

(1)

Number of seats = 50

Number of passengers that have the tickets = 55

random variable $Y \Rightarrow$ Number of ticketed passengers who actually show up for the flight \Rightarrow We have been provided the probability mass function of Y .(a) $P(\text{flight will accomodate all the ticketed passengers who show up}) \Rightarrow P_A$

This can be achieved if at most 50 passengers show up.

$$\begin{aligned}
 P_A &= P(Y \leq 50) \\
 &= P(Y=45) + P(Y=46) + P(Y=47) + P(Y=48) + P(Y=49) \\
 &\quad + P(Y=50) \\
 &= 0.05 + 0.10 + 0.12 + 0.14 + 0.25 + 0.17 \\
 &= \underline{\underline{0.83}} \quad \text{Ans.}
 \end{aligned}$$

(b) $P(\text{not all ticketed passengers who show up can be accommodated}) \Rightarrow P_B$

This is basically the complement of case (a) above,

So,

$$\begin{aligned}
 P_B &= 1 - P(Y \leq 50) \\
 &= 1 - P_A \\
 &= 1 - 0.83 = \underline{\underline{0.17}} \quad \text{Ans.}
 \end{aligned}$$

(c) If I am the first person on the standby list who will get on the plane if there are any seats available,

Then, that would mean that I will be given the seat when at most 49 passengers have shown up.

So, the probability of that happening would be as follows:-

$$\begin{aligned}P_c &= P(Y \leq 49) \\&= P(Y=45) + P(Y=46) + P(Y=47) + P(Y=48) + \\&\quad P(Y=49) \\&= 0.05 + 0.10 + 0.12 + 0.14 + 0.25 \\&= \underline{\underline{0.66}} \text{ Ans.}\end{aligned}$$

Similarly, if I am the third person on the standby list, then the probability for me to get the seat can be computed as when at most 47 passengers have shown up for the flight that had tickets. (3 seats should be left)

Thus,

$$\begin{aligned}P_d &= P(Y \leq 47) \\&= P(Y=45) + P(Y=46) + P(Y=47) \\&= 0.05 + 0.10 + 0.12 \\&= \underline{\underline{0.27}} \text{ Ans.}\end{aligned}$$

2.

One to five forms are required for a building permit.

Y = Number of forms required of the next applicant.

Probability that y forms are required \Rightarrow

$$\boxed{P(y) = Ky} \quad [\text{for } y=1, \dots, 5]$$

a)

To calculate the value of K ,

Sum of all the probabilities should be equal to 1.

(As one-five forms are required, the only cases possible).

$$\sum_{y=1}^5 P(y) = 1$$

$$P(Y=1) + P(Y=2) + P(Y=3) + P(Y=4) + P(Y=5) = 1$$

Putting the values of y in the function for the probability above,

$$1K + 2K + 3K + 4K + 5K = 1$$

$$15K = 1 \Rightarrow \boxed{K = \frac{1}{15}} \quad \underline{\text{Ans.}}$$

b)

for the probability that at most 3 forms are required,

$$\Rightarrow P(Y \leq 3)$$

$$\Rightarrow P(1) + P(2) + P(3)$$

$$\Rightarrow 1K + 2K + 3K = 6K$$

$$\Rightarrow 6 \times \frac{1}{15} = \frac{2}{5} = \underline{\underline{0.4}} \quad \underline{\text{Ans.}}$$

c)

for the probability that between two and four (inclusive) forms are required,

$$\Rightarrow P(2 \leq Y \leq 4) = P(2) + P(3) + P(4)$$

$$= 2K + 3K + 4K = 9K$$

$$\Rightarrow 9 \times \frac{1}{15} = \frac{9}{15} = \frac{3}{5}$$

$$= \underline{\underline{0.6}} \quad \underline{\underline{\text{Ans.}}}$$

d.

for $p(y) = y^2/50$ with y values as 1, 2, 3, 4 or 5

to be the pmf of Y ,

\Rightarrow Sum of all the probabilities in the sample space
should be equal to 1.

$$\sum_{y=1}^5 p(y) = 1 \quad [\text{should be equal to 1}]$$

$$\Rightarrow \sum_{y=1}^5 \frac{y^2}{50}$$

$$\Rightarrow \frac{1^2}{50} + \frac{2^2}{50} + \frac{3^2}{50} + \frac{4^2}{50} + \frac{5^2}{50}$$

$$\Rightarrow \frac{1}{50} [1 + 4 + 9 + 16 + 25]$$

$$\Rightarrow \frac{55}{50} = \frac{11}{10} = \underline{\underline{1.1}}$$

This value comes to be greater than 1.

Thus, this distribution is not possible and

$p(y) = y^2/50$ cannot be the pmf of Y .

Ans.

3. Actual tracking weight of a stereo cartridge is set to track at 3 g.

for a continuous rv X with pdf as shown below,

$$f(x) = \begin{cases} K[1 - (x-3)^2] & 2 \leq x \leq 4 \\ 0 & \text{otherwise} \end{cases}$$

for determining the value of K , we will assume that its a correct pdf. (sum of all values in the domain should be equal to 1).

$$\boxed{\int_{-\infty}^{\infty} f(x) dx = 1}$$

we can break the integration above as,

$$\int_{-\infty}^2 0 dx + \int_2^4 K[1 - (x-3)^2] dx + \int_4^{\infty} 0 dx = 1$$

The first and third term above will become zero.

$$\therefore \int_2^4 K[1 - (x-3)^2] dx = 1$$

$$\Rightarrow K \int_2^4 (1 - (x-3)^2) dx = 1$$

$$\Rightarrow K \left[\int_2^4 1 dx - \int_2^4 (x-3)^2 dx \right] = 1$$

$$\Rightarrow K \left[n \Big|_2^4 - \frac{(x-3)^3}{3} \Big|_2^4 \right] = 1$$

$$\Rightarrow k \left[(4-2) - \left(\frac{(4-3)^3}{3} - \frac{(2-3)^3}{3} \right) \right] = 1$$

$$\Rightarrow k \left[2 - \left(\frac{1}{3} - \frac{(-1)}{3} \right) \right] = 1$$

$$\Rightarrow k \left[2 - \left(\frac{1}{3} + \frac{1}{3} \right) \right] = 1 \Rightarrow k \left[2 - \frac{2}{3} \right] = 1$$

$$\Rightarrow k \left[\frac{6-2}{3} \right] = 1 \Rightarrow k \cdot \frac{4}{3} = 1$$

$$\frac{4k}{3} = 1 \Rightarrow k = \frac{3}{4} \Rightarrow \boxed{k = 0.75}$$

Ans.

(b) Graph of $f(x)$ can be shown as follows:-

firstly, after putting the value of k , $f(x)$ becomes:-

$$f(x) = \begin{cases} 0.75[1-(x-3)^2] & 2 \leq x \leq 4 \\ 0 & \text{otherwise} \end{cases}$$

Let's try to analyze the quadratic function above which looks like a downward facing parabola.

Let's see if it intersects the x -axis,

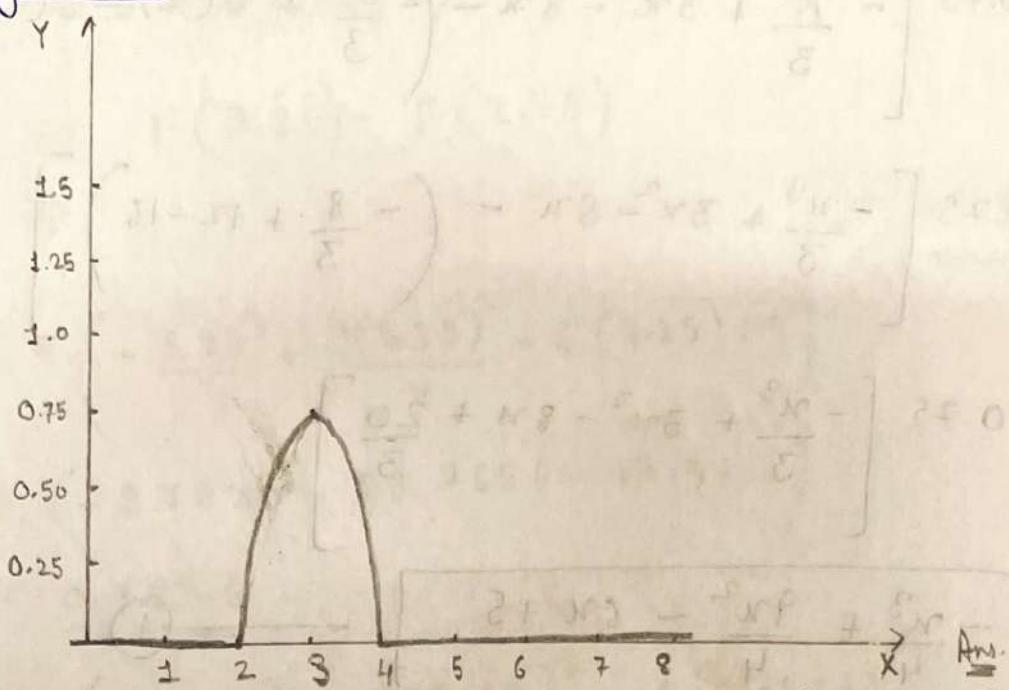
$$0.75[1-(x-3)^2] = 0 \Rightarrow (x-3)^2 = 1$$

$$x^2 - 6x + 9 - 1 = 0 \Rightarrow x^2 - 6x + 8 = 0 \Rightarrow x^2 - 4x - 2x + 8 = 0$$

$$(x-4)(x-2) = 0 \Rightarrow \underline{x=2 \text{ or } 4}$$

\therefore the graph should be a quadratic function, downward facing parabola with root at (maximum value) $x=3$ and intersects the x -axis at $x=2$ and 4 .

Graph of $f(x) \Rightarrow$



c) $P(\text{Actual tracking weight is greater than the prescribed weight}) \Rightarrow P(X > 3)$

But, for computing this, we need to find the cdf
(cumulative distribution) \Rightarrow

$$F(x) = P(X \leq x) = \int_{-\infty}^x f(y) dy$$

[changed
the
variable
for
clarity]

Thus,

$$\int_{-\infty}^x f(y) dy \Rightarrow \int_{-\infty}^2 0 dy + \int_2^x 0.75[1 - (y-3)^2] dy$$

$$\Rightarrow 0.75 \int_2^x (1 - (y-3)^2) dy$$

$$\Rightarrow 0.75 \left[\int_2^x (1 - (y^2 - 6y + 9)) dy \right]$$

$$\Rightarrow 0.75 \int_2^x (-y^2 + 6y - 8) dy \Rightarrow 0.75 \left[-\frac{y^3}{3} + \frac{6y^2}{2} - 8y \right] \Big|_2^x$$

$$\Rightarrow 0.75 \left[-\frac{x^3}{3} + 3x^2 - 8x - \left(-\frac{2^3}{3} + 3(2)^2 - 8(2) \right) \right]$$

$$\Rightarrow 0.75 \left[-\frac{x^3}{3} + 3x^2 - 8x - \left(-\frac{8}{3} + 12 - 16 \right) \right]$$

$$\Rightarrow 0.75 \left[-\frac{x^3}{3} + 3x^2 - 8x + \frac{20}{3} \right]$$

$$\Rightarrow \boxed{-\frac{x^3}{4} + \frac{9x^2}{4} - 6x + 5}$$

Ans.

Now,

for finding $P(X > 3) \Rightarrow$

$$P(X > 3) = 1 - P(X \leq 3)$$

$$= 1 - F(3)$$

$$= 1 - \left[-\frac{3^3}{4} + \frac{9(3)^2}{4} - 6(3) + 5 \right]$$

$$= 1 - \left[-\frac{27}{4} + \frac{81}{4} - 18 + 5 \right]$$

$$= 1 - \left[\frac{-27 + 81 - 52}{4} \right]$$

$$= 1 - \left[\frac{2}{4} \right] = 1 - \frac{1}{2} = \frac{1}{2} = \boxed{0.5}$$

Ans

②

$P(\text{Actual weight is between } 0.25 \text{ g of the prescribed weight of } 3 \text{ g}) \Rightarrow$

$$P(2.75 \leq X \leq 3.25) \Rightarrow$$

$$\Rightarrow F(3.25) - F(2.75)$$

\Rightarrow

[Using equation
① above]

$$F(3.25) = \left(-\frac{3.25^3}{4} + \frac{9(3.25)^2}{4} - 6(3.25) + 5 \right)$$

$$= \left(-8.5820 + 23.7656 - 19.5 + 5 \right)$$

$$= \underline{\underline{0.6836}}$$

$$F(2.75) = \left(-\frac{2.75^3}{4} + \frac{9(2.75)^2}{4} - 6(2.75) + 5 \right)$$

$$= \left(-5.1992 + 17.0156 - 16.5 + 5 \right)$$

$$= \underline{\underline{0.3164}}$$

$$P(2.75 \leq X \leq 3.25) \Rightarrow 0.6836 - 0.3164$$

$$= \underline{\underline{0.3672}}$$

Ans.

② $P(\text{Actual weight is } \overset{\text{over}}{\cancel{\text{below}}} 0.5 \text{ g of the prescribed weight of } 3 \text{ g}) \Rightarrow$ (thus, for $X < 2.5$ and $X > 3.5$)

Thus, we need to find $P(2.5 \leq X \leq 3.5)$, such that,

$$P(\{X < 2.5\} \cup \{X > 3.5\}) = P(X \leq 2.5) + P(X \geq 3.5)$$

$$= \underline{\underline{1 - P(2.5 \leq X \leq 3.5)}}$$

80, for $P(2.5 \leq X \leq 3.5) \Rightarrow F(3.5) - F(2.5)$

\Rightarrow

$$F(3.5) \Rightarrow \left[-\frac{3.5^3}{4} + \frac{9(3.5)^2}{4} - 6(3.5) + 5 \right]$$

$$= \left[-10.7188 + 27.5625 - 21 + 5 \right]$$

$$= \underline{\underline{0.8437}}$$

$$F(2.5) \Rightarrow \left[-\frac{2.5^3}{4} + \frac{9(2.5)^2}{4} - 6(2.5) + 5 \right]$$

$$= \left[-3.9063 + 14.0625 - 15 + 5 \right]$$

$$= \underline{\underline{0.1562}}$$

$$\therefore P(2.5 \leq X \leq 3.5) = F(3.5) - F(2.5)$$

$$= 0.8437 - 0.1562$$

$$= \underline{\underline{0.6875}}$$

\therefore The required probability \Rightarrow

$$P(\{X < 2.5\} \cup \{X > 3.5\})$$

$$= 1 - P(2.5 \leq X \leq 3.5)$$

$$= 1 - 0.6875$$

$$= \underline{\underline{0.3125}}$$

Ans.

④ $X \rightarrow$ amount of time a book on two-hour reserve is actually checked out.

cdf:-

$$F(x) = \begin{cases} 0 & x < 0 \\ x^2/4 & 0 \leq x < 2 \\ 1 & x \geq 2 \end{cases}$$

② $P(X \leq 1) \Rightarrow$

As we are already provided with the cumulative distribution:-

$$\text{So, } P(X \leq 1) = F(1)$$

$$= \frac{(1)^2}{4}$$

$$= \underline{\underline{0.25}} \quad \text{Ans.}$$

[As x lies in the range $0 \leq x < 2$
we take $F(x) = x^2/4$.]

③ $P(0.5 \leq X \leq 1) \Rightarrow$

As we are already provided with the cumulative distribution:-

so,

$$P(0.5 \leq X \leq 1) \Rightarrow F(1) - F(0.5)$$

According to the cdf provided above, we will take

$$F(x) = x^2/4$$

$$\Rightarrow \frac{(1)^2}{4} - \frac{(0.5)^2}{4} = \frac{1}{4} - \frac{0.25}{4} = \frac{0.75}{4}$$

$$= \underline{\underline{0.1875}} \quad \text{Ans}$$

④ for calculating $P(X > 1.5) \Rightarrow$

$$\text{we can do} \Rightarrow 1 - P(X \leq 1.5)$$

$$= 1 - F(1.5)$$

$$= 1 - \frac{(1.5)^2}{4} = 1 - 0.5625$$

(As sum of all probabilities should be 1
 $F(x) = 1$ for $x \geq 2$)

$$\Rightarrow 1 - 0.5625 = \underline{\underline{0.4375}} \quad \underline{\text{Ans.}}$$

(d.)

for calculating the median checkout duration,

$$F(\tilde{\mu}) = 0.5$$

As for $x < 0$, it's 0 and $x \geq 2$, it's 1, median should lie somewhere in $0 \leq \tilde{\mu} < 2$, so,

$$\Rightarrow \frac{\tilde{\mu}^2}{4} = 0.5$$

$$\Rightarrow \tilde{\mu}^2 = 0.5 \times 4 = 2.0$$

$$\Rightarrow \tilde{\mu} = \sqrt{2} = \underline{\underline{1.4142}} \quad \underline{\text{Ans.}}$$

(e.)

for obtaining the density function $f(x)$,

as we have already been provided with the cdf, we would need to differentiate it,

for $x < 0$,

$$f(x) = F'(x) = \frac{d}{dx}(0) = \underline{\underline{0}} \quad \underline{\text{Ans.}}$$

for $0 \leq x < 2$,

$$\begin{aligned} f(x) &= F'(x) = \frac{d}{dx}\left(\frac{x^2}{4}\right) = \frac{1}{4} \frac{d}{dx}(x^2) \\ &= \frac{1}{4} \cdot 2x = \frac{x}{2} = \underline{\underline{0.5x}} \quad \underline{\text{Ans.}} \end{aligned}$$

for $x \geq 2$,

$$f(x) = F'(x) = \frac{d}{dx}(1) = \underline{\underline{0}} \quad \underline{\text{Ans.}}$$

So,

$$f(x) = \begin{cases} 0.5x & 0 \leq x < 2 \\ 0 & \text{otherwise} \end{cases} \quad \underline{\text{Ans.}}$$

5. Probability that a goblet selected has cosmetic flaws and must be classified as "seconds" (p) \Rightarrow

$$\underline{p = 0.1} \quad [10\%]$$

a) If six goblets are randomly selected, i.e., $n = 6$.

Probability that only one of the 6 goblets selected is a "second" \Rightarrow

$$P(\text{only one goblet is a second}) \Rightarrow$$

[As this seems like a ^{binomial} ~~bernoulli~~ distribution where a goblet either is a "second" or is not, so, probability that a goblet is not a second will be $\Rightarrow 1-p = 1-0.1 = \underline{0.9}$]

For only one of six goblets to be "second", others will not be a second \Rightarrow

$$P(X=1) \Rightarrow \binom{6}{1} (0.1)^1 (0.9)^{6-1}$$

$$= \binom{6}{1} (0.1) (0.9)^5$$

$$= \frac{6!}{5! 1!} (0.1) (0.9)^5$$

$$= 6 (0.1) (0.9)^5$$

$$= 6 (0.1) (0.59049)$$

$$= \underline{\underline{0.3543}} \quad \underline{\underline{\text{Ans.}}}$$

b) If six goblets are randomly selected ($n=6$), probability that at least two are seconds \Rightarrow

$P(\text{at least two of the goblets are seconds}) \Rightarrow$

$$= P(X \geq 2)$$

$$= 1 - P(X < 2)$$

[As it will be easier
to compute
 $P(X < 2)$]

$$= 1 - \left[\underbrace{P(X=0)}_{\downarrow} + \underbrace{P(X=1)}_{\downarrow} \right]$$

\Rightarrow no goblet is a "second" only one goblet is a "second".

$$= 1 - \left[\binom{6}{0} (0.1)^0 (0.9)^{6-0} + \binom{6}{1} (0.1)^1 (0.9)^{6-1} \right]$$

$$= 1 - \left[\frac{6!}{6!0!} \cdot 1 \cdot (0.9)^6 + \frac{6!}{5!6!} \cdot (0.1)^1 (0.9)^5 \right]$$

$$= 1 - \left[(0.9)^6 + 0.3543 \right]$$

$$= 1 - [0.5314 + 0.3543] = 1 - 0.8857$$

$$= 0.1143$$

Ans.

(c) If goblets are selected one by one,

The probability that at most five must be selected to find four that are not "seconds" \Rightarrow

This would have 2 cases as follows:-

Case ①:- four goblets are selected such that none of them are not "seconds" \Rightarrow

$$P(X=0) \Rightarrow \binom{4}{0} (0.1)^0 (0.9)^{4-0} = \frac{6!}{6!0!} \cdot 1 \cdot (0.9)^4$$

$$\Rightarrow 1 \times 1 \times (0.9)^4 = (0.9)^4 = \underline{\underline{0.6561}}$$

$$\Rightarrow (0.9)^6 = \underline{\underline{0.5314}}$$

Case (2) :- five goblets are selected such that only one of them is a "second".

$$\begin{aligned} P(X=1) &= \binom{5}{1} (0.1)^1 (0.9)^{5-1} \\ &= \frac{5!}{4!1!} (0.1) (0.9)^4 \\ &= 5 (0.1) (0.9)^4 = 0.32805 \\ &\approx \underline{\underline{0.3281}} \end{aligned}$$

Adding the probabilities for the 2 cases above would give us the final probability.

$$\Rightarrow 0.6561 + 0.3281$$

Case (2):- As we have to stop at the 5th one when we get a goblet that is not a "second".

This would mean that there will be only (at most) one second in the first four goblets selected.

If there needs to be one second among the first four, that would mean that it can occur at the 1st - 4th place, thus, we will have 4 different combination for it. (and 5th one will be not a "second").

$$\therefore \text{Probability} \Rightarrow \left[\binom{4}{1} (0.1)^1 (0.9)^{4-1} \right] \times \boxed{0.9}$$

$$\Rightarrow \left[\frac{4!}{3!1!} \times (0.1)^1 (0.9)^3 \right] \times 0.9$$

$$= 4 \times 0.1 \times (0.9)^4 = \underline{\underline{0.2624}}$$

Adding the probabilities for the 2 cases above will give us the final probability \Rightarrow

$$= [0.6561 + 0.2624]$$

$$= \underline{\underline{0.9185}}$$

Ans.

(6.)

Probability that a randomly selected box of cereal has a particular prize (p) = 0.2

The event of getting a box with the prize is deemed as success here, $\zeta(x)$

As we are purchasing one box after the other until we obtain two of those prizes,

it looks like a negative binomial distribution, where n is unknown (number of trials),

probability of getting the prized box = $p = \underline{0.2}$.
and number of successes (prized boxes) needed
 $= \underline{\underline{n=2}}$.

(a)

\therefore probability that we purchase n boxes that do not have the desired prize \Rightarrow

\Rightarrow following the formula for negative binomial distribution,

$$nb(n; r, p) = nb(n; 2, 0.2)$$

$$= \binom{n+2-1}{2-1} (0.2)^2 (1-0.2)^n$$

$$= \binom{n+1}{1} (0.2)^2 (0.8)^n$$

$$= \frac{(n+1)!}{n! 1!} (0.2)^2 (0.8)^n$$

$$= \underline{\underline{(n+1)(0.2)^2(0.8)^n}}, \text{ where } n = 0, 1, 2, \dots$$

Anw

b)

The probability that we purchase 4 boxes \Rightarrow

As we stop when we have the 2nd prized box, that would mean that we come across $(4-2)=2$ non-prized boxes \Rightarrow

Using the formula for negative binomial distribution



$$nb(n; r, p) = nb(2; 2, p) = nb(2; 2, 0.2)$$

$$= \binom{2+2-1}{2-1} (0.2)^2 (0.8)^1$$

$$= \binom{3}{1} (0.2)^2 (0.8)^1$$

$$= \frac{3!}{2!1!} (0.2)^2 (0.8)^1$$

$$= 3 (0.2)^2 (0.8)^1$$

$$= \underline{\underline{0.0768}} \quad \underline{\underline{\text{Ans}}}$$

c)

The probability that we purchase at most 4 boxes \Rightarrow

i.e. we can get at most 2 non-prized boxes \Rightarrow

$$P(X \leq 2) = P(X=0) + P(X=1) + P(X=2)$$

Using the formula for negative binomial distribution,

$$P(X \leq 2) = nb(0; 2, 0.2) + nb(1; 2, 0.2) + nb(2; 2, 0.2)$$

$$\begin{aligned}
 P(X \leq 2) &= \binom{0+2-1}{2-1} (0.2)^2 (1-0.2)^0 \\
 &\quad + \binom{1+2-1}{2-1} (0.2)^2 (1-0.2)^1 \\
 &\quad + \binom{2+2-1}{2-1} (0.2)^2 (1-0.2)^2 \\
 &= 1 \times (0.2)^2 \times 1 + 2 \times (0.2)^2 \times (0.8) + 3 \times (0.2)^2 \times (0.8)^2 \\
 &= (0.2)^2 + 2(0.2)^2(0.8) + 3(0.2)^2(0.8)^2 \\
 &= 0.04 + 0.064 + 0.0768 \\
 &= \underline{\underline{0.1808}} \quad \text{Ans.}
 \end{aligned}$$

(7)

X = number of children in the family

family decides to have children until it has three children of the same gender.

Given,

$$P(B) = P(G) = 0.5$$

probability of a boy probability of a girl

We need to find the pmf. of X .

As there are 2 genders (B and G) here, that would mean that it can have at most 5 children and 5th will have to be the third child of one of the genders.

(2 of each B and G + 1 of either B/G)

And,

as it needs atleast 3 children of the same gender, X cannot be less than 3.

Thus, $\underline{3 \leq X \leq 5}$ and $X \in I$,

so, either of the 3 cases hold,

$$\underline{X = 3, 4 \text{ or } 5}$$

To calculate the pmf of X , I will have to cover those 3 cases then, (furthermore, we will divide each case in 2 more cases as having a boy is success or having a girl is success).

Case (1):- $\underline{X=3}$,

When we get all boys $\Rightarrow P_B = (0.5)^3$
(BBB) $= \underline{\underline{0.125}}$

When we get all girls $\Rightarrow P_g = (0.5)^3$
 $(G G G) = 0.125$

Adding the 2 sub-cases above,

$$P(X=3) = 0.125 + 0.125 \\ = \underline{\underline{0.25}} \text{ Ans}$$

Case ②:- $X=4$,

- ⓐ When having a boy is a success and we know that third child is a boy when we have stopped. ($r=3, k=2$) ~~K → number of successes.~~ ~~(last one is a boy)~~.

Using
negative
binomial
distribution

$r=3, n=1$, third child is a boy.

$$P_B = \binom{1+3-1}{3-1} (0.5)^3 (0.5)^1$$

$$= \frac{3!}{2!1!} (0.5)^3 (0.5)^1$$

$$= 3 (0.5)^3 (0.5)^1 = 3 (0.5)^4 \\ = \underline{\underline{0.1875}}$$

- ⓑ When having a girl is a success and we know that third child is a girl when we have stopped.

Using negative binomial distribution,

$r=3, n=1$, third child is a girl.

$$P_g = \binom{1+3-1}{3-1} (0.5)^3 (0.5)^1 = 3 (0.5)^4 \\ = \underline{\underline{0.1875}}$$

for case (2),

we will add probabilities for the 2 subcases above $(P_B + P_G)$

$$\Rightarrow 0.1875 + 0.1875$$
$$= \underline{\underline{0.375}} \quad (\underline{\underline{P(X=4)}})$$

Case (3):-

$x=5$,

- (a) When having a boy is a success and we know that third child is a boy,

Using negative binomial distribution,

$$r=3, n=2, p=0.5.$$

$$P_B = nb(2; 3, 0.5) = \binom{2+3-1}{3-1} (0.5)^3 (0.5)^2$$
$$= \binom{4}{2} (0.5)^5$$
$$= \frac{4!}{2! 2!} (0.5)^5 = 3 \times 2 \times (0.5)^5$$
$$= \underline{\underline{6 (0.5)^5}}$$

- (b) When having a girl is a success and we know that third child is a girl,

Using negative binomial distribution,

$$r=3, n=2, p=0.5$$

$$P_G = nb(2; 3, 0.5) = \binom{2+3-1}{3-1} (0.5)^3 (0.5)^2$$
$$= \frac{4!}{2! 2!} (0.5)^5 = \underline{\underline{6 (0.5)^5}}$$

for case ③,
we add the probabilities,

$$\begin{aligned}
 P(X=5) &= P_3 + P_4 \\
 &= 6(0.5)^5 + 6(0.5)^5 \\
 &= 12(0.5)^5 \\
 &= \underline{\underline{0.375}} \quad \text{Ans.}
 \end{aligned}$$

∴ The pmf can be shown as:-

X	3	4	5
P(x)	0.25	0.375	0.375

Ans

$$\frac{E(X)}{n} = \frac{3+4+5}{3} = \frac{12}{3} = 4$$

Now, having aim to find a fixed cost
Estimation of number of fixed costs
has become a major concern.

$$(E(X)-1) \cdot (E(X)) \left(\frac{1}{n}\right)$$

$$E(X^2) = \text{Var}(X) + (E(X))^2 = (E(X)-1) \cdot (E(X)) \left(\frac{1}{n}\right)$$

(8)

Information operators available = 5

Receiving requests according to a Poisson process
with rate $\lambda = 2/\text{min}$.

(a)

Let $X \Rightarrow$ a Poisson random variable signifying the number of requests in a 1-minute interval.

As the period is taken to be 1-min and

$$\underline{\lambda = 2/\text{min}}$$

$$\begin{aligned} \therefore \lambda &= \lambda t \\ &= (2)(1) = \underline{\underline{2}}. \end{aligned}$$

(a)

Probability that during a given 1-min period, the 1st operator receives no requests \Rightarrow

$$P(X=0) \Rightarrow \frac{e^{-\lambda} \lambda^0}{0!} \quad [\text{as it's a poisson process!}]$$

$$= \frac{e^{-2} 2^0}{0!} = e^{-2} = \underline{\underline{0.1353}}$$

(b)

Probability that during a given 1-min period, exactly 4 out of 5 operators receive no requests \Rightarrow

If we consider receiving no requests a success event,

then,

Using binomial distribution formulation,
and $p(\text{receiving no requests}) = 0.1353$ and $n=5$
 $K=4$.

$$\Rightarrow \binom{5}{4} (0.1353)^4 (1 - 0.1353)^1$$

$$\Rightarrow 5 (0.1353)^4 (0.8647) = 0.001448 \approx \underline{\underline{0.0015}}$$

(c)

If all the operators receive exactly the same number of requests (n),

Probability of this happening during the 1-min period \Rightarrow

As this happens for all the operators simultaneously, we can calculate the probability of this happening to one of the operators & then multiply all.

$P(\text{Operator } i \text{ receives } n \text{ requests}) \Rightarrow$

$$= \frac{e^{-\lambda} \lambda^n}{n!}$$

$$= \underline{\underline{\left(\frac{e^{-2} \cdot 2^n}{n!} \right)}}$$

$\therefore P(\text{All operators receive } n \text{ requests}) \Rightarrow$

$$= \prod_{i=1}^5 P(\text{Operator } i \text{ receives } n \text{ requests})$$

[As all are independent]

$$= [P(\text{operator } i \text{ receives } n \text{ requests})]^5$$

And as n can vary, we need to sum those.

$$= \sum_{n=0}^{\infty} \underline{\underline{\left[\frac{e^{-2} \cdot 2^n}{n!} \right]}}^5$$

(all the possibilities)

Ans.

9.
Q

(a)

$X \Rightarrow$ geometric distribution,
 $p \in (0, 1)$

Negative binomial when $r=1$,

pmf \Rightarrow

$$nb(n; 1, p) = (1-p)^n p, n=0, 1, 2, \dots$$

$n \Rightarrow$ number of failures.

$$\text{To prove: } P(X \leq n) = 1 - (1-p)^{n+1}$$

Now, we will handle the LHS:-

$$\begin{aligned} P(X \leq n) &= 1 - P(X > n) \\ &= 1 - \sum_{n=0}^{\infty} p(1-p)^n \end{aligned}$$

Using negative
binomial
distribution.

$\therefore P(X > n)$ is the
sum over all $P(X)$
from $n+1$ to ∞ .

Opening the sum (sigma) as follows:-

$$= 1 - \left[p(1-p)^{n+1} + p(1-p)^{n+2} + p(1-p)^{n+3} + \dots \right]$$

Taking $p(1-p)^{n+1}$ common from the bracket,

$$\Rightarrow 1 - p(1-p)^{n+1} \underbrace{\left[1 + (1-p) + (1-p)^2 + (1-p)^3 + \dots \right]}_{\text{Infinite geometric progression with common ratio, } (1-p)}$$

Infinite geometric progression with common ratio, $(1-p)$

As Sum of an infinite GP is given as:-

$$P(X \leq n) = 1 - p(1-p)^{n+1} \cdot \frac{1}{1-(1-p)} = 1 - (1-p)^{n+1}$$

$$\therefore \boxed{P(X \leq n) = 1 - (1-p)^{n+1}}$$

Hence, proved!
Ans.

(b) $X_n \rightarrow$ has a geometric distribution with parameter λ/n ,
 $\lambda \in (0, 1)$.

To prove:-

$$P\left(\frac{X_n}{n} \leq x\right) \rightarrow P(X \leq x)$$

for the distribution function of an exponential random variable X with parameter λ ,

for $x > 0$,

$P(X \leq x) \Rightarrow$ Cumulative distribution of exponential random variable X is. \Rightarrow

$$\boxed{(1 - e^{-\lambda x})}$$

$$\underline{F(x) = P(X \leq x) = (1 - e^{-\lambda x})}$$

we have the RHS above!

Let's try and solve the LHS part and see if it converges to the RHS.

for distribution function of $\boxed{\frac{X_n}{n}}$ ^{random variable}

where parameter is λ/n ,

and for $x > 0$,

$$\begin{aligned} P\left(\frac{X_n}{n} \leq x\right) &= P(X_n \leq nx) \\ &= \underline{1 - P(X_n > nx)} \quad \text{--- } \textcircled{1} \end{aligned}$$

\Rightarrow As the pmf of geometric distribution
is $\Rightarrow (1-p)^n p$.

for finding the cdf of the geometric distribution,
we can do that as follows:-

$$\Rightarrow \sum_{-\infty}^{\infty} (1-p)^n p$$

As $n > 0$ and we are trying to find, $P(X_n \leq x)$

$$\Rightarrow \sum_{n=0}^{\infty} (1-p)^n p$$

$$\Rightarrow P \left[(1-p)^0 + (1-p)^1 + (1-p)^2 + \dots + (1-p)^n \right]$$

$$\Rightarrow P \left[\underbrace{1 + (1-p)^1 + (1-p)^2 + \dots + (1-p)^n}_{\text{geometric progression}} \right]$$

↓
geometric progression,

the sum of a geometric progression when $r < 1$,
(here $1-p < 1$)
as $p > 0$.

$$\text{can be computed as } \Rightarrow \frac{a(1-r^n)}{(1-r)}$$

$$\Rightarrow P \left[\frac{1 - (1-p)^n}{1 - (1-p)} \right] = \underline{\underline{1 - (1-p)^n}}$$

Then,
 $P(X_n \leq n) = 1 - (1-p)^n$

So,
 $P(X_n > n) = 1 - P(X_n \leq n) = 1 - (1 - (1-p)^n)$
 $= \underline{\underline{(1-p)^n}}$.

Using the formulae here,
and putting it in equation ①,
where parameter = λ/n .

$$\text{So, } P(X_n > nx) = \left(1 - \frac{\lambda}{n}\right)^{nx} \quad [\because nx \text{ is a natural number}]$$

$$\text{So, } \Rightarrow 1 - \left(1 - \frac{\lambda}{n}\right)^{nx}$$

Now, for computing $P\left(\frac{X_n}{n} \leq x\right) \Rightarrow$

$$= \lim_{n \rightarrow \infty} \left(1 - \left(1 - \frac{\lambda}{n}\right)^{nx}\right)$$

$$= 1 - \lim_{n \rightarrow \infty} \left(1 - \frac{\lambda}{n}\right)^{nx}$$

$$\therefore \lim_{n \rightarrow \infty} \left(1 + \frac{x}{n}\right)^n = e^x$$

So,

$$= 1 - \underline{e^{-\lambda x}}$$

\therefore LHS converges to RHS when $n \rightarrow \infty$.

$$P\left(\frac{X_n}{n} \leq x\right) \text{ when } n \rightarrow \infty \text{ becomes } \Rightarrow (1 - e^{-\lambda x})$$

and $P(X \leq x) = (1 - e^{-\lambda x})$ when X is an exponential random variable.

Both of these are equal!

Hence, proved. Ans.

(c)

* Geometric distribution models in a binomial process:-

As we know, binomial distribution describes the number of successes K achieved in n trials, where probability of success is p .

There is a special case of binomial distribution, negative binomial distribution which describes the number of ~~failures~~ until observing r ~~failures~~ ^{successes}, where probability of success is p .

furthermore, geometric distribution is also a special case of negative binomial distribution, where the experiment is stopped at first ~~failure~~ ^{success} ($r=1$).

Thus, geometric distribution models the case when 1st success is achieved after $(n-1)$ trials resulted in failures and n th trial resulted in a successful one.

* Exponential distribution models in a Poisson process:-

As we know, poisson distribution deals with the number of ~~successes~~ occurrences in a fixed period of time, while the exponential distribution deals with the time between occurrences of successive events as the time flows by continuously.

Thus, the exponential distribution models the parameter of the poisson distribution which is based on the time intervals between the occurrences of the successive events. (t/n). This parameter, in turn, helps us to calculate the number of occurrences.

Also, as $n \rightarrow \infty$, the poisson distribution fits the exponential function really well and vice-versa is also true.

Ans.