# Set 13 - pvalues

STAT 401 (Engineering) - Iowa State University

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## Statistical hypothesis testing

#### Definition

A (classical) hypothesis test consists of two hypotheses:

- null hypothesis  $(H_0)$  and
- an alternative hypothesis  $(H_A)$

which make a claim about parameters in a model and a decision to either

- reject the null hypothesis or
- fail to reject the null hypothesis.

We reject the null hypothesis if our pvalue is less than a pre-determined significance level a where the pvalue is the probability when the data are considered random of observing a test statistic as or more extreme than that observed if the null hypothesis is true.

#### Binomial model

If  $Y \sim Bin(n, \theta)$ , then the standard hypotheses are

- $H_0: \theta = \theta_0 = 0.5$  and
- $H_A: \theta \neq \theta_0$ .

In this case, the

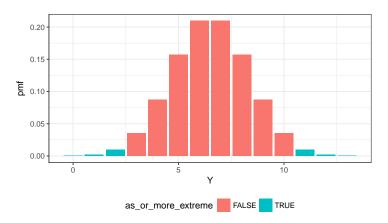
- test statistic is Y,
- its sampling distribution when the null hypothesis is true is  $Y \sim Bin(n, \theta_0)$ , and
- the as or more extreme region is values farther from  $n\theta_0$  than y.

So the pvalue is

$$pvalue = P(|Y - n\theta_0| \ge |y - n\theta_0|)$$

where y is the observed successes.

```
library(dplyr); library(ggplot2)
n <- 13; y <- 2; theta0 <- 0.5
d <- data.frame(Y = 0:n) %>%
mutate(pmf = dbinom(Y, n, theta0),
    as_or_more_extreme = abs(Y-n*theta0) >= abs(y-n*theta0))
ggplot(d, aes(Y, pmf, fill=as_or_more_extreme)) + geom_bar(stat = "identity") +
    theme_bw() + theme(legend.position="bottom")
```



## Binomial example

If  $Y \sim Bin(n,\theta)$  with n=13 and y=2 and we are testing

- $H_0: \theta = 0.5 \text{ versus}$
- $H_A: \theta \neq 0.5$ ,

then the pvalue is

$$pvalue = \sum_{y=0}^{2} P(Y = y | \theta = 0.5) + \sum_{y=0}^{13} P(Y = y | \theta = 0.5)$$

which is

```
(p <- sum(dbinom(c(0:2,11:13), size = 13, prob = 0.5)))
[1] 0.02246094
```

Thus, we would *reject the null hypothesis* for any significance level greater than 0.0224609.

#### binom.test

binom.test(2.13)

#### The R function 'binom.test' can perform this test for us:

```
Exact binomial test

data: 2 and 13
number of successes = 2, number of trials = 13, p-value = 0.02246
alternative hypothesis: true probability of success is not equal to 0.5
95 percent confidence interval:
0.01920667 0.45447106
sample estimates:
probability of success
0.1538462
```

### One-sided pvalues

If  $Y \sim Bin(n, \theta)$ , a one-sided hypothesis test is

- $\bullet \ H_0: \theta \geq \theta_0 = 0.5 \ \text{and}$
- $H_A: \theta < \theta_0$ .

In this case, the

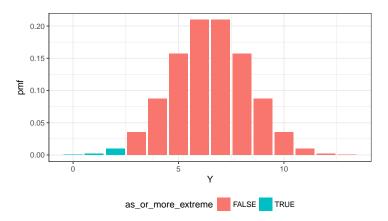
- $\bullet$  test statistic is Y,
- its sampling distribution when the null hypothesis is true is  $Y \sim Bin(n,\theta_0)$ , and
- the as or more extreme region is values farther from  $n\theta_0$  than y in the direction of  $H_A$ .

So the pvalue is

$$pvalue = P(Y - n\theta_0 \le y - n\theta_0) = P(Y \le y)$$

where y is the observed successes.

```
library(dplyr); library(ggplot2)
n <- 13; y <- 2; theta0 <- 0.5
d <- data.frame(Y = 0:n) %>%
  mutate(pmf = dbinom(Y, n, theta0),
         as_or_more_extreme = Y <= y)
ggplot(d, aes(Y, pmf, fill=as_or_more_extreme)) + geom_bar(stat = "identity") +
  theme_bw() + theme(legend.position="bottom")
```



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## Binomial example

If  $Y \sim Bin(n, \theta)$  with n = 13 and y = 2 and we are testing

- $H_0: \theta \geq 0.5$  versus
- $H_A: \theta < 0.5$ ,

then the pvalue is

$$pvalue = \sum_{y=0}^{2} P(Y = y | \theta = 0.5)$$

which is

```
(p <- sum(dbinom(0:2, size = 13, prob = 0.5)))
[1] 0.01123047
```

Thus, we would *reject the null hypothesis* for any significance level greater than 0.0112305.

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### binom.test()

### The R function 'binom.test()' can perform this test for us:

```
binom.test(2, 13, alternative="less")

Exact binomial test

data: 2 and 13
number of successes = 2, number of trials = 13, p-value = 0.01123
alternative hypothesis: true probability of success is less than 0.5

95 percent confidence interval:
0.0000000 0.4100986
sample estimates:
probability of success
0.1538462
```

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## Asymptotic pvalues

If we have an asymptotically normal estimator  $\hat{\theta}=\hat{\theta}(Y)$ , i.e.

$$\hat{\theta}(Y) \stackrel{.}{\sim} N(E[\hat{\theta}], Var[\hat{\theta}])$$

then we can calculate pvalues using this approximate sampling distribution.

- $H_0: \theta = \theta_0 \implies p = P(|\hat{\theta}(Y) E[\hat{\theta}]| \ge |\hat{\theta}(y) E[\hat{\theta}]|)$
- $H_0: \theta \ge \theta_0 \implies p = P(\hat{\theta}(Y) \le \hat{\theta}(y))$
- $H_0: \theta \le \theta_0 \implies p = P(\hat{\theta}(Y) \ge \hat{\theta}(y))$

where

- ullet  $\hat{ heta}(Y)$  is the random estimator and
- $\hat{\theta}(y)$  is the observed estimator.

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### Binomial example

If  $Y \sim Bin(n, \theta)$  and n is large (and y is not close to 0 or n), then

$$Y \stackrel{.}{\sim} N(n\theta, n\theta(1-\theta)).$$

If we have

$$H_0: \theta = \theta_0$$
 versus  $H_A: \theta \neq \theta_0$ ,

then we our pvalue is

$$pvalue = P(|Y - n\theta_0| \ge |y - n\theta_0|)$$

$$= 2P\left(\frac{Y - n\theta_0}{Var[\theta]} < \frac{-|y - n\theta_0|}{SE[\hat{\theta}]}\right)$$

$$\approx 2P\left(Z < \frac{-|y - n\theta_0|}{\sqrt{n\theta_0(1 - \theta_0)}}\right)$$

```
n = 10000; y = 4900; theta0 = 0.5
2*pnorm(-abs(y-n*theta0)/sqrt(n*theta0*(1-theta0)))
```

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[1] 0.04550026

### prop.test()

For the binomial distribution, the prop.test() function performs these hypothesis tests. For example, if  $Y \sim Bin(n,\theta)$  and you want to test  $H_0: \theta = 0.5$  vs  $H_A: \theta \neq 0.5$  when observing y = 4900 successes out of  $n = 10^4$  attempts, the code is

```
prop.test(y, n, p = theta0, correct = FALSE)

1-sample proportions test without continuity correction

data: y out of n, null probability theta0

X-squared = 4, df = 1, p-value = 0.0455
alternative hypothesis: true p is not equal to 0.5

95 percent confidence interval:
0.4802079 0.4997998
sample estimates:
p
0.49
0.49
```

#### But you should use the continuity correction:

```
prop.test(y, n, p = theta0, correct = TRUE)$p.value
[1] 0.04659094
```

### Normal mean

Let  $Y_i \stackrel{ind}{\sim} N(\mu, \sigma^2)$ , then

$$T = \frac{\overline{Y} - \mu}{S/\sqrt{n}} \sim t_{n-1}(0, 1)$$

is our test statistic and its sampling distribution. We have the following null hypothesis tests and pvalues

- $H_0: \mu = \mu_0$  and  $pvalue = P(|T| \ge |t|) = 2P(T < -|t|)$
- $H_0: \mu \geq \mu_0$  and  $pvalue = P(T \leq t) = P(T < t)$
- $H_0: \mu \le \mu_0$  and  $pvalue = P(T \ge t) = 1 P(T < t)$

where

$$t = \frac{\overline{y} - \mu}{s / \sqrt{n}}$$

is the observed value of our test statistic. This is called a one-sample t-test.

#### t.test

```
set.seed(1); y <- rnorm(15, mean = 1)
t.test(y)
One Sample t-test
data: y
t = 4.1894, df = 14, p-value = 0.0009091
alternative hypothesis: true mean is not equal to 0
95 percent confidence interval:
0.5372624 1.6644233
sample estimates:
mean of x
1.100843
t.test(y, mu = 1, alternative = "greater")
One Sample t-test
data: y
t = 0.38377, df = 14, p-value = 0.3535
alternative hypothesis: true mean is greater than 1
95 percent confidence interval:
0.6380276
                Inf
sample estimates:
mean of x
1.100843
```

## Relationship to confidence intervals

There is a one-to-one correspondence between pvalues and confidence intervals. Consider the following null hypotheses and corresponding confidence intervals (CIs)

- $H_0: \theta = \theta_0$  (two-sided CI),
- $H_0: \theta \geq \theta_0$  (one-sided lower), and
- $H_0: \theta \leq \theta_0$  (one-sided upper),

#### **Theorem**

The appropriate (two-sided vs one-sided in the correct direction) 100(1-a)% confidence interval contains  $\theta_0$  if and only if the pvalue is greater than a.

### Interpreting pvalues

We teach students to say the phrases

- ullet reject the null hypothesis if pvalue < a or
- fail to reject the null hypothesis if  $pvalue \ge a$ .

But this is incorrect.

According to the American Statistical Association Statement on Pvalues:

Pvalues can indicate how incompatible the data are with a specific statistical model.

The specific statistical model is the model associated with the null hypothesis. So we're going to say

- ullet the data are incompatible with the null hypothesis if pvalue < a or
- there is insufficient evidence to indicate the data are incompatible with the null hypothesis if  $pvalue \ge a$ .