

Towards Efficient Random-Order Enumeration for Join Queries

Technical Report

0.1 Proof of Lemma 1.

PROOF. To simplify the analysis, we assume that $\text{RRACCESS}^{Q,\varphi}(i)$ returns the interval $[i, i]$ when $\varphi_Q(i) = \perp$. This choice of a minimal trivial interval, while less efficient, enables us to derive upper bounds on both the total running time and the delay of Algorithm 2.

Total running time. Since $\log |B| \leq \log N \leq O(\log |Q|)$, all of $B.\text{ban}$, $B.\text{pick}$ and RRACCESS run in $O(\log |Q|)$ amortized time. Moreover, since they run at most N times, the total running time of Algorithm 2 is at most $O(N \log |Q|)$.

Enumeration delay. Let $N_i (0 \leq i \leq |\text{Res}(Q)|)$ be the random variable of the number of times that $\text{RRACCESS}^{Q,\varphi}$ returns \perp in the process of outputting the first i result tuples (the $(i + N_i)$ -th run of $\text{RRACCESS}^{Q,\varphi}$ returns the i -th result tuple, and in particular, $N_0 = 0$). By the definition of φ_Q , for each tuple $t \in \text{Res}(Q)$, there exists only one integer $i \in \mathbb{N}[1, N]$ such that $\varphi_Q(i) = t$. Then there are exactly $|\text{Res}(Q)|$ non-trivial integers in $\mathbb{N}[1, N]$, and for each $0 \leq i \leq |\text{Res}(Q)|$, N_i is the number of trivial integers obtained by sampling (without replacement) from $\mathbb{N}[1, N]$ until exactly i non-trivial integers are picked. This implies that for each $1 \leq i \leq |\text{Res}(Q)|$, N_i follows the negative hypergeometric distribution [1] with parameters N , $N - |\text{Res}(Q)|$ and i , and its expectation is:

$$E[N_i] = \frac{i(N - |\text{Res}(Q)|)}{|\text{Res}(Q)| + 1}. \quad (1)$$

Therefore, the expected number of calls of $\text{RRACCESS}^{Q,\varphi}$ between the output of the $(i - 1)$ -th result (the start of the algorithm when $i = 1$) and the output of the i -th result is

$$E[N_i - N_{i-1} + 1] = \frac{N + 1}{|\text{Res}(Q)| + 1} = O\left(\frac{N}{|\text{Res}(Q)| + 1}\right). \quad (2)$$

And the expected running time of $\text{RRACCESS}^{Q,\varphi}$ between the output of the last result tuple and the end of the algorithm is

$$E[N - N_{|\text{Res}(Q)|} - |\text{Res}(Q)|] = \frac{N - |\text{Res}(Q)|}{|\text{Res}(Q)| + 1} = O\left(\frac{N}{|\text{Res}(Q)| + 1}\right). \quad (3)$$

Then, since $\text{RRACCESS}^{Q,\varphi}$ runs in $O(\log^2 |Q|)$ worst-case time, the expected delay of Algorithm 2 is $O(\frac{N}{|\text{Res}(Q)| + 1} \log^2 |Q|)$. \square

0.2 Proof of Lemma 2.

PROOF. We prove the contrapositive: if $\varphi_Q^*(i) \neq \text{false}$, then $i \leq \text{upp}(\psi_r)$. By the definition of φ^* , for each integer $i > 0$ satisfying $\varphi_Q^*(i) \neq \text{false}$, there is a tuple $t \in \text{Res}(Q)$ along with a root-to-leaf path u_1, \dots, u_h ($u_1 = r, u_h = u_t$) such that (11) holds. Then, let

$$\begin{aligned} \text{sum}_j &= \sum_{\substack{\psi \prec \psi_{u_{j+1}} \\ \psi \in \text{children}(\psi_{u_j})}} \text{upp}(\psi) \\ &\leq \sum_{\psi \in \text{children}(\psi_{u_j})} \text{upp}(\psi) - \text{upp}(\psi_{u_{j+1}}) \\ &\leq \text{upp}(\psi_{u_j}) - \text{upp}(\psi_{u_{j+1}}), \end{aligned} \quad (4)$$

therefore

$$\begin{aligned} i &= \sum_{j=1}^{h-1} \text{sum}_j + 1 \leq \sum_{j=1}^{h-1} (\text{upp}(\psi_{u_j}) - \text{upp}(\psi_{u_{j+1}})) + 1 \\ &= \text{upp}(\psi_{u_1}) - \text{upp}(\psi_{u_h}) + 1 = \text{upp}(\psi_r). \end{aligned} \quad (5)$$

\square

0.3 Proof of Theorem 2.

PROOF. For $\forall \epsilon > 0$, if there exists a random-order enumeration algorithm \mathcal{A} for join queries with expected $\tilde{O}(\frac{|Q|^{\rho^* - \epsilon}}{|\text{Res}(Q)| + 1})$ delay after a $\tilde{O}(|Q|)$ -time preprocessing, we can define \mathcal{B} as the random algorithm obtained by executing \mathcal{A} until the first result tuple is output and then stopping the enumeration. Then \mathcal{B} is a uniform sampling algorithm for join queries, and according to Markov's inequality, \mathcal{B} runs in $\tilde{O}(|Q| + \frac{|Q|^{\rho^* - \epsilon}}{|\text{Res}(Q)|})$ time with high probability. This contradicts Theorem 1. \square

0.4 Proof of Lemma 4.

PROOF. For each tuple $t \in R$ satisfying ψ , we have $l_i \leq t(x_i) \leq h_i$ for $\forall i \in \{r_1, \dots, r_d\}$. This implies $t_l^\psi \preceq t \preceq t_h^\psi$, and consequently $t \in R[t_l^\psi, t_h^\psi]$. Therefore, $R|_\psi \subseteq R[t_l^\psi, t_h^\psi]$. For each tuple $t \in R[t_l^\psi, t_h^\psi]$, we have $t_l^\psi \preceq t \preceq t_h^\psi$. We need to prove that for each $x_i \in \text{att}(R)$, $l_i \leq t(x_i) \leq h_i$. Let s be a split position of ψ , then:

- (1) For $\forall x_i \in \text{att}(R)$ with $\forall 1 \leq i < s$, since $l_i = h_i$, it follows that $l_i = t(x_i) = h_i$. Otherwise, one of the conditions $t \succeq t_l^\psi$ or $t \preceq t_h^\psi$ would be violated.
- (2) If $x_s \in \text{att}(R)$ then $l_s \leq t(x_s) \leq h_s$, otherwise $t_l^\psi \preceq t \preceq t_h^\psi$ would be violated.
- (3) For $\forall x_i \in \text{att}(R)$ with $\forall s < i \leq n$, since $l_i = \min_Q(x_i)$ and $h_i = \max_Q(x_i)$, it follows that $l_i \leq t(x_i) \leq h_i$.

Then t satisfies ψ , i.e., $t \in R|_\psi$. Therefore, $R[t_l^\psi, t_h^\psi] \subseteq R|_\psi$ and then we have $R[t_l^\psi, t_h^\psi] = R|_\psi$, which implies that $R.\text{cnt}(\psi) = |R|_\psi$. \square

0.5 Proof of Lemma 6.

PROOF. Since $\text{upp}(\psi) \leq 1$, two cases need to be discussed. Firstly, for the case $\text{upp}(\psi) = 0$, we have $|\text{Res}(Q|_\psi)| \leq \text{upp}(\psi) = 0$, which implies that $\text{Res}(Q|_\psi) = \emptyset$. Secondly, for the case $\text{upp}(\psi) = 1$, we have $|\text{Res}(Q|_\psi)| \leq 1$. Let $s = \max\{i | l_i \neq h_i\}$, by Property 1, there exists at most one integer $p \in \mathbb{N}[l_s, h_s]$ such that

- $\text{upp}([l_1, h_1], \dots, [l_s, p - 1], \dots, [l_n, h_n]) = 0$,
- $\text{upp}([l_1, h_1], \dots, [p, p], \dots, [l_n, h_n]) = 1$,
- $\text{upp}([l_1, h_1], \dots, [p + 1, h_s], \dots, [l_n, h_n]) = 0$.

If $\text{upp}([l_1, h_1], \dots, [p, p], \dots, [l_n, h_n]) = 0$ follows for every $p \in \mathbb{N}[l_s, h_s]$, we have $\text{Res}(Q|_\psi) = \emptyset$. Otherwise, it is able to calculate

the only integer $p \in \mathbb{N}[l_s, h_s]$ by a multi-head binary search (introduced in Section 4.4) in $O(\log |Q|)$ time. Then apply the same procedure to the next unfixed position of $[[l_1, h_1], \dots, [p, p], \dots, [l_n, h_n]]$, and continue this process iteratively on other positions as needed. If this process finally obtains $\psi^* = [[l_1^*, h_1^*], \dots, [l_n^*, h_n^*]]$ such that $\text{upp}(\psi^*) = 1$ and $\forall 1 \leq i \leq n, l_i^* = h_i^*$, then (l_1^*, \dots, l_n^*) is the only tuple in $\text{Res}(Q|_{\psi})$. Otherwise, we have $\text{Res}(Q|_{\psi}) = \emptyset$. The time it takes to decide whether $\text{Res}(Q|_{\psi}) = \emptyset$ and calculate ψ^* (if exists) is at most $O(n \log |Q|)$, which is $O(\log |Q|)$ in data complexity. \square

0.6 Proof of Lemma 7.

PROOF. Otherwise, there exists $s < s' \leq n$ such that $[l_{s'}, h_{s'}] \neq [\min_Q(x_{s'}), \max_Q(x_{s'})]$, then for $\forall i \neq s'$, i is not a split position of ψ . Moreover, s' is not a split position since $l_s \neq h_s$, then ψ is not a prefix range filter, which leads to a conflict. \square

0.7 Proof of Lemma 8.

PROOF. For the convenience of the proof, a dummy element $+\infty$ is appended to the end of each sequence, that is, $A_i[|A_i|] = +\infty$ for each $1 \leq i \leq k$. If $F(|A_1|, \dots, |A_k|) \leq T$, the algorithm obviously returns p^* . Then for the case where $F(|A_1|, \dots, |A_k|) > T$, we show that Algorithm 4 maintains the invariant

$$A_i[l_i] \leq p^* \leq A_i[r_i], \quad \forall 1 \leq i \leq k, \quad (6)$$

and then prove that Algorithm 4 correctly returns p^* after at most $O(\log \sum_{i=1}^k |A_i|)$ iterations.

Base case. Initially, for each $1 \leq i \leq k$ we set $l_i = 0$ and $r_i = |A_i|$. Since each A_i is sorted,

$$A_i[0] \leq \min A_i \leq p^* \leq \max A_i \leq A_i[n], \quad (7)$$

then the invariant holds before the first iteration.

Inductive step. Suppose at the start of some iteration we have $A_i[l_i] \leq p^* \leq A_i[r_i]$ for any $1 \leq i \leq k$. We aim to demonstrate that the invariant is preserved after the update performed in the current iteration. Without loss of generality, we assume that for any $1 \leq i \leq k$, $r_i - l_i > 1$ satisfies. (If some i already has $r_i - l_i \leq 1$ and the algorithm has not yet terminated, then by the invariant we have $A_i[l_i] < p^* \leq A_i[r_i]$ and $N_i(A_i[r_i]) \leq N_i(A_i[l_i] + 1) = r_i$, then $N_i(p^*) = r_i = m_i$, and the problem reduces to the remaining $k - 1$ sequences.) In each iteration, the algorithm modifies exactly one of the two bounds: it either updates the lower bound $l_{i_{\min}}$ for some index i_{\min} , or updates the upper bound $r_{i_{\max}}$ for some index i_{\max} . We discuss the two cases:

Case 1: $F(m_1, \dots, m_k) \leq T$.

Let $p = A_{i_{\min}}[m_{i_{\min}}] = \min_{r_i - l_i > 1} A_i[m_i]$. Since

$$N_{i_{\min}}(p) = |\{x \in A_{i_{\min}} \mid x < A_{i_{\min}}[m_{i_{\min}}]\}| \leq m_{i_{\min}}, \quad (8)$$

and for each $i \neq i_{\min}$,

$$N_i(p) \leq N_i(A_i[m_i]) = |\{x \in A_i \mid x < A_i[m_i]\}| \leq m_i, \quad (9)$$

it follows that

$$F(N_1(p), \dots, N_k(p)) \leq F(m_1, \dots, m_k) \leq T. \quad (10)$$

Hence $p^* \geq p$, then after updating $l_{i_{\min}} \leftarrow m_{i_{\min}}$, we still have $A_{i_{\min}}[l_{i_{\min}}] \leq p^* \leq A_{i_{\min}}[r_{i_{\min}}]$. For all $i \neq i_{\min}$, neither l_i nor r_i changes, so the invariant is preserved after the update.

Case 2: $F(m_1, \dots, m_k) > T$.

Let $p = A_{i_{\max}}[m_{i_{\max}}] = \max_{r_i - l_i > 1} A_i[m_i]$. Since

$$N_{i_{\max}}(p + 1) = |\{x \in A_{i_{\max}} \mid x \leq A_{i_{\max}}[m_{i_{\max}}]\}| \geq m_{i_{\max}}, \quad (11)$$

and for each $i \neq i_{\max}$,

$$N_i(p + 1) \geq N_i(A_i[m_i] + 1) = |\{x \in A_i \mid x \leq A_i[m_i]\}| \geq m_i, \quad (12)$$

it follows that

$$F(N_1(p + 1), \dots, N_k(p + 1)) \geq F(m_1, \dots, m_k) > T. \quad (13)$$

Thus $p^* < p + 1$ which implies $p^* \leq p$, then after updating $r_{i_{\max}} \leftarrow m_{i_{\max}}$, we still have $A_{i_{\max}}[l_{i_{\max}}] \leq p^* \leq A_{i_{\max}}[r_{i_{\max}}]$. For all $i \neq i_{\max}$, neither l_i nor r_i changes, so the invariant is preserved after the update.

Convergence and termination. In each iteration, an update is performed on sequence A_i only if $r_i - l_i > 1$, and every such update reduces the length of the interval by at least one. Consequently, the length of each interval $[l_i, r_i]$ shrinks to at most 2 after at most $O(\log |A_i|)$ updates on A_i . Since the number of sequences is a constant, the total number of iterations of the algorithm is bounded by $O(\log \sum_{i=1}^k |A_i|)$.

Now we show that when Algorithm 4 terminates, it returns p^* correctly. If it terminates at Line 4 or Line 13, which implies

$$F(N_1(A_i[l_i] + 1), \dots, N_k(A_i[l_i] + 1)) > T, \quad (14)$$

it follows that $p^* \leq A_i[l_i]$, then by the invariant $p^* \geq A_i[l_i]$, we have $p^* = A_i[l_i]$.

If Algorithm 4 terminates at line 15, which implies $r_i - l_i \leq 1$ and $F(N_1(A_i[l_i] + 1), \dots, N_k(A_i[l_i] + 1)) \leq T$ satisfies for every $1 \leq i \leq k$. Let $p = \min_{1 \leq i \leq k} A_i[r_i]$ and $q = \max_{1 \leq i \leq k} A_i[l_i]$, then

$$\begin{aligned} F(N_1(p), \dots, N_k(p)) &\leq F(N_1(A_1[r_1]), \dots, N_k(A_k[r_k])) \\ &\leq F(N_1(A_1[l_1] + 1), \dots, N_k(A_k[l_k] + 1)) \\ &\leq F(N_1(q + 1), \dots, N_k(q + 1)) \leq T. \end{aligned} \quad (15)$$

Thus $p^* \geq p$, and by the invariant, $p^* \leq p$, we conclude $p^* = p = \min_{1 \leq i \leq k} A_i[r_i]$.

In summary, since i_{\min} , i_{\max} , and F can be computed in $O(1)$ time, and computing each N_i takes at most $O(\log |A_i|)$ time, Algorithm 4 returns p^* in $O(\log \sum_{i=1}^k |A_i|)$ time. \square

0.8 Proof of Lemma 9.

PROOF. Let $\psi = [[l_1, h_1], \dots, [l_n, h_n]]$, since $s = \min\{i \mid l_i \neq h_i\}$ is a split position of ψ , it follows that

- (1) $\forall 1 \leq i < s, l_i = h_i$,
- (2) $\forall s < i \leq n, l_i = \min_Q(x_i)$ and $h_i = \max_Q(x_i)$.

If ψ_{left} is not an empty range, then $l_p \leq \text{mid}_p - 1$, which implies that ψ_{left} is a prefix range filter with split position s . Similarly, if ψ_{left} is not an empty range, then $\text{mid}_p + 1 \leq h_p$, which implies that ψ_{right} is a prefix range filter with split position s . Finally, since $\text{mid}_p \leq \text{mid}_p$, ψ_{mid} is a prefix range filter with split position s . \square

0.9 Proof of Theorem 3.

PROOF. We prove the theorem by analyzing Algorithm 2 under the setting where $N = \text{upp}(\psi_r)$ and $\text{upp}(\psi) = \lfloor \text{AGM}_{c^*}(Q|_{\psi}) \rfloor$ for any range filter ψ . First, the indexes of the relation tables can be built in $O(|Q| \log |Q|)$ time. After the indexes have been built, the computation of c^* and ψ_r takes a constant time. Then, according to Lemma 2, $\varphi_Q^*(i) = \perp$ for each $i > \text{upp}(\psi_r)$. Since

$upp(\psi) = \lfloor AGM_{c^*}(Q|\psi) \rfloor \leq AGM(Q)$, by Lemma 3, Algorithm 3 is a relaxed random-access algorithm running in $O(\log^2 |Q|)$ worst-case time and $O(\log |Q|)$ amortized time. Lastly, since $N = upp(\psi_r)$, by Lemma 1, the expected enumeration delay of Algorithm 2 is $O(\frac{upp(\psi_r)}{|Res(Q)|+1} \log^2 |Q|) = O(\frac{AGM(Q)}{|Res(Q)|+1} \log^2 |Q|)$, and its total running time is at most $O(AGM(Q) \log |Q|)$. \square

0.10 Proof of Lemma 10.

PROOF. Let $\psi = [[l_1, h_1], \dots, [l_n, h_n]]$ be any prefix range filter with split position s . Let I_1, \dots, I_k be a partition of the interval $[l_s, h_s]$. For $\forall 1 \leq i \leq k$, let $\psi_i = [[l_1, h_1], \dots, I_i, \dots, [l_n, h_n]]$, then

$$\sum_{i=1}^k upp(\psi_i) \leq \sum_{i=1}^k upp_{i^*}(\psi_i) \leq upp_{i^*}^*(\psi) = upp^*(\psi), \quad (16)$$

where $i^* = \arg \min_{i=1}^k upp_i(\psi)$. \square

0.11 Proof of Lemma 11.

PROOF. If $\text{att}(Q_s) = \text{att}(Q)$, then $Res(Q|\psi) \subseteq Res(Q_s|\psi)$, which implies $|Res(Q|\psi)| \leq |Res(Q_s|\psi)|$. If $\text{att}(Q_s) \subsetneq \text{att}(Q)$, then

$$\begin{aligned} |Res(Q|\psi)| &= |Res(Q_s|\psi) \bowtie Res(Q_r|\psi)| \\ &\leq \sum_{t \in Res(Q_s|\psi)} |Res((Q_r \bowtie t)|\psi)| \\ &\leq |Res(Q_s|\psi)| \max_{t \in Res(Q_s|\psi)} |Res((Q_r \bowtie t)|\psi)| \quad (17) \\ &\leq |Res(Q_s|\psi)| \cdot |Res(Q_r^*|\psi)| \\ &\leq |Res(Q_s|\psi)| \cdot AGM_c(Q_r^*|\psi). \end{aligned}$$

\square

0.12 Skewness of Datasets.

Our experimental study **includes datasets with skewed distributions**. To quantify the skewness of these datasets, we computed two widely used skewness metrics on their attributes (*i.e.*, on the value-frequency distributions of the columns).

- (1) **Max/Mean Frequency Ratio.** Given an attribute A , let its distinct values be v_1, v_2, \dots, v_k , and let f_i denote the number of occurrences of value v_i . Define

$$MM = \frac{\max_{i=1}^k f_i}{\frac{1}{k} \sum_{i=1}^k f_i}. \quad (18)$$

A larger Max/Mean ratio indicates stronger skew (one or only a few values dominate), whereas a ratio close to 1 suggests a near-uniform distribution.

- (2) **Coefficient of Variation (CV).** Using the same frequencies f_1, \dots, f_k , let $\mu = \frac{1}{k} \sum_{i=1}^k f_i$ and $\sigma = \sqrt{\frac{1}{k} \sum_{i=1}^k (f_i - \mu)^2}$. Then

$$CV = \frac{\sigma}{\mu}. \quad (19)$$

CV measures the relative dispersion of the frequency distribution; a higher CV implies greater skewness.

Table 1 shows the skewness metrics (Max/Mean Frequency Ratio and Coefficient of Variation) computed for each dataset. For the Twitter dataset, the statistics are computed on a 2-million-edge

subgraph uniformly sampled from the original graph, which was used for the Q_Δ experiments.

Table 1: Skewness metrics for key attributes in each dataset.

Dataset	Attribute	MM	CV
TPC-DS+	follow.src/dst	5.38	0.85
	web_sales.ws_bill_customer_sk	2.69	0.52
	web_sales.ws_item_sk	5.40	0.55
Twitter	follow.src/dst	11589.56	24.62

These skewness metrics demonstrate that our experimental study includes a highly skewed real-world dataset, *i.e.*, Twitter, as well as a dataset with mild skew, *i.e.*, TPC-DS+. This highlights that **our evaluation comprehensively covers a wide spectrum of data distributions, ensuring that the proposed methods are tested under both extreme and relatively uniform scenarios.**

0.13 Queries in the Experiments.

Query 1: Q_A on TPC-DS+

```
SELECT *
FROM follow f, web_sales w1, web_sales w2, web_sales w3,
     web_sales w4
WHERE f.src = w1.ws_bill_customer_sk
      AND f.src = w2.ws_bill_customer_sk
      AND f.dst = w3.ws_bill_customer_sk
      AND f.dst = w4.ws_bill_customer_sk
      AND w1.ws_item_sk = w3.ws_item_sk
      AND w2.ws_item_sk = w4.ws_item_sk
```

Query 2: Q_Δ on Twitter

```
SELECT *
FROM follow A, follow B, follow C
WHERE A.dst = B.src AND B.dst = C.src AND C.dst = A.src
```

Query 3: Q_S on Twitter

```
SELECT *
FROM follow A, follow B, follow C, follow D, follow E,
     follow F
WHERE A.dst = B.src AND B.dst = C.src
      AND C.dst = D.src AND D.dst = A.src
      AND E.src = A.src AND E.dst = B.dst
      AND F.dst = A.dst AND F.src = D.src
```

REFERENCES

- [1] Eugene F. Schuster and William R. Sype and. 1987. On the negative hypergeometric distribution. *International Journal of Mathematical Education in Science and Technology* 18, 3 (1987), 453–459. <https://doi.org/10.1080/0020739870180316> arXiv:<https://doi.org/10.1080/0020739870180316>