# PROBABILITY REVIEW(?)

# Predicting amount of rainfall



https://esan108.com/%E0%B8%9E%E0%B8%A3%E0%B8%B0%E0%B9%82%E0%B8%84%E0%B8%81%E0%B8%B4%E0%B8% 99%E0%B8%AD%E0%B8%B0%E0%B9%84%E0%B8%A3-%E0%B8%AB%E0%B8%A1%E0%B8%B2%E0%B8%A2%E0%B8%96

# (Linear) Regression

$$h_{\theta}(x) = \theta_0 + \theta_1 x_1 + \theta_2 x_2 + \theta_3 x_3 + \theta_4 x_4 + \theta_5 x_5$$

• θs are the parameter (or weights)

Assume x<sub>0</sub> is always 1

We can rewrite

$$h_{\theta}(x) = \sum_{i=0}^{n} \theta_i x_i = \theta^T \mathbf{x}$$

- Notation: vectors are bolded
- Notation: vectors are column vectors

### LMS regression with gradient descent

$$\frac{\partial J}{\partial \theta_i} = -\sum_{i=1}^m (y_i - \theta^T \mathbf{x}_i) x_i^{(j)}$$

$$\theta_j \leftarrow \theta_j + r\sum_{i=1}^m (y_i - \theta^T \mathbf{x}_i) x_i^{(j)}$$

Interpretation?

# Logistic Regression

Pass  $\theta^T \mathbf{x}$  through the logistic function

$$h_{\theta}(x) = \frac{1}{1 + e^{-\theta^T x}}$$

# Logistic Regression update rule

$$\theta_j \leftarrow \theta_j - r\sum_{i=1}^m (y_i - h_\theta(x_i))x_i^{(j)}$$

Update rule for linear regression

$$\theta_j \leftarrow \theta_j - r \sum_{i=1}^m (y_i - \theta^T \mathbf{x}_i) x_i^{(j)}$$

# What is Probability?

- Frequentist
   Probability = rate of occurrence in a large number of trials
- Bayesian
   Probability = uncertainty in your knowledge of the world



# Bayesian vs Frequentist



# Bayesian vs Frequentist

Toss a coin

- Frequentist
  - P(head) =  $\theta$ ,  $\theta$  = #heads/#tosses
- Bayesian
  - P(head) =  $\theta$ ,  $\theta$ ~U(0.6,1.0)
  - Parameters of distributions can now have probabilities
  - Bayesian interpretation can gives prior knowledge to the phenomena – subjective view of the world
  - Prior knowledge can be updated according to the observed frequency

# Bayesian statistics

- Coin with P(head) = p
- Observed frequency of heads  $\hat{p} = \#heads/\#n$
- In Bayesian view, we can talk about P(p | β̂) by using Bayes's rule

$$P(p|\hat{p}) = rac{P(\hat{p}|p)P(\hat{p})}{P(\hat{p})}$$

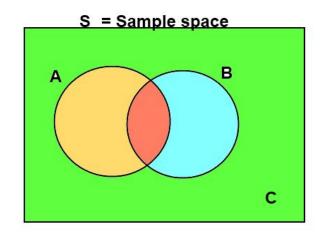
### Important concepts

- Conditional probability
  - Independence
- Bayes' Rule
- Expected Value and Variance
- CDFs
- Sum of RVs
- Gaussian Random Variable
  - Multivariate Gaussian

# Conditional probability

P(A|B) probability of A given B has occurred

$$P(A \mid B) = \frac{P(A \cap B)}{P(B)}$$



- A student posts a facebook status after finishing Pattern Recognition homework
- P(he is happy)
- P(he is happy | the post starts with "#\$@#\$!@#\$")

### Independence

• Two events are independent (statistically independent or stochastically independent) if the occurrence of one does not affect the probability of occurrence of the other.

$$P(A \cap B) = P(A)P(B) \Leftrightarrow P(B) = P(B \mid A)$$

 P(he is happy | His friend posted a cat picture on instragram )

# Bayes' Rule (Bayes's theorem or Bayes' law)

$$P(A \mid B) = \frac{P(B \mid A) P(A)}{P(B)}.$$

Usefulness: We can find P(A|B) from P(B|A) and vice versa

### Expected value

Expected value

$$E[x] = \int_{-\infty}^{\infty} xp(x)dx$$

$$E[g(x)] = \int_{-\infty}^{\infty} g(x)p(x)dx$$

Variance (σ²) (Standard Deviation = σ)

$$Var[x] = E[(x - E[x])^{2}] = \sigma^{2} = \int_{-\infty}^{\infty} (x - E[x])^{2} p(x) dx$$

$$E[(x-E[x])^2] = E[x^2]-(E[x])^2$$

### Expected Value notes

- It's a weighted sum.
- Something can have a high expected value but low probability of occurring
- Lottery: P(win) = 10^(-20), winner gets 10^30
  P(loss) = 1-P(win), loser gets -10
  E(Lottery earnings) = 10^(-20)10^30+ (1-P(win)) (-10)
  = 10^10 10
- Humans are not good at gauging probability at extreme points

# Expected value and Variance properties

- E[a] = a; a is a constant.
- $\bullet \ \mathsf{E}[\mathsf{a}X + \mathsf{b}] = \mathsf{a}\mathsf{E}[X] + \mathsf{b}$
- $\bullet \ \mathsf{E}[X+Y] = \mathsf{E}[X] + \mathsf{E}[Y]$
- Var[a] = 0
- $Var[aX+b] = a^2Var[X]$

Conditional Expected Value

$$E[x \mid A] = \int_{-\infty}^{\infty} xp(x \mid A)dx$$
$$E[g(x) \mid A] = \int_{-\infty}^{\infty} g(x)p(x \mid A)dx$$

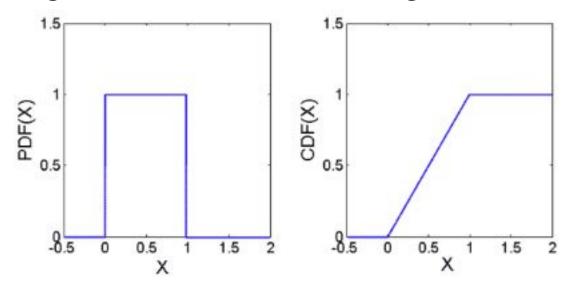
# Cumulative Distribution Functions CDFs

Probability that the RV is less than a certain amount

$$F_X(x_0) = P(X \le x_0) = \int_{-\inf}^{x_0} p(x)dx$$

CDF is the integral of PDF. Differentiating CDF wrt x gives

the PDF



#### Joint distributions

- If we want to monitor how two events are jointly occurring, we consider the joint distribution p<sub>x y</sub>(x,y)
- $p_{X,Y}(x,y) = p_X(x)p_Y(y)$  if x and y are independent

$$P(A) = \iint_A p_{XY}(x, y) dxdy$$

$$p_X(x) = \int_{-\infty}^{\infty} p_{XY}(x, y) dy$$

$$p_{Y}(y) = \int_{-\infty}^{\infty} p_{XY}(x, y) dx$$

#### Sum of Random variables

- $\cdot Z = Y + X$
- What is the pdf of Z? Where Y and X continuous RVs

$$p_{X+Y}(z) = (p_X * p_Y)(z) = (p_Y * p_X)(z)$$

# Central Limit Theorem (CLT)

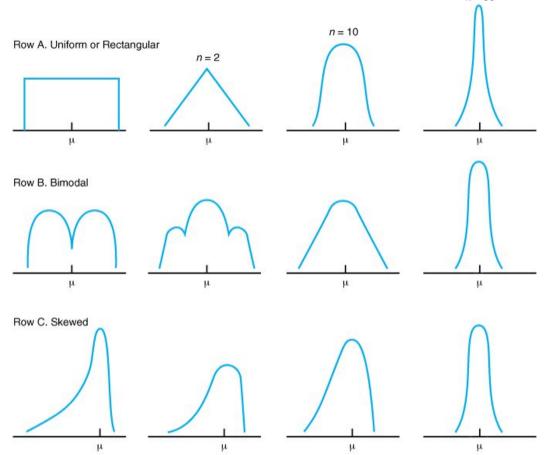
• Suppose X<sub>1</sub>,X<sub>2</sub>,... is a sequence of iid (independent and identically distributed) RVs. As n approaches infinity the sum of the sequence converge in distribution to a Normal distribution

$$\sqrt{n}\left(\left(rac{1}{n}\sum_{i=1}^{n}X_{i}
ight)-\mu
ight)\overset{d}{
ightarrow}N\left(0,\sigma^{2}
ight)$$

 Other variants of CLT exists, without the dependence or identically distributed assumption

# **CLT** implications

• A sum of RVs tends to become Normally distributed very quickly



# Gaussian distribution (normal distribution)

• X is normal (Gaussian):  $X \sim N(\mu, \sigma^2)$ 

$$p(x) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-(x-\mu)^2/2\sigma^2} E[x] = \mu$$

$$Var[x] = \sigma^2$$

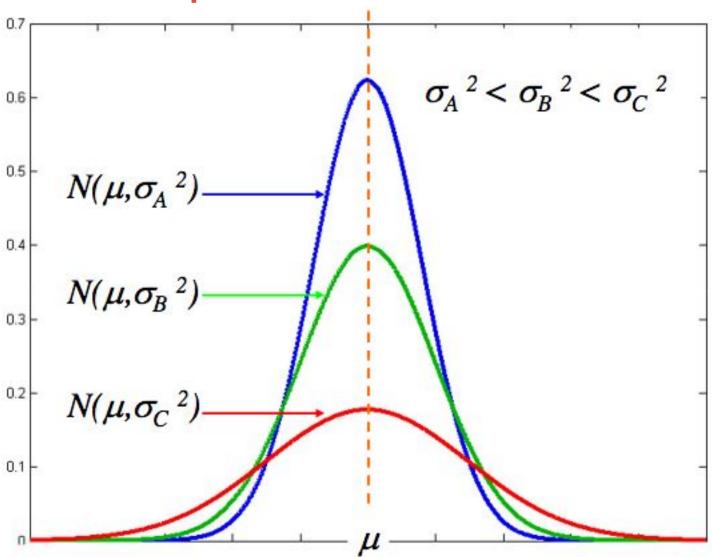
• X is Standard normal (Standard Gaussian):  $X \sim N(0,1)$  when  $\mu=0$ ,  $\sigma^2=1$ 

$$p(x) = \frac{1}{\sqrt{2\pi}} e^{-(x-\mu)^2/2}$$

$$E[x] = 0$$

$$Var[x] = 1$$

# Gaussian pdf



# Linear transformation of Gaussian RV

- Normality is preserved by linear transformation. Calculation involving the normal variable is usually done in terms of standard normal.
- Let Y=aX+b, if  $X\sim N(\mu,\sigma^2) \rightarrow Y\sim N(a\mu+b,a^2\sigma^2)$
- Let  $Z=(X-\mu)/\sigma$ , if  $X\sim N(\mu,\sigma^2) \rightarrow Z\sim N(0,1)$ : Standard Normal

Can you prove this?

#### Summation of 2 Gaussian RVs

- X mean m<sub>1</sub> variance σ<sub>1</sub><sup>2</sup>
- Y mean  $m_2$  variance  $\sigma_2^{\ 2}$
- X and Y are independent

• X+Y is normally distributed with mean  $m_1^{} + m_2^{}$  variance  $\sigma_1^{\; 2} + \sigma_2^{\; 2}$ 

#### Expectation of multivariate distributions

$$E[g(X_1, X_2)] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} g(x_1, x_2) p_{X_1, X_2}(x_1, x_2) dx_1 dx_2$$

$$E[g(X_1)h(X_2)] = E[g(X_1)]E[h(X_2)]$$

If  $X_1$  and  $X_2$  independent

#### Covariance of multivariate distributions

$$-cov(X_1,X_2) = E[(X_1-m_1)(X_2-m_2)]$$

$$-cov(X_1,X_2) = E[(X_1)(X_2)] - m_1m_2$$

- Covariance with itself is just the Variance
- Correlation

$$\rho = \frac{cov(X_1, X_2)}{\sqrt{V(X_1)V(X_2)}}$$

#### Covariance matrix

- Given a set of RVs, X<sub>1</sub> X<sub>2</sub> ... X<sub>n</sub>
- The covariance matrix is a matrix which has the covariance of the i and j RV in position (i,j)

$$\Sigma = \begin{bmatrix} \mathrm{E}[(X_1 - \mu_1)(X_1 - \mu_1)] & \mathrm{E}[(X_1 - \mu_1)(X_2 - \mu_2)] & \cdots & \mathrm{E}[(X_1 - \mu_1)(X_n - \mu_n)] \\ \mathrm{E}[(X_2 - \mu_2)(X_1 - \mu_1)] & \mathrm{E}[(X_2 - \mu_2)(X_2 - \mu_2)] & \cdots & \mathrm{E}[(X_2 - \mu_2)(X_n - \mu_n)] \\ \vdots & \vdots & \ddots & \vdots \\ \mathrm{E}[(X_n - \mu_n)(X_1 - \mu_1)] & \mathrm{E}[(X_n - \mu_n)(X_2 - \mu_2)] & \cdots & \mathrm{E}[(X_n - \mu_n)(X_n - \mu_n)] \end{bmatrix}.$$

# Understanding the Covariance matrix

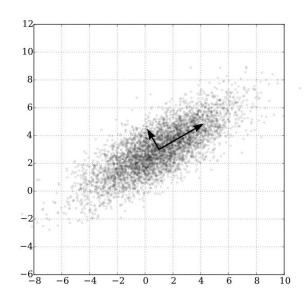
$$Cov(X, X)$$
  $Cov(Y, X)$   $Cov(Y, Y)$ 

Which statements are true?

$$C = B$$

B < 0

A < D



#### Covariance matrix observations

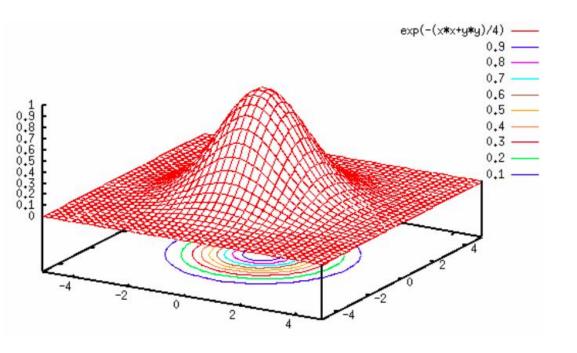
$$\Sigma = \Sigma^T$$

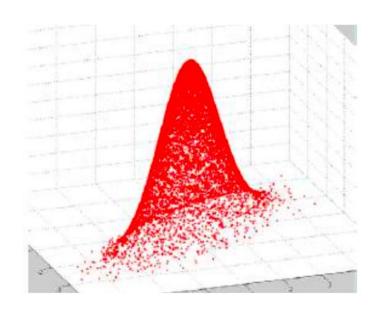
- If the covariance matrix is diagonal, all RVs are mutually independent.
- Covariance matrix is positive-semidefinite
  - Every positive definite matrix is invertible

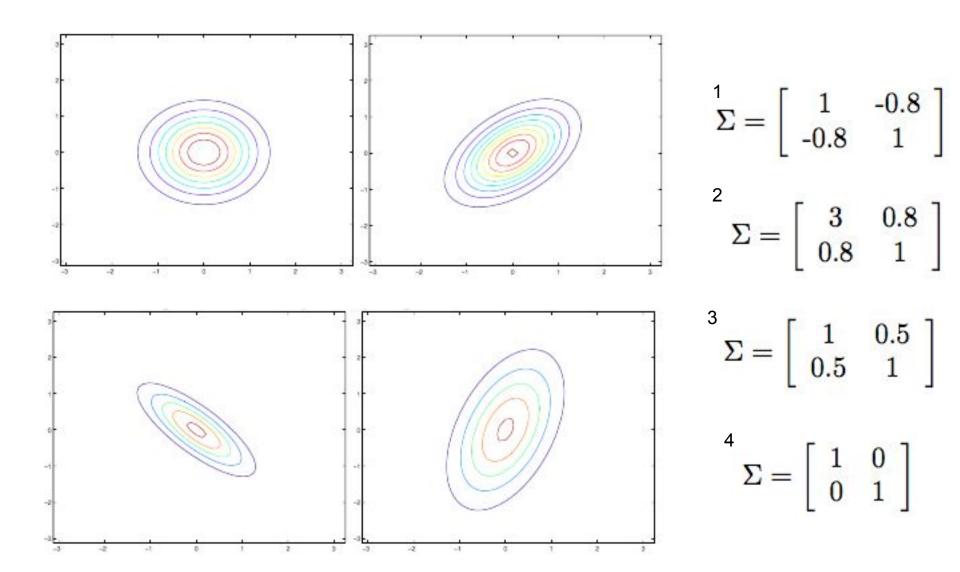
#### Multivariate Gaussian distribution

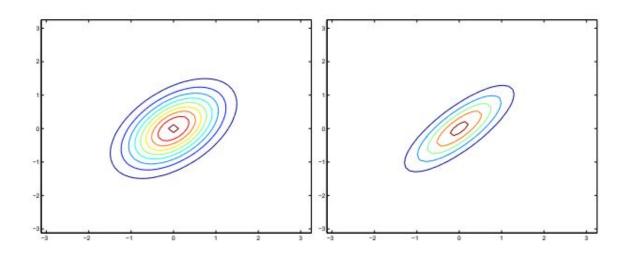
Put X<sub>1</sub>,X<sub>2</sub>,X<sub>3</sub>...X<sub>n</sub> into a vector x

$$p(\mathbf{x}; \mu, \Sigma) = \frac{1}{(2\pi)^{n/2} |\Sigma|^{1/2}} exp[-\frac{1}{2} (\mathbf{x} - \mu)^t \Sigma^{-1} (\mathbf{x} - \mu))]$$









$$\Sigma = \begin{bmatrix} 1 & 0.5 \\ 0.5 & 1 \end{bmatrix} \quad \Sigma = \begin{bmatrix} 1 & 0.8 \\ 0.8 & 1 \end{bmatrix}.$$

# Affine transformation of multivariate Gaussians

y = Ax+b , Assuming A has full rank (invertible)

$$\mathbf{x} \sim N(\mu, \Sigma)$$

$$\mathbf{y} \sim N(A\mu + b, A\Sigma A^T)$$

### Important concepts

- Conditional probability
  - Independence
- Bayes' Rule
- Expected Value and Variance
- CDFs
- Sum of RVs
  - CLT
- Gaussian Random Variable
  - Multivariate Gaussian

## Distribution parameter estimation

- P(head) =  $\theta$ ,  $\theta$  = #heads/#tosses
- HHTTH

- $L(\theta) = P(X; \theta) = P(HHTTH; \theta)$
- Maximum Likelihood Estimate (MLE)

## Linear Regression Revisit

$$h_{\theta}(x) = \theta_0 + \theta_1 x_1 + \theta_2 x_2 + \theta_3 x_3 + \theta_4 x_4 + \theta_5 x_5$$

- θs are the parameter (or weights)
- We can rewrite

$$h_{\theta}(x) = \sum_{i=0}^{n} \theta_i x_i = \theta^T \mathbf{x}$$

- Notation: vectors are bolded
- Notation: vectors are column vectors

# Probabilistic Interpretation of linear regression

- Real world data is our model plus some error term
  - Noise in the data
  - Something that we do not model (features we are missing)
- Let's assume the error is normally distributed with mean zero and variance σ<sup>2</sup>
  - Why Gaussian?
  - Why saying mean is zero is a valid assumption?

$$y_i = \theta^T \mathbf{x}_i + \epsilon_i$$

# Probabilistic view of Linear regression

- Find θ
- Maximize Likelihood of seeing x and y in training
- From our assumption we know that

$$y_i = \theta^T \mathbf{x}_i + \epsilon_i$$

$$p(y_i | \mathbf{x}_i; \theta) = \frac{1}{\sqrt{2\pi}\sigma} exp(-\frac{(y_i - \theta^T \mathbf{x}_i)^2}{2\sigma^2})$$

Error term is normally distributed with mean 0 and variance  $\sigma^2$ 

What is the assumption here? Is it accurate?

## Maximizing Likelihood

-Max 
$$L(\theta) = \prod_{i=1}^m p(y_i|\mathbf{x}_i;\theta)$$

• We use the log likelihood instead  $log(L(\theta)) = I(\theta)$ 

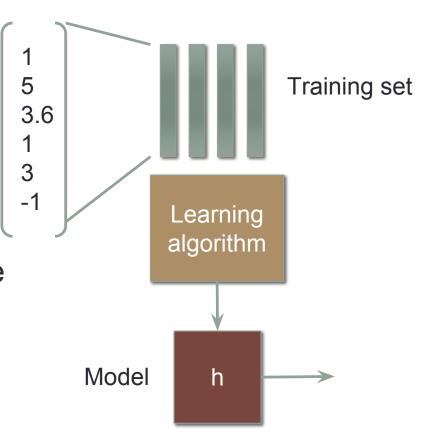
From our previous lecture 
$$Min \quad J(\theta) = \frac{1}{m} \Sigma_{i=1}^m (y_i - \theta^T \mathbf{x_i})^2$$

## Mean square error solution and MLE solution

- Turns out MLE and MSE gets to the same solution
  - This justifies our choice of MSE as the Loss for linear regression
  - This does not mean MSE is the best Loss for regression, but you can at least justify it with a probabilistic reasoning
- Note how our choice of variance  $\sigma^2$  falls out of the maximization, so this derivation is true regardless of which assumption for variance is.

#### Flood or no flood

- What would be the output?
- y = 0 if not flooded
- y = 1 if flooded
- Anything in between is a score for how likely it is to flood



Training phase

## Can we use regression?

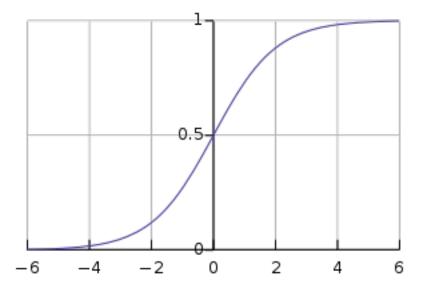
Yes

$$h_{\theta}(x) = \theta_0 + \theta_1 x_1 + \theta_2 x_2 + \theta_3 x_3 + \theta_4 x_4 + \theta_5 x_5$$

- But
- What does it mean when h is higher than 1?
- Can h be negative? What does it mean to have a negative flood value?

## Logistic function

- Let's force h to be between 0 and 1 somehow
- Introducing the logistic function (sigmoid function)



$$f(x) = rac{1}{1+e^{-x}} \ = rac{e^x}{1+e^x}$$

### Logistic Regression

Pass  $\theta^T \mathbf{x}$  through the logistic function

$$h_{\theta}(x) = \frac{1}{1 + e^{-\theta^T x}}$$

#### Loss function?

- MSE error no longer a good candidate
- Let's turn to use probabilistic argument for logistic regression

### Logistic Function derivative

The derivative has a nice property by design.

This is also why many algorithm we'll learn later in class also uses the logistic function

$$g'(z) = \frac{d}{dz} \frac{1}{1 + e^{-z}}$$

$$= \frac{1}{(1 + e^{-z})^2} (e^{-z})$$

$$= \frac{1}{(1 + e^{-z})} \cdot \left(1 - \frac{1}{(1 + e^{-z})}\right)$$

$$= g(z)(1 - g(z)).$$

# Probabilistic view of Logistic Regression

 Let's assume, we'll classify as 1 with probability in accordance to the output of

$$h_{\theta}(x) = \frac{1}{1 + e^{-\theta^T x}}$$

$$P(y = 1 \mid x; \theta) = h_{\theta}(x)$$
  
 
$$P(y = 0 \mid x; \theta) = 1 - h_{\theta}(x)$$

or

$$p(y \mid x; \theta) = (h_{\theta}(x))^{y} (1 - h_{\theta}(x))^{1-y}$$

## Maximizing log likelihood

$$p(y \mid x; \theta) = (h_{\theta}(x))^{y} (1 - h_{\theta}(x))^{1-y}$$

$$g'(z) = \frac{d}{dz} \frac{1}{1 + e^{-z}}$$

$$= \frac{1}{(1 + e^{-z})^2} (e^{-z})$$

$$= \frac{1}{(1 + e^{-z})} \cdot \left(1 - \frac{1}{(1 + e^{-z})}\right)$$

$$= g(z)(1 - g(z)).$$

$$\theta_j \leftarrow \theta_j + r\sum_{i=1}^m (y_i - h_\theta(x_i))x_i^{(j)}$$

## Logistic Regression update rule

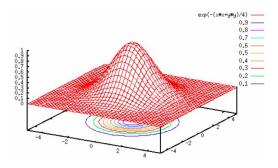
$$\theta_j \leftarrow \theta_j + r\sum_{i=1}^m (y_i - h_\theta(x_i))x_i^{(j)}$$

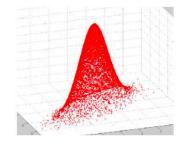
Update rule for linear regression

$$\theta_j \leftarrow \theta_j + r \sum_{i=1}^m (y_i - \theta^T \mathbf{x}_i) x_i^{(j)}$$

## Summary

- Conditional probability
  - Independence
- Bayes' Rule
- Expected Value and Variance
- CDFs
- Sum of RVs
  - CLT
- Gaussian Random Variable
  - Multivariate Gaussian
- Probabilistic view for Regression setups





#### Next time

- HW1 due Monday, Quiz 1 starts 13:10 ends 13:20. HW2 out
- MLE on more things!
- Maximum A Posteriori Estimation (MAP estimate)
- Classification based on probabilistic models

#### Homework

Can I just use scikit-learn?

You can use pandas and numpy for this homework

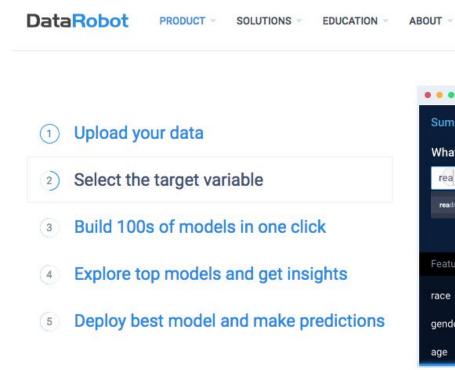


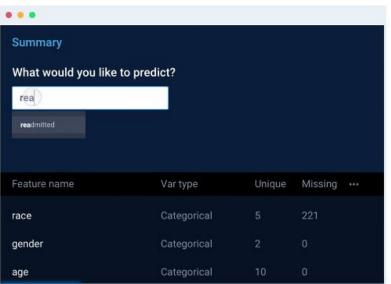
In particular, if you understand something clearly, you should be able to describe it in precise algorithmic terms to a computer: you should be able to implement it from scratch (as a simulation, as a framework, etc).

Creator of Keras

#### One button machines

- Machine learning as a tool for non-experts
- Can a non-expert just provide the data and let the machine decides how to proceed





**CONTACT US** 

## Reinforcement Learning for Model Selection

- Tuning a network takes time
- Let machine learning learns how to tune a network
- Matches or outperforms ML experts performance

