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8. Classification

Data Analysis for Networks - DataNets'19 Anastasios Giovanidis

Sorbonne-LIP6







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Classification Setting

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We have seen how to fit models to data when the response y_i to the input x_i is quantitative (e.g. "0.57", "24", "-24.3", etc.)

Question: How do we choose models and define their accuracy, when y_i 's are qualitative?

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Examples: ("Yes", "No"), ("Red", "Blue", "Green"), ("Malaria", "Yellow Fever", "Flu") or more generally:
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("Class 1", "Class 2", ..., "Class M")
```

Application A: IoT Classifier

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Example application: Internet-of-Things (IoT) for home networks. "Device identification assistant." from [B.2]

- ► Home devices can be controlled from distance. (Camera, Light, Sensor, Mobile, Switch, Alarm, Tablet, Speaker, TV.)
- For better quality-of-service these devices need to be identified by type from the network.
- Massive number of devices with heterogeneous functionality!

Use supervised learning to train an object classifier.

Input data:

- (a) the data-flow information per device, i.e. traffic characteristics.
- (b) a selected list of attributes (features).

A. IoT Features

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Once a device is connected, a MAC address is attributed.

Feature set to use for classification:

- Flow-based statistics:
 - Packet size (mean, max, min)
 - Mean inter-arrival packet time in a flow.
 - Flow-size measured in number of packets.
 - Protocol type: HTTP, HTTPS, SSDP, mDNS, TFTP, etc.
- Textual attributes (Bag-of-words): 0 or 1 per word per object?
 - ▶ Fabrication mark from MAC address.
 - Model and Type from HTTP.

A. IoT Implementation

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- WiFi access connected to an Ethernet switch.
- A measurement computer is connected at the switch to trace traffic.
- ▶ The computer collects data from the new IoT device during 1 min.
- The computer contains the trained classifier, which decides the most relavant class the IoT device belongs to. The decision is probabilistic.

Types of classifier: K-Nearest Neighbours, Naive Bayes, Random Forest, Tree-based classifier, etc.

Application B: Classifying Video QoE

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How to detect video streaming QoE issues from encrypted traffic? (see [B.3])

► Use predictive models to detect different levels of QoE degradation, due to: stalling, average video quality, quality variations.

Labels:

- ► Stalling: (None, Mild, Severe)
- ▶ Video Quality: (Low, Medium, High)
- Quality Switch: use frequency and amplitude of switches.

Features:

- (a) Chunk size percentiles, and average.
- (b) Packet retransmissions, (c) Bandwidth-Delay Product (BDP),
- (d) Bytes-In-Flight (BIF).

Training Accuracy

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Suppose we have training observations:

$$D_n = \{(x_1, y_1), (x_2, y_2), \dots, (x_n, y_n)\}, \text{ with } y_1, \dots, y_n \text{ qualitative.}$$

Consider a fitting model with an estimate $\hat{y}_i = \hat{f}(x_i)$. We use the training error rate:

$$\frac{1}{n}\sum_{i=1}^{n}\mathbf{1}\left(y_{i}\neq\hat{y}_{i}\right).$$

This is the fraction of incorrect classifications:

- \hat{y}_i is the predicted class label for the i-th observation using \hat{f} .
- ▶ $\mathbf{1}(y_i \neq \hat{y}_i) = 0$ for correct classification, else 1.
- ▶ Similar to MSE_{train} in regression!

Test Accuracy

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Most interested in the error rates of the classifier to test observations $(x_o, y_o) \notin D_n$, not used in training.

Again for an estimate $\hat{y}_o = \hat{f}(x_o)$ we use the test error rate:

Ave
$$(\mathbf{1}(y_o \neq \hat{y}_o))$$
.

A good classifier is the one for which the test error is smallest!

Bayes Classifier

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Optimal Classifier: (If all misclassifications are equally important) Assign each observation to the most likely class, given its predictor values:

$$\max_{1 \le j \le M} Pr(Y = j \mid X = x_o)$$

• We consider *conditional probabilities* given the observed x_o .

In a two-class problem

Decision boundary $Pr(Y = 1 \mid X = x_0) = Pr(Y = 2 \mid X = x_0)$

$$Pr(Y = 1 \mid X = x_o) + Pr(Y = 2 \mid X = x_o) = 1:$$

Class 1, if $Pr(Y = 1 \mid X = x_o) > 0.5$

Class 2, if $Pr(Y = 2 \mid X = x_o) > 0.5$

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Bayes example

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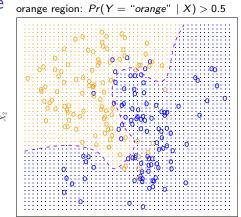


Figure: Bayes classifier: D_{100} data-set and 2 classes (blue, orange). ¹

 X_1

¹Source [B.1]

Bayes classifier cont'd

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- ▶ Orange shaded region: $Pr(Y = "orange" \mid X) > 0.5$.
- ▶ Blue shaded region: $Pr(Y = "blue" \mid X) > 0.5$.
- The dashed line: Bayes decision boundary.
- Circles that fall in regions with different colour: misclassifications

Bayes classifier produces lowest test error rate (irreducible)!

Test
$$Error(x_o) = 1 - \max_j Pr(Y = j \mid X = x_o)$$

Drawback...

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There is one problem however: For real data we do not know the conditional distribution P(Y|X),

(unless we have generated data ourselves, in which case we know the joint distribution P(X, Y)).

Bayes classifier serves as an unreachable gold standard!

If we do not know exactly P(Y|X) we can try to estimate it.

Naive Bayes

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The Naive Bayes classifier can use the data set \mathcal{D}_n .

- ▶ It assumes that the *K* features are independent.
- It uses a simple MAP or ML estimator

$$P(Y \mid \mathcal{D}_n) \propto P(\mathcal{D}_n \mid Y)P(Y)$$
 [MAP]
 $P(Y \mid \mathcal{D}_n) \propto P(\mathcal{D}_n \mid Y)$ [ML]

where Y is the class label.

We choose MAP or ML, depending on the prior information over the class distribution Y.

Naive Bayes with continuous features

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 \varnothing Suppose that X contains K continuous state features. Example: suppose we have M=2 classes ('True-False')

- Let us use a uniform prior over classes, i.e. P(j = 1) = P(j = 2) = 0.5.
- Suppose that the distribution per class j for each feature k is Gaussian $\mathcal{N}(\mu_{j,k}, \sigma_{i,k}^2)$.

ML estimates for mean and variance per feature k per class j:

$$\overline{X}_{j,k} = \frac{1}{n} \sum_{t \in \mathcal{D}_n} X_{j,k}(t),$$

$$\overline{S}_{j,k}^2 = \frac{1}{n} \sum_{t \in \mathcal{D}} (X_{j,k}(t) - \overline{X}_{j,k})^2.$$

Naive Bayes with continuous features (II) A. Giovanidis 2019

Given a Test sample (x_o, y_o) , the estimated class is the one which maximizes the ML (or MAP) estimator, i.e. the maximum between

$$\begin{split} &\prod_{k=1}^K \frac{1}{(2\pi\overline{S}_{\text{Cl},k}^2)^{1/2}} \exp\left(-\frac{(\textbf{x}_o - \overline{X}_{\text{Cl},k})^2}{2\overline{S}_{\text{Cl}}^2}\right) \quad \textit{for} \quad \textit{Class} 1 \\ &\prod_{k=1}^K \frac{1}{(2\pi\overline{S}_{\text{C2},k}^2)^{1/2}} \exp\left(-\frac{(\textbf{x}_o - \overline{X}_{\text{C2},k})^2}{2\overline{S}_{\text{C2}}^2}\right) \quad \textit{for} \quad \textit{Class} 2 \end{split}$$

Naive Bayes with discrete features

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 $\ensuremath{\mathscr{O}}$ Suppose that X contains K binary state features. Then $X_k(t)$ says that feature k appears or not in the t-th data sample from \mathcal{D}_n . Example: suppose we have M=2 classes ('True-False')

- We will estimate from data the prior distribution over classes, i.e. P(j = 1) and P(j = 2).
- Suppose that the distribution per class for each feature is Bernoulli: Bernoulli($p_{i,k}$).

MAP posteriors:

$$P(Y = j \mid \mathcal{D}_n) = P(\mathcal{D}_n \mid Y = j) \cdot P(Y = j)$$

$$= \prod_{t \in \mathcal{D}_n} \left(\prod_{k=1}^K p_{j,k}^{X_k(t)} (1 - p_{j,k})^{1 - X_k(t)} \right) \cdot P(Y = j)$$

Naive Bayes with discrete features (II)

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The parameters $p_{j,k}$ and the prior probability P(Y = j) are unknown and can be estimated by the data set \mathcal{D}_n , as ML Bernoulli estimates.

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	History	Science	F1:'king'	F2:'food'	F3:'equals'	F4:'proof'
Text 1	No	Yes	No	Yes	Yes	Yes
Text 2	No	Yes	No	No	Yes	No
Text 3	Yes	No	Yes	Yes	No	Yes
Text n	No	Yes	Yes	No	Yes	Yes

$$P(j = "History") = \frac{1}{n} \sum_{t=1}^{n} \mathbf{1}_{t,History}(Yes).$$

$$p_{\textit{History},F1} = \frac{\sum_{t=1}^{n} \mathbf{1}_{t,\textit{History}}(\texttt{Yes}) \mathbf{1}_{t,F1}(\texttt{Yes})}{\sum_{t=1}^{n} \mathbf{1}_{t,\textit{History}}(\texttt{Yes})},$$

$$X_{F1}(t) = 1, \text{ if Yes for F1 in Text } t.$$

Classifiers

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We will further consider in this lecture the following classifiers:

- ▶ ★ K-Nearest-Neighbours classifier (KNN)
- ► ★ Logistic Regression (LR)
- ► Linear Discriminant Analysis (LDA)
- Quadratic Discriminant Analysis (QDA)

KNN classifier

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How does the KNN classifier work?

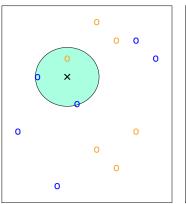
- ▶ Choose a positive integer K > 0.
- ▶ Given a test observation $x_o \notin D_n$, the KNN classifier identifies the K points in the training data-set closest to x_o , it is the set $\mathcal{N}_K(x_o)$.
- ▶ The conditional probability for class j at x_o is estimated as:

$$Pr(Y = j \mid X = x_o) = \frac{1}{K} \sum_{i \in \mathcal{N}_K(x_o)} \mathbf{1}(y_i = j).$$

- ▶ Calculate the estimates for all classes j = 1, ..., M and
- ► Finally, apply Bayes: classify x_o to the class with the largest estimated probability.

KNN example





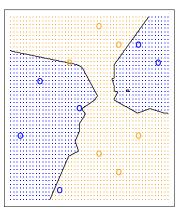


Figure: KNN classifier (K=3): D_{12} data-set and 2 classes. ²

²Source [B.1]

Optimal Choice of K

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Despite its simplicity KNN can give classifiers surprisingly close to Bayes. Choice of K is important:

- If K = 1, very flexible decision boundary → Low Training Error (= 0) but! High Test Error.
- ► As K increases (less flexibility) Training Error increases but the Test Error may not!
- ► Find optimal K* with minimum Test Error (U shape)
- ▶ If K = 100 decision boundary close to linear.

Variance vs Bias Tradeoff or Flexibility vs Interpretability

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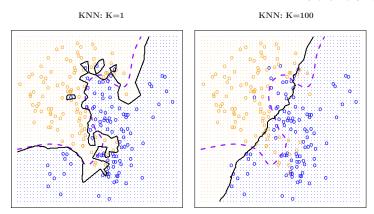


Figure: KNN with K = 1 (left) and K = 100 (right). ³

³Source [B.1]

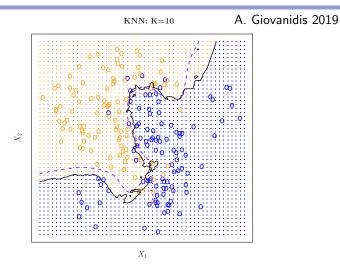


Figure: KNN with K = 10 close to Bayes optimal. ⁴

⁴Source [B.1]

Variance vs Bias Tradeoff

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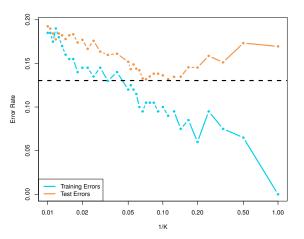


Figure: Training/Test Error Rate. ⁵

What if... Linear Regression?

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Suppose we have again two classes: 'Class 1', 'Class 2'.

- ▶ What if we used Linear Regression for the P(Y|X)?
- Let 'Class 1': Y = 0 and 'Class 2': Y = 1.
- ightharpoonup We assume that the linear model describes the 0/1 data,

$$y_i = \beta_0 + \beta_1 x_i + \epsilon$$

and we look for the regression line

$$\mathbb{E}\left[y_i|x_i\right] = \hat{\beta}_0 + \hat{\beta}_1 x_i$$

Since
$$y_i \in \{0,1\}$$
 then $\mathbb{E}\left[y_i|x_i\right] = Pr\left(y_i = 1|x_i\right) = \hat{\beta}_0 + \hat{\beta}_1 x_i$.

Wrong Shape! less than 0, more than 1 A. Giovanidis 2019

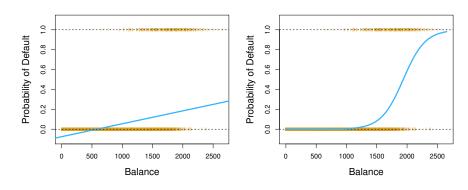


Figure: Pr(Y = 1|X). Linear vs Sigmoidal fit. ⁶

⁶Source [B.1]

Logistic Regression

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Suppose for the two-class problem Pr(Y = 1|X) follows the logistic function.

$$p(X) := Pr(Y = 1|X) = \frac{e^{\beta_o + \beta_1 X}}{1 + e^{\beta_o + \beta_1 X}}$$

- ▶ For $X \to -\infty$: $p(X) \to 0$
- ▶ For $X \to +\infty$: $p(X) \to 1$
- ► It is an S-shaped curve.

We need to fit β_o , β_1 in the non-linear logistic function.

Logistic fit

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We consider a Training data-set D_n with $Y_n = (0, 0, 1, \dots, 0, 1)$.

- We don't want to use MSE fit → complicated expressions.
- Better use: log-likelihood function.

What is the likelihood $g(D_n)$ of the data-sample?

$$g(D_n) = \prod_{i:y_i=1} p(x_i) \prod_{i':y_{i'}=0} (1 - p(x_{i'}))$$

because we assumed that for any X

$$Y = \begin{cases} 1, & p(X) \\ 0, & 1 - p(X) \end{cases}$$

and for all $x_i \in D_n$ we know what is the y_i answer.

Log-likelihood maximization

The log-likelihood function, is then equal to

$$\begin{array}{ll} \ell(\beta_{0},\beta_{1};D_{n}) & = & \log(g(D_{n})) \\ & = & \sum_{i:y_{i}=1} \log p(x_{i}) + \sum_{i':y_{i'}=0} \log (1-p(x_{i'})) \\ & = & \sum_{i=1}^{n} \left\{ y_{i} \log p(x_{i}) + (1-y_{i}) \log (1-p(x_{i})) \right\} \\ & = & \sum_{i=1}^{e^{\beta_{0}+\beta_{1}X}} \sum_{i=1}^{n} \left\{ y_{i} \left(\beta_{0}+\beta_{1}x_{i}\right) - \log \left(1+e^{\beta_{0}+\beta_{1}X}\right) \right\} \end{array}$$

We want to $\max_{\beta_0,\beta_1} \ell(\beta_0,\beta_1; D_n)$.

Newton's algorithm

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We follow standard process:

▶ Hence the log-likelihood logistic function is strictly concave.

$$\begin{bmatrix} \beta_0^{(k+1)} \\ \beta_1^{(k+1)} \end{bmatrix} = \begin{bmatrix} \beta_0^{(k)} \\ \beta_1^{(k)} \end{bmatrix} - (\nabla^2 \ell(\beta_0, \beta_1; D_n))^{-1} \cdot \nabla \ell(\beta_0, \beta_1; D_n)$$

"What are the odds?"

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One can see the logistic expression of the predictions from a different point-of-view:

$$q(x_i) := \frac{p(x_i)}{1 - p(x_i)} = e^{(\beta_0 + \beta_1 x_i)}.$$

odds function: often used in... Horse-racing!

"What are the odds?"

- If $q(x_i) = 1/4$, then $p(x_i = 1) = 0.2$
- If $q(x_i) = 9/1$, then $p(x_i = 1) = 0.9$.

The logits (or log-odds)

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$$Q(x_i) := \log \left(\frac{p(x_i)}{1 - p(x_i)} \right) = \beta_0 + \beta_1 x_i.$$

Here we come back to the expression for the Linear Regression!

Separating hyperplane: For p = 0.5, we get the "linear" boundary

$$0 = \beta_0 + \beta_1 x_{i,1} \quad \left(+ \beta_2 x_{i,2} + \ldots + \beta_K x_{i,K} \right), \qquad \text{for } K \geq 1.$$

e.g. for K=1, it is a point $x_{bound}=-\beta_0/\beta_1$. (left: 1, right: 0)

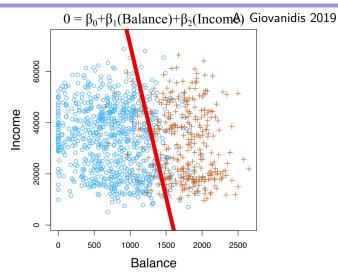


Figure: The boundary separates "blue" from "orange". ⁷

7_{Source [B.1]} 34 / 64

Test Data (Logistic)

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If we have test input data $x_o \notin D_n$, how do we choose its Class? Say $x_o = (x_{o,1}, x_{o,2}, \dots, x_{o,K})$.

Use the fitted values of $\beta_0, \beta_1, \dots, \beta_K$

- ► Either calculate $p(x_o) = \frac{e^{\beta_0 + \beta_1 x_o, 1 + \dots + \beta_K x_o, K}}{1 + e^{\beta_0 + \beta_1 x_o, 1 + \dots + \beta_K x_o, K}}$ and check if >, =, < 0.5,
- or check the position of x_o related to the boundary: $\beta_0 + \beta_1 x_{o,1} + \beta_2 x_{o,2} + \ldots + \beta_K x_{o,K} >, =, < 0$.

e.g.
$$\beta_0 + \beta_1 x_{o,1} + \beta_2 x_{o,2} + \ldots + \beta_K x_{o,K} > 0 \Rightarrow p(x_o) > 0.5$$

™ We need not always use the value of 0.5 for the boundary...

Multiple Logistic Regression

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We have implied that the Logistic Regression is generalised to higher than 1 dimension:

$$\log\left(\frac{p(X)}{1-p(X)}\right) = \beta_0 + \beta_1 X_1 + \ldots + \beta_K X_K,$$

where $X = (X_1, \dots, X_K)$ are K predictors.

Equivalently,

$$p(X) = \frac{e^{\beta_0 + \beta_1 X_1 + \ldots + \beta_K X_K}}{1 + e^{\beta_0 + \beta_1 X_1 + \ldots + \beta_K X_K}}.$$

 β_0, \ldots, β_K are estimated by the maximum likelihood method.

Example

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Using the data set $\operatorname{Default}$ we want to decide, whether an individual is likely to default on its bank account, or not.

X =(balance, income, student[Yes]), so K = 3.

Y = default[Yes].

• First consider only balance, K = 1.

	Coefficient	Std. error	Z-statistic	P-value
Intercept	-10.6513	0.3612	-29.5	< 0.0001
balance	0.0055	0.0002	24.9	< 0.0001

■ 1-unit increase in balance is associated to $\beta_1 = 0.0055$ units increase in log-odds of default.

Example (predictions)

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 $\operatorname{default}[\operatorname{Yes}]$ probability for an individual with $\operatorname{balance} = 1000$ EUR

$$\hat{\rho}(\text{balance} = 1000) \quad = \quad \frac{e^{-10.6513 + 0.0055 \times 1000}}{1 + e^{-10.6513 + 0.0055 \times 1000}} = 0.00576$$

• Now consider binary student[Yes], K = 1.

	Coefficient	Std. error	Z-statistic	P-value
Intercept	-3.5041	0.0707	-49.55	< 0.0001
student[Yes]	0.4049	0.1150	3.52	0.0004

$$\hat{p}(\text{student}[\text{Yes}] = 1) = 0.0431 > \hat{p}(\text{student}[\text{Yes}] = 0) = 0.0292$$

Conclusion 1: Students are more likely to default.

Example (multiple)

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• Now consider the entire X vector, K = 3.

	Coefficient	Std. error	Z-statistic	P-value
Intercept	-10.8690	0.4923	-22.08	< 0.0001
balance	0.0057	0.0002	24.74	< 0.0001
income	0.0030	0.0082	0.37	0.7115
student[Yes]	-0.6468	0.2362	-2.74	0.0062

Paradox: Conclusion 2: Students are less likely to default !!!! $(\beta_{\text{student}[Yes]} < 0)$

Why? The student[Yes] and balance predictors are correlated.

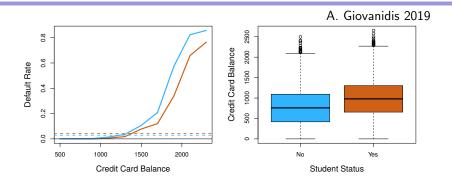


Figure: Students tend to have higher debts in the US/GB/D. 8

Conclusion 1: For the same credit-card balance a student is less likely to default.

⁸Source [B.1]

Logistic Regression for > 2 Classes

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We can easily generalise to M classes:

$$\log \frac{Pr(\textit{Class} = 1|X = x)}{Pr(\textit{Class} = M|X = x)} = \beta_{1,0} + \beta_1^T x$$

$$\dots$$

$$\log \frac{Pr(\textit{Class} = M - 1|X = x)}{Pr(\textit{Class} = M|X = x)} = \beta_{M-1,0} + \beta_{M-1}^T x$$

$$Pr(\textit{Class} = M|X = x) = \frac{1}{1 + \sum_{m=1}^{M-1} \exp(\beta_{m,0} + \beta_m^T x)}$$

- We need M-1 log-odds. The probabilities sum-up to 1.
- The choice of denominator class is arbitrary. Max likelihood.

For multiple classes, discriminant analysis is more popular...

Confusion Matrix

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		True	default	status
		No	Yes	Total
Predicted	No	9,644	252	9,896
$default\ status$	Yes	23	81	104
	Total	9,667	333	10,000

Figure: Confusion Matrix: Predicted vs True default status. ⁹

Two types of errors (True Negative, and False Positive)

- Incorrectly assign an individual who defaults to the "No default".
- ▶ Incorrectly assign an individual who does not default to "Default".

⁹ Source [B.1]

Class-specific Errors

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False Positive (Type I Error / False Alarm):

$$\textit{Error} \left[\widehat{\mathrm{Default}} = \text{``Yes''} | \mathrm{Default} = \text{``No''} \right] \quad = \quad 23/9667 \approx 0.2\%$$

True Negative (Type II Error / Mis-detection):

$$Error\left[\widehat{\mathrm{Default}} = "No" | \mathrm{Default} = "Yes" \right] = 252/333 \approx 75.7\%$$

While the overall error rate is low, the error rate among individuals who defaulted is very high!

If you were an insurance company, how would you rate these results?

Sensitivity and Specificity

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Sensitivity: The percentage of true defaulters that are identified

Sensitivity =
$$\left[\widehat{\text{Default}} = \text{"Yes"} | \text{Default} = \text{"Yes"} \right] =$$

True Positive = $\frac{81}{333} = 24.3\%$.

Specificity: The percentage of non-defaulters correctly identified

$$\begin{array}{lll} {\rm Specificity} & = & \left[\widehat{\rm Default} = "No" | {\rm Default} = "No" \right] = \\ & {\it False Negative} & = & \frac{9644}{9667} = 99.8\%. \end{array}$$

Very low sensitivity...

Improve Sensitivity?

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The Bayes classification assigns an observation to the default, if

$$Pr(\text{default} = \text{Yes} \mid X = x) > 0.5$$

If we are more concerned to detect individuals that actually default we should reduce True Negative!

Then **change the classification rule** to declare default when:

$$Pr(\text{default} = \text{Yes} \mid X = x) > 0.2.$$

New Confusion Matrix

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		True default status		
		No	Yes	Total
Predicted	No	9,432	138	9,570
$default\ status$	Yes	235	195	430
	Total	9,667	333	10,000

Figure: New Confusion Matrix: Predicted vs True default status. 10

- (New False Positive): $\frac{235}{9667} = 3.73\%$ (increased)
- (New True Negative): $\frac{138}{333} = 41.4\%$ (decreased)

¹⁰ Source [B.1]

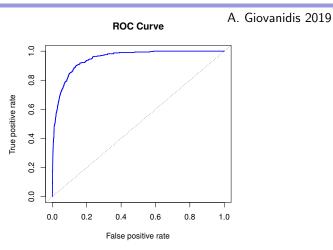


Figure: ROC Curve. Classifier performance = AUC 11

¹¹Source [B.1]

Definitions

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		Predicted class		
		– or Null	+ or Non-null	Total
True	– or Null	True Neg. (TN)	False Pos. (FP)	N
class	+ or Non-null	False Neg. (FN)	True Pos. (TP)	Р
	Total	N*	P*	

Name	Definition	Synonyms
False Pos. rate	FP/N	Type I error, 1—Specificity
True Pos. rate	TP/P	1—Type II error, power, sensitivity, recall
Pos. Pred. value	TP/P^*	Precision, 1—false discovery proportion
Neg. Pred. value	TN/N*	

Figure: Confusion Matrix and nomenclature¹²

¹² Source [B.1]

Linear Discriminant Analysis (LDA)

A. Giovanidis 2019

For classification of two or multiple classes, we often use the LDA classifier:

- ► Again, the class boundaries are linear.
- ▶ Instead of modelling Pr(Y = k | X = x) directly as in LR, it does this indirectly by modelling Pr(X = x | Y = k).
- It makes use of the Bayes' Theorem and the Bayes classifier.
- It assumes that the distribution of X's is approximately Normal, (or Gaussian).

Bayes' Theorem in Classification

A. Giovanidis 2019

We want to calculate the conditional probability for each class

$$Pr(Y = k | X = x) \stackrel{Bayes'}{=} \frac{Pr(X = x | Y = k) Pr(Y = k)}{Pr(X = x)}$$

$$\stackrel{Total}{=} \frac{Pr(X = x | Y = k) Pr(Y = k)}{\sum_{m=1}^{M} Pr(X = x | Y = m) Pr(Y = m)}$$

$$= \frac{f_k(x) \cdot \pi_k}{\sum_{m=1}^{M} f_m(x) \cdot \pi_m}$$
(1)

 \square We need the conditional probability of X given the class, and the frequency of each class.

Given these, we can choose for $X = x_o$, the class with $\max_{1 \le j \le M} Pr(Y = j | X = x_o)$ (Bayes classifier).

LDA for 1 predictor K = 1

A. Giovanidis 2019

We can **assume** that $f_k(x)$ is normal or Gaussian.

▶ For K = 1:

$$f_k(x) = \frac{1}{\sigma_k \sqrt{2\pi}} \exp\left(-\frac{1}{2\sigma_k^2}(x-\mu_k)^2\right),$$

 μ_k and σ_k^2 are the mean and variance for the k-th class.

- Let us further assume that $\sigma_1^2 = \sigma_2^2 = \dots = \sigma_M^2 = \sigma^2$, hence there is a shared variance among all classes.
- ▶ The π_m 's are also called prior probabilities.

Q: Is the gaussian assumption reasonable?

LDA
$$(K=1)$$

Plugging in (1), we get:

$$Pr(Y = k | X = x) = \frac{\frac{1}{\sigma\sqrt{2\pi}} \exp\left(-\frac{1}{2\sigma^2}(x - \mu_k)^2\right) \cdot \pi_k}{\sum_{m=1}^{M} \frac{1}{\sigma\sqrt{2\pi}} \exp\left(-\frac{1}{2\sigma^2}(x - \mu_m)^2\right) \cdot \pi_m}$$

Unknowns: π_m , μ_m , $\forall m$, and σ .

We take the log in the above expression. We then assign for X = x, the class m^* such that

$$m^* = \arg \max_{1 \le m \le M} \Pr(Y = m | X = x)$$

$$= \arg \max_{1 \le m \le M} \log \Pr(Y = m | X = x)$$

$$= \arg \max_{1 \le m \le M} \left\{ x \cdot \frac{\mu_m}{\sigma^2} - \frac{\mu_m^2}{2\sigma^2} + \log(\pi_m) \right\}$$

$$= \arg \max_{1 \le m \le M} \left\{ x \cdot c_1 + c_0 \right\} \quad \text{(linear!)}$$

Estimating the decision function

A. Giovanidis 2019

For each m we have the linear discriminant function function of x:

$$\delta_m(x) = x \cdot \frac{\mu_m}{\sigma^2} - \frac{\mu_m^2}{2\sigma^2} + \log(\pi_m),$$

and to calculate it from the dataset D_n we use the estimates:

$$\hat{\mu}_m = \frac{1}{n_m} \sum_{i:y_i=m} x_i,$$

$$\hat{\sigma}^2 = \frac{1}{n-M} \sum_{m=1}^M \sum_{i:y_i=m} (x_i - \hat{\mu}_m)^2,$$

$$\hat{\pi}_m = \frac{n_m}{n}.$$

2-class example

A. Giovanidis 2019

In the case of M=2 classes, suppose $\pi_1=\pi_2$ additionally. Then the discriminant functions become:

$$\delta_1(x) = x \cdot \frac{\mu_1}{\sigma^2} - \frac{\mu_1^2}{2\sigma^2} + \log(\pi_1)$$

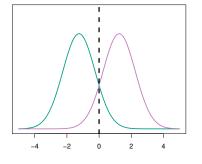
$$\delta_2(x) = x \cdot \frac{\mu_2}{\sigma^2} - \frac{\mu_2^2}{2\sigma^2} + \log(\pi_2)$$

so that x is assigned class 1, if $\delta_1(x) > \delta_2(x)$ or,

$$2x(\mu_1 - \mu_2) > \mu_1^2 - \mu_2^2$$

The decision boundary are the points x, s.t.

$$x = \frac{\mu_1 + \mu_2}{2}.$$



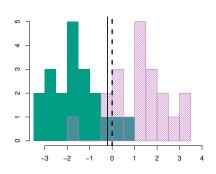


Figure: Two normal density functions and decision boundary. ¹³

¹³ Source [B.1]

LDA for K > 1 dimensions

A. Giovanidis 2019

How does the LDA perform, when the predictors X have more than 1 dimension? say $X = (X_1, \dots, X_K)$.

Assume a multivariate Gaussian distribution instead of a 1-dimensional $X \sim \mathcal{N}(\mu, \mathbf{\Sigma})$.

$$f(x) = \frac{1}{(2\pi)^{K/2} |\mathbf{\Sigma}|^{1/2}} \exp\left(-\frac{1}{2}(x-\mu)^T \mathbf{\Sigma}^{-1}(x-\mu)\right).$$

• mean $\mu = (\mu_1, \dots, \mu_K)$, • common covariance matrix Σ .

Linear Discriminant Function:

$$\delta_k(x) = x^T \mathbf{\Sigma}^{-1} \mu_k - \frac{1}{2} \mu_k^T \mathbf{\Sigma}^{-1} \mu_k + \log(\pi_k)$$

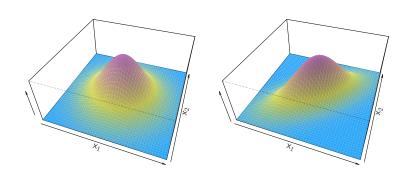


Figure: Examples of binormal distributions. 14

¹⁴Source [B.1]

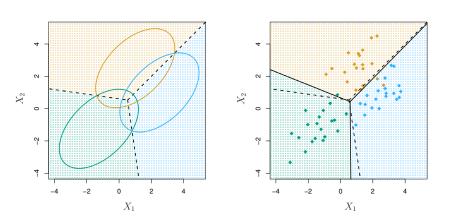


Figure: Classification for M=3 classes and K=2 dimensions. ¹⁵

¹⁵Source [B.1]

Quadratic Discriminant Analysis (QDA)

A. Giovanidis 2019

LDA assumed for each class a different mean μ_k and same covariance matrix Σ .

 \mathbb{Q} QDA assumes different covariance matrix per class. That is, an observation from the k-th class is of the form $X \sim \mathcal{N}(\mu_k, \Sigma_k)$.

Quadratic Discriminant Function:

$$\delta_k(x) = -\frac{1}{2}x^T \mathbf{\Sigma}_k^{-1} x + x^T \mathbf{\Sigma}_k^{-1} \mu_k - \frac{1}{2} \mu_k^T \mathbf{\Sigma}_k^{-1} \mu_k - \frac{1}{2} \log |\mathbf{\Sigma}_k| + \log(\pi_k)$$

QDA is more flexible than LDA: Bias vs Variance tradeoff!

QDA examples



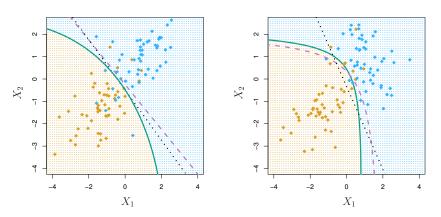


Figure: (left:) Truth common Σ , (right:) Truth different Σ_1 , Σ_2 . ¹⁶

¹⁶ Source [B.1]

Method comparison: linear

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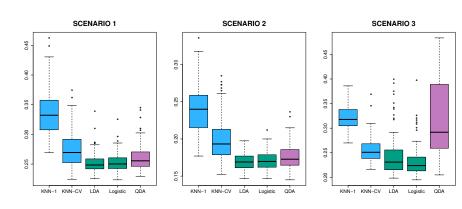


Figure: (1) uncorr., \mathcal{N} , $\mu_1 \neq \mu_2$, (2) corr., \mathcal{N} , (3) uncorr., t-distr. ¹⁷

¹⁷Source [B.1]

Method comparison: non-linear

A. Giovanidis 2019

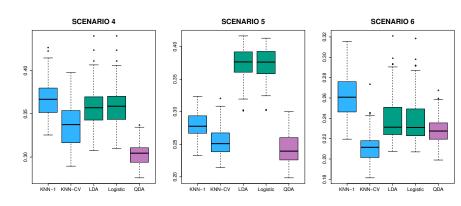


Figure: (4) corr. \mathcal{N} , $\Sigma_1 \neq \Sigma_2$, (5) logistic $X_1^2, X_2^2, X_1 X_2$ (6) more-NL. ¹⁸

¹⁸ Source [B.1]

END