

Application of Variational EM to Wireless system

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Machine Learning for Wireless Communications (EE798L)

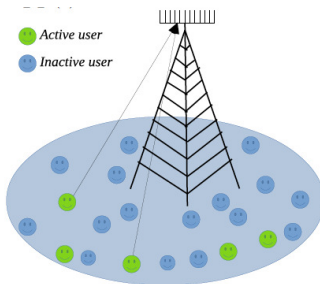
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Recap of last lecture and today's agenda

- Recap of last class
 - Discussed variational EM (VEM) algorithm
- Today's agenda
 - Apply VEM algorithm to wireless system
 - Ref: Variational Approximation for Bayesian Inference [Life after the EM algorithm], IEEE Signal Processing Magazine Nov. 2008, Dimitris G. Tzikas, Aristidis C. Likas, and Nikolaos P. Galatsanos.

5G mMTC systems model (recap)

- Consider a mMTC system with M single-antenna mMTC devices and N -antenna base-station (BS)



- Only few mMTC active devices transmit data which BS need to process
- BS does not know which devices are active. All active M mMTC devices transmit simultaneously
- Total number of mMTC devices $M \gg N$ and number of **active** mMTC devices $K < N \ll M$
- Received signal assuming all devices are active $\mathbf{y} = \mathbf{H}\mathbf{x} + \mathbf{n}$
- Tx signal $\mathbf{x} = [x_1, \dots, x_M]^T$, rx signal $\mathbf{y} = [y_1, \dots, y_N]^T$, and noise $\mathbf{n} = [n_1, \dots, n_N]^T$
- Sparse transmit vector \mathbf{x} contains only $K \ll M$ non-zero values $\mathbf{x} = [1, 0, \dots, 1, \dots, 0, 0, 0 \dots, 0]^T$

EM for 5G mMTC systems

- 5G mMTC data model is $\mathbf{y} = \mathbf{H}\mathbf{x} + \mathbf{n}$ with unknown noise variance β^{-1}
- Likelihood is Gaussian $p(\mathbf{y}|\mathbf{H}, \mathbf{x}, \beta) = \mathcal{N}(\mathbf{H}\mathbf{x}, \beta^{-1}\mathbf{I})$.
- We assumed \mathbf{x} as latent variable, and α and β as parameters
- We assumed following **sparsity-promoting** prior on latent variable \mathbf{x}

$$p(\mathbf{x}|\alpha) = \mathcal{N}(0, (\text{diag}(\alpha))^{-1}) = \prod_{m=1}^M \mathcal{N}(0, \alpha_m^{-1})$$

- First calculated posterior distribution of \mathbf{x} , which **we recall from lecture 10**, was
 - Gaussian with $p(\mathbf{x}|\mathbf{y}, \mathbf{H}, \beta) = \mathcal{N}(\boldsymbol{\mu}_{\mathbf{x}}, \boldsymbol{\Sigma}_{\mathbf{x}})$ covariance matrix and mean (**note them**)

$$\boldsymbol{\Sigma}_{\mathbf{x}} = \left(\beta \mathbf{H}^T \mathbf{H} + \text{diag}(\alpha) \right)^{-1} \text{ and } \boldsymbol{\mu}_{\mathbf{x}} = \beta \boldsymbol{\Sigma}_{\mathbf{x}} \mathbf{H}^T \mathbf{y}$$

- Computed $\mathbb{E}_{\mathbf{x}}$ of log of CDLL under pos. dis. $p(\mathbf{x}|\mathbf{y}, \mathbf{H}, \beta)$ and maximized it to calculate α , β

$$\begin{aligned} \alpha_i^{new} &= \frac{1}{\boldsymbol{\Sigma}_{\mathbf{x}}(i, i) + (\boldsymbol{\mu}_{\mathbf{x}}(i))^2} \\ (\beta^{new})^{-1} &= \frac{1}{N} (||\mathbf{y} - \mathbf{H}\boldsymbol{\mu}_{\mathbf{x}}||^2 + \text{Tr}[\mathbf{H}^T \mathbf{H} \boldsymbol{\Sigma}_{\mathbf{x}}]) \end{aligned}$$

Variational EM for 5G mMTC systems (1)

- Our 5G mMTC data model is $\mathbf{y} = \mathbf{H}\mathbf{x} + \mathbf{n}$ with unknown noise variance β^{-1}
- Likelihood is Gaussian $p(\mathbf{y}|\mathbf{H}, \mathbf{x}, \beta) = \mathcal{N}(\mathbf{H}\mathbf{x}, \beta^{-1}\mathbf{I})$
- We now not only assume \mathbf{x} as latent variable, but also α and β
 - Allows freedom in system design, which further encourages sparse solution
- We still assume the same prior on \mathbf{x}

$$p(\mathbf{x}|\alpha) = \mathcal{N}(0, (\text{diag}(\alpha))^{-1}) = \prod_{m=1}^M \mathcal{N}(0, \alpha_m^{-1})$$

- Assume Gamma prior over α and β , which we recall is conjugate for the precision of a Gaussian

$$\begin{aligned} p(\alpha) &\stackrel{(a)}{=} \prod_{m=1}^M \text{Gamma}(\alpha_m|a, b) = \prod_{m=1}^M \frac{b^a}{\Gamma(a)} (\alpha_m)^{a-1} e^{-b\alpha_m} \\ p(\beta) &= \text{Gamma}(\beta|c, d) \end{aligned}$$

- Equality (a) is because we assume each α_m is independently distributed
- a, b, c, d are parameter, but we fix them, and do not calculate them
- Treat \mathbf{x} , α and β as latent variables, such that $\mathbf{Z} = \{\mathbf{x}, \alpha, \beta\}$ and $\theta = \{\}$

Variational EM for 5G mMTC systems (2)

- If the current estimate for the parameters is denoted θ^{old} , then variational EM algorithm is

- 1 E step: use current parameter values θ^{old} to find posterior of latent variables

$$p(\mathbf{Z}|\mathbf{X}, \theta^{old}) = q^*(\mathbf{Z}) = \prod_{i=1}^M q_i^*(\mathbf{Z}_i), \text{ where}$$
$$\log q_j^*(\mathbf{Z}_j) = \mathbb{E}_{i \neq j} [\log p(\mathbf{X}, \mathbf{Z}|\theta)] + \text{constant}$$

- 2 Use $p(\mathbf{Z}|\mathbf{X}, \theta^{old})$ to find expectation of CDLL evaluated for some general θ

$$\sum_{\mathbf{Z}} p(\mathbf{Z}|\mathbf{X}, \theta^{old}) \log p(\mathbf{X}, \mathbf{Z}|\theta) = \mathcal{Q}(\theta, \theta^{old})$$

- 3 M step: determine the revised parameter estimate θ^{new} by maximizing expected value of CDLL

$$\theta^{new} = \underset{\theta}{\operatorname{argmax}} \mathcal{Q}(\theta, \theta^{old}).$$

- Since, we do not have θ , there is only variational E step, and no M step

- Recall $\mathbf{Z} = \{\mathbf{x}, \alpha, \beta\}$ and $\theta = \{\}$, and $\mathbf{X} = \{\mathbf{y}, \mathbf{H}\}$
- We accordingly need to calculate $p(\mathbf{Z}|\mathbf{X}, \theta^{old})$, which is now joint distribution $p(\mathbf{x}, \alpha, \beta|\mathbf{y}, \mathbf{H})$
 - Its calculation can be easily shown to be intractable
- Using variational (mean field) approximation, with independent $\mathbf{x}, \alpha, \beta$, posterior is
$$p(\mathbf{x}, \alpha, \beta|\mathbf{y}, \mathbf{H}) \approx q(\mathbf{x}, \alpha, \beta) = q(\mathbf{x})q(\alpha)q(\beta)$$

Variational E step for calculating $q(\mathbf{x})$ (1)

- According to variational EM algorithm derived earlier, we have

$$\log q_j^*(\mathbf{Z}_j) = \mathbb{E}_{i \neq j} [\log p(\mathbf{X}, \mathbf{Z} | \theta)] + \text{constant}$$

- In this case, it reduces to

$$\begin{aligned} \log q(\mathbf{x}) &= \mathbb{E}_{\alpha, \beta} [\log p(\mathbf{y}, \mathbf{H}, \mathbf{x}, \alpha, \beta)] + \text{constant} \\ &= \mathbb{E}_{\alpha, \beta} [\log p(\mathbf{y} | \mathbf{H}, \mathbf{x}, \alpha, \beta) + \log p(\mathbf{x} | \alpha) + \underbrace{\log p(\alpha) + \log p(\beta)}_{\text{remove terms independent of } \mathbf{x}}] + \text{constant} \\ &= \mathbb{E}_{\alpha, \beta} [\log p(\mathbf{y} | \mathbf{H}, \mathbf{x}, \beta) + \log p(\mathbf{x} | \alpha)] + \text{constant} \\ &= \mathbb{E}_{\alpha, \beta} [\log \mathcal{N}(\mathbf{H}\mathbf{x}, \beta^{-1}\mathbf{I}) + \log \mathcal{N}(\mathbf{0}, \text{diag}(\alpha)^{-1})] \\ &= \mathbb{E}_{\alpha, \beta} \left[-\frac{\beta}{2} (\mathbf{y} - \mathbf{H}\mathbf{x})^T (\mathbf{y} - \mathbf{H}\mathbf{x}) - \frac{1}{2} \sum_{m=1}^M \alpha_m x_m^2 \right] + \text{constant} \\ &\stackrel{(a)}{=} -\frac{\mathbb{E}_{\beta}}{2} [\beta] (\mathbf{y}^T \mathbf{y} - 2\mathbf{y}^T \mathbf{H}\mathbf{x} + \mathbf{x}^T \mathbf{H}^T \mathbf{H}\mathbf{x}) - \frac{1}{2} \sum_{m=1}^M \mathbb{E}_{\alpha} [\alpha_m] x_m^2 + \text{constant} \\ &\stackrel{(b)}{=} -\frac{1}{2} \mathbf{x}^T [\mathbb{E}_{\beta} [\beta] \mathbf{H}^T \mathbf{H} + \mathbb{E}_{\alpha} [\mathbf{A}]] \mathbf{x} + \mathbb{E} [\beta] \mathbf{y}^T \mathbf{H}\mathbf{x} + \text{constant} \end{aligned}$$

Variational E step for calculating $q(\mathbf{x})$ (2)

- Here $\mathbf{A} = \text{diag}(\alpha)$
- Equality (a): Mean field approximation according to which α and β are independently distributed
- Equality (b): By rearranging the terms and by ignoring terms independent of \mathbf{x}
- We have from the last slide

$$\begin{aligned} q(\mathbf{x}) &= -\frac{1}{2} \mathbf{x}^T [\mathbb{E}_\beta[\beta] \mathbf{H}^T \mathbf{H} + \mathbb{E}_\alpha[\mathbf{A}]] \mathbf{x} + \mathbb{E}[\beta] \mathbf{y}^T \mathbf{H} \mathbf{x} + \text{constant} \\ &= -\frac{1}{2} \mathbf{x}^T [\mathbb{E}_\beta[\beta] \mathbf{H}^T \mathbf{H} + \mathbb{E}_\alpha[\mathbf{A}]] \mathbf{x} + \mathbb{E}[\beta] \mathbf{x}^T \mathbf{H}^T \mathbf{y} + \text{constant} \end{aligned}$$

- By comparing it with standard Gaussian expression,¹ we get

$$\begin{aligned} q(\mathbf{x}) &= \mathcal{N}(\mathbf{x} | \boldsymbol{\mu}_x, \boldsymbol{\Sigma}_x), \text{ with} \\ \boldsymbol{\Sigma}_x &= [\mathbb{E}_\beta[\beta] \mathbf{H}^T \mathbf{H} + \mathbb{E}_\alpha[\mathbf{A}]]^{-1}, \quad \boldsymbol{\mu}_x = \mathbb{E}_\beta[\beta] \boldsymbol{\Sigma}_x \mathbf{H}^T \mathbf{y} \end{aligned}$$

- These updates are same as the EM-SBL updates derived earlier
 - Extra expectation over α and β because we assumed them as random variables
- To calculate the expectations, we need posterior $q(\alpha)$ and $q(\beta)$ which we now calculate

¹ $-\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu}_x)^T \boldsymbol{\Sigma}_x^{-1}(\mathbf{x} - \boldsymbol{\mu}_x) = -\frac{1}{2} \mathbf{x}^T \boldsymbol{\Sigma}_x^{-1} \mathbf{x} + \mathbf{x}^T \boldsymbol{\Sigma}_x^{-1} \boldsymbol{\mu}_x + \text{constant}$

Variational E step for calculating $q(\alpha)$ (1)

- For calculating $q(\alpha)$, we again take expectation of log of CDLL

$$\begin{aligned}\log q(\alpha) &= \mathbb{E}_{\mathbf{x}, \beta}[\log p(\mathbf{y}, \mathbf{H}, \mathbf{x}, \alpha, \beta)] + \text{constant} \\ &= \mathbb{E}_{\mathbf{x}, \beta}[\log p(\mathbf{y}|\mathbf{H}, \mathbf{x}, \beta) + \log p(\mathbf{x}|\alpha) + \log p(\alpha) + \log p(\beta)] + \text{constant}\end{aligned}$$

- Recalling that $p(\alpha) = \prod_{m=1}^M \text{Gamma}(\alpha_m | a, b) = \prod_{m=1}^M \frac{b^a}{\Gamma(a)} (\alpha_m)^{a-1} e^{-b\alpha_m}$
- Also, recall $p(\mathbf{x}|\alpha) = \prod_{m=1}^M \mathcal{N}(0, \alpha_m^{-1})$. By retaining terms which are **dependent on α** , we get

$$\begin{aligned}\log q(\alpha) &= \mathbb{E}_{\mathbf{x}, \beta}[\log p(\mathbf{x}|\alpha) + \log p(\alpha)] + \text{constant} \\ &= \frac{1}{2} \left(\sum_{m=1}^M \log \alpha_m - \sum_{m=1}^M \alpha_m \mathbb{E}[|x_m|^2] \right) + (a-1) \sum_{m=1}^M \log \alpha_m - b \sum_{m=1}^M \alpha_m + \text{constant} \\ &= \left(a - \frac{1}{2} \right) \sum_{m=1}^M \log \alpha_m - \sum_{m=1}^M \alpha_m \left(\frac{1}{2} \mathbb{E}[|x_m|^2] + b \right) + \text{constant} \\ &= (\tilde{a} - 1) \sum_{m=1}^M \log \alpha_m - \sum_{m=1}^M \alpha_m \tilde{b}_m + \text{constant}\end{aligned}$$

- Here $\tilde{a} = a + 1/2$ and $\tilde{b}_m = \frac{1}{2} \mathbb{E}[|x_m|^2] + b$

Variational E step for calculating $q(\alpha)$ (2)

- We recall the standard form of a Gamma distribution

$$\begin{aligned} p(\alpha) &= \prod_{m=1}^M \text{Gamma}(\alpha_m | a, b) = \prod_{m=1}^M \frac{b^a}{\Gamma(a)} (\alpha_m)^{a-1} e^{-b\alpha_m} \\ \Rightarrow \log(p(\alpha)) &= (a-1) \sum_{m=1}^M \log \alpha_m - \sum_{m=1}^M b_m \alpha_m + \text{constant} \end{aligned} \quad (1)$$

- Our posterior expression $q(\alpha)$ from last slide is

$$\log q(\alpha) = \tilde{a} \sum_{m=1}^M \log \alpha_m - \sum_{m=1}^M \tilde{b}_m \alpha_m + \text{constant} \quad (2)$$

- Comparing (1) and (2), posterior of α is $q(\alpha) = \prod_{m=1}^M \text{Gamma}(\alpha_m | \tilde{a}, \tilde{b}_m)$, where

$$\begin{aligned} \tilde{a} &= a + \frac{1}{2} \\ \tilde{b}_m &= b + \frac{1}{2} \mathbb{E}[|x_m|^2] \end{aligned}$$

Variational E step for calculating $q(\beta)$

- Similarly, the posterior of β is $q(\beta) = \text{Gamma}(\beta | \tilde{c}, \tilde{d})$, with

$$\begin{aligned}\tilde{c} &= c + \frac{N}{2} \\ \tilde{d} &= d + \frac{1}{2} \mathbb{E}[||\mathbf{y} - \mathbf{H}\mathbf{x}||^2]\end{aligned}$$

Variational EM algorithm for 5G mMTC system

- 1 Calculate posterior $q(\mathbf{x})$ as follows

$$\begin{aligned} q(\mathbf{x}) &= \mathcal{N}(\mathbf{x} | \boldsymbol{\mu}_{\mathbf{x}}, \boldsymbol{\Sigma}_{\mathbf{x}}), \text{ with} \\ \boldsymbol{\Sigma}_{\mathbf{x}} &= [\mathbb{E}_{\beta}[\beta] \mathbf{H}^T \mathbf{H} + \mathbb{E}_{\alpha}[\mathbf{A}]]^{-1}, \quad \boldsymbol{\mu}_{\mathbf{x}} = \mathbb{E}_{\beta}[\beta] \boldsymbol{\Sigma}_{\mathbf{x}} \mathbf{H}^T \mathbf{y} \end{aligned}$$

- 2 Calculate $q(\alpha) = \prod_{m=1}^M \text{Gamma}(\alpha_m | \tilde{a}, \tilde{b}_m)$ with $\mathbb{E}_{\alpha_m}[\alpha_m] = \tilde{a} / \tilde{b}_m$
 - To calculate \tilde{b}_m , we will use posterior $q(\mathbf{x})$
- 3 Calculate posterior of β is $q(\beta) = \text{Gamma}(\beta | \tilde{c}, \tilde{d})$, with $\mathbb{E}_{\beta}[\beta] = \tilde{c} / \tilde{d}$
 - To calculate \tilde{d} , we will use posterior $q(\mathbf{x})$
- 4 Keep repeating above step till convergence

Proof of pudding is in the eating

