02477 - Bayesian Machine Learning: Lecture 11

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Outline

1 The ELBO objective

2 Free-form vs fixed-form variational inference

Hyperparameter estimation in VI

4 Scaling Gaussian processes using variational inference



Variational inference: big picture

Our goal is to approximate a posterior distribution of interest

$$p \equiv p(\mathbf{w}|\mathbf{y}) = \frac{p(\mathbf{y}|\mathbf{w})p(\mathbf{w})}{p(\mathbf{y})}$$

- Variational inference in three steps
 - 1. Define collection of "simple" approximate probability distributions Q (the variational family)
 - 2. Define a measure of "distance" between probability distributions $\mathbb{D}[q||p]$
 - 3. Search for the distribution $q \in \mathcal{Q}$ that resembles the exact posterior p as close as possible as measured by $\mathbb{D}[q||p]$
- The variational approximation q for the target distribution $p \approx q$ is defined as

$$q_* = \arg\min_{q \in \mathcal{Q}} \mathbb{D}\left[q||p
ight]$$

Understanding the ELBO objective

Suppose our model of interest is

$$p(\mathbf{y}, \mathbf{w}) = \prod_{n=1}^{N} p(y_n | \mathbf{w}) p(\mathbf{w})$$

■ We optimize the ELBO \mathcal{L} to minimize the KL divergence (between approximation and target)

$$\mathcal{L}[q] = \mathbb{E}_q[\ln p(\mathbf{y}, \mathbf{w})] - \mathbb{E}_q[\ln q(\mathbf{w})]$$

Re-writing the lower bound

$$\mathcal{L}[q] \equiv \mathbb{E}_q \left[\ln \prod_{n=1}^N p(y_n | \boldsymbol{w}) p(\boldsymbol{w}) \right] - \mathbb{E}_q \left[\ln q(\boldsymbol{w}) \right]$$

$$= \sum_{n=1}^N \mathbb{E}_q \left[\ln p(y_n | \boldsymbol{w}) \right] + \mathbb{E}_q \left[\ln p(\boldsymbol{w}) \right] - \mathbb{E}_q \left[\ln q(\boldsymbol{w}) \right]$$

$$= \sum_{n=1}^N \mathbb{E}_q \left[\ln p(y_n | \boldsymbol{w}) \right] - \mathbb{E}_q \left[\ln \frac{q(\boldsymbol{w})}{p(\boldsymbol{w})} \right]$$

$$= \sum_{n=1}^N \mathbb{E}_q \left[\ln p(y_n | \boldsymbol{w}) \right] - \text{KL}[q(\boldsymbol{w}) | | p(\boldsymbol{w})]$$

■ The first term (expected log likelihood) is a *data-fit* term, while the KL term encourages q to be close to the prior p(w) (regularization term)



Free-form variational inference

■ Factorized approximation for approximating the target distribution $p \equiv p(w|y)$

$$q(\mathbf{w}) = \prod_{j=1}^J q(\mathbf{w}_j), \quad ext{where} \quad \mathbf{w} = [\mathbf{w}_1, \mathbf{w}_2, ..., \mathbf{w}_J]$$

Optimality condition for mean-field families

$$\ln q^*(\boldsymbol{w}_k) = \mathbb{E}_{i\neq k} \left[\ln p(\boldsymbol{y}, \boldsymbol{w}) \right] + K$$

- Strategy for free-form variational inference
 - 1. Deduce optimal form for each factor and derive parameters
 - 2. Derive fixed-point algorithm based on the optimal parameters
- + Optimal functional form given assumptions
- + Fast optimization

- Requires model-specific derivations
- Required integrals may be intractable
- Optimal forms may not be "known" distributions

Fixed-form variational inference

- **Fixed-form** variational inference as an alternative to *free-form*: We give up an the optimal functional form and simply *assume* some family of distributions q_{ψ} with parameters ψ
- \blacksquare We often refer to ψ as variational parameters
- **Example**: Consider a model p(y, w), where $w \in \mathbb{R}^D$
 - Full-rank Gaussians: $m \in \mathbb{R}^D$, $V \in \mathbb{R}^{D \times D}$ $g(w) = \mathcal{N}(w|m, V)$.
 - Low-rank Gaussians: $m \in \mathbb{R}^D$, $B \in \mathbb{R}^{D \times K}$, and $C \in \mathbb{R}^{D \times D}$ is diagonal. $q(w) = \mathcal{N}(w|m, BB^T + C^2)$,
 - Mean-field Gaussians: $m{m} = ig[m_1, \dots, m_Dig] \in \mathbb{R}^D$ and $m{v} = ig[v_1, \dots, v_Dig] = \mathbb{R}^D$ $q(m{w}) = \prod_{i=1}^D \mathcal{N}(w_i|m_i, v_i),$
- Non-restricted to Gaussians: Gamma, Gauss, Beta, Dirichlet, Bernouilli etc

Fixed-form variational inference II

- Fixed-form variational inference as an alternative to free-form
- **Example**: Consider a model p(y, w), where $w \in \mathbb{R}^D$

■ Suppose we choose a full-rank Gaussian family

$$q(\mathbf{w}) = \mathcal{N}(\mathbf{w}|\mathbf{m}, \mathbf{V}),$$

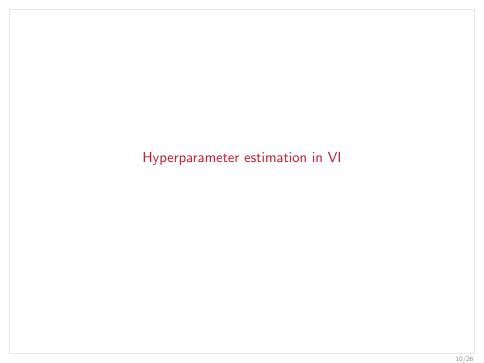
such that the variational family ${\mathcal Q}$ consists of all multivariate Gaussian distributions

lacksquare q $_{m{\psi}}$ is now parametrized by the variational parameters $m{\psi} = \{m{m}, m{V}\}$

$$\begin{split} \mathcal{L}\left[q_{\boldsymbol{\psi}}\right] &= \mathbb{E}_{q_{\boldsymbol{\psi}}}\left[\ln p(\boldsymbol{y}, \boldsymbol{w})\right] - \mathbb{E}_{q_{\boldsymbol{\psi}}}\left[\ln q_{\boldsymbol{\psi}}(\boldsymbol{w})\right] \\ &= \mathbb{E}_{\mathcal{N}(\boldsymbol{w}|\boldsymbol{m}, \boldsymbol{V})}\left[\ln p(\boldsymbol{y}, \boldsymbol{w})\right] - \mathbb{E}_{\mathcal{N}(\boldsymbol{w}|\boldsymbol{m}, \boldsymbol{V})}\left[\ln \mathcal{N}(\boldsymbol{w}|\boldsymbol{m}, \boldsymbol{V})\right] \end{split}$$

■ Fitting the approximation using gradient-based methods

$$\boldsymbol{q}^* = \arg\max_{\boldsymbol{q} \in \mathcal{O}} \mathcal{L}\left[\boldsymbol{q}\right] \quad \Longleftrightarrow \quad \boldsymbol{\psi}^* = \arg\max_{\boldsymbol{\psi}} \mathcal{L}\left[\boldsymbol{q}_{\boldsymbol{\psi}}\right] \quad \Longleftrightarrow \quad \boldsymbol{m}^*, \boldsymbol{V}^* = \arg\max_{\boldsymbol{m}, \boldsymbol{V}} \mathcal{L}\left[\boldsymbol{q}_{\boldsymbol{\psi}}\right]$$



Dealing with hyperparameters for variational inference

 \blacksquare Consider a model with data y, parameters θ , and hyperparameters ξ

$$p(\theta|\mathbf{y}, \boldsymbol{\xi}) = \frac{p(\mathbf{y}|\theta, \boldsymbol{\xi})p(\theta|\boldsymbol{\xi})}{p(\mathbf{y}|\boldsymbol{\xi})}$$

Examples

1. Linear regression

 $\theta = \mathbf{w}$ would be the regression weights $\boldsymbol{\xi} = \{\alpha, \beta\}$ would be prior and noise precision.

2. Gaussian process regression

 $\begin{aligned} \boldsymbol{\theta} &= \boldsymbol{f} \text{ would be the latent function values} \\ \boldsymbol{\xi} &= \left\{ \sigma^2, \kappa, \ell \right\} \text{ would be noise variance and kernel hyperparameters} \end{aligned}$

■ Hyperparameter estimation via the marginal likelihood (MLII/MAPII)

$$\hat{oldsymbol{\xi}} = rg \max_{oldsymbol{\xi}} \log p(oldsymbol{y} | oldsymbol{\xi})$$

But what to do for non-conjugate model, where we cannot compute the marginal likelihood?

$$p(\theta|\mathbf{y},\boldsymbol{\xi}) \approx q(\theta)$$

Dealing with hyperparameters for variational inference

Consider a model with data y, parameters θ , and hyperparameters ξ

$$p(\theta|\mathbf{y},\boldsymbol{\xi}) = \frac{p(\mathbf{y}|\theta,\boldsymbol{\xi})p(\theta|\boldsymbol{\xi})}{p(\mathbf{y}|\boldsymbol{\xi})}$$

■ Variational approximation $p(\theta|\mathbf{y}, \boldsymbol{\xi}) \approx q(\theta)$ by optimizing the ELBO

$$\mathcal{L}_{oldsymbol{\xi}}\left[q
ight] = \mathbb{E}_{q}\left[\ln p(oldsymbol{y},oldsymbol{ heta}|oldsymbol{\xi})
ight] - \mathbb{E}_{q}\left[\ln q(oldsymbol{ heta})
ight]$$

■ The ELBO is a lowerbound on the log marginal likelihood

$$\log p(\mathbf{y}|\mathbf{\xi}) \geq \mathcal{L}_{\mathbf{\xi}}[q]$$

■ and hence, we can do hyperparameter estimation via ELBO

$$\hat{oldsymbol{\xi}} = rg \max_{oldsymbol{\xi}} p(oldsymbol{y} | oldsymbol{\xi}) pprox rg \max_{oldsymbol{\xi}} \mathcal{L}_{oldsymbol{\xi}}\left[q
ight]$$

Key take-away: We can optimize the ELBO wrt. *variational parameters* ψ and *hyperparameters* ξ simultaneously

$$\hat{\psi}^*, \hat{\pmb{\xi}}^* = \arg\max_{\pmb{\xi}} \mathcal{L}_{\pmb{\xi}} \left[q_{\pmb{\psi}} \right]$$

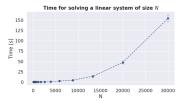


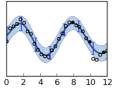
Speeding up Gaussian process inference

■ Gaussian process priors, $f \sim \mathcal{GP}(0, k(\mathbf{x}, \mathbf{x}'))$ can be extremely powerful, but scales poorly: $\mathcal{O}(N^3)$ in computational complexity and $\mathcal{O}(N^2)$ in memory footprint

$$\begin{split} p(y^*|\mathbf{y}) &= \mathcal{N}\left(y^*|\mu_{y^*|\mathbf{y}}, \sigma_{y^*|\mathbf{y}}^2\right) \\ \mu_{y^*|\mathbf{y}} &= \mathbf{k} \left(\mathbf{K} + \beta^{-1} \mathbf{I}\right)^{-1} \mathbf{y} \\ \sigma_{y^*|\mathbf{y}}^2 &= \mathbf{c} - \mathbf{k} \left(\mathbf{K} + \beta^{-1} \mathbf{I}\right)^{-1} \mathbf{k}^T \end{split}$$

- Obviously, using a subset of data would be faster, but can we do something more clever? Yes!
- We introduce *inducing points* z_i for i = 1, ..., M for $M \ll N$
- Combining *variational inference* with *inducing points* allows us to approximate the posterior much faster, i.e. $\mathcal{O}(NM^2)$, and even faster, $\mathcal{O}(M^2)$, using *mini-batching* (next week)





Hensman et al, 2019: Gaussian Processes for Big Data

Introducing inducing points

■ Suppose we have dataset $\mathcal{D} = \{x_n, y_n\}$, where $x_n \in \mathbb{R}^D$ and $y_n \in \mathbb{R}$ $y_n = f(x_n) + e_n$

■ Assuming Gaussian noise and a GP prior, i.e. $f(x) \sim \mathcal{GP}(\mathbf{0}, k(x, x'))$ yields the joint distribution

$$p(\mathbf{y}, \mathbf{f}) = p(\mathbf{y}|\mathbf{f})p(\mathbf{f}) = \mathcal{N}(\mathbf{y}|\mathbf{f}, \sigma^2\mathbf{I})\mathcal{N}(\mathbf{f}|\mathbf{0}, \mathbf{K}_{ff})$$

Goal: fast way compute p(f|y)

$$f_n = f(\mathbf{x}_n)$$

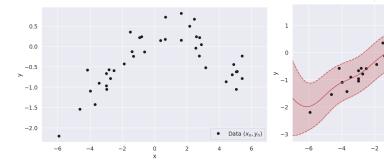
- We introduce M inducing points $z_i \in \mathbb{R}^D$ for m = 1, ..., M such that $u_i = f(z_i)$
- Extended joint distribution for data \mathbf{y} and latent function values for both \mathbf{f} and \mathbf{u} $p(\mathbf{y}, \mathbf{f}, \mathbf{u}) = p(\mathbf{y}|\mathbf{f})p(\mathbf{f}|\mathbf{u})p(\mathbf{u}) = \mathcal{N}(\mathbf{y}|\mathbf{f}, \sigma^2\mathbf{I})\mathcal{N}(\mathbf{f}|\mathbf{K}_{fu}\mathbf{K}_{uu}^{-1}\mathbf{u}, \mathbf{K}_{ff} \mathbf{K}_{fu}\mathbf{K}_{uu}^{-1}\mathbf{K}_{fu})\mathcal{N}(\mathbf{u}|\mathbf{0}, \mathbf{K}_{uu})$
- We can always marginalize to get the original model back

$$p(\mathbf{y}, \mathbf{f}) = \int p(\mathbf{y}, \mathbf{f}, \mathbf{u}) d\mathbf{u}$$

Intuition

Our extended joint distribution

$$p(\mathbf{y}, \mathbf{f}, \mathbf{u}) = p(\mathbf{y}|\mathbf{f})p(\mathbf{f}|\mathbf{u})p(\mathbf{u}) = \mathcal{N}(\mathbf{y}|\mathbf{f}, \sigma^2\mathbf{I})\mathcal{N}(\mathbf{f}|\mathbf{K}_{\mathrm{fu}}\mathbf{K}_{\mathrm{uu}}^{-1}\mathbf{u}, \mathbf{K}_{\mathrm{ff}} - \mathbf{K}_{\mathrm{fu}}\mathbf{K}_{\mathrm{uu}}^{-1}\mathbf{K}_{\mathrm{fu}})\mathcal{N}(\mathbf{u}|\mathbf{0}, \mathbf{K}_{\mathrm{uu}})$$



■ Can we find a set of *inducing points* $\{z_i\}_{i=1}^M$ and associated function values $u_i = f(z_i)$ such that we can absorb all the information from the full data set?

Setting up the approximation and choosing a variational family

Our extended joint distribution

$$\rho(\textbf{y}, \textbf{f}, \textbf{u}) = \rho(\textbf{y}|\textbf{f})\rho(\textbf{f}|\textbf{u})\rho(\textbf{u}) = \mathcal{N}(\textbf{y}|\textbf{f}, \sigma^2\textbf{I})\mathcal{N}(\textbf{f}|\textbf{K}_{\text{fu}}\textbf{K}_{\text{uu}}^{-1}\textbf{u}, \textbf{K}_{\text{ff}} - \textbf{K}_{\text{fu}}\textbf{K}_{\text{uu}}^{-1}\textbf{K}_{\text{fu}})\mathcal{N}(\textbf{u}|\textbf{0}, \textbf{K}_{\text{uu}})$$

- We will make a variational approximation: $p(f, u|y) \approx q(f, u)$
- A clever choice for the variational family

$$q(\mathbf{f}, \mathbf{u}) = p(\mathbf{f}|\mathbf{u})q(\mathbf{u}) = \mathcal{N}(\mathbf{f}|\mathbf{K}_{fu}\mathbf{K}_{uu}^{-1}\mathbf{u}, \mathbf{K}_{ff} - \mathbf{K}_{fu}\mathbf{K}_{uu}^{-1}\mathbf{K}_{fu})\mathcal{N}(\mathbf{u}|\mathbf{m}_{u}, \mathbf{S}_{u})$$

where

 $\mathbf{m}_u \in \mathbb{R}^M$ and $\mathbf{S}_u \in \mathbb{R}^{M \times M}$ are variational parameters to be estimated

 K_{ff} is the prior covariance matrix for f

 $K_{\mu\nu}$ is the prior covariance matrix for u

 K_{fu} is the prior covariance between f and u

Recall p(f|u) is the prior conditional distribution derived from the prior p(f, u).

The approximate posterior distribution

Our choice of variational family:

$$q(\mathbf{f}, \mathbf{u}) = p(\mathbf{f}|\mathbf{u})q(\mathbf{u}) = \mathcal{N}(\mathbf{f}|\mathbf{K}_{fu}\mathbf{K}_{uu}^{-1}\mathbf{u}, \mathbf{K}_{ff} - \mathbf{K}_{fu}\mathbf{K}_{uu}^{-1}\mathbf{K}_{fu})\mathcal{N}(\mathbf{u}|\mathbf{m}_{u}, \mathbf{S}_{u})$$

implies a marginal variational approximation for q(f) (linear Gaussian system)

$$q(f) = \int p(f|u)q(u)du$$

$$= \int \mathcal{N}(f|K_{fu}K_{uu}^{-1}u, K_{ff} - K_{fu}K_{uu}^{-1}K_{fu})\mathcal{N}(u|m_u, S_u)du$$

$$= \mathcal{N}(f|\underbrace{K_{fu}K_{uu}^{-1}m_u}_{m_f}, \underbrace{K_{ff} + K_{fu}K_{uu}^{-1}(S_u - K_{uu})K_{uu}^{-1}K_{uf}}_{S_f})$$

$$= \mathcal{N}(f|m_f, S_f)$$

 \blacksquare ... and similarly we can make predictions for new input points to get $f^* = f(x^*)$

$$p(f^*|\mathbf{y}) = \int p(f^*|\mathbf{u})p(\mathbf{u}|\mathbf{y})d\mathbf{u}$$

$$\approx \int p(f^*|\mathbf{u})q(\mathbf{u})d\mathbf{u}$$

$$= \mathcal{N}(f^*|\underbrace{\mathbf{K}_{f^*u}\mathbf{K}_{uu}^{-1}\mathbf{m}_u}_{m_{f^*}},\underbrace{\mathbf{K}_{f^*f^*}+\mathbf{K}_{f^*u}\mathbf{K}_{uu}^{-1}(\mathbf{S}_u - \mathbf{K}_{uu})\mathbf{K}_{uu}^{-1}\mathbf{K}_{uf^*}}_{\mathbf{S}_{f^*S}})$$

lacksquare So all we need to do is to estimate the mean and covariance for $q(oldsymbol{u}) = \mathcal{N}(oldsymbol{u} | oldsymbol{m}_u, oldsymbol{S}_u)$

Calculating the ELBO

We substitute our model and variational approximation into the ELBO

$$\mathcal{L}\left[q\right] = \mathbb{E}_{p(f|u)q(u)}\left[\log p(y, f, u)\right] - \mathbb{E}_{p(f|u)q(u)}\left[\log q(f, u)\right]$$

$$= \mathbb{E}_{p(f|u)q(u)}\left[\log p(y|f)p(f|u)p(u)\right] - \mathbb{E}_{p(f|u)q(u)}\left[\log p(f|u)q(u)\right]$$

$$= \mathbb{E}_{p(f|u)q(u)}\left[\log p(y|f)\right] + \mathbb{E}_{p(f|u)q(u)}\left[\log p(f|u)\right] + \mathbb{E}_{p(f|u)q(u)}\left[\log p(u)\right]$$

$$- \mathbb{E}_{p(f|u)q(u)}\left[\log p(f|u)\right] - \mathbb{E}_{p(f|u)q(u)}\left[\log q(u)\right]$$

$$= \mathbb{E}_{p(f|u)q(u)}\left[\log p(y|f)\right] + \mathbb{E}_{p(f|u)q(u)}\left[\log p(u)\right] - \mathbb{E}_{p(f|u)q(u)}\left[\log q(u)\right]$$

$$= \mathbb{E}_{p(f|u)q(u)}\left[\log p(y|f)\right] + \mathbb{E}_{q(u)}\left[\log p(u)\right] - \mathbb{E}_{q(u)}\left[\log q(u)\right]$$

$$= \mathbb{E}_{p(f|u)q(u)}\left[\log p(y|f)\right] - \mathsf{KL}\left[q(u)||p(u)\right]$$

Next, recall that
$$p(\mathbf{y}|\mathbf{f}) = \mathcal{N}(\mathbf{y}|\mathbf{f}, \sigma^2\mathbf{I}) = \prod_{n=1}^{N} \mathcal{N}(\mathbf{y}_n|f_n, \sigma^2)$$

$$\mathcal{L}[q] = \sum_{n=1}^{N} \mathbb{E}_{p(f|\boldsymbol{u})q(\boldsymbol{u})} \left[\log \mathcal{N}(y_n|f_n, \sigma^2) \right] - \mathsf{KL}[q(\boldsymbol{u})||p(\boldsymbol{u})]$$
$$= \sum_{n=1}^{N} \mathbb{E}_{p(f_n|\boldsymbol{u})q(\boldsymbol{u})} \left[\log \mathcal{N}(y_n|f_n, \sigma^2) \right] - \mathsf{KL}[q(\boldsymbol{u})||p(\boldsymbol{u})]$$

Calculating the first term

The KI term

■ The two terms in the lowerbound

$$\mathcal{L}[q] = \sum_{n=1}^{N} \mathbb{E}_{p(f_n|\boldsymbol{u})q(\boldsymbol{u})} \left[\log \mathcal{N}(y_n|f_n, \sigma^2) \right] - \mathsf{KL}\left[q(\boldsymbol{u})||p(\boldsymbol{u})\right]$$

The KL divergence between two multivariate Gaussians can be computed in closed-form:

$$\begin{split} \mathsf{KL}\left[\mathcal{N}(\pmb{u}|\pmb{m}_0,\pmb{\Sigma}_0)||\mathcal{N}(\pmb{u}|\pmb{m}_1,\pmb{\Sigma}_1)\right] \\ &= \frac{1}{2}\left[\mathsf{trace}(\pmb{\Sigma}_1^{-1}\pmb{\Sigma}_0) + (\pmb{\mu}_1 - \pmb{\mu}_0)^T\pmb{\Sigma}_1^{-1}(\pmb{\mu}_1 - \pmb{\mu}_0) - D + \log\frac{|\pmb{\Sigma}_1|}{|\pmb{\Sigma}_0|}\right] \end{split}$$

■ So for $q(u) = \mathcal{N}(u|m_u, S_u)$ and $p(u) = \mathcal{N}(u|0, K_{uu})$, we get

$$\mathsf{KL}\left[q(\boldsymbol{u})||p(\boldsymbol{u})\right] = \frac{1}{2}\left[\mathsf{trace}(\boldsymbol{K}_{uu}^{-1}\boldsymbol{S}_{u}) + \boldsymbol{m}_{u}^{T}\boldsymbol{K}_{uu}^{-1}\boldsymbol{m}_{u} - \boldsymbol{M} + \log\frac{|\boldsymbol{K}_{uu}|}{|\boldsymbol{S}_{u}|}\right]$$

Combining everything

We derived

$$\mathcal{L}[q] = \sum_{n=1}^{N} \mathbb{E}_{p(f_n|\boldsymbol{u})q(\boldsymbol{u})} \left[\log \mathcal{N}(y_n|f_n, \sigma^2) \right] - \mathsf{KL}[q(\boldsymbol{u})||p(\boldsymbol{u})]$$

■ ... and then showed that the first term simplifies to

$$\mathbb{E}_{p(f_n|\boldsymbol{u})q(\boldsymbol{u})}\left[\log \mathcal{N}(y_n|f_n,\sigma^2)\right] = \log \mathcal{N}(y_n|m_{f_n},\sigma^2) \mathrm{d}f_n - \frac{1}{2\sigma}\sigma_{f_n}^2$$

■ Combining yields

$$\begin{split} \mathcal{L}\left[q\right] &= \sum_{n=1}^{N} \left[\log \mathcal{N}(y_n | m_{f_n}, \sigma^2) \mathrm{d}f_n - \frac{1}{2\sigma} \sigma_{f_n}^2 \right] - \mathsf{KL}\left[q(\boldsymbol{u}) || \rho(\boldsymbol{u})\right] \\ &= \sum_{n=1}^{N} \log \mathcal{N}(y_n | m_{f_n}, \sigma^2) - \frac{1}{2\sigma} \sum_{n=1}^{N} \sigma_{f_n}^2 - \mathsf{KL}\left[q(\boldsymbol{u}) || \rho(\boldsymbol{u})\right] \\ &= \log \mathcal{N}(\boldsymbol{y} | m_f, \sigma^2 \boldsymbol{I}) - \frac{1}{2\sigma} \mathsf{trace}(\boldsymbol{S}_f) - \mathsf{KL}\left[q(\boldsymbol{u}) || \rho(\boldsymbol{u})\right] \end{split}$$

where

$$\mathbf{m}_f = \mathbf{K}_{fu} \mathbf{K}_{uu}^{-1} \mathbf{m}_u$$

$$\mathbf{S}_f = \mathbf{K}_{ff} + \mathbf{K}_{fu} \mathbf{K}_{uu}^{-1} (\mathbf{S}_u - \mathbf{K}_{uu}) \mathbf{K}_{uu}^{-1} \mathbf{K}_{uf}$$

■ We can optimize this bound wrt. the *variational parameters* \mathbf{m}_u and \mathbf{S}_u and wrt. the *hyperparameters* simultaneously!

One step more step

■ It turns out we can optimize the bound wrt. m_u and S_u analytically

$$\begin{split} \mathcal{L}\left[q\right] &= \sum_{n=1}^{N} \left[\log \mathcal{N}(y_n|m_{f_n},\sigma^2) \mathrm{d}f_n - \frac{1}{2\sigma} \sigma_{f_n}^2\right] - \mathsf{KL}\left[q(\boldsymbol{u})||p(\boldsymbol{u})\right] \\ &= \sum_{n=1}^{N} \log \mathcal{N}(y_n|m_{f_n},\sigma^2) - \frac{1}{2\sigma} \sum_{n=1}^{N} \sigma_{f_n}^2 - \mathsf{KL}\left[q(\boldsymbol{u})||p(\boldsymbol{u})\right] \\ &= \log \mathcal{N}(\boldsymbol{y}|\boldsymbol{m}_f,\sigma^2\boldsymbol{I}) - \frac{1}{2\sigma} \mathsf{trace}(\boldsymbol{S}_f) - \mathsf{KL}\left[q(\boldsymbol{u})||p(\boldsymbol{u})\right] \end{split}$$

■ ... to get

$$egin{aligned} oldsymbol{S}_u^{-1} &= rac{1}{\sigma^2} oldsymbol{K}_{uu}^{-1} oldsymbol{K}_{uf} oldsymbol{K}_{fu} oldsymbol{K}_{uu}^{-1} + oldsymbol{K}_{uu}^{-1} \ oldsymbol{m}_u &= rac{1}{\sigma^2} oldsymbol{S}_u oldsymbol{K}_{uu}^{-1} oldsymbol{K}_{uf} oldsymbol{y} \end{aligned}$$

... which leads to the collapsed lowerbound

$$\mathcal{L}[q] = \log \mathcal{N}(\mathbf{y}|\mathbf{0}, \mathbf{K}_{fu}\mathbf{K}_{uu}\mathbf{K}_{uf} + \sigma^2\mathbf{I}) - \frac{1}{2\sigma} \text{trace}(\mathbf{K}_{ff} - \mathbf{K}_{fu}\mathbf{K}_{uu}^{-1}\mathbf{K}_{uf})$$

The big picture and how to use it in practice

Goal: fast way compute p(f|y)

- 1. Choose a set of *inducing points* z_i , where $u_i = f(z_i)$
- 2. Optimize the *collapsed bound* wrt. our hyperparameters θ

$$\hat{\boldsymbol{\theta}}^* = \arg\max_{\boldsymbol{\theta}} \mathcal{L}_{\boldsymbol{\theta}}\left[\boldsymbol{q}\right] = \log\mathcal{N}(\boldsymbol{y}|\boldsymbol{0}, \boldsymbol{K}_{\mathit{fu}}\boldsymbol{K}_{\mathit{uu}}\boldsymbol{K}_{\mathit{uf}} + \sigma^2\boldsymbol{I}) - \frac{1}{2\sigma}\mathsf{trace}(\boldsymbol{K}_{\mathit{ff}} - \boldsymbol{K}_{\mathit{fu}}\boldsymbol{K}_{\mathit{uu}}^{-1}\boldsymbol{K}_{\mathit{uf}})$$

3. Compute the posterior mean and covariance for latent function values $q(\mathbf{u}) = \mathcal{N}(\mathbf{u}|\mathbf{m}_u, \mathbf{S}_u)$

$$egin{aligned} oldsymbol{S}_u^{-1} &= rac{1}{\sigma^2} oldsymbol{K}_{uu}^{-1} oldsymbol{K}_{uf} oldsymbol{K}_{fu} oldsymbol{K}_{uu}^{-1} + oldsymbol{K}_{uu}^{-1} \ oldsymbol{m}_u &= rac{1}{\sigma^2} oldsymbol{S}_u oldsymbol{K}_{uu}^{-1} oldsymbol{K}_{uf} oldsymbol{y} \end{aligned}$$

4. Compute the approximate posterior distribution for $q(f) = \mathcal{N}(f|m_f, S_f)$

$$\mathbf{m}_f = \mathbf{K}_{fu} \mathbf{K}_{uu}^{-1} \mathbf{m}_u$$

$$\mathbf{S}_f = \mathbf{K}_{ff} + \mathbf{K}_{fu} \mathbf{K}_{uu}^{-1} (\mathbf{S}_u - \mathbf{K}_{uu}) \mathbf{K}_{uu}^{-1} \mathbf{K}_{uf}$$

5. Use $p(f|y) \approx q(f)$ to make predictions

Airline delays dataset

- Flight arrival and departure times for every commercial flight in the USA from Jan. 2008 to April 2008
- 2 million flights: 700000 flight for training, 100000 for testing
- Target variable: Delay in minutes:
- Arr D=8 features: age of aircraft, flight distance, airtime, departure time, arrival time, day of the week, day of the month, month
- Squared exponential kernel with separate lengthscale $\ell_i > 0$ for each dimension

$$k(\mathbf{x}, \mathbf{x}') = \kappa^2 \exp \left[-\frac{1}{2} \sum_{i=1}^{D} \ell_i^{-1} |\mathbf{x}_i - \mathbf{x}_i'|^2 \right] + \tau$$

■ Using M = 1000 inducing points

Hensman et al: Gaussian Processes for Big Data (2013)

Airline delays dataset

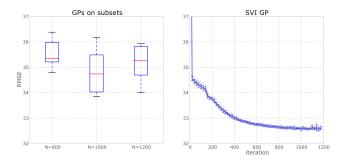


Figure 7: Root mean squared errors in predicting flight delays using information about the flight.

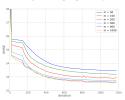


Figure 8: Root mean square errors for models with different numbers of inducing variables.