R06 - ANOVA and F-tests

STAT 401 (Engineering) - Iowa State University

April 2, 2018

One-way ANOVA model/assumptions

The one-way ANOVA (ANalysis Of VAriance) model is

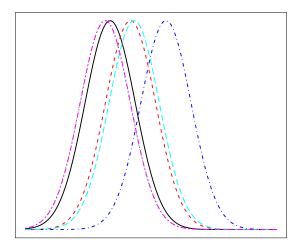
$$Y_{ij} \overset{ind}{\sim} N\left(\mu_j, \sigma^2\right)$$
 or $Y_{ij} = \mu_j + \epsilon_{ij}, \, \epsilon_{ij} \overset{iid}{\sim} N(0, \sigma^2)$

for
$$j = 1, \ldots, J$$
 and $i = 1, \ldots, n_j$.

Assumptions:

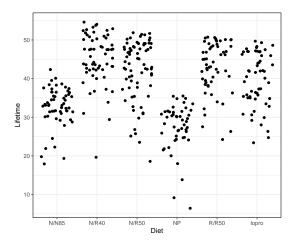
- Errors are normally distributed.
- Errors have a common variance.
- Errors are independent.

ANOVA assumptions graphically



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Consider the mice data set



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One-way ANOVA F-test

Are any of the means different?

Hypotheses in English:

 H_0 : all the means are the same

 H_1 : at least one of the means is different

Statistical hypotheses:

$$\begin{array}{ll} H_0: & \mu_j = \mu \text{ for all } j & Y_{ij} \stackrel{iid}{\sim} N(\mu,\sigma^2) \\ H_1: & \mu_j \neq \mu_{j'} \text{ for some } j \text{ and } j' & Y_{ij} \stackrel{ind}{\sim} N\left(\mu_j,\sigma^2\right) \end{array}$$

An ANOVA table organizes the relevant quantities for this test and computes the pvalue.

ANOVA table

A start of an ANOVA table:

Source of variation	Sum of squares	d.f.	Mean square
		J-1	
Error (Within groups)	$SSE = \sum_{j=1}^{J} \sum_{i=1}^{n_j} (Y_{ij} - \overline{Y}_j)^2$	n-J	$\frac{SSE}{n-J} \left(= \hat{\sigma}^2 \right)$
Total	$SST = \sum_{i=1}^{J} \sum_{i=1}^{n_j} (Y_{ij} - \overline{Y})^2$	n-1	

where

- J is the number of groups,
- n_j is the number of observations in group j,
- $n = \sum_{j=1}^{J} n_j$ (total observations),
- ullet $\overline{Y}_j = rac{1}{n_j} \sum_{i=1}^{n_j} Y_{ij}$ (average in group j),
- and $\overline{Y} = \frac{1}{n} \sum_{j=1}^{J} \sum_{i=1}^{n_j} Y_{ij}$ (overall average).

ANOVA table

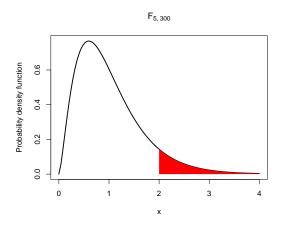
An easier to remember ANOVA table:

Source of variation	Sum of squares	df	Mean square	F-statistic	p-value
Factor A (between groups)	SSA	J-1	MSA = SSA/J - 1	MSA/MSE	(see below)
Error (within groups)	SSE	n - J	MSE = SSE/n - J		,
Total	SST=SSA+SSE	n-1			

Under H_0 ($\mu_j = \mu$),

- the quantity MSA/MSE has an F-distribution with J-1 numerator and n-J denominator degrees of freedom,
- ullet larger values of MSA/MSE indicate evidence against H_0 , and
- the p-value is determined by $P(F_{J-1,n-J} > MSA/MSE)$.

F-distribution



One-way ANOVA F-test (by hand)

```
# A tibble: 7 x 4
 Diet.
          n mean
<chr> <int> <dbl> <dbl>
1 N/N85
       57 32.7 5.13
2 N/R40
       60 45.1 6.70
3 N/R50
       71 42.3 7.77
    49 27.4 6.13
4 NP
5 R/R50 56 42.9 6.68
6 lopro 56 39.7 6.99
7 Total
        349 38.8 NA
```

So

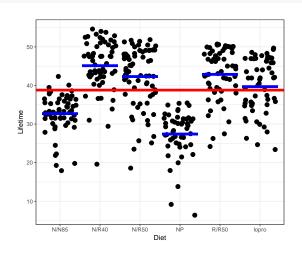
```
SSA = 57 \times (32.7 - 38.8)^2 + 60 \times (45.1 - 38.8)^2 + 71 \times (42.3 - 38.8)^2 + 49 \times (27.4 - 38.8)^2 \\ +56 \times (42.9 - 38.8)^2 + 56 \times (39.7 - 38.8)^2 = 12734
SST = (35.5 - 38.8)^2 + (35.4 - 38.8)^2 + (34.9 - 38.8)^2 + \dots + (19.6 - 38.8)^2 + (47.6 - 38.8)^2 = 28031
SSE = SST - SSA = 28031 - 12734 = 15297
J - 1 = 5
n - J = 349 - 6 = 343
n - 1 = 348
MSA = SSA/J - 1 = 12734/5 = 2547
MSE = SSE/n - J = 15297/343 = 44.6 = \hat{\sigma}^2
F = MSA/MSE = 2547/44.6 = 57.1
```

 $P(F_{5,343} > 57.1) < 0.0001$

p =

Graphical comparison

Warning: Ignoring unknown aesthetics: y



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R code and output for one-way ANOVA

General F-tests

The one-way ANOVA F-test is an example of a general hypothesis testing framework that uses F-tests. This framework can be used to test

- composite alternative hypotheses or, equivalently,
- a full vs a reduced model.

The general idea is to balance the amount of variability remaining when moving from the reduced model to the full model measured using the sums of squared errors (SSEs) relative to the amount of complexity, i.e. parameters, added to the model.

Testing full vs reduced models

If $Y_{ij} \stackrel{ind}{\sim} N(\mu_j, \sigma^2)$ for $j = 1, \dots, J$ and we want to test the hypotheses

- $H_0: \mu_j = \mu$ for all j
- $H_1: \mu_j \neq \mu_{j'}$ for some j and j'

think about this as two models:

- $H_0: Y_{ij} \stackrel{ind}{\sim} N(\mu, \sigma^2)$ (reduced)
- $H_1: Y_{ij} \stackrel{ind}{\sim} N(\mu_j, \sigma^2)$ (full)

We can use an F-test to calculate a p-value for tests of this type.

Nested models: full vs reduced

Definition

Two models are nested if the reduced model is a special case of the full model.

For example, consider the full model

$$Y_{ij} \stackrel{ind}{\sim} N(\mu_j, \sigma^2).$$

One special case of this model occurs when $\mu_j=\mu$ and thus

$$Y_{ij} \stackrel{ind}{\sim} N(\mu, \sigma^2).$$

is a reduced model and these two models are nested.

Calculating the sum of squared residuals (errors)

Model	Full	Reduced
Assumption	$H_1: Y_{ij} \stackrel{ind}{\sim} N\left(\mu_j, \sigma^2\right)$	$H_0: Y_{ij} \stackrel{iid}{\sim} N(\mu, \sigma^2)$
Mean	$\hat{\mu}_j = \overline{Y}_j = \frac{1}{n_j} \sum_{i=1}^{n_j} Y_{ij}$	$\hat{\mu} = \overline{Y} = \frac{1}{n} \sum_{j=1}^{J} \sum_{i=1}^{n_j} Y_{ij}$
Residual	$r_{ij} = Y_{ij} - \hat{\mu}_j = Y_{ij} - \overline{Y}_j$	$r_{ij} = Y_{ij} - \hat{\mu} = Y_{ij} - \overline{Y}$
SSE	$\sum_{j=1}^{J} \sum_{i=1}^{n_j} r_{ij}^2$	$\sum_{j=1}^{J} \sum_{i=1}^{n_j} r_{ij}^2$

General F-tests

Do the following

Calculate

2. Calculate

Extra degrees of freedom = # of mean parameters (full) - # of mean parameters (reduced)

3. Calculate F-statistics

$${\rm F} = \frac{{\rm Extra~sum~of~squares~/~Extra~degrees~of~freedom}}{\hat{\sigma}^2}$$

- 4. A pvalue is $P(F_{ndf,ddf} > \mathsf{F})$
 - ullet numerator degrees of freedom (ndf) = Extra degrees of freedom
 - denominator degrees of freedom (ddf): df associated with $\hat{\sigma}^2$

Mice lifetimes

Consider the hypothesis that all diets have a common mean lifetime except NP.

Let

$$Y_{ij} \stackrel{ind}{\sim} N(\mu_j, \sigma^2)$$

with j=1 being the NP group then the hypotheses are

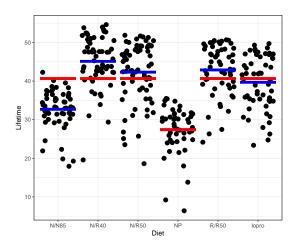
- $H_0: \mu_j = \mu \text{ for } j \neq 1$
- $H_1: \mu_j \neq \mu_{j'}$ for some $j, j' = 2, \dots, 6$

As models:

- $H_0: Y_{i1} \stackrel{iid}{\sim} N(\mu_1, \sigma^2)$ and $Y_{ij} \stackrel{iid}{\sim} N(\mu, \sigma^2)$ for $j \neq 1$
- $H_1: Y_{ij} \stackrel{ind}{\sim} N(\mu_i, \sigma^2)$

As a picture

```
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```



Making R do the calculations

```
case0501$NP = factor(case0501$Diet == "NP")
modR = lm(Lifetime~NP, case0501) # (R)educed model
modF = lm(Lifetime~Diet, case0501) # (F)ull model
anova (modR, modF)
Analysis of Variance Table
Model 1: Lifetime ~ NP
Model 2: Lifetime ~ Diet
  Res.Df RSS Df Sum of Sq F Pr(>F)
  347 20630
    343 15297 4 5332.2 29.89 < 2.2e-16 ***
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

Lack-of-fit F-test

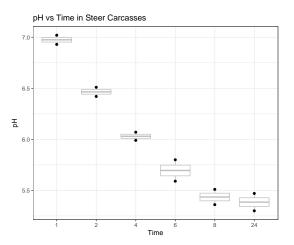
Let Y_{ij} be the i^{th} observation from the j^{th} group where the group is defined by those observations having the same explanatory variable value (X_i) .

Two models:

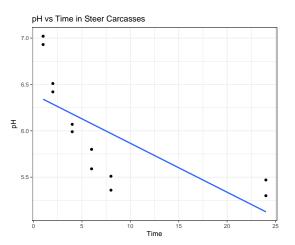
ANOVA:
$$Y_{ij} \stackrel{ind}{\sim} N(\mu_j, \sigma^2)$$
 (full) Regression: $Y_{ij} \stackrel{ind}{\sim} N(\beta_0 + \beta_1 X_j, \sigma^2)$ (reduced)

- Regression model is reduced:
 - ANOVA has J parameters for the mean
 - Regression has 2 parameters for the mean
 - Set $\mu_i = \beta_0 + \beta_1 X_i$.
- Small pvalues indicate a lack-of-fit, i.e. the regression (reduced) model is not adequate.
- Lack-of-fit F-test requires multiple observations at a few X_i values!

pH vs Time - ANOVA



pH vs Time - Regression





Lack-of-fit F-test in R

```
# Use as.factor to turn a continuous variable into a categorical variable
m_anova = lm(pH ~ as.factor(Time), Sleuth3::ex0816)
m_reg = lm(pH ~ Time , Sleuth3::ex0816)
anova(m_reg, m_anova)

Analysis of Variance Table

Model 1: pH ~ Time
Model 2: pH ~ as.factor(Time)
Res.Df RSS Df Sum of Sq F Pr(>F)
1 10 1.97289
2 6 0.05905 4 1.9138 48.616 0.0001048 ***
---
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

There is evidence the data are incompatible with the null hypothesis that states the means of each group fall along a line.

Summary

- Use F-tests for comparison of full vs reduced model
 - One-way ANOVA F-test
 - General F-tests
 - Lack-of-fit F-tests

Think about F-tests as comparing models.