

## PROBABILITY THEORY: LECTURE NOTES 7

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**Disclaimer:** These notes are created from a variety of great references in the field, most notably, Varadhan's lecture notes and Dembo's lecture notes.

**Definition 1.** A sequence of random variables  $\{X_n\}$  is said to be *stationary* if the distribution of  $\{X_1, \dots, X_n\}$  is identical to that of  $\{X_{1+k}, \dots, X_{n+k}\}$  for all  $n \geq 1$  and  $k \geq 1$ .

**Definition 2.** Let  $(\Omega, \mathcal{F}, P)$  be a probability space. A measurable mapping  $T : \Omega \rightarrow \Omega$  is said to be *measure-preserving* if  $P(T^{-1}(A)) = P(A)$  for all  $A \in \mathcal{F}$ .

A measure-preserving map naturally gives rise to a stationary sequence as follows: let  $X(\omega)$  be a random variable; define  $X_n(\omega) = X(T^n(\omega))$ ,  $n \geq 0$ , where  $T^0(\omega) := \omega$ .

To see that the above process is stationary, define  $B := \{\omega : (X_1, \dots, X_n) \in A\}$ . Then note that  $T^{-k}(B) = \{\omega : (X_{1+k}, \dots, X_{n+k}) \in A\}$ . Since  $T$  is measure-preserving, we are done.

*Remark 1.* That any stationary process of real-valued random variables is induced by such a measure-preserving transformation is a consequence of Kolmogorov's extension theorem (why?).

This part of the notes will focus on measure-preserving transformations and hence  $T$  will always be assumed to be measure-preserving.

**Definition 3.** A set  $A$  is said to be *strictly-invariant* if  $T^{-1}(A) = A$ , while a set  $A$  is said to be *invariant* if  $P(T^{-1}(A) \Delta A) = 0$ .

**Exercise 7.1.** Show the following basic properties of mappings and sets.

- (1)  $T^{-1}(\cup_i A_i) = \cup_i T^{-1}(A_i)$ .
- (2)  $T(T^{-1}(A)) = A$ ,  $T^{-1}(T(A)) \supseteq A$ .
- (3)  $A \Delta B = A^c \Delta B^c$ .
- (4)  $\cup_i (A_i \Delta B_i) \supseteq (\cup_i A_i) \Delta (\cup_i B_i)$ .
- (5)  $T^{-1}(A^c) = (T^{-1}(A))^c$ .
- (6)  $T^{-1}(A \Delta B) = T^{-1}(A) \Delta T^{-1}(B)$ .
- (7)  $A \Delta (\cup_{i=1}^{\infty} B_i) \subseteq (A \Delta B_1) \cup (\cup_{i=1}^{\infty} (B_i \Delta B_{i+1}))$

From the above properties, it is immediate that  $\mathcal{I}$  - the collection of invariant sets - is a  $\sigma$ -algebra. we will call this to be the *invariant- $\sigma$ -algebra*. Similarly, we can also define the *strictly-invariant- $\sigma$ -algebra*.

Given an invariant set  $A$ , let  $B := \cup_{n=0}^{\infty} T^{-n}(A)$ . Note that  $A \subseteq B$  and  $T^{-1}(B) = \cup_{n=1}^{\infty} T^{-n}(A) \subseteq B$ . Define  $C := \cap_{n=0}^{\infty} T^{-n}(B)$ . Note that  $T^{-1}(C) = \cap_{n=1}^{\infty} T^{-n}(B)$ ; however since  $B \cap T^{-1}(B) = T^{-1}(B)$  we have  $T^{-1}(C) = C$ . Argue that  $P(A \Delta C) = 0$ .

**Lemma 1.** *If  $X$  is  $\mathcal{I}$ -measurable then  $X(T(\omega)) = X(\omega)$  almost surely.*

*Proof.* If  $A$  is invariant then  $T^{-1}(A)$  is also invariant (use (6) in Exercise along with measure-preserving property). Therefore  $X(T(\omega))$  is also  $\mathcal{I}$ -measurable. Given two rational numbers  $p < q$  let  $A_{p,q} = \{\omega : X(\omega) < p, X(T(\omega)) > q\}$ , and let  $B_p = \{\omega : X(\omega) < p\}$ . It is clear that  $A_{p,q} \subseteq B_p \Delta T^{-1}(B_p)$  and hence  $P(A_{p,q}) = 0$ . Now the lemma follows immediately.  $\square$

**Definition 4.** A measure-preserving transformation associated with a stationary process is called *ergodic* if  $A \in \mathcal{I}$ , the invariant  $\sigma$ -algebra, implies that  $P(A) = 0$  or  $P(A) = 1$ .

**Lemma 2** (Maximal Ergodic Lemma). *Let  $X_j(\omega) = X(T^j(\omega))$ ,  $S_k = \sum_{i=0}^{k-1} X_i(\omega)$ , and  $M_k(\omega) = \max(0, S_1(\omega), \dots, S_k(\omega))$ . Then  $E(X1_{M_k > 0}) \geq 0$ .*

*Proof.* If  $j \leq k$  then  $M_k(T(\omega)) \geq S_j(T(\omega))$ , implying

$$X(\omega) \geq S_{j+1}(\omega) - M_k(T(\omega)), \quad j = 0, \dots, k$$

Therefore

$$\begin{aligned} E(X(\omega)1_{M_k > 0}) &\geq \int_{M_k > 0} \max\{S_1(\omega), \dots, S_k(\omega)\} - M_k(T(\omega)) dP \\ &= \int_{M_k > 0} M_k(\omega) - M_k(T(\omega)) dP \geq 0. \end{aligned}$$

The last inequality is due to the following observation:  $M_k(\omega) = 0$  we have  $M_k(T(\omega)) \geq 0$ . However since integrals of  $M_k(\omega)$  and  $M_k(T(\omega))$  are same (measure-preserving property of  $T$ ), the inequality follows.  $\square$

**Theorem 1** (Birkhoff's Ergodic Theorem). *For any  $X \in L^1$ ,*

$$\frac{1}{n} \sum_{m=0}^{n-1} X(T^m(\omega)) \rightarrow E(X|\mathcal{I}) \text{ a.s.}$$

*and in  $L^1$ .*

*Proof.* Since  $E(X|\mathcal{I})$  is invariant under  $T$  (see lemma above) w.l.o.g. we can center  $X$  and assume  $E(X|\mathcal{I}) = 0$ . Let  $\bar{X} = \limsup \frac{S_n}{n}$  and let  $D = \{\omega : \bar{X} > \epsilon\}$ . Since  $\bar{X}(T(\omega)) = \bar{X}(\omega)$ , we have that  $D \in \mathcal{I}$ .

Define a new sequence of random variables  $Y(\omega) = (X(\omega) - \epsilon)1_D$  and let  $U_n = Y_0 + \dots + Y_{n-1}$ . Let  $M_n(\omega) = \max(0, U_1(\omega), \dots, U_n(\omega))$ . Observe that  $M_n \uparrow$  and  $\lim_n M_n > 0$  on  $D$ . Let  $E_n = \{\omega : M_n > 0\}$ . Hence  $E_n \uparrow D$ . Since  $|Y| \leq |X| + \epsilon$  we have

$$0 \leq E(Y1_{E_n}) \rightarrow E(Y1_D).$$

where the inequality comes from Maximal ergodic lemma. Hence

$$E((X(\omega) - \epsilon)1_D) \geq 0 \implies E(E(((X(\omega) - \epsilon)1_D)|\mathcal{I})) = E(1_D E(X|\mathcal{I}) - \epsilon 1_D) \geq 0.$$

This shows that  $P(D) = 0$ , similarly working with  $-X$ , completes the almost sure convergence.

To show convergence in  $L_1$ , let  $X_M = X1_{|X| < M}$ . Almost sure convergence above and bounded convergence theorem says that

$$E \left| \frac{1}{n} \sum_{m=0}^{n-1} X_M(T^m(\omega)) - E(X_M|\mathcal{I}) \right| \rightarrow 0.$$

Let  $\hat{X}_M = X - X_M$ . Since

$$E \left| \frac{1}{n} \sum_{m=0}^{n-1} \hat{X}_M(T^m(\omega)) \right| \leq E(|\hat{X}_M|),$$

and since  $|E(\hat{X}_M|\mathcal{I})| \leq E(|\hat{X}_M|)$ , triangle inequality completes the  $L^1$  convergence.  $\square$

Given a measurable transformation  $T$ , let  $\mathcal{M}$  denote the convex set of all probability measures that is  $T$ -invariant (this could be empty).

**Theorem 2.** *A probability measure  $P \in \mathcal{M}$  is ergodic if and only if it is an extreme point of  $\mathcal{M}$ .*

*Proof.* Assume that  $P$  is ergodic and yet  $P = aP_1 + (1-a)P_2$  for  $0 < a < 1$ . Since  $P$  is ergodic, it implies that  $P_1 = P_2$  on  $\mathcal{I}$ , hence  $P_1$  and  $P_2$  are also ergodic. Let  $f$  be any bounded measurable function on  $(\omega, \mathcal{F})$ . Define

$$h(\omega) = \lim_{n \rightarrow \infty} \frac{1}{n} (f(\omega) + f(T\omega) + \cdots + f(T^{n-1}\omega))$$

when it exists. From Ergodic theorem, we know that the limit exists on a set  $E$  with  $P_1(E) = P_2(E) = 1$ . Further, from bounded convergence theorem we also know that

$$E_{P_i}(h) = \int f dP_i \quad i = 1, 2.$$

However since  $h$  is  $\mathcal{I}$  measurable and  $P_1 = P_2$  on  $\mathcal{I}$ , we see that  $\int f dP_1 = \int f dP_2$  for any bounded measurable function (in particular indicator functions). Hence  $P_1 = P_2$  on  $\mathcal{F}$ .

If  $P$  is an extreme point of  $\mathcal{M}$  and  $P$  is not ergodic, then there exists  $A \in \mathcal{I}$  with  $0 < P(A) < 1$ . Define  $P_1(E) = \frac{P(E \cap A)}{P(A)}$  and  $P_2(E) = \frac{P(E \cap A^c)}{P(A^c)}$ . Note that  $P_1, P_2$  belong to  $\mathcal{M}$  and if  $\frac{P(E \cap A)}{P(A)} = \frac{P(E \cap A^c)}{P(A^c)} \forall E \in \mathcal{F}$ ; however this cannot happen for  $E = A$ . Hence  $P$  (which can be written as a non-trivial convex combination of  $P_1, P_2$ ) is not extremal in  $\mathcal{M}$ .  $\square$

**Lemma 3.** *For any stationary measure  $P$ , the regular conditional probability of  $P$  given  $\mathcal{I}$ , denoted by  $Q(\omega, \cdot)$ , is stationary and ergodic.*

*Proof.* We know that almost surely

$$Q(\omega, A) = E(1_A | \mathcal{I}).$$

We need to show that  $Q(\omega, A) = Q(\omega, TA)$ . Suffices to show that for all  $I \in \mathcal{I}$

$$\int_I 1_A dP = \int_I 1_{TA} dP$$

or in other words  $P(A \cap I) = P(TA \cap I)$  which is immediate due to invariance of  $P$ .

To show ergodicity, we need to show that  $Q(\omega, I) = 0$  or  $1$ , for  $I \in \mathcal{I}$  for almost all  $\omega$ . This is again immediate. (note that the issue of throwing away too many null sets was covered during definition of regular conditional probabilities).  $\square$

**Theorem 3.** *Any invariant measure  $P \in \mathcal{M}$  can be written as a convex combination of ergodic measures, i.e.*

$$P = \int_{\mathcal{M}_e} Q \mu_P(dQ).$$

*Proof.* By regular conditional probabilities

$$P = \int Q(w, \cdot) dP.$$

By previous lemma  $Q(\omega, \cdot) \in \mathcal{M}_e$  and hence we have a induced measure  $\mu$  on measures in  $\mathcal{M}_e$ . By changing the integration with respect to that measure, we are done.  $\square$

## 8. BACKWARDS MARTINGALES, EXCHANGEABLE PROCESSES, AND DE FINETTI'S THEOREM

**Definition 5.** A *backwards martingale* is a sequence of random variables  $\{X_n\}, n \leq 0$ , adapted to a filtration,  $\{\mathcal{F}_n\}_{n \leq 0}$  defined by

$$X_n := E(X_0 | \mathcal{F}_n), n \leq 0,$$

where  $X_0$  is an integrable random variable.

**Lemma 4.** Let  $(\Omega, \mathcal{F}, P)$  be a probability space and  $X$  be an integrable random variable. Consider the collection

$$\{Y : Y = E(X | \mathcal{G}), \text{ for some } \mathcal{G} \subset \mathcal{F}\}.$$

Then the collection of random variables is uniformly integrable.

*Remark:* Formally the collection contains versions of conditional expectation.

*Proof.* Let  $\epsilon > 0$  be given. Let  $c_\delta = \sup_{A \in \mathcal{F}: P(A) \leq \delta} E(|X| 1_A)$ . We know from dominated convergence theorem that  $c_\delta \downarrow 0$  and  $\delta \downarrow 0$ . Choose  $\delta_0$  such that  $c_{\delta_0} \leq \epsilon$ . Choose  $M$  such that  $\frac{1}{M} E(|X|) \leq \delta_0$ .

Jensen's inequality says that  $|Y| \leq E(|X| | \mathcal{G})$  a.s. In particular  $E(|Y|) \leq E(|X|)$ . Hence  $P(|Y| > M) \leq \frac{1}{M} E(|Y|) \leq \delta_0$ . From the definition of conditional expectation

$$E(|Y|; |Y| > M) \leq E(|X|; |Y| > M) \leq \epsilon.$$

$\square$

**Theorem 4.** The limit  $X_{-\infty} = \lim_{n \rightarrow -\infty} X_n$  exists a.s. and in  $L^1$ .

*Proof.* Doob's upcrossing inequality for the number of upcrossings,  $U_n$ , between  $[a, b]$  made by  $X_{-n}, \dots, X_0$  yields  $(b-a) E(U_n) \leq E(X_0 - a)_+ < \infty$ . Hence the limit exists almost surely. Since the collection is uniformly integrable (above lemma), we have convergence in  $L^1$  (another lemma established before).  $\square$

Given a collection of random variables  $X_1, X_2, \dots$ , let  $\mathcal{E}_n$  be the events that are invariant under permutations of  $\{1, 2, \dots, n\}$  but leave  $n+1, \dots$  fixed. Let  $\mathcal{E} = \bigcap_n \mathcal{E}_n$  be the exchangeable  $\sigma$ -algebra.

**Definition 6.** A sequence  $X_1, X_2, \dots$  is said to be *exchangeable* if for every  $n$  and for every permutation  $\pi$  of  $\{1, \dots, n\}$  the distribution of  $(X_1, \dots, x_n)$  and  $X_{\pi(1)}, \dots, X_{\pi(n)}$  are the same.

**Theorem 5** (de Finetti). If  $X_1, X_2, \dots, X_n$  are exchangeable the conditioned on  $\mathcal{E}$ ,  $X_1, \dots$  are independent and identically distributed.

*Proof.* Let  $f$  be a bounded measurable function. Define

$$A_n(f) := \frac{1}{\binom{n}{k}} \sum_{i \subset [n]} f(X_{i_1}, \dots, X_{i_k}).$$

Since  $A_n$  is exchangeable, we have

$$A_n(f) = E(A_n(f) | \mathcal{E}_n) = \frac{1}{\binom{n}{k}} \sum_{i \subset [n]} E(f(X_{i_1}, \dots, X_{i_k}) | \mathcal{E}_n) = E(f(X_1, \dots, X_k) | \mathcal{E}_n).$$

From backwards martingale theorem  $A_n(f) \rightarrow A_\infty(f) = E(f(X_1, \dots, X_k) | \mathcal{E})$ .

Consider  $\phi(x_1, \dots, x_k) = f(x_1, \dots, x_{k-1})g(x_k)$  and define  $\phi_j(x_1, \dots, x_{k-1}) = f(x_1, \dots, x_{k-1})g(x_j)$  for  $1 \leq j \leq k-1$ . Then observe that

$$\begin{aligned} (n)_{k-1} A_n(f) n A_n(g) &= (n)_k A_n(\phi) + (n)_{k-1} \sum_{j=1}^{k-1} A_n(\phi_j) && \Longleftrightarrow \\ A_n(f) A_n(g) &= \frac{n-k+1}{n} A_n(\phi) + \frac{1}{n} \sum_{j=1}^{k-1} A_n(\phi_j). \end{aligned}$$

Taking  $n \rightarrow \infty$  we obtain  $A_\infty(f) A_\infty(g) = A_\infty(\phi)$ , or

$$E(f(X_1, \dots, X_{k-1})g(X_k) | \mathcal{E}) = E(f(X_1, \dots, X_{k-1}) | \mathcal{E}) E(g(X_k) | \mathcal{E}).$$

□