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Feature Selection / Regularization

Data Analysis for Networks - DataNets'19 Anastasios Giovanidis

Sorbonne-LIP6







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Bibliography

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Intro

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In the multiple-regression setting, we assumed that the linear model with additive noise:

$$Y = \beta_0 + \beta_1 X_1 + \ldots + \beta_p X_p + \epsilon$$

describes the relationship between a response Y and a set of $p \ge 1$ predictor variables X_1, X_2, \dots, X_p .

- The model fit uses least squares (LSs) to estimate the $\hat{\beta}_i$'s.

But, is it always a good fit? Are there any ways to improve this fit?

Feature Selection, Regularization, and Dimension Reduction.

Main idea

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Either shrink the coefficients for some feature variables or remove them completely!

Why?

- ▶ Prediction Accuracy: If n >> p then LSs do have low variance. But when e.g. $n \leq p$ the model is highly variable!
- Model Interpretability: Some variables used as predictors may not be relevant with the response. Better remove them to reduce model complexity.

Methods

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We will present three main methods that modify the LSs:

- ▶ Subset Selection: Identify a subset of the original p predictors to be relevant (say $p_s < p$). Then apply LSs fit.
- Shrinkage: Fit the model with all p features, but shrink some coefficients even to zero.
 - \rightarrow This method reduces variance.
- ▶ Dimension Reduction: This method projects the p predictors to an M

Best subset selection

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To find the best set one needs to perform LSs for all possible combinations fo the p predictors:

- ▶ All models with 1 predictor: p.
- ► All models with 2 predictors: $\begin{pmatrix} p \\ 2 \end{pmatrix} = \frac{p(p-1)}{2}$.
- etc.

In total 2^p possibilities.

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Algorithm 6.1 Best subset selection

- 1. Let \mathcal{M}_0 denote the *null model*, which contains no predictors. This model simply predicts the sample mean for each observation.
- 2. For $k = 1, 2, \dots p$:
 - (a) Fit all $\binom{p}{k}$ models that contain exactly k predictors.
 - (b) Pick the best among these $\binom{p}{k}$ models, and call it \mathcal{M}_k . Here best is defined as having the smallest RSS, or equivalently largest R^2 .
- 3. Select a single best model from among $\mathcal{M}_0, \dots, \mathcal{M}_p$ using cross-validated prediction error, C_p (AIC), BIC, or adjusted R^2 .

Pick models with smallest Train RSS, Select with smallest Test RSS.

¹ Source [B.1]

Many possibilities to test...

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The method needs to test too many feature combinations:

- for p = 10, approx 1,000 models,
- for p = 20, over 1,000,000 possibilities!
- etc.

The Best subset selection becomes computationally infeasible for large sets of features.

We need to find other ways to select good subsets stepwise.

Forward Stepwise Selection

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The method:

- Begins with a model without predictors,
- adds predictors to the model one-at-a-time,
- until all predictors are in the model.

At each step the variable that gives the greatest additional improvement to the fit is added.

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Algorithm 6.2 Forward stepwise selection

- 1. Let \mathcal{M}_0 denote the *null* model, which contains no predictors.
- 2. For $k = 0, \ldots, p 1$:
 - (a) Consider all p-k models that augment the predictors in \mathcal{M}_k with one additional predictor.
 - (b) Choose the *best* among these p k models, and call it \mathcal{M}_{k+1} . Here *best* is defined as having smallest RSS or highest R^2 .
- 3. Select a single best model from among $\mathcal{M}_0, \ldots, \mathcal{M}_p$ using cross-validated prediction error, C_p (AIC), BIC, or adjusted R^2 .

Pick models with smallest Train RSS, Select with smallest Test RSS.

²Source [B.1]

Advantages

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The method is computationally advantageous compared to Best selection:

- Instead of 2^p fitting models, it needs to compute only $1 + \sum_{k=0}^{p-1} (p-k) = 1 + p(p+1)/2$ models.
- e.g. for p = 20, fit 211 models instead of 1,048,576!
- ▶ Can be used also when n < p (stops at n features)

It is not guaranteed to find the best possible model out of the 2^{p} .

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# Variables	Best subset	Forward stepwise
One	rating	rating
Two	rating, income	rating, income
Three	rating, income, student	rating, income, student
Four	cards, income,	rating, income,
	student, limit	student, limit

Source [B.1]

•

Algorithm 6.3 Backward stepwise selection

- 1. Let \mathcal{M}_p denote the full model, which contains all p predictors.
- 2. For $k = p, p 1, \dots, 1$:
 - (a) Consider all k models that contain all but one of the predictors in \mathcal{M}_k , for a total of k-1 predictors.
 - (b) Choose the *best* among these k models, and call it \mathcal{M}_{k-1} . Here best is defined as having smallest RSS or highest R^2 .
- 3. Select a single best model from among $\mathcal{M}_0, \ldots, \mathcal{M}_p$ using cross-validated prediction error, C_p (AIC), BIC, or adjusted R^2 .

⁴Source [B.1]

Choosing the Optimal Model

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All methods hand-pick a small number of models based on a small value of Test RSS or high value of R^2 .

To choose exactly one model among these, we need to find the one with smallest Test error. To do so we can:

- Estimate the Test Error, by adjusting the Train Error to account for Bias.
- Directly estimate the Test Error using a validation set or cross-validation.

Adjustment of the Train Error

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We already know the Cross-Validation concept.

We can also adjust the Train error to select the model with best Test prediction:

- C_{ρ} -estimate: $C_{\rho} = \frac{1}{n} \left(RSS + 2d\hat{\sigma}^2 \right)$.
- ► Akaike Information Criterion: $AIC = \frac{1}{n\hat{\sigma}^2}(RSS + 2d\hat{\sigma}^2)$.
- ▶ Bayesian Information Criterion: $BIC = \frac{1}{n\hat{\sigma}^2}(RSS + \log(n)d\hat{\sigma}^2)$.
- Adjusted $R^2 = 1 \frac{RSS/(n-d-1)}{TSS/(n-1)}$, where $TSS = \sum (y_i \bar{y})^2$.

Shrinkage

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We have seen methods to optimally select a subset of appropriate features, leaving the rest out.

As an alternative, we can keep all p features, but use a technique that constraints or regularizes the coefficient estimates.

Estimates can be shrunk towards zero! This technique can significantly reduce variance.

Ridge Regression and the Lasso.

Ridge Regression

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Similar to LSs fit, the Ridge Regression solves

$$\min_{\beta} \qquad \sum_{i=1}^{n} \left(y_i - \beta_0 - \sum_{j=1}^{p} \beta_j x_{ij} \right)^2 + \lambda \sum_{j=1}^{p} \beta_j^2$$

where the lefthand side is just the RSS, and $\lambda \ge 0$ is a tuning parameter, to be determined.

The second term is called shrinkage penalty.

Properties

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- ▶ When $\lambda = 0$: it is just the Least-Squares fit.
- ▶ When $\lambda \to \infty$ β 's will approach zero.
- ▶ Find the "best" set of parameters β .

 \blacksquare Each choice of λ produces a different set of estimates $\hat{\beta}_{\lambda}$.

Note 1: The shrinkage penalty is **not** applied to the intercept β_0 . **Note 2:** Best apply ridge-regression after standardizing the predictors (all on the same scale / standard deviation 1)

$$\tilde{x}_{ij} = \frac{x_{ij}}{\sqrt{\frac{1}{n} \sum_{i=1}^{n} (x_{ij} - \bar{x}_j)^2}}.$$

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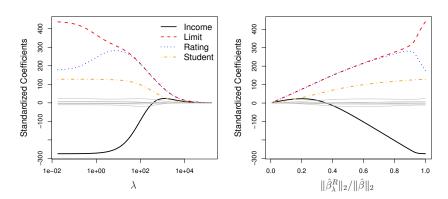


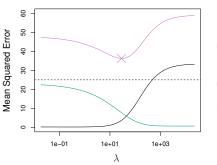
Figure: Change of Ridge Regression coefficients vs λ .⁵

⁵Source [B.1]

Improvement over LSs

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 \square As λ increases, the flexibility of the ridge regression fit decreases: decreased variance but increased Bias.



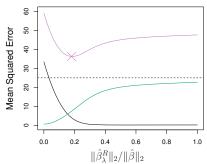


Figure: Bias vs Variance tradeoff and Test MSE.⁶

⁶Source [B.1]

The Lasso

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Similar to Ridge Regression, the Lasso solves a different problem

$$\min_{\beta} \qquad \sum_{i=1}^{n} \left(y_i - \beta_0 - \sum_{j=1}^{p} \beta_j x_{ij} \right)^2 + \lambda \sum_{j=1}^{p} |\beta_j|$$

where the lefthand side is just the RSS, and $\lambda \geq 0$ is again a tuning parameter, to be determined.

The second term is the lasso penalty (uses ℓ_1 -norm instead of ℓ_2).

Advantages

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As formulation, the Lasso is similar to Ridge Regression, with a penalty that uses a different norm.

What is new here?

- ► The Lasso penalty can force some estimates to be exactly zero
 → performs Variable Selection.
- Lasso's models are sparse involving a subset of variables.
- Simple, more interpretable models.

Example Lasso

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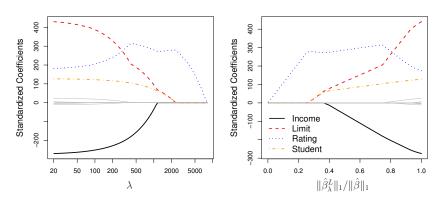


Figure: Change of Lasso coefficients vs λ .⁷

⁷Source [B.1]

Equivalent Problems

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The Ridge, Lasso, and Best subset selection are each equivalent to:

$$\begin{split} &\min_{\beta} \left\{ \sum_{i=1}^{n} \left(y_i - \beta_0 - \sum_{j=1}^{p} \beta_j x_{ij} \right)^2 \right\} \quad \textit{s.t.} \quad \sum_{j=1}^{p} \beta_j^2 \leq \textit{s} \quad (\textit{Ridge}) \\ &\min_{\beta} \left\{ \sum_{i=1}^{n} \left(y_i - \beta_0 - \sum_{j=1}^{p} \beta_j x_{ij} \right)^2 \right\} \quad \textit{s.t.} \quad \sum_{j=1}^{p} |\beta_j| \leq \textit{s} \quad (\textit{Lasso}) \\ &\min_{\beta} \left\{ \sum_{i=1}^{n} \left(y_i - \beta_0 - \sum_{j=1}^{p} \beta_j x_{ij} \right)^2 \right\} \quad \textit{s.t.} \quad \sum_{j=1}^{p} \mathbf{1} \left(\beta_j \neq 0 \right) \leq \textit{s} \quad (\textit{Best}). \end{split}$$

Illustrative Explanation

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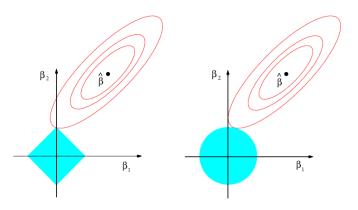


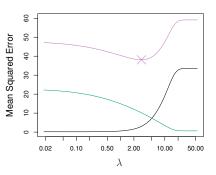
Figure: Why does Lasso lead to estimates equal to 0?8

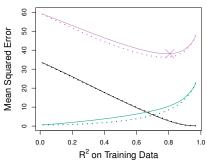
⁸Source [B.1]

Ridge > Lasso

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Here: Ridge needs all 45 coefficients \neq 0. Lasso chose 2-out-of-45 features.⁹



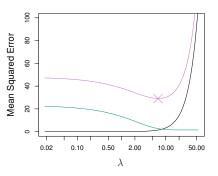


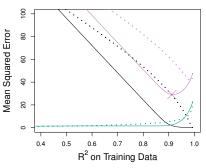
⁹Source [B.1]

Ridge < Lasso

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Here: True response is a function of only 2 predictors and the rest are irrelevant. Ridge needs again all 45 coefficients \neq 0. Lasso chose 2-out-of-45 features. ¹⁰





¹⁰Source [B.1]

Special Case n = p

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 \square Data centred around \bar{x} , no need for intercept.

▶ Least Squares: min $\sum_{j=1}^{p} (y_j - \beta_j)^2$. Solution:

$$\hat{\beta}_j = y_j.$$

► Ridge Regression: min $\sum_{j=1}^{\rho} (y_j - \beta_j)^2 + \lambda \sum_{j=1}^{\rho} \beta_j^2$. Solution:

$$\hat{\beta}_j^{(R)} = y_j/(1+\lambda).$$

▶ Lasso: min $\sum_{j=1}^{p} (y_j - \beta_j)^2 + \lambda \sum_{j=1}^{p} |\beta_j|$. Solution:

$$\hat{\beta}_{j}^{(L)} = \begin{cases} y_{j} - \lambda/2, & \text{if } y_{j} > \lambda/2 \\ y_{j} + \lambda/2, & \text{if } y_{j} < -\lambda/2 \\ 0, & \text{if } |y_{j}| \leq \lambda/2 \end{cases}$$

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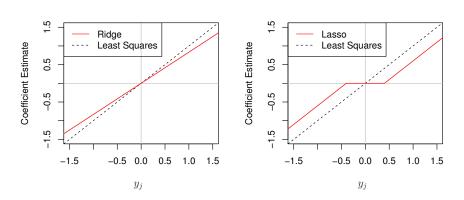


Figure: Ridge and Lasso coefficients over λ , compared to LSs. 11

¹¹Source [B.1]

How to select parameter λ ?

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For both Ridge and Lasso the tuning parameter λ (equivalently s) needs to be determined.

Again find the minimum Test MSE using Cross-Validation!

- ▶ Choose a grid of λ values.
- ▶ Compute the cross-validation error for each value of λ .
- Select the tuning parameter value with minimum CV error.
- ▶ Finally, re-fit the model using all observations and the chosen λ .

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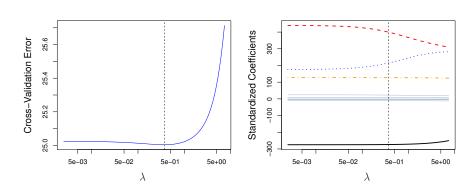


Figure: Ridge parameter tuning and comparison with LSs. 12

¹² Source [B.1]

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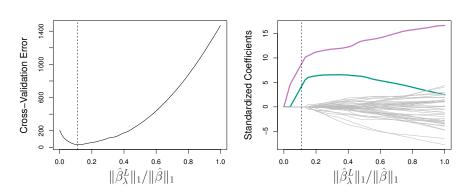


Figure: Lasso parameter tuning and comparison with LSs. 13

¹³Source [B.1]

Dimension Reduction

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Until here, methods control variance in 2 ways. From the original X_1, \ldots, X_p predictors,

- either choose a subset of the original variables, or
- shrink some of their coefficients to zero.

A new method!

- 1. first, Transform the predictors to a lower dimension Z_1, \ldots, Z_M , with M < p,
- 2. then Fit the LS model using the M predictors.

Dimension Reduction cont'd

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 Z_1, Z_2, \ldots, Z_M : linear combinations $(M \leq p)$ of our original p predictors.

$$Z_m = \sum_{j=1}^p \phi_{j,m} X_j.$$

Unknowns: $\phi_{1,m}, \ldots, \phi_{p,m}$, for $m = 1, \ldots, M$

Then we can fit the linear regression model:

$$y_i = \theta_0 + \sum_{m=1}^M \theta_m z_{i,m} + \epsilon_i, \quad i = 1, \dots, n$$

using the Least Squares. The method can outperform simple LSs!

Dimension Reduction cont'd II

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We need not estimate the p+1 coefficients: $\beta_0, \beta_1, \ldots, \beta_p$, but rather the less M+1 coefficients: $\theta_0, \theta_1, \ldots, \theta_M$.

$$\sum_{m=1}^{M} \theta_{m} z_{i,m} = \sum_{m=1}^{M} \theta_{m} \sum_{j=1}^{p} \phi_{j,m} x_{i,j}$$
$$= \sum_{j=1}^{p} \sum_{m=1}^{M} \theta_{m} \phi_{j,m} x_{i,j} = \sum_{j=1}^{p} \beta_{j} x_{i,j}.$$

so that, the β_i 's need to take the form:

$$\beta_j := \sum_{m=1}^M \theta_m \phi_{j,m}.$$

Works well when p is large compared to n data available.

Partial Least Squares (PLS)

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PLS is a dimension reduction method:

- 1. First, identifies a new set of features $Z_1, \ldots, Z_M, M < p$ in a supervised way,
- 2. then fits a linear model via LSs using the M features.

Supervised way: means uses the response Y to identify new features, related to the response.

PLS cont'd

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PLS steps:

- Standardise and centre the p predictors.
- ▶ Z_1 : Do LSs for all pairs $(X_j, Y) \to \text{use } \phi_{j,1} := \beta_{j,1}, \ i = 1, \dots, p$. In $Z_1 = \sum_{j=1}^p \phi_{j,1} X_1$, PLS places the highest weight on the variables most strongly related to the response Y.
- ▶ Z_2 : Do LSs for all pairs (X_j, Z_1) and find new $X_j^{(2)} = X_j \gamma_{j,1} Z_1$. Then do LSs for all pairs $(X_j^{(2)}, Y) \rightarrow \text{use } \phi_{j,1} := \beta_{j,1}^{(2)}, \ i = 1, \dots, p$.
- ► Z₃: ...

The PLS seeks directions that have high variance and high correlation with the response.

Principal Components Regression (PCR)

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Steps of the PCR Method:

- 1. Standardise $X_j/\sqrt{Var(X_j)}$ and then centre $(X_j \bar{X}_j)$ all p features.
- 2. Choose a number $M \leq p$ for the components.
- 3. Using Principal Component Analysis (PCA) define Z_1, \ldots, Z_M using the coefficients $\phi_{j,m}$, $j=1,\ldots,p$.
- 4. Use Z_1, \ldots, Z_M as the predictors in a linear regression model that is fit using LSs.
- Using different number of components M choose the model which minimises the Test MSE.

The PCR is not a feature selection method! (uses linear combinations of all *p* original features)

PCR example

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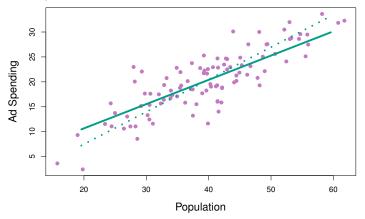


Figure: PCR direction Z_1 for p = 2.14

¹⁴Source [B.1]

What is Principal Components Analysis?

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PCA is a technique to reduce the dimension of a $(n \times p)$ data matrix **X**:

$$X = UDV^T$$

where **U** is $(n \times p)$, **V** is $(p \times p)$ and **D** is a $p \times p$ diagonal matrix, with diagonal entries $d_1 \geq d_2 \geq \ldots \geq d_p \geq 0$, called the singular values of **X**.

Intuition: the right matrix V and the square of the singular value matrix D^2 also determine the eigen-decomposition of the sample covariance matrix $S = X^T X/n$,

$$\mathbf{X}^{T}\mathbf{X} = (\mathbf{U}\mathbf{D}\mathbf{V}^{T})^{T}(\mathbf{U}\mathbf{D}\mathbf{V}^{T})$$

$$= \mathbf{V}\mathbf{D}\mathbf{U}^{T}\mathbf{U}\mathbf{D}\mathbf{V}^{T}$$

$$= \mathbf{V}\mathbf{D}^{2}\mathbf{V}^{T}$$

PCA cont'd

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The principal components directions of X are the columns of V.

The first principal component direction has the property that $\mathbf{Z}_1 = \mathbf{X} \mathbf{v}_1$ has the largest sample variance, equal to d_1^2/n .

- ▶ This practically means that along the direction v_1 , the data vary the most (if projected).
- ▶ Also, this direction defines the line which is as close as possible to the data!

PCA Example

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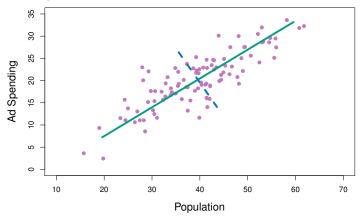


Figure: Principal components directions for p = 2.¹⁵

¹⁵Source [B.1]

PCA Example cont'd

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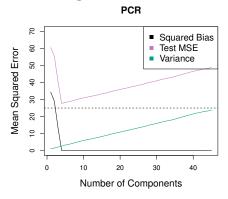
The two PCA directions are:

$$Z_1 = 0.839 \times (pop - \overline{pop}) + 0.544 \times (ad - \overline{ad})$$

$$Z_2 = 0.544 \times (pop - \overline{pop}) - 0.839 \times (ad - \overline{ad})$$

- ▶ With p = 2, we can construct at most M = 2 linear combinations.
- ▶ The directions (0.839, 0.544) and (0.544, -0.839) are orthogonal!
- ▶ For the Regression using PCA, we need to choose the sets $\{Z_1\}$ and $\{Z_1, Z_2\}$ and find which one leads to smallest Test MSE.

PCR vs Ridge and Lasso



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Ridge Regression and Lasso

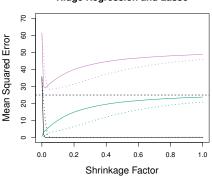


Figure: p = 45 features: PCR vs Lasso (solid) and Ridge (dotted). ¹⁶

¹⁶Source [B.1]

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Dimension Reduction
Principal Components

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END