

Amazon Reviews

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STAT 544 - Iowa State University

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Amazon Reviews - Upright, bagless, cyclonic vacuum cleaners

	Number of ratings							
product_id	n1	n2	n3	n4	n5	n_total	mean	sd
B000REMVGK	21	17	2	8	7	55	2.33	1.44
B001EFMD8W	40	34	28	77	347	526	4.25	1.26
B001PB51GQ	14	12	13	31	69	139	3.93	1.36
B002DGSJVG	22	8	3	6	10	49	2.47	1.63
B002G9UQZC	8	0	1	1	1	11	1.82	1.47
B002GHBRX4	18	8	9	14	27	76	3.32	1.61
B002HF66BI	9	5	2	2	3	21	2.29	1.49
B003OA77MC	15	7	8	24	42	96	3.74	1.47
B003OAD24Y	7	7	4	9	19	46	3.57	1.53
B003Y3AA3C	20	3	1	2	2	28	1.68	1.28
B0043EW354	40	25	25	60	163	313	3.90	1.44
B00440EO8G	2	1	1	1	7	12	3.83	1.64
B004R9197I	9	1	1	9	26	46	3.91	1.58
B008L5F4H0	3	1	2	12	7	25	3.76	1.27

Model for Amazon Reviews

Let y_{ij} be the j th review for the i th product.

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Model for Amazon Reviews

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$$y_{ij} \stackrel{\text{ind}}{\sim} N(\theta_i, \sigma^2)$$

and

$$\theta_i \stackrel{\text{ind}}{\sim} N(\mu, \tau^2)$$

and

$$p(\mu, \tau, \sigma) \propto Ca^+(\sigma; 0, 1)Ca^+(\tau; 0, 1)$$

Normal hierarchical model in Stan

```
normal_model = "  
data {  
  int <lower=1> n;  
  int <lower=1> n_products;  
  int <lower=1,upper=5> stars[n];  
  int <lower=1,upper=n_products> product_id[n];  
}  
  
parameters {  
  real mu;                // implied uniform prior  
  real<lower=0> sigma;  
  real<lower=0> tau;  
  real theta[n_products];  
}  
  
model {  
  // Prior  
  sigma ~ cauchy(0,1);  
  tau ~ cauchy(0,1);  
  
  // Hierarchical model  
  theta ~ normal(mu,tau);  
  
  // Data model  
  for (i in 1:n) stars[i] ~ normal(theta[product_id[i]], sigma);  
}  
"
```

Fit model

```
m = stan_model(model_code = normal_model)
```

```
In file included from file24d05f9c1dec.cpp:8:
```

```
In file included from /Library/Frameworks/R.framework/Versions/3.4/Resources/library/StanHeaders/include/src/stan/m
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```
/Library/Frameworks/R.framework/Versions/3.4/Resources/library/BH/include/boost/config/compiler/clang.hpp:200:1
```

```
# define BOOST_NO_CXX11_RVALUE_REFERENCES
```

```
^
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```
<command line>:6:9: note: previous definition is here
```

```
#define BOOST_NO_CXX11_RVALUE_REFERENCES 1
```

```
^
```

```
1 warning generated.
```

```
dat = list(n = nrow(d),
           n_products = nlevels(d$product_id),
           stars = d$stars,
           product_id = as.numeric(d$product_id))
r = sampling(m, dat)
```

```
SAMPLING FOR MODEL '03148bf3617900613206f68b66119d86' NOW (CHAIN 1).
```

```
Gradient evaluation took 0.000253 seconds
```

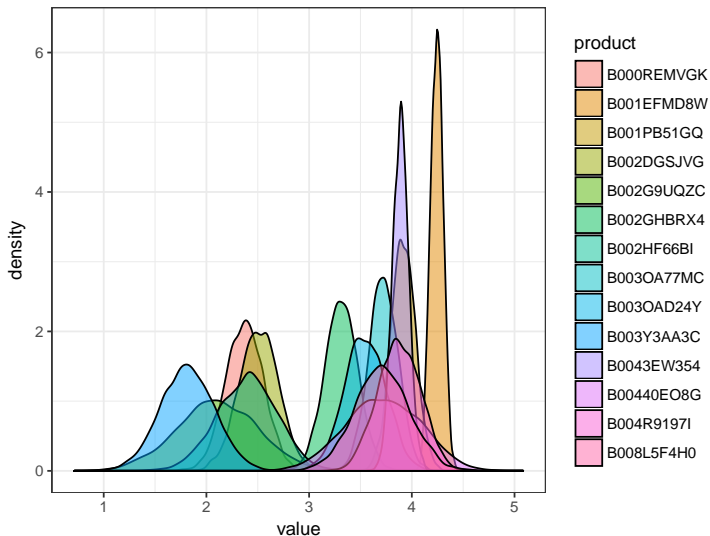

Tabular summary

Inference for Stan model: 03148bf3617900613206f68b66119d86.
 4 chains, each with iter=2000; warmup=1000; thin=1;
 post-warmup draws per chain=1000, total post-warmup draws=4000.

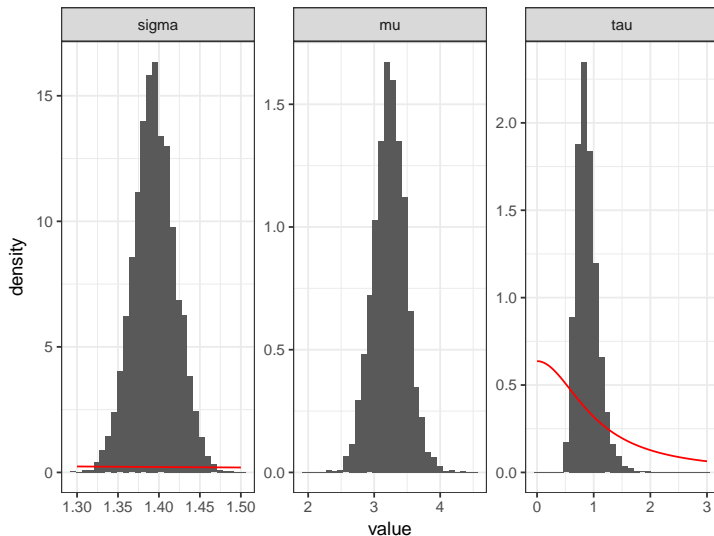
	mean	se_mean	sd	2.5%	25%	50%	75%	97.5%	n_eff	Rhat
mu	3.23	0.00	0.26	2.73	3.07	3.23	3.40	3.73	4000	1
sigma	1.39	0.00	0.03	1.34	1.38	1.39	1.41	1.45	4000	1
tau	0.89	0.00	0.19	0.58	0.75	0.86	0.99	1.34	4000	1
theta[1]	2.37	0.00	0.18	2.02	2.25	2.37	2.49	2.72	4000	1
theta[2]	4.24	0.00	0.06	4.13	4.20	4.25	4.29	4.36	4000	1
theta[3]	3.92	0.00	0.12	3.68	3.84	3.91	3.99	4.15	4000	1
theta[4]	2.51	0.00	0.19	2.14	2.38	2.51	2.64	2.88	4000	1
theta[5]	2.10	0.01	0.39	1.33	1.84	2.10	2.37	2.86	4000	1
theta[6]	3.31	0.00	0.16	3.00	3.21	3.31	3.42	3.63	4000	1
theta[7]	2.40	0.00	0.29	1.82	2.20	2.40	2.59	2.95	4000	1
theta[8]	3.72	0.00	0.14	3.45	3.63	3.72	3.82	4.00	4000	1
theta[9]	3.54	0.00	0.20	3.15	3.41	3.54	3.68	3.93	4000	1
theta[10]	1.81	0.00	0.26	1.30	1.63	1.81	1.99	2.33	4000	1
theta[11]	3.89	0.00	0.08	3.74	3.84	3.89	3.94	4.05	4000	1
theta[12]	3.72	0.01	0.36	3.01	3.47	3.72	3.98	4.42	4000	1
theta[13]	3.88	0.00	0.21	3.47	3.73	3.87	4.02	4.28	4000	1
theta[14]	3.71	0.00	0.27	3.19	3.53	3.71	3.89	4.23	4000	1
lp__	-1207.37	0.07	2.87	-1213.62	-1209.10	-1207.11	-1205.33	-1202.55	1515	1

Samples were drawn using NUTS(diag_e) at Thu Mar 1 09:41:07 2018.
 For each parameter, n_eff is a crude measure of effective sample size,
 and Rhat is the potential scale reduction factor on split chains (at
 convergence, Rhat=1).

Vacuum cleaner mean posteriors (θ_i)



Other parameter posteriors



A quick rating

Suppose a new vacuum cleaner comes on the market and there are two Amazon reviews both with 5 stars.

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Suppose a new vacuum cleaner comes on the market and there are two Amazon reviews both with 5 stars. What do you think the average star rating will be (in the future) for this new product?

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$$E[\theta^* | \bar{y}^*, n^*, \sigma, \mu, \tau] = \frac{\frac{n^*}{\sigma^2}}{\frac{n^*}{\sigma^2} + \frac{1}{\tau^2}} \bar{y}^* + \frac{\frac{1}{\tau^2}}{\frac{n^*}{\sigma^2} + \frac{1}{\tau^2}} \mu$$

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$$\begin{aligned}
 E[\theta^* | \bar{y}^*, n^*, \sigma, \mu, \tau] &= \frac{\frac{n^*}{\sigma^2}}{\frac{n^*}{\sigma^2} + \frac{1}{\tau^2}} \bar{y}^* + \frac{\frac{1}{\tau^2}}{\frac{n^*}{\sigma^2} + \frac{1}{\tau^2}} \mu \\
 &= \frac{n^*}{n^* + \frac{\sigma^2}{\tau^2}} \bar{y}^* + \frac{\frac{\sigma^2}{\tau^2}}{n^* + \frac{\sigma^2}{\tau^2}} \mu \\
 &= \frac{n^*}{n^* + m} \bar{y}^* + \frac{m}{n^* + m} \mu
 \end{aligned}$$

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$$\begin{aligned} E[\theta^* | \bar{y}^*, n^*, \sigma, \mu, \tau] &= \frac{\frac{n^*}{\sigma^2}}{\frac{n^*}{\sigma^2} + \frac{1}{\tau^2}} \bar{y}^* + \frac{\frac{1}{\tau^2}}{\frac{n^*}{\sigma^2} + \frac{1}{\tau^2}} \mu \\ &= \frac{n^*}{n^* + \frac{\sigma^2}{\tau^2}} \bar{y}^* + \frac{\frac{\sigma^2}{\tau^2}}{n^* + \frac{\sigma^2}{\tau^2}} \mu \\ &= \frac{n^*}{n^* + m} \bar{y}^* + \frac{m}{n^* + m} \mu \end{aligned}$$

where $m = \sigma^2/\tau^2$ is a measure of how many *prior* samples there are.

IMDB rating

From <http://www.imdb.com/chart/top.html>:

$$\text{weighted rating (WR)} = (v / (v+m)) R + (m / (v+m)) C$$

Where:

R = average for the movie (mean) = (Rating)

v = number of votes for the movie = (votes)

m = minimum votes required to be listed in the Top 250
(currently 25000)

C = the mean vote across the whole report (currently 7.1)

Thus IMDB uses a Bayesian estimate for the rating for each movie where
 $m = \sigma^2 / \tau^2 = 25,000$.

Clearly incorrect model

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An alternative model is

$$z_{ij} \stackrel{\text{ind}}{\sim} \text{Bin}(4, \theta_i)$$

where $z_{ij} = y_{ij} - 1$ is the j th star rating minus 1 of product i

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The idea behind this model would be that product i the probability of earning each star is θ_i and each star is independent.

Binomial hierarchical model in Stan

```
binomial_model = "  
data {  
  int <lower=1> n;  
  int <lower=1> n_products;  
  int <lower=1,upper=5> stars[n];  
  int <lower=1,upper=n_products> product_id[n];  
}  
  
transformed data {  
  int <lower=0, upper=4> z[n];  
  for (i in 1:n) z[i] = stars[i]-1;  
}  
  
parameters {  
  real<lower=0> alpha;  
  real<lower=0> beta;  
  real<lower=0,upper=1> theta[n_products];  
}  
  
model {  
  // Prior  
  target += -5*log(alpha+beta)/2; // improper prior  
  
  // Hierarchical model  
  theta ~ beta(alpha,beta);  
  
  // Data model  
  for (i in 1:n) z[i] ~ binomial(4, theta[product_id[i]]);  
}  
"
```

Fit model

```
m = stan_model(model_code = binomial_model)
```

```
In file included from file24d05140e485.cpp:8:
```

```
In file included from /Library/Frameworks/R.framework/Versions/3.4/Resources/library/StanHeaders/include/src/stan/m
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```
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```

```
<command line>:6:9: note: previous definition is here
```

```
#define BOOST_NO_CXX11_RVALUE_REFERENCES 1
```

```
^
```

```
1 warning generated.
```

```
dat = list(n = nrow(d),
```

```
      n_products = nlevels(d$product_id),
```

```
      stars = d$stars,
```

```
      product_id = as.numeric(d$product_id))
```

```
r = sampling(m, dat)
```

```
SAMPLING FOR MODEL 'e26b5a276955604814aba1dc21dc3cbe' NOW (CHAIN 1).
```

```
Gradient evaluation took 0.000362 seconds
```

Tabular summary

Inference for Stan model: e26b5a276955604814aba1dc21dc3cbe.

4 chains, each with iter=2000; warmup=1000; thin=1;

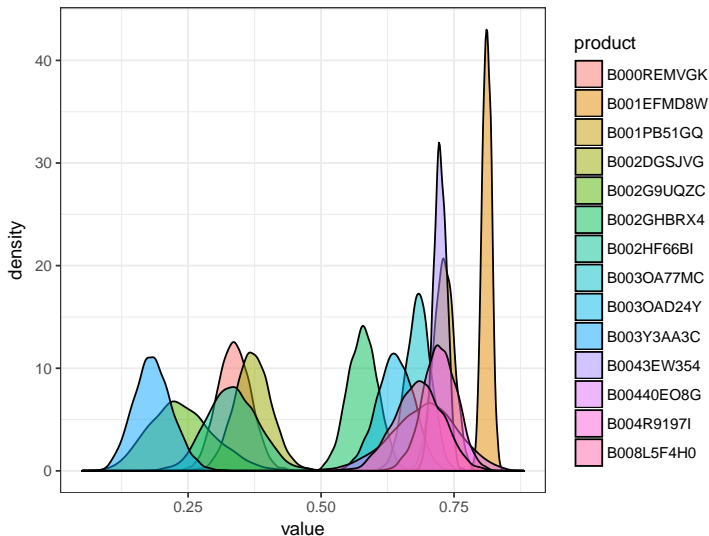
post-warmup draws per chain=1000, total post-warmup draws=4000.

	mean	se_mean	sd	2.5%	25%	50%	75%	97.5%	n_eff	Rhat
alpha	2.70	0.02	1.07	1.07	1.93	2.57	3.29	5.30	4000	1
beta	2.27	0.01	0.87	0.95	1.65	2.13	2.74	4.37	4000	1
theta[1]	0.34	0.00	0.03	0.28	0.32	0.34	0.36	0.40	4000	1
theta[2]	0.81	0.00	0.01	0.79	0.81	0.81	0.82	0.83	4000	1
theta[3]	0.73	0.00	0.02	0.69	0.72	0.73	0.74	0.77	4000	1
theta[4]	0.37	0.00	0.03	0.30	0.35	0.37	0.39	0.44	4000	1
theta[5]	0.24	0.00	0.06	0.13	0.20	0.23	0.28	0.36	4000	1
theta[6]	0.58	0.00	0.03	0.52	0.56	0.58	0.60	0.63	4000	1
theta[7]	0.33	0.00	0.05	0.24	0.30	0.33	0.37	0.43	4000	1
theta[8]	0.68	0.00	0.02	0.64	0.67	0.68	0.70	0.73	4000	1
theta[9]	0.64	0.00	0.03	0.57	0.62	0.64	0.66	0.70	4000	1
theta[10]	0.18	0.00	0.04	0.12	0.16	0.18	0.21	0.26	4000	1
theta[11]	0.72	0.00	0.01	0.70	0.72	0.72	0.73	0.75	4000	1
theta[12]	0.69	0.00	0.06	0.56	0.65	0.70	0.73	0.81	4000	1
theta[13]	0.72	0.00	0.03	0.66	0.70	0.72	0.75	0.79	4000	1
theta[14]	0.68	0.00	0.05	0.59	0.65	0.68	0.71	0.77	4000	1
lp__	-3265.15	0.07	2.86	-3271.42	-3266.93	-3264.79	-3263.04	-3260.59	1632	1

Samples were drawn using NUTS(diag_e) at Thu Mar 1 09:42:54 2018.

For each parameter, `n_eff` is a crude measure of effective sample size, and `Rhat` is the potential scale reduction factor on split chains (at convergence, `Rhat`=1).

Review mean posteriors (θ_i)



Other parameter posteriors

Recall that

- α is the prior success
- β is the prior failures

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- $E[\theta_i | \alpha, \beta] = \frac{\alpha}{\alpha + \beta}$ is the prior expectation for the probability

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$$E[\text{stars}_{*j} | \alpha, \beta] = E[z_{*j} + 1 | \alpha, \beta]$$

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$$E[\text{stars}_{*j} | \alpha, \beta] = E[z_{*j} + 1 | \alpha, \beta] = E[z_{*j} | \alpha, \beta] + 1$$

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$$\begin{aligned} E[\text{stars}_{*j} | \alpha, \beta] &= E[z_{*j} + 1 | \alpha, \beta] = E[z_{*j} | \alpha, \beta] + 1 \\ &= E[E[z_{*j} | \theta^*] | \alpha, \beta] + 1 \end{aligned}$$

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$$\begin{aligned} E[\text{stars}_{*j} | \alpha, \beta] &= E[z_{*j} + 1 | \alpha, \beta] = E[z_{*j} | \alpha, \beta] + 1 \\ &= E[E[z_{*j} | \theta^*] | \alpha, \beta] + 1 = E[4\theta^* | \alpha, \beta] + 1 \end{aligned}$$

Other parameter posteriors

Recall that

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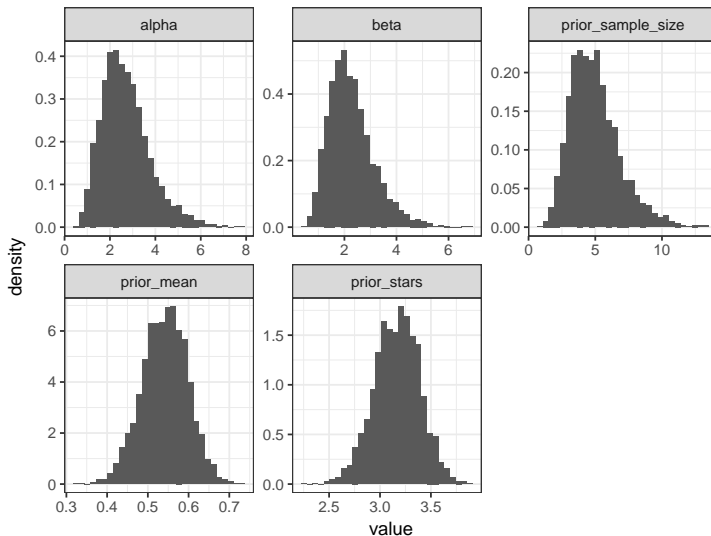
So

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$$\begin{aligned} E[\text{stars}_{*j} | \alpha, \beta] &= E[z_{*j} + 1 | \alpha, \beta] = E[z_{*j} | \alpha, \beta] + 1 \\ &= E[E[z_{*j} | \theta^*] | \alpha, \beta] + 1 = E[4\theta^* | \alpha, \beta] + 1 \\ &= 4 \frac{\alpha}{\alpha + \beta} + 1 \end{aligned}$$

Other parameter posteriors



Uniform use of star ratings

This binomial model has the proper support $\{0, 1, 2, 3, 4\}$ for stars minus 1,

Uniform use of star ratings

This binomial model has the proper support $\{0, 1, 2, 3, 4\}$ for stars minus 1, but does it have the correct proportion of observations in each star category?

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As an example, $\hat{\theta}_2 = 0.81$.

Uniform use of star ratings

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But this ignores the uncertainty in θ_2 (95% CI is (0.79, 0.83)), so perhaps this difference is due to this uncertainty.

Posterior predictive pvalue

To assess this model fit, we will simulate posterior predictive star ratings for product 2 and compare to the observed ratings:

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Thus the following procedure will simulation from the joint distribution for the predictive ratings:

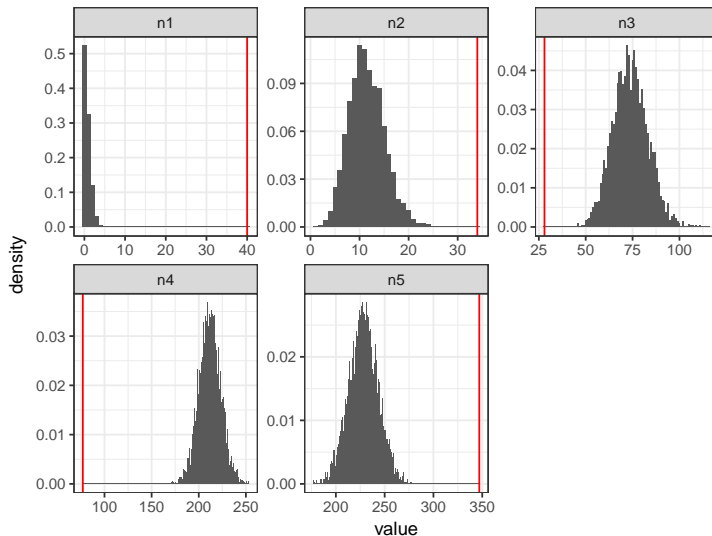
1. $\theta_2 \sim p(\theta_2|z)$,
2. For $j = 1, \dots, 526$, $z_{2j} \stackrel{ind}{\sim} \text{Bin}(4, \theta_2)$, and
3. $\text{star}_{2j} = z_{2j} + 1$.

Posterior predictive distribution in R

```
theta2 = as.numeric(draws$theta[,2])
ztilde2 = adply(theta2, 1, function(x) {
  ztilde = rbinom(526, 4, x) + 1
  data.frame(n1 = sum(ztilde==1),
             n2 = sum(ztilde==2),
             n3 = sum(ztilde==3),
             n4 = sum(ztilde==4),
             n5 = sum(ztilde==5))
})
head(ztilde2)
```

	X1	n1	n2	n3	n4	n5
1	1	0	14	63	208	241
2	2	0	11	70	206	239
3	3	0	16	81	209	220
4	4	0	7	76	187	256
5	5	0	8	89	221	208
6	6	0	8	64	220	234

Posterior predictive distribution in R



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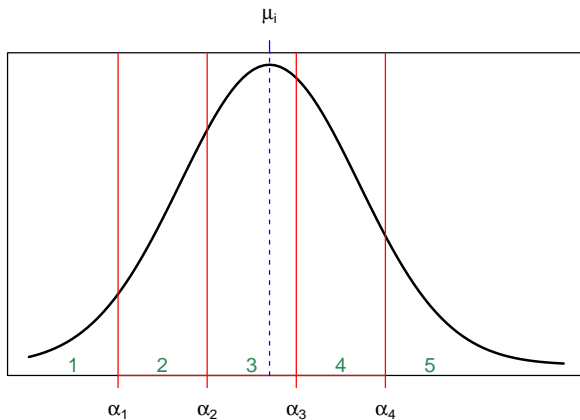
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where $\alpha_0 = -\infty$, $\alpha_1 = 0$, and $\alpha_5 = \infty$, and Φ is the standard normal cumulative distribution function (cdf).

Visualizing the model



Hierarchical model

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with a prior

$$p(\eta, \tau) \propto Ca(\tau; 0, 1).$$

```

ordinal_model = "
data {
  int <lower=1> n_products;
  int <lower=0> s[n_products,5]; // summarized count by product
}

parameters {
  real<lower=0> alpha_diff[3];
  real mu[n_products];
  real eta;
  real<lower=0> tau;
}

transformed parameters {
  ordered[4] alpha;           // cut points
  simplex[5] theta[n_products]; // each theta vector sums to 1

  alpha[1] = 0; for (i in 1:3) alpha[i+1] = alpha[i] + alpha_diff[i];

  for (p in 1:n_products) {
    theta[p,1] = Phi(-mu[p]);
    for (j in 2:4)
      theta[p,j] = Phi(alpha[j]-mu[p]) - Phi(alpha[j-1]-mu[p]);
    theta[p,5] = 1-Phi(alpha[4]-mu[p]);
  }
}

model {
  tau ~ cauchy(0,1);
  mu ~ normal(eta, tau);
  for (p in 1:n_products) s[p] ~ multinomial(theta[p]); // n_reviews[p] is implicit
}
"
```

Fit model

```
m = stan_model(model_code = ordinal_model)
```

```
In file included from file23897f9b2fb4.cpp:8:
```

```
In file included from /Library/Frameworks/R.framework/Versions/3.3/Resources/library/StanHeaders/include/src/stan/m
```

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```
In file included from /Library/Frameworks/R.framework/Versions/3.3/Resources/library/BH/include/boost/math/tool
```

```
In file included from /Library/Frameworks/R.framework/Versions/3.3/Resources/library/BH/include/boost/config.hp
```

```
/Library/Frameworks/R.framework/Versions/3.3/Resources/library/BH/include/boost/config/compiler/clang.hpp:196:1
```

```
# define BOOST_NO_CXX11_RVALUE_REFERENCES
```

```
^
```

```
<command line>:6:9: note: previous definition is here
```

```
#define BOOST_NO_CXX11_RVALUE_REFERENCES 1
```

```
^
```

```
1 warning generated.
```

```
dat = list(n_products = nrow(for_table),
```

```
        s = as.matrix(for_table[,2:6]))
```

```
r = sampling(m, dat, pars = c("alpha", "eta", "tau", "mu"))
```

```
SAMPLING FOR MODEL 'cfd399bb3e758fc22eaf105a07c2068f' NOW (CHAIN 1).
```

```
Chain 1, Iteration: 1 / 2000 [ 0%] (Warmup)
```

```
Chain 1, Iteration: 200 / 2000 [ 10%] (Warmup)
```

```
Chain 1, Iteration: 400 / 2000 [ 20%] (Warmup)
```

Fit model

r

Inference for Stan model: cfd399bb3e758fc22eaf105a07c2068f.

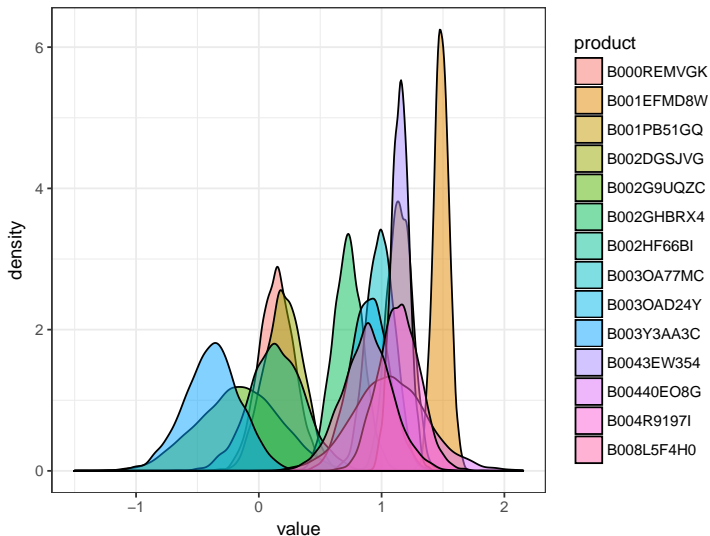
4 chains, each with iter=2000; warmup=1000; thin=1;

post-warmup draws per chain=1000, total post-warmup draws=4000.

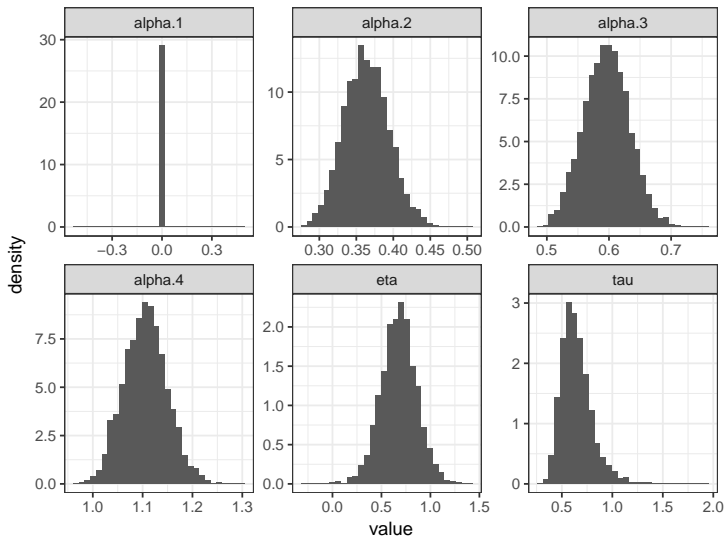
	mean	se_mean	sd	2.5%	25%	50%	75%	97.5%	n_eff	Rhat
alpha[1]	0.00	0.00	0.00	0.00	0.00	0.00	0.00	0.00	4000	NaN
alpha[2]	0.36	0.00	0.03	0.31	0.34	0.36	0.38	0.43	4000	1
alpha[3]	0.60	0.00	0.04	0.53	0.57	0.60	0.62	0.67	4000	1
alpha[4]	1.11	0.00	0.04	1.03	1.08	1.11	1.13	1.19	4000	1
eta	0.68	0.00	0.18	0.31	0.56	0.68	0.80	1.04	4000	1
tau	0.65	0.00	0.15	0.42	0.54	0.62	0.73	1.02	4000	1
mu[1]	0.15	0.00	0.14	-0.13	0.05	0.15	0.24	0.43	4000	1
mu[2]	1.49	0.00	0.06	1.36	1.45	1.49	1.53	1.61	4000	1
mu[3]	1.15	0.00	0.10	0.95	1.08	1.15	1.22	1.33	4000	1
mu[4]	0.20	0.00	0.16	-0.10	0.10	0.20	0.31	0.50	4000	1
mu[5]	-0.17	0.01	0.33	-0.81	-0.39	-0.16	0.06	0.46	4000	1
mu[6]	0.73	0.00	0.12	0.48	0.64	0.73	0.81	0.97	4000	1
mu[7]	0.15	0.00	0.22	-0.30	0.00	0.14	0.30	0.59	4000	1
mu[8]	0.99	0.00	0.11	0.77	0.91	0.99	1.07	1.21	4000	1
mu[9]	0.90	0.00	0.16	0.59	0.79	0.90	1.00	1.22	4000	1
mu[10]	-0.38	0.00	0.23	-0.84	-0.53	-0.37	-0.23	0.06	4000	1
mu[11]	1.15	0.00	0.07	1.01	1.10	1.15	1.20	1.29	4000	1
mu[12]	1.07	0.00	0.30	0.48	0.87	1.06	1.26	1.69	4000	1
mu[13]	1.14	0.00	0.17	0.83	1.03	1.14	1.26	1.47	4000	1
mu[14]	0.88	0.00	0.21	0.47	0.75	0.89	1.02	1.30	4000	1
lp__	-1835.69	0.09	3.13	-1842.60	-1837.59	-1835.37	-1833.46	-1830.53	1206	1

Samples were drawn using NUTS(diag_e) at Thu Feb 23 11:33:47 2017.

Review mean posteriors (θ_i)



Other parameter posteriors



Visualizing the model

