I09 - Comparing means

STAT 587 (Engineering) Iowa State University

October 4, 2020

Consider the model $Y_i \overset{ind}{\sim} N(\mu, \sigma^2)$.

Consider the model $Y_i \overset{ind}{\sim} N(\mu, \sigma^2)$. We have discussed a number of statistical procedures to draw inferences about μ :

• Frequentist: based on distribution of $\frac{\overline{Y} - \mu}{s/\sqrt{n}}$

- Frequentist: based on distribution of $\frac{\overline{Y} \mu}{s/\sqrt{n}}$
 - ullet p-value for a hypothesis test, e.g. $H_0: \mu = \mu_0$,

- Frequentist: based on distribution of $\frac{\overline{Y} \mu}{s/\sqrt{n}}$
 - $\bullet \ p\text{-value}$ for a hypothesis test, e.g. $H_0: \mu = \mu_0$,
 - ullet confidence interval for μ ,

- ullet Frequentist: based on distribution of $rac{\overline{Y}-\mu}{s/\sqrt{n}}$
 - $\bullet \ p\text{-value}$ for a hypothesis test, e.g. $H_0: \mu = \mu_0$,
 - ullet confidence interval for μ ,
- \bullet Bayesian: based on posterior for μ
 - ullet credible interval for μ ,

- \bullet Frequentist: based on distribution of $\frac{\overline{Y}-\mu}{s/\sqrt{n}}$
 - ullet p-value for a hypothesis test, e.g. $H_0: \mu=\mu_0$,
 - ullet confidence interval for μ ,
- ullet Bayesian: based on posterior for μ
 - ullet credible interval for μ ,
 - posterior model probability, e.g. $p(H_0|y)$,

- \bullet Frequentist: based on distribution of $\frac{\overline{Y}-\mu}{s/\sqrt{n}}$
 - ullet p-value for a hypothesis test, e.g. $H_0: \mu = \mu_0$,
 - ullet confidence interval for μ ,
- ullet Bayesian: based on posterior for μ
 - ullet credible interval for μ ,
 - ullet posterior model probability, e.g. $p(H_0|y)$, and
 - posterior probabilities, e.g. $P(\mu < \mu_0|y)$.

Consider the model $Y_i \overset{ind}{\sim} N(\mu, \sigma^2)$. We have discussed a number of statistical procedures to draw inferences about μ :

- \bullet Frequentist: based on distribution of $\frac{\overline{Y} \mu}{s/\sqrt{n}}$
 - ullet p-value for a hypothesis test, e.g. $H_0: \mu = \mu_0$,
 - ullet confidence interval for μ ,
- ullet Bayesian: based on posterior for μ
 - ullet credible interval for μ ,
 - ullet posterior model probability, e.g. $p(H_0|y)$, and
 - posterior probabilities, e.g. $P(\mu < \mu_0|y)$.

Now, we will consider what happens when you have multiple μs .

Consider the model

$$Y_{g,i} \stackrel{ind}{\sim} N(\mu_g, \sigma_g^2)$$

for g=1,2 and $i=1,\ldots,n_g$.

Consider the model

$$Y_{g,i} \stackrel{ind}{\sim} N(\mu_g, \sigma_g^2)$$

for g=1,2 and $i=1,\ldots,n_g$. and you are interested in the relationship between μ_1 and μ_2 .

Consider the model

$$Y_{g,i} \stackrel{ind}{\sim} N(\mu_g, \sigma_g^2)$$

for g=1,2 and $i=1,\ldots,n_g$. and you are interested in the relationship between μ_1 and μ_2 .

$$\frac{\overline{Y}_1 - \overline{Y}_2 - (\mu_1 - \mu_2)}{\sqrt{\frac{S_1^2}{n_1} + \frac{S_2^2}{n_2}}}$$

Consider the model

$$Y_{g,i} \stackrel{ind}{\sim} N(\mu_g, \sigma_g^2)$$

for g=1,2 and $i=1,\ldots,n_g$. and you are interested in the relationship between μ_1 and μ_2 .

• Frequentist: based on distribution of

$$\frac{\overline{Y}_1 - \overline{Y}_2 - (\mu_1 - \mu_2)}{\sqrt{\frac{S_1^2}{n_1} + \frac{S_2^2}{n_2}}}$$

• p-value for a hypothesis test, e.g. $H_0: \mu_1 = \mu_2$,

Consider the model

$$Y_{g,i} \stackrel{ind}{\sim} N(\mu_g, \sigma_g^2)$$

for g=1,2 and $i=1,\ldots,n_g$. and you are interested in the relationship between μ_1 and μ_2 .

$$\frac{\overline{Y}_1 - \overline{Y}_2 - (\mu_1 - \mu_2)}{\sqrt{\frac{S_1^2}{n_1} + \frac{S_2^2}{n_2}}}$$

- p-value for a hypothesis test, e.g. $H_0: \mu_1 = \mu_2$,
- confidence interval for $\mu_1 \mu_2$,

Consider the model

$$Y_{g,i} \stackrel{ind}{\sim} N(\mu_g, \sigma_g^2)$$

for g=1,2 and $i=1,\ldots,n_g$. and you are interested in the relationship between μ_1 and μ_2 .

$$\frac{\overline{Y}_1 - \overline{Y}_2 - (\mu_1 - \mu_2)}{\sqrt{\frac{S_1^2}{n_1} + \frac{S_2^2}{n_2}}}$$

- p-value for a hypothesis test, e.g. $H_0: \mu_1 = \mu_2$,
- confidence interval for $\mu_1 \mu_2$,
- Bayesian: posterior for μ_1, μ_2 , i.e. $p(\mu_1, \mu_2|y)$
 - credible interval for $\mu_1 \mu_2$,

Consider the model

$$Y_{g,i} \stackrel{ind}{\sim} N(\mu_g, \sigma_g^2)$$

for g=1,2 and $i=1,\ldots,n_g$. and you are interested in the relationship between μ_1 and μ_2 .

$$\frac{\overline{Y}_1 - \overline{Y}_2 - (\mu_1 - \mu_2)}{\sqrt{\frac{S_1^2}{n_1} + \frac{S_2^2}{n_2}}}$$

- p-value for a hypothesis test, e.g. $H_0: \mu_1 = \mu_2$,
- confidence interval for $\mu_1 \mu_2$,
- Bayesian: posterior for μ_1, μ_2 , i.e. $p(\mu_1, \mu_2|y)$
 - credible interval for $\mu_1 \mu_2$,
 - posterior model probability, e.g. $p(H_0|y)$,

Consider the model

$$Y_{g,i} \stackrel{ind}{\sim} N(\mu_g, \sigma_g^2)$$

for g=1,2 and $i=1,\ldots,n_g$. and you are interested in the relationship between μ_1 and μ_2 .

$$\frac{\overline{Y}_1 - \overline{Y}_2 - (\mu_1 - \mu_2)}{\sqrt{\frac{S_1^2}{n_1} + \frac{S_2^2}{n_2}}}$$

- p-value for a hypothesis test, e.g. $H_0: \mu_1 = \mu_2$,
- confidence interval for $\mu_1 \mu_2$,
- Bayesian: posterior for μ_1, μ_2 , i.e. $p(\mu_1, \mu_2|y)$
 - credible interval for $\mu_1 \mu_2$,
 - posterior model probability, e.g. $p(H_0|y)$, and
 - probability statements, e.g. $P(\mu_1 < \mu_2 | y)$.

Consider the model

$$Y_{g,i} \stackrel{ind}{\sim} N(\mu_g, \sigma_g^2)$$

for g=1,2 and $i=1,\ldots,n_g$. and you are interested in the relationship between μ_1 and μ_2 .

• Frequentist: based on distribution of

$$\frac{\overline{Y}_1 - \overline{Y}_2 - (\mu_1 - \mu_2)}{\sqrt{\frac{S_1^2}{n_1} + \frac{S_2^2}{n_2}}}$$

- p-value for a hypothesis test, e.g. $H_0: \mu_1 = \mu_2$,
- confidence interval for $\mu_1 \mu_2$,
- ullet Bayesian: posterior for μ_1,μ_2 , i.e. $p(\mu_1,\mu_2|y)$
 - credible interval for $\mu_1 \mu_2$,
 - ullet posterior model probability, e.g. $p(H_0|y)$, and
 - probability statements, e.g. $P(\mu_1 < \mu_2 | y)$.

where $y = (y_{1,1}, \dots, y_{1,n_1}, y_{2,1}, \dots, y_{2,n_2}).$

Data example

Suppose you have two manufacturing processes to produce sensors and you are interested in the average sensitivity of the sensors.

Data example

Suppose you have two manufacturing processes to produce sensors and you are interested in the average sensitivity of the sensors.

So you run the two processes and record the sensitivity of each sensor in units of mV/V/mm Hg (http://www.ni.com/white-paper/14860/en/).

Data example

Suppose you have two manufacturing processes to produce sensors and you are interested in the average sensitivity of the sensors.

So you run the two processes and record the sensitivity of each sensor in units of mV/V/mm Hg (http://www.ni.com/white-paper/14860/en/). And you have the following summary statistics:

```
# A tibble: 2 x 4
process n mean sd
<chr> (-int) <dbl> (-dbl> 1
P1 22 7.74 1.87
P2 P2 34 9.24 2.26
```

Because there is no indication that you have any expectation regarding the sensitivities of process 1 compared to process 2, we will conduct a two-sided two-sample t-test assuming the variances are not equal.

Because there is no indication that you have any expectation regarding the sensitivities of process 1 compared to process 2, we will conduct a two-sided two-sample t-test assuming the variances are not equal, i.e.

$$Y_{g,i} \stackrel{ind}{\sim} N(\mu_g, \sigma_g^2)$$

Because there is no indication that you have any expectation regarding the sensitivities of process 1 compared to process 2, we will conduct a two-sided two-sample t-test assuming the variances are not equal, i.e.

$$Y_{g,i} \stackrel{ind}{\sim} N(\mu_g, \sigma_g^2)$$

and

$$H_0: \mu_1=\mu_2$$
 and $H_A: \mu_1
eq \mu_2$

Because there is no indication that you have any expectation regarding the sensitivities of process 1 compared to process 2, we will conduct a two-sided two-sample t-test assuming the variances are not equal, i.e.

$$Y_{g,i} \stackrel{ind}{\sim} N(\mu_g, \sigma_g^2)$$

and

$$H_0: \mu_1 = \mu_2$$
 and $H_A: \mu_1 \neq \mu_2$

```
t.test(sensitivity ~ process, data = d2)
Welch Two Sample t-test

data: sensitivity by process
t = -2.6932, df = 50.649, p-value = 0.009571
alternative hypothesis: true difference in means is not equal to 0
95 percent confidence interval:
    -2.610398 -0.380530
sample estimates:
mean in group P1 mean in group P2
    7.743761    9.239224
```

Posterior for μ_1, μ_2

Assume

$$Y_{g,i} \stackrel{ind}{\sim} N(\mu_g, \sigma_g^2)$$
 and $p(\mu_1, \mu_2, \sigma_1^2, \sigma_2^2) \propto \frac{1}{\sigma_1^2} \frac{1}{\sigma_2^2}$.

Then

$$\mu_g|y \stackrel{ind}{\sim} t_{n_g-1}(\overline{y}_g, s_g^2/n_g)$$

Posterior for μ_1, μ_2

Assume

$$Y_{g,i} \stackrel{ind}{\sim} N(\mu_g, \sigma_g^2)$$
 and $p(\mu_1, \mu_2, \sigma_1^2, \sigma_2^2) \propto \frac{1}{\sigma_1^2} \frac{1}{\sigma_2^2}$.

Then

$$\mu_g|y \stackrel{ind}{\sim} t_{n_g-1}(\overline{y}_g, s_g^2/n_g)$$

and a draw for μ_g can be obtained by taking

$$\overline{y}_g + T_{n_g-1} s_g / \sqrt{n_g}, \quad T_{n_g-1} \stackrel{ind}{\sim} t_{n_g-1}(0,1).$$

Posterior for μ_1, μ_2

Assume

$$Y_{g,i} \stackrel{ind}{\sim} N(\mu_g, \sigma_g^2)$$
 and $p(\mu_1, \mu_2, \sigma_1^2, \sigma_2^2) \propto \frac{1}{\sigma_1^2} \frac{1}{\sigma_2^2}$.

Then

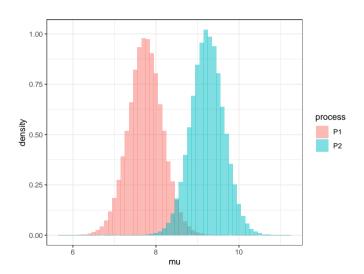
$$\mu_g|y \stackrel{ind}{\sim} t_{n_g-1}(\overline{y}_g, s_g^2/n_g)$$

and a draw for μ_g can be obtained by taking

$$\overline{y}_g + T_{n_g-1} s_g / \sqrt{n_g}, \quad T_{n_g-1} \stackrel{ind}{\sim} t_{n_g-1}(0,1).$$

Simulations:

We can use these draws to compare the posteriors



Credible interval for the difference

To obtain statistical inference on the difference, we use the samples and take the difference

```
d3 <- sims %>%
 spread(process, mu) %>%
 mutate(diff = P1-P2)
# Bayes estimate for the difference
mean(d3$diff)
[1] -1.493267
# Estimated 95% equal-tail credible interval
quantile(d3$diff, c(.025,.975))
      2.5%
                97.5%
-2 6339752 -0 3483025
# Estimate of the probability that mul is larger than mu2
mean(d3\$diff > 0)
Γ1] 0.00591
```

Now, let's consider the more general problem of

$$Y_{g,i} \stackrel{ind}{\sim} N(\mu_g, \sigma_g^2)$$

for
$$g=1,2,\dots,G$$
 and $i=1,\dots,n_g$

Now, let's consider the more general problem of

$$Y_{g,i} \stackrel{ind}{\sim} N(\mu_g, \sigma_g^2)$$

for $g=1,2,\ldots,G$ and $i=1,\ldots,n_g$ and you are interested in the relationship amongst the μ_g .

Now, let's consider the more general problem of

$$Y_{g,i} \stackrel{ind}{\sim} N(\mu_g, \sigma_g^2)$$

for $g=1,2,\ldots,G$ and $i=1,\ldots,n_g$ and you are interested in the relationship amongst the μ_g

We can perform the following statistical procedures:

- Frequentist:
 - p-value for test of $H_0: \mu_g = \mu$ for all g,

Now, let's consider the more general problem of

$$Y_{g,i} \stackrel{ind}{\sim} N(\mu_g, \sigma_g^2)$$

for $g=1,2,\ldots,G$ and $i=1,\ldots,n_g$ and you are interested in the relationship amongst the μ_g

We can perform the following statistical procedures:

- Frequentist:
 - ullet p-value for test of $H_0: \mu_g = \mu$ for all g,
 - ullet confidence interval for $\mu_g \mu_{g'}$,

Now, let's consider the more general problem of

$$Y_{g,i} \stackrel{ind}{\sim} N(\mu_g, \sigma_g^2)$$

for $g=1,2,\ldots,G$ and $i=1,\ldots,n_g$ and you are interested in the relationship amongst the μ_g

We can perform the following statistical procedures:

- Frequentist:
 - p-value for test of $H_0: \mu_g = \mu$ for all g,
 - confidence interval for $\mu_g \mu_{g'}$,
- Bayesian: based on posterior for μ_1, \ldots, μ_G
 - ullet credible interval for $\mu_q \mu_{q'}$,

Three or more means

Now, let's consider the more general problem of

$$Y_{g,i} \stackrel{ind}{\sim} N(\mu_g, \sigma_g^2)$$

for $g=1,2,\ldots,G$ and $i=1,\ldots,n_g$ and you are interested in the relationship amongst the μ_g

We can perform the following statistical procedures:

- Frequentist:
 - p-value for test of $H_0: \mu_g = \mu$ for all g,
 - confidence interval for $\mu_g \mu_{g'}$,
- Bayesian: based on posterior for μ_1, \ldots, μ_G
 - ullet credible interval for $\mu_g \mu_{g'}$,
 - posterior model probability, e.g. $p(H_0|y)$,

Three or more means

Now, let's consider the more general problem of

$$Y_{g,i} \stackrel{ind}{\sim} N(\mu_g, \sigma_g^2)$$

for $g=1,2,\ldots,G$ and $i=1,\ldots,n_g$ and you are interested in the relationship amongst the μ_g

We can perform the following statistical procedures:

- Frequentist:
 - p-value for test of $H_0: \mu_g = \mu$ for all g,
 - confidence interval for $\mu_a \mu_{a'}$,
 - Bayesian: based on posterior for μ_1, \ldots, μ_G
 - ullet credible interval for $\mu_g \mu_{g'}$,
 - ullet posterior model probability, e.g. $p(H_0|y)$, and
 - probability statements, e.g. $P(\mu_g < \mu_{g'}|y)$

where g and g' are two different groups.

Data example

Suppose you have three manufacturing processes to produce sensors and you are interested in the average sensitivity of the sensors.

Data example

Suppose you have three manufacturing processes to produce sensors and you are interested in the average sensitivity of the sensors.

So you run the three processes and record the sensitivity of each sensor in units of mV/V/mm Hg (http://www.ni.com/white-paper/14860/en/).

Data example

Suppose you have three manufacturing processes to produce sensors and you are interested in the average sensitivity of the sensors.

So you run the three processes and record the sensitivity of each sensor in units of mV/V/mm Hg (http://www.ni.com/white-paper/14860/en/). And you have the following summary statistics:

```
# A tibble: 3 x 4

process n mean sd

<chr> <int> <int> <ib/> <dbl> <dbl>

1 P1 22 7.74 1.87

2 P2 34 9.24 2.26

3 P3 7 10.8 1.96
```

p-values

When there are lots of means, the first null hypothesis is typically

$$H_0: \mu_g = \mu \,\forall \, g$$

p-values

When there are lots of means, the first null hypothesis is typically

$$H_0: \mu_g = \mu \,\forall \, g$$

```
oneway.test(sensitivity ~ process, data = d)
```

One-way analysis of means (not assuming equal variances)

```
data: sensitivity and process
```

F = 7.6287, num df = 2.000, denom df = 17.418, p-value = 0.004174

Pairwise differences

Then we typically look at pairwise differences:

When

$$Y_{g,i} \stackrel{ind}{\sim} N(\mu_g, \sigma_g^2),$$

When

$$Y_{g,i} \stackrel{ind}{\sim} N(\mu_g, \sigma_g^2),$$

we have

$$\mu_g|y \stackrel{ind}{\sim} t_{n_g-1}(\overline{y}_g, s_g^2/n_g)$$

When

$$Y_{g,i} \stackrel{ind}{\sim} N(\mu_g, \sigma_g^2),$$

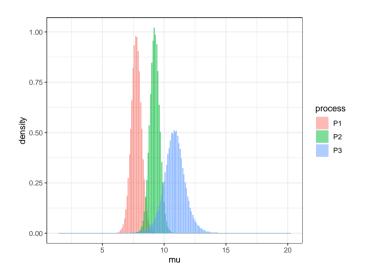
we have

$$\mu_g|y \stackrel{ind}{\sim} t_{n_g-1}(\overline{y}_g, s_g^2/n_g)$$

and that a draw for μ_g can be obtained by taking

$$\overline{y}_g + T_{n_g-1} s_g / \sqrt{n_g}, \quad T_{n_g-1} \stackrel{ind}{\sim} t_{n_g-1}(0,1).$$

Compare posteriors



Credible intervals for differences

Use the simulations to calculate posterior probabilities and credible intervals for differences.

```
# Estimate of the probability that one mean is larger than another
sims %>%
 spread(process, mu) %>%
 mutate(mu1-mu2) = P1-P2,
           mu1-mu3 = P1-P3.
           mu2-mu3 = P2-P3) \%>\%
 select(`mu1-mu2`, `mu1-mu3`, `mu2-mu3`) %>%
 gather(comparison, diff) %>%
 group by(comparison) %>%
 summarize(probability = mean(diff>0) %>% round(4).
            lower = quantile(diff, .025) %>% round(2),
           upper = quantile(diff, .975) %>% round(2)) %>%
 mutate(credible interval = paste("(".lower.".".upper.")". sep="")) %>%
 select(comparison, probability, credible_interval)
# A tibble: 3 v 3
 comparison probability credible interval
 <chr>>
               <dbl> <chr>
                 0.0059 (-2.63,-0.35)
1 mii1-mii2
2 m111-m113
                 0.0037 (-5.06.-1.11)
3 m112-m113
                 0.0493 (-3.56,0.37)
```

In the model

$$Y_{g,i} \stackrel{ind}{\sim} N(\mu_g, \sigma_g^2)$$

we can calculate a p-value for the following null hypothesis:

$$H_0:\sigma_g=\sigma\quad ext{for all}\quad g$$

In the model

$$Y_{g,i} \stackrel{ind}{\sim} N(\mu_g, \sigma_g^2)$$

we can calculate a $\emph{p}\text{-value}$ for the following null hypothesis:

$$H_0: \sigma_g = \sigma \quad \text{for all} \quad g$$

bartlett.test(sensitivity ~ process, data = d)

Bartlett test of homogeneity of variances

data: sensitivity by process

Bartlett's K-squared = 0.90949, df = 2, p-value = 0.6346

In the model

$$Y_{g,i} \stackrel{ind}{\sim} N(\mu_g, \sigma_g^2)$$

we can calculate a p-value for the following null hypothesis:

$$H_0:\sigma_g=\sigma\quad ext{for all}\quad g$$

```
bartlett.test(sensitivity ~ process, data = d)
Bartlett test of homogeneity of variances
data: sensitivity by process
Bartlett's K-squared = 0.90949, df = 2, p-value = 0.6346
```

This may give us reason to proceed as if the variances is the same in all groups, i.e.

$$Y_{g,i} \stackrel{ind}{\sim} N(\mu_g, \sigma^2).$$

In the model

$$Y_{g,i} \stackrel{ind}{\sim} N(\mu_g, \sigma_g^2)$$

we can calculate a p-value for the following null hypothesis:

$$H_0:\sigma_g=\sigma\quad ext{for all}\quad g$$

```
bartlett.test(sensitivity ~ process, data = d)
Bartlett test of homogeneity of variances

data: sensitivity by process
Bartlett's K-squared = 0.90949, df = 2, p-value = 0.6346
```

This may give us reason to proceed as if the variances is the same in all groups, i.e.

$$Y_{g,i} \stackrel{ind}{\sim} N(\mu_g, \sigma^2).$$

This assumption is common when the number of observations in the groups is small.

Comparing means when the variances are equal

Assuming $Y_{g,i} \stackrel{ind}{\sim} N(\mu_g, \sigma^2)$, we can test

$$H_0: \mu_g = \mu \,\forall \, g$$

Comparing means when the variances are equal

Assuming $Y_{g,i} \overset{ind}{\sim} N(\mu_g, \sigma^2)$, we can test

$$H_0: \mu_g = \mu \,\forall \, g$$

oneway.test(sensitivity ~ process, data = d, var.equal = TRUE)

One-way analysis of means

data: sensitivity and process

F = 6.7543, num df = 2, denom df = 60, p-value = 0.002261

Comparing means when the variances are equal

Assuming $Y_{g,i} \stackrel{ind}{\sim} N(\mu_g, \sigma^2)$, we can test

F = 6.7543, num df = 2, denom df = 60, p-value = 0.002261

$$H_0: \mu_g = \mu \,\forall \, g$$

```
oneway.test(sensitivity ~ process, data = d, var.equal = TRUE)
One-way analysis of means
data: sensitivity and process
```

Then we typically look at pairwise differences, i.e. $H_0: \mu_q = \mu_{q'}$.

```
pairwise.t.test(d$sensitivity, d$process, p.adjust.method = "none")
Pairwise comparisons using t tests with pooled SD
data: d$sensitivity and d$process
    P1    P2
P2 0.0116 -
P3 0.0012 0.0720
```

If $Y_{g,i} \stackrel{ind}{\sim} N(\mu_g, \sigma^2)$ and we use the prior $p(\mu_1, \dots, \mu_G, \sigma^2) \propto 1/\sigma^2$,

If $Y_{g,i} \stackrel{ind}{\sim} N(\mu_g, \sigma^2)$ and we use the prior $p(\mu_1, \dots, \mu_G, \sigma^2) \propto 1/\sigma^2$, then

$$\mu_g|y,\sigma^2 \stackrel{ind}{\sim} N(\overline{y}_g,\sigma^2/n_g) \quad \sigma^2|y \sim IG\left(\frac{n-G}{2},\frac{1}{2}\sum_{g=1}^G\sum_{i=1}^{n_g}(y_{g,i}-\overline{y}_g)^2\right)$$

where
$$n = \sum_{g=1}^{G} n_g$$
.

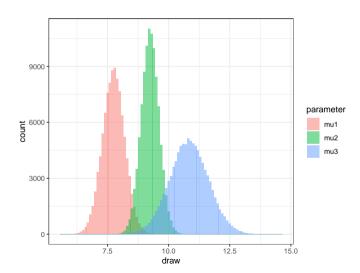
If $Y_{g,i} \stackrel{ind}{\sim} N(\mu_g, \sigma^2)$ and we use the prior $p(\mu_1, \dots, \mu_G, \sigma^2) \propto 1/\sigma^2$, then

$$\mu_g|y,\sigma^2 \stackrel{ind}{\sim} N(\overline{y}_g,\sigma^2/n_g) \quad \sigma^2|y \sim IG\left(\frac{n-G}{2},\frac{1}{2}\sum_{g=1}^G\sum_{i=1}^{n_g}(y_{g,i}-\overline{y}_g)^2\right)$$

where $n=\sum_{g=1}^G n_g$ and thus, we obtain joint samples for μ by performing the following

- 1. $\sigma^{2(m)} \sim p(\sigma^2|y)$
- 2. For g = 1, ..., G, $\mu_g \sim p(\mu_g | y, \sigma^{2(m)})$.

Compare posteriors



Credible interval for the differences

To compare the means, we compare the samples drawn from the posterior.

```
sims %>%
 mutate(`mu1-mu2` = mu1-mu2,
        mu1-mu3 = mu1-mu3.
        mu2-mu3 = mu2-mu3) \%>\%
 select('mu1-mu2', 'mu1-mu3', 'mu2-mu3') %>%
 gather(comparison, diff) %>%
 group_by(comparison) %>%
 summarize(probability = mean(diff>0) %>% round(4).
           lower = quantile(diff, .025) %>% round(2),
           upper = quantile(diff, .975) %>% round(2)) %>%
 mutate(credible_interval = paste("(",lower,",",upper,")", sep="")) %>%
 select(comparison, probability, credible_interval)
# A tibble: 3 x 3
 comparison probability credible_interval
 <chr>
           <dbl> <chr>
1 mu1-mu2 0.0059 (-2.65,-0.35)
          0.0007 (-4.92,-1.26)
2 mu1-mu3
3 mu2-mu3
                0.036 (-3.34,0.15)
```

Summary

Multiple (independent) normal means

Summary

Multiple (independent) normal means

- p-values
- confidence intervals
- posterior densities
- credible intervals
- posterior probabilities