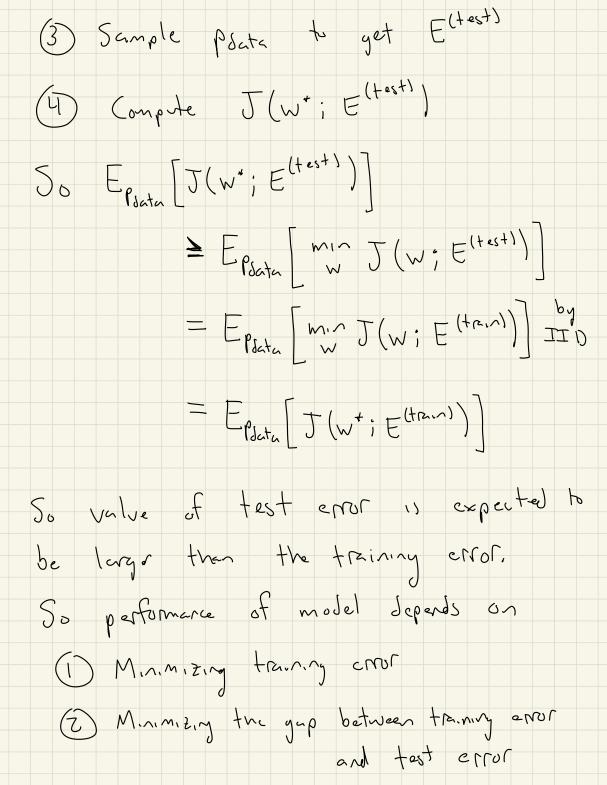
Question How to grarantoe statistically that

m.n.mizing J(train) minimizes J(test)? ITO Assumptions: Both truining set & test set
are gotten by independently randomly sampling tre same probability sistribution Polata. Usually T is accomplished by Finding Some
parameter w* which minimize cost fraction J (w; E(train)) = (E(train) means) (measures the)
average error
in E(trin) and neasony performance Us. og J (w*; E(+0s+)) The process yous: lu get E(train) () Randomly sample Polata (2) F.n. w = argmin J (w; E(+ra.n))



If you struggle with (1) this is a problem of "underfithing" (1) => "overfitting" (Informal) The capacity of a learning algo is the "size" of the space of functions the algo can fit. hypothesis space Note: To make this formal, we need statistical learning theory (e.g. VC dinersium) E.g. Linear regression with I variable Case I no polynomial tems hypothess predict with [[]] predict with $f(x) = w \times tb$ (capacity) the space of all such functions has dim 2 Casc 2 poly tams up to Jegree 4 predict with f(x) = W4 X4 + N3 X3 + W2 X2+W, X+b Now capacity is 5

Intaset looks like Suppose (training error) versus capacity - (apacity optimal occan's capacity (razor)

Bayes Error (Suppose we have Posata) Use Plata to make predictions e.g. $f(\vec{x}) = argmax \rho_{sata}(y|\vec{x})$ (spansal) Such a model provides a theoretical minimal cost for any prediction model. Define Bayes Error = min [Plata [] (f; x,y)] where $J(f; \vec{x}, y)$ is the cost of prediction $f(\vec{x}) = y$

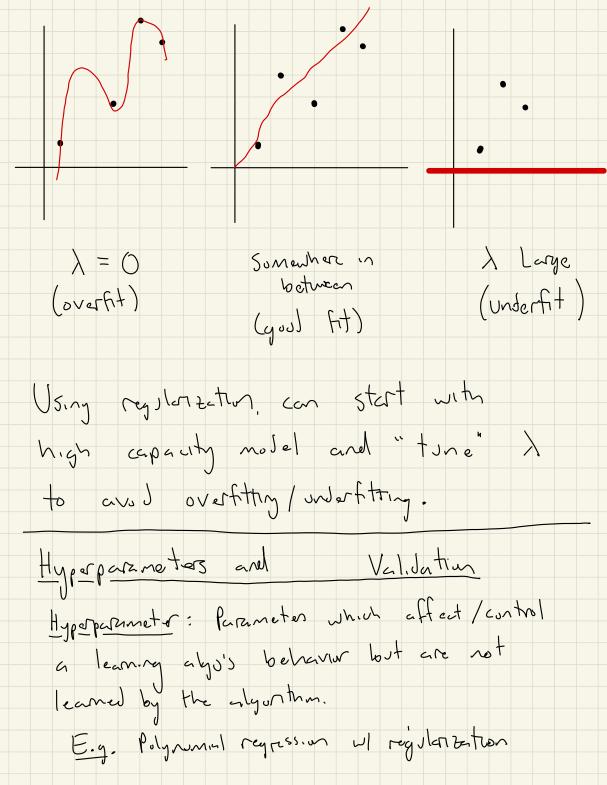
Note: Bayes Error can be nonzos - mapping x' to y can be nondetoministic - X' can be mapped to y in a determnistic way but Using more variables than we have Regularization: I dea: Add a new term to cost function

John Former restricting to hypothes.s

Space Example (Linear regression) M.s.G $J(\vec{w}) = M \hat{\chi} w - y \|_{2}^{2} + \lambda \|w\|_{2}^{2}$ nininize J la find w*, this

promotes coefficient of w* to be Since we extra tem Small.) dictates how strongly this is promoted.

E.y Us.my poly terms up to degree 10



n, & are hyperparameters max degree of poly terms Hyperparameters are usually timed by hand to minimite test error. Issue: When we do this, the test set is no longer a test set (model parameters are fit to the test set) Solution Introduce a new "test set" called validation set The process 80% 10% 10% 5plit data E into Etan Eval Etest (1) Given hyperparameter X train model with Etran by Finding $\vec{v}_{\vec{x}} = argmin J(\vec{w}, \vec{\lambda}; E^{+ain})$ 2) Evalute performance by looking et J(Wx; Eval) (without regularisation)

(3) Repeat (D&(2) with Jifferent choices of X until find lowest value J (Wxx; Eval) best } (tuning) (1) Approximate generalization error by computing J(Wz, Etest) (without regularization) Tip: If E is too small, may want to

(read ty k-fold cross validation

(about this) (allows for puttry more in test set) in python cross_val_score (skleam) Bias & Variance: Def: A point estimator for a parameter of associated to some distribution p is any function of random variables XIIII (distributed by p)

write
$$\hat{\Theta} = g(x_1, ..., x_n)$$

estimation for Θ

The bias of $\hat{\Theta}$ is

 $b_{1}a_{1}s_{2}(\hat{\Theta}) = E_{p}[\hat{\Theta}] - \Theta$

The variance of $\hat{\Theta}$ is

 $Var(\hat{\Theta}) = E_{p}[(\hat{\Theta} - E_{p})]$

$$Var(\hat{\theta}) = E_{\rho}[(\hat{\theta} - E_{\rho}[\hat{\theta}])^2]$$

E.g. XIIIIX IID random variables

Jistrboted by Barroulli distribution
$$\Omega = \{0, 1\}$$

$$p(x=1) = \theta$$

$$p(x=x) = \theta^{x} (1-\theta)^{(1-x)}$$

Estimate
$$\theta$$
 by $\hat{G} = \frac{\{i \mid x_i = 1\}}{n} = \frac{\sum_{i=1}^{n} x_i}{n}$

bias
$$(\hat{\Theta}) = \mathbb{E}_{p}[\hat{\Theta}] - \Theta$$

$$= \Gamma[\mathcal{D}^{\hat{A}}] \times \mathbb{E}_{p}[\hat{\Theta}]$$

$$= \mathbb{E}_{\rho} \left[\frac{\sum_{i=1}^{n} \times_{i}}{n} \right] - \Theta$$

$$=\frac{1}{n}\sum_{i=1}^{n}E_{p}[x_{i}]-\theta$$

$$=\frac{1}{n}\sum_{i=1}^{n}\Theta-\Theta$$

$$=\frac{n\Theta}{n}-\Theta=O$$

$$Var(\hat{\theta}) = Var(\frac{L}{n} Z_{i=1}^{n} \times i)$$

$$=\frac{1}{N^2}\sum_{i=1}^{\infty}\bigvee(X_i)$$

$$=\frac{1}{N^2}\sum_{i=1}^{n}\Theta(1-\Theta)$$

$$= \frac{\theta(1-\theta)}{n} \longrightarrow 0 \quad \text{as} \quad n \to \infty$$

Exoruses (1) Gaussian Sistribution
$$\mathcal{N}(x; \mathcal{M}, r^2)$$
estimate \mathcal{M} by IID random variables x_1, \dots, x_n

$$\hat{\mathcal{M}} = \underbrace{x_1 + \dots + x_n}_{n}$$
Show B_{i} as $(\hat{\mathcal{M}}) = 0$
(2) Estimate r^2 by $\hat{r}^2 = \frac{1}{n} \sum_{i=1}^n (x_i - \hat{\mathcal{M}})^2$
Show B_{i} as $(\hat{r}^2) = -\frac{r^2}{n}$

$$Var(\hat{r}^2) = ?$$
(3) Estimate r^2 by $\hat{r}^2 = \frac{1}{n-1} \sum_{i=1}^n (x_i - \hat{\mathcal{M}})^2$

$$Show bias (\hat{r}^2) = 0 \quad Var(\hat{r}^2) = ?$$
(all ω) the industry variance estimator)
$$Def: \text{The mean squard error } (MSE) \text{ of } \hat{\sigma}$$
15 $MSE(\hat{\sigma}) = E_p[(\hat{\theta} - \theta)^2]$

$$\underbrace{\operatorname{Lemma}}_{\mathsf{Po}} \mathsf{MSE}(\hat{\theta}) = \operatorname{Var}(\hat{\theta}) + \operatorname{Usas}(\hat{\theta})^{2}$$

$$\underbrace{\operatorname{Po}}_{\mathsf{Po}} \cdot \mathsf{F} : \mathsf{MSE}(\hat{\theta}) = \operatorname{Ep}[(\hat{\theta} - \theta)^{2}]$$

$$= \operatorname{Ep}[(\hat{\theta} - \operatorname{Ep}(\hat{\theta}))^{2} + (\operatorname{E}(\hat{\theta}) - \theta)^{2}]$$

$$+ 2(\hat{\theta} - \operatorname{E}(\hat{\theta}))(\operatorname{E}(\hat{\theta}) - \theta)^{2}$$

$$+ 2 \operatorname{E}[(\hat{\theta} - \operatorname{E}(\hat{\theta}))(\operatorname{E}(\hat{\theta}) - \theta)^{2}]$$

$$+ 2 \operatorname{E}[(\hat{\theta} - \operatorname{E}(\hat{\theta}))(\operatorname{E}(\hat{\theta}) - \theta)^{2}]$$

$$+ 2 \operatorname{E}[(\hat{\theta} - \operatorname{E}(\hat{\theta}))(\operatorname{E}(\hat{\theta}) - \theta)^{2}]$$

$$+ 2 (\operatorname{E}[\hat{\theta}] - \theta) : \quad \text{constart}$$

$$= \operatorname{Var}(\hat{\theta}) + (\operatorname{E}[\hat{\theta}] - \theta)^{2}$$

$$+ 2 (\operatorname{E}[\hat{\theta}] - \theta) : \quad \text{E}[(\hat{\theta} - \operatorname{E}[\hat{\theta})]$$

$$= \operatorname{E}[(\hat{\theta}) + \operatorname{E}[(\hat{\theta})] - \theta)^{2}$$

$$= \operatorname{Var}(\hat{\theta}) + \operatorname{bas}(\hat{\theta})^{2}$$

Maximum Likelihoud Estimators Maximum

Back to Usual Setup: Plata generally

X = X +rain

X +est

X has rows x'(1),..., x'(m)

+rated as random

variables distributed

by Plata Gal: Approximate Plata(X). Let Prodel (x, 5) be an approximation for Plata depending on parameter & Egg. If we expect Plates to be gassian Prodel (E, O) = N(x; mo, Eo) Def The maximum likelihood estimator Fr 6 15

