

Application of EM Algorithm to Wireless system

Rohit Budhiraja

Machine Learning for Wireless Communications (EE798L)

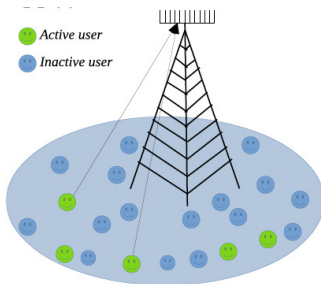
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Recap of last lecture and today's agenda

- Recap of last class
 - Discussed an alternative view of EM algorithm
 - Proved that EM maximizes log likelihood while maximizing the lower bound
- Today's agenda
 - Apply EM to 5G wireless mMTC systems - sparse Bayesian learning

5G mMTC systems model (recap)

- Consider a mMTC system with M single-antenna mMTC devices and N -antenna base-station (BS)



- Only few mMTC active devices transmit data which BS need to process
- BS does not know which devices are active. All active M mMTC devices transmit simultaneously
- Total number of mMTC devices $M \gg N$ and number of **active** mMTC devices $K < N \ll M$
- Received signal assuming all devices are active $\mathbf{y} = \mathbf{H}\mathbf{x} + \mathbf{n}$
- Tx signal $\mathbf{x} = [x_1, \dots, x_M]^T$, rx signal $\mathbf{y} = [y_1, \dots, y_N]^T$, and noise $\mathbf{n} = [n_1, \dots, n_N]^T$
- Sparse transmit vector \mathbf{x} contains only $K \ll M$ non-zero values $\mathbf{x} = [1, 0, \dots, 1, \dots, 0, 0, 0 \dots, 0]^T$

EM algorithm for SBL in 5G mMTC systems (1)

- If the current estimate of parameters is denoted θ^{old} , then EM algorithm is
 - ① E step: use current parameter values θ^{old} to find posterior of latent variables $p(\mathbf{Z}|\mathbf{X}, \theta^{old})$
 - ② Use $p(\mathbf{Z}|\mathbf{X}, \theta^{old})$ to find expectation of log of CDLL evaluated for some general θ

$$\sum_{\mathbf{Z}} p(\mathbf{Z}|\mathbf{X}, \theta^{old}) \log p(\mathbf{X}, \mathbf{Z}|\theta) = \mathcal{Q}(\theta, \theta^{old})$$

- ③ M step: determine the revised parameter estimate θ^{new} by maximizing this function

$$\theta^{new} = \underset{\theta}{\operatorname{argmax}} \mathcal{Q}(\theta, \theta^{old}).$$

- 5G mMTC data model is $\mathbf{y} = \mathbf{H}\mathbf{x} + \mathbf{n}$ with unknown noise variance β^{-1}
- We assume
 - $\mathbf{X} = \mathbf{y}$ and \mathbf{H} as observed variable
 - \mathbf{x} as latent variable \mathbf{z} with prior $p(\mathbf{x}|\alpha) = \mathcal{N}(\mathbf{0}, (\operatorname{diag}(\alpha))^{-1})$ and $\operatorname{diag}(\alpha) = \operatorname{diag}(\alpha_1, \dots, \alpha_M)$
 - Gaussian prior promotes sparsity in \mathbf{x} – Note α has M dimensions
 - Parameter $\theta = \{\alpha, \beta\}$
- EM will calculate
 - posterior distribution $p(\mathbf{x}|\mathbf{y}, \mathbf{H}, \alpha, \beta)$ in Step 1
 - expected value (using above posterior) of log of CDLL $p(\mathbf{y}, \mathbf{H}, \mathbf{x}|\alpha, \beta)$ in Step 2
 - Point estimate of the parameters by maximizing above expectation

EM algorithm Step 1 – calculation of posterior distribution

- For, 5G mMTC data model $\mathbf{y} = \mathbf{H}\mathbf{x} + \mathbf{n}$
- EM in E Step will first calculate posterior distribution $p(\mathbf{x}|\mathbf{y}, \mathbf{H}, \alpha, \beta)$
- Recall from lecture 10, the posterior density of \mathbf{x} is Gaussian with $p(\mathbf{x}|\mathbf{y}, \mathbf{H}, \beta) = \mathcal{N}(\boldsymbol{\mu}_{\mathbf{x}}, \boldsymbol{\Sigma}_{\mathbf{x}})$
 - posterior covariance matrix and mean

$$\boldsymbol{\Sigma}_{\mathbf{x}} = \left(\beta \mathbf{H}^T \mathbf{H} + \text{diag}(\alpha) \right)^{-1} \text{ and } \boldsymbol{\mu}_{\mathbf{x}} = \beta \boldsymbol{\Sigma}_{\mathbf{x}} \mathbf{H}^T \mathbf{y}$$

- We also have

$$\boldsymbol{\Sigma}_{\mathbf{x}} = \mathbb{E}_{\mathbf{x}}[\mathbf{x}\mathbf{x}^T] - \boldsymbol{\mu}_{\mathbf{x}}\boldsymbol{\mu}_{\mathbf{x}}^T \Rightarrow \mathbb{E}_{\mathbf{x}}[\mathbf{x}\mathbf{x}^T] = \boldsymbol{\Sigma}_{\mathbf{x}} + \boldsymbol{\mu}_{\mathbf{x}}\boldsymbol{\mu}_{\mathbf{x}}^T$$

- EM in E Step will then first calculate expectation of log of CDLL

$$\begin{aligned} \log p(\mathbf{y}, \mathbf{H}, \mathbf{x} | \alpha, \beta) &= \log \{ p(\mathbf{y} | \mathbf{H}, \mathbf{x}, \alpha, \beta) p(\mathbf{H}, \mathbf{x} | \alpha, \beta) \} = \log \{ p(\mathbf{y} | \mathbf{H}, \mathbf{x}, \alpha, \beta) p(\mathbf{x} | \alpha) \} \\ &= \log p(\mathbf{y} | \mathbf{H}, \mathbf{x}, \beta) + \log p(\mathbf{x} | \alpha) \end{aligned}$$

- For model $\mathbf{y} = \mathbf{H}\mathbf{x} + \mathbf{n}$, we have $p(\mathbf{y} | \mathbf{H}, \mathbf{x}, \beta) = \mathcal{N}(\mathbf{H}\mathbf{x}, \beta^{-1}\mathbf{I})$ and prior $\mathcal{N}(\mathbf{0}, \text{diag}(\alpha)^{-1})$
 - Note \mathbf{y} is N dimensional Gaussian and prior on \mathbf{x} is M dimensional Gaussian

EM algorithm Step 1 – calculation of expectation wrt posterior

- EM computes $\mathbb{E}_{\mathbf{x}}$ of log of CDLL under pos. dis. $p(\mathbf{x}|\mathbf{y}, \mathbf{H}, \beta)$ and maximizes it to calculate α, β

$$\begin{aligned}\mathbb{E}_{\mathbf{x}} \log p(\mathbf{y}, \mathbf{H}, \mathbf{x}|\alpha, \beta) &= \mathbb{E}_{\mathbf{x}}[\log p(\mathbf{y}|\mathbf{H}, \mathbf{x}, \beta) + \log p(\mathbf{x}|\alpha)] \\&= \mathbb{E}_{\mathbf{x}}[\log \mathcal{N}(\mathbf{H}\mathbf{x}, \beta^{-1}\mathbf{I}) + \log \mathcal{N}(\mathbf{0}, \text{diag}(\alpha)^{-1})] \\&= \frac{1}{2}\mathbb{E}_{\mathbf{x}} \left[N \log \beta + \sum_{i=1}^M \log \alpha_i - \beta \|\mathbf{y} - \mathbf{H}\mathbf{x}\|^2 - \mathbf{x}^T \text{diag}(\alpha) \mathbf{x} \right] \\&= \frac{1}{2}\mathbb{E}_{\mathbf{x}} \left[N \log \beta + \sum_{i=1}^M \log \alpha_i - \beta \|\mathbf{y} - \mathbf{H}\mathbf{x}\|^2 - \text{Tr}(\text{diag}(\alpha) \mathbf{x} \mathbf{x}^T) \right] \\&= \frac{1}{2} N \log \beta + \sum_{i=1}^M \log \alpha_i - \underbrace{\beta \mathbb{E}_{\mathbf{x}} [\|\mathbf{y} - \mathbf{H}\mathbf{x}\|^2]}_{T_1} - \underbrace{\mathbb{E}_{\mathbf{x}} [\text{Tr}(\text{diag}(\alpha) \mathbf{x} \mathbf{x}^T)]}_{T_2} \quad (1)\end{aligned}$$

Calc. of hyper-parameter α and noise precision β using EM (1)

- Simplifying T_1 :

$$\begin{aligned} T_1 &= \mathbb{E}_{\mathbf{x}}[||\mathbf{y} - \mathbf{H}\mathbf{x}||^2] = \mathbb{E}_{\mathbf{x}} [\mathbf{y}^T \mathbf{y} - 2\mathbf{y}^T \mathbf{H}\mathbf{x} + \mathbf{x}^T \mathbf{H}^T \mathbf{H}\mathbf{x}] \\ &= \mathbf{y}^T \mathbf{y} - 2\mathbf{y}^T \mathbf{H} \mathbb{E}_{\mathbf{x}}\{\mathbf{x}\} + \text{Tr} [\mathbf{H} \mathbb{E}_{\mathbf{x}}\{\mathbf{x}\mathbf{x}^T\} \mathbf{H}^T] \\ &= \mathbf{y}^T \mathbf{y} - 2\mathbf{y}^T \mathbf{H} \boldsymbol{\mu}_{\mathbf{x}} + \text{Tr} [(\mathbf{H}(\boldsymbol{\mu}_{\mathbf{x}}\boldsymbol{\mu}_{\mathbf{x}}^T + \boldsymbol{\Sigma}_{\mathbf{x}})\mathbf{H}^T)] \\ &= \mathbf{y}^T \mathbf{y} - 2\mathbf{y}^T \mathbf{H} \boldsymbol{\mu}_{\mathbf{x}} + \text{Tr} (\mathbf{H}\boldsymbol{\mu}_{\mathbf{x}}\boldsymbol{\mu}_{\mathbf{x}}^T \mathbf{H}^T) + \text{Tr} (\mathbf{H}\boldsymbol{\Sigma}_{\mathbf{x}}\mathbf{H}^T) \\ &= \mathbf{y}^T \mathbf{y} - 2\mathbf{y}^T \mathbf{H} \boldsymbol{\mu}_{\mathbf{x}} + \text{Tr} ((\mathbf{H}\boldsymbol{\mu}_{\mathbf{x}})^T (\mathbf{H}\boldsymbol{\mu}_{\mathbf{x}})) + \text{Tr} (\mathbf{H}\boldsymbol{\Sigma}_{\mathbf{x}}\mathbf{H}^T) \\ &= \mathbf{y}^T \mathbf{y} - 2\mathbf{y}^T \mathbf{H} \boldsymbol{\mu}_{\mathbf{x}} + (\mathbf{H}\boldsymbol{\mu}_{\mathbf{x}})^T (\mathbf{H}\boldsymbol{\mu}_{\mathbf{x}}) + \text{Tr} (\mathbf{H}\boldsymbol{\Sigma}_{\mathbf{x}}\mathbf{H}^T) \\ &= ||\mathbf{y} - \mathbf{H}\boldsymbol{\mu}_{\mathbf{x}}||^2 + \text{Tr} [\mathbf{H}^T \mathbf{H} \boldsymbol{\Sigma}_{\mathbf{x}}] \end{aligned} \tag{2}$$

- Simplifying T_2 :

$$\begin{aligned} T_2 &= \mathbb{E}_{\mathbf{x}} [\text{Tr} (\text{diag}(\alpha)\mathbf{x}\mathbf{x}^T)] = \text{Tr} [\text{diag}(\alpha)\mathbb{E}_{\mathbf{x}} (\mathbf{x}\mathbf{x}^T)] = \text{Tr} [\text{diag}(\alpha) (\boldsymbol{\mu}_{\mathbf{x}}\boldsymbol{\mu}_{\mathbf{x}}^T + \boldsymbol{\Sigma}_{\mathbf{x}})] \\ &= \text{Tr} [\boldsymbol{\mu}_{\mathbf{x}}^T \text{diag}(\alpha)\boldsymbol{\mu}_{\mathbf{x}} + \text{diag}(\alpha)\boldsymbol{\Sigma}_{\mathbf{x}}] = \sum_{i=1}^M \alpha_i ((\boldsymbol{\mu}_{\mathbf{x}}(i))^2 + \boldsymbol{\Sigma}_{\mathbf{x}}(i, i)) \end{aligned} \tag{3}$$

Calc. of hyper-parameter α and noise precision β using EM (2)

- On substituting (2) and (3) in (1), we get

$$\begin{aligned}\mathbb{E}_{\mathbf{x}}[\log p(\mathbf{y}, \mathbf{H}, \mathbf{x} | \alpha, \beta)] &= \frac{1}{2} [N \log \beta + \sum_{i=1}^M \log \alpha_i - \beta \|\mathbf{y} - \mathbf{H} \boldsymbol{\mu}_{\mathbf{x}}\|^2 - \beta \text{Tr}(\mathbf{H}^T \mathbf{H} \boldsymbol{\Sigma}_{\mathbf{x}}) \\ &\quad - \sum_{i=1}^M \alpha_i ((\boldsymbol{\mu}_{\mathbf{x}}(i))^2 + \boldsymbol{\Sigma}_{\mathbf{x}}(i, i))] \end{aligned} \quad (4)$$

- Maximizes it w.r.t to α_i by differentiating (4) and setting the result equal to zero:

$$\frac{1}{2\alpha_i^{new}} - \frac{1}{2} [\boldsymbol{\Sigma}_{\mathbf{x}}(i, i) + (\boldsymbol{\mu}_{\mathbf{x}}(i))^2] = 0 \Rightarrow \alpha_i^{new} = \frac{1}{\boldsymbol{\Sigma}_{\mathbf{x}}(i, i) + (\boldsymbol{\mu}_{\mathbf{x}}(i))^2}$$

- Maximizes it w.r.t to β by differentiating (4) w.r.t to β and setting it equal to zero:

$$\begin{aligned}\frac{N}{2\beta^{new}} - \frac{1}{2} \|\mathbf{y} - \mathbf{H} \boldsymbol{\mu}_{\mathbf{x}}\|^2 - \frac{1}{2} \text{Tr}[\mathbf{H}^T \mathbf{H} \boldsymbol{\Sigma}_{\mathbf{x}}] &= 0 \\ (\beta^{new})^{-1} &= \frac{1}{N} (\|\mathbf{y} - \mathbf{H} \boldsymbol{\mu}_{\mathbf{x}}\|^2 + \text{Tr}[\mathbf{H}^T \mathbf{H} \boldsymbol{\Sigma}_{\mathbf{x}}]) \end{aligned} \quad (5)$$

Limitations of EM algorithm

- EM assumes in E step, tractability in calculating
 - posterior distribution of latent variable $q(\mathbf{Z}) = p(\mathbf{Z}|\mathbf{X}, \theta)$
- Variational inference helps when they are not tractable
 - Bypasses the requirement of exactly knowing $p(\mathbf{Z}|\mathbf{X}, \theta)$, by assuming an appropriate $q(\mathbf{Z})$