# STAT 401A - Statistical Methods for Research Workers Modeling assumptions

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# Normality assumptions

In the paired t-test, we assume

$$D_i \stackrel{iid}{\sim} N(\mu, \sigma^2).$$

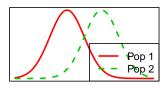
In the two-sample t-test, we assume

$$Y_{ij} \stackrel{ind}{\sim} N(\mu_j, \sigma^2).$$

### Paired t-test

# 0

### Two-sample t-test



Difference

Distribution

# Normality assumptions

In the paired t-test, we assume

$$D_i \stackrel{iid}{\sim} N(\mu, \sigma^2).$$

In the two-sample t-test, we assume

$$Y_{ij} \stackrel{ind}{\sim} N(\mu_j, \sigma^2).$$

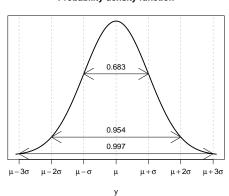
Key features of the normal distribution assumption:

- Centered at the mean (expectation)  $\mu$
- Standard deviation describes the spread
- Symmetric around  $\mu$  (no skewness)
- Non-heavy tails, i.e. outliers are rare (no kurtosis)

# Normality assumptions

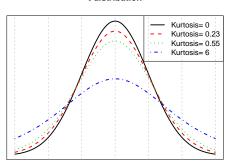
### Probability density function





### t distribution

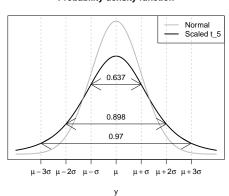
Probability density function, f(y)

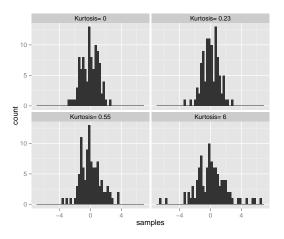


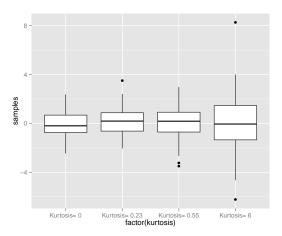
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### Probability density function





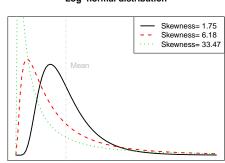




# **Skewness**

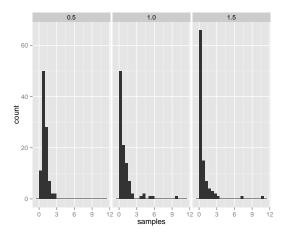
### Log-normal distribution





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# Samples from skewed distributions



### Robustness

### Definition

A statistical procedure is robust to departures from a particular assumption if it is valid even when the assumption is not met.

Remark If a 95% confidence interval is robust to departures from a particular assumption, the confidence interval should cover the true value about 95% of the time.

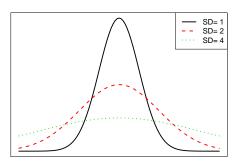
### Robustness to skewness and kurtosis

Percentage of 95% confidence intervals that cover the true difference in means in an equal-sample two-sample t-test with non-normal populations (where the distributions are the same other than their means).

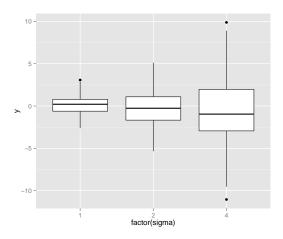
sample size	strongly skewed	moderately skewed	mildly skewed	heavy-tailed	short-tailed
5	95.5	95.4	95.2	98.3	94.5
10	95.5	95.4	95.2	98.3	94.6
25	95.3	95.3	95.1	98.2	94.9
50	95.1	95.3	95.1	98.1	95.2
100	94.8	95.3	95.0	98.0	95.6

# Differences in variances

### Normal distribution



# Differences in variances



### Robustness to differences in variances

Percentage of 95% confidence intervals that cover the true difference in means in an equal-sample two-sample t-test  $(r = \sigma_1/\sigma_2)$ .

n1	n2	r=1/4	r=1/2	r=1	r=2	r=4
10	10	95.2	94.2	94.7	95.2	94.5
10	20	83.0	89.3	94.4	98.7	99.1
10	40	71.0	82.6	95.2	99.5	99.9
100	100	94.8	96.2	95.4	95.3	95.1
100	200	86.5	88.3	94.8	98.8	99.4
100	400	71.6	81.5	95.0	99.5	99.9

### **Outliers**

### Definition

A statistical procedure is resistant if it does not change very much when a small part of the data changes, perhaps drastically.

### Identify outliers:

- If recording errors, fix.
- ② If outlier comes from a different population, remove and report.
- If results are the same with and without outliers, report with outliers.
- If results are different, use resistant analysis or report both analyses.

# Common ways for independence to be violated

- Cluster effect
  - e.g. pigs in a pen
- Correlation effect
  - e.g. measurements in time with drifting scale
- Spatial effect
  - e.g. corn yield plots (drainage)

## Common transformations for data

From: http://en.wikipedia.org/wiki/Data\_transformation\_(statistics)

### Definition

In statistics, data transformation refers to the application of a deterministic mathematical function to each point in a data set that is, each data point  $y_i$  is replaced with the transformed value  $z_i = f(y_i)$ , where f is a function.

The most common transformations to

- If  $y \in (0,1)$ , then  $f(y) = \sin^{-1}(\sqrt{y})$ .
- If y is a count, then  $f(y) = \sqrt{y}$ .
- If y is positive and right-skewed, then f(y) = log(y), the natural logarithm of y.

**Remark** Since  $log(0) = -\infty$ , the logarithm cannot be used directly when some  $y_i$  are zero. In these cases, use log(y+c) where c is something small relative to your data, e.g. half of the minimum non-zero value.

# Log transformation

If  $z_{ij} = log(y_{ij})$  and we run a two-sample t-test on the z's, then we assume

$$Z_{ij} \stackrel{ind}{\sim} N(\mu_j, \sigma^2)$$

and we calculate the quantity  $\overline{Z}_2 - \overline{Z}_1$  and  $\exp\left(\overline{Z}_2 - \overline{Z}_1\right) = e^{\overline{Z}_2 - \overline{Z}_1}$  estimates

Median of population 2
Median of population 1

on the original scale or, equivalently, it estimates the multiplicative effect of moving from population 1 to population 2.

# Log transformation interpretation

If we have a randomized experiment:

**Remark** It is estimated that the response of an experimental unit to treatment 2 will be  $\exp\left(\overline{Z}_2 - \overline{Z}_1\right)$  times as large as its response to treatment 1.

If we have an observational study:

**Remark** It is estimated that the median for population 2 is  $\exp\left(\overline{Z}_2 - \overline{Z}_1\right)$  times as large as the median for population 1.

# Confidence intervals with log transformation

If  $z_{ij} = log(y_{ij})$  and we assume

$$Z_{ij} \stackrel{ind}{\sim} N(\mu_j, \sigma^2),$$

then a  $100(1-\alpha)\%$  two-sided confidence interval for  $\mu_2-\mu_1$  is

$$(L,U) = \overline{Z}_2 - \overline{Z}_1 \pm t_{n_1+n_2-2}(1-\alpha/2)SE(\overline{Z}_2 - \overline{Z}_1).$$

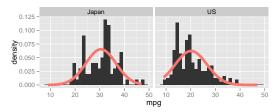
A  $100(1-\alpha)\%$  confidence interval for

Median of population 2
Median of population 1

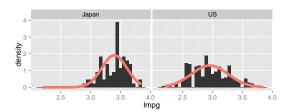
is  $(e^L, e^U)$ .

# Miles per gallon data

### Untransformed:

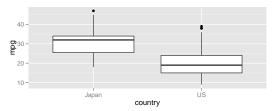


### Logged:

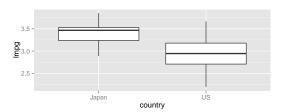


# Miles per gallon data

### Untransformed:



### Logged:



# Equal variances?

We might also be concerned about the assumption of equal variances.

### Untransformed:

	country	n	mean	sd
1	Japan	79	30.48	30.48
2	US	249	20.14	20.14

the ratio of the standard deviations is around 1.5 and there are 3 times as many observations in the US.

### Logged:

	country	n	mean	sd
1	Japan	79	3.40	3.40
2	US	249	2.96	2.96

Now the ratio of standard deviations is only 1.15.

# Using R for t-test using logarithms

```
t = t.test(log(mpg)~country, d, var.equal=TRUE)
exp(t$estimate)
mean in group Japan mean in group US
              29.85
                                 19.21
exp(-diff(t$estimate)) # I had to put in the negative sign
mean in group US
           1.554
exp(t$conf.int)
[1] 1.445 1.672
attr(, "conf.level")
[1] 0.95
```

# SAS code for t-test using logarithms

```
DATA mpg;
INFILE 'mpg.csv' DELIMITER=',' FIRSTOBS=2;
INPUT mpg country $;

PROC TTEST DATA=mpg TEST=ratio;
CLASS country;
VAR mpg;
run;
```

# SAS output for t-test using logarithms

The TTEST Procedure

Variable: mpg

	country	Geomet		efficient Variation	Min	imum	Maximum	
	Japan	79 29.8	525	0.2111	18.	0000	47.0000	
		249 19.2		0.3147			39.0000	
	Ratio (1/2)		544	0.2928				
		Geometric			Coe	fficient		
country	Method	Mean	95% CI	Mean	of V	ariation	95% CI	CV
Japan		29.8525	28.4887	31.2817		0.2111	0.1820	0.2514
JS <sup>*</sup>		19.2051	18.4825	19.9560		0.3147	0.2882	0.3467
Ratio (1/2)	Pooled	1.5544	1.4452	1.6719		0.2928	0.2712	0.3183
Ratio (1/2)	Satterthwaite	1.5544	1.4636	1.6508				
			Coet	ficients				
	Method	of Varia	tion	DF t	Value	Pr >  t	1	
	Pooled	Equal		326	11.91	<.000	1	
	Satterthwait		193	3.33	14.46	<.000	1	
		Equal	ity of Vari	iances				
	Method	Num DF	Den DF	F Value	e Pr	> F		

Folded F

248

2.17

0.0001

78

### Conclusion

Japanese median miles per gallon is 1.55 [95% CI (1.46,1.65)] times as large as US median miles per gallon.