Data Analysis

Moscow 2025

Lecture 6 Linear Regression

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Regression analysis

Regression analysis is a set of statistical methods used to estimate relationships between a dependent variable and one or more independent variables. It depicts how dependent variable will change when one or more independent variable changes. The results of regression analysis can be used to forecast the values of the dependent variable. One point to keep in mind with regression analysis is that causal relationships among the variables cannot be determined. While the terminology is such that we say that X "predicts" Y, we cannot say that X "causes" Y. Wile forecasting we should also consider how external factors could affect the values of the dependent variable.



Linear regression

Linear regression analysis rests on the assumption that the dependent variable is continuous. The independent variables can be either continuous or dichotomous. Independent variables with more than two levels can also be used in regression analyses, but they first must be converted into variables that have only two levels. This is called dummy coding.



Simple linear regression

The linear relationship between the variables is expressed using an equation of a straight line. In simple regression we predict an outcome based on a single predictor.

$$Y = a + b * X$$

where:

- X is an independent variable (predictor);
- Y is the dependent variable;
- a, b are constant values (model parameters).

The model parameters are determined using the least squares method.

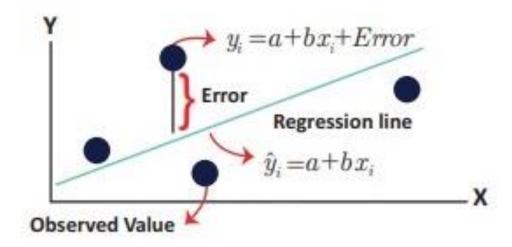


Method of least squares

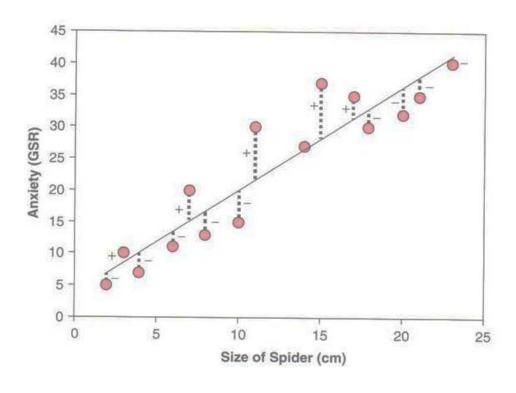
Method of least squares is a way of finding the line that best fits the data (i.e. the line that goes through, or is close to, as many of the data points as possible).

$$E(a,b) = \sum_{i=1}^{n} (y_i - \hat{y}_i)^2$$

i.e.,
$$E(a,b) = \sum_{i=1}^{n} (y_i - a - bx_i)^2$$
.



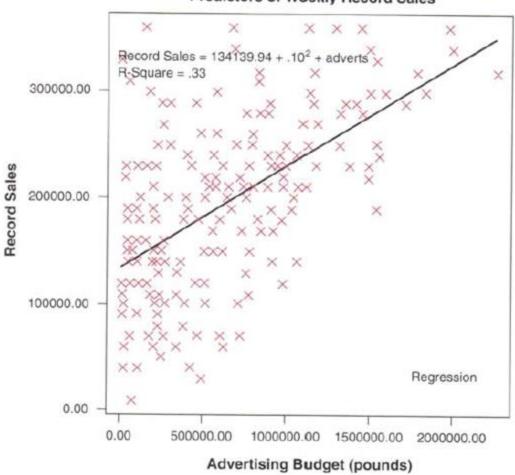
Simple Linear Regression Model



This graph shows a scatterplot of some data with a line representing the general trend. The vertical lines (dashed) represent the differences (or residuals) between the line and the actual data. The method of least squares works by selecting the line that has the lowest sum of squared differences.

Data Analysis

Predictors of Weekly Record Sales



To evaluate how well a straight line expresses the relationship between variables, it is useful to examine the scatterplot to see if there is a linear relationship and to detect outliers that can significantly skew the result.



Sum of squares

Residual Sum of Squares (RSS)

$$RSS = \sum_{i=1}^{n} (y_i - \hat{y}_i)^2$$

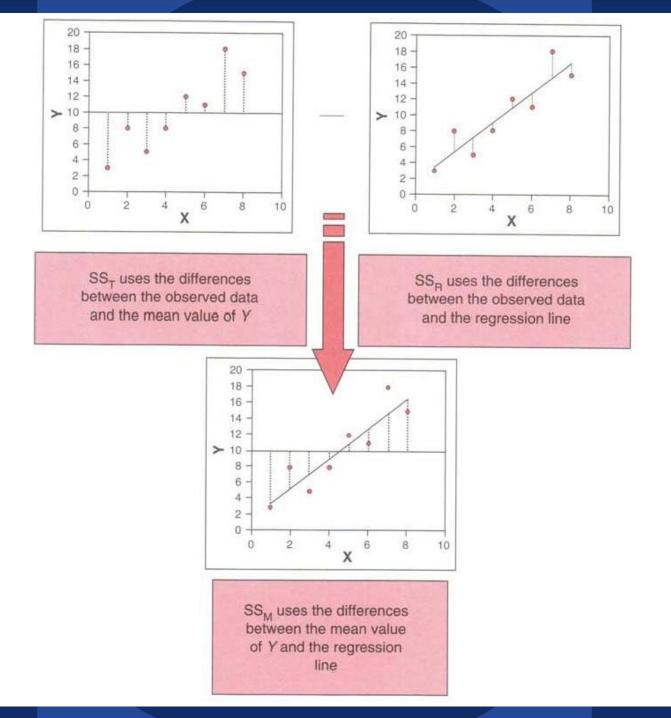
Total Sum of Squares (TSS)

$$TSS = \sum_{i=1}^{n} (y_i - \overline{y})^2$$

Explained Sum of Squares (ESS)

$$ESS = \sum_{i=1}^{n} (\hat{y}_i - \overline{y})^2$$

TSS = RSS + ESS

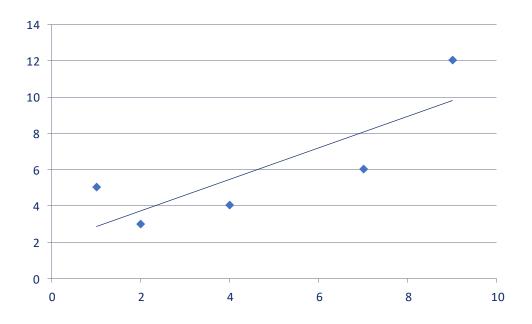




Simple linear regression: example

X	Y
1	5
2	3
4	4
7	6
9	12

$$b = \frac{\sum (x_i - \bar{x})(y_i - \bar{y})}{\sum (x_i - \bar{x})^2}$$
$$a = \bar{y} - b\bar{x}$$



$$Y = 0.8*X + 2$$

Multiple Linear Regression

Lecture 6

The multiple regression model supposes that we have more than one predictor. Each predictor variable has its own coefficient, and the outcome variable is predicted from a combination of all the variables multiplied by their respective coefficients plus a residual term.

$$Y_i = (b_0 + b_1X_1 + b_2X_2 + ... + b_nX_n) + e_i$$



Regression Equation

$$Y_i = (b_0 + b_1X_1 + b_2X_2 + ... + b_nX_n) + e_i$$

In the equation:

- b₁ is the coefficient of the first predictor (X₁);
- b_n is the coefficient of the n^{th} predictor (X_n) ;
- e_i is the difference between the predicted and the observed value of Y for the ith participant.



Selection of predictors

The predictors included and the way in which they are entered into the model can have an impact on the result. It is recommended to select predictors based on past research (but it's necessary to be sure that the past research was done appropriately). The selection could be also based of the theoretical framework of the research.



Adding categorical predictors into the model

If the variable is dichotomous, it could be added into the model without any preliminary transformation.

If the categorical variable has more than two values, it should be recoded into several dummy variables. Each group of cases, defined by the values of the categorical variable, should contain not less then 15% of cases.



Creating dummy variables

If the categorical variable has n values, we will enter into the model n-1 dichotomous variables.

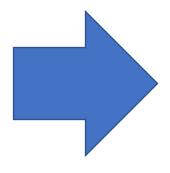
One category is selected as a "basic" or "reference" and the other categories will be compared with this "basic" category. It could be a category, which contains the biggest number of cases.



Creating dummy variables

The initial variable "Educational level" takes the following values: no (1), secondary (2), higher (3), postgraduate (4). We can choose "higher" as a reference group.

Educational level
higher
no
secondary
postgraduate
higher
secondary



Ed_no	Ed_secondary	Ed_postgraduate	
0	0	0	
1	0	0	
0	1	0	
0	0	1	
0	0	0	
0	1	0	



Coefficient of determination (R-squared)

R-squared is the proportion of variation in the dependent variable that is accounted for by the model. Value of the coefficient varies from 0 to 1. An R-squared of 1 indicates that the regression predictions perfectly fit the data.

$$R^{2} = 1 - \frac{SS_{RES}}{SS_{TOT}} = 1 - \frac{\sum_{i} (y_{i} - \hat{y}_{i})^{2}}{\sum_{i} (y_{i} - \overline{y}_{i})^{2}}$$



R-squared

Regression models with low R-squared values can be perfectly good models. Some fields of study have an inherently greater amount of unexplainable variation. In these areas R² values are bound to be lower. For example, studies that try to explain human behavior generally have R² values less than 50%. People are just harder to predict than things like physical processes. If we have a low R-squared value but the independent variables are statistically significant, we can still draw important conclusions about the relationships between the variables. There is a scenario where small R-squared values can cause problems. If we need to generate predictions that are relatively precise, a low R² can be a showstopper.



Adjusted R-squared

Adjusted R-squared is a modified version of R-squared that has been adjusted for the number of predictors in the model. The adjusted R-squared increases when the new term improves the model more than would be expected by chance. It decreases when a predictor improves the model by less than expected. Adjusted R-squared could be used to compare models that have a different number of variables.

adjusted
$$R^2 = 1 - \left[\left(\frac{n-1}{n-k-1} \right) \left(\frac{n-2}{n-k-2} \right) \left(\frac{n+1}{n} \right) \right] (1-R^2)$$

where n — number of cases, k — number of predictors



F-statistics

The F-test compares the current model with zero predictor variables (the intercept only model) and decides whether the added coefficients significantly improved the model. If we get a significant result, then whatever coefficients we included in the model improved the model's fit.

$$F = \frac{R^2}{1 - R^2} \frac{(n - m - 1)}{m}$$



Mean Absolute Error (MAE)

In the fields of statistics and machine learning, the Mean Absolute Error (MAE) is a frequently employed metric. It's a measurement of the typical absolute discrepancies between a dataset's actual values and projected values.

$$MAE = rac{1}{n} \sum_{i=1}^{n} |x_i – y_i|$$

Where:

- x_i represents the actual or observed values for the i-th data point.
- y_i represents the predicted value for the i-th data point.



Mean Squared Error (MSE)

A popular metric in statistics and machine learning is the Mean Squared Error (MSE). It measures the square root of the average discrepancies between a dataset's actual values and projected values. MSE is frequently utilized in regression issues and is used to assess how well predictive models work.

$$MSE = rac{1}{n} \sum_{i=1}^n (x_i – y_i)^2$$

where:

- x_i represents the actual or observed value for the i-th data point.
- y_i represents the predicted value for the i-th data point.



Root Mean Squared Error (RMSE)

RMSE is a usually used metric in regression analysis and machine learning to measure the accuracy or goodness of fit of a predictive model, especially when the predictions are continuous numerical values. The RMSE quantifies how well the predicted values from a model align with the actual observed values in the dataset.

$$RMSE = \sqrt{rac{1}{n}\sum_{i=1}^n (x_i \! - \! y_i)^2}$$

Where:

- RMSE is the Root Mean Squared Error.
- x_i represents the actual or observed value for the i-th data point.
- y_i represents the predicted value for the i-th data point.



Evaluating the accuracy of the regression model

When we have produced a model based on a sample of data two important questions should be considered:

- 1. Does the model fit the observed data well, or it's influenced by a small number of cases?
- 2. Can the model generalize to other samples?

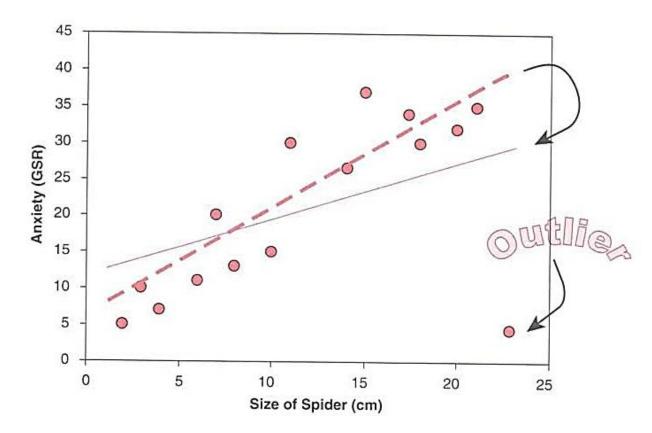


Diagnostics of the model: outliers and influential cases

An **outlier** is a case which differs substantially from the main trend of the data. Outliers can cause the model to be biased because they affect the values of the estimated regression coefficients.

An **influential case** is a case which has serious influence over the parameters of the model. If we delete it the regression coefficients will change.

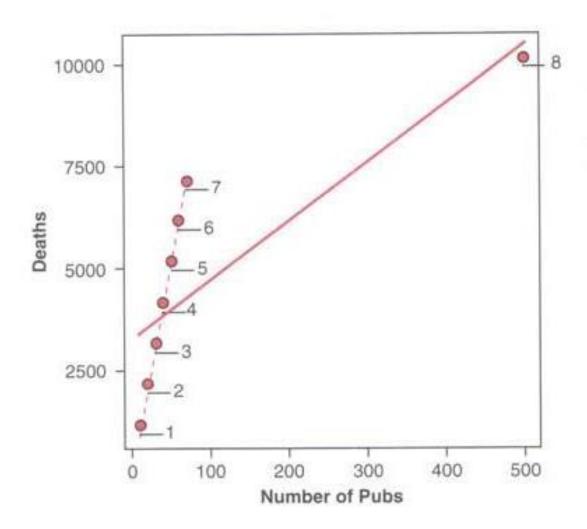
Outliers



The change in one point had a dramatic effect on the regression model: the gradient reduced (the line becomes flatter) and the intercept increases (the line crosses the Y-axis at a higher point).



Influential cases



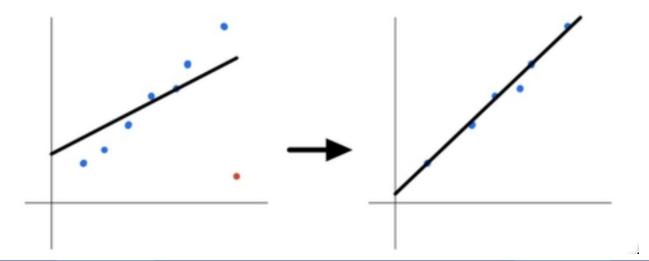
An influential case is any case that significantly alters the value of a regression coefficient whenever it is deleted from an analysis. If the deletion particular cases in an analysis alters the parameters of the regression equation significantly, then these cases represent influential cases.



How to identify an influential case?

Influential statistics: DfBeta(s) shows how the regression coefficients change if the influential case is excluded.

The case could be an influential if the values of these statistics are greater than 1.





Diagnostics of the model: residuals

Residuals are the differences between the values of the outcome observed in the sample and the values of the outcome predicted by the model. If the model fits the sample data well then all residuals will be small and their distribution will be not different from normal. If a particular case has a large residual, then it could be an outlier. Potential outlier is a case with standardized residual greater than 3 or less than -3.

Multicollinearity

Multicollinearity is the occurrence of high intercorrelations among two or more independent variables in a multiple regression model. Multicollinearity can lead to skewed or misleading results when a researcher or analyst attempts to determine how well each independent variable can be used most effectively to predict or understand the dependent variable in a statistical model. It is a situation when there is a strong correlation (r > 0.7) between two or more predictors in a regression model.

VIF (variance-inflation factor) indicates whether a predictor has a strong linear relationship with the other predictor(s). If VIF is greater than 5 there could be a multicollinearity.

$$VIF x_i = \frac{1}{Tolerance} = \frac{1}{1 - R_i^2}$$



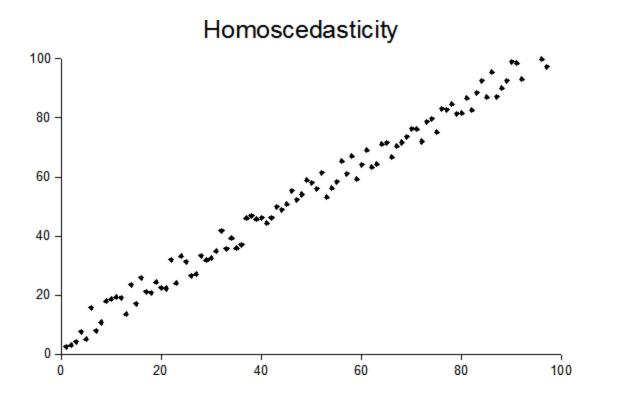
Heteroscedasticity

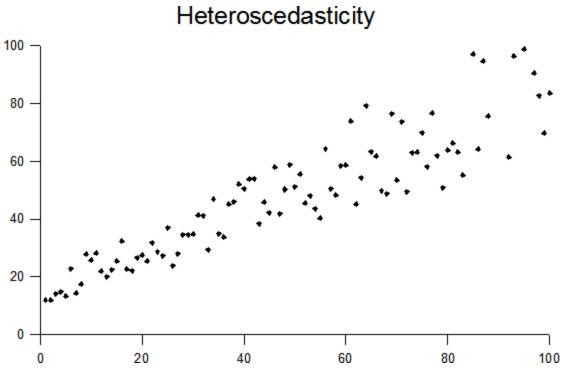
It is supposed that in good regression models the variance of the residuals is homogeneous across levels of the predicted values (homoscedasticity). If the model is well-fitted, there should be no pattern to the residuals plotted against the fitted values. If the variance of the residuals is non-constant, then there is a heteroscedasticity.

The graphical analysis could detect heteroscedasticity. A commonly used graphical method is to use the plot to show the residuals versus fitted (predicted) values.



Heteroscedasticity in a simple regression model

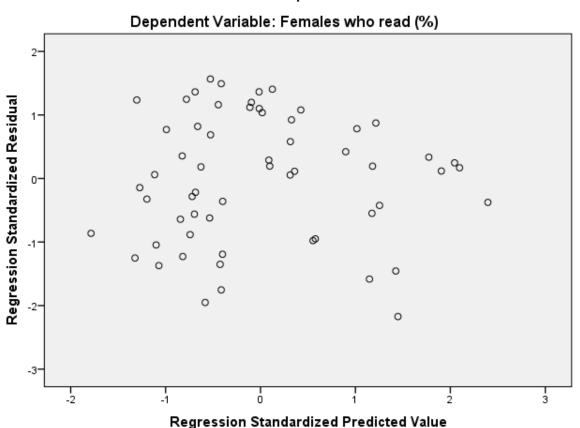






Heteroscedasticity in a multiple regression model

Scatterplot





Gauss Markov Theorem

The **Gauss Markov theorem** tells us that if a certain set of assumptions are met, the ordinary least squares estimate for regression coefficients gives you the best linear unbiased estimate possible.

There are five Gauss Markov assumptions:

- Linearity: the parameters we are estimating using the OLS method must be themselves linear.
- Random: our data must have been randomly sampled from the population.
- Non-Collinearity: the regressors being calculated aren't perfectly correlated with each other.
- Exogeneity: the regressors aren't correlated with the error term.
- Homoscedasticity: no matter what the values of our regressors might be, the error of the variance is constant.



Akaike Information Criterion (AIC)

A common way to compare models is by using the so-called information criterion. It is a way to balance bias and variance or accuracy (fit) and simplicity (parsimony).

 $AIC_p = n * ln(\frac{SSE_p}{n}) + 2 * p$

p is the number of estimated parameters (including the constant), n is the number of observations,

SSE is the residual sum of squares.

The smaller the AIC the better. A model is going to be better when the sample size is larger, the unexplained variance is lower and we use the fewer parameters. AIC is a relative measure that compares one model to another to choose the one that loses less information. It's NOT a measure of how good a model is.

R Data Analysis

OLS Regression Results							
Dep. Variable:	SalePrice	R-squared:		0.681			
Model:	OLS	Adj. R-squared:		0.681			
Method:	Least Squares	F-statistic:		1037.			
Date:	Fri, 25 Feb 2022	Prob (F-statistic):		0.00			
Constant – value of the	15:31:21	Log-Likelihood:		-17709.			
	1460	AIC:		3.543e+04			
dependent variable if all	1456	BIC:		3.545e+04			
the predictors are equal to	3						
zero	nonrobust						
	oef std err	t P> t	[0.025	0.975			

	coef	std err	t	P> t	[0.025	0.975]
const	-3.132e+04	3992.921	-7.844	0.000	-3.92e+04	-2.35e+04
GrLivArea	65.6106	2.667	24.600	0.000	60.379	70.842
GarageCars	3.365e+04	1854.413	18.146	0.000	3e+04	3.73e+04
TotalBsmtSF	50.4508	3.136	16.087	0.000	44.299	56.603
Omnibus:		520.28	0 Durbin-			1.978
Prob (Omnibus	3):	0.00	0 Jarque-	Bera (JB):		31276.464
Skew:		-0.82	2 Prob(JE	3):		0.00
Kurtosis:		25.61	5 Cond. N	Jo.		6.64e+03



			91E3	======			
Dep. Variabl	 e:	SalePr	ice	R-squ	 ared:		0.681
Model:			OLS	Adj.	R-squared:		0.681
Method:		Least Squa	res	F-sta	tistic:		1037.
Date:	Fri	i, 25 Feb 2	022	Prob	(F-statistic)	:	0.00
Time:		15:31	:21	Log-I	ikelihood:		-17709.
Dograssian coeffi	signt for a corta	in variable	460	AIC:			3.543e+04
Regression coeffi			456	BIC:			3.545e+04
means that one-un			3				
will lead to an in			ıst				
dependent v	ariable by 65.6	units	====				
	coei	std err		t	P> t	[0.025	0.975]
const	- <u>132e+04</u>	3992.921		-7.844	0.000	-3.92e+04	-2.35e+04
GrLivArea	65.6106	2.667	i	24.600	0.000	60.379	70.842
GarageCars	3.365e+U4	1854.413		18.146	0.000	3e+04	3.73e+04
TotalBsmtSF	50.4508	3.136		16.087	0.000	44.299	56.603
Omnibus:		520.	280	Durbi	.n-Watson:		1.978
Prob(Omnibus):	0.	000	Jarqu	ue-Bera (JB):		31276.464
Skew:		-0.	822	Prob (JB):		0.00
Kurtosis:		25.	615	Cond.	No.		6.64e+03

Linear Regression

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56.603

OLS Regression Results

Dep. Variable:	SalePrice	R-squared:	0.681
Model:	OLS	Adj. R-squared:	0.681
Method:	Least Squares	F-statistic:	1037.
Date:	Fri, 25 Feb 2022	Prob (F-statistic):	0.00
Time:	15:31:21	Log-Likelihood:	-17709.

No. Observations:

TotalBsmtSF

Regression coefficients

nonr covariance rype:

coef

Regression Equation: SalePrice = -31320 + 65,6*GrLivArea + 33650*GarageCars + 50,5*TotalBsmtSF

44.299

0.000

-3.132e+04 3992.921 -7.844 0.000 -3.92e+04 -2.35e+04const 70.842 65.6106 2.667 24.600 0.000 60.379 GrLivArea 3.365e+04 1854.413 18.146 0.000 3e+04 3.73e+04 GarageCars 50.4508 3.136

16.087

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std er

Omnibus:	520.280	Durbin-Watson:	1.9/8
Prob(Omnibus):	0.000	Jarque-Bera (JB):	31276.464
Skew:	-0.822	Prob(JB):	0.00
Kurtosis:	25.615	Cond. No.	6.64e+03

SalePrice	R-squar	ed:		0.681
OLS	Adj. R-	squared:		0.681
east Squares	F-stati	stic:		1037.
25 Feb 2022	Prob (F	-statistic	:):	0.00
15:31:21	Log-Lik	elihood:		-17709.
1460	_	- 1		543e+04
1456	v3:00:	7 1	\(\sigma\)	545e+04
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nonrobust	J(01) —	$\sqrt{n-2}$	$\sum (x_i - \bar{x})^2$	2
std err	t	P> t	[0.025	0.975]
3992.921	-7.844	0.000	-3.92e+04	-2.35e+04
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	OLS east Squares 25 Feb 2022 15:31:21 1460 1456 3 nonrobust std err 3992.921 2.667 1854.413 3.136 520.280 0.000 -0.822	OLS Adj. R- east Squares F-stati 25 Feb 2022 Prob (F 15:31:21 Log-Lik 1460 1456 3 s(b_1) = nonrobust std err t 3992.921 -7.844 2.667 24.600 1854.413 18.146 3.136 16.087 520.280 Durbin- 0.000 Jarque0.822 Prob (JB	OLS Adj. R-squared: east Squares F-statistic: 25 Feb 2022 Prob (F-statistic) 15:31:21 Log-Likelihood: 1460 1456 3 nonrobust $s(b_1) = \sqrt{\frac{1}{n-2}} *$ std err t P> t 3992.921 -7.844 0.000 2.667 24.600 0.000 1854.413 18.146 0.000 3.136 16.087 0.000 520.280 Durbin-Watson: 0.000 Jarque-Bera (JB): -0.822 Prob(JB):	OLS Adj. R-squared: east Squares F-statistic: 25 Feb 2022 Prob (F-statistic): 15:31:21 Log-Likelihood: $s(b_1) = \sqrt{\frac{1}{n-2} * \frac{\sum (y_i - \hat{y}_i)^2}{\sum (x_i - \bar{x})^2}}$ std err t P> t [0.025] $s(b_1) = \sqrt{\frac{1}{n-2} * \frac{\sum (y_i - \hat{y}_i)^2}{\sum (x_i - \bar{x})^2}}$ std err t P> t [0.025] $s(b_1) = \sqrt{\frac{1}{n-2} * \frac{\sum (y_i - \hat{y}_i)^2}{\sum (x_i - \bar{x})^2}}$ std err t P> t [0.025] $s(b_1) = \sqrt{\frac{1}{n-2} * \frac{\sum (y_i - \hat{y}_i)^2}{\sum (x_i - \bar{x})^2}}$ std err t P> t [0.025] $s(b_1) = \sqrt{\frac{1}{n-2} * \frac{\sum (y_i - \hat{y}_i)^2}{\sum (x_i - \bar{x})^2}}$ std err t P> t [0.025] $s(b_1) = \sqrt{\frac{1}{n-2} * \frac{\sum (y_i - \hat{y}_i)^2}{\sum (x_i - \bar{x})^2}}$ std err t P> t [0.025] $s(b_1) = \sqrt{\frac{1}{n-2} * \frac{\sum (y_i - \hat{y}_i)^2}{\sum (x_i - \bar{x})^2}}$ std err t P> t [0.025] $s(b_1) = \sqrt{\frac{1}{n-2} * \frac{\sum (y_i - \hat{y}_i)^2}{\sum (x_i - \bar{x})^2}}$ std err t P> t [0.025] $s(b_1) = \sqrt{\frac{1}{n-2} * \frac{\sum (y_i - \hat{y}_i)^2}{\sum (x_i - \bar{x})^2}}$ std err t P> t [0.025] $s(b_1) = \sqrt{\frac{1}{n-2} * \frac{\sum (y_i - \hat{y}_i)^2}{\sum (x_i - \bar{x})^2}}$ std err t P> t [0.025] $s(b_1) = \sqrt{\frac{1}{n-2} * \frac{\sum (y_i - \hat{y}_i)^2}{\sum (x_i - \bar{x})^2}}$ std err t P> t [0.025] $s(b_1) = \sqrt{\frac{1}{n-2} * \frac{\sum (y_i - \hat{y}_i)^2}{\sum (x_i - \bar{x})^2}}$ std err t P> t [0.025] $s(b_1) = \sqrt{\frac{1}{n-2} * \frac{\sum (y_i - \hat{y}_i)^2}{\sum (x_i - \bar{x})^2}}}$ std err t P> t [0.025]

Linear Regression

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		OLS Re	gression ke 	:suits 		
Dep. Variabl	le:	SalePr	-	ared: R-squared:		0.681
Method:		Least Squa:		tistic:		1037.
Date:	Fri	i, 25 Feb 2	022 Prob	(F-statistic	:):	0.00
Time:		15:31	:21 Log-L	ikelihood:		-17709.
No. Observ		1.	460 AIC:			3.543e+04
Df Residua	CO	ef 1	456 BIC:			3.545e+04
Df Model:	t = -		3			
Covariance	ctd	err cob	ust			
=======	coet	sto err	t	P> t	[0.025	0.975]
const	-3.132e+04	3992.921	-7.844	0.000	-3.92e+04	-2.35e+04
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Kurtosis:		25.	615 Cond.	No.		6.64e+03

6.64e+03



Kurtosis:

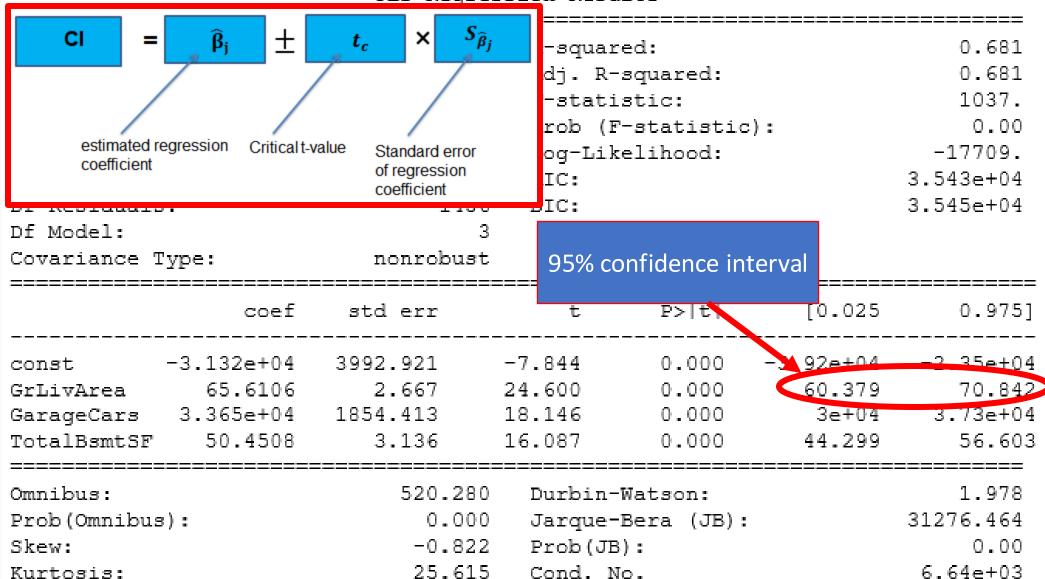
	 	OLS Reg	 ression Resu 	ılts		
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Df Residual	s:					3.545e+04
Df Model:		Statist	ical significan	ce of		
Covariance	rype:	regre:	ssion coefficie	ents		
	coef	sta err	E	Palti	[0.025	0.975
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TotalBsmtSF	50.4508	3.136	16.087	0.000	44.299	56.60
Omnibus:		520.2	 80 Durbin-	-Watson:		1.978
Prob(Omnibu	s):	0.0	00 Jarque-	Bera (JB):		31276.464
Skew:	•	-0.8				0.00

Cond. No.

Linear Regression

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OLS	Regre	ssion	Results





=========							
Dep. Variable: SalePrice		-			0.681		
Model: OLS		3 Adj.	R-squared:		0.681 1037.		
		Least Square:	s F-sta	F-statistic:			
		i, 25 Feb 202:	2 Prob	(F-statistic	:):	0.00	
Time:		15:31:23	l Log-I			-17709.	
No. Observat	ions:	146) AIC:	R-squared	, the model	3.543e+04	
Df Residuals	:	145	5 BIC:	explains 68%	% of variation	3.545e+04	
Df Model:		:	3		ependent		
Covariance T	ype:	nonrobus	_		•		
=========	=========	==========	=======	variable	e's values	=======	
	coef	std err	t	P> t	[0.025	0.975]	
const	-3.132e+04	3992.921	-7.844	0.000	-3.92e+04	-2.35e+04	
GrLivArea	65.6106	2.667	24.600	0.000	60.379	70.842	
GarageCars	3.365e+04	1854.413	18.146	0.000	3e+04	3.73e+04	
TotalBsmtSF	50.4508	3.136	16.087	0.000	44.299	56.603	
Omnibus:		520.28	Durbi	.n-Watson:		1.978	
Prob (Omnibus):	0.00) Jarqu	e-Bera (JB):		31276.464	
Skew:		-0.82	2 Prob(JB):		0.00	
Kurtosis:		25.61	o Cond.	No -		6.64e+03	

Data Analysis

			=======					
Dep. Variabl	.e:	SalePric	e R-sq	uared:			0.	.681
Model:		OI	s Adj.	R-square	ed:		0	.681
Method:		Least Square	s F-st	atistic:			1(037.
Date:	Fr	i, 25 Feb 202	2 Prob	(F-stati	istic):	ı		0.00
Time:		15:31:2	1 Log-	Likelihoo	od:		- 17'	709.
No. Observat	ions:	146	O AIC:				3.543	e+04
Df Residuals	s:	145	6 BIC:				3.545	e+04
Df Model:			3					
Covariance I	'ype:	nonrobus	t		-	The model is	S	
	coef	std err	t	P>	statis	stically signi	ficant	.975]
const	-3.132e+04	3992.921	-7.844	0.0	000 -	-3.92e+04	-2.3	5e+04
GrLivArea	65.6106	2.667	24.600	0.0	000	60.379	7(0.842
GarageCars	3.365e+04	1854.413	18.146	0.0	000	3e+04	3.7	3e+04
TotalBsmtSF	50.4508	3.136	16.087	0.0	000	44.299	5(6.603
Omnibus:		520.28	0 Durb:	in-Watsor	-===== 1:		1	.978
Prob (Omnibus	:):	0.00		ue-Bera (31276.	
Skew:		-0.82	_	(JB):				0.00
Kurtosis:		25.61		. No.			6.64	e+03



OLS	Reare:	ssion	Results

		=======	========		========
Dep. Variable:	SalePrice	R-squa	red:		0.681
Model:	OLS	Adj. R	-squared:		0.681
Method:	Least Squares	F-stat	istic:		1037.
Date:	Fri, 25 Feb 2022	Prob (F-statistic	e):	0.00
Time:	15:31:21	Log-Li	kelihood:		-17709.
No. Observations:	1460	AIC:			3.543e+04
Df Residuals:	1456	BIC:			3.545e+04
Df Model:	3				
Covariance Type:	nonrobust				
		t t	P> t	[0.025	0.975
	nality test for residuals:	0.4.4	0.000	-3.92e+04	-2 35e+0
GrLiv residuals are r	not normally distributed	600	0.000		70.84
GarageCars 3.365e	+04 1854.413	18.146	0.000	3e+04	
TotalBsmtSF 50.4		16.087	0.000	44.299	56.60
======================================	520.280	Durbin	======== -Watson:	========	 1.978
Prob(Omnibus):	0.000	Jarque	-Bera (JB):	:	31276.464
Skew:	-0.822	Prob(J			0.00
Kurtosis:	25.615	Cond.	•		6.64e+03

Linear Regression

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OLS	Regression	Results

		OLS Req	ression	. Re 	:sults 		
Dep. Variable:		SalePri		_	ared:		0.681
Model:					R-squared:		0.681 1037.
Method:		Least Squar			tistic:	_	
Date:	rrı	•			(F-statistic)	:	0.00
Time:		15:31:		-	ikelihood:		-17709.
No. Observations:			160 AI				3.543e+04
Df Residuals:		14	156 BI	C:			3.545e+04
Df Model:			3				
Covariance Type:		nonrobu	ıst				
	coef	std err	=====	=== t	P> t	[0.025	0.975]
const			-7.8	44	0.000	-3.92e+04	-2.35e+04
GrLiv Skewness and	Kurtosi	s calculated	24.6	00	0.000	60.379	70.842
Compo	residual		18.1	46	0.000	3e+04	3.73e+04
Total	residual		16.0	87	0.000	44.299	56.603
Omnibus:	=====	520.2	80 Du	=== rhi	 .n-Watson:	=======	1.978
Prob(Omnibus):		0.00			ne-Bera (JB):		31276.464
Skew:			322 Pr	_			0.00
Kurtosis:		25.6	_		No.		6.64e+03

Data Analysis

	OLS Regre	ssion Res	ults 		
Dep. Variable:	SalePrice	R-squa	red:		0.681
Model:	OLS	: Adj.R	-squared:		0.681
Method:	Least Squares	F-stat	istic:		1037.
Date:	Fri, 25 Feb 2022	Prob (F-statistic)	:	0.00
Time:	15:31:21	Log-Li	kelihood:		-17709.
No. Observations:	1460	AIC:			3.543e+04
Df Residuals:	1456	BIC:			3.545e+04
Df Model:	3				
Covariance Type:	nonrobust	:			
	ef std err	t			
2 122-1	0.4 2002 021	7 044		•	the values var
	04 3992.921	-7.844 24.600	from 1 to 2 t	here is a ho	moscedasticity
GrLivArea 65.61 GarageCars 3.365e+		18.146	0.000	3e 04	3.73e+04
TotalBsmtSF 50.45		16.087	0.000	44.293	56.603
Omnibus:	 520.280	Durbin	======== -Watson:		1.978
Prob(Omnibus):	0.000	Jarque	-Bera (JB):		31276.464
Skew:	-0.822	-			0.00
Kurtosis:	25.615	•	r		6.64e+03

Linear Regression

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			========				
Dep. Variabl	.e:	SalePric	e R-sq	uared:			0.68
Model:		OL	S Adj.	R-squar	ed:		0.68
Method:		Least Square	s F-st	atistic:			1037
Date:	Fr.	i, 25 Feb 202	2 Prob	(F-stat	istic)):	0.0
Time:		15:31:2	1 Log-:	Likeliho	od:		-17709
No. Observat	ions:	146	0 AIC:				3.543e+0
Df Residuals	; :	145	6 BIC:				3.545e+0
Df Model:			3		<i>a i</i>	~ ^ T	
Covariance I	lype:	nonrobus	t	ŀ	$\mathcal{H}_0: \mathcal{A}$	S=0, K	= 31
	coef	std err	t	P>	· t	[0.025	0.97
const	-3.132e+04	3992.921	-7.844		Nan		+
GrLivArea	65.6106	2.667	24.600		NOT	mality test fo	r the 8
GarageCars	3.365e+04	1854.413	18.146			residuals	+
TotalBsmtSF	50.4508	3.136	16.087	0.	000	44.299	56.6
Omnibus:		520.28	======= 0 Durb:	======= in-Watso	n:	========	1.97
Prob (Omnibus	;):	0.00		ue-Bera			31276.46
Skew:	r	-0.82	-	(JB):	. ,		0.0
Kurtosis:		25.61		. No.			6.64e+0

Useful links

Data Analysis

- https://dss.princeton.edu/online help/analysis/regression intro.htm
- https://datatofish.com/statsmodels-linear-regression/
- https://mlu-explain.github.io/linear-regression/

Thank you for your attention!