear to me.

**d)**

Let's first compute the LHS:

$$\begin{aligned}\lim_{n \rightarrow \infty} \int_0^\pi f_n d\mu &= \lim_{n \rightarrow \infty} \int_0^\pi f_n dx \\ &\stackrel{a)}{=} \lim_{n \rightarrow \infty} 2 = 2\end{aligned}$$

We have also established in **b)** that  $f_n \rightarrow 0$ . So

$$\int_0^\pi \lim_{n \rightarrow \infty} f_n d\mu = \int_0^\pi 0 d\mu = 0$$

So:

$$\lim_{n \rightarrow \infty} \int_{[0, \pi]} f_n d\mu \neq \int_{[0, \pi]} \lim_{n \rightarrow \infty} f_n d\mu$$

## Exercise 17

**a)**

Before calculating the measure integral we are going to prove:  $X_n \rightarrow 0$ .

Let  $x \in \mathbb{R}$ . Then it holds that for any  $\varepsilon > 0$ :

$$\exists \hat{n} \in \mathbb{N} : f_n(x) = \frac{1}{n} = \left| \frac{1}{n} - 0 \right| < \varepsilon \quad \forall n \geq \hat{n}$$

So  $f_n(x) \rightarrow 0 \quad \forall x \in \mathbb{R}$ .

Therefore the measure integral on the LHS is:

$$\begin{aligned}\mathbb{E}(\lim_{n \rightarrow \infty} X_n) &= \int_{\mathbb{R}} \lim_{n \rightarrow \infty} X_n d\lambda \\ &= \int_{\mathbb{R}} 0 d\lambda = 0\end{aligned}$$

But the integral on the RHS is:

$$\begin{aligned}
\lim_{n \rightarrow \infty} \mathbb{E}(X_n) &= \lim_{n \rightarrow \infty} \int_{\mathbb{R}} X_n d\lambda \\
&= \lim_{n \rightarrow \infty} \frac{1}{2n} \lambda([-n, n] \cap \mathbb{R}) \\
&= \lim_{n \rightarrow \infty} \frac{1}{2n} \lambda([-n, n]) \\
&= \lim_{n \rightarrow \infty} \frac{1}{2n} n - (-n) = \lim_{n \rightarrow \infty} \frac{2n}{2n} = 1
\end{aligned}$$

**b)**

Because  $\mathbb{E}(\lim_{n \rightarrow \infty} X_n) \neq \lim_{n \rightarrow \infty} \mathbb{E}(X_n)$  it is save to say that neither the monotone convergence or the dominated convergence theorem apply.

But for sake of completeness:

$X_n \leq X_{n+1}$  i.e.  $X_n(x) \leq X_{n+1}(x) \quad \forall x \in \mathbb{R}$  must hold for monotone convergence theorem must hold. But if we look at  $x = 0$  and  $X_1$  as well as  $X_2$  then:

$$X_1(0) = \frac{1}{2} > \frac{1}{4} = X_2(0)$$

This acts as a counterexample where monotonicity does not hold.

**c)**

We can define  $Y$  in a way, where  $Y(x) \geq X_n(x) \quad \forall x \in \mathbb{R}, n \in \mathbb{N}$ :

$$Y(x) = \sup\{X_n(x) | n \in \mathbb{N}\} \Leftrightarrow Y(x) = \sum_{n=1}^{\infty} \frac{1}{2n} 1_{[n-1, n] \cup [-n, -(n-1)]}(x)$$

Furthermore let:

$$s_n = \sum_{i=1}^n \frac{1}{2i} 1_{[i-1, i] \cup [-i, -(i-1)]}(x)$$

Then  $s_n$  is always simple and  $s_n \rightarrow Y$  and  $s_n \leq s_{n+1}$  (so monotone convergence applies):

$$\begin{aligned}
\int_{\mathbb{R}} Y d\lambda &= \int_{\mathbb{R}} \lim_{n \rightarrow \infty} s_n d\lambda \\
&= \lim_{n \rightarrow \infty} \int_{\mathbb{R}} s_n d\lambda \\
&= \lim_{n \rightarrow \infty} \sum_{i=1}^n \frac{1}{2i} \lambda([i-1, i] \cup [-i, -(i-1)]) \cap \mathbb{R}) \\
&= \lim_{f.a. \ n \rightarrow \infty} \sum_{i=1}^n \frac{1}{2i} (\lambda([i-1, i]) + \lambda([-i, -(i-1)])) \\
&= \lim_{n \rightarrow \infty} \sum_{i=1}^n \frac{1}{2i} (i - (i-1)) + (-(i-1) + i) \\
&= \lim_{n \rightarrow \infty} \sum_{i=1}^n \frac{1}{2i} * 2 = \lim_{n \rightarrow \infty} \sum_{i=1}^n \frac{1}{i} = \infty
\end{aligned}$$

## Exercise 18

a)

We have to show that:

$$X_n \xrightarrow{P} X \not\Rightarrow \mathbb{E}(\|X_n - X\|^2) \rightarrow 0$$

The **counter example** for this proof was inspired by **exercise 16 but strongly simplified**. We will define:

$$f_n := \begin{cases} \sqrt{n} & \text{if } 0 \leq x \leq \frac{1}{n} \\ 0 & \text{if } \frac{1}{n} < x \leq 1 \end{cases}$$

Then  $f_n : [0, 1] \rightarrow \mathbb{R}_+$  holds and we can consider the measuring space  $([0, 1], \mathcal{B}([0, 1]), \lambda)$  for the domain. And because  $\lambda([0, 1]) = 1$  holds, this measuring space is also a probability space.

Now let's quickly prove measurability:

Let  $[a, b] \subset \mathbb{R}_+$  with  $a, b \in \mathbb{R}_+$  then:

$$f_n^{-1}([a, b]) = \begin{cases} [0, 1/n] & \text{if } \{0, \sqrt{n}\} \cap [a, b] = \{\sqrt{n}\} \\ [1/n, 1] & \text{if } \{0, \sqrt{n}\} \cap [a, b] = \{0\} \\ [0, 1] & \text{if } \{0, \sqrt{n}\} \cap [a, b] = \{0, \sqrt{n}\} \\ \emptyset & \text{else} \end{cases}$$

So  $f_n^{-1}([a, b]) \in \mathcal{B}([0, 1])$  holds for every  $[a, b] \subset \mathbb{R}_+$ . And therefore  $f_n$  is measurable and a real r.v..

Now we can argue  $f_n \xrightarrow{P} 0$  because for any  $\varepsilon > 0$  it holds that:

$$P(|f_n - 0| > \varepsilon) = P\left(\left[0, \frac{1}{n}\right]\right) = \lambda\left(\left[0, \frac{1}{n}\right]\right) = \frac{1}{n} \xrightarrow{n \rightarrow \infty} 0$$

But trying to prove convergence in quadratic mean leads to:

$$\begin{aligned} \mathbb{E}(|f_n - 0|^2) &= \mathbb{E}(|f_n|^2) = \int_0^1 f_n^2 d\lambda = \int_0^1 f_n(x)^2 dx \\ &= \int_0^{\frac{1}{n}} n dx = \frac{n}{n} = 1 \end{aligned}$$

And so:

$$\lim_{n \rightarrow \infty} \mathbb{E}(|f_n - 0|^2) = \lim_{n \rightarrow \infty} 1 = 1 \neq 0$$

**b)**

According to the lecture we can say:

$$X_n \xrightarrow{P} b \implies X_n \xrightarrow{d} b$$

So let  $f(x)$  be a continuous and bounded function. Specifically let this function be:

$$f(x) = \begin{cases} x - b & \text{if } x \in [b - 1, b + 1] \\ b - 1 & \text{if } x < b - 1 \\ b + 1 & \text{if } x > b + 1 \end{cases}$$

Then convergence in distribution implies that:

$$\mathbb{E}(f(X_n)) \xrightarrow{n \rightarrow \infty} \mathbb{E}(f(b)) = \mathbb{E}(b - b) = 0$$

So:

$$\mathbb{E}(X_n - b) \xrightarrow[n \rightarrow \infty]{} 0 \Leftrightarrow \mathbb{E}(X_n) - b \xrightarrow[n \rightarrow \infty]{} 0$$

But since  $b$  is a constant we can say:

$$\mathbb{E}(X_n) \xrightarrow[n \rightarrow \infty]{} b$$

## Exercise 19

**a)**

We start by saying for any  $\varepsilon > 0$  it holds that:

$$P(|X_n - 1| > \varepsilon) = P(X_n \notin [1 - \varepsilon, 1 + \varepsilon]) = P(X_n < 1 - \varepsilon)$$

We can continue this equation with:

$$\begin{aligned} P(X_n < 1 - \varepsilon) &= P(\max_{0 \leq i \leq n} X_i < 1 - \varepsilon) \\ &\stackrel{iid}{=} P(X_1 < 1 - \varepsilon) \cdot \dots \cdot P(X_n < 1 - \varepsilon) \\ &= \lambda([0, 1 - \varepsilon)) \cdot \dots \cdot \lambda([0, 1 - \varepsilon)) \\ &= (1 - \varepsilon)^n \end{aligned}$$

Now we can look at behavior against infinity

$$\lim_{n \rightarrow \infty} P(|X_n - 1| > \varepsilon) = \lim_{n \rightarrow \infty} (1 - \varepsilon)^n = 0$$

So it is proven that:

$$X_n \xrightarrow{P} 1$$

**b)**

If  $X_n \sim \text{Exp}(\lambda_n)$  then  $P(X_n \leq x) = 1 - e^{-\lambda_n x}$ . Also:

$$\begin{aligned} P(|X_n - 0| > \varepsilon) &= P(X_n > \varepsilon) \\ &= 1 - P(X_n \leq \varepsilon) \\ &= 1 - 1 + e^{-\lambda_n \varepsilon} = e^{-\lambda_n \varepsilon} \end{aligned}$$



And so:

$$\lim_{n \rightarrow \infty} P(|X_n - 0| > \varepsilon) = \lim_{n \rightarrow \infty} e^{-\lambda_n x} \underset{\lambda_n \rightarrow \infty}{=} 0$$

c)

$$\begin{aligned} \lim_{n \rightarrow \infty} P(|X_n - 0| > \varepsilon) &= \lim_{n \rightarrow \infty} P(X_n > \varepsilon) \\ &= \lim_{\varepsilon > 0} \lim_{n \rightarrow \infty} P(X_n = n) \\ &= \lim_{n \rightarrow \infty} P(U \geq n) \\ &= \lim_{n \rightarrow \infty} 1 - P(U < n) \\ &= \lim_{n \rightarrow \infty} 1 - 1 + e^{-n} \\ &= \lim_{n \rightarrow \infty} e^{-n} = 0 \end{aligned}$$

d)

$$P(|X_n - X| > \varepsilon) = P(|Y_n + X - X| > \varepsilon) = P(|Y_n| > \varepsilon)$$

In this case we can use **Chebychevs Inequality**:

$$\begin{aligned} \lim_{n \rightarrow \infty} P(|Y_n| > \varepsilon) &\leq \lim_{n \rightarrow \infty} P(|Y_n| \geq \varepsilon) \\ &\underset{\mathbb{E}(Y_n) \rightarrow 0}{=} \lim_{n \rightarrow \infty} P(|Y_n - \mathbb{E}(Y_n)| \geq \varepsilon) \\ &\leq \lim_{c.i.} \lim_{n \rightarrow \infty} \frac{Var(Y_n)}{\varepsilon^2} \\ &= \lim_{n \rightarrow \infty} \frac{\sigma^2}{\varepsilon^2 n} = 0 \end{aligned}$$

And since  $P$  is a positive mapping, we can state:

$$P(|Y_n| > \varepsilon) \xrightarrow[n \rightarrow \infty]{} 0$$