## Midterm review

Dr. Jarad Niemi

Iowa State University

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## Chapters

- Probability and inference (Ch 1)
- Single-parameter models (Ch 2)

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- Introduction to multiparameter models (Ch 3)

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- Stan

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  - Inference via simulations

### General

- Priors
  - Conjugate (Sec 2.4)
  - Default Jeffreys (Sec 2.8)
  - Weakly informative (Sec 2.9)
- Posteriors
  - Compromise between data and prior (2.2)
  - Point estimation
  - Credible intervals (Sec 2.3)

## Specific models

- Binomial (Sec 2.1–2.4)
- Normal, unknown mean (Sec 2.5)
- Normal, unknown variance (Sec 2.6)
- Poisson (Sec 2.6)
- Exponential (Sec 2.6)
- Poisson with exposure (Sec 2.7)

### Additional comments:

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• Posterior decomposition, e.g.

$$p(\theta_1,\ldots,\theta_n|y)=p(\theta_1|y)\prod_{i=2}^n p(\theta_i|\theta_{1:i-1},y)$$

where  $1: i-1=1, 2, \ldots, i-1$ .

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• Conditional independence, e.g.

$$p(\theta_i|\theta_{1:i-1}, y) = p(\theta_i|\theta_{i-1}, y)$$

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where 
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Limiting distribution:

$$\theta|y \stackrel{d}{\to} N\left(\hat{\theta}, \frac{1}{n}I(\hat{\theta})^{-1}\right)$$

- Not unique to Bayesian statistics
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  - Number of parameters increase with sample size
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• Hierarchical normal model (Sec 5.4)

$$y_{ij} \stackrel{\text{iid}}{\sim} N(\mu_j, \sigma_j^2) \quad \mu_j \stackrel{\text{iid}}{\sim} N(\eta, \tau^2) \quad \sigma_j^2 \stackrel{\text{iid}}{\sim} Ga(\alpha, \beta)$$

From a Bayesian perspective,

Simple: 
$$H_i: \theta = \theta_i$$
 Composite:  $H_i: \theta \in (\theta_i, \theta_{i+1}]$ 

Treat all simple (or all composite) hypotheses as formal Bayesian parameter estimation. Treat a mix of simple and composite hypotheses as formal Bayesian tests.

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- the marginal likelihood,  $p(y|H_i)$ .

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• Data replications

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- Graphical posterior predictive checks (Sec 6.4)
- Posterior predictive pvalues (Sec 6.3)

$$p_B = P(T(y^{rep}, \theta) \ge T(y, \theta)|y)$$

for a test statistic  $T(y, \theta)$ .

#### Stan

```
model = "
data {
  int<lower=0> N;
  int<lower=0> n[N];
  int<lower=0> y[N];
  real s;
parameters {
  real<lower=0,upper=1> mu;
  real<lower=0> eta;
transformed parameters {
  real<lower=0> alpha;
  real<lower=0> beta;
  alpha <- eta * mu;
  beta <- eta * (1-mu);
model {
      ~ beta(20,30);
  eta ~ lognormal(0,s);
      ~ beta_binomial(n,alpha,beta);
generated quantities {
  real<lower=0,upper=1> theta[N];
  for (i in 1:N) theta[i] <- beta_rng(alpha+y[i], beta+n[i]-y[i]);</pre>
```