

### 02477 - Bayesian Machine Learning: Lecture 2

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#### Outline



- Quick re-cap of last week
- 2 Probabilistic machine learning
- 3 The plug-in approximation
- 4 Grid approximations for non-conjugate models
- 5 Introduction to exercise: towards logistic regression

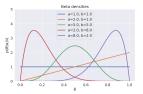


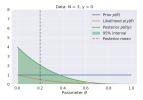
### Quick re-cap of last week



 Bayes' rule provides a systematic way to combine data with prior knowledge

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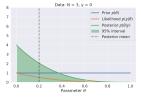
■ The beta-binomial model is a conjugate model

$$p(\theta) = \text{Beta}(\theta|a_0, b_0)$$
 (Prior)

$$p(y|\theta) = {N \choose y} \theta^{y} (1-\theta)^{N-y}$$
 (Likelihood)

$$p(\theta|y) = \text{Beta}(\theta|y + a_0, N - y + b_0)$$
 (Posterior)







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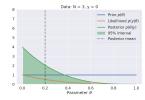
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**E**stimate  $\theta$  using the *mean* of the *posterior* distribution

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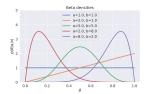
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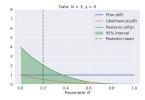
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 .. and use credibility intervals of the posterior to quantify the uncertainty

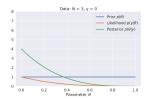
$$P(\theta \in [0.01, 0.60] | y) = 0.95$$







**Example continued**: suppose we have this website ad with N=3 views and y=0 clicks.





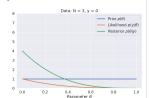


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(Prior)

■ Using a *uniform prior*, i.e.  $p(\theta) = \text{Beta}(\theta|1,1)$ 

$$\begin{split} \rho(\theta) &= \mathsf{Beta}(\theta|1,1) & (\textit{Prior}) \\ \rho(y|\theta) &= {3 \choose 0} \theta^0 (1-\theta)^3 & (\textit{Likelihood}) \\ \rho(\theta|y) &= \mathsf{Beta}(\theta|1,4) & (\textit{Posterior}) \end{split}$$







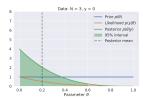
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■ Summarize our knowledge using posterior

$$\mathbb{E}[\theta|y] = \frac{1}{5}, \qquad P(\theta \in [0.01, 0.60]|y) \approx 0.95$$





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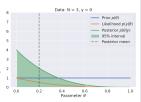
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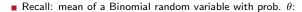
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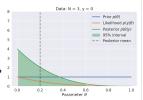
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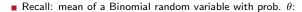
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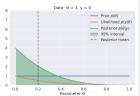
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Which value of  $\theta$  to use?





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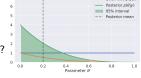
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Data: N = 3, v = 0

**Recall**: mean of a Binomial random variable with prob.  $\theta$ :

$$\mathbb{E}\left[\mathbf{v}^*\right] = \mathbf{N}^*\theta = 50\theta$$

Which value of  $\theta$  to use? How do we use the *posterior knowledge* to make predictions?



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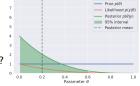
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Which value of  $\theta$  to use? How do we use the *posterior knowledge* to make predictions? How do we take the *uncertainty* into account?



# Probabilistic machine learning





| Product rule         | Sum rule   | Conditional   | Conditional independence  |
|----------------------|--|---|---|
| p(a, b) = p(b a)p(a) | $p({m b}) = \int p({m a},{m b}) \mathrm{d}{m a}$ | $p(\boldsymbol{a} \boldsymbol{b}) = \frac{p(\boldsymbol{a},\boldsymbol{b})}{p(\boldsymbol{b})}$ | $p(\boldsymbol{a}, \boldsymbol{b} \boldsymbol{c}) = p(\boldsymbol{a} \boldsymbol{c})p(\boldsymbol{b} \boldsymbol{c})$ |

- A probabilistic model is *completely specified* by its *joint distribution*
- Consider a model with two *random variables*: y (data) and  $\theta$  (unknown parameter)





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| p(a, b) =    | $p(\mathbf{b} \mathbf{a})p(\mathbf{a})$ |  |  |  |

Sum rule 
$$p(b) = \int p(a, b) da$$

Conditional 
$$p(a|b) = \frac{p(a,b)}{p(b)}$$

Conditional independence 
$$p(\mathbf{a}, \mathbf{b}|\mathbf{c}) = p(\mathbf{a}|\mathbf{c})p(\mathbf{b}|\mathbf{c})$$

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- The joint distribution of all random variables can be expressed via the product rule

$$p(\theta, y) = p(y|\theta)p(\theta)$$





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### Probabilistic machine learning I



Product rule p(a, b) = p(b|a)p(a)

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■ The evidence p(y) can be obtained from the joint distribution via the sum rule

$$p(y) = \int p(y,\theta) d\theta = \int p(y|\theta)p(\theta) d\theta$$

■ Hence, in theory, we can derive all quantities of interest from the joint distribution

### Probabilistic machine learning II



Product rule p(a, b) = p(b|a)p(a)

Sum rule

 $p(\mathbf{b}) = \int p(\mathbf{a}, \mathbf{b}) d\mathbf{a}$ 

Conditional  $p(a|b) = \frac{p(a,b)}{p(b)}$ 

Conditional independence  $p(\mathbf{a}, \mathbf{b}|\mathbf{c}) = p(\mathbf{a}|\mathbf{c})p(\mathbf{b}|\mathbf{c})$ 

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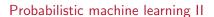
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■ A probabilistic model is *completely specified* by its *joint distribution* 

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■ What if we have more than one observed variable, e.g.  $y_1$  and  $y_2$ ?





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- What if we have more than one observed variable, e.g.  $y_1$  and  $y_2$ ?
- For a broad class of models the likelihood can be further decomposed using conditional independence:

$$p(y_1, y_2|\theta) = p(y_1|\theta)p(y_2|\theta)$$

### Probabilistic machine learning II



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lacksquare ... or more generally for  $m{y} = \begin{bmatrix} y_1 & y_2 & \dots & y_N \end{bmatrix}$ 

$$p(\mathbf{y}|\theta) = p(y_1|\theta)p(y_2|\theta)\dots p(y_N|\theta) = \prod_{n=1}^N p(y_n|\theta)$$

### Probabilistic machine learning II



Product rule p(a, b) = p(b|a)p(a) Sum rule

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A probabilistic model is completely specified by its joint distribution

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■ The *ioint distribution* becomes

$$p(\theta, \mathbf{y}) = p(\mathbf{y}|\theta)p(\theta) = \prod_{n=1}^{N} p(y_n|\theta)p(\theta)$$





■ Example continued: Your website ad has been shown N=123 times and generated y=12 clicks. Suppose you pay for another  $N^*=50$  views, how many clicks  $y^*$  should you expect *given the observed data*?

### Website ad example continued: Making predictions



- **Example continued:** Your website ad has been shown N = 123 times and generated y = 12 clicks. Suppose you pay for another  $N^* = 50$  views, how many clicks  $y^*$  should you expect *given the observed data*?
- lacktriangled Assuming each click can be modelled using *conditionally independent* Bernoulli trials with the *same probability* heta

$$p(y|\theta) = Bin(y|N,\theta)$$

# Website ad example continued: Making predictions



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■ The assumption of *conditional independence* implies

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$$p(y, y^*|\theta) = p(y|\theta)p(y^*|\theta) = Bin(y|N, \theta)Bin(y|N^*, \theta)$$





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 $\blacksquare$  Completing the model by *imposing a prior* for  $\theta$ 

$$p(\theta) = \text{Beta}(\theta|a_0, b_0)$$

■ Goal: compute *predictive distribution* of  $y^*$  given we have observed y = 12, i.e.  $p(y^*|y = 12)$ .

### A probabilistic perspective on making predictions



Product rule  $p(\pmb{a},\pmb{b}) = p(\pmb{b}|\pmb{a})p(\pmb{a}) \qquad \qquad p(\pmb{b}) = \int p(\pmb{a},\pmb{b})\mathrm{d}\pmb{a}$ 

Conditional  $p(a|b) = \frac{p(a,b)}{p(b)}$ 

Conditional independence  $p(\mathbf{a}, \mathbf{b}|\mathbf{c}) = p(\mathbf{a}|\mathbf{c})p(\mathbf{b}|\mathbf{c})$ 

**Goal**: Given some data y, what can we say about a new observation  $y^*$ ?





Product rule
$$p(a, b) = p(b|a)p(a)$$

Sum rule 
$$p(\mathbf{b}) = \int p(\mathbf{a}, \mathbf{b}) d\mathbf{a}$$

Conditional 
$$p(a|b) = \frac{p(a,b)}{p(b)}$$

Conditional independence 
$$p(\mathbf{a}, \mathbf{b}|\mathbf{c}) = p(\mathbf{a}|\mathbf{c})p(\mathbf{b}|\mathbf{c})$$

**Goal**: Given some data y, what can we say about a new observation  $y^*$ ?

■ Step 1: Formulate joint distribution for all variables of interests

$$p(y^*, y, \theta) = p(y^*, y|\theta)p(\theta) =$$





Product rule p(a, b) = p(b|a)p(a)

Sum rule  $p(\mathbf{b}) = \int p(\mathbf{a}, \mathbf{b}) d\mathbf{a}$ 

Conditional  $p(a|b) = \frac{p(a,b)}{p(b)}$ 

Conditional independence  $p(\mathbf{a}, \mathbf{b}|\mathbf{c}) = p(\mathbf{a}|\mathbf{c})p(\mathbf{b}|\mathbf{c})$ 

**Goal**: Given some data y, what can we say about a new observation  $y^*$ ?

■ Step 1: Formulate joint distribution for all variables of interests

$$p(y^*, y, \theta) = p(y^*, y|\theta)p(\theta) = p(y^*|\theta)p(y|\theta)p(\theta)$$





Product rule 
$$p(a, b) = p(b|a)p(a)$$

Sum rule 
$$p(b) = \int p(a, b) da$$

Conditional 
$$p(\mathbf{a}|\mathbf{b}) = \frac{p(\mathbf{a},\mathbf{b})}{p(\mathbf{b})}$$

Conditional independence 
$$p(\mathbf{a}, \mathbf{b}|\mathbf{c}) = p(\mathbf{a}|\mathbf{c})p(\mathbf{b}|\mathbf{c})$$

■ Step 1: Formulate joint distribution for all variables of interests

$$p(y^*, y, \theta) = p(y^*, y|\theta)p(\theta) = p(y^*|\theta)p(y|\theta)p(\theta)$$

■ Step 2: Condition on the observed data y

$$p(y^*,\theta|y) = \frac{p(y*,y,\theta)}{p(y)} =$$





Product rule 
$$p(a, b) = p(b|a)p(a)$$

Sum rule 
$$p(b) = \int p(a, b) da$$

Conditional 
$$p(\mathbf{a}|\mathbf{b}) = \frac{p(\mathbf{a},\mathbf{b})}{p(\mathbf{b})}$$

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$$p(y^*, y, \theta) = p(y^*, y|\theta)p(\theta) = p(y^*|\theta)p(y|\theta)p(\theta)$$

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$$p(y^*, \theta|y) = \frac{p(y^*, y, \theta)}{p(y)} = \frac{p(y^*|\theta)p(y|\theta)p(\theta)}{p(y)}$$





Product rule 
$$p(a, b) = p(b|a)p(a)$$

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$$p(y^*, \theta|y) = \frac{p(y*, y, \theta)}{p(y)} = \frac{p(y^*|\theta)p(y|\theta)p(\theta)}{p(y)}$$

$$p(y^*|y) = \int p(y^*, \theta|y) d\theta = \int \frac{p(y^*|\theta)p(y|\theta)p(\theta)}{p(y)} d\theta =$$





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$$p(\mathbf{a}|\mathbf{b}) = \frac{p(\mathbf{a},\mathbf{b})}{p(\mathbf{b})}$$

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$$p(a, b|c) = p(a|c)p(b|c)$$

■ Step 1: Formulate joint distribution for all variables of interests

$$p(y^*, y, \theta) = p(y^*, y|\theta)p(\theta) = p(y^*|\theta)p(y|\theta)p(\theta)$$

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$$p(y^*, \theta|y) = \frac{p(y*, y, \theta)}{p(y)} = \frac{p(y^*|\theta)p(y|\theta)p(\theta)}{p(y)}$$

$$p(y^*|y) = \int p(y^*, \theta|y) d\theta = \int \frac{p(y^*|\theta)p(y|\theta)p(\theta)}{p(y)} d\theta = \int p(y^*|\theta)p(\theta|y) d\theta$$





Product rule 
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$$p(y^*|y) = \int p(y^*, \theta|y) d\theta = \int \frac{p(y^*|\theta)p(y|\theta)p(\theta)}{p(y)} d\theta = \int p(y^*|\theta)p(\theta|y) d\theta = \mathbb{E}_{p(\theta|y)} \left[ p(y^*|\theta) \right]$$

## A probabilistic perspective on making predictions



Product rule 
$$p(a, b) = p(b|a)p(a)$$

Sum rule 
$$p(\mathbf{b}) = \int p(\mathbf{a}, \mathbf{b}) d\mathbf{a}$$

Conditional 
$$p(\mathbf{a}|\mathbf{b}) = \frac{p(\mathbf{a},\mathbf{b})}{p(\mathbf{b})}$$

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$$p(\mathbf{a}, \mathbf{b}|\mathbf{c}) = p(\mathbf{a}|\mathbf{c})p(\mathbf{b}|\mathbf{c})$$

**Goal**: Given some data y, what can we say about a new observation  $y^*$ ?

■ Step 1: Formulate joint distribution for all variables of interests

$$p(y^*, y, \theta) = p(y^*, y|\theta)p(\theta) = p(y^*|\theta)p(y|\theta)p(\theta)$$

■ Step 2: Condition on the observed data y

$$p(y^*, \theta|y) = \frac{p(y^*, y, \theta)}{p(y)} = \frac{p(y^*|\theta)p(y|\theta)p(\theta)}{p(y)}$$

 $\blacksquare$  Step 3: Marginalize out parameter  $\theta$  using the sum rule to get the posterior predictive distribution

$$p(y^*|y) = \int p(y^*, \theta|y) d\theta = \int \frac{p(y^*|\theta)p(y|\theta)p(\theta)}{p(y)} d\theta = \int p(y^*|\theta)p(\theta|y) d\theta = \mathbb{E}_{p(\theta|y)} \left[ p(y^*|\theta) \right]$$

■ **Key take-away**: To reason about  $y^*$  given y, we need to average the likelihood for  $y^*$  wrt. to the posterior distribution  $p(\theta|y)$ .

#### Quiz time



Take the quiz called Lecture 2: Prior, likelihood, posterior, posterior predictive to test your understanding.



■ Example continued: Your website ad has been shown N=123 times and generated y=12 clicks. Suppose you pay for another  $N^*=50$  views, how many clicks  $y^*$  should you expect *given the observed data*?



- Example continued: Your website ad has been shown N = 123 times and generated y = 12 clicks. Suppose you pay for another  $N^* = 50$  views, how many clicks  $y^*$  should you expect given the observed data?
- We already defined the model

$$p(y|\theta) = \text{Bin}(y|N,\theta)$$
 (Likelihood)  
 $p(y^*|\theta) = \text{Bin}(y^*|N^*,\theta)$  (Predictive likelihood)  
 $p(\theta) = \text{Beta}(\theta|a_0,b_0)$  (Prior)



- **Example continued:** Your website ad has been shown N = 123 times and generated y = 12 clicks. Suppose you pay for another  $N^* = 50$  views, how many clicks  $y^*$  should you expect *given the observed data*?
- We already defined the model

$$p(y|\theta) = \text{Bin}(y|N,\theta)$$
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 $p(\theta) = \text{Beta}(\theta|a_0,b_0)$  (Prior)

■ We know how to compute the *posterior distribution* 

$$p(\theta|y) = \text{Beta}(\theta|y + a_0, N - y + b_0)$$



- **Example continued:** Your website ad has been shown N = 123 times and generated y = 12 clicks. Suppose you pay for another  $N^* = 50$  views, how many clicks  $y^*$  should you expect *given the observed data*?
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■ We know how to compute the *posterior distribution* 

$$p(\theta|y) = \text{Beta}(\theta|y + a_0, N - y + b_0)$$

■ Next, we want to compute the *posterior predictive distribution* 

$$p(y^*|y) = \int p(y^*|\theta)p(\theta|y)d\theta =$$



- **Example continued:** Your website ad has been shown N = 123 times and generated y = 12 clicks. Suppose you pay for another  $N^* = 50$  views, how many clicks  $y^*$  should you expect *given the observed data*?
- We already defined the model

$$p(y|\theta) = \text{Bin}(y|N,\theta)$$
 (Likelihood)  
 $p(y^*|\theta) = \text{Bin}(y^*|N^*,\theta)$  (Predictive likelihood)  
 $p(\theta) = \text{Beta}(\theta|a_0,b_0)$  (Prior)

■ We know how to compute the *posterior distribution* 

$$p(\theta|y) = \text{Beta}(\theta|y + a_0, N - y + b_0)$$

■ Next, we want to compute the posterior predictive distribution

$$p(y^*|y) = \int p(y^*|\theta)p(\theta|y)d\theta = \int \mathsf{Bin}(y|N^*,\theta)\mathsf{Beta}(\theta|y+a_0,N-y+b_0)d\theta$$

■ Intuition: Instead of plugging in a single value for the parameter estimate, we plug in all possible values for  $\theta$  and weight the result according to  $p(\theta|y)$ 



$$p(y^* = k|y) = \int \mathsf{Bin}(y = k|N^*, \theta) \mathsf{Beta}(\theta|y + a_0, N - y + b_0) \mathsf{d}\theta$$



$$\begin{split} p(y^* = k|y) &= \int \mathsf{Bin}(y = k|N^*, \theta) \mathsf{Beta}(\theta|y + a_0, N - y + b_0) \mathsf{d}\theta \\ &= \int \binom{N^*}{k} \theta^k (1 - \theta)^{N^* - k} \frac{1}{B(\alpha, \beta)} \theta^{\alpha - 1} (1 - \theta)^{\beta - 1} \mathsf{d}\theta \end{split}$$



$$\begin{split} \rho(y^* = k|y) &= \int \mathsf{Bin}(y = k|N^*, \theta) \mathsf{Beta}(\theta|y + a_0, N - y + b_0) \mathrm{d}\theta \\ &= \int \binom{N^*}{k} \theta^k (1 - \theta)^{N^* - k} \frac{1}{B(\alpha, \beta)} \theta^{\alpha - 1} (1 - \theta)^{\beta - 1} \mathrm{d}\theta \\ &= \binom{N^*}{k} \frac{1}{B(\alpha, \beta)} \int \theta^k (1 - \theta)^{N^* - k} \theta^{\alpha - 1} (1 - \theta)^{\beta - 1} \mathrm{d}\theta \end{split} \quad \text{(Linearity)}$$





$$\begin{split} \rho(y^* = k|y) &= \int \mathsf{Bin}(y = k|N^*, \theta) \mathsf{Beta}(\theta|y + a_0, N - y + b_0) \mathsf{d}\theta \\ &= \int \binom{N^*}{k} \theta^k (1 - \theta)^{N^* - k} \frac{1}{B(\alpha, \beta)} \theta^{\alpha - 1} (1 - \theta)^{\beta - 1} \mathsf{d}\theta \\ &= \binom{N^*}{k} \frac{1}{B(\alpha, \beta)} \int \theta^k (1 - \theta)^{N^* - k} \theta^{\alpha - 1} (1 - \theta)^{\beta - 1} \mathsf{d}\theta \qquad \text{(Linearity)} \\ &= \binom{N^*}{k} \frac{1}{B(\alpha, \beta)} \int \theta^{k + \alpha - 1} (1 - \theta)^{\beta + N^* - k - 1} \mathsf{d}\theta \qquad \text{(Simplify)} \end{split}$$





$$\begin{split} \rho(y^* = k|y) &= \int \mathsf{Bin}(y = k|N^*, \theta) \mathsf{Beta}(\theta|y + a_0, N - y + b_0) \mathsf{d}\theta \\ &= \int \binom{N^*}{k} \theta^k (1 - \theta)^{N^* - k} \frac{1}{B(\alpha, \beta)} \theta^{\alpha - 1} (1 - \theta)^{\beta - 1} \mathsf{d}\theta \\ &= \binom{N^*}{k} \frac{1}{B(\alpha, \beta)} \int \theta^k (1 - \theta)^{N^* - k} \theta^{\alpha - 1} (1 - \theta)^{\beta - 1} \mathsf{d}\theta \qquad \text{(Linearity)} \\ &= \binom{N^*}{k} \frac{1}{B(\alpha, \beta)} \int \theta^{k + \alpha - 1} (1 - \theta)^{\beta + N^* - k - 1} \mathsf{d}\theta \qquad \text{(Simplify)} \\ &= \binom{N^*}{k} \frac{1}{B(\alpha, \beta)} \int \theta^{k + \alpha - 1} (1 - \theta)^{\beta + N^* - k - 1} \mathsf{d}\theta \end{split}$$



■ Compute posterior predictive distribution

$$\begin{split} \rho(y^* = k|y) &= \int \mathsf{Bin}(y = k|N^*, \theta) \mathsf{Beta}(\theta|y + a_0, N - y + b_0) \mathsf{d}\theta \\ &= \int \binom{N^*}{k} \theta^k (1 - \theta)^{N^* - k} \frac{1}{B(\alpha, \beta)} \theta^{\alpha - 1} (1 - \theta)^{\beta - 1} \mathsf{d}\theta \\ &= \binom{N^*}{k} \frac{1}{B(\alpha, \beta)} \int \theta^k (1 - \theta)^{N^* - k} \theta^{\alpha - 1} (1 - \theta)^{\beta - 1} \mathsf{d}\theta \qquad \text{(Linearity)} \\ &= \binom{N^*}{k} \frac{1}{B(\alpha, \beta)} \int \theta^{k + \alpha - 1} (1 - \theta)^{\beta + N^* - k - 1} \mathsf{d}\theta \qquad \text{(Simplify)} \\ &= \binom{N^*}{k} \frac{1}{B(\alpha, \beta)} \int \theta^{k + \alpha - 1} (1 - \theta)^{\beta + N^* - k - 1} \mathsf{d}\theta \end{split}$$

We recognize the terms in green as the functional form of a Beta density, and hence, we know how to compute the integral

$$p(y^* = k|y) = {N^* \choose k} \frac{B(\alpha + k, \beta + N^* - k)}{B(\alpha, \beta)}$$

for  $\alpha = v + a_0$  and  $\beta = N - v + b_0$ .





**Example continued:** Your website ad has been shown N = 123 times and generated y = 12 clicks. Suppose you pay for another  $N^* = 50$  views, how many clicks  $y^*$  should you expect *given the observed data*?

$$p(\theta) = \operatorname{Beta}(\theta|1,1)$$
 (Prior)  
 $p(y|\theta) = \operatorname{Bin}(y|123,\theta)$  (Likelihood)  
 $p(y^*|\theta) = \operatorname{Bin}(y^*|50,\theta)$  (Predictive likelihood)  
 $p(\theta|y) = \operatorname{Beta}(\theta|13,112)$  (Posterior)





**Example continued:** Your website ad has been shown N = 123 times and generated y = 12 clicks. Suppose you pay for another  $N^* = 50$  views, how many clicks  $y^*$  should you expect *given the observed data*?

$$\begin{aligned} p(\theta) &= \text{Beta}(\theta|1,1) & (\textit{Prior}) \\ p(y|\theta) &= \text{Bin}(y|123,\theta) & (\textit{Likelihood}) \\ p(y^*|\theta) &= \text{Bin}(y^*|50,\theta) & (\textit{Predictive likelihood}) \\ p(\theta|y) &= \text{Beta}(\theta|13,112) & (\textit{Posterior}) \end{aligned}$$

■ Distribution of clicks  $y^*$  based on views  $N^* = 50$  views

$$p(y^* = k|y) = {N^* \choose k} \frac{B(\alpha + k, \beta + N^* - k)}{B(\alpha, \beta)} =$$





**Example continued:** Your website ad has been shown N=123 times and generated y=12 clicks. Suppose you pay for another  $N^*=50$  views, how many clicks  $y^*$  should you expect *given the observed data*?

$$\begin{aligned} & p(\theta) = \text{Beta}(\theta|1,1) & (\textit{Prior}) \\ & p(y|\theta) = \text{Bin}(y|123,\theta) & (\textit{Likelihood}) \\ & p(y^*|\theta) = \text{Bin}(y^*|50,\theta) & (\textit{Predictive likelihood}) \\ & p(\theta|y) = \text{Beta}(\theta|13,112) & (\textit{Posterior}) \end{aligned}$$

■ Distribution of clicks  $y^*$  based on views  $N^* = 50$  views

$$p(y^* = k|y) = {N^* \choose k} \frac{B(\alpha + k, \beta + N^* - k)}{B(\alpha, \beta)} = {50 \choose k} \frac{B(13 + k, 162 - k)}{B(13, 112)}$$





**Example continued:** Your website ad has been shown N=123 times and generated y=12 clicks. Suppose you pay for another  $N^*=50$  views, how many clicks  $y^*$  should you expect *given the observed data*?

$$p(\theta) = \operatorname{Beta}(\theta|1,1) \qquad (Prior) \qquad 0.14$$

$$p(y|\theta) = \operatorname{Bin}(y|123,\theta) \qquad (Likelihood) \qquad 0.12$$

$$p(y^*|\theta) = \operatorname{Bin}(y^*|50,\theta) \qquad (Predictive \ likelihood) \stackrel{>}{\geq} 0.08$$

$$p(\theta|y) = \operatorname{Beta}(\theta|13,112) \qquad (Posterior) \qquad 0.04$$

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■ Distribution of clicks  $y^*$  based on views  $N^* = 50$  views

$$p(y^* = k|y) = {N^* \choose k} \frac{B(\alpha + k, \beta + N^* - k)}{B(\alpha, \beta)} = {50 \choose k} \frac{B(13 + k, 162 - k)}{B(13, 112)}$$





**Example continued:** Your website ad has been shown N = 123 times and generated y = 12 clicks. Suppose you pay for another  $N^* = 50$  views, how many clicks  $y^*$  should you expect *given the observed data*?

$$p(\theta) = \operatorname{Beta}(\theta|1,1) \qquad (Prior) = \begin{pmatrix} 0.16 \\ 0.14 \\ 0.14 \\ 0.14 \\ 0.14 \\ 0.14 \\ 0.15 \\ 0.16 \\ 0.16 \\ 0.16 \\ 0.18 \\ 0.18 \\ 0.18 \\ 0.18 \\ 0.19 \\$$

■ Distribution of clicks  $y^*$  based on views  $N^* = 50$  views

$$p(y^* = k|y) = {N^* \choose k} \frac{B(\alpha + k, \beta + N^* - k)}{B(\alpha, \beta)} = {50 \choose k} \frac{B(13 + k, 162 - k)}{B(13, 112)}$$

■ The expected number of clicks given the data is

$$\mathbb{E}_{p(y^*|y)}[y^*] = \sum_{k=0}^{50} kp(y^* = k|y) \approx 5.2$$



# The plug-in approximation

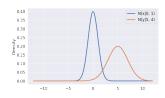


DTU

Univariate Gaussians

■ The *normal distribution* (also known as the Gaussian) is distribution over  $x \in \mathbb{R}$  with density

$$\mathcal{N}(x|\mu,\sigma^2) = rac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-rac{(x-\mu)^2}{2\sigma^2}
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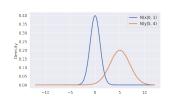


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 $\blacksquare$  Two parameters:  $\mu \equiv \mathbb{E}\left[x\right]$  and  $\sigma^2 \equiv \mathbb{V}\left[x\right]$ 







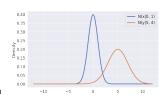
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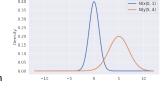


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■ We will talk more about Gaussians later in this course

## A few prerequisites first

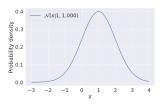
DTU

Dirac's delta function





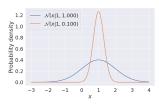
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DTU

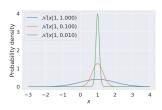
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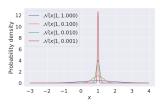
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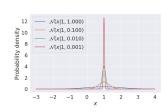
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DTU

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DTU

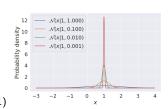
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$$\delta(x - \mu) = \begin{cases} \infty & \text{if } x = \mu \\ 0 & \text{otherwise} \end{cases}$$
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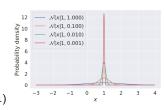
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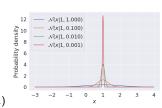
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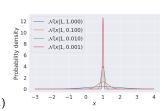
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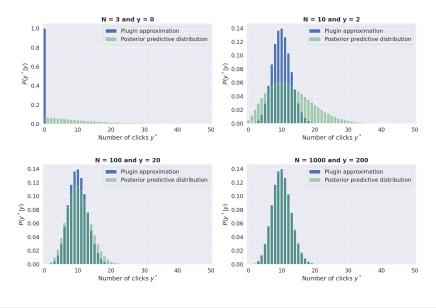
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- This is how we make predictions in deep learning...



### The posterior predictive distribution and plugin approximations





Grid approximations for non-conjugate models



Big picture so far...

- lacktriangle We studied the binomial model for estimating proportions and imposed a Beta prior for heta for Bayesian inference
- We derived the *posterior* and *posterior predictive* distributions *analytically*. This is possible due to *conjugacy* of the Beta prior and binomial likelihood



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■ To compute the posterior mean, variance etc. we need the evidence p(y)

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 Unfortunately, we cannot evaluate the evidence, i.e. p(y) analytically intractable for most models of practical interest...

## Approximate inference methods



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  - 1. Grid approximations
  - 2. Laplace approximations
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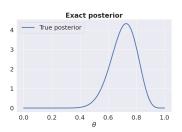
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- **This week**: We will focus on *grid approximations*, which are easy to understand and apply, and they will help build our intuition about marginalization.

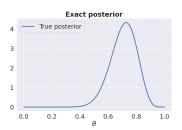


■ Constructing the grid approximation for  $p(\theta|y)$ 



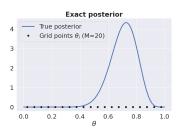


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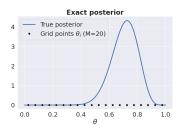
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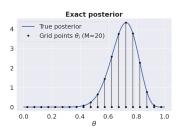
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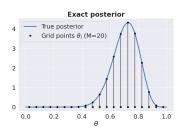




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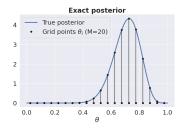


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- 4. Compute normalized probabilities  $\pi_i = \frac{1}{Z}\tilde{\pi}_i$  to get the grid approximation

$$q( heta) = \sum_{i=1}^{M} \pi_i \delta( heta - heta_i)$$



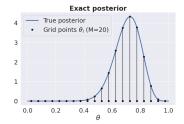


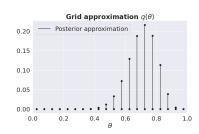
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- 4. Compute normalized probabilities  $\pi_i = \frac{1}{Z}\tilde{\pi}_i$  to get the grid approximation

$$q( heta) = \sum_{i=1}^{M} \pi_i \delta( heta - heta_i)$$





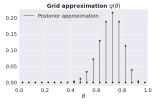




 The grid approximation is a discrete distribution, so computing summaries is easy, e.g the posterior mean

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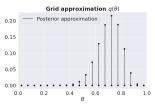




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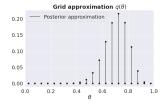
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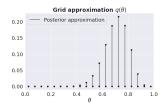
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■ Example: Computing post. probabilities for  $\theta < 0.6$ 

$$p(\theta < 0.6|y) \approx q(\theta < 0.6)$$

$$= \sum_{i=1}^{M} \mathbb{I} \left[\theta_{i} < 0.6\right] \pi_{i}$$

$$q( heta) = \sum_{i=1}^{M} \pi_i \delta( heta - heta_i)$$



Posterior summaries

 The grid approximation is a discrete distribution, so computing summaries is easy, e.g the posterior mean

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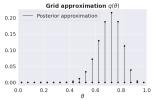
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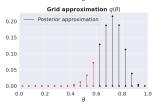
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$$= \sum_{i=1}^{M} \mathbb{I} \left[ \theta_j < 0.6 \right] \pi_j$$

$$= \sum_{i=1}^{j} \pi_i, \quad j = \max \left\{ i | \theta_i < 0.6 \right\}$$

$$q(\theta) = \sum_{i=1}^{M} \pi_i \delta(\theta - \theta_i)$$







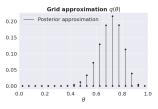


The posterior predictive distribution

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The posterior predictive distribution

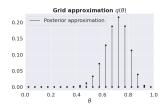
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DTU

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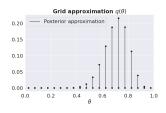
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DTU

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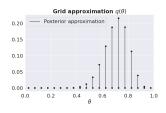
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# The grid approximation

DIU

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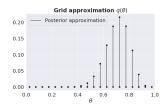
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■ To make predictions we literally compute a weighted sum of all possible parameter values

# The grid approximation

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A few practical considerations

- Choosing the grid range
  - Range for  $\theta \in [0, 1]$  is easy
  - For a different parameter  $\alpha \in \mathbb{R}$ , we need to choose an interval [a, b] for the grid. Often identified visually.
- Scaling with model dimensionality
  - Suppose we use M = 20 points for each dimension

1D: 20 evaluations

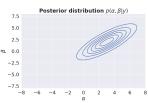
2D:  $20^2 = 400$  evaluations

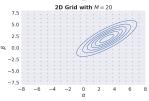
3D:  $20^3 = 8000$  evaluations

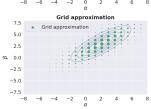
4D:  $16000^4 = 160000$  evaluations

- Grid approximations do not scale well beyond 3-4 dimensions
- Number of grid points M
  - M is balance between computational cost and accuracy
  - Grid approximation is zero when evaluated outside the grid points
  - Often diminishing returns as M increases (next slide)

Technical University of Denmark, DTU Compute, Department of Applied Math and Computer Science

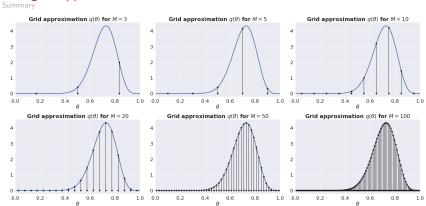






# The grid approximation





- Pros: Simple, easy and intuitive.
- Cons: Suffers from curse of dimensionality and does not scale beyond 3-4 dimensions



Introduction to exercise: towards logistic regression



- So far we focussed on modelling *proportions*, i.e.  $\theta \in [0,1]$  given data about y successes in N conditionally independent trials
- The binomial likelihood is also often used in *dose-response* models, which is key for determining "safe" dosages for drugs, pollution, foods etc.



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- Example: A company wants to study side effects of their new drug

| x (Dose in mg) | y (# side effects) | N (# patients) |
|----------------|--------------------|----------------|
| 80             | 0                  | 69             |
| 160            | 4                  | 832            |
| 320            | 13                 | 835            |
| 480            | 20                 | 459            |
| 640            | 12                 | 324            |
| 800            | 6                  | 103            |



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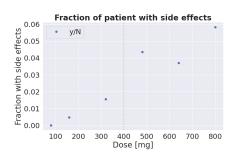
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- We could analyze the data for each dose independently using a beta-binomial model, but we would like to ...
  - 1. understand how dose affect the probability of side effects
  - 2. make predictions for new dosages  $x^*$
  - 3. borrow "statistical strength" across dosages



Example & motivation II

| x (Dose in mg) | y (# side effects) | N (# patients) | y/N   |
|----------------|--------------------|----------------|-------|
| 80             | 0                  | 69             | 0     |
| 160            | 4                  | 832            | 0.005 |
| 320            | 13                 | 835            | 0.016 |
| 480            | 20                 | 459            | 0.044 |
| 640            | 12                 | 324            | 0.037 |
| 800            | 6                  | 103            | 0.058 |



■ How accurate can we predict the probability of side effects for  $x^* = 400 mg$ ?

# Towards logistic regression Setting up the likelihood



For each dose  $x_i$ , we assume

$$y_i|x_i \sim \text{Bin}(y_i|N_i,\theta_i), \quad \theta_i \in [0,1]$$



Setting up the likelihood

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■ We model the probability  $\theta_i$  as function of the dose  $x_i$ , i.e.

$$\theta_i \equiv \theta(x_i) = \sigma(\alpha + \beta x),$$

where  $\sigma(x): \mathbb{R} \to [0,1]$  is a sigmoid function and  $\alpha, \beta \in \mathbb{R}$  are model parameters.



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Assuming conditional independence we can write the joint likelihood

$$p(\mathbf{y}|\mathbf{x},\alpha,\beta) = \prod_{i=1}^{M} p(y_i|x_i,\alpha,\beta) = \prod_{i=1}^{M} \text{Bin}(y_i|N_i,\theta_i),$$

where 
$$\mathbf{y} = [y_1, y_2, \dots y_6]$$
 and similar for  $\mathbf{x} = [x_1, x_2, \dots x_6]$ 



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■ The predictive likelihood for  $y^*$  is

$$p(y^*|x^*, \alpha, \beta) = Bin(y^*|N_i, \theta^*)$$

where  $\theta^* \equiv \theta(x^*)$ 



Setting up the prior

lacktriangle The model parameters are lpha (intercept) and eta (slope) of the generalized linear model



Setting up the prior

- $\blacksquare$  The model parameters are  $\alpha$  (intercept) and  $\beta$  (slope) of the generalized linear model
- Prior information: we have no prior information about the sign of the parameters. Hence, we choose a zero-mean Gaussian distributions

$$p(\alpha, \beta) = \mathcal{N}(\alpha|0, \sigma_{\alpha}^2)\mathcal{N}(\beta|0, \sigma_{\beta}^2), \qquad \sigma_{\alpha}^2, \sigma_{\beta}^2 > 0$$



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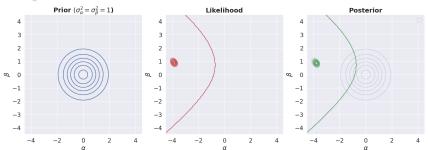
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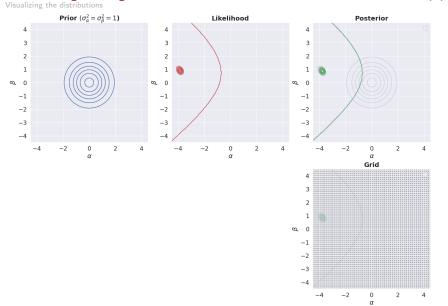
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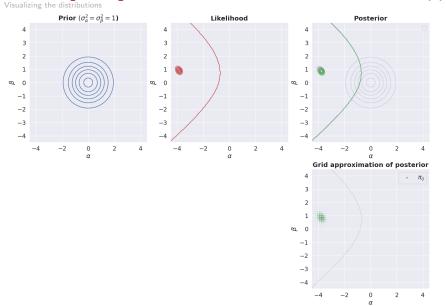
■ After obtaining the posterior of  $\alpha$ ,  $\beta$ , we can *propagate* the posterior uncertainty of the parameters to any quantity that depends on  $\alpha$ ,  $\beta$ , i.e.  $\theta(x) = \sigma(\alpha + \beta x)$ , the fraction of people with side effects  $y^*/N$  etc.

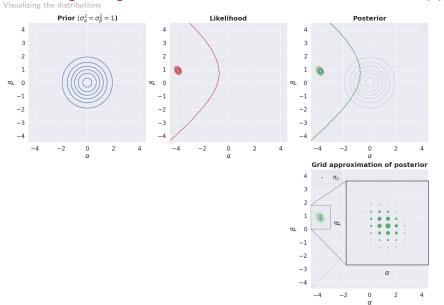


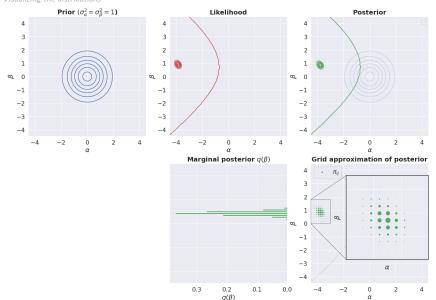
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Making predictions

$$p(y^* = k | \mathbf{y}, \mathbf{x}, x^*) = \iint \underbrace{p(y^* = k | x^*, \alpha, \beta)}_{\text{Burlined for } \mathbf{x}^*} \underbrace{p(\alpha, \beta | \mathbf{y}, \mathbf{x})}_{\text{posterior distribution}} d\alpha d\beta$$
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$$= \mathbb{E}_{p(\alpha, \beta | \mathbf{y}, \mathbf{x}, x^*)} [p(y^* = k | x^*, \alpha, \beta)] \qquad \text{(Integrals as expectation)}$$



Making predictions

$$\begin{split} \rho(y^* = k | \mathbf{y}, \mathbf{x}, x^*) &= \iint \underbrace{\rho(y^* = k | x^*, \alpha, \beta)}_{\text{likelihood for } y^*} \underbrace{\rho(\alpha, \beta | \mathbf{y}, \mathbf{x})}_{\text{posterior distribution}} \, \mathrm{d}\alpha \mathrm{d}\beta \qquad \text{(Sum rule)} \\ &= \mathbb{E}_{\rho(\alpha, \beta | \mathbf{y}, \mathbf{x}, x^*)} \left[ p(y^* = k | x^*, \alpha, \beta) \right] \qquad \text{(Integrals as expectation)} \\ &= \mathbb{E}_{\rho(\alpha, \beta | \mathbf{y}, \mathbf{x}, x^*)} \left[ \mathrm{Bin}(y^* | N^*, \theta^*) \right] \qquad \text{(Inserting dist.)} \end{split}$$



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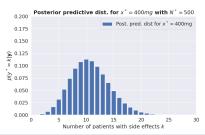
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Making predictions

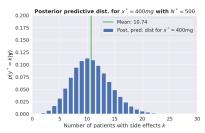
$$\begin{split} \rho(y^* = k | \mathbf{y}, \mathbf{x}, \mathbf{x}^*) &= \iint \underbrace{\rho(y^* = k | \mathbf{x}^*, \alpha, \beta)}_{\text{likelihood for } y^*} \underbrace{\rho(\alpha, \beta | \mathbf{y}, \mathbf{x})}_{\text{posterior distribution}} \text{d}\alpha\text{d}\beta \qquad \text{(Sum rule)} \\ &= \mathbb{E}_{\rho(\alpha, \beta | \mathbf{y}, \mathbf{x}, \mathbf{x}^*)} \left[ p(y^* = k | \mathbf{x}^*, \alpha, \beta) \right] \qquad \text{(Integrals as expectation)} \\ &= \mathbb{E}_{\rho(\alpha, \beta | \mathbf{y}, \mathbf{x}, \mathbf{x}^*)} \left[ \text{Bin}(y^* | N^*, \theta^*) \right] \qquad \text{(Inserting dist.)} \\ &\approx \mathbb{E}_{q(\alpha, \beta)} \left[ \text{Bin}(y^* | N^*, \theta^*) \right] \qquad \text{(Grid approx.)} \\ &= \sum_{i} \text{Bin}(y^* | N^*, \sigma(\alpha_i + \beta_j x^*)) \pi_{ij} \end{split}$$





Making predictions

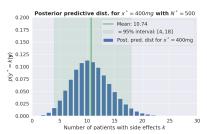
$$\begin{split} \rho(y^* = k | \mathbf{y}, \mathbf{x}, x^*) &= \int \int \underbrace{\rho(y^* = k | x^*, \alpha, \beta)}_{\text{likelihood for } y^*} \underbrace{\rho(\alpha, \beta | \mathbf{y}, \mathbf{x})}_{\text{posterior distribution}} \text{d}\alpha \text{d}\beta \qquad \text{(Sum rule)} \\ &= \mathbb{E}_{p(\alpha, \beta | \mathbf{y}, \mathbf{x}, x^*)} \left[ p(y^* = k | x^*, \alpha, \beta) \right] \qquad \text{(Integrals as expectation)} \\ &= \mathbb{E}_{p(\alpha, \beta | \mathbf{y}, \mathbf{x}, x^*)} \left[ \text{Bin}(y^* | N^*, \theta^*) \right] \qquad \text{(Inserting dist.)} \\ &\approx \mathbb{E}_{q(\alpha, \beta)} \left[ \text{Bin}(y^* | N^*, \theta^*) \right] \qquad \text{(Grid approx.)} \\ &= \sum_{i: j} \text{Bin}(y^* | N^*, \sigma(\alpha_i + \beta_j x^*)) \pi_{ij} \end{split}$$





Making predictions

$$\begin{split} \rho(y^* = k | \mathbf{y}, \mathbf{x}, \mathbf{x}^*) &= \iint \underbrace{\rho(y^* = k | \mathbf{x}^*, \alpha, \beta)}_{\text{likelihood for } y^*} \underbrace{\rho(\alpha, \beta | \mathbf{y}, \mathbf{x})}_{\text{posterior distribution}} \text{d}\alpha \text{d}\beta \qquad \text{(Sum rule)} \\ &= \mathbb{E}_{\rho(\alpha, \beta | \mathbf{y}, \mathbf{x}, \mathbf{x}^*)} \left[ p(y^* = k | \mathbf{x}^*, \alpha, \beta) \right] \qquad \text{(Integrals as expectation)} \\ &= \mathbb{E}_{\rho(\alpha, \beta | \mathbf{y}, \mathbf{x}, \mathbf{x}^*)} \left[ \text{Bin}(y^* | N^*, \theta^*) \right] \qquad \text{(Inserting dist.)} \\ &\approx \mathbb{E}_{q(\alpha, \beta)} \left[ \text{Bin}(y^* | N^*, \theta^*) \right] \qquad \text{(Grid approx.)} \\ &= \sum_{i, j} \text{Bin}(y^* | N^*, \sigma(\alpha_i + \beta_j x^*)) \pi_{ij} \end{split}$$



#### Intro to exercise



- On DTU Learn you will find an exercise for each week in notebook format
- We will spend all 4 hours from 13-17 working with the exercises
- In this exercise you will
  - Dive deeper into the Bayesian framework
  - Study and implement the probabilistic model for logistic regression for the Challenger Distaster dataset
  - Study and implement the grid approximations
  - Practice probabilistic reasoning
- Mix of pen&paper, programming and discussion questions
- Feel free to collaborate with your peers
- Ask for help!
  - Ask for help when stuck
  - Use teachers/TAs to check your understanding
  - Engage in discussion to practice
- Feedbacks persons: Meet at 16:45