#### A. Giovanidis 2019

### Feature Selection / Regularization

Data Analysis for Networks - DataNets'19 Anastasios Giovanidis

Sorbonne-LIP6







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### **Bibliography**

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B.1 Gareth James, Daniela Witten, Trevor Hastie, Robert Tibshirani. "An introduction to statistical learning: with applications in R". Springer Texts in Statistics. ISBN 978-1-4614-7137-0 Chapter 6 DOI 10.1007/978-1-4614-7138-7

### Intro

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In the multiple-regression setting, we assumed that the linear model with additive noise:

$$Y = \beta_0 + \beta_1 X_1 + \ldots + \beta_p X_p + \epsilon$$

describes the relationship between a response Y and a set of  $p \ge 1$  predictor variables  $X_1, X_2, \dots, X_p$ .

- The model fit uses least squares (LSs) to estimate the  $\hat{\beta}_i$ 's.

But, is it always a good fit? Are there any ways to improve this fit?

Feature Selection, Regularization, and Dimension Reduction.

### Main idea

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Either shrink the coefficients for some feature variables or remove them completely!

#### Why?

- ▶ Prediction Accuracy: If n >> p then LSs do have low variance. But when e.g.  $n \leq p$  the model is highly variable!
- Model Interpretability: Some variables used as predictors may not be relevant with the response. Better remove them to reduce model complexity.

### Methods

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We will present three main methods that modify the LSs:

- ▶ Subset Selection: Identify a subset of the original p predictors to be relevant (say  $p_s < p$ ). Then apply LSs fit.
- Shrinkage: Fit the model with all p features, but shrink some coefficients even to zero.
  - $\rightarrow$  This method reduces variance.
- ▶ Dimension Reduction: This method projects the p predictors to an M

### Best subset selection

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To find the best set one needs to perform LSs for all possible combinations fo the p predictors:

- ▶ All models with 1 predictor: p.
- ▶ All models with 2 predictors:  $\begin{pmatrix} p \\ 2 \end{pmatrix} = \frac{p(p-1)}{2}$ .
- etc.

In total 2<sup>p</sup> possibilities.

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#### Algorithm 6.1 Best subset selection

- 1. Let  $\mathcal{M}_0$  denote the *null model*, which contains no predictors. This model simply predicts the sample mean for each observation.
- 2. For  $k = 1, 2, \dots p$ :
  - (a) Fit all  $\binom{p}{k}$  models that contain exactly k predictors.
  - (b) Pick the best among these  $\binom{p}{k}$  models, and call it  $\mathcal{M}_k$ . Here best is defined as having the smallest RSS, or equivalently largest  $R^2$ .
- 3. Select a single best model from among  $\mathcal{M}_0, \dots, \mathcal{M}_p$  using cross-validated prediction error,  $C_p$  (AIC), BIC, or adjusted  $R^2$ .

Pick models with smallest Train RSS, Select with smallest Test RSS.

<sup>1</sup> Source [B.1]

### Many possibilities to test...

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The method needs to test too many feature combinations:

- for p = 10, approx 1,000 models,
- for p = 20, over 1,000,000 possibilities!
- etc.

The Best subset selection becomes computationally infeasible for large sets of features.

We need to find other ways to select good subsets stepwise.

### Forward Stepwise Selection

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#### The method:

- Begins with a model without predictors,
- adds predictors to the model one-at-a-time,
- until all predictors are in the model.

At each step the variable that gives the greatest additional improvement to the fit is added.

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#### Algorithm 6.2 Forward stepwise selection

- 1. Let  $\mathcal{M}_0$  denote the *null* model, which contains no predictors.
- 2. For  $k = 0, \ldots, p 1$ :
  - (a) Consider all p-k models that augment the predictors in  $\mathcal{M}_k$  with one additional predictor.
  - (b) Choose the *best* among these p k models, and call it  $\mathcal{M}_{k+1}$ . Here *best* is defined as having smallest RSS or highest  $R^2$ .
- 3. Select a single best model from among  $\mathcal{M}_0, \ldots, \mathcal{M}_p$  using cross-validated prediction error,  $C_p$  (AIC), BIC, or adjusted  $R^2$ .

Pick models with smallest Train RSS, Select with smallest Test RSS.

<sup>&</sup>lt;sup>2</sup>Source [B.1]

The method is computationally advantageous compared to Best selection:

- Instead of  $2^p$  fitting models, it needs to compute only  $1 + \sum_{k=0}^{p-1} (p-k) = 1 + p(p+1)/2$  models.
- e.g. for p = 20, fit 211 models instead of 1,048,576!
- ▶ Can be used also when n < p (stops at n features)

It is not guaranteed to find the best possible model out of the  $2^{p}$ .

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# Variables	Best subset	Forward stepwise
One	rating	rating
Two	rating, income	rating, income
Three	rating, income, student	rating, income, student
Four	cards, income,	rating, income,
	student, limit	student, limit

Source [B.1]

# **Backward Stepwise Selection**

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#### Algorithm 6.3 Backward stepwise selection

- 1. Let  $\mathcal{M}_p$  denote the full model, which contains all p predictors.
- 2. For  $k = p, p 1, \dots, 1$ :
  - (a) Consider all k models that contain all but one of the predictors in  $\mathcal{M}_k$ , for a total of k-1 predictors.
  - (b) Choose the *best* among these k models, and call it  $\mathcal{M}_{k-1}$ . Here *best* is defined as having smallest RSS or highest  $R^2$ .
- 3. Select a single best model from among  $\mathcal{M}_0, \ldots, \mathcal{M}_p$  using cross-validated prediction error,  $C_p$  (AIC), BIC, or adjusted  $R^2$ .

<sup>&</sup>lt;sup>4</sup>Source [B.1]

### Choosing the Optimal Model

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All methods hand-pick a small number of models based on a small value of Test RSS or high value of  $R^2$ .

To choose exactly one model among these, we need to find the one with smallest Test error. To do so we can:

- Estimate the Test Error, by adjusting the Train Error to account for Bias.
- Directly estimate the Test Error using a validation set or cross-validation.

### Adjustment of the Train Error

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We already know the Cross-Validation concept.

We can also adjust the Train error to select the model with best Test prediction:

- $C_{\rho}$ -estimate:  $C_{\rho} = \frac{1}{n} \left( RSS + 2d\hat{\sigma}^2 \right)$ .
- ► Akaike Information Criterion:  $AIC = \frac{1}{n\hat{\sigma}^2}(RSS + 2d\hat{\sigma}^2)$ .
- ▶ Bayesian Information Criterion:  $BIC = \frac{1}{n\hat{\sigma}^2}(RSS + \log(n)d\hat{\sigma}^2)$ .
- Adjusted  $R^2 = 1 \frac{RSS/(n-d-1)}{TSS/(n-1)}$ , where  $TSS = \sum (y_i \bar{y})^2$ .

### Shrinkage

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We have seen methods to optimally select a subset of appropriate features, leaving the rest out.

As an alternative, we can keep all p features, but use a technique that constraints or regularizes the coefficient estimates.

Estimates can be shrunk towards zero! This technique can significantly reduce variance.

Ridge Regression and the Lasso.

### Ridge Regression

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Similar to LSs fit, the Ridge Regression solves

$$\min_{\beta} \qquad \sum_{i=1}^{n} \left( y_i - \beta_0 - \sum_{j=1}^{p} \beta_j x_{ij} \right)^2 + \lambda \sum_{j=1}^{p} \beta_j^2$$

where the lefthand side is just the RSS, and  $\lambda \ge 0$  is a tuning parameter, to be determined.

The second term is called shrinkage penalty.

### **Properties**

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- ▶ When  $\lambda = 0$ : it is just the Least-Squares fit.
- ▶ When  $\lambda \to \infty$   $\beta$ 's will approach zero.
- ▶ Find the "best" set of parameters  $\beta$ .

 $\blacksquare$  Each choice of  $\lambda$  produces a different set of estimates  $\hat{\beta}_{\lambda}$ .

**Note 1:** The shrinkage penalty is **not** applied to the intercept  $\beta_0$ . **Note 2:** Best apply ridge-regression after standardizing the predictors (all on the same scale / standard deviation 1)

$$\tilde{x}_{ij} = \frac{x_{ij}}{\sqrt{\frac{1}{n} \sum_{i=1}^{n} (x_{ij} - \bar{x}_j)^2}}.$$

### Ridge Example

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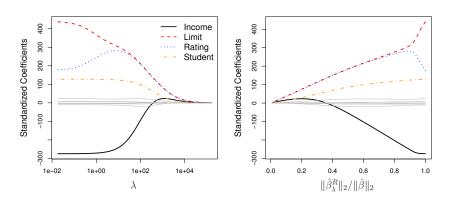


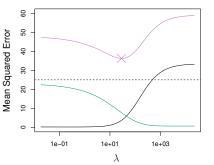
Figure: Change of Ridge Regression coefficients vs  $\lambda$ .<sup>5</sup>

<sup>&</sup>lt;sup>5</sup>Source [B.1]

### Improvement over LSs

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 $\square$  As  $\lambda$  increases, the flexibility of the ridge regression fit decreases: decreased variance but increased Bias.



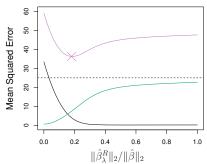


Figure: Bias vs Variance tradeoff and Test MSE.<sup>6</sup>

<sup>&</sup>lt;sup>6</sup>Source [B.1]

### The Lasso

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Similar to Ridge Regression, the Lasso solves a different problem

$$\min_{\beta} \qquad \sum_{i=1}^{n} \left( y_i - \beta_0 - \sum_{j=1}^{p} \beta_j x_{ij} \right)^2 + \lambda \sum_{j=1}^{p} |\beta_j|$$

where the lefthand side is just the RSS, and  $\lambda \geq 0$  is again a tuning parameter, to be determined.

The second term is the lasso penalty (uses  $\ell_1$ -norm instead of  $\ell_2$ ).

### Advantages

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As formulation, the Lasso is similar to Ridge Regression, with a penalty that uses a different norm.

What is new here?

- ► The Lasso penalty can force some estimates to be exactly zero
  → performs Variable Selection.
- Lasso's models are sparse involving a subset of variables.
- Simple, more interpretable models.

# Example Lasso

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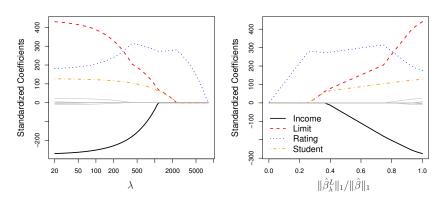


Figure: Change of Lasso coefficients vs  $\lambda$ .<sup>7</sup>

<sup>&</sup>lt;sup>7</sup>Source [B.1]

# **Equivalent Problems**

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The Ridge, Lasso, and Best subset selection are each equivalent to:

$$\begin{split} &\min_{\beta} \left\{ \sum_{i=1}^{n} \left( y_i - \beta_0 - \sum_{j=1}^{p} \beta_j x_{ij} \right)^2 \right\} \quad \textit{s.t.} \quad \sum_{j=1}^{p} \beta_j^2 \leq \textit{s} \quad (\textit{Ridge}) \\ &\min_{\beta} \left\{ \sum_{i=1}^{n} \left( y_i - \beta_0 - \sum_{j=1}^{p} \beta_j x_{ij} \right)^2 \right\} \quad \textit{s.t.} \quad \sum_{j=1}^{p} |\beta_j| \leq \textit{s} \quad (\textit{Lasso}) \\ &\min_{\beta} \left\{ \sum_{i=1}^{n} \left( y_i - \beta_0 - \sum_{j=1}^{p} \beta_j x_{ij} \right)^2 \right\} \quad \textit{s.t.} \quad \sum_{j=1}^{p} \mathbf{1} \left( \beta_j \neq 0 \right) \leq \textit{s} \quad (\textit{Best}). \end{split}$$

# Illustrative Explanation

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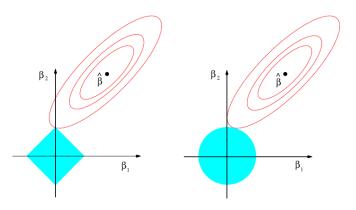


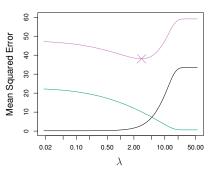
Figure: Why does Lasso lead to estimates equal to 0?8

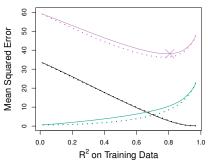
<sup>&</sup>lt;sup>8</sup>Source [B.1]

# Ridge > Lasso

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Here: Ridge needs all 45 coefficients  $\neq$  0. Lasso chose 2-out-of-45 features.



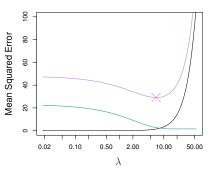


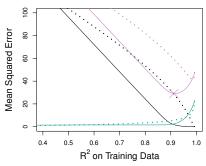
<sup>9</sup>Source [B.1]

### Ridge < Lasso

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Here: True response is a function of only 2 predictors and the rest are irrelevant. Ridge needs again all 45 coefficients  $\neq$  0. Lasso chose 2-out-of-45 features. <sup>10</sup>





<sup>&</sup>lt;sup>10</sup>Source [B.1]

# Special Case n = p

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 $\square$  Data centred around  $\bar{x}$ , no need for intercept.

▶ Least Squares: min  $\sum_{j=1}^{p} (y_j - \beta_j)^2$ . Solution:

$$\hat{\beta}_j = y_j.$$

► Ridge Regression: min  $\sum_{j=1}^{\rho} (y_j - \beta_j)^2 + \lambda \sum_{j=1}^{\rho} \beta_j^2$ . Solution:

$$\hat{\beta}_j^{(R)} = y_j/(1+\lambda).$$

▶ Lasso: min  $\sum_{j=1}^{p} (y_j - \beta_j)^2 + \lambda \sum_{j=1}^{p} |\beta_j|$ . Solution:

$$\hat{\beta}_{j}^{(L)} = \begin{cases} y_{j} - \lambda/2, & \text{if } y_{j} > \lambda/2 \\ y_{j} + \lambda/2, & \text{if } y_{j} < -\lambda/2 \\ 0, & \text{if } |y_{j}| \leq \lambda/2 \end{cases}$$

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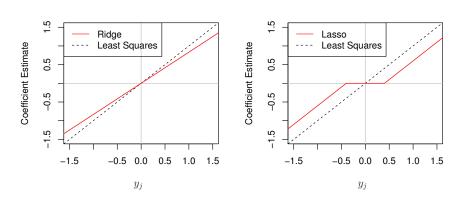


Figure: Ridge and Lasso coefficients over  $\lambda$ , compared to LSs. 11

<sup>&</sup>lt;sup>11</sup>Source [B.1]

# How to select parameter $\lambda$ ?

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For both Ridge and Lasso the tuning parameter  $\lambda$  (equivalently s) needs to be determined.

Again find the minimum Test MSE using Cross-Validation!

- ▶ Choose a grid of  $\lambda$  values.
- ▶ Compute the cross-validation error for each value of  $\lambda$ .
- ► Select the tuning parameter value with minimum CV error.
- ▶ Finally, re-fit the model using all observations and the chosen  $\lambda$ .

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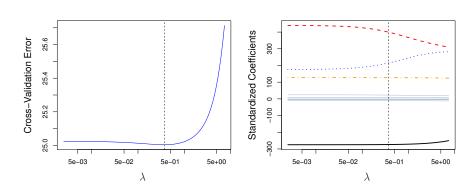


Figure: Ridge parameter tuning and comparison with LSs. 12

<sup>12</sup> Source [B.1]

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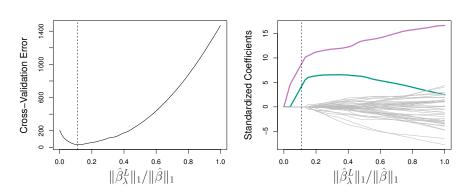


Figure: Lasso parameter tuning and comparison with LSs. 13

<sup>&</sup>lt;sup>13</sup>Source [B.1]

### **Dimension Reduction**

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Until here, methods control variance in 2 ways. From the original  $X_1, \ldots, X_p$  predictors,

- either choose a subset of the original variables, or
- shrink some of their coefficients to zero.

#### A new method!

- 1. first, Transform the predictors to a lower dimension  $Z_1, \ldots, Z_M$ , with M < p,
- 2. then Fit the LS model using the M predictors.

### Dimension Reduction cont'd

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 $Z_1, Z_2, \ldots, Z_M$ : linear combinations  $(M \le p)$  of our original p predictors.

$$Z_m = \sum_{j=1}^p \phi_{j,m} X_j.$$

**Unknowns:**  $\phi_{1,m}, \ldots, \phi_{p,m}$ , for  $m = 1, \ldots, M$ 

Then we can fit the linear regression model:

$$y_i = \theta_0 + \sum_{m=1}^M \theta_m z_{i,m} + \epsilon_i, \quad i = 1, \dots, n$$

using the Least Squares. The method can outperform simple LSs!

### Dimension Reduction cont'd II

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We need not estimate the p+1 coefficients:  $\beta_0, \beta_1, \ldots, \beta_p$ , but rather the less M+1 coefficients:  $\theta_0, \theta_1, \ldots, \theta_M$ .

$$\sum_{m=1}^{M} \theta_{m} z_{i,m} = \sum_{m=1}^{M} \theta_{m} \sum_{j=1}^{p} \phi_{j,m} x_{i,j}$$
$$= \sum_{j=1}^{p} \sum_{m=1}^{M} \theta_{m} \phi_{j,m} x_{i,j} = \sum_{j=1}^{p} \beta_{j} x_{i,j}.$$

so that, the  $\beta_i$ 's need to take the form:

$$\beta_j := \sum_{m=1}^M \theta_m \phi_{j,m}.$$

Works well when p is large compared to n data available.

# Principal Components Regression (PCR)

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Steps of the PCR Method:

- 1. Standardise  $X_j/\sqrt{Var(X_j)}$  and then centre  $(X_j-\bar{X}_j)$  all p features.
- 2. Choose a number  $M \leq p$  for the components.
- 3. Using Principal Component Analysis (PCA) define  $Z_1, \ldots, Z_M$  using the coefficients  $\phi_{j,m}$ ,  $j=1,\ldots,p$ .
- 4. Use  $Z_1, \ldots, Z_M$  as the predictors in a linear regression model that is fit using LSs.
- Using different number of components M choose the model which minimises the Test MSE.

The PCR is not a feature selection method! (uses linear combinations of all *p* original features)

# What is Principal Components Analysis?

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PCA is a technique to reduce the dimension of a  $(n \times p)$  data matrix **X**:

$$X = UDV^T$$

where **U** is  $(n \times p)$ , **V** is  $(p \times p)$  and **D** is a  $p \times p$  diagonal matrix, with diagonal entries  $d_1 \geq d_2 \geq \ldots \geq d_p \geq 0$ , called the singular values of **X**.

Intuition: the right matrix V and the square of the singular value matrix  $D^2$  also determine the eigen-decomposition of the sample covariance matrix  $S = X^T X/n$ ,

$$\mathbf{X}^{T}\mathbf{X} = (\mathbf{U}\mathbf{D}\mathbf{V}^{T})^{T}(\mathbf{U}\mathbf{D}\mathbf{V}^{T})$$

$$= \mathbf{V}\mathbf{D}\mathbf{U}^{T}\mathbf{U}\mathbf{D}\mathbf{V}^{T}$$

$$= \mathbf{V}\mathbf{D}^{2}\mathbf{V}^{T}$$

### PCA cont'd

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The principal components directions of X are the columns of V.

The first principal component direction has the property that  $\mathbf{Z}_1 = \mathbf{X} v_1$  has the largest sample variance, equal to  $d_1^2/n$ .

- ▶ This practically means that along the direction  $v_1$ , the data vary the most (if projected).
- ▶ Also, this direction defines the line which is as close as possible to the data!

# PCA Example

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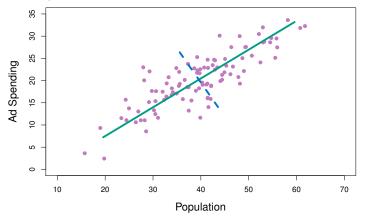


Figure: Principal components directions for p = 2.14

<sup>&</sup>lt;sup>14</sup>Source [B.1]

# PCA Example cont'd

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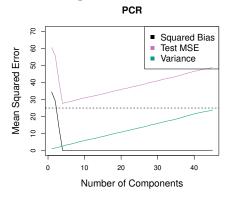
The two PCA directions are:

$$Z_1 = 0.839 \times (pop - \overline{pop}) + 0.544 \times (ad - \overline{ad})$$

$$Z_2 = 0.544 \times (pop - \overline{pop}) - 0.839 \times (ad - \overline{ad})$$

- ▶ With p = 2, we can construct at most M = 2 linear combinations.
- ▶ The directions (0.839, 0.544) and (0.544, -0.839) are orthogonal!
- ▶ For the Regression using PCA, we need to choose the sets  $\{Z_1\}$  and  $\{Z_1, Z_2\}$  and find which one leads to smallest Test MSE.

### PCR vs Ridge and Lasso



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#### Ridge Regression and Lasso

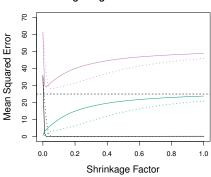


Figure: p = 45 features: PCR vs Lasso (solid) and Ridge (dotted). <sup>15</sup>

<sup>&</sup>lt;sup>15</sup>Source [B.1]

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Dimension Reduction
Principal Components

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# **END**