## An Introduction to Empirical Likelihood

## **Textbooks**

- Greene, W.H., 2012, Econometric Analysis (7th ed.), Chapter 12.3.2.
- Cameron, A.C. and Trivedi, P.K., 2005, Microeconometrics: Methods and Applications, Chapter 6.8.
- Hansen, B.E., 2013, Econometrics, Chapter 14.
- Anatolyev, S. and Gospodinov, N., 2011, Methods for Estimation and Inference in Modern Econometrics, Chapter 2.

### Section 1

Introduction

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- Method of Moments vs. Maximum Likelihood.
- Many economic models can be written in the form

$$E[g(Z_i,\beta)]=0$$

where  $\beta \in \mathcal{B}$  is a k-dimensional parameter of interest, and  $g(\cdot, \cdot)$  is a  $\mathbb{R}^m$  valued function.  $k \leq m$ .

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# Examples I

#### Linear model with endogenous variables and instruments

- Structural equation  $Y_i = X_i \beta_0 + \epsilon_i$
- $X_i$  is an endogenous variable.
- W<sub>i</sub> is a vector of instruments.
- We estimate  $\beta_0$  via the moment condition

$$E\left[\epsilon_{i}W_{i}\right]=E\left[\left(y_{i}-X_{i}\beta\right)W_{i}\right]=0$$

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Empirical Likelihood

# Examples II

Asset pricing in Dynamic Rational Expectation Model.

- Hansen and Singleton (1982).
- The rational expectation theory implies

$$E_t\left[b\frac{U'(C_{t+j},\beta_0)}{U'(C_t,\beta_0)}X_{t+j}-1\right]=0$$

where b is a discount factor,  $U(\cdot)$  is a utility function known up to a parameter  $\beta$ , and  $X_{t+j} = (P_{t+j} + D_{t+j})/P_{t+j}$  is the interest rate.

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Empirical Likelihood

## **Outline**

- Review GMM and introduce Empirical Likelihood (EL)
- @ Generalize EL and develop its asymptotic theory
- Sextend EL into high-dimensional moments

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### Section 2

## **Empirical Likelihood**

## Generalized Method of Moments

- Let  $E_n[\cdot] = n^{-1} \sum_{i=1}^n \cdot$  be the empirical mean.
- The GMM estimator

$$\widehat{\beta}_{\mathrm{GMM}} = \arg\min_{\beta \in \mathcal{B}} E_n[g(Z_i,\beta)]' W \underline{E}_n[g(Z_i,\beta)],$$

#### where W is a positive-definite weighting matrix.

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## Likelihood Approach

- Consider a random sample from population distribution F(x) with density f(x).
- The joint likelihood function is

$$\prod_{i=1}^n f(x_i).$$

• The parametric method assumes that  $f(\cdot)$  is known up to a finite-dimensional parameter.

## Nonparametric Likelihood

- A nonparametric method assigns probability  $p_i$  to each observation  $x_i$  for i = 1, ..., n.
- It can be viewed as a sampling experiment from the multinomial population concentrated at the sample points.
- The Nonparametric Maximum Likelihood Estimator (NPMLE) solves

$$\max_{\mathbf{p}} \prod_{i=1}^{n} p_{i} \ \text{ s.t. } p_{i} \geq 0, \ \sum_{i=1}^{n} p_{i} = 1,$$

or equivalently

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### Solution

• To solve the constrained optimization problem,

$$\mathcal{L}(\mathbf{p}, \mu) = \frac{1}{n} \sum_{i=1}^{n} \log p_i - \mu \left( \sum_{i=1}^{n} p_i - 1 \right).$$

The first-order condition gives

$$\partial \mathcal{L}/\partial p_i = (np_i)^{-1} - \mu = 0,$$

for all i = 1, ..., n.

• As  $\sum_{i=1}^{n} p_i = 1$ , we have  $\mu = 1$ , so that

$$\widehat{p}_i = 1/n$$
.

This is exactly the Empirical Distribution Function.

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## **EL with Moment Constraints**

- Again, the model is  $E[g(Z_i, \beta)] = 0$ .
- The EL problem is

$$\max_{\beta, \mathbf{p}} \sum_{i=1}^{n} \log p_{i}$$
s.t. 
$$\sum_{i=1}^{n} p_{i} g(Z_{i}, \beta) = 0 \text{ and } \sum_{i=1}^{n} p_{i} = 1.$$

• If k = m, then

$$\widehat{\beta} = \arg_{\beta} \left( \frac{1}{n} \sum_{i=1}^{n} g(Z_i, \beta) = 0 \right),$$

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### Solution when k < m

The Lagrangian

$$\mathcal{L}(\beta, \mathbf{p}, \lambda, \mu) = \frac{1}{n} \sum_{i=1}^{n} \log p_i - \lambda' \sum_{i=1}^{n} p_i g(Z_i, \beta) - \mu \left( \sum_{i=1}^{n} p_i - 1 \right).$$

• The first order condition with respect to  $p_i$ ,

$$(np_i)^{-1} - \lambda' g(Z_i, \beta) - \mu = 0.$$
 (1)

Multiple both sides by p<sub>i</sub> and sum over i,

$$1 - \lambda' \sum_{i=1}^{n} p_i g(Z_i, \beta) - \mu \sum_{i=1}^{n} p_i = 1 - \mu = 0.$$



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## Solution when k < m (Continue)

- Plug in  $\mu = 1$  into (1),  $p_i = [n(1 + \lambda' g(Z_i, \beta))]^{-1}$ , which is slightly different from 1/n.
- Substitute the expression of  $p_i$  into the criterion function, we obtain the EL estimator  $\widehat{\beta}$  and multiplier  $\widehat{\lambda}$  via the saddlepoint problem

$$\max_{\beta \in \mathcal{B}} \min_{\lambda} - \sum_{i=1}^{n} \log (1 + \lambda' g(Z_i, \beta)).$$

## Computation

• The inner loop and the outer loop

$$\max_{\beta} \left\{ \min_{\lambda(\beta)} \left[ -\sum_{i=1}^{n} \log \left( 1 + \lambda(\beta)' g(Z_i, \beta) \right) \right] \right\}.$$

- The inner loop is globally convex in  $\lambda$
- The outer loop is neither convex nor concave in general in  $\beta$ .

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Empirical Likelihood

## Differences from GMM

- GMM: Two-step, iterative, continuous-updating (Hansen, Heaton and Yaron, 1996). Basic idea: stick to  $p_i = 1/n$ .
- Altonji and Segal (1996): large bias in two-step GMM
- EL: more flexible **p**. One-step estimator. Self-normalize.
- High-order improvement
  - Kitamura(2001): Asymptotic optimality under the generalized Neyman-Pearson criterion
  - ▶ Newey and Smith (2004): High-order bias correction

## Example (Imbens, 1997)

- The logarithm of hourly wage of 827 men in 1971–1978.
- The model is borrowed from Card (1994)

$$\ln y_{it} = \mu_t + \omega_i + u_{it} + \varepsilon_{it} 
u_{it} = \alpha u_{it-1} + \eta_{it}$$

where  $\mu_t$  is a common time-varying component,  $\omega_i$  is the individual fixed effect,  $\eta_{it}$  is the shock to the autoregressive component.

- Unknown parameters:  $\mu_1, \ldots, \mu_T$ ;  $\sigma^2_{\eta,1}, \ldots, \sigma^2_{\eta,T}$ ;  $\sigma^2_{\varepsilon}$ ;  $\sigma^2_{\omega}$  and  $\alpha$ .
- Imbens (1997) constructs 44 moments to estimate the 19 parameters. He compares two-step GMM, iterated GMM and EL.



## Example (Imbens, 1997)

#### Real data estimates

	2S-GMM	GMM	EL	s.e.
$\sigma_{\omega}^{2}$	0.110	0.111	0.123	0.053
$\sigma_{\omega}^{2}$	0.040	0.039	0.041	0.003
$\alpha$	0.913	0.912	0.899	0.057

• Simulation: true value and bias divided by the average of s.e.

	True value	2S-GMM	GMMi	EL
$\sigma_{\omega}^2$	0.10	-0.28	-0.28	-0.16
$\sigma_{\varepsilon}^2$	0.05	-0.31	-0.30	-0.22
α	0.50	0.05	0.06	0.02

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### Literature

- Owen (1988): Original paper on EL
- Qin and Lawless (1993): EL with estimating equations
- Imbens (1997): Introduce EL into econometrics
- Kitamura (1997): EL in time series
- Newey and Smith (2004): Generalized EL

### Section 3

Generalized Empirical Likelihood

## Generalized Empirical Likelihood

• Let  $\rho(\cdot)$  be a smooth scalar function that satisfies

$$\rho(0) = 0$$
 and  $\partial \rho(0)/\partial v = \partial^2 \rho(0)/\partial v^2 = -1$ .

GEL estimator solves

$$\min_{\beta} \sup_{\lambda} \sum_{i=1}^{n} \rho \left( \lambda' g(Z_{i}, \beta) \right).$$

- The Generalized Empirical Likelihood (GEL) contains several important estimators as special cases.
  - ▶ EL:  $\rho(v) = \log(1 v)$ ,
  - continuous-updating:  $\rho(v) = -\frac{1}{2}v^2 v$ ,
  - ▶ Exponential Tilting (Kitamura and Stutzer, 1997):  $\rho(v) = 1 \exp(v)$ .



Empirical Likelihood

# **Asymptotic Properties**

- Let  $J \equiv E \left[ \partial g(Z, \beta_0) / \partial \beta' \right]$  and  $V \equiv E \left[ g(Z, \beta_0) g(Z, \beta_0)' \right]$ .
- With random sampling,

$$\sqrt{n}\left(\widehat{\beta}-\beta_0\right)\Rightarrow N\left(0,\left(J'V^{-1}J\right)^{-1}\right).$$

• First-order asymptotically equivalent to the efficient GMM.

# Asymptotic Tests

- Wald test, LM test and overidentification test.
- Moreover, it inherits a *likelihood ratio*-type test. Let  $(\widehat{\beta}, \widehat{\lambda})$  and  $(\widetilde{\beta}, \widetilde{\lambda})$  be the estimates without and with constraints, respectively. Then

$$GELR = 2\sum_{i=1}^{n} \left[ \rho \left( \tilde{\lambda}' g(Z_i, \tilde{\beta}) \right) - \rho \left( \hat{\lambda}' g(Z_i, \hat{\beta}) \right) \right].$$

Under the null,

$$GELR \Rightarrow \chi_q^2$$

where q is the number of restrictions.

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### Example

- Overidentification Test for EL.  $\rho(v) = \log(1 v)$ .
- $H_0$ : Correct model specification.  $E[g(Z, \beta_0)] = 0$ .
  - ► Likelihood without restriction:  $-n \log n$
  - Likelihood with restriction:  $-n \log n \sum_{i=1}^{n} \log \left(1 + \widehat{\lambda}' g\left(Z_i, \widehat{\beta}\right)\right)$ .
- Test statistic

$$ELR = 2\sum_{i=1}^{n} \log \left(1 + \widehat{\lambda}' g\left(Z_{i}, \widehat{\beta}\right)\right) \Rightarrow \chi_{m}^{2}.$$



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#### Section 4

#### Relaxed Empirical Likelihood

# **High-Dimensional Moments**

- Large data, large model vs. parsimonious modeling.
  - ► Altonji, Smith and Vidangos (2013)
  - ► Eaton, Kortum and Kramarz (2011)
  - ► Han, Orea, Schmidt (2005)
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Problem of optimally-weighted GMM.

$$\mathbb{E}_{n}\left[g_{i}\left(\beta\right)\right]'$$
  $\mathbb{W}_{GMM}$   $\mathbb{E}_{n}\left[g_{i}\left(\beta\right)\right]$ 

• Problem of EL. For any  $\beta \in \mathcal{B}$ , the constraints

$$\sum_{i=1}^{n} p_i = 1 \text{ and } \sum_{i=1}^{n} p_i g_i(\beta) = 0_m$$

• Problem of optimally-weighted GMM.

$$W_{\text{GMM}} = \widehat{V}^{-1}$$
  
 $\widehat{V}: m \times m$ , rank  $n$ 

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$$m \text{ equations}$$

$$n \text{ free parameters}$$

#### Extension: Relaxed EL

- Shi (2013) proposes the first asymptotically normal estimator that allows m > n.
- Relax equality constraints of the standard EL problem by inequality constraints

$$\max_{\beta,\mathbf{p}} \sum_{i=1}^{n} \log p_{i}$$
s.t. 
$$\sum_{i=1}^{n} p_{i} = 1,$$

$$\left| \sum_{i=1}^{n} p_{i} h_{j}(Z_{i}, \beta) \right| \leq \tau, \forall j = 1, \dots, n.$$

where  $\tau$  is a tuning parameter, and  $h_j = g_j/\widehat{\sigma}_j$  is the scale-standardized function.

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### Summary

- What is EL and why
- GEL and the asymptotic theory
- REL for high-dimensional moments