1. (a) $$I(X,Y) = \sum_y \sum_x p(x,y) \log \frac{p(x,y)}{p(x)p(y)}$$

$$= \sum_y \sum_x p(x,y) \left[\log \frac{p(x,y)}{p(x)} - \log p(y)\right]$$

$$= \sum_y \sum_x p(x,y) \left[\log p(y|x) - \log p(y)\right]$$

$$= \sum_y \sum_x p(x,y) (-\log p(y)) + \sum_y \sum_x p(x,y) \log p(y|x)$$

$$= -\sum_y \log p(y) \left(\sum_x p(x,y)\right) - \left[-\left(\sum_y \sum_x p(y|x) p(x) \log p(y|x)\right)\right]$$

$$= -\sum_y \log p(y) \cdot p(y) - \left[-\sum_x p(x) \sum_y p(y|x) \log p(y|x)\right]$$

$$= H(Y) - \sum_x p(x) H(Y|X=x) = H(Y) - H(Y|X)$$

Since $I(X,Y)$ is symmetric for X, Y, Similarly we can get $I(X,Y) = H(X) - H(X|Y)$

(b) From (a) we know, $I(X,Y) = H(X) - H(X|Y)$, $\Rightarrow$

$I(X,Y) = H(X) + \sum_x p(x) \sum_y p(y|x) \log p(y|x)$. Since $X = f(Y)$, then

$p(y|x) = 0$ or 1, for each case, $\sum_x p(x) \sum_y p(y|x) \log p(y|x) = 0$, then

$I(X,Y) = H(X) + 0 = H(X)$. Similarly, $I(X,Y) = H(Y)$

(c) $$\min_\theta D_{KL}(\hat{p}\|q) = H(\hat{p},q) - H(\hat{p}) \propto H(\hat{p},q)$$

(Noticing only q contains θ here)

$$H(\hat{p},q) = -\sum_x \hat{p}(x) \log q(x|\theta)$$

$$= -\sum_x \frac{1}{N} \sum_{i=1}^N 1_{\{x=x_i\}} \log q(x|\theta)$$

$$= -\frac{1}{N} \sum_{i=1}^N \log(q(x_i|\theta))$$

Thus $$\min_\theta D_{KL}(\hat{p}\|q) \Leftrightarrow \min_\theta -\frac{1}{N} \sum_{i=1}^N \log q(x_i|\theta)$$

$$\Leftrightarrow \max_\theta \sum_{i=1}^N \log q(x_i|\theta)$$

$$\Leftrightarrow \max_\theta \prod_{i=1}^N q(x_i|\theta) \longrightarrow \text{likelyhood}$$

Thus the minimum Kullback-Leibler divergence is obtained by the maximum likelihood.

(d) For continuous variable,

we have $H(x) = -\int p(x) \ln p(x)\, dx$,

From the question, it has three constraint:

$$\text{s.t.} \begin{cases} \int p(x)\, dx = 1 \\ \int x\, p(x)\, dx = \mu \\ \int (x-\mu)^2 p(x)\, dx = \sigma^2 \end{cases}$$

Use Lagrange multipliers to max $H(x)$:

$$L = -\int p(x) \ln p(x) dx + \lambda_1 \left(\int p(x) dx - 1\right) + \lambda_2 \left(\int x\, p(x) dx - \mu\right) + \lambda_3 \left(\int (x-\mu)^2 p(x) dx - \sigma^2\right)$$

$$\frac{df}{dp(x)} = -\ln p(x) - 1 + \lambda_1 + \lambda_2 x + \lambda_3 (x-\mu)^2 = 0$$

$$\Rightarrow p(x) = e^{-1+\lambda_1+\lambda_2 x+\lambda_3 (x-\mu)^2}$$

$$\text{with} \begin{cases} \int e^{-1+\lambda_1+\lambda_2 x+\lambda_3 (x-\mu)^2} dx = 1 \\ \int x e^{-1+\lambda_1+\lambda_2 x+\lambda_3 (x-\mu)^2} dx = \mu \\ \int (x-\mu)^2 e^{-1+\lambda_1+\lambda_2 x+\lambda_3 (x-\mu)^2} dx = \sigma^2 \end{cases} \Rightarrow p(x) = \frac{1}{(2\pi\sigma^2)^{\frac{1}{2}}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$$

$\therefore$ $p = N(\mu, \sigma^2)$ is the solution of max $H(x)$

$\Leftrightarrow$ $H(q) \leq H(p)$ for q be any probability density with mean μ and variance σ^2.

2. (a) $\text{Dirichlet}(p|\alpha) = \dfrac{\Gamma(\sum_1^m \alpha_k)}{\prod_{k=1}^m \Gamma(\alpha_k)} \prod_{k=1}^m p_k^{\alpha_k - 1}$

$$= \exp\left(\log \frac{\Gamma(\sum_1^m \alpha_k)}{\prod_{k=1}^m \Gamma(\alpha_k)} \prod_{k=1}^m p_k^{\alpha_k-1} \right)$$

$$= \exp\left[\log\Gamma(\sum_1^m \alpha_k) + \sum_1^m (\alpha_k - 1)\log p_k - \sum_1^m \log\Gamma(\alpha_k) \right]$$

$$= \exp\left[\sum_1^m (\alpha_k - 1)\log p_k - \left(\sum_1^m \log\Gamma(\alpha_k) - \log\Gamma(\sum_1^m \alpha_k) \right) \right]$$

Thus, set $\eta(\alpha) = \begin{pmatrix} \alpha_1 - 1 \\ \alpha_2 - 1 \\ \vdots \\ \alpha_m - 1 \end{pmatrix}$, $T(p) = \begin{pmatrix} \log p_1 \\ \log p_2 \\ \\ \log p_m \end{pmatrix}$,

$A(\alpha) = \sum_1^m \log\Gamma(\alpha_k) - \log\Gamma(\sum_1^m \alpha_k)$. Then $D(p|\alpha) = \exp[\eta(\alpha)^T T(p) - A(\alpha)]$

(b) The Dirichlet log-likehood function $F(\alpha) = \log p(D|\alpha) = \sum_{j=1}^N \log D(p^j|\alpha)$

$$\therefore F(\alpha) = \sum_{j=1}^N \left[\sum_{k=1}^m (\alpha_k - 1)\log p_k^j + A(\alpha) \right]$$

$$= \sum_{k=1}^m (\alpha_k - 1) \sum_{j=1}^N \log p_k^j + N A(\alpha)$$

$$= N \sum_{k=1}^m (\alpha_k - 1) \frac{1}{N} \sum_{j=1}^N \log p_k^j + N\left[\log\Gamma(\sum_{k=1}^m \alpha_k) - \sum_{k=1}^m \log\Gamma(\alpha_k)\right]$$

$$= N\left[\sum_{k=1}^m (\alpha_k - 1)\hat{t}_k + \log\Gamma(\sum_{k=1}^m \alpha_k) - \sum_{k=1}^m \log\Gamma(\alpha_k)\right], \quad \hat{t}_k = \frac{1}{N}\sum_{j=1}^N \log p_k^j$$

(c) $$\frac{\partial F}{\partial \alpha_k} = N\left(\hat{t}_k + \frac{\partial \log\Gamma(\sum_{k=1}^m \alpha_k)}{\partial \sum_{k=1}^m \alpha_k} \cdot 1 - \frac{\partial \log\Gamma(\alpha_k)}{\partial \alpha_k} \right)$$

$$= N\left(\hat{t}_k + \Psi(\sum_{k=1}^m \alpha_k) - \Psi(\alpha_k) \right), \quad k = 1, \cdots m$$

$$\therefore \frac{\partial F}{\partial \alpha} = \left(\frac{\partial F}{\partial \alpha_1}, \frac{\partial F}{\partial \alpha_2}, \cdots \frac{\partial F}{\partial \alpha_m} \right)$$

(d) $\dfrac{\partial^2 F}{\partial \alpha_k^2} = N\left(\Psi'(\sum_{k=1}^m \alpha_k) - \Psi'(\alpha_k) \right)$, $\dfrac{\partial^2 F}{\partial \alpha_k \alpha_j} = N\Psi'(\sum_{k=1}^m \alpha_k)$

Thus set $q_{kk} = -N\Psi'(\alpha_k)$, $c = N\Psi'(\sum_{k=1}^m \alpha_k)$.

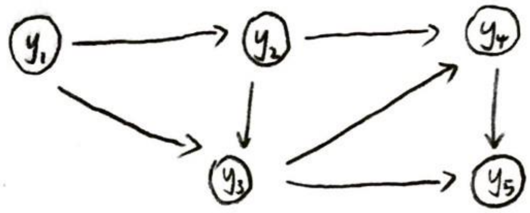
then $\nabla_\alpha^2 F(\alpha) = \left(\dfrac{\partial^2 F}{\partial \alpha_k \alpha_j} \right)_{k,j=1,\cdots m} = Q + c\mathbb{1}\mathbb{1}^T$

(e) $$\alpha^{new} = \alpha^{old} - [H_F(\alpha^{old})]^{-1} \nabla F(\alpha^{old})$$

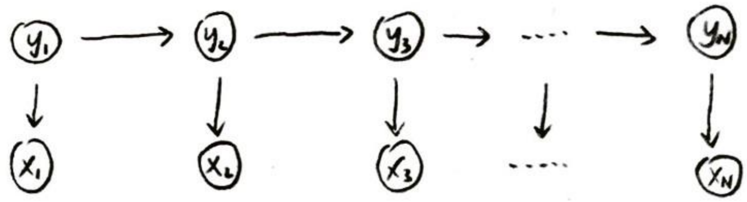
$$= \alpha^{old} - [Q + c\mathbb{1}\mathbb{1}^T]^{-1} \cdot \nabla F(\alpha^{old})$$

$$= \alpha^{old} - \left(Q^{-1} - \frac{Q^{-1} c \mathbb{1}\mathbb{1}^T Q^{-1}}{1 + \mathbb{1}^T Q^{-1} c \mathbb{1}} \right) \nabla F(\alpha^{old})$$

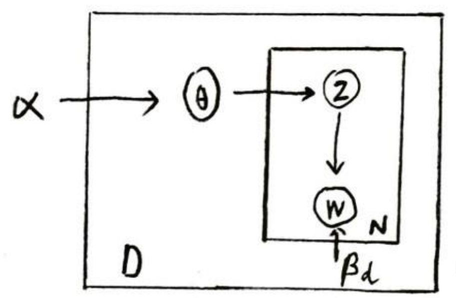
3. (a) (i) $p(y_1, y_2, y_3, y_4, y_4, y_5) = p(y_1)p(y_2)\prod_{k=3}^{5} p(y_k \mid y_{k-1}, y_{k-2})$



(ii) $p(x_1, \cdots x_N, y_1, \cdots y_N) = p(y_1)\prod_{k=2}^{N} p(y_k \mid y_{k-1})\prod_{k=1}^{N} p(x_k \mid y_k)$



(b)



(c) ①: $p(r, \theta, \phi, z) = \prod_{i=1}^{M}\left[p(r_i)\, p(\theta_i \mid r_i)\prod_{j=1}^{N} p(\phi_{ij})\, p(z_{ij} \mid \phi_{ij})\right]$

②: $p(z, w) = \prod_{i=1}^{M}\left[p(z_i)\prod_{j=1}^{N} p(w_{ij} \mid z_i)\right]$

5. (a) $f(y|x;\theta) = \sum_{k=1}^{K} \pi_k \phi(y; w_k^T x + b_k, \sigma_k^2)$

$$\Rightarrow f(y|\underline{x};\theta) = \prod_{i=1}^{N} \sum_{k=1}^{K} \pi_k \phi(y_i; w_k^T x_i + b_k, \sigma_k^2)$$

$$\therefore \log f(\underline{y}|\underline{x};\theta) = \sum_{i=1}^{N} \log \sum_{k=1}^{K} \pi_k \phi(y_i; w_k^T x_i + b_k, \sigma_k^2)$$

(b) $f(y|x, z=k;\theta) = \phi(y; w_k^T x + b_k, \sigma_k^2)$,

$f(y, z|x;\theta) = \pi_z \phi(y; w_z^T x + b_z, \sigma_z^2)$, Let $\Delta_{ik} = 1_{\{z_i=k\}}$

Then $\log f(y, z|x;\theta) = \sum_{k=1}^{K} \sum_{i=1}^{N} 1_{\{z_i=k\}} \log \pi_{z_i} \phi(y_i; w_{z_i}^T x_i + b_{z_i}, \sigma_{z_i}^2)$

$$= \sum_{k=1}^{K} \sum_{i=1}^{N} \Delta_{ik} \log \pi_k \phi(y_i; w_k^T x_i + b_k, \sigma_k^2)$$

(c) $Q(\theta, \theta^{old}) = E_z[\log f(y, z|x;\theta) | y, x; \theta^{old}]$

Let $r_{ik} = p(z_i=k | y_i, x_i; \theta^{old})$

$$r_{ik} = \frac{f(z_i, y_i|x_i;\theta^{old})}{f(y_i|x_i;\theta^{old})} = \frac{\pi_k^{old} \phi(y_i, w_k^{T\,old} x_i + b_k^{old}, \sigma_k^{old})}{\sum_{k=1}^{K} \pi_k^{old} \phi(y_i, w_k^{T\,old} x_i + b_k^{old}, \sigma_k^{old})}$$

$$Q(\theta, \theta^{old}) = \sum_{i=1}^{N} \sum_{k=1}^{K} E_q[1_{\{z_i=k\}} \log \pi_k \phi(y, w_k^T x + b_k, \sigma_k^2)]$$

$$= \sum_{i=1}^{N} \sum_{k=1}^{K} r_{ik} \log \pi_k \phi(y, w_k^T x + b_k, \sigma_k^2)$$

$$= \sum_{i=1}^{N} \sum_{k=1}^{K} r_{ik} \log \pi_k + \sum_{i=1}^{N} \sum_{k=1}^{K} r_{ik} \log \phi(y, w_k^T x + b_k, \sigma_k^2)$$

(d) $\theta^{new} = \arg\max_{\theta} Q(\theta, \theta^{old})$, s.t. $\sum_{k=1}^{K} \pi_k = 1$, $\forall \pi_k > 0$ for $k = 1, \cdots K$

let $f = Q(\theta, \theta^{old}) + \lambda(\sum_{k=1}^{K} \pi_k - 1) - \sum_{k=1}^{K} \alpha_k \pi_k$, $\alpha_k = 0$, $\forall k = 1, \cdots K$

$$\frac{\partial f}{\partial \pi_k} = \sum_{i=1}^{N} \frac{r_{ik}}{\pi_k} + \lambda = 0 \Rightarrow \pi_k = \frac{\sum_{i=1}^{N} r_{ik}}{-\lambda}$$

$$\frac{\partial f}{\partial \lambda} = \sum_{k=1}^{K} \pi_k - 1 = 0 \Rightarrow \sum_{k=1}^{K} \sum_{i=1}^{N} r_{ik} / -\lambda - 1 = 0 \Rightarrow \lambda = -N \Rightarrow \pi_k = \frac{\sum^{N} r_{ik}}{K}$$

$$\frac{\partial f}{\partial w_k} = \sum_{i=1}^{N} r_{ik}^2 \left(\frac{\log \phi(y, w_k^T x + b_k, \sigma_k^2)}{\partial w_k} \right) = \sum_{i=1}^{N} r_{ik} \frac{\partial \left(\log \frac{1}{\sqrt{2\pi}\sigma_k} + \frac{y_i - (w_k^T x_i + b_k))^2}{2\sigma_k} \right)}{\partial w_k}$$

$$= \sum_{i=1}^{N} r_{ik} \frac{2(y_i - w_k^T x_i + b_k)}{2\sigma_k} (-x_i) = 0$$

$$\frac{\partial f}{\partial b_k} = \sum_{i=1}^{N} r_{ik} \frac{\partial \left(\log \frac{1}{\sqrt{2\pi}\sigma_k} + \frac{(y_i - (w_k^T x_i + b_k))^2}{2\sigma_k} \right)}{\partial b_k} = \sum_{i=1}^{N} r_{ik} \frac{y_i - (w_k^T x_i + b_k)}{\sigma_k} \cdot (-1) = 0$$

Set $\tilde{w}_k = \begin{bmatrix} w_k \\ b_k \end{bmatrix}$, $\tilde{x} = \begin{bmatrix} x_1 & 1 \\ \vdots & \vdots \\ x_N & 1 \end{bmatrix} = [x, 1]$, $y = \begin{bmatrix} y_1 \\ \vdots \\ y_N \end{bmatrix}$

then we can combine $\frac{\partial f}{\partial w_k}$ & $\frac{\partial f}{\partial b_k}$ to get $\sum_{i=1}^{N} r_{ik} (y_i - \tilde{w}_k^T \tilde{x}_i) \tilde{x}_i = 0$

Use closed form to represent the whole question, we can set the objective function as

$\sum_{i=1}^{N} r_{ik}(y_i - \tilde{w}_k^T \tilde{x}_i)^2$ can be written as $(\tilde{X}\tilde{w}_k - y)^T R(\tilde{X}\tilde{w}_k - y)$

where $R = \mathrm{diag}(r_{1k}, r_{2k}, \cdots r_{Nk})$

$$\therefore \sum_{i=1}^{N} r_{ik}(y_i - \tilde{w}_k^T \hat{x}_i)\tilde{x}_i = 2X^T R X w - 2X^T R y = 0$$

$$\Rightarrow \tilde{w}_k = (X^T R X)^{-1} X^T R y$$

$$\frac{\partial f}{\partial \sigma_k^2} = \frac{\partial \sum_{i=1}^{N} r_{ik}\left(\log \frac{1}{\sqrt{2\pi\sigma_k^2}} + \left(-\frac{1}{2\sigma_k^2}(y_i - \tilde{w}_k^T \tilde{x}_i)^2\right)\right)}{\partial \sigma_k^2}$$

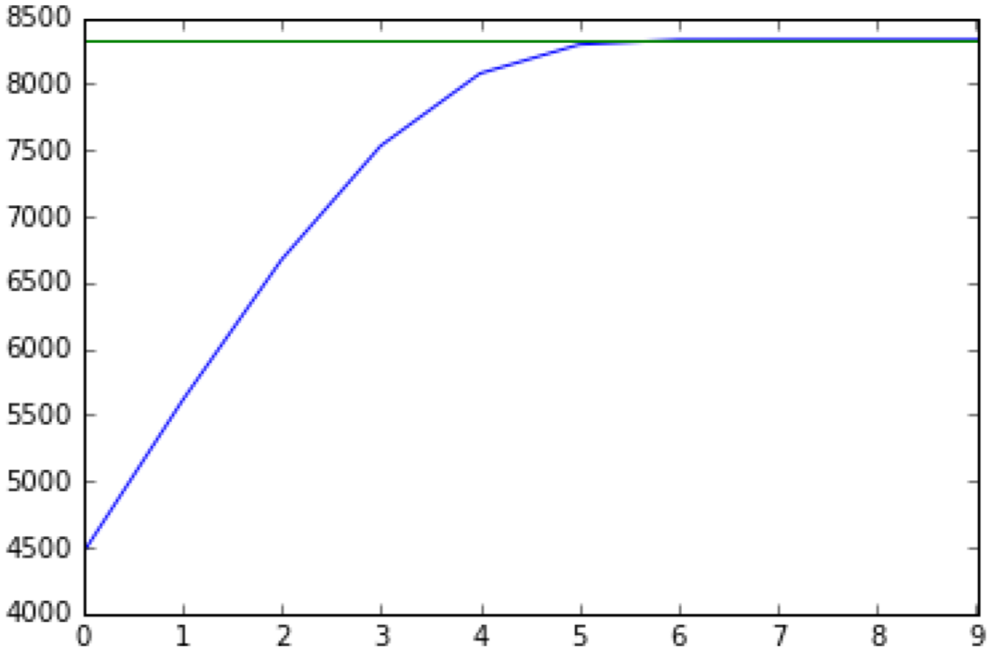
$$= \sum_{i=1}^{N} r_{ik} + \frac{1}{\sigma_k^2}\sum_{i=1}^{N}(y_i - \hat{w}_k^T \tilde{x}_i)^2 r_{ik} = 0$$

$$\Rightarrow \sigma_k^2 = \frac{1}{\sum_{i=1}^{N} r_{ik}} \sum_{i=1}^{N}(y_i - \tilde{w}_k^T \tilde{x}_i)^2 r_{ik}$$

$$= \frac{1}{\sum_{i=1}^{N} r_{ik}} (\tilde{X}\tilde{w}_k - y)^T R(\tilde{X}\tilde{w}_k - y)$$

2.(f)

The plot of the log-likelihood vs iteration is as follows:

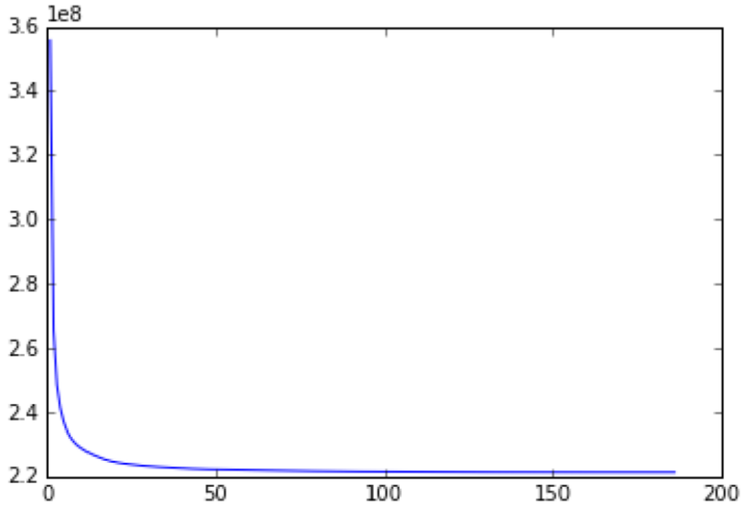


The estimated model parameters are:

9.96979256, 4.90985978, 14.86165346, 19.66495057, 48.97431779

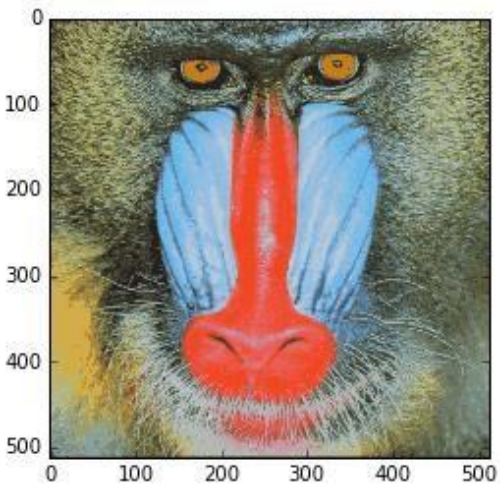
4.(a)

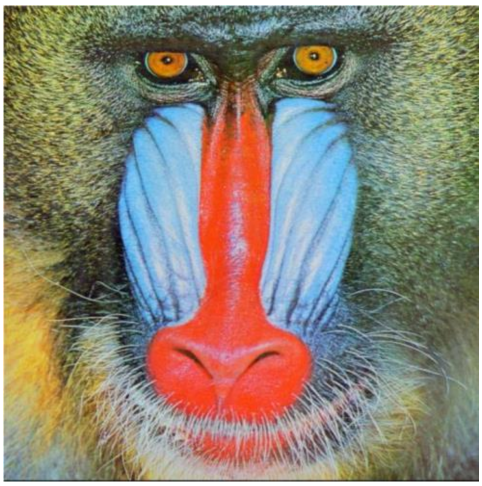
The plot of objective function is



Put the two pictures together:

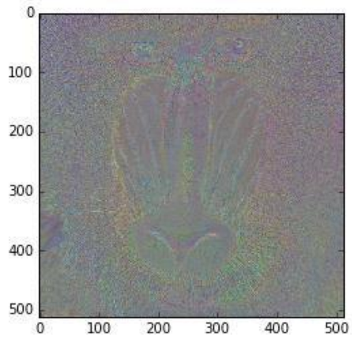
Compressed **Original**





Comparing the compressed picture with the original one, we can find the large parts with single color are best preserved. For the border parts where the color changes from one to another, they are generally not preserved well.

The picture of the difference:



The compression ratio is :

$$\mathrm{r} = \frac{24 * 2^2 * 64 + 216^2 * log_2 64}{24 * 512^2} = 6.3477\%$$

The relative mean absolute error of the compresses image is: 0.05

4.(b)

The number of 24 bits per pixel needed:

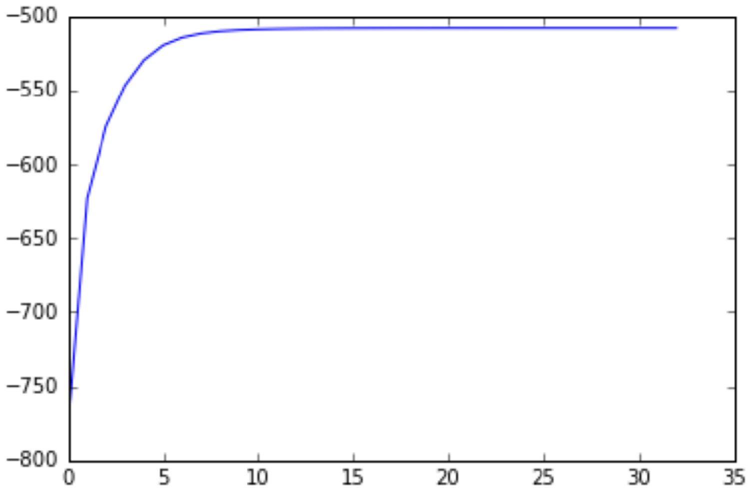
$$\mathrm{bpp} = \frac{24 * M^2 * K + \left(\frac{N}{M}\right)^2 * log_2 K}{N^2}$$

the compress ratio is:

$$\mathrm{ratio} = \mathrm{bpp}/24 = \frac{24 * M^2 * K + \left(\frac{N}{M}\right)^2 * log_2 K}{24N^2}$$

5.(e)

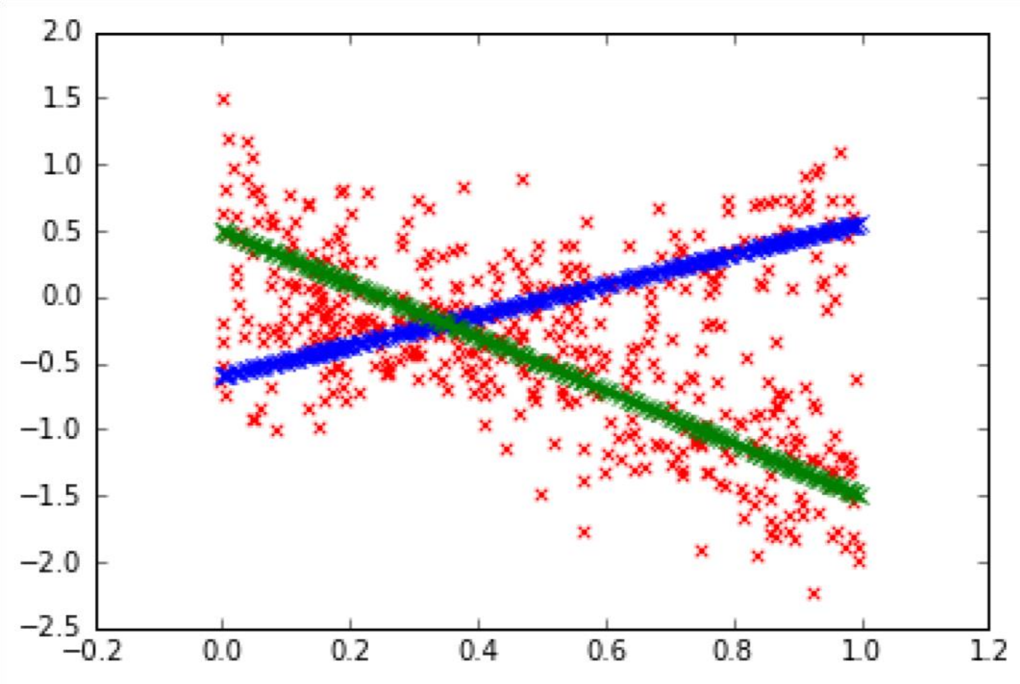
The plot of the log-likelihood is :



The estimated model parameters:

$$\pi = (0.293, 0.707), \quad (w, b) = \begin{pmatrix} 0.960 & -0.474 \\ -2.063 & 0.492 \end{pmatrix}, \qquad \sigma^2 = (0.407,\ 0.336)$$

The plot showing the data and estimated lines is :



Codes:

```
#2.f import numpy as np from matplotlib import
pyplot as plt from scipy.special import
gammaln, polygamma
from __future__ import division
N = 1000 m
= 5
alpha = np.array([10, 5, 15, 20, 50]) P
= np.random.dirichlet(alpha, N)
t=np.sum(np.log(P),axis=0)/N
t=t.reshape((5,1))
U=np.ones((5,1))
alpha0=np.ones((5,1))
loglihd=np.zeros((1000,1)) for
i in range(1000):
    dalpha1=N*(t+polygamma(0,np.sum(alpha0))*U-polygamma(0,alpha0));
g1=-N*polygamma(1,alpha0)     Q=np.diag(np.diag(np.ones((5,5))*g1))
    C=N*polygamma(1,np.sum(alpha0))
    Q1=np.linalg.inv(Q)     alphan=alpha0-(Q1-
(Q1.dot(C*U.dot(U.T)).dot(Q1))/(1+C*U.T.dot(Q1).dot(U))).dot(dalpha1)
loglihd[i]=N*((alphan-
```

```
1).T.dot(t)+gammaln(np.sum(alphan))np.sum(gammaln(alphan)))     if
i==0:
        alpha0=alphan     elif i > 0 and np.abs(loglihd[i]-
loglihd[i-1])>0.0001:
        alpha0=alphan     else:
        alphafnl=alphan        break
lstd=N*((alpha-1).T.dot(t)+gammaln(np.sum(alpha))-np.sum(gammaln(alpha)))
x=np.linspace(0,9,10);y=loglihd[0:10];stdl=lstd*np.ones((10,1))
plt.plot(x,y);plt.plot(x,stdl)
```

```
#4.a import numpy as np from
matplotlib import pyplot as plt from
__future__ import division from
scipy.ndimage import imread
mandrill = imread('mandrill.png', mode='RGB').astype(float)
N = int(mandrill.shape[0])
M = 2;k = 64
X = np.zeros((N**2//M**2, 3*M**2))
for i in range(N//M):     for j in
range(N//M):
        X[i*N//M+j,:] = mandrill[i*M:(i+1)*M,j*M:(j+1)*M,:].reshape(3*M**2)
Jf=np.zeros((300,1))
#calculating  tagfet  value  J
def calcJ(data,centers):
    diffsq=(centers[:,np.newaxis,:]-data)**2
    return np.sum(np.min(np.sum(diffsq,axis=2),axis=0))
#implement k means  def
kmeans(data,k):
#initializing centers and list J
    centers=data[np.random.choice(range(data.shape[0]),k,replace=False),:]
    J=[];
#closest center for each sample
for itera in range(300):
        sqdistances=np.sum((centers[:,np.newaxis,:]-data)**2,axis=2)
closest=np.argmin(sqdistances,axis=0)
#calculate J and append to list
        J.append(calcJ(data,centers))
        Jf[itera]=calcJ(data,centers)
#update clusters        for i in range(k):
centers[i,:]=data[closest==i,:].mean(axis=0)
```

```
#decide whether stopping        if itera>0 and
np.abs(Jf[itera]-Jf[itera-1])==0:
X=centers[closest,:]
          break        else:
          continue
   J.append(calcJ(data,centers))
   return J,centers,closest


JF,centersf,closestf=kmeans(X,64)
Xnew=centersf[closestf,:]
mandrillnew=np.zeros(512*512*3)
mandrillnew=mandrillnew.reshape(512,512,3)
for i in range(256):     for j in range(256):
       mandrillnew[i*2:(i+1)*2,j*2:(j+1)*2,:]=Xnew[i*256+j,:].reshape(2,2,3)
plt.imshow(mandrillnew/255)
plt.show()
mandrillgrey=mandrillnew-mandrill+128*np.ones(512*512*3).reshape(512,512,3)
plt.imshow(mandrillgrey/255)
plt.show() Jf1=Jf[Jf>0]
x1=np.linspace(1,186,186)
plt.plot(x1,Jf1)
error=np.sum(np.abs(mandrillnew-mandrill))/(3*255*512**2)



#5.e
from __future__ import division import
numpy as np
from matplotlib import pyplot as plt from
scipy.stats import norm as norm
# Generate the data according to the specification in the homework description


N = 500
x =
  np.random.rand(N)


pi0 = np.array([0.7, 0.3])
w0 = np.array([-2, 1]) b0
= np.array([0.5, -0.5])
sigma0 = np.array([.4, .3])


y = np.zeros_like(x) for
  i in range(N):
```

```
    k = 0 if np.random.rand() < pi0[0] else 1
    y[i] = w0[k]*x[i] + b0[k] + np.random.randn()*sigma0[k]

ccpiest = np.array([0.5, 0.5]);cwest = np.array([1, -1]) cbest =
np.array([0, 0]);sigmaest = np.array([np.std(y), np.std(y)]) cwest2
= np.array([cwest,cbest]).T;x2 = np.array([x,np.ones(N)]) p1 =
ccpiest[0]*norm(cwest2[0].dot(x2),sigmaest[0]).pdf(y) p2 =
ccpiest[1]*norm(cwest2[1].dot(x2),sigmaest[1]).pdf(y) r1 =
p1/(p1+p2);r2 = p2/(p1+p2);r = np.array([r1,r2])
Q1 = np.sum(r1*np.log(ccpiest[0]*norm(cwest2[0].dot(x2),sigmaest[0]).pdf(y)))
Q2 = np.sum(r2*np.log(ccpiest[1]*norm(cwest2[1].dot(x2),sigmaest[1]).pdf(y)))
Q = Q1+Q2;diff = 1;ll = [];ite = [];count = 0

for i in range(100):
    ll.append(Q);ite.append(count);tmp = Q     ccpiest
= np.array([np.mean(r1),np.mean(r2)])
    w1 = np.linalg.inv(x2.dot(np.diag(r1)).dot(x2.T)).dot(x2).dot(np.diag(r1)).dot(y)
w2 = np.linalg.inv(x2.dot(np.diag(r2)).dot(x2.T)).dot(x2).dot(np.diag(r2)).dot(y)
cwest2 = np.array([w1,w2])
    sigmaest1 = ((np.sum(r1*(y-cwest2[0].dot(x2))**2))/np.sum(r1))**0.5
sigmaest2 = ((np.sum(r2*(y-cwest2[1].dot(x2))**2))/np.sum(r2))**0.5     sigmaest =
np.array([sigmaest1,sigmaest2])
    p1 = ccpiest[0]*norm(cwest2[0].dot(x2),sigmaest[0]).pdf(y)     p2
= ccpiest[1]*norm(cwest2[1].dot(x2),sigmaest[1]).pdf(y)     r1 =
p1/(p1+p2);r2 = p2/(p1+p2);r = np.array([r1,r2])
    Q1 = np.sum(r1*np.log(ccpiest[0]*norm(cwest2[0].dot(x2),sigmaest[0]).pdf(y)))
    Q2 = np.sum(r2*np.log(ccpiest[1]*norm(cwest2[1].dot(x2),sigmaest[1]).pdf(y)))
Q = Q1+Q2;diff = Q-tmp;count = count+1     if diff>=1e-4:        continue     else:
        break
plt.plot(x, cwest2[0].dot(x2), c='b', marker='x') plt.plot(x,
cwest2[1].dot(x2), c='g', marker='x')
plt.scatter(x, y, c='r', marker='x') plt.plot(ite,ll)
```