- When easy is not preferred: A discounting paradigm to assess load-independent task preference
- Josephine Zerna<sup>†,1</sup>, Christoph Scheffel<sup>†,1</sup>, Corinna Kührt<sup>1</sup>, & Alexander Strobel<sup>1</sup>
- <sup>1</sup> Faculty of Psychology, Technische Universität Dresden, 01069 Dresden, Germany

# 5 Author Note

- The authors made the following contributions. Josephine Zerna: Conceptualization,
- <sup>7</sup> Methodology, Funding acquisition, Formal analysis, Investigation, Project administration,
- 8 Software, Visualization, Writing original draft preparation, Writing review & editing;
- 9 Christoph Scheffel: Conceptualization, Methodology, Funding acquisition, Investigation,
- Project administration, Software, Writing review & editing; Corinna Kührt: Formal
- analysis, Writing review & editing, Visualization; Alexander Strobel: Conceptualization,
- Funding acquistion, Writing review & editing. † Josephine Zerna and Christoph Scheffel
- 13 contributed equally to this work.
- 14 Correspondence concerning this article should be addressed to Josephine Zerna,
- <sup>15</sup> Zellescher Weg 17, 01069 Dresden, Germany. E-mail: josephine.zerna@tu-dresden.de

16 Abstract

When individuals set goals, they consider the subjective value (SV) of the anticipated reward and the required effort, a trade-off that is of great interest to psychological research. One approach to quantify the SVs of levels of a cognitive task is the Cognitive Effort Discounting Paradigm by Westbrook and colleagues (2013). However, it fails to acknowledge the highly subjective nature of effort, as it assumes a unidirectional, inverse relationship between task load and SVs. Therefore, it cannot map differences in effort perception that arise from traits like Need for Cognition, since individuals who enjoy effortful cognitive activities likely do not prefer the easiest level. We replicated the analysis

(CAD) Paradigm, which quantifies SVs without assuming that the easiest level is preferred, thereby enabling the quantification of SVs for tasks without objective order of task load.

of Westbrook and colleagues with our adaptation, the Cognitive and Affective Discounting

Keywords: effort discounting, registered report, specification curve analysis, need for cognition, n-back

30 Word count: 4,300

When easy is not preferred: A discounting paradigm to assess load-independent task preference

Introduction

33

In everyday life, effort and reward are closely intertwined<sup>1</sup>. With each decision a
person makes, they have to evaluate whether the effort required to reach a goal is worth
being exerted, given the reward they receive when reaching the goal. A reward is
subjectively more valuable if it is obtained with less effort, so the required effort is used as
a reference point for estimating the reward value<sup>1</sup>. However, the cost of the effort itself is
also subjective, and research has not yet established which function best describes the
relationship between effort and cost<sup>2</sup>. Investigating effort and cost is challenging because
"effort is not a property of the target task alone, but also a function of the individual's
cognitive capacities, as well as the degree of effort voluntarily mobilized for the task, which
in turn is a function of the individual's reward sensitivity" (p. 209)<sup>2</sup>.

One task that is often used to investigate effort is the *n*-back task, a working memory task in which a continuous stream of stimuli, e.g. letters, is presented on screen.

Participants indicate via button press whether the current stimulus is the same as *n* stimuli before, with *n* being the level of difficulty between one and six<sup>3</sup>. The *n*-back task is well suited to investigate effort because it is an almost continuous manipulation of task load as has been shown by monotonic increases in error rates, reaction times<sup>4</sup>, and brain activity in areas associated with working memory<sup>5,6</sup>. However, its reliability measures are mixed, and associations of *n*-back performance and measures such as executive functioning and fluid intelligence are often inconsistent<sup>4</sup>.

A way to quantify the subjective cost of each *n*-back level has been developed by
Westbrook, Kester, and Braver<sup>7</sup>, called the Cognitive Effort Discounting Paradigm
(COG-ED). First, the participants complete the *n*-back levels to familiarize themselves
with the task. Then, 1-back is compared with each more difficult level by asking the

participants to decide between receiving a fixed 2\$ for the more difficult level or the flexible starting value of 1\$ for 1-back. If they choose the more difficult level, the reward for 1-back increases by 0.50\$, if they choose 1-back, it decreases by 0.50\$. This is repeated five more times, with each adjustment of the 1-back reward being half of the previous step, while the reward for the more difficult level remains fixed at 2\$. The idea is to estimate the point of subjective equivalence, i.e., the monetary ratio at which both offers are equally preferred. The subjective value (SV) of each more difficult level is then calculated by dividing the final reward value of 1-back by the fixed 2\$ reward. Westbrook et al. used these SVs to investigate inter-individual differences in effort discounting. Younger participants showed lower effort discounting, i.e., they needed a lower monetary incentive for choosing the more difficult levels over 1-back.

The individual degree of effort discounting in the study by Westbrook et al. was also 68 associated with the participants' scores in Need for Cognition (NFC), a personality trait describing an individual's tendency to actively seek out and enjoy effortful cognitive 70 activities<sup>8</sup>. Westbrook et al.<sup>7</sup> conceptualized NFC as a trait measure of effortful task 71 engagement, providing a subjective self-report of effort discounting for each participant which could then be related to the SVs as an objective measure of effort discounting. On the surface, this association stands to reason, as individuals with higher NFC are more motivated to mobilize cognitive effort because they perceive it as intrinsically rewarding. Additionally, it has been shown that individuals avoid cognitive effort only to a certain degree, possibly to retain a sense of self-control<sup>9</sup>, a trait more prominent in individuals 77 with high NFC<sup>10-12</sup>. However, the relation of NFC and SVs might be confounded, since other studies utilizing the COG-ED paradigm found the association of NFC and SVs to 79 disappear after correcting for performance<sup>13</sup> or found no association of NFC and SVs at all<sup>14</sup>. On the other hand, task load has been shown to be a better predictor of SVs than task performance<sup>7,15,16</sup>, so more research is needed to shed light on this issue.

With the present study, we alter one fundamental assumption of the original

83

COG-ED paradigm: That the easiest n-back level has the highest SV. We therefore adapted the COG-ED paradigm in a way that allows the computation of SVs for different 85 n-back levels without presuming that all individuals inherently prefer the easiest level. Since we also aim to establish this paradigm for the assessment of tasks with no objective 87 task load, e.g., emotion regulation tasks<sup>17</sup>, we call it the Cognitive and Affective Discounting Paradigm (CAD). In the present study, we validated the CAD paradigm by conceptually replicating the findings of Westbrook et al. 7. Additionally, we compared the effort discounting behavior of participants regarding the n-back task and an emotion regulation task. The full results of the latter are published in a second Registered Report<sup>17</sup>. The COG-ED paradigm has been applied to tasks in different domains before, showing 93 that SVs across task domains correlate<sup>14</sup>, but these tasks had an objective order of task load, which is not the case for the choice of emotion regulation strategies or other paradigms where there is no objective order of task load.

Our hypotheses were derived from the results of Westbrook et al.<sup>7</sup>. As a manipulation 97 check, we hypothesized that with increasing n-back level the (1a) the signal detection 98 parameter d' declines, while (1b) reaction time and (1c) perceived task load increase. 99 Regarding the associations of task load and effort discounting we hypothesized that (2a) 100 SVs decline with increasing n-back level, and (2b) they do so even after controlling for 101 declining task performance. And finally, we hypothesized that the CAD paradigm can 102 show inter-individual differences in effort discounting, such that participants with higher 103 NFC have (3a) lower SVs for 1-back but higher SVs for 2- and 3-back, (3b) lower perceived 104 task load across all levels, and (3c) higher aversion against 1-back but lower aversion 105 against 2- and 3-back. Each hypothesis is detailed in the Design Table in the Appendix. 106

107 Methods

We report how we determined our sample size, all data exclusions (if any), all manipulations, and all measures in the study<sup>cf. 18</sup>. The paradigm was written and

presented using  $Psychopy^{19}$ . We used R with R  $Studio^{20,21}$  with the main packages  $afex^{22}$  and  $BayesFactor^{23}$  for all our analyses.

### 112 Ethics information

The study protocol complies with all relevant ethical regulations and was approved by the ethics committee of the Technische Universität Dresden (reference number SR-EK-50012022). Prior to testing, written informed consent was obtained. Participants received 24€ in total or course credit for participation.

#### 117 Design

**CAD Paradigm.** Figure 1 illustrates how different modifications of the COG-ED 118 paradigm<sup>7</sup> return SVs that do or do not reflect the true preference of a hypothetical 119 participant, who likes 2-back most, 3-back less, and 1-back least (for reasons of clarity there 120 are only three levels in the example). The COG-ED paradigm, which compares every more 121 difficult level with 1-back sets the SV of 1-back to 1, regardless of the response pattern. 122 Adding a comparison of the more difficult levels with each other allows the SVs of those two 123 levels to be more differentiated, but leaves the SV of 1-back unchanged. Adding those same 124 pairs again, but with the opposite assignment of fixed and flexible level, does approach the 125 true preference, but has two disadvantages. First, the SVs are still quite alike across levels 126 due to the fact that every more difficult level has only been compared with the easiest 127 level, and second, having more task levels than just three would lead to an exponential 128 increase in comparisons. Therefore, the solution lies in reducing the number of necessary comparisons by presenting only one effort discounting round for each possible pair of levels after determining for each pair which level should be fixed and which should be flexible. 131 This is determined by presenting each possible pair of levels on screen with the question 132 "Would you prefer 1 € for level A or 1 € for level B?". Participants respond by clicking the 133 respective on-screen button. Each pair is presented three times, resulting in 18 presented 134

pairs, which are fully randomized in order and in the assignment of which level is on the left 135 or right of the screen. For each pair, the level that was chosen by the participant at least 136 two out of three times will be used as the level with a flexible value, which starts at  $1 \in$ 137 and is changed in every iteration. The other level in the pair will be set to a fixed value of 138 2 €. Then, the effort discounting sensu Westbrook et al. begins, but with all possible pairs 139 and with the individually determined assignment of fixed and flexible level. The order in 140 which the pairs are presented is fully randomized, and each pair goes through all iteration 141 steps of adding/subtracting  $0.50 \in 0.25 \in 0.13 \in 0.06 \in 0.03 \in 0.02 \in to/from the$ flexible level's reward (each adjustment half of the previous one, rounded to two decimals) 143 before moving on to the next one. This procedure allows to compute SVs based on actual 144 individual preference instead of objective task load. For each pair, the SV of the flexible 145 level is 1, as it was preferred when faced with equal rewards, and the SV of the fixed level is the final reward of the flexible level divided by 2 €. Each level's "global" SV is calculated as the mean of this level's SVs from all pairs in which it appeared. If the participant has a clear preference for one level, this level's SV will be 1. If not, then no level's SV will be 1, but each level's SV can still be interpreted as an absolute and relative value, so each 150 participant's effort discounting behaviour can still be quantified. The interpretation of SVs 151 in Westbrook et al. was "The minimum relative reward required for me to choose 1-back 152 over this level". So if the SV of 3-back was 0.6, the participant would need to be rewarded 153 with at least 60 % of what they are being offered for doing 3-back to do 1-back instead, 154 forgoing the higher reward for 3-back. In this study, the SV can be interpreted as "The 155 minimum relative reward required for me to choose any other level over this level". 156 Therefore, an SV of 1 indicates that this level is preferred over all others, while SVs lower 157 than 1 indicate that in at least one pair, a different level was preferred over this one.

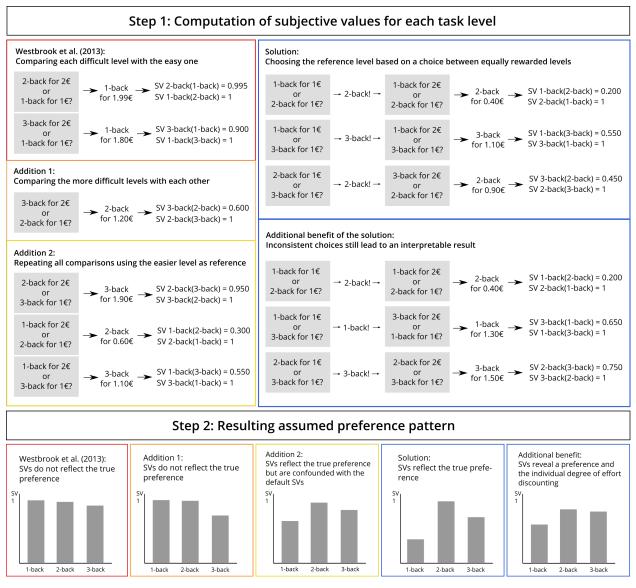


Figure 1. An example for subjective values for an n-back task with three levels, returned by different modifications of the COG-ED paradigm for a hypothetical participant with the true preference 2-back > 3-back > 1-back. The grey boxes are the choice options shown to the participant. The participant's final reward value of the flexible level is displayed after the first arrow. The resulting subjective value of each level is displayed after the second arrow, in the notation "SV 3-back(1-back)" for the subjective value of 3-back when 1-back is the other choice. The Solution and Additional Benefit panel follow the same logic, but are preceded by a choice between equal rewards, and the participant's first choice indicated by an exclamation mark.

Study procedure. Healthy participants aged 18 to 30 years were recruited using the software  $ORSEE^{24}$ . Participants completed the personality questionnaires online and

then visited the lab for two sessions one week apart. NFC was assessed using the 16-item 162 short form of the Need for Cognition Scale<sup>25,26</sup>. Responses to each item (e.g., "Thinking is 163 not my idea of fun", recoded) were recorded on a 7-point Likert scale. The NFC scale 164 shows comparably high internal consistency (Cronbach's  $\alpha > .80$ )<sup>26,27</sup>. Several other 165 personality questionnaires were used in this study but are the topic of the Registered 166 Report for the second lab session<sup>17</sup>. A full list of measures can be found in our Github 167 repository. In the first session, participants provided informed consent and demographic 168 data before completing the computer-based paradigm. The paradigm started with the 169 n-back levels one to four, presented sequentially with two runs per level, consisting of 64 170 consonants (16 targets, 48 non-targets) per run. The levels were referred to by color 171 (1-back black, 2-back red, 3-back blue, 4-back green) to avoid anchor effects in the effort 172 discounting procedure. To assess perceived task load, we used the 6-item NASA Task Load Index (NASA-TLX)<sup>28</sup>, where participants evaluate their subjective perception of mental 174 load, physical load, effort, frustration, performance, and time pressure during the task on a 20-point scale. At the end of each level, participants filled out the NASA-TLX on a tablet, 176 plus an item with the same response scale, asking them how aversive they found this 177 n-back level. After the n-back task, participants completed the CAD paradigm on screen 178 and were instructed to do so as realistically as possible, even though the displayed rewards 179 were not paid out on top of their compensation. They were told that one of their choices 180 would be randomly picked for the final run of n-back, the data of which was not analyzed 181 as it only served to incentivise truthful behavior and stay close to the design of Westbrook 182 et al.<sup>7</sup>. After the CAD paradigm, participants filled out a short questionnaire on the tablet, 183 indicating whether they adhered to the instructions (yes/no) and what the primary 184 motivation for their decisions during the effort discounting procedure was (avoid 185 boredom/relax/avoid effort/seek challenge/other). 186

The second session consisted of an emotion regulation task with negative pictures and the instruction to suppress facial reactions, detach cognitively from the picture content,

and distract oneself, respectively. The paradigm followed the same structure of task and
effort discounting procedure, but participants could decide which strategy they wanted to
reapply in the last block. Study data was collected and managed using REDCap electronic
data capture tools hosted at Technische Universität Dresden<sup>29,30</sup>.

# 193 Sampling plan

Sample size determination was mainly based on the results of the analyses of 194 Westbrook et al.<sup>7</sup> (see Design Table). The hypothesis that yielded the largest necessary 195 sample size was a repeated measures, within-factors ANOVA of reaction time and n-back 196 level. Sample size analysis with  $G^*Power^{31,32}$  indicated that we should collect data from at 197 least 53 participants, assuming  $\alpha = .05$  and  $\beta = .95$ . However, the sample size analysis for 198 the hypotheses of the second lab session revealed a larger necessary sample size of 85 190 participants to find an effect of d = -0.32 of emotion regulation on facial muscle activity 200 with  $\alpha = .05$  and  $\beta = .95$ . To account for technical errors, noisy physiological data, or 201 participants who indicate that they did not follow the instructions, we aimed to collect 202 about 50% more data sets than necessary, N = 120 in total. 203

### $_{04}$ Analysis plan

Data collection and analysis were not performed blind to the conditions of the 205 experiments. We excluded the data of a participant from all analyses, if the participant 206 stated that they did not follow the instructions, if the investigator noted that the 207 participant misunderstood the instructions, or if the participant withdrew their consent. No data was replaced. The performance measure d' was computed as the difference of the z-transformed hit rate and the z-transformed false alarm rate<sup>33</sup>. Reaction time (RT) data 210 was trimmed by excluding all trials with responses faster than 100 ms, as the relevant 211 cognitive processes cannot have been completed before<sup>34,35</sup>. Aggregated RT values were 212 described using the median and the median of absolute deviation (MAD) as robust 213

estimates of center and variability, respectively<sup>36</sup>. Error- and post-error trials were
excluded, because RT in the latter is longer due to more cautious behavior<sup>37,38</sup>. To test our
hypotheses, we performed a series of rmANOVAs and an MLM with orthogonal
sum-to-zero contrasts in order to meaningfully interpret results<sup>39</sup>.

218 Manipulation check. Declining performance was investigated by calculating an
219 rmANOVA with six paired contrasts comparing d' between two levels of 1- to 4-back at a
220 time. Another rmANOVA with six paired contrasts was computed to compare the median
221 RT between two levels of 1- to 4-back at a time. To investigate changes in NASA-TLX
222 ratings, six rmANOVAs were computed, one for each NASA-TLX subscale, and each with
223 six paired contrasts comparing the ratings between two levels of 1- to 4-back at a time.

Subjective values. For each effort discounting round, the SV of the fixed level was 224 calculated by adding or subtracting the last adjustment of 0.02 € from the last monetary 225 value of the flexible level, depending on the participant's last choice, and dividing this value 226 by  $2 \in$ . This yielded an SV between 0 and 1 for the fixed compared with the flexible level, 227 while the SV of the flexible level was 1. The closer the SV of the fixed level is to 0, the 228 stronger the preference for the flexible level. All SVs of each level were averaged to compute 229 one "global" SV for each level. An rmANOVA with four different contrasts were computed 230 to investigate the association of SVs and the *n*-back levels: Declining linear (3,1,-1,-3), 231 ascending quadratic (-1,1,1,-1), declining logistic (3,2,-2,-3), and positively skewed normal 232 (1,2,-1,-2). Depending on whether the linear or one of the other three contrasts fit the curve best, we applied a linear or nonlinear multi-level model in the next step, respectively.

To determine the influence of task performance on the association of SVs and n-back level, we performed MLM. We applied restricted maximum likelihood (REML) to fit the model. As an effect size measure for random effects we first calculated the intraclass correlation (ICC), which displays the proportion of variance that is explained by differences between persons. Second, we estimated a random slopes model of n-back level (level 1,

fixed, and random factor: 0-back, 1-back, 2-back, 3-back) predicting SV nested within
subjects. As Mussel et al. 40 could show, participants with high versus low NFC not only
have a more shallow decline in performance with higher n-back levels, but show a
demand-specific increase in EEG theta oscillations, which has been associated with mental
effort. We controlled for performance, i.e., d' (level 1, fixed factor, continuous), median RT
(level 1, fixed factor, continuous) in order to eliminate a possible influence of declining
performance on SV ratings.

$$SV \sim level + d' + medianRT + (level|subject)$$

Level-1-predictors were centered within cluster as recommended by Enders & Tofighi<sup>41</sup>. By this, the model yields interpretable parameter estimates. If necessary, we will adjusted the 248 optimization algorithm to improve model fit. We visually inspected the residuals of the 249 model for evidence to perform model criticism. This was done by excluding all data points 250 with absolute standardized residuals above 3 SD. As effect size measures, we calculated pseudo  $R^2$  for our model and  $f^2$  to estimate the effect of n-back level according to Lorah<sup>42</sup>. To ensure the validity of the association of SVs and n-back level, we conducted a 253 specification curve analysis<sup>43</sup>, which includes 63 possible preprocessing pipelines of the RT data. These pipelines specify which transformation was applied (none, log, inverse, or 255 square-root), which outliers were excluded (none, 2, 2.5, or 3 MAD from the median, RTs 256 below 100 or 200 ms), and across which dimensions the transformations and exclusions 257 were applied (across/within subjects and across/within n-back levels). The rmANOVA was 258 run with each of the 63 pipelines, which will also include our main pipeline (untransformed 259 data, exclusion of RTs below 100 ms). The ratio of pipelines that lead to significant versus 260 non-significant effects provides an indication of how robust the effect actually is. The 261 specification curve analysis was linked to the MLM in the initial submission, and was 262 assigned to hypothesis 3a during the review process and in the Stage 1 report. Afterwards, 263

we noticed that 3a does not contain any RT data, so the specification curve analysis was reassigned to the MLM with the agreement of the editor.

The association of SVs and NFC was examined with an rmANOVA. We subtracted the SV of 1- from 2-back and 2- from 3-back, yielding two SV difference scores per participant. The sample was divided into participants with low and high NFC using a median split. We then computed an rmANOVA with the within-factor n-back level and the between-factor NFC group to determine whether there is a main effect of level and/or group, and/or an interaction between level and group on the SV difference scores. Post-hoc tests were computed depending on which effect reached significance at p < .01.

The association of subjective task load with NFC was examined similarly. We 273 calculated NASA-TLX sum scores per participant per level, computed an rmANOVA with 274 the within-factor n-back level and the between-factor NFC group, and applied post-hoc 275 tests based on which effect reached significance at p < .01. And the association of 276 subjective aversiveness of the task with NFC was examined with difference scores as well, 277 since we expected this curve to mirror the SV curve, i.e. as the SV rises, the aversiveness 278 declines, and vice versa. We subtracted the aversiveness ratings of 1- from 2-back and 2-279 from 3-back, yielding two aversiveness difference scores per participant. Then, we computed an rmANOVA with the within-factor n-back level and the between-factor NFC group, and applied post-hoc tests based on which effect reached significance at p < .01.

The results of each analysis was assessed on the basis of both p-value and the Bayes factor  $BF_{10}$ , calculated with the BayesFactor package<sup>23</sup> using the default prior widths of the functions anovaBF, lmBF and ttestBF. We considered a  $BF_{10}$  close to or above 3/10 as moderate/strong evidence for the alternative hypothesis, and a  $BF_{10}$  close to or below .33/.10 as moderate/strong evidence for the null hypothesis<sup>44</sup>.

### Pilot data

The sample of the pilot study consisted of N=15 participants (53.3% female, 289  $M = 24.43 \ (SD = 3.59)$  years old). One participant's data was removed because they 290 misunderstood the instruction. Due to a technical error the subjective task load data of one participant was incomplete, so the hypotheses involving the NASA Task Load Index were analyzed with n = 14 data sets. The results showed increases in subjective and 293 objective task load measures with higher n-back level. Importantly, SVs were lower for higher n-back levels, but not different between 1- and 2-back, which shows that the easiest level is not universally preferred. The LMM revealed n-back level as a reliable predictor of SV, even after controlling for declining task performance (d' and median RT). NASA-TLX 297 scores were higher with higher n, and lower for the group with lower NFC scores, but NFC 298 and n-back level did not interact. All results are detailed in the Supplementary Material. 299

# 300 Data availability

The data of this study can be downloaded from osf.io/vnj8x/.

# OZ Code availability

The paradigm code as well as the R Markdown file used to analyze the data and write this document are available at github.com/ChScheffel/CAD.

#### 305 Protocol registration

The Stage 1 Registered Report protocol has been approved and is available at osf.io/qa2bg/.

Results

Data was collected between the 16th of August 2022 and the 3rd of February 2023.

All of the N = 124 participants who filled out the online questionnaires came to the first

lab session. Based on the experimenters' notes, we excluded the data of seven participants

from analysis for misunderstanding the instruction of the n-back task, and the data of one

participant who reported that they confused the colours of the levels during effort

discounting. Our final data set therefore included N = 116 participants, which is 2.20

times more than what the highest sample size calculation required.

# 316 Manipulation checks

The performance measure d' did not change across n-back levels  $(F(2.85, 327.28) = 0.01, p = .999, \hat{\eta}_G^2 = .000, 90\%$  CI  $[.000, .000], BF_{10} = 3.31 \times 10^{-3}), but$  the median RT did  $(F(2.46, 283.05) = 98.67, p < .001, \hat{\eta}_G^2 = .192, 90\%$  CI [.130, .248], BF<sub>10</sub> =  $2.28 \times 10^{34}$ ). Specifically, the median RT was higher for the more difficult level in every contrast, with two exceptions: It did not differ between 2- and 4-back, and it was higher for 3- than for 4-back (Table 1).

Table 1
Paired contrasts for the rmANOVA comparing the median reaction time between n-back levels

Contrast	Estimate	SE	df	t	p	BF10	$\eta_p^2$	95%CI
1 - 2	-0.12	0.01	345.00	-11.80	0.000	$1.75 \times 10^{30}$	0.29	[0.22, 1.00]
1 - 3	-0.16	0.01	345.00	-16.20	0.000	$8.80 \times 10^{45}$	0.43	[0.37, 1.00]
1 - 4	-0.12	0.01	345.00	-12.50	0.000	$4.79 \times 10^{34}$	0.31	[0.25, 1.00]
2 - 3	-0.04	0.01	345.00	-4.47	0.000	$5,\!538.45$	0.05	[0.02, 1.00]
2 - 4	-0.01	0.01	345.00	-0.71	0.894	0.10	1.45e-03	[0.00, 1.00]
3 - 4	0.04	0.01	345.00	3.76	0.001	$6.35 \times 10^{6}$	0.04	[0.01, 1.00]

Note. The column Contrast contains the n of the n-back levels. SE = standard error, df = degrees of freedom, t = t-statistic, p = p-value, CI = confidence interval.

All NASA-TLX subscale scores increased across n-back levels. The effort subscale 323  $(F(2.20, 253.06) = 203.82, p < .001, \hat{\eta}_G^2 = .316, 90\% \text{ CI } [.250, .375], \text{BF}_{10} = 2.47 \times 10^{34})$ 324 increased across all levels, but the magnitude of change decreased from 1- to 2-back 325  $(t(345) = -12.35, p_{\text{Tukey}(4)} < .001, BF_{10} = 4.24 \times 10^{19})$  to 3- to 4-back (t(345) = -2.72, 0.001)326  $p_{\text{Tukey}(4)} = .035$ , BF<sub>10</sub> = 174.38). Three subscales had significant differences between all 327 contrasts except for 3- versus 4-back: While ratings on the frustration and time subscales 328 were higher for more difficult levels  $(F(2.50, 287.66) = 68.06, p < .001, \hat{\eta}_G^2 = .172, 90\%$  CI 329 [.112, .227], BF<sub>10</sub> =  $5.26 \times 10^{15}$ , and F(2.21, 254.65) = 51.08, p < .001,  $\hat{\eta}_G^2 = .117$ , 90% CI 330 [.065, .168], BF<sub>10</sub> =  $3.94 \times 10^9$ , respectively), ratings on the performance subscale decreased 331 with higher n  $(F(2.49, 285.97) = 95.33, p < .001, \hat{\eta}_G^2 = .241, 90\%$  CI [.176, .299], 332  $BF_{10} = 1.55 \times 10^{24}$ ). Ratings on the mental subscale consistently increased across all levels 333  $(F(1.99, 228.35) = 274.47, p < .001, \hat{\eta}_G^2 = .375, 90\% \text{ CI } [.309, .432], \text{ BF}_{10} = 1.64 \times 10^{43}).$ 334 Ratings on the physical subscale were higher for more difficult levels 335  $(F(1.68, 192.93) = 15.91, p < .001, \hat{\eta}_G^2 = .041, 90\% \text{ CI } [.009, .075], \text{ BF}_{10} = 60.54), \text{ apart}$ 336 from the contrasts 2- versus 3-back  $(t(345) = -2.34, p_{\text{Tukey}(4)} = .092, BF_{10} = 10.45)$  and 3-337 versus 4-back  $(t(345) = -1.07, p_{\text{Tukey}(4)} = .705, BF_{10} = 0.47).$ 338

# 339 Decline of subjective values

348

When asking participants what motivated their decisions in the effort discounting,
11.2% stated that they wanted to avoid boredom, 22.4% stated that they wanted a
challenge, 34.5% stated that they wanted to avoid effort, and 4.3% stated that they wanted
to relax. The remaining 27.6% of participants used the free text field and provided reasons
such as "I wanted a fair relation of effort and reward.", "I wanted the fun that I had in the
more challenging levels.", "I wanted to maximize reward first and minimize effort second.",
or "I did not want to perform poorly when I was being paid for it." Figure XX in the
Supplement shows the different motivations in the context of the SVs per n-back level.

The rmANOVA showed a significant difference between the SVs across n-back levels

 $(F(1.98, 227.98) = 65.65, p < .001, \hat{\eta}_G^2 = .288, 90\%$  CI [.222, .347], BF<sub>10</sub> = 1.58 × 10<sup>64</sup>). All four pre-defined contrasts reached significance (Table 2), so a purely linear contrast can be rejected.

Table 2
Contrasts for the rmANOVA comparing the subjective values between n-back levels

Contrast	Estimate	SE	df	t	p	$\eta_p^2$	95%CI
Declining Linear	1.11	0.08	345.00	13.40	<.001	0.34	[0.28, 1.00]
Ascending Quadratic	0.15	0.04	345.00	4.14	<.001	0.05	[0.02, 1.00]
Declining Logistic	1.22	0.09	345.00	13.00	<.001	0.33	[0.26, 1.00]
Positively Skewed Normal	0.75	0.06	345.00	12.70	<.001	0.32	[0.26, 1.00]

Note. SE = standard error, df = degrees of freedom, t = t-statistic, p = p-value, CI = confidence interval.

The declining logistic contrast had the highest effect estimate (t(345) = 12.97, p < .001), suggesting a shallow decline of SVs between 1- and 2-back, and 3- and 4-back, respectively, and a steeper decline of SVs between 2- and 3-back.

Consequently, we had to adapt the multi level model to incorporate this non-linear trend. To apply the contrast to the *n*-back levels, we had to turn the variables into a factor, with two consequences: Centered variables cannot be turned into factors, so we entered the variable level in its raw form, and factors cannot be used as random slopes, so the model is now defined as:

$$SV \sim level + d' + medianRT + (1|subject)$$

This means that the intercept still varies between subjects, but there are no random slopes anymore. To provide more than one observation per factor level, we used the two rounds per n-back level per subject, rather than n-back levels per subject. The ICC of the null model indicated that there is a correlation of r = .10 between the SVs of a subject, i.e. that 9.59% of variance in SVs can be explained by differences between participants. We did not

use an optimization algorithm to improve the fit of the random intercept model. A total of
9 data points from 6 participants were excluded, because the residuals exceeded 3 SD
above the mean. The results of the final model are displayed in Table 3.

Table 3
Results of the multi level model on the influence of n-back level (as a declining logistic contrast) and task performance on subjective values.

Parameter	Beta	SE	df	t-value	<i>p</i> -value	$f^2$	Random Effects (SD)
Intercept	0.95	0.02	507.45	59.45	<.001		0.09
n-back level	-0.04	0.02	800.15	-2.36	<.001	0.64	
ď,	0.02	0.00	798.75	5.60	<.001	0.04	
median RT	0.02	0.07	798.58	0.30	0.768	0.00	

Note: SE = standard error, df = degrees of freedom, SD = standard deviation.

An exploratory was used ANOVA to compare the fit of the final model with a linear 368 random intercept model, confirming that the two models were different from each other 369  $(\chi^2(2) = 34.48, p < .001)$ , and with an Akaike Information Criterion of AIC = -492.61370 and a Bayesian Information Criterion of BIC = -454.02 the declining logistic model was superior to the linear model (AIC = -462.12, BIC = -433.18). The final model had an effect size of  $f^2 = 0.64$  for the *n*-back levels and  $f^2 = 0.04$  for d', which are considered large 373 and small, respectively<sup>45</sup>. This means that the n-back level explained 64.20% and d'explained 3.95% of variance in SVs relative to the unexplained variance, respectively. The 375 beta coefficient indicated that with every 1-unit increase in d', the SV increased by 0.02. 376 The effect size of the median RT was  $f^2 = 0.00$ . 377

To investigate the dependency of the model results on the RT preprocessing, we conducted a specification curve analysis (Figure 2).

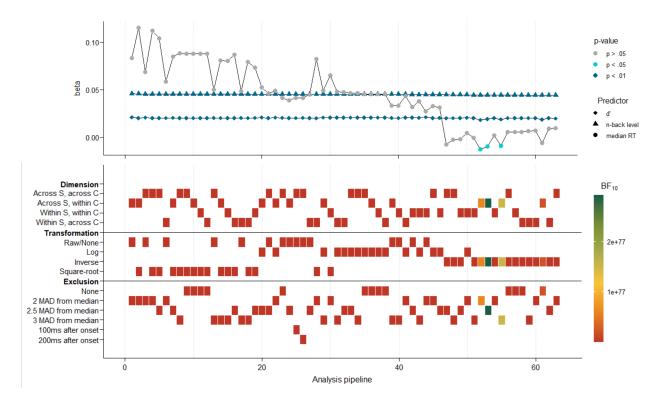


Figure 2. Results of the multi level model for each of the 63 preprocessing pipelines. The lower panel indicates the type of preprocessing, the upper panel shows the beta coefficient of each predictor and its p-value. The colourbar indicates the BF10. The pipelines are sorted in descending order of the magnitude of the n-back level beta.

Regardless of the preprocessing pipeline, *n*-back level and *d'* were significant predictors of SVs, and had stable effect estimates across all pipelines. The only pipelines in which the median RT was a significant predictor of SVs, were the three pipelines with the highest Bayes Factors. These three pipelines contain data that has been inverse transformed across subjects but within conditions, i.e. within the round of an *n*-back level.

#### 6 Differences between NFC groups

Figure 3 shows a scatterplot of SVs per n-back level, colored depending on the participant's NFC score. There is a concentration of participants who have assigned their highest SV to 1-back, and this concentration fades across n-back levels. At the same time, there is a subtle separation of SVs across n-back levels, depending on the participant's

NFC score: While the SVs of those with higher NFC scores remain elevated, the SVs of those with lower NFC scores decline more strongly. Specifically, n = 71 participants had an absolute preference for 1-back, n = 18 for 2-back, n = 9 for 3-back, and n = 13 for 4-back. There were n = 5 participants who did not have an absolute preference for any n-back level, i.e. none of their SVs was 1.

### INSERT FIGURE 3 HERE

396

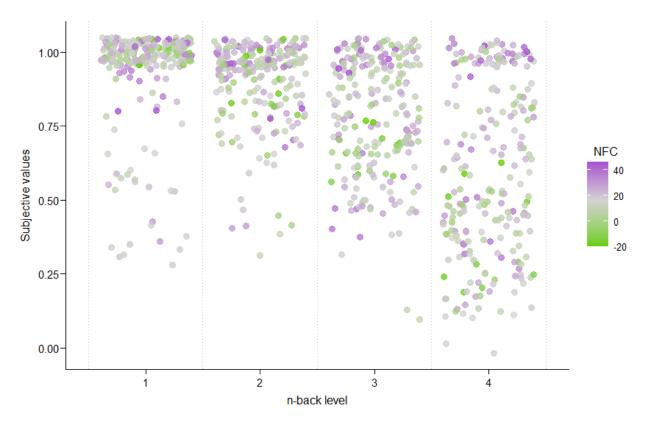


Figure 3. Subjective values per n-back level. Each dot indicates a participant, the colours indicate their Need for Cognition score. N=116. There is a horizontal jitter of 0.4 and a vertical jitter of 0.05 for visual clarity.

The median NFC was 16, with n=57 subjects below and n=59 above the median.

We used an rmANOVA to investigate whether the difference between the SVs of 1- and

2-back, and 2- and 3-back, respectively, depended on whether a participant's NFC score

was above or below the median. There was a main effect of the n-back level

( $F(1,114) = 9.13, p = .003, \hat{\eta}_G^2 = .040, 90\%$  CI [.002, .115]), but neither a main effect of the

NFC group  $(F(1, 114) = 3.18, p = .077, \hat{\eta}_G^2 = .013, 90\%$  CI [.000, .068]) nor an interaction of NFC group and n-back level  $(F(1, 114) = 0.46, p = .499, \hat{\eta}_G^2 = .002, 90\%$  CI [.000, .037]).

Post-hoc tests showed that the difference between the SVs of 2- and 3-back is slightly more negative than the difference between 1- and 2-back (t(114) = 3.02, p = .003), but there were large inter-individual differences, especially for 2- and 3-back (Figure 4). This means that across the whole sample, there was a steeper decline in SVs from 2- to 3-back than from 1- to 2-back, but some participants showed a completely opposite pattern.

### INSERT FIGURE 4 HERE

409

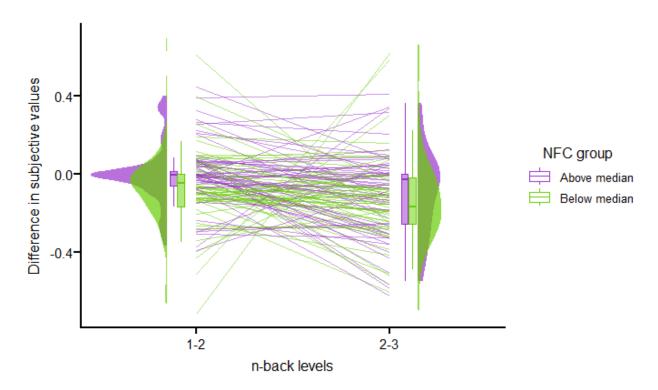


Figure 4. Difference scores for subjective values when subtracting 2- from 1-back and 3-from 2-back. Each line indicates a participant, the colours indicate their Need for Cognition score. N=116.

The rmANOVA on the association between NFC scores and NASA-TLX scores revealed a main effect of n-back level  $(F(2.10, 239.56) = 154.50, p < .001, \hat{\eta}_G^2 = .223, 90\%$  CI [.159, .282], BF<sub>10</sub> =  $2.22 \times 10^{45}$ ) and an interaction between n-back level and NFC scores  $(F(2.10, 239.56) = 4.93, p = .007, \hat{\eta}_G^2 = .009, 90\%$  CI [.000, .025]), but no main effect

of NFC scores  $(F(1, 114) = 3.22, p = .075, \hat{\eta}_G^2 = .022, 90\%$  CI [.000, .084], 414  $\mathrm{BF}_{10}=1.75\times10^2$ ). Post-hoc tests showed that the participants with NFC scores below the 415 median had higher NASA-TLX scores for 3-back  $(t(114) = -2.15, p = .033, BF_{10} = 11.15)$ 416 and for 4-back  $(t(114) = -2.89, p = .005, BF_{10} = 336.88)$  than those with NFC scores 417 above the median. Regardless of NFC scores, NASA-TLX scores were higher for the more 418 difficult level in each pair of n-back levels (Figure 5). 419

### INSERT FIGURE 5 HERE

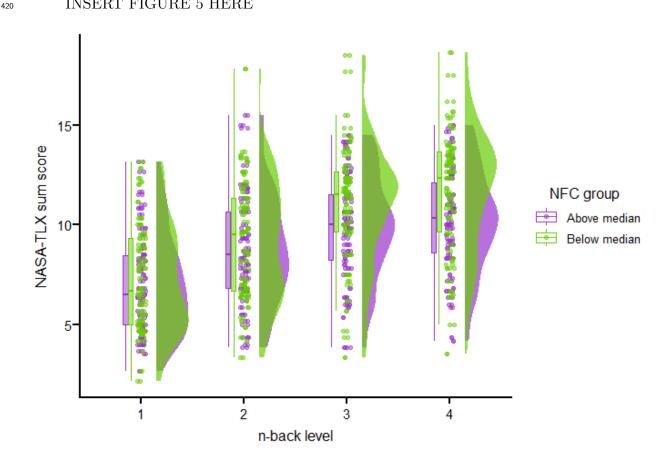


Figure 5. NASA-TLX sum scores for each n-back level. Colours indicate Need for Cognition (NFC) score above or below the median. N = 116.

With another rmANOVA we investigated whether the difference between the 421 aversiveness scores of 1- and 2-back, and 2- and 3-back, respectively, depended on whether 422 a participant's NFC score was above or below the median. There was a main effect of NFC 423 group  $(F(1,114) = 8.43, p = .004, \hat{\eta}_G^2 = .043, 90\%$  CI [.003, .119], BF<sub>10</sub> = 14.26) and a 424

main effect of the *n*-back level  $(F(1,114) = 10.21, p = .002, \hat{\eta}_G^2 = .034, 90\%$  CI [.000, .105], but no interaction. Post-hoc tests revealed that participants with NFC scores below the median reported higher aversiveness than participants with NFC scores above the median (t(114) = -2.90, p = .004). Regardless of NFC, the difference of the aversiveness scores of 2- and 3-back was smaller and more negative than that of 1- and 2-back (t(114) = -3.20, p = .002), but again, there were large inter-individual differences.

431 Discussion

This Registered Report aimed to adapt the Cognitive Effort Discounting Paradigm 432 (COG-ED) paradigm by Westbrook et al.<sup>7</sup>, which can estimate subjective values of 433 different n-back levels, into the Cognitive and Affective Discounting (CAD) paradigm, 434 which can estimate subjective values of tasks without assuming that the easiest level is 435 inherently preferred. For this purpose, we changed the way in which the discounting 436 options are presented to the participants, basing the anchor on their own choices rather 437 than on the objective task load. The analyses were closely aligned with those in Westbrook 438 et al.  $^7$  to demonstrate the changes in SVs brought about by the new paradigm. This study 439 also applied the CAD paradigm to an emotion regulation task, the results of which are detailed in Scheffel et al..

# 442 Manipulation checks

The performance measure d' did not differ across n-back levels, but the RT increased from 1- to 2- to 3-back and then remained on a high level for 4-back. This points to three important characteristics of the n-back task in this context. Firstly, RT as a valid group-level indicator of performance might only be useful for levels up to n = 3, and could be used to investigate inter-individual differences for n > 3. Secondly, there is a speed-accuracy tradeoff in the first three levels, that might even re-emerge in higher levels, where d' would decline and RT would remain stable. And lastly, the fact that neither accuracy and nor speed is an informative performance measure by itself has been observed before  $^{46}$  and both show different associations with various measures of intelligence  $^4$ , suggesting that they should always be reported as separate indices. Additionally, d' might not have differed across n-back levels because the manipulation of task load is not strictly continuous. Several participants said that they perceived 3-back as more difficult than d-back because they found it is easier to remember chunks of stimuli when d-back as no even number than when d-back as no odd number.

All NASA-TLX subscales differed across n-back levels, but the effort and mental load subscales were the only ones to consistently increase across all levels. This would support the notion of the n-back task offering a continuous manipulation of task load, as least subjectively. Ratings on the frustration and time subscales increased and ratings on the performance subscale decreased until 3-back and then remained stable. This pattern is akin to the RT, which also increased and then remained stable. Ratings on the physical load subscale increased with n-back levels, but not between 2- and 3-back and 3- and 4-back, respectively.

### 65 Decline of subjective values

The rmANOVA with different pre-defined contrasts showed that all fit the SVs to a 466 different degree, and that the SVs do not simply decline linearly across n-back levels. The 467 best fit was a declining logistic curve, reflecting that the majority of participants preferred 468 1-back and that SVs for 2-back were also high, before having more inter-individual variance 469 for 3- and 4-back. The MLM with the logistic contrast showed that the n-back level explained the majority of variance in SVs, while the performance measure d' also explained 471 some variance in SVs, albeit less. With increasing n-back level and decreasing d', the SV decreased. The median RT was not a significant predictor in this model, which was somewhat surprising because RT but not d' yielded significant differences across levels in 474 the manipulation checks. However, participants might have deliberately or subconsciously

used the feedback they received at the end of each round, i.e. twice per n-back level, as an 476 anchor during the effort discounting. This feedback was based on correct responses and not 477 on RT, so if participants based their effort discounting choices at least partly on this 478 feedback, they were either motivated to repeat a task in which they performed well and/or 479 they were reluctant to accept a larger reward for a task in which they did not perform well. 480 Since more participants reported effort avoidance as their motivation in the effort 481 discounting than those who reported seeking a challenge, we can assume that they were 482 more motivated to repeat a task in which they performed well because their good 483 performance coincided with low effort. 484

The declining logistic n-back levels and d' remained significant predictors of SVs 485 throughout all 63 preprocessing pipelines in the specification curve analysis, with betas 486 that varied by less than 0.01. In contrast to this stood the variability of the median RT 487 betas, which ranged from about 0.11 to -0.01, and reached significance in only three 488 pipelines. These three pipelines had the highest  $BF_{10}$  and applied inverse transformation 480 to the RT data, across subjects but within conditions, and excluded data based on the 490 MAD. Interestingly, the curve of median RT betas in the upper panel of Figure 2 mirrored 491 the rectangular pipeline indicators in the transformation rows of the lower panel, so the 492 transformation choice influenced the median RT much more than the dimension or the 493 exclusion choice did. As Fernandez et al. 47 found, applying more than one preprocessing 494 step to the reaction time data of a Stroop task increased the risk of false positives beyond 495  $\alpha = .05$ , and transformation choices inflated this risk more than outlier exclusion or aggregation choices did. Our data seems to corroborate this finding for n-back tasks as well. Surprisingly, the d' betas appear almost unaffected by the preprocessing pipeline, even though d' was computed after the outlier exclusion. This indicates that researchers who are interested in the correctness rather than the speed of responses can choose a simple 500 preprocessing pipeline without risking false positives through elaborate transformations. 501

# Differences between NFC groups

The majority of participants (61.20 %) had an absolute preference for 1-back over the other levels, but that also means that there were 34.50 % who had an absolute preference for 2-, 3-, or 4-back, and 4.30 % who preferred no specific level over all others. It shows that when given the choice, there is a large number of participants who do not prefer the easiest level, confirming the necessity of an effort discounting paradigm that works independent of the objective task load. The CAD paradigm provides the means to depict these preferences.

Despite the visual separation of the SVs of participants with very low and high NFC 500 scores in higher n-back levels, the NFC group did not reach significance in predicting SVs. 510 This was likely due to the bandwidth of SVs of participants with NFC scores around the 511 median, and due to the fact that the difference appeared most pronounced for 4-back, and 512 we only analyzed the difference scores between 1- and 2-back and 2- and 3-back. The 513 analysis of NASA-TLX scores showed that the sum score increased with every n-back level, 514 and that participants with NFC scores below the median had higher NASA-TLX scores for 515 3- and 4-back than those below the median. This demonstrates that higher n-back levels have a higher discriminatory power regarding inter-individual differences in subjective effort perception. This was also supported by the fact that higher n-back levels were 518 perceived as more aversive, and participants with NFC scores below the median reported higher aversion than those with NFC scores above the median.

### 521 Limitations

When developing a new paradigm, it is challenging to decide on the optimal analysis strategy, as every hypothesis is based on expected data patterns rather than previous findings. While the Stage 1 review process made the analyses as robust as possible, there were still unknown factors that should be addressed by future studies. For instance, the differences between participants with higher and lower NFC should be investigated with

extreme groups rather than a median split, especially in academic samples where NFC can 527 be expected to be higher on average and more narrow in range. Additionally, we expected 528 the SVs of participants with lower NFC scores to peak at 1-back and the SVs of those with 529 higher scores to peak at 2-back, but the way the SVs of both groups appeared to drift 530 apart in the higher n-back levels suggests that an analysis of those levels would be more 531 fruitful in determining group differences. Another important point is the instruction, not 532 just for the n-back task, but for the effort discounting as well. We had to exclude several 533 participants for misunderstanding the task instruction, so we will add a visual instruction 534 or a training next time. And even though the participants were instructed to do the effort 535 discounting with the aim to be satisfied with their choices instead of trying to increase the 536 rewards, we cannot be sure that they did so. Nevertheless, we exceeded the largest required 537 sample size by 2.20 times, which gives our analyses high statistical power, and we adhered to the agreed upon analyses of Stage 1.

### 40 Conclusion

Effort and reward are two highly subjective concepts with relevance in daily life. 541 With each decision an individual makes, they must weigh the required effort against the 542 expected reward to decide if and how to behave in that situation. So far, effort discounting paradigms have relied on the assumption that the task that is objectively easiest is the one 544 that is preferred by everyone, and each more difficult task is simply being devalued compared to the easy one. However, effort-related traits such as NFC suggest that this is not the case. Therefore, we developed a paradigm that allows effort discounting independent of objective task load, which we tested using an n-back task. The results showed that many participants indeed preferred a more or even the most difficult n-back level. Spanning the entire sample, these preferences took the shape of a declining logistic 550 curve across n-back levels. While the subjective value declined with increasing levels, it 551 increased with better performance as measured in d', and was unaffected by the reaction 552

time. Participants with NFC scores above the median reported lower subjective task load
in and less aversion to more difficult levels. However, they did not have higher subjective
values per se, which was likely due to our choice of median split and our assumption that
these group differences would emerge in lower levels. In fact, the reaction time and
self-report data suggest that individual differences emerge especially from 3-back upwards,
emphasizing the need for a flexible effort discounting paradigm. The CAD paradigm offers
this flexibility, and we encourage future studies to question traditional assumptions in the
field of effort discounting in the light of these findings.

561 References

- 562 1. Botvinick, M. M., Huffstetler, S. & McGuire, J. T. Effort discounting in human nucleus accumbens. Cognitive, affective & behavioral neuroscience 9, 16–27 (2009).
- 564 2. Kool, W. & Botvinick, M. Mental labour. *Nature Human Behaviour* **2**, 899–908 (2018).
- Mackworth, J. F. Paced memorizing in a continuous task. *Journal of Experimental Psychology* **58**, 206–211 (1959).
- Jaeggi, S. M., Buschkuehl, M., Perrig, W. J. & Meier, B. The concurrent validity of the N-back task as a working memory measure. *Memory* 18, 394–412 (2010).
- 5. Jonides, J. et al. Verbal Working Memory Load Affects Regional Brain Activation as Measured by PET. Journal of Cognitive Neuroscience 9, 462–475 (1997).
- 6. Owen, A. M., McMillan, K. M., Laird, A. R. & Bullmore, E. N-back working memory paradigm: A meta-analysis of normative functional neuroimaging studies. *Human Brain Mapping* **25**, 46–59 (2005).
- 7. Westbrook, A., Kester, D. & Braver, T. S. What is the subjective cost of cognitive effort? Load, trait, and aging effects revealed by economic preference. *PLOS ONE* 8, e68210 (2013).
- Social Psychology 42, 116–131 (1982).

  8. Cacioppo, J. T. & Petty, R. E. The Need for Cognition. Journal of Personality and
  Social Psychology 42, 116–131 (1982).
- Wu, R., Ferguson, A. & Inzlicht, M. Do humans prefer cognitive effort over doing nothing? https://psyarxiv.com/d2gkf/ (2021) doi:10.31234/osf.io/d2gkf.
- 580 10. Bertrams, A. & Dickhäuser, O. Passionate thinkers feel better. *Journal of Individual*581 Differences **33**, 69–75 (2012).

- 11. Nishiguchi, Y., Takano, K. & Tanno, Y. The Need for Cognition mediates and moderates the association between depressive symptoms and impaired Effortful Control.

  \*Psychiatry Research 241, 8–13 (2016).
- 12. Xu, P. & Cheng, J. Individual differences in social distancing and mask-wearing in the pandemic of COVID-19: The role of need for cognition, self-control and risk attitude.

  Personality and Individual Differences 175, 110706 (2021).
- 13. Kramer, A.-W., Van Duijvenvoorde, A. C. K., Krabbendam, L. & Huizenga, H. M. Individual differences in adolescents' willingness to invest cognitive effort: Relation to need for cognition, motivation and cognitive capacity. *Cognitive Development* 57, 100978 (2021).
- 588 14. Crawford, J. L., Eisenstein, S. A., Peelle, J. E. & Braver, T. S. Domain-general cognitive motivation: Evidence from economic decision-making. *Cognitive Research:*Principles and Implications 6, 4 (2021).
- 590 15. Culbreth, A., Westbrook, A. & Barch, D. Negative symptoms are associated with an increased subjective cost of cognitive effort. *Journal of Abnormal Psychology* **125**, 528–536 (2016).
- Westbrook, A., Lamichhane, B. & Braver, T. The subjective value of cognitive effort is encoded by a domain-general valuation network. *The Journal of Neuroscience* **39**, 3934–3947 (2019).
- 594 17. Scheffel, C., Zerna, J., Gärtner, A., Dörfel, D. & Strobel, A. Estimating individual subjective values of emotion regulation strategies. (2022).
- 596 18. Simmons, J. P., Nelson, L. D. & Simonsohn, U. A 21 word solution. (2012) doi:10.2139/ssrn.2160588.
- 19. Peirce, J. et al. PsychoPy2: Experiments in behavior made easy. Behavior Research

  Methods 51, 195–203 (2019).

- 20. R Core Team. R: A language and environment for statistical computing. (R Foundation for Statistical Computing, 2020).
- 602 21. RStudio Team. RStudio: Integrated development for R. (2020).

603

- Singmann, H., Bolker, B., Westfall, J., Aust, F. & Ben-Shachar, M. S. Afex: Analysis of factorial experiments. (2021).
- Morey, R. D. & Rouder, J. N. BayesFactor: Computation of Bayes factors for common designs. (2021).
- Greiner, B. Subject pool recruitment procedures: Organizing experiments with ORSEE. Journal of the Economic Science Association 1, 114–125 (2015).
- Cacioppo, J. T., Petty, R. E. & Kao, C. F. The Efficient Assessment of Need for
   Cognition. Journal of Personality Assessment 48, 306–307 (1984).
- 26. Bless, H., Wänke, M., Bohner, G., Fellhauer, R. F. & Schwarz, N. Need for Cognition:
  Eine Skala zur Erfassung von Engagement und Freude bei Denkaufgaben. Zeitschrift
  für Sozialpsychologie 25, (1994).
- Fleischhauer, M. et al. Same or different? Clarifying the relationship of need for cognition to personality and intelligence. Personality & Social Psychology Bulletin 36, 82–96 (2010).
- Hart, S. G. & Staveland, L. E. Development of NASA-TLX (Task Load Index): Results of empirical and theoretical research. **52**, 139–183 (1988).
- Harris, P. A. et al. Research electronic data capture (REDCap)—A metadata-driven methodology and workflow process for providing translational research informatics support. Journal of Biomedical Informatics 42, 377–381 (2009).
- 620 30. Harris, P. A. et al. The REDCap consortium: Building an international community of software platform partners. Journal of Biomedical Informatics 95, 103208 (2019).

- 522 31. Faul, F., Erdfelder, E., Lang, A.-G. & Buchner, A. G\*Power 3: A flexible statistical power analysis program for the social, behavioral, and biomedical sciences. Behavior Research Methods 39, 175–191 (2007).
- Faul, F., Erdfelder, E., Buchner, A. & Lang, A.-G. Statistical power analyses using G\*Power 3.1: Tests for correlation and regression analyses. *Behavior Research Methods* 41, 1149–1160 (2009).
- Macmillan, N. A. & Creelman, C. D. Response bias: Characteristics of detection theory, threshold theory, and "nonparametric" indexes. *Psychological Bulletin* **107**, 401–413 (1990).
- Whelan, R. Effective Analysis of Reaction Time Data. The Psychological Record 58, 475–482 (2008).
- Berger, A. & Kiefer, M. Comparison of Different Response Time Outlier Exclusion

  Methods: A Simulation Study. Frontiers in Psychology 12, 2194 (2021).
- 532 36. Lachaud, C. M. & Renaud, O. A tutorial for analyzing human reaction times: How to filter data, manage missing values, and choose a statistical model. Applied Psycholinguistics 32, 389–416 (2011).
- Dutilh, G. et al. Testing theories of post-error slowing. Attention, Perception, & Psychophysics 74, 454–465 (2012).
- Houtman, F., Castellar, E. N. & Notebaert, W. Orienting to errors with and without immediate feedback. *Journal of Cognitive Psychology* **24**, 278–285 (2012).
- 39. Singmann, H. & Kellen, D. An introduction to mixed models for experimental psychology. in *New methods in cognitive psychology* 4–31 (Routledge, 2019). doi:10.4324/9780429318405-2.
- 40. Mussel, P., Ulrich, N., Allen, J. J. B., Osinsky, R. & Hewig, J. Patterns of theta oscillation reflect the neural basis of individual differences in epistemic motivation.

  Scientific Reports 6, (2016).

641		
642	41.	Enders, C. K. & Tofighi, D. Centering predictor variables in cross-sectional multilevel

models: A new look at an old issue. Psychological Methods 12, 121–138 (2007).

- 42. Lorah, J. Effect size measures for multilevel models: Definition, interpretation, and 644 TIMSS example. Large-scale Assessments in Education 6, (2018). 645
- 43. Simonsohn, U., Simmons, J. P. & Nelson, L. D. Specification curve analysis. Nature 646 Human Behaviour 4, 1208–1214 (2020). 647
- 44. Wetzels, R., Ravenzwaaij, D. van & Wagenmakers, E.-J. Bayesian analysis. 1–11 648 (2015) doi:10.1002/9781118625392.wbecp453. 649
- Cohen, J. A power primer. Psychological Bulletin 112, 155–159 (1992). 45. 650
- 46. Meule, A. Reporting and interpreting working memory performance in n-back tasks. 652 Frontiers in Psychology 8, (2017). 653
- 47. Fernández, L. M. & Vadillo, M. A. Flexibility in reaction time analysis: Many roads 654 to a false positive? Royal Society Open Science 7, 190831 (2020). 655

651

643

656

662

667

# Acknowledgements

This research is partly funded by the German Research Foundation (DFG) as part of the Collaborative Research Center (CRC) 940, and partly funded by centralized funds of the Faculty of Psychology at Technische Universität Dresden. The funders have/had no role in study design, data collection and analysis, decision to publish or preparation of the manuscript.

### **Author Contributions**

JZ, CS, and AS conceptualized the study and acquired funding. JZ and CS developed the methodology, investigated, administered the project, and wrote the software. JZ, CS, and CK did the formal analysis. JZ visualized the results. JZ and CK prepared the original draft. All authors reviewed, edited, and approved the final version of the manuscript.

# Competing Interests

The authors declare no competing interests.

# Figures and figure Captions

# INSERT FIGURE 1 HERE

669

670

Figure 1. An example for subjective values for an n-back task with three levels, 671 returned by different modifications of the COG-ED paradigm for a hypothetical participant 672 with the true preference 2-back > 3-back > 1-back. The grey boxes are the choice options 673 shown to the participant. The participant's final reward value of the flexible level is displayed after the first arrow. The resulting subjective value of each level is displayed after 675 the second arrow, in the notation "SV 3-back(1-back)" for the subjective value of 3-back 676 when 1-back is the other choice. The Solution and Additional Benefit panel follow the 677 same logic, but are preceded by a choice between equal rewards, and the participant's first 678 choice indicated by an exclamation mark. 679

Design Table

INSERT DESIGN TABLE HERE

Supplement

INSERT SUPPLEMENT HERE