

# Probability and Statistics: Lecture-23

Monsoon-2020

by Pawan Kumar (IIIT, Hyderabad)

on October 5, 2020

## » Checklist for online class

1. Turn off your microphone, when you are listening
2. Turn on microphone only when you have question
3. Attend tutorials to practice problems or to discuss solutions or doubts
4. Chat is not always reliable, I may not look at chat

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## » Special Distribution: Exponential Distributions...

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$$f_X(x) = \begin{cases} \lambda e^{-\lambda x} & x > 0 \\ 0 & \text{otherwise} \end{cases}$$

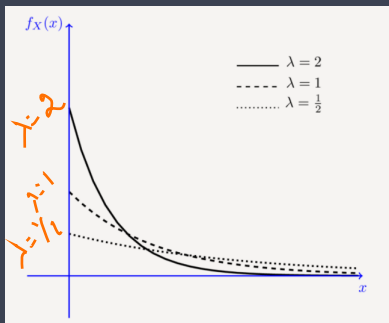
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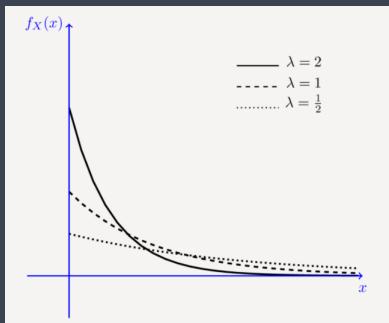
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The **CDF** is given by

$$F_X(x) = \int_0^x \lambda e^{-\lambda t} dt = 1 - e^{-\lambda x}$$



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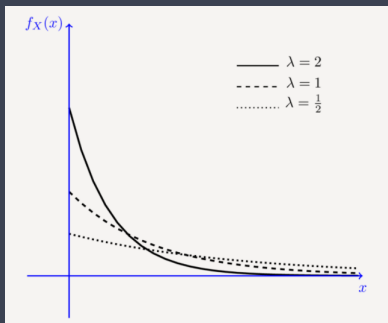
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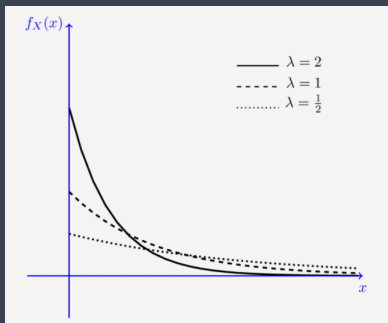


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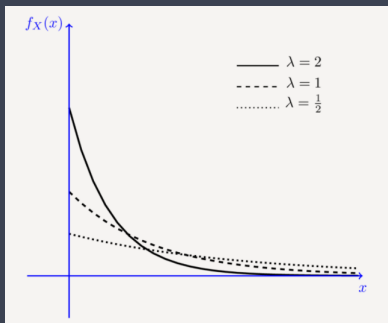
$$\begin{aligned} E[X] &= \int_0^{\infty} x \lambda e^{-\lambda x} dx = \frac{1}{\lambda} \int_0^{\infty} y e^{-y} dy \\ &= \frac{1}{\lambda} [-e^{-y} - y e^{-y}]_0^{\infty} = \frac{1}{\lambda} \end{aligned}$$

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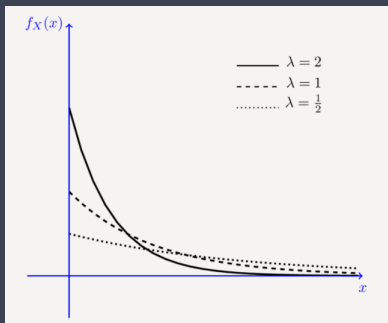


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Var(X) is given by:

$E[X^2] - (E[X])^2$   $\lambda x = y$

$$E[X^2] = \int_0^{\infty} x^2 \lambda e^{-\lambda x} dx = \frac{1}{\lambda^2} \int_0^{\infty} y^2 e^{-y} dy$$

int. by parts

$$= \frac{1}{\lambda^2} \left[ -2e^{-y} - 2ye^{-y} - y^2 e^{-y} \right]_0^{\infty} = \frac{2}{\lambda^2}$$

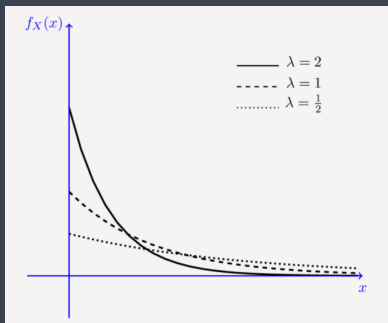
verify

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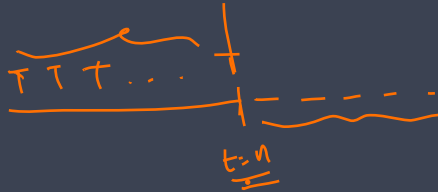


$\text{Var}(X)$  is given by:

$$\begin{aligned} E[X^2] &= \int_0^{\infty} x^2 \lambda e^{-\lambda x} dx = \frac{1}{\lambda^2} \int_0^{\infty} y^2 e^{-y} dy \\ &= \frac{1}{\lambda^2} [-2e^{-y} - 2ye^{-y} - y^2 e^{-y}]_0^{\infty} = \frac{2}{\lambda^2} \end{aligned}$$

$$\text{Var}(X) = \underbrace{E[X^2]} - (\underbrace{E[X]})^2 = \frac{2}{\lambda^2} - \frac{1}{\lambda^2} = \frac{1}{\lambda^2}$$

## » Exponential Distribution is Memoryless...



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$$P(X > x + a \mid X > a) = P(X > x), \quad \text{for } a, x \geq 0.$$

$$\begin{aligned} \underline{P(X > x + a \mid X > a)} &= \frac{P(X > x + a, X > a)}{P(X > a)} \\ &= \frac{P(\underline{X > x + a})}{P(\underline{X > a})} = \frac{1 - F_X(x + a)}{1 - F_X(a)} \\ &= \frac{e^{-\lambda(x+a)}}{e^{-\lambda a}} = e^{-\lambda x} \\ &= \underline{P(X > x)} \end{aligned}$$

$$\begin{aligned} F_X(x) &= \int_{-\infty}^x \lambda e^{-\lambda t} dt \\ P(X \leq x) \\ \Rightarrow P(X > x + a) \\ &= 1 - P(X \leq x + a) \\ &= 1 - F_X(x + a) \\ &= 1 - e^{-\lambda(x+a)} \\ F_X(x) &= 1 - e^{-\lambda x} \\ e^{-\lambda x} &= 1 - F_X(x) \\ &= P(X > x) \end{aligned}$$



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$$f_Z(z) = \frac{1}{\sqrt{2\pi}} e^{-\frac{z^2}{2}}, \quad \text{for all } z \in \mathbb{R}$$

Handwritten notes:

$$\int_{-\infty}^{\infty} f_Z(z) dz = 1 \quad \text{check}$$
$$\int_{-\infty}^{\infty} e^{-z^2/2} dz = \sqrt{2\pi}$$

Indefinite integral

## » Normal (Gaussian) Distribution...

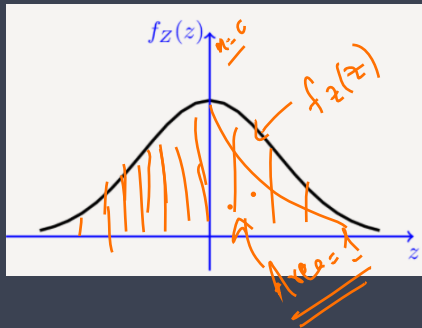
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← even or odd

$$\underline{\underline{f(-z) = f(z)}}$$

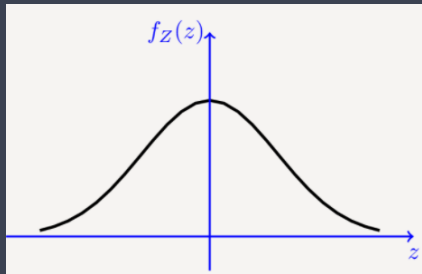


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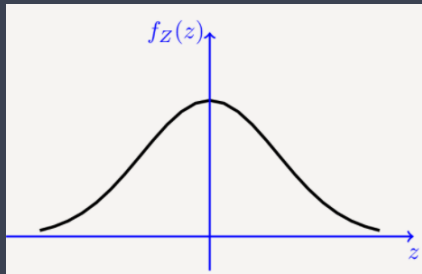
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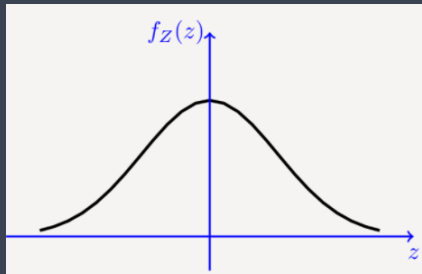


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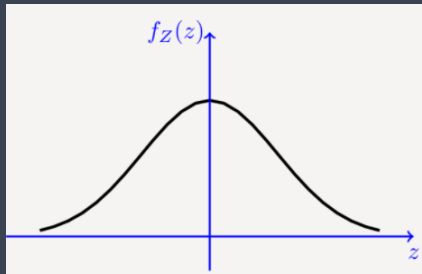
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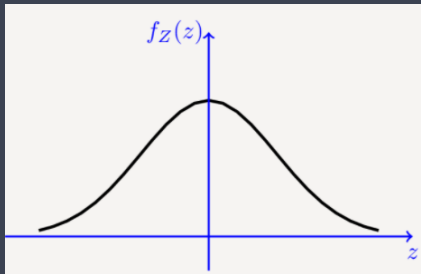
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- \* Here  $\frac{1}{\sqrt{2\pi}}$  is there to make area under curve 1

## » Mean and Variance of Standard Normal Distribution...

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Let  $Z$  be a **normal distribution**, i.e.,  $Z \sim N(0, 1)$ , then  $E[Z] = 0$  and  $\text{Var}(Z) = 1$ .

#### Recall

If  $g(u) : \mathbb{R} \rightarrow \mathbb{R}$ . If  $g(u)$  is an **odd function**, i.e.,  $g(-u) = -g(u)$ , and

$$\left| \int_0^{\infty} g(u) du \right| < \infty,$$

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» Answer to previous problem...

$$z \sim N(0,1)$$

$$E[z] = \int_{-\infty}^{\infty} x \cdot \underbrace{\frac{1}{\sqrt{2\pi}} e^{-x^2/2}}_{\text{odd fn.}} dx$$

$$= 0$$

$$E[z^2] = \int_{-\infty}^{\infty} x^2 \underbrace{\frac{1}{\sqrt{2\pi}} e^{-x^2/2}}_{\text{even fn.}} dx$$

Int. by parts

$$\int_{-\infty}^{\infty} x^2 e^{-x^2/2} dx = \sqrt{2\pi}$$

$$= \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} x^2 e^{-x^2/2} dx$$

$$= \frac{1}{\sqrt{2\pi}} 2 \int_0^{\infty} x^2 e^{-x^2/2} dx$$

$$I = \int_0^{\infty} \underbrace{x^2}_{1^{st}} \underbrace{e^{-x^2/2}}_{2^{nd}} dx$$

$$= \left[ x^2 \int_0^{\infty} e^{-x^2/2} dx \right]_0^{\infty} - \int_0^{\infty} 2x \int_0^{\infty} e^{-x^2/2} dx$$

» Answer to previous problem...

$$= 2 \int_0^{\infty} \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx =$$

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$$= \frac{2}{\sqrt{2\pi}} \int_0^{\infty} e^{-x^2/2} dx$$

$$= \frac{2}{\sqrt{2\pi}} \cdot \frac{1}{2} \int_{-\infty}^{\infty} e^{-x^2/2} dx = \frac{1}{\sqrt{2\pi}}$$



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$Z \sim N(0,1)$

$$\Phi(x) = P(Z \leq x) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^x e^{-\frac{u^2}{2}} du$$

*Handwritten notes: Arrows point from  $\Phi(x)$  to  $F(x)$  and from  $P(Z \leq x)$  to  $F(x)$ . The term  $e^{-\frac{u^2}{2}}$  is circled.*

- \* The integral **does not** have a **closed** form solution!
- \* However, values of  $F(Z)$  have been tabulated

## » CDF of Standard Normal Distribution...

$$F_Z(x) = \underline{P(Z \leq x)}$$

### Definition of CDF of Standard Normal Distribution

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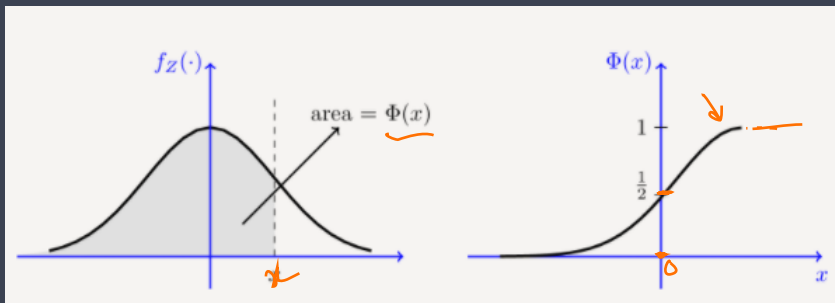
\* The integral **does not** have a **closed** form solution!

\* However, values of  $F(Z)$  have been **tabulated**

→ \* The **CDF** of any **normal** distribution can be written in terms of  $\Phi$  function

## » CDF of Standard Normal Distribution...

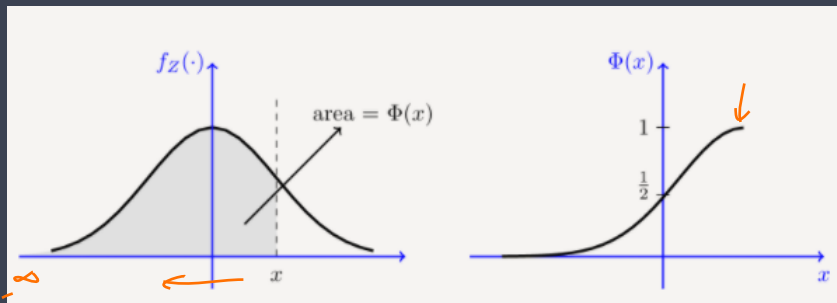
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The  $\Phi$  function satisfies the following properties:



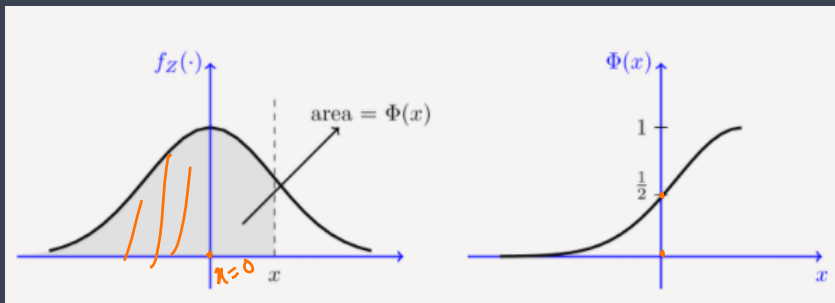
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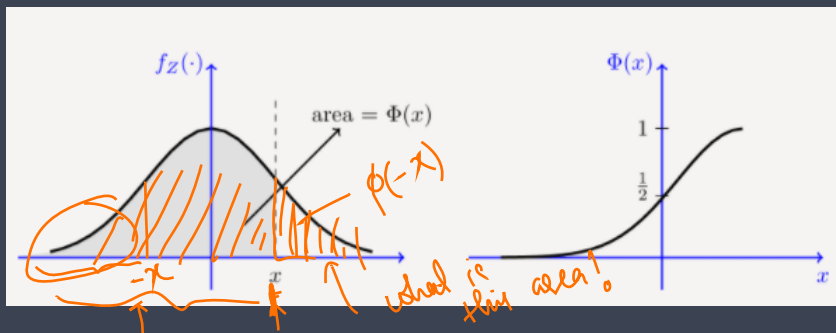
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- \*  $\lim_{x \rightarrow \infty} \Phi(x) = 1$ ,  $\lim_{x \rightarrow -\infty} \Phi(x) = 0$

- \*  $\Phi(0) = \frac{1}{2}$

- \*  $\Phi(-x) = 1 - \Phi(x)$  for all  $x \in \mathbb{R}$

## » Bound for $\Phi$ Function...

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### Bound for $\Phi$ Function

Let  $Z \sim N(0, 1)$ . We recall that

$$\Phi(x) = P(Z \leq x).$$

$$\begin{aligned} 1 - \Phi(x) &= 1 - P(Z \leq x) \\ &= P(Z > x). \end{aligned}$$

For all  $x \geq 0$ , the  $\Phi$ -function satisfies the following bound

$$\frac{1}{\sqrt{2\pi}} \frac{x}{x^2 + 1} e^{-x^2/2} \leq \underbrace{1 - \Phi(x)}_{P(Z > x)} \leq \frac{1}{\sqrt{2\pi}} \frac{1}{x} e^{-x^2/2}$$

» Answer to previous problem...

To show upper bound:

$$P(z > x) = \frac{1}{\sqrt{2\pi}} \int_x^\infty \underbrace{e^{-u^2/2}}_{g(u)} du$$

$$\leq \frac{1}{\sqrt{2\pi}} \int_x^\infty \underbrace{\frac{u}{x} e^{-u^2/2}}_{g(u)} du$$

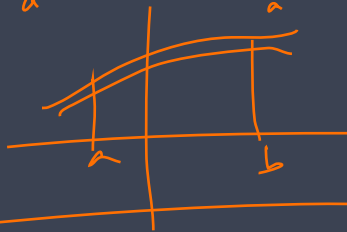
$$= \frac{1}{\sqrt{2\pi}} \cdot \frac{1}{x} \int_x^\infty u e^{-u^2/2} du$$

Note: obviously  
 $u > x \geq 0$

Recall

$$f(x) \leq g(x) \quad x \in [a, b]$$

$$\int_a^b f(x) dx \leq \int_a^b g(x) dx$$



$$= \frac{1}{\sqrt{2\pi}} \frac{1}{x} \left[ -e^{-u^2/2} \right]_x^\infty = \frac{1}{\sqrt{2\pi}} \frac{1}{x} e^{-x^2/2}$$