STAT 576 Bayesian Analysis

Lecture 13: Nonparametric Models

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In this lecture, we will discuss nonparametric models, where the prior is placed directly on the function μ within a function space.

- ► A **stochastic process** is a collection of random variables indexed by some set, e.g. time or space.
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- ► A **Gaussian process** is a stochastic process such that any finite collection of r.v.s. has a multivariate normal distribution.
- ▶ Specifically, if $\{\mu(x): x \in \mathcal{X}\}$ is a Gaussian process, then for any finite set of indices $x_1, \ldots, x_n \in \mathcal{X}$, the random vector $(\mu(x_1), \ldots, \mu(x_n))$ has a multivariate normal distribution.
- lacktriangle As a special case, for any $x\in\mathcal{X}$, $\mu(x)$ is a normal random variable.

- A Gaussian process is completely specified by its mean function $m(x) = E[\mu(x)]$ and covariance function $k(x, x') = \text{Cov}(\mu(x), \mu(x'))$.
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- ▶ The joint distribution of $\mu(x_1), \ldots, \mu(x_n)$ is given by

$$\begin{pmatrix} \mu(x_1) \\ \vdots \\ \mu(x_n) \end{pmatrix} \sim \mathcal{N} \left(\begin{pmatrix} m(x_1) \\ \vdots \\ m(x_n) \end{pmatrix}, K(x_1, \dots, x_n) \right).$$

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- Consistent defintion: the distribution of $\mu(x_1), \ldots, \mu(x_m)$ derived from the joint distribution of $\mu(x_1), \ldots, \mu(x_n)$ is the same for any choice of x_{m+1}, \ldots, x_n .
- $ightharpoonup K(x_1,\ldots,x_n)$ is positive definite for any choice of x_1,\ldots,x_n .

Gaussian Process — Covariance

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WLOG, we assume $\tau^2 = l^2 = 1$. To show K is positive definite, we need to show that for any vector $u = (u_1, \dots, u_n)$, $u^T K u \ge 0$.

Notice that

$$k(x_i, x_j) = \exp\left(-\frac{1}{2}|x_i - x_j|^2\right) = \mathbb{E}\left[e^{i|x_i - x_j|Z}\right]$$

for $Z \sim \mathcal{N}(0,1)$.

► Therefore,

$$u^T K u = \sum_{i,j} u_i u_j k(x_i, x_j) = \sum_{i,j} u_i u_j \mathbb{E}\left[e^{i|x_i - x_j|Z}\right] = \mathbb{E}\left[\left(\sum_i u_i e^{ix_i Z}\right)^2\right] \ge 0.$$

Gaussian Process — Basis Functions

The Gaussian Process can also be constructed by basis functions:

$$\mu(x) = \sum_{h=1}^{H} \beta_h b_h(x), \quad \beta = (\beta_1, \dots, \beta_H) \sim \mathcal{N}(\beta_0, \Sigma_\beta)$$

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Then μ is a Gaussian process with

$$m(x) = \boldsymbol{b}(x)^T \boldsymbol{\beta}_0, \quad k(x, x') = \boldsymbol{b}(x)^T \boldsymbol{\Sigma}_{\beta} \boldsymbol{b}(x').$$

Suppose, we have observed the data $(x_1,y_1),\ldots,(x_n,y_n)$, and we assume that

$$y_i = \mu(x_i) + \epsilon_i, \quad \epsilon_i \sim \mathcal{N}(0, \sigma^2).$$

Given a new point \tilde{x} , we want to estimate the expected value of y at \tilde{x} , i.e. $\mu(\tilde{x})$.

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We assume that μ is a Gaussian process with mean function m(x)=0 and covariance function $k(x,x')=\tau^2\exp\left(-|x-x'|^2/(2l^2)\right)$.

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- ▶ The joint distribution of $y = (y_1, ..., y_n)$ and $\mu(\tilde{x})$ is given by

$$\begin{pmatrix} y \\ \mu(\tilde{x}) \end{pmatrix} \sim \mathcal{N} \left(\begin{pmatrix} 0 \\ 0 \end{pmatrix}, \begin{pmatrix} K(x,x) + \sigma^2 I & K(x,\tilde{x}) \\ K(\tilde{x},x) & K(\tilde{x},\tilde{x}) \end{pmatrix} \right).$$

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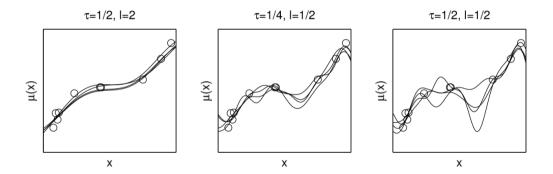
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With the properties of conditional distribution of multivariate normal, we can derive the posterior distribution of $\mu(\tilde{x})$ given y.

$$\mu(\tilde{x}) \mid x, y, \tau^2, \sigma^2, l^2$$

$$\sim \mathcal{N}\left(K(\tilde{x}, x)(K(x, x) + \sigma^2 I)^{-1} y, K(\tilde{x}, \tilde{x}) - K(\tilde{x}, x)(K(x, x) + \sigma^2 I)^{-1} K(x, \tilde{x})\right)$$

Gaussian Process — Example



For a Bayesian procedure, we need to specify the prior distributions for the hyperparameters τ^2 , σ^2 , and l^2 .

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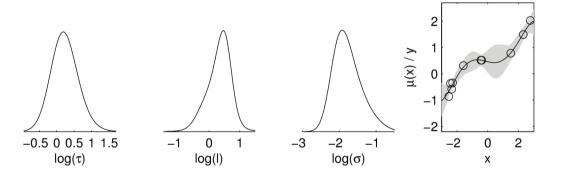
A common choice is

$$p(\log \tau) \propto 1$$
, $p(\log \sigma) \propto 1$, $p(\log l) \propto 1$.

The log-likelihood is

$$\log p(y \mid x, \tau^2, \sigma^2, l^2) = -\frac{1}{2} y^T (K(x, x) + \sigma^2 I)^{-1} y - \frac{1}{2} \log |K(x, x) + \sigma^2 I| - \frac{n}{2} \log(2\pi).$$

The posterior is now straightforward.



In this example, we analyze the patterns of birthdays in the United States. The data is the number of births on each day of the year from 1969 to 1988.

- ▶ This is a time series data, where the index is the number of days from 1969-01-01.
- ▶ The series contains periodic patterns, e.g. yearly and weekly patterns.
- The series also contains long term trends.

We model the time series as an additive model:

$$y(t) = f_1(t) + f_2(t) + f_3(t) + f_4(t) + f_5(t) + \epsilon_t,$$

Long-term trend:

$$f_1(t) \sim \mathcal{GP}(0, k_1), \quad k_1(t, t') = \sigma_1^2 \exp\left(-\frac{|t - t'|^2}{2l_1^2}\right)$$

► Short-term trend:

$$f_2(t) \sim \mathcal{GP}(0, k_2), \quad k_2(t, t') = \sigma_2^2 \exp\left(-\frac{|t - t'|^2}{2l_2^2}\right)$$

We model the time series as an additive model:

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► Weekly pattern:

$$f_3(t) \sim \mathcal{GP}(0, k_3), \quad k_3(t, t') = \sigma_3^2 \exp\left(-\frac{2\sin^2(\pi(t - t')/7)}{l_{3,1}^2}\right) \exp\left(-\frac{|t - t'|^2}{2l_{3,2}^2}\right)$$

Yearly pattern:

$$f_4(t) \sim \mathcal{GP}(0, k_4), \quad k_4(t, t') = \sigma_4^2 \exp\left(-\frac{2\sin^2(\pi(t - t')/365.25)}{l_{4,1}^2}\right) \exp\left(-\frac{|t - t'|^2}{2l_{4,2}^2}\right)$$

We model the time series as an additive model:

$$y(t) = f_1(t) + f_2(t) + f_3(t) + f_4(t) + f_5(t) + \epsilon_t,$$

Special days and its interaction with weekends:

$$f_5(t) = I_{s.d.}(t)\beta_a + I_{s.d.}(t)I_{w.e.}(t)\beta_b$$

where $I_{s.d.}(t)$ is an indicator function for special days (13 holidays), and $I_{w.e.}(t)$ is an indicator function for weekends.

 $ightharpoonup \epsilon_t \sim \mathcal{N}(0, \sigma^2)$ is the residual.

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- ▶ log-unifom prior for other parameters.
- ► The model can be further extended by considering weekdays v.s. weekends. See textbook Ch. 21.2.

