

BASIC ALGEBRA tlw

Ufuk Cem Birbiri

Ex-1)

a) Let $y(v) = Sv$, so that

$$f(v, y(v)) = U^T \cdot y(v).$$

$$\rightarrow f'(v, y(v)) = \frac{\partial f}{\partial v} + \frac{\partial f}{\partial y} \cdot y'(v)$$

The first term $\frac{\partial f}{\partial v}$ is the derivative of inner product. If we hold $y(v)$ fixed, then the derivative of $U^T \cdot y(v)$ is $y(v)^T$.

Likewise, in the second term, $\frac{\partial f}{\partial y}$ is equal to U^T .

And $y'(v) = \frac{\partial y}{\partial v}$ is the derivative of a matrix times a vector.

$$\text{So, } y'(v) = \frac{\partial (Sv)}{\partial v} = S.$$

$$\Rightarrow f'(v, y(v)) = y(v)^T + U^T S$$

Since $y(v) = Sv$, we have $y(v)^T = U^T \cdot S^T$

$$f'(v, y(v)) = U^T \cdot S^T + U^T S = U^T (S^T + S)$$

Since S is symmetric, $S^T = S$

$$f'(v, y(v)) = U^T (S^T + S) = U^T (2S) = 2 U^T S$$

(1)

$$b) g(\mathbf{v}) := \mathbf{v}^T \mathbf{v}$$

$$\nabla g(\mathbf{v}) = \nabla(\mathbf{v}^T \mathbf{v}).$$

Let's say $\mathbf{v} = (v_1, v_2, \dots, v_n)$
Let's write $g(\mathbf{v}) = \mathbf{v}^T \mathbf{v} = \sum_{i=1}^n v_i^2$.

$$\frac{d}{dv_1} (\mathbf{v}^T \mathbf{v}) = \frac{d}{dv_1} \left(\sum_{i=1}^n v_i^2 \right) = \frac{d}{dv_1} v_1^2 = 2v_1$$

We can compute the derivative for each term of \mathbf{v} .

$$\text{So } \frac{d}{dv} (\mathbf{v}^T \mathbf{v}) = 2\mathbf{v}$$

a)

Theorem: Max and Min of ~~$f(\mathbf{v})$~~ given $\|\mathbf{v}\|=1$

- Let A be a symmetric matrix that defines a quadratic form $f(\mathbf{v}) = \mathbf{v}^T A \mathbf{v}$. Then, under the condition that \mathbf{v} is a unit vector, the maximum value of $f(\mathbf{v})$ is the largest eigenvalue ~~λ~~ λ_{\max} , and the minimum value of $f(\mathbf{v})$ is the smallest eigenvalue λ_{\min} .

$$\star \max \{ f(\mathbf{v}) : \|\mathbf{v}\|=1 \} = \lambda_{\max}$$

$$\star \min \{ f(\mathbf{v}) : \|\mathbf{v}\|=1 \} = \lambda_{\min}.$$

(2)

In our case, S is a positive definite ~~symmetric~~ matrix.

$$S = Q \Lambda Q^T.$$

$$f(u) = u^T S u = u^T Q \Lambda Q^T u$$

- If $\|u\|=1$, then $\lambda_{\max} = \max \{f(u)\}$
or $\lambda_{\min} = \min \{f(u)\}$.

- we know that $u \in L$ where $L := \{u \in \mathbb{R}^D \mid \|u\|_2=1\}$.

So, we can write $f(u) = u^T Q \Lambda Q^T u = u^T S u$

Here, S is a ~~symmetric~~ matrix where $S \in \mathbb{R}^{D \times D}$.

Also, Λ is a symmetric matrix where it has eigenvalues of S in the main diagonal in decreasing order.

Let the new form of $f(u) = u^T Q \Lambda Q^T u = z^T \Lambda z$
where $z = Q^T u$.

~~Norm of the z should be equal to 1~~

~~since $f(u)$ didn't change and~~
~~only S is multiplied by Λ , which ~~is~~ is symmetric~~
and $\Lambda \in \mathbb{R}^{D \times D}$.

$$\text{So, } \|Q^T u = z\| = 1$$

b) If the $\|u\|=1$, ~~then~~ $\max \{f(u) : \|u\|=1\} = \lambda_{\max}$.

$$\text{So, } \max_{u \in L} (f(u)) = \max_{u \in L} (u^T S u) \leq \lambda_{\max} = \lambda_1$$

where $\lambda_1 > \lambda_2 > \dots > \lambda_D$

(3)

Ex-2

a Vector space V of dim. N , and a sub-vector space $W \subset V$ with an orthonormal basis $\{v_1, \dots, v_k\}$.

Orthogonal Projection: $P_W : V \rightarrow W$.

a) Show that the projection matrix associated to P_W can be written in the form $M_{P_W} = UU^T$. What're the columns of U ?

- If $M_{P_W} = U \cdot U^T$ is a projection matrix, then the columns of U are orthonormal, we have $U^T U = I$.

Clearly, $(UU^T)^T = (U^T)^T \cdot U^T = U \cdot U^T$

- ~~If U_1, \dots, U_m are orthonormal vectors spanning N -dimensional subspace,~~
~~then~~

- The columns of U are ~~the~~ U_i 's \rightarrow orthonormal basis $= \{v_1, \dots, v_k\}$

b) $M_{P_W} = U \cdot U^T$ is symmetric since:

$$(UU^T)^T = (U^T)^T U^T = U \cdot U^T \text{ since } U^T U = I$$

It is idempotent since:

$$(UU^T)^2 = (UU^T)(UU^T) = U(U^T U)U^T = UIU^T = U \cdot U^T$$

(since $U^T U = I$)

(4)

Question-3

a)

$$y_i = \mathbf{U}^T \cdot \mathbf{x}_i$$

$$E[y_1] = E[\mathbf{U}^T \mathbf{x}_1]$$

$$E[y_1] = \mathbf{U}^T - E[\mathbf{x}_1] \rightarrow \text{Since Expectation is a linear operator.}$$

//
 $\bar{y} = \mathbf{U}^T \cdot \bar{\mathbf{x}}$

(5)

Question-3

b) We are given N data points $x_1 \dots x_N$ in \mathbb{R}^D .

Let's put these data points into the rows of an $N \times D$ matrix X :

$$X = \begin{bmatrix} -x_1^T - \\ \vdots \\ -x_N^T - \end{bmatrix} \in \mathbb{R}^{N \times D}$$

For PCA to work correctly, we will need to subtract off the sample mean from each row, in particular denote by \bar{x} the sample mean:

$$\bar{x} = \frac{1}{N} \sum_{i=1}^N x_i,$$

and for each $i \in [N]$, denote by \tilde{x}_i the i -th data point after subtracting mean:

$$\tilde{x}_i = x_i - \bar{x}.$$

Let denote the resulting mean-centered data matrix:

$$\tilde{X} = \begin{bmatrix} -\tilde{x}_1^T - \\ \vdots \\ -\tilde{x}_N^T - \end{bmatrix} \in \mathbb{R}^{N \times D}$$

The $D \times D$ sample covariance matrix S can be written in terms of \tilde{X} as:

$$S = \frac{1}{N-1} \sum_{i=1}^N (x_i - \bar{x})(x_i - \bar{x})^T = \frac{1}{N-1} \tilde{X}^T \cdot \tilde{X}$$

Now, first consider projecting to ~~#~~ one dimension, say along a direction v_1 for some unit vector $v_1 \in \mathbb{R}^D$.

For any direction v , the sample variance of the projection data points

$v^T \tilde{X}_1, \dots, v^T \tilde{X}_N$ is given by:

$$\frac{1}{N-1} \sum_{i=1}^N (v^T \tilde{X}_i)^2 = \frac{1}{N-1} \| \tilde{X} v \|_2^2 = \frac{1}{N-1} v^T \tilde{X}^T \tilde{X} v = v^T S v.$$

- We wish to find $v_1 \in \mathbb{R}^D$ (with $\|v_1\|_2 = 1$) s.t. the sample variance of the projected data points is maximized:

$$v_1 \in \underset{v \in \mathbb{R}^D: \|v\|_2=1}{\operatorname{argmax}} v^T S v$$

Here, in our case $v = \underline{\text{ }}$

- It is well known that this maximum is achieved by choosing v_1 to be an eigenvector of S corresponding to the largest eigenvalue of S . Indeed, introducing a Lagrange multiplier λ for the constraint $\|v\|_2^2 = 1$,

~~(7)~~ →

, we obtain the Lagrangian function

$$-\mathbf{v}^T \mathbf{S} \mathbf{v} + \lambda (\|\mathbf{v}\|_2^2 - 1);$$

minimizing this over \mathbf{v} gives:

$$\mathbf{S} \mathbf{v} = \lambda \mathbf{v}$$

showing that \mathbf{v} must be an eigenvector of \mathbf{S} . Moreover, we have that the sample variance of the projected data points is then given by

$$\mathbf{v}^T \mathbf{S} \mathbf{v} = \lambda \|\mathbf{v}\|_2^2 = \lambda, \quad \text{the eigenvalue corresponding to } \mathbf{v}$$

to this eigenvector. Thus, the variance is maximized by taking λ to be the largest eigenvalue λ_1 of \mathbf{S} , and \mathbf{v} to be the corresponding (normalized) eigenvector

$$\mathbf{v}_1.$$

According to explanation above,

(a) $\mathbf{U}_1, \dots, \mathbf{U}_k$'s are columns of \mathbf{U} . In the explanation

$$\mathbf{v} = \mathbf{U}.$$

 Setting the derivative of J :

$$J = -\mathbf{v}^T \mathbf{S} \mathbf{v} + \lambda (\|\mathbf{v}\|_2^2 - 1) \quad \text{wrt } \mathbf{v} \text{ to 0, we}$$

get a stationary point \rightarrow

 (8)

$$\frac{\partial J}{\partial V} = (-2SV) + 2\lambda V = 0$$

$$SV = \lambda V.$$

This shows that at the stationary point, V must be an eigenvector of S , and λ be an eigenvalue corresponding to the eigenvector V .

- Left-multiplying the equation,

$J = -V^T SV + \lambda (\|V\|_2^2 - 1)$ with V^T , we can see that the maximum variance is equal to the eigenvalue λ .

$$V^T S V = V^T \lambda V$$

$$V^T S V = \lambda V^T V \rightarrow \text{since } \lambda \text{ is a scalar}$$

$$V^T S V = \lambda \rightarrow \text{since } V^T V = 1$$

Again, in this explanation $V = U$ and U_1, \dots, U_k are the columns of $V = U$.