# STAT 401A - Statistical Methods for Research Workers Inference Using *t*-Distributions

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## Random variables

From: http://www.stats.gla.ac.uk/steps/glossary/probability\_distributions.html

#### Definition

A random variable is a function that associates a unique numerical value with every outcome of an experiment.

#### Definition

A discrete random variable is one which may take on only a countable number of distinct values such as 0, 1, 2, 3, 4,... Discrete random variables are usually (but not necessarily) counts.

#### Definition

A continuous random variable is one which takes an infinite number of possible values. Continuous random variables are usually measurements.

## Random variables

#### Examples:

- Discrete random variables
  - Coin toss: Heads (1) or Tails (0)
  - Die roll: 1, 2, 3, 4, 5, or 6
  - Number of Ovenbirds at a 10-minute point count
  - RNAseq feature count
- Continuous random variables
  - Pig average daily (weight) gain
  - Corn yield per acre

## Statistical notation

Let Y be 1 if the coin toss is heads and 0 if tails, then

$$Y \sim Bin(n, p)$$

which means

Y is a binomial random variable with n trials and probability of success p

For example, if Y is the number of heads observed when tossing a fair coin ten times, then  $Y \sim Bin(10, 0.5)$ .

Later we will be constructing  $100(1-\alpha)\%$  confidence intervals, these intervals are constructed such that if n of them are constructed then  $Y \sim Bin(n, 1-\alpha)$  will cover the true value.

## Statistical notation

Let  $Y_i$  be the average daily (weight) gain in pounds for the ith pig, then

$$Y_i \stackrel{iid}{\sim} N(\mu, \sigma^2)$$

which means

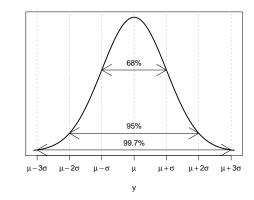
 $Y_i$  are independent and identically distributed normal (Gaussian) random variables with expected value  $E[Y_i] = \mu$  and variance  $V[Y_i] = \sigma^2$  (standard deviation  $\sigma$ ).

For example, if a litter of pigs is expected to gain 2 lbs/day with a standard deviation of 0.5 lbs/day and the knowledge of how much one pig gained does not affect what we think about how much the others have gained, then  $Y_i \stackrel{iid}{\sim} N(2, 0.5^2)$ .

# Normal (Gaussian) distribution

A random variable Y has a normal distribution, i.e.  $Y \sim N(\mu, \sigma^2)$ , with mean  $\mu$  and variance  $\sigma^2$  if draws from this distribution follow a bell curve centered at  $\mu$  with spread determined by  $\sigma^2$ :

#### Probability density function

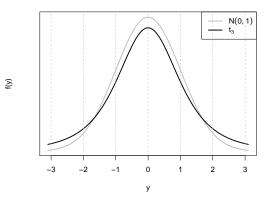


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## t-distribution

A random variable Y has a t-distribution, i.e.  $Y \sim t_{\nu}$ , with degrees of freedom  $\nu$  if draws from this distribution follow a similar bell shaped pattern:

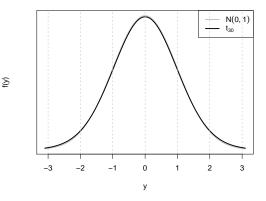
#### Probability density function



## t-distribution

As  $v \to \infty$ , then  $t_v \overset{d}{\to} N(0,1)$ , i.e. as the degrees of freedom increase, a t distribution gets closer and closer to a standard normal distribution, i.e. N(0,1). If v>30, the differences is negligible.

#### Probability density function

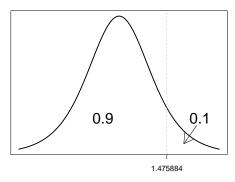


## t critical value

#### Definition

If  $T \sim t_v$ , a  $t_v(1 - \alpha/2)$  critical value is the value such that  $P(T < t_v(1 - \alpha/2)) = 1 - \alpha/2$  (or  $P(T > t_v(1 - \alpha)) = \alpha/2$ ).

#### Probability density function t<sub>5</sub>



t

# **Hypotheses**

#### Three key features:

- a test statistic calculated from data
- a sampling distribution for the test statistic under the null hypothesis
- a region that is as or more extreme (one-sided vs two-sided hypotheses)

Calculate probability of being in the region:

#### Definition

A pvalue is the probability of observing a test statistic as or more extreme than that observed, if the null hypothesis is true.

- ullet If pvalue is less than or equal to lpha, we reject the null hypothesis.
- If pvalue is greater than  $\alpha$ , we fail to reject the null hypothesis.

## **Z**-statistics

Let's assume, we have

- calculated a test statistic z that
- ullet has a  $Z \sim N(0,1)$  sampling distribution if the null hypothesis is true.

We could easily replace z with t and have a sampling distribution that is  $t_v$  for some v.

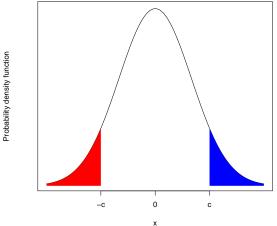
Now, we can have one of three types of hypotheses:

- Two-sided: P(|Z| > |z|) = P(Z > |z|) + P(Z < -|z|)= 2P(Z < -|z|)
- One-sided:
  - P(Z > z) = P(Z < -z)
  - P(Z < z)

F(c) = P(Z < c) is the cumulative distribution function for the standard normal.

# Symmetric distributions

The standard normal and t distributions are both symmetric around zero.



$$P(Z > c) = P(Z < -c)$$
 blue area is equal to red area

# Paired t-test example

In a paired t-test, we might have

- a one-sided hypothesis that the difference is greater than zero
- a test statistic t = 2.43
- which has a t distribution with 7 degrees of freedom when the null hypothesis is true.

So we need to calculate

$$P(t_7 > 2.43) = 1 - P(t_7 < 2.43)$$

where  $t_7$  represents a t-distribution with 7 degrees of freedom.

# Using SAS or R

```
In SAS,
PROC IML;
  p = 1-CDF('T', 2.43, 7);
  PRINT p;
  QUIT;
In R.
p = 1-pt(2.43,7)
```

Both obtain p=0.0227.

# Paired t-test example

In a paired t-test, we might have

- a one-sided hypothesis that the difference is greater than zero
- a test statistic t = 2.43
- which has a t distribution with 7 degrees of freedom when the null hypothesis is true.

So we need to calculate

$$P(t_7 > 2.43) = 1 - P(t_7 < 2.43) = 0.02$$

So, we reject the null hypothesis that there is no difference, but how big is the difference?

## Confidence interval construction

Key steps in confidence interval construction:

- Calculate statistic
- Calculate standard error of the statistic
- Find the appropriate critical value
- Construct the interval
  - Two-sided interval:

statistic  $\pm$  critical value imes standard error

One-sided interval:

 $statistic - critical\ value \times standard\ error$ 

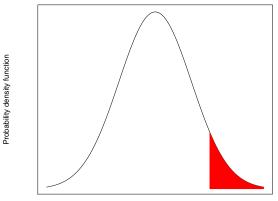
or

statistic + critical value × standard error

## Critical values

A related quantity are critical values for confidence interval construction, e.g.

$$(\overline{Y}_2 - \overline{Y}_1) \pm t_v (1 - \alpha/2) SE(\overline{Y}_2 - \overline{Y}_1).$$



х

# Using SAS or R

```
If \alpha = 0.05, then 1 - \alpha/2 = 0.975.
In SAS.
PROC IML;
  q = QUANTILE('T', 0.975, 7);
  PRINT q;
  QUIT;
In R,
q = qt(0.975,7)
```

Both obtain q=2.364.

## Summary

Two main approaches to statistical inference:

- Statistical hypothesis (hypothesis test)
- Statistical question (confidence interval)

# Cedar-apple rust

Cedar-apple rust is a (non-fatal) disease that affects apple trees. Its most obvious symptom is rust-colored spots on apple leaves. Red cedar trees are the immediate source of the fungus that infects the apple trees. If you could remove all red cedar trees within a few miles of the orchard, you should eliminate the problem. In the first year of this experiment the number of affected leaves on 8 trees was counted; the following winter all red cedar trees within 100 yards of the orchard were removed and the following year the same trees were examined for affected leaves.

- Statistical hypothesis:
  - $H_0$ : Removing red cedar trees increases or maintains the same mean number of rusty leaves.
  - $H_1$ : Removing red cedar trees decreases the mean number of rusty leaves.
- Statistical question:

What is the reduction of rusty leaves in our sample between year 1 and year 2 (perhaps due to removal of red cedar trees?

## Data

#### Here are the data

```
y1 = c(38, 10, 84, 36, 50, 35, 73, 48)
y2 = c(32,16,57,28,55,12,61,29)
leaves = data.frame(year1=y1, year2=y2, diff=y1-y2)
leaves
  year1 year2 diff
     38
          32
    10
         16
               -6
               27
        28
    50
        55
              -5
    73
         61
              12
     48
          29
              19
mean(leaves$diff)
Γ17 10.5
```

Is this a statistically significant difference?

# Assumptions

#### Let

- $Y_{1j}$  be the number of rusty leaves on tree j in year 1
- $Y_{2j}$  be the number of rusty leaves on tree j in year 2

#### Assume

$$D_j = Y_{1j} - Y_{2j} \stackrel{iid}{\sim} N(\mu, \sigma^2)$$

#### Then

$$H_0$$
:  $\mu = 0 \ (\mu \le 0)$ 

*H*<sub>1</sub>: 
$$\mu > 0$$

# Paired t-test pvalue

Test statistic

$$t = \frac{\overline{D_j} - \mu}{SE(\overline{D_j})}$$

where  $SE(\overline{D_i}) = s/\sqrt{n}$  with

- n being the number of observations (differences) and
- s being the sample standard deviation of the differences.

If  $H_0$  is true, then  $\mu=0$  and  $t\sim t_{n-1}$ . The pvalue is  $P(t_{n-1}>t)$  since this is a one-sided test.

For these data,

- $\overline{D_j} = 10.5$
- $SE(\overline{D_j}) = 4.31$
- t = 2.43
- p = 0.02

## Confidence interval

The  $100(1-\alpha)\%$  confidence interval has lower endpoint

$$\overline{D_j} - t_{n-1}(1-\alpha)SE(\overline{D_j})$$

and upper endpoint at infinity

For these data at 95% confidence, the lower endpoint is

$$10.5 - 1.89 \cdot 4.31 = 2.33$$

So we are 95% confident that the true difference in the number of rusty leaves is greater than 2.33.

# SAS code for paired t-test

```
DATA leaves;
  INPUT tree year1 year2;
  DATALINES;
1 38 32
2 10 16
3 84 57
4 36 28
5 50 55
6 35 12
7 73 61
8 48 29
PROC TTEST DATA=leaves SIDES=U;
    PAIRED year1*year2;
    RUN;
```

,

# SAS output for paired t-test

#### The TTEST Procedure

Difference: year1 - year2

N Mean Std Dev Std Err Minimum

8 10.5000 12.2007 4.3136 -6.0000

8 10.5000 12.2007 4.3136 -6.0000 27.0000

Mean 95% CL Mean Std Dev 95% CL Std Dev

10.5000 2.3275 Infty 12.2007 8.0668 24.8317

DF t Value Pr > t
7 2.43 0.0226

Maximum

# R output for paired t-test

```
t.test(leaves$year1, leaves$year2, paired=TRUE, alternative="greater")

Paired t-test

data: leaves$year1 and leaves$year2

t = 2.434, df = 7, p-value = 0.02257
alternative hypothesis: true difference in means is greater than 0

95 percent confidence interval:
2.328 Inf
sample estimates:
mean of the differences

10.5
```

## Statistical Conclusion

Removal of red cedar trees within 100 yards is associated with a significant reduction in rusty apple leaves (paired t-test t=2.43, p=0.023). The mean reduction in rust color leaves is  $10.5 [95\% \text{ CI } (2.33,\infty)]$ .

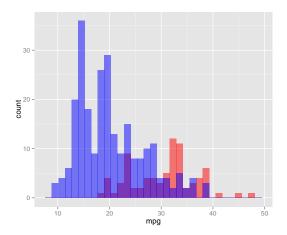
# Do Japanese cars get better mileage than American cars?

- Statistical hypothesis:
  - $H_0$ : Mean mpg of Japanese cars is the same as mean mpg of American cars.
  - *H*<sub>1</sub>: Mean mpg of Japanese cars is different than mean mpg of American cars.
- Statistical question:

What is the difference in mean mpg between Japanese and American cars?

- Data collection:
  - Collect a random sample of Japan/American cars

```
mpg = read.csv("mpg.csv")
library(ggplot2)
ggplot(mpg, aes(x=mpg))+
geom_histogram(data=subset(mpg,country=="Japan"), fill="red", alpha=0.5)+
geom_histogram(data=subset(mpg,country=="US"), fill="blue", alpha=0.5)
```



# Assumptions

#### Let

- $Y_{1i}$  represent the jth Japanese car
- $Y_{2j}$  represent the jth American car

#### Assume

$$Y_{1j} \stackrel{iid}{\sim} N(\mu_1, \sigma^2) \qquad Y_{2j} \stackrel{iid}{\sim} N(\mu_2, \sigma^2)$$

## Restate the hypotheses using this notation

 $H_0$ :  $\mu_1 = \mu_2$ 

 $H_1: \mu_1 \neq \mu_2$ 

Alternatively

 $H_0$ :  $\mu_1 - \mu_2 = 0$ 

 $H_1: \mu_1 - \mu_2 \neq 0$ 

## Test statistic

The test statistic we use here is

$$\frac{\overline{Y}_1 - \overline{Y}_2 - (\mu_1 - \mu_2)}{SE(\overline{Y}_1 - \overline{Y}_2)}$$

where

- $\bullet$   $\overline{Y}_1$  is the sample average mpg of the Japanese cars
- ullet  $\overline{Y}_2$  is the sample average mpg of the American cars

and

$$SE(\overline{Y}_1 - \overline{Y}_2) = s_p \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}$$
  $s_p = \sqrt{\frac{(n_1 - 1)s_1^2 + (n_2 - 1)s_2^2}{(n_1 + n_2 - 2)}}$ 

where

- ullet  $s_1$  is the sample standard deviation of the mpg of the Japanese cars
- $s_2$  is the sample standard deviation of the mpg of the American cars

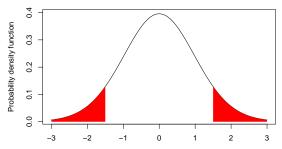
## **Pvalue**

If  $H_0$  is true, then  $\mu_1=\mu_2$  and the test statistic

$$t = \frac{\overline{Y}_1 - \overline{Y}_2 - (\mu_1 - \mu_2)}{\mathsf{SE}(\overline{Y}_1 - \overline{Y}_2)} \sim t_{n_1 + n_2 - 2}$$

where  $t_v$  is a t-distribution with df degrees of freedom.

Pvalue is  $P(|t_{n_1+n_2-2}| > |t|) = P(t_{n_1+n_2-2} > |t|) + P(t_{n_1+n_2-2} < -|t|)$  or as a picture



## Hand calculation

To calculate the quantity by hand, we need 6 numbers:

Car	Ν	Mean	SD	
Japanese	79	30.5	6.11	
American	249	20.1	6.41	

Calculate

$$s_p = \sqrt{\frac{(79-1)\cdot 6.11^2 + (249-1)\cdot 6.41^2}{79+249-2}} = 6.34$$

$$SE(\overline{Y}_1 - \overline{Y}_2) = 6.34\sqrt{\frac{1}{279} + \frac{1}{249}} = 0.82$$

$$t = \frac{30.5 - 20.1}{0.82} = 12.6$$

Finally, we are interested in finding  $P(|t_{326}| > |12.6|) < 0.0001$  which is found using a table or software.

## Confidence interval

Alternatively, we can construct a  $100(1-\alpha)\%$  confidence interval. The formula is

$$\overline{Y}_1 - \overline{Y}_2 \pm t_{n_1+n_2-2}(1-\alpha/2)SE(\overline{Y}_1 - \overline{Y}_2)$$

where  $\pm$  indicates plus and minus and  $t_v(1-\alpha/2)$  is the value such that  $P(t_v < t_v(1-\alpha/2)) = 1-\alpha/2$ . If  $\alpha = 0.05$  and df = 326, then  $t_v(1-\alpha/2) = 1.97$ .

The 95% confidence interval is

$$30.5 - 20.1 \pm 1.97 \cdot 0.82 = (8.73, 11.9)$$

We are 95% confident that, on average, Japanese cars get between 8.73 and 11.9 more mpg than American cars.

# SAS code for two-sample t-test

```
DATA mpg;
    INFILE 'mpg.csv' DELIMITER=',' FIRSTOBS=2;
    INPUT mpg country $;
PROC TTEST DATA=mpg;
    CLASS country;
    VAR mpg;
    RUN;
```

#### The TTEST Procedure

Variable: mpg

coun	try	N	Mean	Std Dev	S	td Err	Minimum	Maximum	
Japa	n	79	30.4810	6.1077		0.6872	18.0000	47.0000	
US		249	20.1446	6.4147		0.4065	9.0000	39.0000	
Diff	(1-2)		10.3364	6.3426		0.8190			
country	Method		Mean	95%	CL Me	an	Std Dev	95% CL 8	Std Dev
Japan			30.4810	29.113	30 31	.8491	6.1077	5.2814	7.2429
US			20.1446	19.343	39 20	.9452	6.4147	5.8964	7.0336
Diff (1-2)	Pooled		10.3364	8.725	2 11	.9477	6.3426	5.8909	6.8699
Diff (1-2)	Sattert	hwaite	10.3364	8.757	6 11	.9152			
Method Pooled		od	Variance	s	DF	t Value	Pr >  t		
		Equal		326	12.62	<.0001			
	Satt	erthwaite	e Unequal	136	8.87	12.95	<.0001		
			Equali	ty of Va	rianc	es			
		Method	Num DF	Den DF	F	Value	Pr > F		
		Folded H	248	78	3	1.10	0.6194		

# R code/output for two-sample t-test

```
t.test(mpg^country, data=mpg, var.equal=TRUE)

Two Sample t-test

data: mpg by country
t = 12.62, df = 326, p-value < 2.2e-16
alternative hypothesis: true difference in means is not equal to 0
95 percent confidence interval:
8.725 11.948
sample estimates:
mean in group Japan mean in group US
30.48 20.14</pre>
```

## Conclusion

Mean miles per gallon of Japanese cars is significantly different than mean miles per gallon of American cars (two-sample t-test t=12.62, p < 0.0001). Japanese cars get an average of 10.3 [95% CI (8.7,11.9)] more miles per gallon than American cars.