

Statistics 452: Statistical Learning and Prediction

Chapter 4, Part 2: Linear Discriminant Analysis Classifier

Brad McNeney

2017-09-28

Linear Discriminant Analysis (LDA) Classifier

Logistic Regression vs LDA Classifier

- ▶ Logistic regression models $Pr(Y = 1|X = x) = p(x)$.
 - ▶ More generally, polytomous regression models $Pr(Y = k|X = x)$.
- ▶ LDA models $Pr(X = x|Y = k)$, which relates to $Pr(Y = k|X = x)$ via Bayes' rule:

$$Pr(Y = k|X = x) = \frac{Pr(X = x|Y = k)Pr(Y = k)}{\sum_j Pr(X = x|Y = j)Pr(Y = j)}.$$

Notation and Terminology

- ▶ Notation: $f_k(X) = Pr(X = x|Y = k)$, $\pi_k = Pr(Y = k)$, $p_k(x) = Pr(X = x|Y = k)$, so that

$$p_k(x) = \frac{f_k(x)\pi_k}{\sum_j f_j(x)\pi_j}.$$

- ▶ Terminology:
 - ▶ π_k is the *prior* probability of class k
 - ▶ $p_k(x)$ is the *posterior* probability of class k given data x .

Bayes Classifier

- ▶ We saw in Chapter 2 that the classifier with the lowest error rate is the one that chooses the class with the highest posterior probability.
- ▶ If we have good estimates, $\hat{p}_k(x)$, we can choose the class with the highest **estimated** posterior probability.
- ▶ Estimate $p_k(x)$ by estimating
 - ▶ π_k : If training sample is a population random sample, use the sample proportions.
 - ▶ $f_k(x)$: LDA model is parametric, so we estimate f_k by estimating its parameters.

LDA Classifier: Model for $f_k(x)$ when $p = 1$

- ▶ Assume $f_k(x)$ is normal with mean μ_k and variance σ_k^2 .
 - ▶ Then estimating f_k amounts to estimating μ_k and σ_k^2 .
 - ▶ Further assume σ_k is the same for all k and let σ denote this common value.
- ▶ For a particular x , the class with highest posterior probability is the one with highest value of

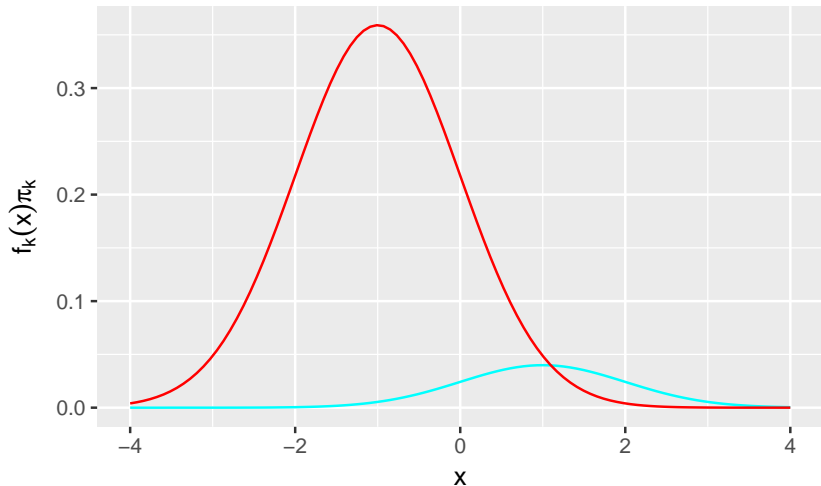
$$\delta_k(x) = x \frac{\mu_k}{\sigma^2} - \frac{\mu_k^2}{2\sigma^2} + \log \pi_k.$$

Example

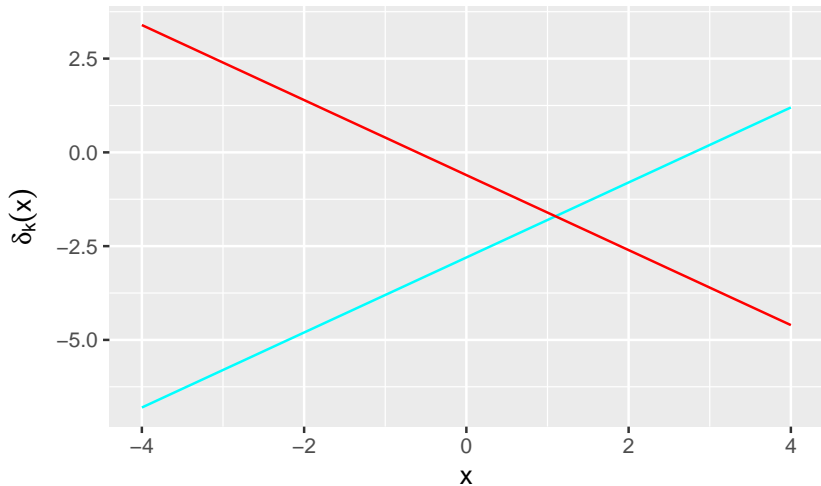
- Suppose two classes, $\pi_1 = 0.1$, $\pi_2 = 0.9$, $\mu_1 = 1$, $\mu_2 = -1$, $\sigma = 1$.

```
mu1<-1; mu2<-(-1); sigma<-1; pi1<-0.1;pi2<-0.9
x <-seq(from=-4,to=4,length=100)
f1 <- dnorm(x,mean=mu1,sd=sigma)
f2 <- dnorm(x,mean=mu2,sd=sigma)
delta <- function(x,mu,sigma,pi) {
  x*mu/sigma^2 - mu^2/(2*sigma^2) + log(pi)
}
delta1 <- delta(x,mu1,sigma,pi1)
delta2 <- delta(x,mu2,sigma,pi2)
dd <- data.frame(x=x,f1=f1,f2=f2,delta1=delta1,delta2=delta2)
```

```
library(tidyverse)
ggplot(dd,aes(x=x))+
  geom_line(aes(y=f1*pi1),color="cyan")+
  geom_line(aes(y=f2*pi2),color="red") +
  labs(y=expression(f[k](x)*pi[k]))
```




```
ggplot(dd,aes(x=x))+  
  geom_line(aes(y=delta1),color="cyan")+  
  geom_line(aes(y=delta2),color="red") +  
  labs(y=expression(delta[k](x)))
```



Decision Boundary

- ▶ The point where the δ_k lines cross is the decision boundary.
- ▶ Can show that the boundary is

$$\frac{\mu_1 + \mu_2}{2} + \frac{\sigma^2 \log(\pi_2/\pi_1)}{\mu_1 - \mu_2}.$$

- ▶ In this example the boundary is
 $(1 + (-1))/2 + \log(.9/.1)/2 \approx 1.1$

```
(mu1 + mu2)/2 + sigma^2*log(pi2/pi1)/(mu1-mu2)
```

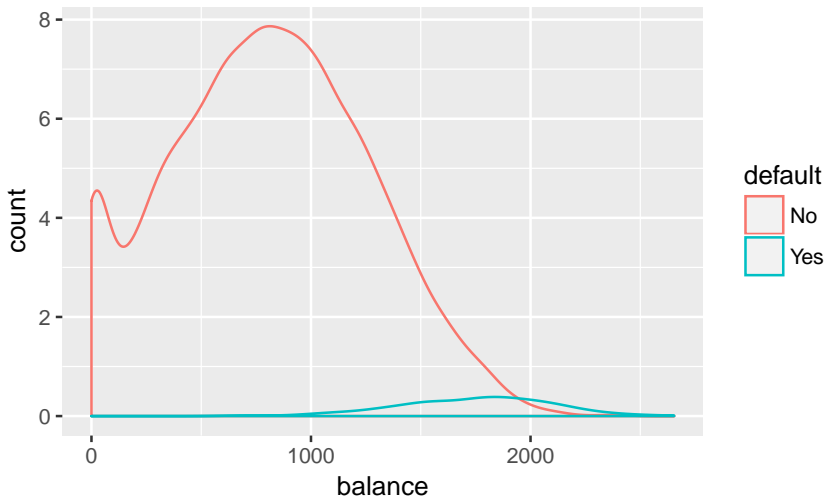
```
## [1] 1.098612
```

Estimated Distributions

- ▶ Replace means and variance in $\delta_k(x)$'s with estimates (equation 4.15 of text) to get the *discriminant functions* $\hat{\delta}_k$.
- ▶ Illustrate with Default data.

```
library(ISLR)  
data(Default)
```

```
ggplot(Default,aes(x=balance,color=default)) +  
  geom_density(aes(y=..count..))
```



Software Note

- ▶ Specifying `color=default` groups the data by default.
- ▶ The density geom plots the estimated densities $\hat{f}_k(x)$ by default, but also computes the variable `..count..`, which for group k is $\hat{f}_k(x)n_k$.
- ▶ We are interested in $\hat{f}_k(x)\hat{\pi} = \hat{f}_k(x)n_k/n$, so `count` is proportional to what we want to plot.

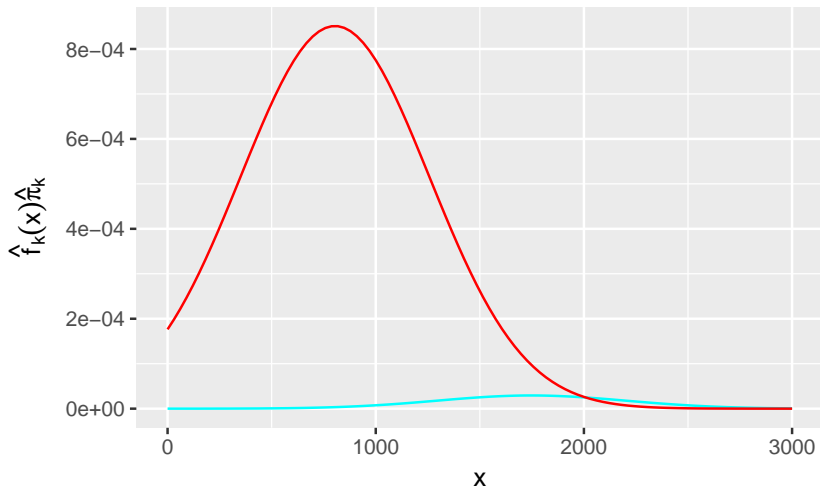
Estimated Densities and Discriminant Functions

```
n <- 10000; K <- 2
pi1 <- with(Default,mean(default=="Yes")); pi2 <- 1-pi1
defBalance <- with(Default,balance[default=="Yes"])
nodefBalance <- with(Default,balance[default=="No"])
mu1 <- mean(defBalance)
mu2 <- mean(nodefBalance)
sigma <- sqrt((sum((defBalance-mu1)^2) + sum((nodefBalance-mu2)^2))/(n-K))
#x <- seq(from=min(Default$balance),to=max(Default$balance),length=100)
with(Default,range(balance))
```

```
## [1] 0.000 2654.323
```

```
x <- seq(from=0,to=3000,length=100)
f1 <- dnorm(x,mean=mu1,sd=sigma)
f2 <- dnorm(x,mean=mu2,sd=sigma)
delta1 <- delta(x,mu1,sigma,pi1)
delta2 <- delta(x,mu2,sigma,pi2)
dd <- data.frame(x=x,f1=f1,f2=f2,delta1=delta1,delta2=delta2)
```

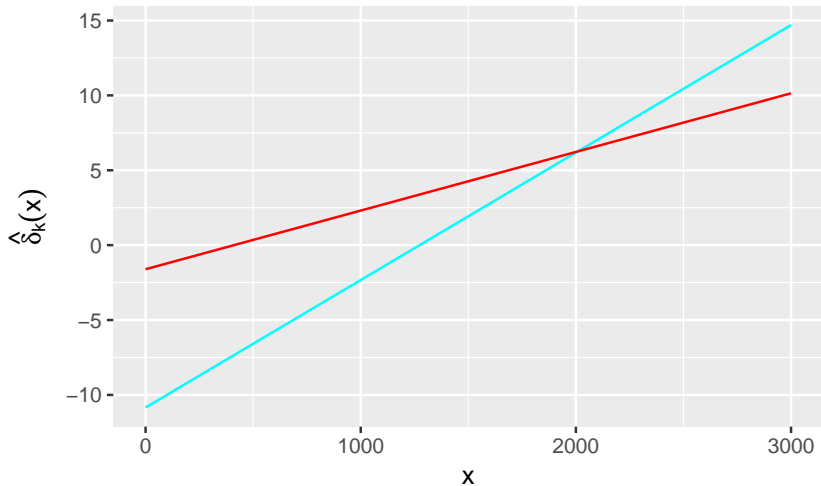
```
ggplot(dd,aes(x=x))+ geom_line(aes(y=f1*pi1),color="cyan")+
  geom_line(aes(y=f2*pi2),color="red") +
  labs(y=expression(hat(f)[k](x)*hat(pi)[k]))
```



```
min(x[f1*pi1>f2*pi2])
```

```
## [1] 2030.303
```

```
ggplot(dd,aes(x=x))+  
  geom_line(aes(y=delta1),color="cyan")+  
  geom_line(aes(y=delta2),color="red") +  
  labs(y=expression(hat(delta)[k](x)))
```



Decision Boundary for Default Data

```
(mu1 + mu2)/2 + sigma^2*log(pi2/pi1)/(mu1-mu2)
```

```
## [1] 2008.584
```

LDA Classifier in R

```
library(MASS)
ll <- lda(default ~ balance, data=Default)
preds <- predict(ll)
head(preds$posterior)
```

```
##           No           Yes
## 1 0.9972130 0.002786981
## 2 0.9958358 0.004164240
## 3 0.9865931 0.013406929
## 4 0.9988882 0.001111757
## 5 0.9963955 0.003604464
## 6 0.9933487 0.006651334
```

```
head(preds$class)
```

```
## [1] No No No No No No
## Levels: No Yes
```

LDA Classifier for $p > 1$

- ▶ Same basic idea of classifying based on highest posterior probabilities.
- ▶ Generalize the model for $f_k(X)$ from a Gaussian distribution to a *multivariate* Gaussian distribution.
- ▶ The mean in class k is now a vector μ_k and the variance is a matrix Σ .
- ▶ Details on the multivariate density and the resulting $\delta_k(x)$ are given on page 143 of the text.

LDA Classifier Using all Default Data

```
ll <- lda(default ~ student + balance + income, data=Default)
preds <- predict(ll)
head(preds$posterior)
```

| ## | | No | Yes |
|------|-----------|-------------|-----|
| ## 1 | 0.9967765 | 0.003223517 | |
| ## 2 | 0.9973105 | 0.002689531 | |
| ## 3 | 0.9852914 | 0.014708600 | |
| ## 4 | 0.9988157 | 0.001184329 | |
| ## 5 | 0.9959768 | 0.004023242 | |
| ## 6 | 0.9957918 | 0.004208244 | |

Predictions vs True Default Status

- ▶ Tabulating predictions and true default status gives the following “confusion matrix”:

```
Default <- data.frame(Default, predicted = preds$class)
xtabs( ~ predicted + default, data=Default)
```

```
##           default
## predicted    No   Yes
##           No  9645  254
##           Yes   22   79
```

- ▶ (Slight difference with results in text.)
- ▶ 9724 correctly predicted, 276 incorrectly predicted, so error rate is 2.76%.
- ▶ But 254/333, or 76% of defaulters missed.

Classification Error Rates

- ▶ Terminology is from diagnostic testing in medicine.
 - ▶ E.G., pap smear to screen for cervical cancer
- ▶ The test can be positive (T^+) or negative (T^-), and an individual may have the disease (D^+) or not (D^-):

| | | Disease Status | |
|-------------|-------|------------------------|------------------------|
| | | D^- | D^+ |
| Test Result | T^- | True Negative (TN) | False Negative (FN) |
| | T^+ | False Positive (FP) | True Positive (TP) |
| | | TN+FP | FN+TP |

- ▶ Replace test with prediction and true disease classification with true classification.

Sensitivity, Specificity, True- and False-Positive Rates

- ▶ The observed true positive rate (TPR) is the proportion of D^+ who are TP, or $TP/(TP+FN)$.
 - ▶ $TP/(TP+FN)$ is an estimate of $P(T^+ | D^+)$, known as the *sensitivity* of the test.
 - ▶ Sensitivity is only 79/333 or 23.7% for Default data.
- ▶ The observed true negative rate is the proportion of D^- who are TN, or $TN/(TN+FP)$.
 - ▶ $TN/(TN+FP)$ is an estimate of the true negative rate $P(T^- | D^-)$, known as the *specificity* of the test. Specificity is 9645/9667 or 99.8% for Default data.
 - ▶ The complement, $FP/(TN+FP)$, is an estimate of the false positive rate (FPR) $P(T^+ | D^-)$. FPR is 22/9667 or 0.2% for Default data.

Receiver Operating Characteristic (ROC) Curves

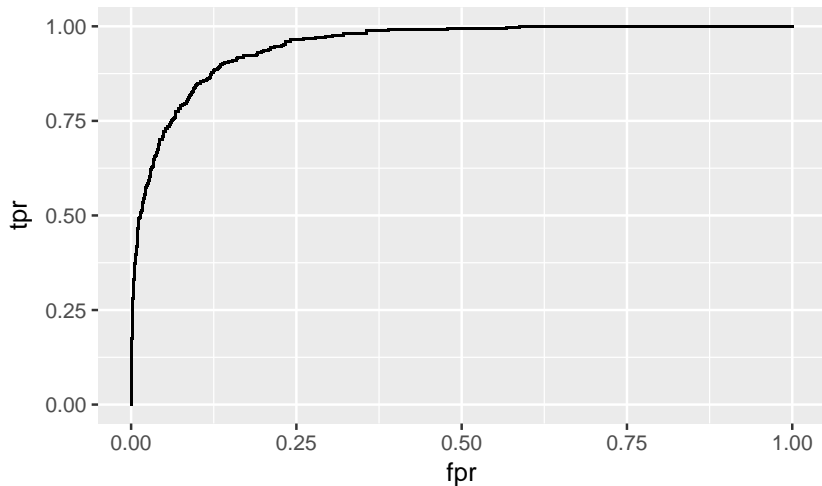
- ▶ The classification of a subject with $X = x_0$ as T^+ or T^- is made by comparing the posterior probability $Pr(T^+|X = x_0)$ to a threshold $t = 1/2$; e.g., $> t$ classify as T^+ , $\leq t$ classify as T^- .
- ▶ Lowering the threshold will increase the number of T^+ , and hence the TPR, $P(T^+ | D^+)$, and the FPR, $P(T^+ | D^-)$.
- ▶ The ROC curve is a plot plot of the TPR *versus* FPR as we vary t .
 - ▶ Try to get a sense of the trade-off between true- and false-positive rates, and possibly choose an “optimal” t .

ROC for LDA Classifier on Default Example

```
library(AUC) # for the roc() function; install.packages("AUC") to install
library(broom) # for the tidy() function
posteriorYes <- preds$posterior[, "Yes"]
trueYes <- # require a binary factor with 1=Yes
  (Default$default=="Yes") %>% as.numeric() %>% factor()
ROCres <- roc(posteriorYes, trueYes)
tidyROCres <- tidy(ROCres)
```

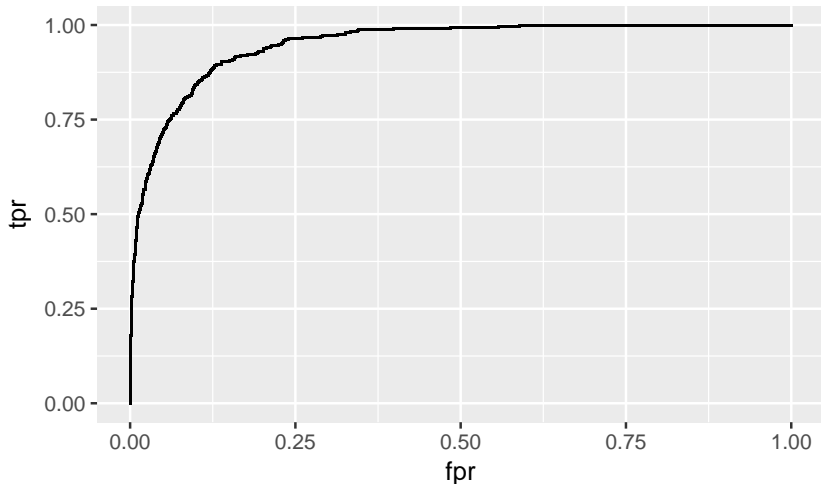
ROC Curve

```
ggplot(tidyROCres, aes(x=fpr, y=tpr)) + geom_point(pch=".")
```



ROC Curve for Logistic Regression

```
lfit <- glm(default ~ ., data=Default, family=binomial())  
posteriorYesLogistic <- fitted(lfit)  
ROClogist <- tidy(roc(posteriorYesLogistic, trueYes))  
ggplot(ROClogist, aes(x=fpr, y=tpr)) + geom_point(pch=".")
```



ROC Interpretation

- ▶ The points on the plot represent the FPR/TPR combinations for each threshold.
- ▶ The ideal test threshold would yield a TPR of one and a FPR of zero, and so would appear in the upper-left corner of the ROC plot.
 - ▶ We usually select the threshold that is closest to the upper-left corner.
 - ▶ No obvious “best” threshold in this example. Perhaps the threshold that gives TPR of about 0.875 and FPR of about 0.125.

```
tidyROCres %>%  
  filter(fpr >= 0.1249, fpr <= 0.1251)
```

| ## | instance | cutoff | fpr | tpr |
|------|----------|------------|-----------|-----------|
| ## 1 | 7644 | 0.04572886 | 0.1249612 | 0.8768769 |
| ## 2 | 8518 | 0.04572196 | 0.1250647 | 0.8768769 |
| ## 3 | 8490 | 0.04569361 | 0.1250647 | 0.8798799 |

Re-classify

```
n <- nrow(Default); thresh <- 0.0457
dclass <- rep("No",n); dclass[posteriorYes>thresh] <- "Yes"
Default <- data.frame(Default,prednew = dclass)
xtabs(~ prednew + default, data=Default)
```

```
##           default
## prednew    No   Yes
##      No  8458   41
##      Yes 1209  292
```

- ▶ Sensitivity is much better (about 87.7%).
- ▶ Many more false-positives (1209)

Area Under Curve (AUC)

- ▶ The area under the ROC curve is a measure of the overall performance of the classifier.
- ▶ If TPR jumps to near one with FPR remaining low, the performance is good.
 - ▶ This would give AUC near 1.
- ▶ A poor classifier would do no better than the null classifier, which for threshold t randomly assigns $t \times 100\%$ to be the success category and has $AUC=1/2$.

```
auc(ROCrres)
```

```
## [1] 0.9495202
```

Quadratic Discriminant Analysis (QDA)

- ▶ Back to the $p = 1$ case.
- ▶ If we let the variances differ by class k , then it can be shown that the Bayes classifier assigns an observation with $X = x$ to the class with the largest

$$\delta_k(x) = -\frac{1}{2} \frac{(x - \mu_k)^2}{\sigma^2} - \frac{1}{2} \log \sigma_k^2 + \log \pi_k,$$

which is a quadratic function of x .

- ▶ See the text, page 149, for the multivariate version.

QDA for the Default Data

```
qq <- qda(default ~ student + income + balance, data=Default)
preds <- predict(qq)
Default <- data.frame(Default, predicted = preds$class)
xtabs( ~ predicted + default, data=Default)
```

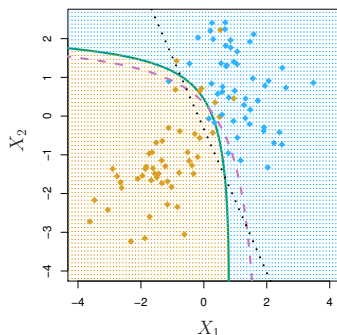
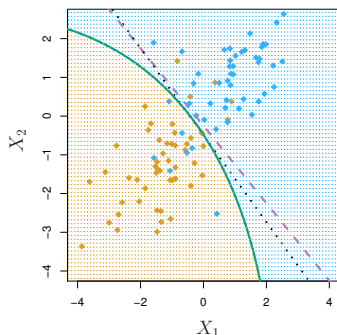
```
##           default
## predicted   No   Yes
##           No  9645  254
##           Yes   22   79
```

- ▶ Same as LDA for these data.

QDA vs LDA: Bias vs Variance

- ▶ When should we assume the K classes have a common variance matrix?
- ▶ LDA has fewer parameters and so lower variance.
 - ▶ When p is large relative to n , the number of parameters in QDA can become large, and variance of the posterior probabilities large.
- ▶ On the other hand, if the variances really **are** different, the LDA posterior probabilities can be biased.
- ▶ Recommendation from the text: If the training data set is large, use QDA. If not, use LDA.

QDA vs LDA: Decision Boundaries



- Text, Figure 4.9: $p=2$, Bayes (purple dashed), LDA (black dotted), QDA (green solid).

Comparison of Classification Methods

- ▶ Performance depends on the nature of the true decision boundary (curve that separates the two classes).
- ▶ One can show that logistic regression leads to linear decision boundaries, just like LDA
 - ▶ Difference is in how the methods estimate the parameters of the boundaries.
 - ▶ Often give very similar results, as in the Default data.
- ▶ QDA gives quadratic boundaries.
- ▶ The KNN classifier is non-parametric and so there are no restrictions on the decision boundary.
- ▶ Performance depends on the nature of the true decision boundary.