

STAT 574 HOMEWORK 1 SOLUTION

1 (Mixed Models). Use the Weight v.s. Height example. For the following cases, write down the mixed-effect models, specify the unknown parameters, and determine the number of parameters to be fitted.

- (a) The weight is linear in height, but the intercepts differ for different families.
- (b) The weight is linear in height, but both the intercepts and the coefficients for height differ for different families.
- (c) In addition to (b), we know the expected coefficient for height is 5.30.
- (d) The weight is linear in height, but the intercepts differ for different families, and the height coefficients differ for different genders.

Solution.

- (a) The model is

$$W_{ij} = \beta_0 + \beta_1 H_{ij} + b_i + \epsilon_{ij},$$

where $b_i \sim \mathcal{N}(0, \sigma^2 d)$ and $\epsilon_{ij} \sim \mathcal{N}(0, \sigma^2)$. The unknown parameters are β_0 , β_1 , σ^2 and d . The number of parameters is, therefore, 4.

- (b) The model is

$$W_{ij} = \beta_0 + \beta_1 H_{ij} + b_{0i} + b_{1i} H_{ij} + \epsilon_{ij},$$

where $(b_{0i}, b_{1i})^T \sim \mathcal{N}(\mathbf{0}, \sigma^2 \mathbf{D})$ and $\epsilon_{ij} \sim \mathcal{N}(0, \sigma^2)$. The unknown parameters are β_0 , β_1 , σ^2 and \mathbf{D} , totaling 6 parameters (because $\mathbf{D} \in \mathbb{R}^{2 \times 2}$ has to be symmetric).

- (c) The model is

$$W_{ij} = \beta_0 + 5.3 H_{ij} + b_{0i} + b_{1i} H_{ij} + \epsilon_{ij}$$

or equivalently,

$$W_{ij} - 5.3 H_{ij} = \beta_0 + b_{0i} + b_{1i} H_{ij} + \epsilon_{ij},$$

where $(b_{0i}, b_{1i})^T \sim \mathcal{N}(\mathbf{0}, \sigma^2 \mathbf{D})$ and $\epsilon_{ij} \sim \mathcal{N}(0, \sigma^2)$. The unknown parameters are β_0 , σ^2 and \mathbf{D} , totaling 5 parameters.

- (d) The model is

$$W_{ijk} = \beta_0 + \beta_1 H_{ijk} + b_{0i} + b_{1j} H_{ijk} + \epsilon_{ijk},$$

where index (i, j, k) means the k -th person in i -th family and j -th gender, and $b_{0i} \sim \mathcal{N}(0, \sigma^2 d_0)$, $b_{1j} \sim \mathcal{N}(0, \sigma^2 d_1)$ and $\epsilon_{ijk} \sim \mathcal{N}(0, \sigma^2)$. The unknown parameters are β_0 , β_1 , σ^2 , d_0 and d_1 , totaling 5 parameters.

2 (Generalized Least Squares). Consider a linear model with dependent observations that

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \boldsymbol{\epsilon},$$

where $\boldsymbol{\epsilon} \sim \mathcal{N}(\mathbf{0}, \boldsymbol{\Sigma})$.

- (a) Write down the log-likelihood function and calculate the MLE $\hat{\boldsymbol{\beta}}^{(MLE)}$.
- (b) Consider the following weighted residual sum-of-squares.

$$\text{RSS}(w_1, \dots, w_n) = \sum_{i=1}^n w_i (y_i - \boldsymbol{\beta}^T \mathbf{x}_i)^2,$$

where \mathbf{x}_i is the covariates for unit i , y_i is the corresponding observation, and $w_i > 0$ is the weight assigned to unit i .

Find $\hat{\boldsymbol{\beta}}^{(WLS)}$ that minimizes $\text{RSS}(w_1, \dots, w_n)$.

- (c) Figure out when you can find the weights w_1, \dots, w_n such that $\hat{\boldsymbol{\beta}}^{(MLE)} = \hat{\boldsymbol{\beta}}^{(WLS)}$.

Solution.

- (a) The log-likelihood function is

$$\ell(\boldsymbol{\beta}) = -\frac{1}{2} [\log |\boldsymbol{\Sigma}| + (\mathbf{y} - \mathbf{X}\boldsymbol{\beta})^T \boldsymbol{\Sigma}^{-1} (\mathbf{y} - \mathbf{X}\boldsymbol{\beta})]$$

By maximizing the log-likelihood function, we have

$$\hat{\boldsymbol{\beta}}^{(MLE)} = (\mathbf{X}^T \boldsymbol{\Sigma}^{-1} \mathbf{X})^{-1} \mathbf{X}^T \boldsymbol{\Sigma}^{-1} \mathbf{y}$$

- (b) Taking the partial derivative of RSS, we have

$$\frac{\partial \text{RSS}}{\partial \boldsymbol{\beta}} = 2 \sum_{i=1}^n w_i \boldsymbol{\beta}^T \mathbf{x}_i \mathbf{x}_i^T - 2 \sum_{i=1}^n w_i y_i \mathbf{x}_i^T$$

By setting the partial derivative to zero, we have

$$\hat{\boldsymbol{\beta}}^{(WLS)} = \left(\sum_{i=1}^n w_i \mathbf{x}_i \mathbf{x}_i^T \right)^{-1} \sum_{i=1}^n w_i y_i \mathbf{x}_i$$

- (c) Let $\mathbf{W} = \text{diag}(w_1, w_2, \dots, w_n)$. We notice that

$$\hat{\boldsymbol{\beta}}^{(WLS)} = (\mathbf{X}^T \mathbf{W} \mathbf{X})^{-1} \mathbf{X}^T \mathbf{W} \mathbf{y}.$$

Therefore, when $\boldsymbol{\Sigma}^{-1} = \mathbf{W}$, that is when $\boldsymbol{\Sigma}$ is diagonal and $w_i \propto 1/\Sigma_{ii}$, where Σ_{ii} is the i -th diagonal of $\boldsymbol{\Sigma}$, MLE and WLS are equal.

3 (Restricted Maximum Likelihood). Consider n i.i.d. observations $x_1, \dots, x_n \sim \mathcal{N}(\mu, \sigma^2)$.

- (a) Find the MLE $\hat{\mu}$ and $\hat{\sigma}^2$.
(b) Construct $(n-1)$ new observations in the following way

$$y_1 = x_1 - x_n, y_2 = x_2 - x_n, \dots, y_{n-1} = x_{n-1} - x_n.$$

Show that (y_1, \dots, y_{n-1}) is independent of μ .

- (c) Determine the distribution of (y_1, \dots, y_{n-1}) .
(d) Use above as the restricted likelihood for $\hat{\sigma}^2$, and find out $\hat{\sigma}_R^2$ that maximizes the restricted likelihood.
(e) Compare $\hat{\sigma}^2$ and $\hat{\sigma}_R^2$.

Solution.

- (a) The log-likelihood function is

$$\ell(\mu, \sigma^2) = -\frac{1}{2} \left[n \log \sigma^2 + \sigma^{-2} \sum_{i=1}^n (x_i - \mu)^2 \right].$$

The MLE is therefore

$$\hat{\mu} = \frac{1}{n} \sum_{i=1}^n x_i = \bar{x}, \quad \hat{\sigma}^2 = \frac{1}{n} \sum_{i=1}^n (x_i - \hat{\mu})^2.$$

- (b) $(y_1, \dots, y_{n-1}, \hat{\mu})$ is a linear combination of (x_1, \dots, x_n) , and is therefore multivariate Gaussian. It suffices to show the covariance is zero. Since

$$\text{Cov}(y_i, \bar{x}) = \text{Cov}(x_i - x_n, x_i/n + x_n/n) = \text{Var}(x_i)/n - \text{Var}(x_n)/n = 0,$$

(y_1, \dots, y_{n-1}) and $\bar{\mu}$ are independent.

- (c) Since (y_1, \dots, y_{n-1}) is multivariate Gaussian, it suffices to determine its mean and variance. Notice that

$$\mathbb{E}[y_i] = \mathbb{E}[x_i] - \mathbb{E}[x_n] = 0$$

$$\text{Var}(y_i) = \text{Var}(x_i) + \text{Var}(x_n) = 2\sigma^2$$

$$\text{Cov}(y_i, y_j) = \text{Var}(x_n) = \sigma^2, \quad \text{for } i \neq j.$$

Hence, we have

$$\begin{bmatrix} y_1 \\ y_2 \\ \vdots \\ y_{n-1} \end{bmatrix} \sim \mathcal{N}(\mathbf{0}, \sigma^2 \mathbf{\Sigma}),$$

where

$$\mathbf{\Sigma} = \begin{bmatrix} 2 & 1 & \cdots & 1 \\ 1 & 2 & \cdots & 1 \\ \vdots & \vdots & \ddots & \vdots \\ 1 & 1 & \cdots & 2 \end{bmatrix} \in \mathbb{R}^{(n-1) \times (n-1)}$$

(d) Notice that, $\Sigma = \mathbb{I} + \mathbf{1}\mathbf{1}^T$, and by Woodbury identity, we have

$$\Sigma^{-1} = \mathbb{I} - \mathbf{1}(\mathbf{1} + \mathbf{1}^T \mathbf{1})^{-1} \mathbf{1}^T = \mathbb{I} - n^{-1} \mathbf{1}\mathbf{1}^T.$$

Hence, the restricted log-likelihood function is

$$\ell_R(\sigma^2) = -\frac{1}{2} [\log |\sigma^2 \mathbf{\Sigma}| + \mathbf{y}^T (\sigma^2 \mathbf{\Sigma})^{-1} \mathbf{y}]$$

The maximum is attained at

$$\hat{\sigma}_R^2 = \frac{1}{n-1} \mathbf{y}^T \Sigma^{-1} \mathbf{y} = \frac{1}{n-1} \sum_{i=1}^n x_i^2 - \frac{1}{n(n-1)} \left(\sum_{i=1}^n x_i \right)^2 = \frac{n}{n-1} \hat{\sigma}^2.$$

(e) The restricted MLE is proportional to the MLE.

4 (Programming). Download the Weight v.s. Height dataset in **Family.txt** from

<https://github.com/eugenedemidenko/mixedmodels>

The file is in **Data/MixedModels/Chapter02/** folder. For the four models in Question 1, fit with the dataset and report your final models (specifying the parameters and the distributions).

Solution.

(a) The model is

$$W_{ij} = -206.83 + 5.35H_{ij} + b_i + \epsilon_{ij},$$

where $b_i \sim \mathcal{N}(0, 14.07^2)$ and $\epsilon_{ij} \sim \mathcal{N}(0, 24.71^2)$.

(b) The model is

$$W_{ij} = -248.83 + 6.01H_{ij} + b_{0i} + b_{1i}H_{ij} + \epsilon_{ij},$$

where

$$\begin{pmatrix} b_{0i} \\ b_{1i} \end{pmatrix} \sim \mathcal{N}\left(\mathbf{0}, \begin{bmatrix} 237.07^2 & -0.998 \times 237.07 \times 3.63 \\ -0.998 \times 237.07 \times 3.63 & 3.63^2 \end{bmatrix}\right)$$

and $\epsilon_{ij} \sim \mathcal{N}(0, 19.81^2)$.

(c) The model is

$$\begin{pmatrix} b_{0i} \\ b_{1i} \end{pmatrix} \sim \mathcal{N}\left(\mathbf{0}, \begin{bmatrix} 223.16^2 & -0.998 \times 223.16 \times 3.42 \\ -0.998 \times 223.16 \times 3.42 & 3.42^2 \end{bmatrix}\right)$$

and $\epsilon_{ij} \sim \mathcal{N}(0, 20.04^2)$.

(d) The model is

$$W_{ijk} = -74.52 + 3.39H_{ijk} + b_{0i} + b_{1j}H_{ijk} + \epsilon_{ijk},$$

where $b_{0i} \sim \mathcal{N}(0, 13.12^2)$, $b_{1j} \sim \mathcal{N}(0, 0.22^2)$ and $\epsilon_{ijk} \sim \mathcal{N}(0, 24.07^2)$.

The code for this problem is pasted below:

```
1 data = read.table("./Data/MixedModels/Chapter02/Family.txt", header=T, stringsAsFactors=F)
2
3 library(nlme)
4 fit.a = lme(fixed=Weight~Height, random=~1|FamilyID, data=data)
5 print(fit.a)
6
7 fit.b = lme(fixed=Weight~Height, random=~1+Height|FamilyID, data=data)
8 print(fit.b)
9
10 data$ynew = data$Weight - 5.3 * data$Height
11 fit.c = lme(fixed=ynew~1, random=~1+Height|FamilyID, data=data)
12 print(fit.c)
13
14 library(lme4)
15 fit.d = lmer(Weight~1+Height+(1|FamilyID)+(0+Height|Sex), data=data)
16 print(fit.d)
```