

## I05 - Confidence intervals

STAT 587 (Engineering)  
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## Exact confidence intervals

The **coverage** of an interval estimator is the probability the interval will contain the true value of the parameter *when the data are considered to be random*. If an interval estimator has  $100(1 - \alpha)\%$  coverage, then we call it a  $100(1 - \alpha)\%$  **confidence interval** and  $1 - \alpha$  is the **confidence level**.

That is, we calculate

$$1 - \alpha = P(L < \theta < U)$$

where  $L$  and  $U$  are random because they depend on the data. Thus **confidence** is a statement about the **procedure**.

## Normal model

If  $Y_i \stackrel{ind}{\sim} N(\mu, \sigma^2)$  and we assume the default prior  $p(\mu, \sigma^2) \propto 1/\sigma^2$ , then a  $100(1 - a)\%$  credible interval for  $\mu$  is given by

$$\bar{y} \pm t_{n-1, a/2} s / \sqrt{n}.$$

When the data are considered random

$$T_{n-1} = \frac{\bar{Y} - \mu}{S/\sqrt{n}} \sim t_{n-1}(0, 1)$$

thus the probability  $\mu$  is within our credible interval is

$$\begin{aligned} & P\left(\bar{Y} - t_{n-1, a/2} S / \sqrt{n} < \mu < \bar{Y} + t_{n-1, a/2} S / \sqrt{n}\right) \\ &= P\left(-t_{n-1, a/2} < \frac{\bar{Y} - \mu}{S/\sqrt{n}} < t_{n-1, a/2}\right) \\ &= P\left(-t_{n-1, a/2} < T_{n-1} < t_{n-1, a/2}\right) \\ &= 1 - a. \end{aligned}$$

Thus, this  $100(1 - a)\%$  credible interval is also a  $100(1 - a)\%$  confidence interval.

## Yield data example

Recall the corn yield example from I04 with 9 randomly selected fields in Iowa whose sample average yield is 186 and sample standard deviation is 22. Then a 95% confidence interval for the mean corn yield on Iowa farms is

$$186 \pm 2.31 \times 22/\sqrt{9} = (169, 202).$$

## Standard error

The **standard error of an estimator** is an *estimate* of the standard deviation of the estimator (when the data are considered random).

If  $Y \sim \text{Bin}(n, \theta)$ , then

$$\hat{\theta} = \frac{Y}{n} \quad \text{has} \quad SE[\hat{\theta}] = \sqrt{\frac{\hat{\theta}(1 - \hat{\theta})}{n}}.$$

If  $Y_i \stackrel{\text{ind}}{\sim} N(\mu, \sigma^2)$ , then

$$\hat{\mu} = \bar{Y} \quad \text{has} \quad SE[\hat{\mu}] = S/\sqrt{n}.$$

# Approximate confidence intervals

If an **unbiased** estimator has an asymptotic normal distribution, then we can construct an **approximate**  $100(1 - \alpha)\%$  confidence interval for  $E[\hat{\theta}] = \theta$  using

$$\hat{\theta} \pm z_{\alpha/2} SE[\hat{\theta}].$$

where  $SE[\hat{\theta}]$  is the **standard error** of the estimator and  $P(Z > z_{\alpha/2}) = \alpha/2$ .

This comes from the fact that if  $\hat{\theta} \sim N(\theta, SE[\hat{\theta}]^2)$ , then

$$\begin{aligned} &P\left(\hat{\theta} - z_{\alpha/2} SE(\hat{\theta}) < \theta < \hat{\theta} + z_{\alpha/2} SE(\hat{\theta})\right) \\ &= P\left(-z_{\alpha/2} < \frac{\hat{\theta} - \theta}{SE(\hat{\theta})} < z_{\alpha/2}\right) \\ &\approx P\left(-z_{\alpha/2} < Z < z_{\alpha/2}\right) \\ &= 1 - \alpha. \end{aligned}$$

## Normal example

If  $Y_i \stackrel{\text{ind}}{\sim} N(\mu, \sigma^2)$  and we have the estimator  $\hat{\mu} = \bar{Y}$ , then

$$E[\hat{\mu}] = \mu \quad \text{and} \quad SE[\hat{\mu}] = S/\sqrt{n}$$

Thus an **approximate**  $100(1 - \alpha)\%$  confidence interval for  $\mu = E[\hat{\mu}]$  is

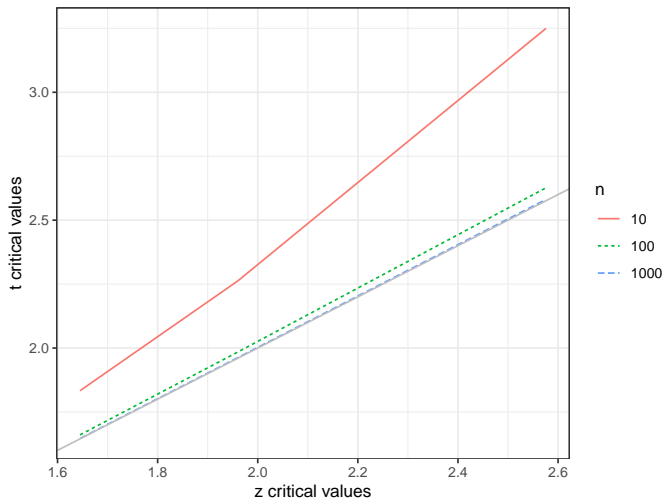
$$\hat{\mu} \pm z_{\alpha/2} SE[\hat{\mu}] = \bar{Y} \pm z_{\alpha/2} S/\sqrt{n}.$$

Note that this is almost identical to the **exact**  $100(1 - \alpha)\%$  confidence interval for  $\mu$ ,

$$\bar{Y} \pm t_{n-1, \alpha/2} S/\sqrt{n}$$

and when  $n$  is large  $z_{\alpha/2} \approx t_{n-1, \alpha/2}$ .

# T critical values vs Z critical values





# Approximate confidence interval for binomial proportion

If  $Y \sim \text{Bin}(n, \theta)$ , then an **approximate**  $100(1 - \alpha)\%$  confidence interval for  $\theta$  is

$$\hat{\theta} \pm z_{\alpha/2} \sqrt{\frac{\hat{\theta}(1 - \hat{\theta})}{n}}.$$

where  $\hat{\theta} = Y/n$  since

$$E[\hat{\theta}] = E\left[\frac{Y}{n}\right] = \theta$$

and

$$SE[\hat{\theta}] = \sqrt{\frac{\hat{\theta}(1 - \hat{\theta})}{n}}.$$

## Gallup poll example

In a Gallup poll dated 2017/02/19, 32.1% of respondents of the 1,500 randomly selected U.S. adults indicated that they were “engaged at work”. Thus an approximate 95% confidence interval for the proportion of all U.S. adults is

$$0.321 \pm 1.96 \times \sqrt{\frac{.321(1 - .321)}{1500}} = (0.30, 0.34).$$

## Confidence interval summary

Model	Parameter	Estimator	Confidence Interval	Type
$Y_i \stackrel{ind}{\sim} N(\mu, \sigma^2)$	$\mu$	$\hat{\mu} = \bar{y}$	$\hat{\mu} \pm t_{n-1, a/2} s / \sqrt{n}$	exact
$Y_i \stackrel{ind}{\sim} N(\mu, \sigma^2)$	$\mu$	$\hat{\mu} = \bar{y}$	$\hat{\mu} \pm z_{a/2} s / \sqrt{n}$	approximate
$Y \sim Bin(n, \theta)$	$\theta$	$\hat{\theta} = y/n$	$\hat{\theta} \pm z_{a/2} \sqrt{\hat{\theta}(1 - \hat{\theta})/n}$	approximate
$Y_i \stackrel{ind}{\sim} Ber(\theta)$	$\theta$	$\hat{\theta} = \bar{y}$	$\hat{\theta} \pm z_{a/2} \sqrt{\hat{\theta}(1 - \hat{\theta})/n}$	approximate

Bayesian credible intervals generally provide approximate confidence intervals.

**Approximate** means that the coverage will get closer to the desired probability, i.e.  $100(1 - a)\%$ , as the sample size gets larger.