# Artificial Intelligence (CS303)

Lecture 9: Inference in Bayesian Networks

### Hints for this lecture

• Calculate/estimate posterior probability distribution.

### Outline of this lecture

Problem Statement

• Exact Inference

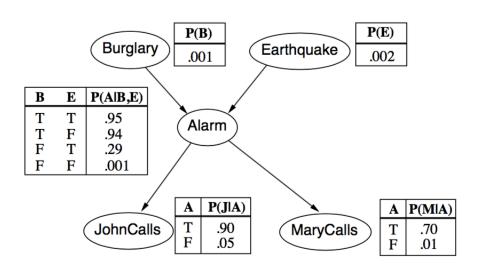
Approximate Inference

### I. Problem Statement

# What are we looking for?

• Given a Bayesian Network, and an (or some) observed events, which specifies the value for evidence variables, we want to know the probability distribution of one (or several) query variables X, P (X | events).

• E  $\mathbf{p}$  P(Burglary | JohnCalls = true, MaryCalls = true)



## More Examples

```
Simple queries: compute posterior marginal P(X_i|\mathbf{E}=\mathbf{e})
   e.g., P(NoGas|Gauge = empty, Lights = on, Starts = false)
Conjunctive queries: P(X_i, X_i | \mathbf{E} = \mathbf{e}) = P(X_i | \mathbf{E} = \mathbf{e})P(X_i | X_i, \mathbf{E} = \mathbf{e})
Optimal decisions: decision networks include utility information;
        probabilistic inference required for P(outcome|action, evidence)
Value of information: which evidence to seek next?
Sensitivity analysis: which probability values are most critical?
Explanation: why do I need a new starter motor?
```

### **II. Exact Inference**

### Enumeration

### Simple query on the burglary network:

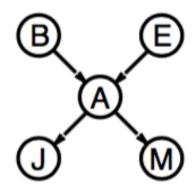
$$\mathbf{P}(B|j,m)$$

$$= \mathbf{P}(B,j,m)/P(j,m)$$

$$= \alpha \mathbf{P}(B,j,m)$$

$$= \alpha \sum_{e} \sum_{a} \mathbf{P}(B,e,a,j,m)$$

Need to consider all values of "hidden variables", e.g., alarm=true, alarm=false



### Rewrite full joint entries using product of CPT entries:

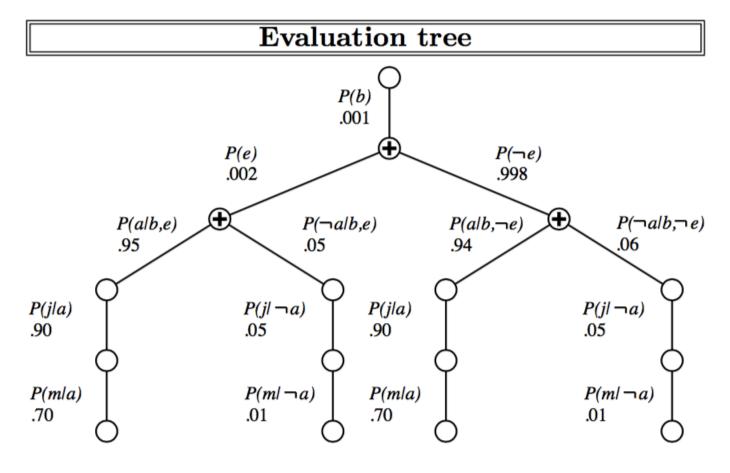
$$\mathbf{P}(B|j,m) = \alpha \sum_{e} \sum_{a} \mathbf{P}(B)P(e)\mathbf{P}(a|B,e)P(j|a)P(m|a)$$

$$= \alpha \mathbf{P}(B) \sum_{e} P(e) \sum_{a} \mathbf{P}(a|B,e)P(j|a)P(m|a)$$

Recursive depth-first enumeration: O(n) space,  $O(d^n)$  time

d=2 for Boolean variables

### Enumeration



Enumeration is inefficient: repeated computation e.g., computes P(j|a)P(m|a) for each value of e

## Enumeration by Variable Elimination

Variable elimination: carry out summations right-to-left, storing intermediate results (factors) to avoid recomputation

```
\begin{aligned} \mathbf{P}(B|j,m) &= \alpha \underbrace{\mathbf{P}(B) \sum_{e} \underbrace{P(e) \sum_{a} \mathbf{P}(a|B,e)}_{A} \underbrace{P(j|a)}_{J} \underbrace{P(m|a)}_{M}}_{J} \\ &= \alpha \mathbf{P}(B) \sum_{e} P(e) \sum_{a} \mathbf{P}(a|B,e) P(j|a) f_{M}(a) \\ &= \alpha \mathbf{P}(B) \sum_{e} P(e) \sum_{a} \mathbf{P}(a|B,e) f_{J}(a) f_{M}(a) \\ &= \alpha \mathbf{P}(B) \sum_{e} P(e) \sum_{a} f_{A}(a,b,e) f_{J}(a) f_{M}(a) \\ &= \alpha \mathbf{P}(B) \sum_{e} P(e) f_{\bar{A}JM}(b,e) \text{ (sum out } A\text{)} \\ &= \alpha \mathbf{P}(B) f_{\bar{E}\bar{A}JM}(b) \text{ (sum out } E\text{)} \\ &= \alpha f_{B}(b) \times f_{\bar{E}\bar{A}JM}(b) \end{aligned}
```

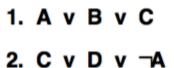
## Complexity of Exact Inference

#### Singly connected networks (or polytrees):

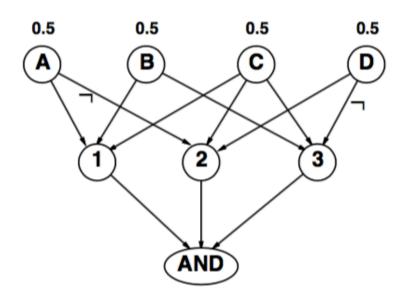
- any two nodes are connected by at most one (undirected) path
- time and space cost of variable elimination are  $O(d^k n)$

#### Multiply connected networks:

- can reduce 3SAT to exact inference ⇒ NP-hard
- equivalent to counting 3SAT models ⇒ #P-complete



3. B v C v ¬D



### III. Approximate Inference

### Basic Idea

• Sampling/Monte Carlo/Stochastic Simulation...

#### Basic idea:

- 1) Draw N samples from a sampling distribution S
- 2) Compute an approximate posterior probability  $\hat{P}$
- 3) Show this converges to the true probability P

#### Outline:

- Sampling from an empty network
- Rejection sampling: reject samples disagreeing with evidence
- Likelihood weighting: use evidence to weight samples
- Markov chain Monte Carlo (MCMC): sample from a stochastic process whose stationary distribution is the true posterior



## Sampling from an empty network

```
function PRIOR-SAMPLE(bn) returns an event sampled from bn inputs: bn, a belief network specifying joint distribution \mathbf{P}(X_1,\ldots,X_n) \mathbf{x}\leftarrow an event with n elements for i=1 to n do x_i\leftarrow a random sample from \mathbf{P}(X_i\mid parents(X_i)) given the values of Parents(X_i) in \mathbf{x} return \mathbf{x}
```

Assume the joint distribution could be easily sampled

# Sampling from an empty network

Probability that PRIORSAMPLE generates a particular event

$$S_{PS}(x_1 \dots x_n) = \prod_{i=1}^n P(x_i|parents(X_i)) = P(x_1 \dots x_n)$$

i.e., the true prior probability

E.g., 
$$S_{PS}(t, f, t, t) = 0.5 \times 0.9 \times 0.8 \times 0.9 = 0.324 = P(t, f, t, t)$$

Let  $N_{PS}(x_1 \ldots x_n)$  be the number of samples generated for event  $x_1, \ldots, x_n$ 

Then we have

$$\lim_{N \to \infty} \hat{P}(x_1, \dots, x_n) = \lim_{N \to \infty} N_{PS}(x_1, \dots, x_n) / N$$

$$= S_{PS}(x_1, \dots, x_n)$$

$$= P(x_1 \dots x_n)$$

That is, estimates derived from PRIORSAMPLE are consistent

Shorthand: 
$$\hat{P}(x_1, \ldots, x_n) \approx P(x_1 \ldots x_n)$$

# Rejection Sampling

 $\hat{\mathbf{P}}(X|\mathbf{e})$  estimated from samples agreeing with  $\mathbf{e}$ 

```
function Rejection-Sampling(X, e, bn, N) returns an estimate of P(X|e) local variables: N, a vector of counts over X, initially zero for j=1 to N do  x \leftarrow \text{Prior-Sample}(bn)  if x is consistent with e then  N[x] \leftarrow N[x] + 1 \text{ where } x \text{ is the value of } X \text{ in } x  return NORMALIZE(N[X])
```

In case distribution of one or more variable is difficult to sample

```
E.g., estimate \mathbf{P}(Rain|Sprinkler=true) using 100 samples 27 samples have Sprinkler=true Of these, 8 have Rain=true and 19 have Rain=false.
```

```
\hat{\mathbf{P}}(Rain|Sprinkler = true) = \text{Normalize}(\langle 8, 19 \rangle) = \langle 0.296, 0.704 \rangle
```

Similar to a basic real-world empirical estimation procedure

```
\hat{\mathbf{P}}(X|\mathbf{e}) = \alpha \mathbf{N}_{PS}(X,\mathbf{e}) (algorithm defn.)

= \mathbf{N}_{PS}(X,\mathbf{e})/N_{PS}(\mathbf{e}) (normalized by N_{PS}(\mathbf{e}))

\approx \mathbf{P}(X,\mathbf{e})/P(\mathbf{e}) (property of PRIORSAMPLE)

= \mathbf{P}(X|\mathbf{e}) (defn. of conditional probability)
```

Hence rejection sampling returns consistent posterior estimates

Problem: hopelessly expensive if  $P(\mathbf{e})$  is small

 $P(\mathbf{e})$  drops off exponentially with number of evidence variables!

# Likelihood Weighting

Idea: fix evidence variables, sample only nonevidence variables, and weight each sample by the likelihood it accords the evidence

Fix the evidence variables to reduce the sampling space

```
function LIKELIHOOD-WEIGHTING(X, e, bn, N) returns an estimate of P(X|e)
   local variables: W, a vector of weighted counts over X, initially zero
   for j = 1 to N do
        \mathbf{x}, w \leftarrow \text{Weighted-Sample}(bn)
        \mathbf{W}[x] \leftarrow \mathbf{W}[x] + w where x is the value of X in x
   return Normalize(W[X])
function WEIGHTED-SAMPLE(bn, e) returns an event and a weight
   \mathbf{x} \leftarrow an event with n elements; w \leftarrow 1
   for i = 1 to n do
        if X_i has a value x_i in e
             then w \leftarrow w \times P(X_i = x_i \mid parents(X_i))
             else x_i \leftarrow a random sample from P(X_i \mid parents(X_i))
   return x, w
```

# Markov Chain Monte Carlo (MCMC)

"State" of network = current assignment to all variables.

Generate next state by sampling one variable given Markov blanket Sample each variable in turn, keeping evidence fixed

```
function MCMC-Ask(X, e, bn, N) returns an estimate of P(X|e) local variables: N[X], a vector of counts over X, initially zero Z, the nonevidence variables in bn x, the current state of the network, initially copied from e initialize x with random values for the variables in Y for j=1 to N do for each Z_i in Z do sample the value of Z_i in x from P(Z_i|mb(Z_i)) given the values of MB(Z_i) in x N[x] \leftarrow N[x] + 1 where x is the value of X in x Markov Blanket return NORMALIZE(N[X])
```

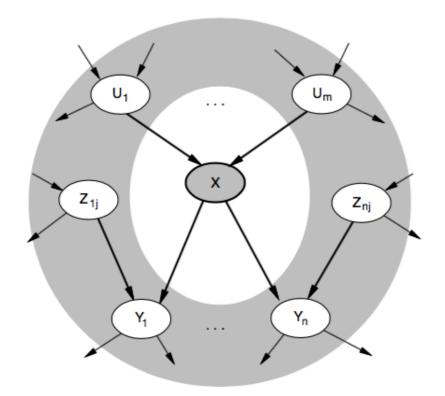
Further reduce the sampling space by only considering variables in Markov blanket

Can also choose a variable to sample at random each time

## Markov Chain Monte Carlo (MCMC)

#### Markov blanket

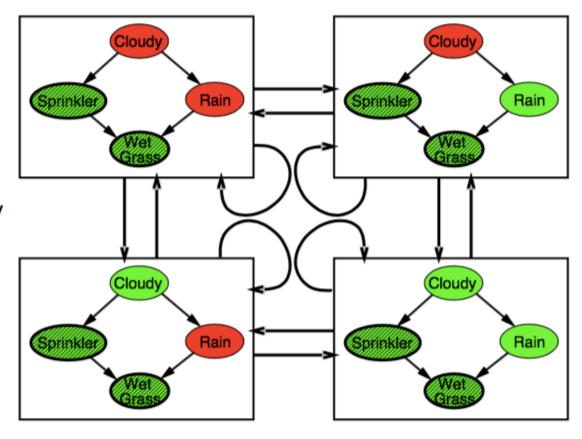
Each node is conditionally independent of all others given its Markov blanket: parents + children + children's parents



# Markov Chain Monte Carlo (MCMC)

With Sprinkler = true, WetGrass = true, there are four states:

Theorem: chain approaches stationary distribution: long-run fraction of time spent in each state is exactly proportional to its posterior probability



Wander about for a while, average what you see

# Summary

### Exact inference by variable elimination:

- polytime on polytrees, NP-hard on general graphs
- space = time, very sensitive to topology

### Approximate inference by LW, MCMC:

- LW does poorly when there is lots of (downstream) evidence
- LW, MCMC generally insensitive to topology
- Convergence can be very slow with probabilities close to 1 or 0
- Can handle arbitrary combinations of discrete and continuous variables

### To be continued