- When easy is not preferred: A discounting paradigm to assess load-independent task 1 preference
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Abstract

When individuals set goals, they consider the subjective value (SV) of the anticipated 17 reward and the required effort, a trade-off that is of great interest to psychological research. 18 One approach to quantify the SVs of levels of a cognitive task is the Cognitive Effort 19 Discounting Paradigm by Westbrook and colleagues (2013). However, it fails to acknowledge the highly subjective nature of effort, as it assumes a unidirectional, inverse 21 relationship between task load and SVs. Therefore, it cannot map differences in effort perception that arise from traits like Need for Cognition, since individuals who enjoy 23 effortful cognitive activities likely do not prefer the easiest level. We replicated the analysis of Westbrook and colleagues with our adaptation, the Cognitive and Affective Discounting (CAD) Paradigm, which quantifies SVs without assuming that the easiest level is preferred, thereby enabling the quantification of SVs for tasks without objective order of task load. Results show that many participants preferred a more or the most difficult level. Variance in SVs was best explained by a declining logistic contrast of the n-back levels and by the 29 accuracy of responses, while reaction time as a predictor was highly volatile depending on 30 the preprocessing pipeline. Participants with higher Need for Cognition scores perceived 31 higher n-back levels as less effortful and found them less aversive. Effects of Need for Cognition on SVs in lower levels did not reach significance, as group differences only 33 emerged in higher levels. The CAD Paradigm appears to be well suited for assessing and 34 analysing task preferences independent of the supposed objective task difficulty. 35 Keywords: effort discounting, registered report, specification curve analysis, need for 36

38 Word count: 7000

cognition, n-back

When easy is not preferred: A discounting paradigm to assess load-independent task preference

Introduction

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In everyday life, effort and reward are closely intertwined¹. With each decision a
person makes, they have to evaluate whether the effort required to reach a goal is worth
being exerted, given the reward they receive when reaching the goal. A reward is
subjectively more valuable if it is obtained with less effort, so the required effort is used as
a reference point for estimating the reward value¹. However, the cost of the effort itself is
also subjective, and research has not yet established which function best describes the
relationship between effort and cost². Investigating effort and cost is challenging because
"effort is not a property of the target task alone, but also a function of the individual's
cognitive capacities, as well as the degree of effort voluntarily mobilized for the task, which
in turn is a function of the individual's reward sensitivity" (p. 209)².

One task that is often used to investigate effort is the *n*-back task, a working memory task in which a continuous stream of stimuli, e.g. letters, is presented on screen.

Participants indicate via button press whether the current stimulus is the same as *n* stimuli before, with *n* being the level of difficulty between one and six³. The *n*-back task is well suited to investigate effort because it is an almost continuous manipulation of task load as has been shown by monotonic increases in error rates, reaction times⁴, and brain activity in areas associated with working memory^{5,6}. However, its reliability measures are mixed, and associations of *n*-back performance and measures such as executive functioning and fluid intelligence are often inconsistent⁴.

A way to quantify the subjective cost of each *n*-back level has been developed by
Westbrook, Kester, and Braver⁷, called the Cognitive Effort Discounting Paradigm
(COG-ED). First, the participants complete the *n*-back levels to familiarize themselves
with the task. Then, 1-back is compared with each more difficult level by asking the

participants to decide between receiving a fixed 2\$ for the more difficult level or the flexible starting value of 1\$ for 1-back. If they choose the more difficult level, the reward for 1-back increases by 0.50\$, if they choose 1-back, it decreases by 0.50\$. This is repeated five more times, with each adjustment of the 1-back reward being half of the previous step, while the reward for the more difficult level remains fixed at 2\$. The idea is to estimate the point of subjective equivalence, i.e., the monetary ratio at which both offers are equally preferred. The subjective value (SV) of each more difficult level is then calculated by dividing the final reward value of 1-back by the fixed 2\$ reward. Westbrook et al. used these SVs to investigate inter-individual differences in effort discounting. Younger participants showed lower effort discounting, i.e., they needed a lower monetary incentive for choosing the more difficult levels over 1-back.

The individual degree of effort discounting in the study by Westbrook et al. was also 76 associated with the participants' scores in Need for Cognition (NFC), a personality trait describing an individual's tendency to actively seek out and enjoy effortful cognitive activities⁸. Westbrook et al.⁷ conceptualized NFC as a trait measure of effortful task engagement, providing a subjective self-report of effort discounting for each participant which could then be related to the SVs as an objective measure of effort discounting. On the surface, this association stands to reason, as individuals with higher NFC are more motivated to mobilize cognitive effort because they perceive it as intrinsically rewarding. Additionally, it has been shown that individuals avoid cognitive effort only to a certain degree, possibly to retain a sense of self-control⁹, a trait more prominent in individuals with high NFC¹⁰⁻¹². However, the relation of NFC and SVs might be confounded, since other studies utilizing the COG-ED paradigm found the association of NFC and SVs to disappear after correcting for performance¹³ or found no association of NFC and SVs at all¹⁴. On the other hand, task load has been shown to be a better predictor of SVs than task performance^{7,15,16}, so more research is needed to shed light on this issue.

With the present study, we alter one fundamental assumption of the original

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COG-ED paradigm: That the easiest n-back level has the highest SV. We therefore adapted the COG-ED paradigm in a way that allows the computation of SVs for different 93 n-back levels without presuming that all individuals inherently prefer the easiest level. Since we also aim to establish this paradigm for the assessment of tasks with no objective 95 task load, e.g., emotion regulation tasks¹⁷, we call it the Cognitive and Affective Discounting Paradigm (CAD). In the present study, we validated the CAD paradigm by 97 conceptually replicating the findings of Westbrook et al. 7. Additionally, we compared the effort discounting behavior of participants regarding the n-back task and an emotion regulation task. The full results of the latter are published in a second Registered Report¹⁷. 100 The COG-ED paradigm has been applied to tasks in different domains before, showing 101 that SVs across task domains correlate¹⁴, but these tasks had an objective order of task 102 load, which is not the case for the choice of emotion regulation strategies or other 103 paradigms where there is no objective order of task load.

Our hypotheses were derived from the results of Westbrook et al.⁷. As a manipulation 105 check, we hypothesized that with increasing n-back level the (1a) the signal detection 106 parameter d' declines, while (1b) reaction time and (1c) perceived task load increase. 107 Regarding the associations of task load and effort discounting we hypothesized that (2a) 108 SVs decline with increasing n-back level, and (2b) they do so even after controlling for 109 declining task performance. And finally, we hypothesized that the CAD paradigm can show 110 inter-individual differences in effort discounting, such that participants with higher NFC 111 have (3a) lower SVs for 1-back but higher SVs for 2- and 3-back, (3b) lower perceived task 112 load across all levels, and (3c) higher aversion against 1-back but lower aversion against 2-113 and 3-back. Each hypothesis is detailed in the Design Table in the Supplementary Material. 114

115 Methods

We report how we determined our sample size, all data exclusions (if any), all manipulations, and all measures in the study^{cf. 18}. The paradigm was written and

presented using $Psychopy^{19}$. We used R with R $Studio^{20,21}$ with the main packages $afex^{22}$ and $BayesFactor^{23}$ for all our analyses.

20 Ethics information

The study protocol complies with all relevant ethical regulations and was approved by the ethics committee of the Technische Universität Dresden (reference number SR-EK-50012022). Prior to testing, written informed consent was obtained. Participants received 24€ in total or course credit for participation.

$_{125}$ Design

CAD Paradigm. Figure 1 illustrates how different modifications of the COG-ED 126 paradigm⁷ return SVs that do or do not reflect the true preference of a hypothetical 127 participant, who likes 2-back most, 3-back less, and 1-back least (for reasons of clarity there 128 are only three levels in the example). The COG-ED paradigm, which compares every more 129 difficult level with 1-back sets the SV of 1-back to 1, regardless of the response pattern. 130 Adding a comparison of the more difficult levels with each other allows the SVs of those two 131 levels to be more differentiated, but leaves the SV of 1-back unchanged. Adding those same 132 pairs again, but with the opposite assignment of fixed and flexible level, does approach the 133 true preference, but has two disadvantages. First, the SVs are still quite alike across levels 134 due to the fact that every more difficult level has only been compared with the easiest 135 level, and second, having more task levels than just three would lead to an exponential 136 increase in comparisons. Therefore, the solution lies in reducing the number of necessary comparisons by presenting only one effort discounting round for each possible pair of levels after determining for each pair which level should be fixed and which should be flexible. 139 This is determined by presenting each possible pair of levels on screen with the question "Would you prefer 1 € for level A or 1 € for level B?". Participants respond by clicking the 141 respective on-screen button. Each pair is presented three times, resulting in 18 presented

pairs, which are fully randomized in order and in the assignment of which level is on the left 143 or right of the screen. For each pair, the level that was chosen by the participant at least 144 two out of three times will be used as the level with a flexible value, which starts at $1 \in$ 145 and is changed in every iteration. The other level in the pair will be set to a fixed value of 146 2 €. Then, the effort discounting sensu Westbrook et al. begins, but with all possible pairs 147 and with the individually determined assignment of fixed and flexible level. The order in 148 which the pairs are presented is fully randomized, and each pair goes through all iteration 149 steps of adding/subtracting $0.50 \in 0.25 \in 0.13 \in 0.06 \in 0.03 \in 0.02 \in to/from the$ 150 flexible level's reward (each adjustment half of the previous one, rounded to two decimals) 151 before moving on to the next one. This procedure allows to compute SVs based on actual 152 individual preference instead of objective task load. For each pair, the SV of the flexible 153 level is 1, as it was preferred when faced with equal rewards, and the SV of the fixed level is the final reward of the flexible level divided by 2 €. Each level's "global" SV is calculated 155 as the mean of this level's SVs from all pairs in which it appeared. If the participant has a 156 clear preference for one level, this level's SV will be 1. If not, then no level's SV will be 1, 157 but each level's SV can still be interpreted as an absolute and relative value, so each 158 participant's effort discounting behaviour can still be quantified. The interpretation of SVs 159 in Westbrook et al. was "The minimum relative reward required for me to choose 1-back 160 over this level". So if the SV of 3-back was 0.6, the participant would need to be rewarded 161 with at least 60 % of what they are being offered for doing 3-back to do 1-back instead, 162 forgoing the higher reward for 3-back. In this study, the SV can be interpreted as "The 163 minimum relative reward required for me to choose any other level over this level". 164 Therefore, an SV of 1 indicates that this level is preferred over all others, while SVs lower 165 than 1 indicate that in at least one pair, a different level was preferred over this one.

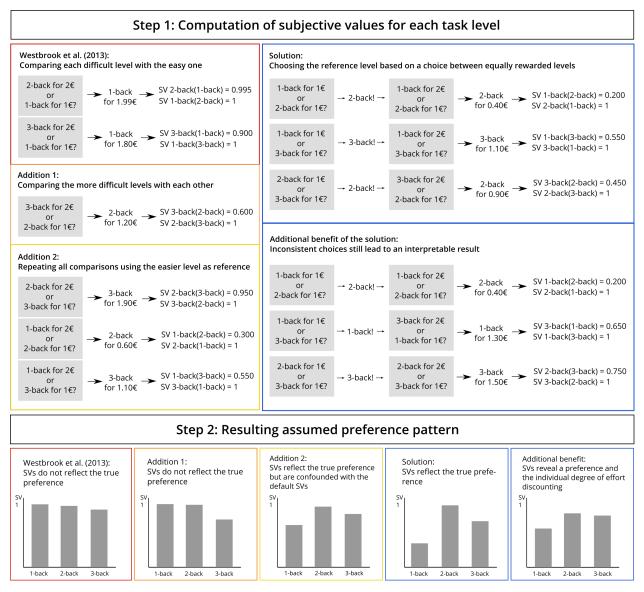


Figure 1. An example for subjective values for an n-back task with three levels, returned by different modifications of the COG-ED paradigm for a hypothetical participant with the true preference 2-back > 3-back > 1-back. The grey boxes are the choice options shown to the participant. The participant's final reward value of the flexible level is displayed after the first arrow. The resulting subjective value of each level is displayed after the second arrow, in the notation "SV 3-back(1-back)" for the subjective value of 3-back when 1-back is the other choice. The Solution and Additional Benefit panel follow the same logic, but are preceded by a choice between equal rewards, and the participant's first choice indicated by an exclamation mark.

Study procedure. Healthy participants aged 18 to 30 years were recruited using the software $ORSEE^{24}$. Participants completed the personality questionnaires online and

then visited the lab for two sessions one week apart. NFC was assessed using the 16-item short form of the Need for Cognition Scale^{25,26}. Responses to each item (e.g., "Thinking is 171 not my idea of fun", recoded) were recorded on a 7-point Likert scale. The NFC scale 172 shows comparably high internal consistency (Cronbach's $\alpha > .80$)^{26,27}. Several other 173 personality questionnaires were used in this study but are the topic of the Registered 174 Report for the second lab session¹⁷. A full list of measures can be found in our Github 175 repository. In the first session, participants provided informed consent and demographic 176 data before completing the computer-based paradigm. The paradigm started with the 177 n-back levels one to four, presented sequentially with two runs per level, consisting of 64 178 consonants (16 targets, 48 non-targets) per run. The levels were referred to by color 179 (1-back black, 2-back red, 3-back blue, 4-back green) to avoid anchor effects in the effort 180 discounting procedure. To assess perceived task load, we used the 6-item NASA Task Load Index (NASA-TLX)²⁸, where participants evaluate their subjective perception of mental 182 load, physical load, effort, frustration, performance, and time pressure during the task on a 183 20-point scale. At the end of each level, participants filled out the NASA-TLX on a tablet, 184 plus an item with the same response scale, asking them how aversive they found this 185 n-back level. After the n-back task, participants completed the CAD paradigm on screen 186 and were instructed to do so as realistically as possible, even though the displayed rewards 187 were not paid out on top of their compensation. They were told that one of their choices 188 would be randomly picked for the final run of n-back, the data of which was not analyzed 189 as it only served to incentivise truthful behavior and stay close to the design of Westbrook 190 et al.⁷. After the CAD paradigm, participants filled out a short questionnaire on the tablet, 191 indicating whether they adhered to the instructions (yes/no) and what the primary 192 motivation for their decisions during the effort discounting procedure was (avoid 193 boredom/relax/avoid effort/seek challenge/other). 194

The second session consisted of an emotion regulation task with negative pictures and the instruction to suppress facial reactions, detach cognitively from the picture content,

and distract oneself, respectively. The paradigm followed the same structure of task and
effort discounting procedure, but participants could decide which strategy they wanted to
reapply in the last block. Study data was collected and managed using REDCap electronic
data capture tools hosted at Technische Universität Dresden^{29,30}.

201 Sampling plan

Sample size determination was mainly based on the results of the analyses of 202 Westbrook et al.⁷ (see Design Table). The hypothesis that yielded the largest necessary 203 sample size was a repeated measures ANOVA with within-between interaction of NFC and 204 n-back level influencing SVs. Sample size analysis with $G^*Power^{31,32}$ indicated that we 205 should collect data from at least 72 participants, assuming $\alpha = .05$ and $\beta = .95$. However, 206 the sample size analysis for the hypotheses of the second lab session revealed a larger 207 necessary sample size of 85 participants to find an effect of d = -0.32 of emotion regulation 208 on facial muscle activity with $\alpha = .05$ and $\beta = .95$. To account for technical errors, noisy 209 physiological data, or participants who indicate that they did not follow the instructions, 210 we aimed to collect about 50% more data sets than necessary, N = 120 in total. 211

212 Analysis plan

Data collection and analysis were not performed blind to the conditions of the 213 experiments. We excluded the data of a participant from all analyses, if the participant 214 stated that they did not follow the instructions, if the investigator noted that the 215 participant misunderstood the instructions, or if the participant withdrew their consent. No data was replaced. The performance measure d' was computed as the difference of the 217 z-transformed hit rate and the z-transformed false alarm rate³³. Reaction time (RT) data 218 was trimmed by excluding all trials with responses faster than 100 ms, as the relevant 219 cognitive processes cannot have been completed before^{34,35}. Aggregated RT values were 220 described using the median and the median of absolute deviation (MAD) as robust 221

estimates of center and variability, respectively³⁶. Error- and post-error trials were
excluded, because RT in the latter is longer due to more cautious behavior^{37,38}. To test our
hypotheses, we performed a series of rmANOVAs and an MLM with orthogonal
sum-to-zero contrasts in order to meaningfully interpret results³⁹.

226 Manipulation check. Declining performance was investigated by calculating an
227 rmANOVA with six paired contrasts comparing d' between two levels of 1- to 4-back at a
228 time. Another rmANOVA with six paired contrasts was computed to compare the median
229 RT between two levels of 1- to 4-back at a time. To investigate changes in NASA-TLX
230 ratings, six rmANOVAs were computed, one for each NASA-TLX subscale, and each with
231 six paired contrasts comparing the ratings between two levels of 1- to 4-back at a time.

Subjective values. For each effort discounting round, the SV of the fixed level was 232 calculated by adding or subtracting the last adjustment of 0.02 € from the last monetary 233 value of the flexible level, depending on the participant's last choice, and dividing this value 234 by $2 \in$. This yielded an SV between 0 and 1 for the fixed compared with the flexible level, 235 while the SV of the flexible level was 1. The closer the SV of the fixed level is to 0, the 236 stronger the preference for the flexible level. All SVs of each level were averaged to compute 237 one "global" SV for each level. An rmANOVA with four different contrasts were computed 238 to investigate the association of SVs and the *n*-back levels: Declining linear (3,1,-1,-3), 239 ascending quadratic (-1,1,1,-1), declining logistic (3,2,-2,-3), and positively skewed normal (1,2,-1,-2). Depending on whether the linear or one of the other three contrasts fit the curve best, we applied a linear or nonlinear multi-level model in the next step, respectively.

To determine the influence of task performance on the association of SVs and *n*-back level, we performed MLM. We applied restricted maximum likelihood (REML) to fit the model. As an effect size measure for random effects we first calculated the intraclass correlation (ICC), which displays the proportion of variance that is explained by differences between persons. Second, we estimated a random slopes model of *n*-back level (level 1,

fixed, and random factor: 0-back, 1-back, 2-back, 3-back) predicting SV nested within
subjects. As Mussel et al. 40 could show, participants with high versus low NFC not only
have a more shallow decline in performance with higher n-back levels, but show a
demand-specific increase in EEG theta oscillations, which has been associated with mental
effort. We controlled for performance, i.e., d' (level 1, fixed factor, continuous), median RT
(level 1, fixed factor, continuous) in order to eliminate a possible influence of declining
performance on SV ratings.

$$SV \sim level + d' + medianRT + (level|subject)$$

Level-1-predictors were centered within cluster as recommended by Enders & Tofighi⁴¹. By
this, the model yields interpretable parameter estimates. If necessary, we will adjusted the
optimization algorithm to improve model fit. We visually inspected the residuals of the
model for evidence to perform model criticism. This was done by excluding all data points
with absolute standardized residuals above 3 SD. As effect size measures, we calculated
pseudo R^2 for our model and f^2 to estimate the effect of n-back level according to Lorah⁴².

The association of SVs and NFC was examined with an rmANOVA. We subtracted 261 the SV of 1- from 2-back and 2- from 3-back, yielding two SV difference scores per 262 participant. The sample was divided into participants with low and high NFC using a 263 median split. We then computed an rmANOVA with the within-factor n-back level and the 264 between-factor NFC group to determine whether there is a main effect of level and/or 265 group, and/or an interaction between level and group on the SV difference scores. Post-hoc tests were computed depending on which effect reached significance at p < .01. To ensure the validity of this association, we conducted a specification curve analysis⁴³, which included 63 possible preprocessing pipelines of the RT data. These pipelines specify which 269 transformation was applied (none, log, inverse, or square-root), which outliers were 270 excluded (none, 2, 2.5, or 3 MAD from the median, RTs below 100 or 200 ms), and across 271

which dimensions the transformations and exclusions were applied (across/within subjects and across/within *n*-back levels). The rmANOVA was run with each of the 63 pipelines, which also included our main pipeline (untransformed data, exclusion of RTs below 100 ms). The ratio of pipelines that lead to significant versus non-significant effects provides an indication of how robust the effect actually is.

The association of subjective task load with NFC was examined similarly. We 277 calculated NASA-TLX sum scores per participant per level, computed an rmANOVA with 278 the within-factor n-back level and the between-factor NFC group, and applied post-hoc 279 tests based on which effect reached significance at p < .01. And the association of 280 subjective aversiveness of the task with NFC was examined with difference scores as well, 281 since we expected this curve to mirror the SV curve, i.e. as the SV rises, the aversiveness declines, and vice versa. We subtracted the aversiveness ratings of 1- from 2-back and 2from 3-back, yielding two aversiveness difference scores per participant. Then, we 284 computed an rmANOVA with the within-factor n-back level and the between-factor NFC 285 group, and applied post-hoc tests based on which effect reached significance at p < .01. 286 The results of each analysis was assessed on the basis of both p-value and the Bayes 287 factor BF_{10} , calculated with the BayesFactor package²³ using the default prior widths of 288 the functions anovaBF, lmBF and ttestBF. We considered a BF_{10} close to or above 3/10 as 289

moderate/strong evidence for the alternative hypothesis, and a BF_{10} close to or below

.33/.10 as moderate/strong evidence for the null hypothesis⁴⁴.

292 Pilot data

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The sample of the pilot study consisted of N=15 participants (53.3% female, M=24.43~(SD=3.59) years old). One participant's data was removed because they
misunderstood the instruction. Due to a technical error the subjective task load data of
one participant was incomplete, so the hypotheses involving the NASA Task Load Index

were analyzed with n = 14 data sets. The results showed increases in subjective and
objective task load measures with higher n-back level. Importantly, SVs were lower for
higher n-back levels, but not different between 1- and 2-back, which shows that the easiest
level is not universally preferred. The LMM revealed n-back level as a reliable predictor of
SV, even after controlling for declining task performance (d' and median RT). NASA-TLX
scores were higher with higher n, and lower for the group with lower NFC scores, but NFC
and n-back level did not interact. All results are detailed in the Supplementary Material.

Data availability

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The data of this study can be downloaded from osf.io/vnj8x/.

6 Code availability

The paradigm code, the R script for analysis, and the R Markdown file used to compile this document are available at osf.io/vnj8x/.

9 Protocol registration

The Stage 1 Registered Report protocol has been approved and is available at osf.io/qa2bg/.

312 Results

313 Adjustments for Stage 2

There were two necessary adjustments of the methods. First, we failed to update the necessary sample size after the analyses changed with the first review round. Instead of the 72 subjects stated above, the largest minimum sample size was actually 53 subjects (see hypothesis 1b in the Design Table in the Supplementary Material). And secondly, we

needed to re-assign the SCA to the MLM in hypothesis 2b, to which it belonged in the initial submission. Following the advice of the reviewers, we had separated the NFC analysis from the MLM, and wanted to apply the SCA to the new NFC analysis (hypothesis 3a), as this was one of our main points of interest. However, since hypothesis 3a does not contain any RT data and the SCA is therefore only useful for the MLM in hypothesis 2b, we applied it to the MLM.

324 Sample

Data was collected between the 16th of August 2022 and the 3rd of February 2023. 325 All of the N=124 participants who filled out the online questionnaires came to the first 326 lab session. Based on the experimenters' notes, we excluded the data of seven participants 327 from analysis for misunderstanding the instruction of the n-back task, and the data of one 328 participant who reported that they confused the colours of the levels during effort 329 discounting. Our final data set therefore included N = 116 participants (83.60% female, 330 $M = 22.40 \ (SD = 3)$ years old), which is 2.2 times more than what the highest sample size 331 calculation required. 332

333 Manipulation checks

The performance measure d' did not change across n-back levels $(F(2.85, 327.28) = 0.01, p = .999, \hat{\eta}_G^2 = .000, 90\%$ CI $[.000, .000], BF_{10} = 3.31 \times 10^{-3}), but the median RT did <math>(F(2.46, 283.05) = 98.67, p < .001, \hat{\eta}_G^2 = .192, 90\%$ CI $[.130, .248], BF_{10} = 2.28 \times 10^{34})$. Specifically, the median RT was higher for the more difficult level in every contrast, with two exceptions: It did not differ between 2- and 4-back, and it was higher for 3- than for 4-back (Table 1).

Table 1 $Paired\ contrasts\ for\ the\ rmANOVA\ comparing\ the\ median\ reaction\ time\ between\ n-back\ levels$

Contrast	Estimate	SE	df	t	p	BF_{10}	η_p^2	95%CI
1 - 2	-0.12	0.01	345.00	-11.80	0.000	1.75×10^{30}	0.29	[0.22, 1.00]
1 - 3	-0.16	0.01	345.00	-16.20	0.000	8.80×10^{45}	0.43	[0.37, 1.00]
1 - 4	-0.12	0.01	345.00	-12.50	0.000	4.79×10^{34}	0.31	[0.25, 1.00]
2 - 3	-0.04	0.01	345.00	-4.47	0.000	$5,\!538.45$	0.05	[0.02, 1.00]
2 - 4	-0.01	0.01	345.00	-0.71	0.894	0.10	1.45e-03	[0.00, 1.00]
3 - 4	0.04	0.01	345.00	3.76	0.001	6.35×10^{6}	0.04	[0.01, 1.00]

Note. The column Contrast contains the n of the n-back levels. SE = standard error, df = degrees of freedom, t = t-statistic, p = p-value, CI = confidence interval.

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All NASA-TLX subscale scores increased across n-back levels. The effort subscale
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    (F(2.20, 253.06) = 203.82, p < .001, \hat{\eta}_G^2 = .316, 90\% \text{ CI } [.250, .375], \text{BF}_{10} = 2.47 \times 10^{34})
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    increased across all levels, but the magnitude of change decreased from 1- to 2-back
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    (t(345) = -12.35, p_{\text{Tukey}(4)} < .001, BF_{10} = 4.24 \times 10^{19}) to 3- to 4-back (t(345) = -2.72, 0.001)
343
    p_{\text{Tukey}(4)} = .035, BF<sub>10</sub> = 174.38). Three subscales had significant differences between all
344
    contrasts except for 3- versus 4-back: While ratings on the frustration and time subscales
345
    were higher for more difficult levels (F(2.50, 287.66) = 68.06, p < .001, \hat{\eta}_G^2 = .172, 90\% CI
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    [.112, .227], BF<sub>10</sub> = 5.26 \times 10^{15}, and F(2.21, 254.65) = 51.08, p < .001, \hat{\eta}_G^2 = .117, 90% CI
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    [.065, .168], BF_{10} = 3.94 \times 10^9, respectively), ratings on the performance subscale decreased
348
    with higher n (F(2.49, 285.97) = 95.33, p < .001, \hat{\eta}_G^2 = .241, 90\% CI [.176, .299],
349
    BF_{10} = 1.55 \times 10^{24}). Ratings on the mental subscale consistently increased across all levels
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    (F(1.99, 228.35) = 274.47, p < .001, \hat{\eta}_G^2 = .375, 90\% \text{ CI } [.309, .432], \text{ BF}_{10} = 1.64 \times 10^{43}).
    Ratings on the physical subscale were higher for more difficult levels
    (F(1.68, 192.93) = 15.91, p < .001, \hat{\eta}_G^2 = .041, 90\% \text{ CI } [.009, .075], \text{ BF}_{10} = 60.54), \text{ apart}
353
    from the contrasts 2- versus 3-back (t(345) = -2.34, p_{\text{Tukey}(4)} = .092, BF_{10} = 10.45) and 3-
354
    versus 4-back (t(345) = -1.07, p_{\text{Tukey}(4)} = .705, BF_{10} = 0.47). The full results of these
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    manipulation checks are listed in Table S.1 to S.8 in the Supplementary Material.
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Decline of subjective values

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When asking participants what motivated their decisions in the effort discounting,
11.2% stated that they wanted to avoid boredom, 22.4% stated that they wanted a
challenge, 34.5% stated that they wanted to avoid effort, and 4.3% stated that they wanted
to relax. The remaining 27.6% of participants used the free text field and provided reasons
such as "I wanted a fair relation of effort and reward.", "I wanted the fun that I had in the
more challenging levels.", "I wanted to maximize reward first and minimize effort second.",
or "I did not want to perform poorly when I was being paid for it.". Figure S.1 in the
Supplementary Material shows the different motivations in the context of the SVs per
n-back level.

The rmANOVA showed a significant difference between the SVs across n-back levels (F(1.98, 227.98) = 65.65, p < .001, $\hat{\eta}_G^2 = .288$, 90% CI [.222, .347], BF₁₀ = 1.58 × 10⁶⁴). All four pre-defined contrasts reached significance (Table 2), so a purely linear contrast can be rejected.

Table 2
Contrasts for the rmANOVA comparing the subjective values between n-back levels

Contrast	Estimate	SE	df	t	p	η_p^2	95%CI
Declining Linear	1.11	0.08	345.00	13.40	<.001	0.34	[0.28, 1.00]
Ascending Quadratic	0.15	0.04	345.00	4.14	<.001	0.05	[0.02, 1.00]
Declining Logistic	1.22	0.09	345.00	13.00	<.001	0.33	[0.26, 1.00]
Positively Skewed Normal	0.75	0.06	345.00	12.70	<.001	0.32	[0.26, 1.00]

Note. SE = standard error, df = degrees of freedom, t = t-statistic, p = p-value, CI = confidence interval.

The declining logistic contrast had the highest effect estimate (t(345) = 12.97, p < .001), suggesting a shallow decline of SVs between 1- and 2-back, and 3- and 4-back, respectively, and a steeper decline of SVs between 2- and 3-back.

Consequently, we had to adapt the multi level model to incorporate this non-linear

trend. To apply the contrast to the *n*-back levels, we had to turn the variables into a factor, with two consequences: Centered variables cannot be turned into factors, so we entered the variable level in its raw form, and factors cannot be used as random slopes, so the model is now defined as:

$$SV \sim level + d' + medianRT + (1|subject)$$

This means that the intercept still varies between subjects, but there are no random slopes anymore. To provide more than one observation per factor level, we used the two rounds per n-back level per subject, rather than n-back levels per subject. The ICC of the null model indicated that there is a correlation of r = .10 between the SVs of a subject, i.e. that 9.59% of variance in SVs can be explained by differences between participants. We did not use an optimization algorithm to improve the fit of the random intercept model. A total of 9 data points from 6 participants were excluded, because the residuals exceeded 3 SD above the mean. The results of the final model are displayed in Table 3.

Table 3
Results of the multi level model on the influence of n-back level (as a declining logistic contrast) and task performance on subjective values.

Parameter Be	eta SE	df	t-value	p-value	f^2	Random Effects (SD)
Intercept 0.9 n-back level -0.0 d' 0.0 median RT 0.0	0.04 0.02 02 0.00	800.15 798.75	-2.36 5.60	<.001 <.001	0.04	0.09

Note. SE = standard error, df = degrees of freedom, SD = standard deviation.

An exploratory ANOVA was used to compare the fit of the final model with a linear random intercept model, confirming that the two models were different from each other ($\chi^2(2) = 34.48, p < .001$), and with an Akaike Information Criterion of AIC = -492.61 and a Bayesian Information Criterion of BIC = -454.02 the declining logistic model was

superior to the linear model (AIC = -462.12, BIC = -433.18). The final model had an effect size of $f^2 = 0.64$ for the n-back levels and $f^2 = 0.04$ for d', which are considered large and small, respectively⁴⁵. This means that the n-back level explained 64.20% and d' explained 3.95% of variance in SVs relative to the unexplained variance, respectively. The beta coefficient indicated that with every 1-unit increase in d', the SV increased by 0.02. The effect size of the median RT was $f^2 = 0.00$.

To investigate the dependency of the model results on the RT preprocessing, we conducted a specification curve analysis (Figure 2).

INSERT FIGURE 2 HERE

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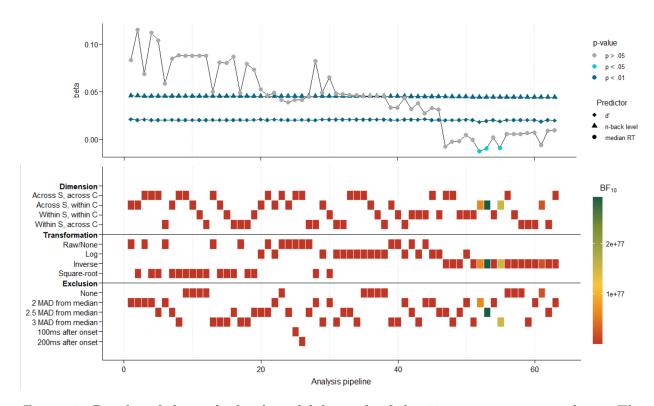


Figure 2. Results of the multi level model for each of the 63 preprocessing pipelines. The lower panel indicates the type of preprocessing, the upper panel shows the beta coefficient of each predictor and its p-value. The colourbar indicates the BF10. The pipelines are sorted in descending order of the magnitude of the n-back level beta.

Regardless of the preprocessing pipeline, n-back level and d' were significant predictors of SVs, and had stable effect estimates across all pipelines. The only pipelines in

which the median RT was a significant predictor of SVs, were the three pipelines with the
highest Bayes Factors. These three pipelines contain data that has been inverse
transformed across subjects but within conditions, i.e. within the round of an *n*-back level.

$_{405}$ Differences between NFC groups

Figure 3 shows a scatterplot of SVs per n-back level, colored depending on the 406 participant's NFC score. There is a concentration of participants who have assigned their 407 highest SV to 1-back, and this concentration fades across n-back levels. At the same time, 408 there is a subtle separation of SVs across n-back levels, depending on the participant's 409 NFC score: While the SVs of those with higher NFC scores remain elevated, the SVs of 410 those with lower NFC scores decline more strongly. Specifically, n = 71 participants had an 411 absolute preference for 1-back, n = 18 for 2-back, n = 9 for 3-back, and n = 13 for 4-back. 412 There were n=5 participants who did not have an absolute preference for any n-back 413 level, i.e. none of their SVs was 1. 414

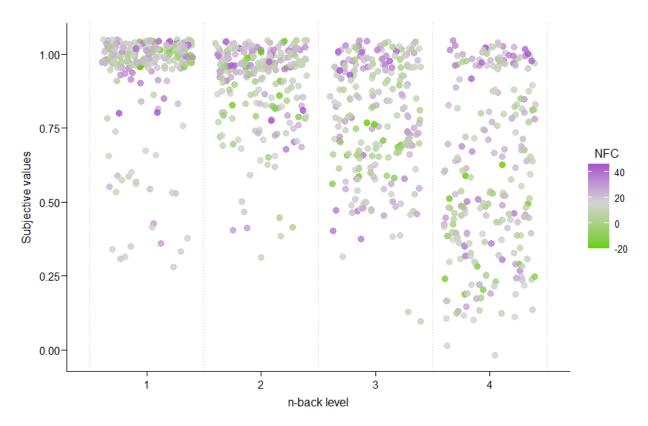


Figure 3. Subjective values per n-back level. Each dot indicates a participant, the colours indicate their Need for Cognition (NFC) score. N = 116. There is a horizontal jitter of 0.4 and a vertical jitter of 0.05 for visual clarity.

The median NFC was 16, with n = 57 subjects below and n = 59 above the median. 416 We used an rmANOVA to investigate whether the difference between the SVs of 1- and 417 2-back, and 2- and 3-back, respectively, depended on whether a participant's NFC score 418 was above or below the median. There was a main effect of the n-back level 419 $(F(1,114) = 9.13, p = .003, \hat{\eta}_G^2 = .040, 90\% \text{ CI } [.002, .115], \text{ BF}_{10} = 12.68), \text{ but neither a}$ 420 main effect of the NFC group $(F(1,114) = 3.18, p = .077, \hat{\eta}_G^2 = .013, 90\% \text{ CI } [.000, .068],$ 421 $BF_{10} = 0.56$) nor an interaction of NFC group and n-back level (F(1, 114) = 0.46, p = .499, p = .499 $\hat{\eta}_G^2 = .002, 90\%$ CI [.000, .037]). Post-hoc tests showed that the difference between the SVs of 2- and 3-back is slightly more negative than the difference between 1- and 2-back 424 (t(114) = -3.02, p = .003), but there were large inter-individual differences, especially for 425 2- and 3-back (left panel of Figure 4). This means that across the whole sample, there was

a steeper decline in SVs from 2- to 3-back than from 1- to 2-back, but some participants 427 showed a completely opposite pattern. 428

INSERT FIGURE 4 HERE

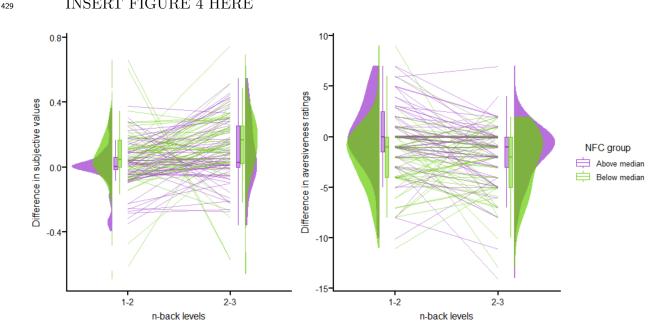


Figure 4. Difference scores for subjective values (left) and aversiveness ratings (right) when subtracting 2- from 1-back and 3- from 2-back. Each line indicates a participant, the colours indicate their Need for Cognition (NFC) score. N = 116.

The rmANOVA on the association between NFC scores and NASA-TLX scores 430 revealed a main effect of *n*-back level $(F(2.10, 239.56) = 154.50, p < .001, \hat{\eta}_G^2 = .223, 90\%$ 431 CI [.159, .282], $\mathrm{BF}_{10} = 2.22 \times 10^{45}$) and an interaction between n-back level and NFC 432 scores $(F(2.10, 239.56) = 4.93, p = .007, \hat{\eta}_G^2 = .009, 90\%$ CI [.000, .025]), but no main effect 433 of NFC scores ($F(1,114)=3.22,\,p=.075,\,\hat{\eta}_G^2=.022,\,90\%$ CI [.000, .084], 434 $BF_{10} = 1.75 \times 10^2$). Post-hoc tests showed that the participants with NFC scores below the 435 median had higher NASA-TLX scores for 3-back $(t(114) = -2.15, p = .033, BF_{10} = 11.15)$ 436 and for 4-back $(t(114) = -2.89, p = .005, BF_{10} = 336.88)$ than those with NFC scores 437 above the median. Regardless of NFC scores, NASA-TLX scores were higher for the more 438 difficult level in each pair of n-back levels (Figure 5). 439

INSERT FIGURE 5 HERE

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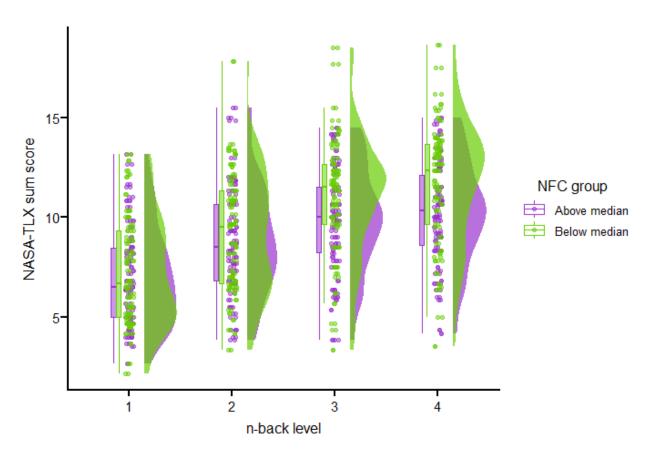


Figure 5. NASA-TLX sum scores for each n-back level. Colours indicate Need for Cognition (NFC) score above or below the median. N = 116.

With another rmANOVA we investigated whether the difference between the 441 aversiveness scores of 1- and 2-back, and 2- and 3-back, respectively, depended on whether a participant's NFC score was above or below the median. There was a main effect of NFC 443 group $(F(1,114)=8.43,\ p=.004,\ \hat{\eta}_G^2=.043,\ 90\%$ CI $[.003,.119],\ \mathrm{BF_{10}}=14.26\pm0.00\%)$ 444 and a main effect of the *n*-back level ($F(1,114)=10.21,\,p=.002,\,\hat{\eta}_G^2=.034,\,90\%$ CI 445 [.000, .105],), but no interaction $(F(1, 114) = 2.59, p = .110, \hat{\eta}_G^2 = .009, 90\%$ CI [.000, .058]). Post-hoc tests revealed that participants with NFC scores below the median reported higher aversiveness than participants with NFC scores above the median (t(114) = 2.90, p = .004) (right panel of Figure 4). Regardless of NFC, the difference of the aversiveness scores of 2- and 3-back was smaller and more negative than that of 1- and 450 2-back (t(114) = 3.20, p = .002), but again, there were large inter-individual differences. 451

The full results of these analyses of NFC group differences can be found in Table S.11 to
S.15 in the Supplementary Material.

Exploratory analysis

To investigate the apparent group difference between the SVs of participants with 455 NFC scores below and above the median in higher n-back levels, we computed an 456 rmANOVA with the within-factor level (1 to 4) and the between-factor NFC group 457 (below/above median). There was no main effect of NFC group (F(1,114) = 2.63,458 $p=.108,\,\hat{\eta}_G^2=.007,\,90\%$ CI [.000, .053], 2.95×10^{-1}), but a main effect of the *n*-back level 459 $(F(2.01,229.39)=67.39,\, p<.001,\, \hat{\eta}_G^2=.295,\, 90\%$ CI [.228, .354], $2.70\times 10^{30})$ and an 460 interaction $(F(2.01, 229.39) = 3.24, p = .041, \hat{\eta}_G^2 = .020, 90\%$ CI [.000, .044]). Post-hoc 461 tests for the main effect of level showed that Svs were lower for the more difficult n-back 462 level in each paired contrast except for 1- versus 2-back. Post-hoc tests for the interaction 463 effect showed that the NFC groups only had a significant difference in SVs for 4-back, 464 where participants below the NFC median had lower scores ($\Delta M = 0.11, 95\%$ CI 465 [0.01, 0.22], t(114) = 2.13, p = .036). Despite not reaching significance, 1-back was the only 466 level in which participants with NFC scores above the median seemed to have lower SVs than those with scores below the median ($\Delta M = -0.05, 95\%$ CI [-0.11, 0.01], t(114) = -1.50, p = .136). The full results of this exploratory analysis of NFC group differences can be found in Table S.16 and S.17 in the Supplementary Material.

471 Discussion

This Registered Report aimed to adapt the Cognitive Effort Discounting Paradigm (COG-ED) paradigm by Westbrook et al.⁷, which can estimate subjective values of different *n*-back levels, into the Cognitive and Affective Discounting (CAD) paradigm, which can estimate SVs of tasks without defaulting to the objective task load as a benchmark. For this purpose, we changed the way in which the discounting options are

presented to the participants, based the anchor on their own choices, and computed SVs across multiple combinations of task levels. The analyses were closely aligned with those in Westbrook et al.⁷ to demonstrate the changes in SVs brought about by the new paradigm.
This study also applied the CAD paradigm to an emotion regulation task, the results of which are detailed in a second Registered Report¹⁷.

482 Manipulation checks

The performance measure d' did not differ across n-back levels, but the RT increased 483 from 1- to 2- to 3-back and then remained on a high level for 4-back. This points to three important characteristics of the n-back task in this context. Firstly, RT as a valid group-level indicator of performance might only be useful for levels up to n=3, and could 486 be used to investigate inter-individual differences for n > 3. Secondly, there is a speed-accuracy tradeoff in the first three levels, that might even re-emerge in higher levels, 488 where d' would decline and RT would remain stable. And lastly, the fact that neither 489 accuracy and nor speed is an informative performance measure by itself has been observed 490 before⁴⁶ and both show different associations with various measures of intelligence⁴, 491 suggesting that they should always be reported as separate indices. Additionally, d' might 492 not have differed across n-back levels because the manipulation of task load is not strictly 493 continuous. Several participants said that they perceived 3-back as more difficult than 494 4-back because they found it is easier to remember chunks of stimuli when n was an even 495 number than when n was an odd number. 496

All NASA-TLX subscales differed across n-back levels, but the effort and mental load subscales were the only ones to consistently increase across all levels. This would support the notion of the n-back task offering a continuous manipulation of task load, as least subjectively. Ratings on the frustration and time subscales increased and ratings on the performance subscale decreased until 3-back and then remained stable. This pattern is akin to the RT, which also increased and then remained stable. Ratings on the physical

load subscale increased with *n*-back levels, but not between 2- and 3-back and 3- and 4-back, respectively.

Decline of subjective values

The rmANOVA with different pre-defined contrasts showed that all fit the SVs to a 506 different degree, and that the SVs do not simply decline linearly across n-back levels. The best fit was a declining logistic curve, reflecting that the majority of participants preferred 508 1-back and that SVs for 2-back were also high, before having more inter-individual variance 509 for 3- and 4-back. Thomson and Oppenheimer⁴⁷ argue that the different effort curves that 510 have been observed for different tasks are likely due to the fact that we still understand 511 quite little about how and why different manipulations of effort work. For example, the 512 n-back task is likely not a continuous manipulation of task load, as discussed above. 513 However, the declining logistic curve is similar to the sigmoidal curve that had the best fit 514 in a different effort paradigm⁴⁸, which the authors explained with the low effect of low 515 energy costs, suggesting there are still common features of effort across different tasks and 516 domains. The MLM with the logistic contrast showed that the n-back level explained the 517 majority of variance in SVs, while the performance measure d' also explained some variance 518 in SVs, albeit less. With increasing n-back level and decreasing d', the SV decreased. The 519 median RT was not a significant predictor in this model, which was somewhat surprising 520 because RT but not d' yielded significant differences across levels in the manipulation 521 checks. However, participants might have deliberately or subconsciously used the feedback 522 they received at the end of each round, i.e. twice per n-back level, as an anchor during the effort discounting. This feedback was based on correct responses and not on RT, so if participants based their effort discounting choices at least partly on this feedback, they 525 were either motivated to repeat a task in which they performed well and/or they were 526 reluctant to accept a larger reward for a task in which they did not perform well. Since 527 more participants reported effort avoidance as their motivation in the effort discounting 528

than those who reported seeking a challenge, we can assume that they were more motivated to repeat a task in which they performed well because their good performance coincided with low effort.

The declining logistic n-back levels and d' remained significant predictors of SVs 532 throughout all 63 preprocessing pipelines in the specification curve analysis, with betas 533 that varied by less than 0.01. In contrast to this stood the variability of the median RT betas, which ranged from about 0.11 to -0.01, and reached significance in only three pipelines. These three pipelines had the highest BF_{10} and applied inverse transformation 536 to the RT data, across subjects but within conditions, and excluded data based on the MAD. Interestingly, the curve of median RT betas in the upper panel of Figure 2 mirrored the rectangular pipeline indicators in the transformation rows of the lower panel, so the 539 transformation choice influenced the median RT much more than the dimension or the 540 exclusion choice did. As Fernandez et al. 49 found, applying more than one preprocessing 541 step to the reaction time data of a Stroop task increased the risk of false positives beyond 542 $\alpha = .05$, and transformation choices inflated this risk more than outlier exclusion or 543 aggregation choices did. Our data seems to corroborate this finding for n-back tasks as 544 well. Surprisingly, the d' betas appear almost unaffected by the preprocessing pipeline, 545 even though d' was computed after the outlier exclusion. This indicates that researchers 546 who are interested in the correctness rather than the speed of responses can choose a simple 547 preprocessing pipeline without risking false positives through elaborate transformations. 548

Differences between NFC groups

The majority of participants (61.20 %) had an absolute preference for 1-back over the other levels, but that also means that there were 34.50 % who had an absolute preference for 2-, 3-, or 4-back, and 4.30 % who preferred no specific level over all others. It shows that when given the choice, there is a large number of participants who do not prefer the easiest level, confirming the necessity of an effort discounting paradigm that works independent of

the objective task load. The CAD paradigm provides the means to depict these preferences.

Despite the visual separation of the SVs of participants with very low and high NFC 556 scores in higher n-back levels, the NFC group did not reach significance in predicting SVs. 557 This was likely due to the bandwidth of SVs of participants with NFC scores around the 558 median, and due to the fact that the difference appeared most pronounced for 4-back, and 559 we only analyzed the difference scores between 1- and 2-back and 2- and 3-back. As the exploratory analysis showed, only 4-back yielded a significant group difference, and SVs of participants with NFC scores above the median were higher for 2- to 4-back and lower for 1-back. The analysis of NASA-TLX scores showed that the sum score increased with every n-back level, and that participants with NFC scores below the median had higher NASA-TLX scores for 3- and 4-back than those below the median. This demonstrates that higher n-back levels have a higher discriminatory power regarding inter-individual 566 differences in subjective effort perception. This was also supported by the fact that higher 567 n-back levels were perceived as more aversive, and participants with NFC scores below the 568 median reported higher aversion than those with NFC scores above the median. 569

570 Limitations

When developing a new paradigm, it is challenging to decide on the optimal analysis 571 strategy, as every hypothesis is based on expected data patterns rather than previous 572 findings. While the Stage 1 review process made the analyses as robust as possible, there 573 were still unknown factors that should be addressed by future studies. For instance, the 574 differences between participants with higher and lower NFC should be investigated with extreme groups rather than a median split, especially in academic samples where NFC can be expected to be higher on average and more narrow in range. Additionally, we expected the SVs of participants with lower NFC scores to peak at 1-back and the SVs of those with higher scores to peak at 2-back, but the way the SVs of both groups appeared to drift 579 apart in the higher n-back levels suggests that an analysis of those levels would be more 580

fruitful in determining group differences. Future studies could create a stronger separation 581 between the concepts investigated in this study, and model the SVs and their task-related 582 influencing factors first, before looking at (non-linear) associations with personality. 583 Another important point is the instruction, not just for the n-back task, but for the effort 584 discounting as well. We had to exclude several participants for misunderstanding the task 585 instruction, so we will add a visual instruction or a training next time. And even though 586 the participants were instructed to do the effort discounting with the aim to be satisfied 587 with their choices instead of trying to increase the rewards, we cannot be sure that they did so. One might also argue that the $2 \in \text{reward range}$ was not large enough to be an 580 incentive for effort expenditure. However, findings by Bialaszek et al.⁵⁰ suggest that 590 participants are actually more sensitive to effort when the reward is small Nevertheless, we 591 exceeded the largest required sample size by 2.20 times, which gives our analyses high statistical power, and we adhered to the agreed upon analyses of Stage 1.

594 Conclusion

Effort and reward are two highly subjective concepts with relevance in daily life. 595 With each decision an individual makes, they must weigh the required effort against the 596 expected reward to decide if and how to behave in that situation. So far, effort discounting 597 paradigms have relied on the assumption that the task that is objectively easiest is the one 598 that is preferred by everyone, and each more difficult task is simply being devalued 599 compared to the easy one. However, effort-related traits such as Need for Cognition suggest that this is not the case. Therefore, we developed a paradigm that allows effort discounting independent of objective task load, which we tested using an n-back task. The results showed that many participants indeed preferred a more or even the most difficult n-back level. Spanning the entire sample, these preferences took the shape of a declining logistic curve across n-back levels. While the subjective value declined with increasing levels, it 605 increased with better performance as measured in d', and was unaffected by the reaction 606

time. Participants with Need for Cognition scores above the median reported lower 607 subjective task load in and less aversion to more difficult levels. However, they did not 608 have higher subjective values per se, which was likely due to our choice of median split and 609 our assumption that these group differences would emerge in lower levels. In fact, the 610 reaction time and self-report data suggest that individual differences emerge especially 611 from 3-back upwards, emphasizing the need for tasks with high discriminatory power and 612 effort discounting paradigms with flexible, participant-centered mechanisms. The CAD 613 paradigm offers this flexibility, and we encourage future studies to question traditional 614 assumptions in the field of effort discounting in the light of these findings, and to re-use 615 this data set for exploratory analyses. 616

References

- 1. Botvinick, M. M., Huffstetler, S. & McGuire, J. T. Effort discounting in human nucleus accumbens. Cognitive, affective & behavioral neuroscience 9, 16–27 (2009).
- Kool, W. & Botvinick, M. Mental labour. Nature Human Behaviour 2, 899–908
 (2018).
- Mackworth, J. F. Paced memorizing in a continuous task. *Journal of Experimental Psychology* **58**, 206–211 (1959).
- 4. Jaeggi, S. M., Buschkuehl, M., Perrig, W. J. & Meier, B. The concurrent validity of the N-back task as a working memory measure. *Memory* 18, 394–412 (2010).
- 5. Jonides, J. et al. Verbal Working Memory Load Affects Regional Brain Activation as Measured by PET. Journal of Cognitive Neuroscience 9, 462–475 (1997).
- 6. Owen, A. M., McMillan, K. M., Laird, A. R. & Bullmore, E. N-back working memory paradigm: A meta-analysis of normative functional neuroimaging studies. *Human Brain Mapping* **25**, 46–59 (2005).
- Westbrook, A., Kester, D. & Braver, T. S. What is the subjective cost of cognitive effort? Load, trait, and aging effects revealed by economic preference. *PLOS ONE* 8, e68210 (2013).
- 632 8. Cacioppo, J. T. & Petty, R. E. The Need for Cognition. Journal of Personality and

 Social Psychology 42, 116–131 (1982).
- 9. Wu, R., Ferguson, A. & Inzlicht, M. Do humans prefer cognitive effort over doing nothing? https://psyarxiv.com/d2gkf/ (2021) doi:10.31234/osf.io/d2gkf.
- Bertrams, A. & Dickhäuser, O. Passionate thinkers feel better. *Journal of Individual*Differences 33, 69–75 (2012).

- 11. Nishiguchi, Y., Takano, K. & Tanno, Y. The Need for Cognition mediates and moderates the association between depressive symptoms and impaired Effortful Control.

 *Psychiatry Research 241, 8–13 (2016).
- Xu, P. & Cheng, J. Individual differences in social distancing and mask-wearing in the pandemic of COVID-19: The role of need for cognition, self-control and risk attitude.
 Personality and Individual Differences 175, 110706 (2021).
- 642 13. Kramer, A.-W., Van Duijvenvoorde, A. C. K., Krabbendam, L. & Huizenga, H. M. Individual differences in adolescents' willingness to invest cognitive effort: Relation to need for cognition, motivation and cognitive capacity. Cognitive Development 57, 100978 (2021).
- 14. Crawford, J. L., Eisenstein, S. A., Peelle, J. E. & Braver, T. S. Domain-general cognitive motivation: Evidence from economic decision-making. *Cognitive Research:*Principles and Implications 6, 4 (2021).
- Culbreth, A., Westbrook, A. & Barch, D. Negative symptoms are associated with an increased subjective cost of cognitive effort. *Journal of Abnormal Psychology* 125, 528–536 (2016).
- Westbrook, A., Lamichhane, B. & Braver, T. The subjective value of cognitive effort is encoded by a domain-general valuation network. *The Journal of Neuroscience* **39**, 3934–3947 (2019).
- Scheffel, C., Zerna, J., Gärtner, A., Dörfel, D. & Strobel, A. Estimating individual subjective values of emotion regulation strategies. (2022).
- 652 18. Simmons, J. P., Nelson, L. D. & Simonsohn, U. A 21 word solution. (2012) doi:10.2139/ssrn.2160588.
- Peirce, J. et al. PsychoPy2: Experiments in behavior made easy. Behavior Research

 Methods 51, 195–203 (2019).

- 656 20. R Core Team. R: A language and environment for statistical computing. (R Foundation for Statistical Computing, 2020).
- 658 21. RStudio Team. RStudio: Integrated development for R. (2020).

- Singmann, H., Bolker, B., Westfall, J., Aust, F. & Ben-Shachar, M. S. Afex: Analysis of factorial experiments. (2021).
- Morey, R. D. & Rouder, J. N. BayesFactor: Computation of Bayes factors for common designs. (2021).
- Greiner, B. Subject pool recruitment procedures: Organizing experiments with ORSEE. Journal of the Economic Science Association 1, 114–125 (2015).
- Cacioppo, J. T., Petty, R. E. & Kao, C. F. The Efficient Assessment of Need for Cognition. Journal of Personality Assessment 48, 306–307 (1984).
- Bless, H., Wänke, M., Bohner, G., Fellhauer, R. F. & Schwarz, N. Need for Cognition:
 Eine Skala zur Erfassung von Engagement und Freude bei Denkaufgaben. Zeitschrift
 für Sozialpsychologie 25, (1994).
- Fleischhauer, M. et al. Same or different? Clarifying the relationship of need for cognition to personality and intelligence. Personality & Social Psychology Bulletin 36, 82–96 (2010).
- Hart, S. G. & Staveland, L. E. Development of NASA-TLX (Task Load Index): Results of empirical and theoretical research. **52**, 139–183 (1988).
- Harris, P. A. et al. Research electronic data capture (REDCap)—A metadata-driven methodology and workflow process for providing translational research informatics support. Journal of Biomedical Informatics 42, 377–381 (2009).
- of software platform partners. Journal of Biomedical Informatics 95, 103208 (2019).

- 578 31. Faul, F., Erdfelder, E., Lang, A.-G. & Buchner, A. G*Power 3: A flexible statistical power analysis program for the social, behavioral, and biomedical sciences. Behavior Research Methods 39, 175–191 (2007).
- Faul, F., Erdfelder, E., Buchner, A. & Lang, A.-G. Statistical power analyses using G*Power 3.1: Tests for correlation and regression analyses. *Behavior Research Methods* 41, 1149–1160 (2009).
- Macmillan, N. A. & Creelman, C. D. Response bias: Characteristics of detection theory, threshold theory, and "nonparametric" indexes. *Psychological Bulletin* **107**, 401–413 (1990).
- Whelan, R. Effective Analysis of Reaction Time Data. The Psychological Record 58, 475–482 (2008).
- Berger, A. & Kiefer, M. Comparison of Different Response Time Outlier Exclusion

 Methods: A Simulation Study. Frontiers in Psychology 12, 2194 (2021).
- Lachaud, C. M. & Renaud, O. A tutorial for analyzing human reaction times: How to filter data, manage missing values, and choose a statistical model. *Applied Psycholinguistics* **32**, 389–416 (2011).
- Dutilh, G. et al. Testing theories of post-error slowing. Attention, Perception, & Psychophysics 74, 454–465 (2012).
- Houtman, F., Castellar, E. N. & Notebaert, W. Orienting to errors with and without immediate feedback. *Journal of Cognitive Psychology* **24**, 278–285 (2012).
- 39. Singmann, H. & Kellen, D. An introduction to mixed models for experimental psychology. in *New methods in cognitive psychology* 4–31 (Routledge, 2019). doi:10.4324/9780429318405-2.
- Mussel, P., Ulrich, N., Allen, J. J. B., Osinsky, R. & Hewig, J. Patterns of theta oscillation reflect the neural basis of individual differences in epistemic motivation.

 Scientific Reports 6, (2016).

- Enders, C. K. & Tofighi, D. Centering predictor variables in cross-sectional multilevel models: A new look at an old issue. *Psychological Methods* **12**, 121–138 (2007).
- TIMSS example. Large-scale Assessments in Education 6, (2018).
- 702 43. Simonsohn, U., Simmons, J. P. & Nelson, L. D. Specification curve analysis. Nature

 Human Behaviour 4, 1208–1214 (2020).
- 44. Wetzels, R., Ravenzwaaij, D. van & Wagenmakers, E.-J. Bayesian analysis. 1–11
 (2015) doi:10.1002/9781118625392.wbecp453.
- ⁷⁰⁶ 45. Cohen, J. A power primer. *Psychological Bulletin* **112**, 155–159 (1992).

707

- Meule, A. Reporting and interpreting working memory performance in n-back tasks.

 Frontiers in Psychology 8, (2017).
- Thomson, K. S. & Oppenheimer, D. M. The "Effort Elephant" in the room: What is effort, anyway? Perspectives on Psychological Science 17, 1633–1652 (2022).
- 48. Klein-Flügge, M. C., Kennerley, S. W., Saraiva, A. C., Penny, W. D. & Bestmann, S. Behavioral modeling of human choices reveals dissociable effects of physical effort and temporal delay on reward devaluation. *PLOS Computational Biology* 11, e1004116 (2015).
- Fernández, L. M. & Vadillo, M. A. Flexibility in reaction time analysis: Many roads to a false positive? Royal Society Open Science 7, 190831 (2020).
- 50. Białaszek, W., Marcowski, P. & Ostaszewski, P. Physical and cognitive effort discounting across different reward magnitudes: Tests of discounting models. *PLOS ONE* **12**, e0182353 (2017).

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Acknowledgements

This research is partly funded by the German Research Foundation (DFG) as part of
the Collaborative Research Center (CRC) 940, and partly funded by centralized funds of
the Faculty of Psychology at Technische Universität Dresden. The funders have/had no
role in study design, data collection and analysis, decision to publish or preparation of the
manuscript.

Author Contributions

JZ and CS contributed equally to this work. JZ, CS, and AS conceptualized the
study and acquired funding. JZ and CS developed the methodology, investigated,
administered the project, and wrote the software. JZ, CS, and CK did the formal analysis.
JZ visualized the results. JZ and CK prepared the original draft. All authors reviewed,
edited, and approved the final version of the manuscript.

Competing Interests

The authors declare no competing interests.