

STAT 401A - Statistical Methods for Research Workers

Multiple regression models

Jarad Niemi (Dr. J)

Iowa State University

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Multiple regression

Recall the simple linear regression model is

$$Y_i \stackrel{\text{iid}}{\sim} N(\beta_0 + \beta_1 X_i, \sigma^2)$$

The **multiple regression model** is

$$Y_i \stackrel{\text{iid}}{\sim} N(\beta_0 + \beta_1 X_{i,1} + \cdots + \beta_p X_{i,p}, \sigma^2)$$

where

- Y_i is the response for observation i and
- $X_{i,p}$ is the p^{th} explanatory variable for observation i .

We may also write

$$Y_i \stackrel{\text{iid}}{\sim} N(\mu_i, \sigma^2) \quad \text{or} \quad Y_i = \mu_i + e_i, e_i \stackrel{\text{iid}}{\sim} N(0, \sigma^2)$$

where

$$\mu_i = \beta_0 + \beta_1 X_{i,1} + \cdots + \beta_p X_{i,p}.$$

Explanatory variables

There is a lot of flexibility in the mean

$$\mu_i = E[Y_i | X_{i,1}, \dots, X_{i,p}] = \beta_0 + \beta_1 X_{i,1} + \dots + \beta_p X_{i,p}$$

as there are many possibilities for the explanatory variables $X_{i,1}, \dots, X_{i,p}$:

- Higher order terms (X^2)
- Additional explanatory variables (X_1 and X_2)
- Dummy/indicator variables for categorical variables ($X_1 = I()$)
- Interactions ($X_1 X_2$)
 - Continuous-continuous
 - Continuous-categorical
 - Categorical-categorical

Interpretation

Model:

$$Y_i \stackrel{ind}{\sim} N(\beta_0 + \beta_1 X_{i,1} + \cdots + \beta_p X_{i,p}, \sigma^2)$$

The interpretation is

- β_0 is the expected value of the response Y_i when **all** explanatory variables are zero.
- β_p , $p \neq 0$ is the expected increase in the response for a one-unit increase in the p^{th} explanatory variable **when all other explanatory variables are held constant**.
- R^2 is the proportion of the variance in the response explained by the model

Higher order terms (X^2)

Let

- Y_i be the distance for the i^{th} run of the experiment and
- H_i be the height for the i^{th} run of the experiment.

Simple linear regression assumes

$$Y_i \stackrel{ind}{\sim} N(\beta_0 + \beta_1 H_i, \sigma^2)$$

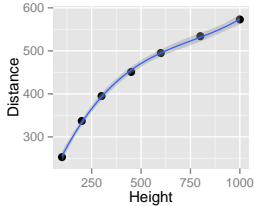
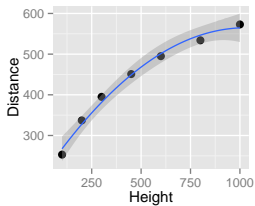
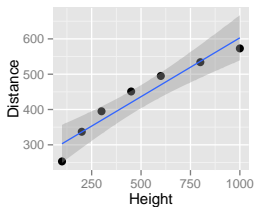
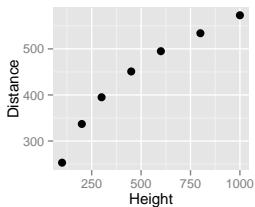
The quadratic multiple regression assumes

$$Y_i \stackrel{ind}{\sim} N(\beta_0 + \beta_1 H_i + \beta_2 H_i^2, \sigma^2)$$

The cubic multiple regression assumes

$$Y_i \stackrel{ind}{\sim} N(\beta_0 + \beta_1 H_i + \beta_2 H_i^2 + \beta_3 H_i^3, \sigma^2)$$

Case1001



SAS code and output

```
DATA case1001;
  INFILE 'case1001.csv' DSD FIRSTOBS=2;
  INPUT distance height;
  height2 = height*height;
  height3 = height*height2;

# PROC REG allows multiple MODEL statements
PROC REG DATA=case1001;
  MODEL distance = height;
  MODEL distance = height height2;
  MODEL distance = height height2 height3;
RUN;
```

Parameter Estimates

Variable	DF	Parameter Estimate	Standard Error	t Value	Pr > t
Intercept	1	269.71246	24.31239	11.09	0.0001
height	1	0.33334	0.04203	7.93	0.0005
Intercept	1	199.91282	16.75945	11.93	0.0003
height	1	0.70832	0.07482	9.47	0.0007
height2	1	-0.00034369	0.00006678	-5.15	0.0068
Intercept	1	155.77551	8.32579	18.71	0.0003
height	1	1.11530	0.06567	16.98	0.0004
height2	1	-0.00124	0.00013842	-8.99	0.0029
height3	1	5.477104E-7	8.327329E-8	6.58	0.0072

SAS code and output

```
DATA case1001;
  INFILE 'case1001.csv' DSD FIRSTOBS=2;
  INPUT distance height;
  height2 = height ** 2;
  height3 = height ** 3;

PROC GLM DATA=case1001;
  MODEL distance = height height2 height3;

/* PROC GLM allows the variable construction within the MODEL statement
   and provides nicer output (not shown here) */
DATA case1001;
  INFILE 'case1001.csv' DSD FIRSTOBS=2;
  INPUT distance height;

/* This shorthand puts in H, H^2, and H^3 */
PROC GLM DATA=case1001;
  MODEL distance = height|height|height;

/* This only puts H^3 */
PROC GLM DATA=case1001;
  MODEL distance = height*height*height;
```


R code and output

```
# Construct the variables by hand
case1001$Height2 = case1001$Height^2
case1001$Height3 = case1001$Height^3

m1 = lm(Distance~Height,          case1001)
m2 = lm(Distance~Height+Height2,  case1001)
m3 = lm(Distance~Height+Height2+Height3, case1001)

coefficients(m1)

(Intercept)      Height
    269.7125      0.3333

coefficients(m2)

(Intercept)      Height      Height2
    1.999e+02      7.083e-01     -3.437e-04

coefficients(m3)

(Intercept)      Height      Height2      Height3
    1.558e+02      1.115e+00     -1.245e-03      5.477e-07
```

R code and output

```
# Let R construct the variables for you
m = lm(Distance~poly(Hight, 3, raw=TRUE), case1001)
summary(m)
```

Call:

```
lm(formula = Distance ~ poly(Hight, 3, raw = TRUE), data = case1001)
```

Residuals:

```
      1      2      3      4      5      6      7
-2.4036  3.5809  1.8917 -4.4688 -0.0804  2.3216 -0.8414
```

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)	
(Intercept)	1.56e+02	8.33e+00	18.71	0.00033	***
poly(Hight, 3, raw = TRUE)1	1.12e+00	6.57e-02	16.98	0.00044	***
poly(Hight, 3, raw = TRUE)2	-1.24e-03	1.38e-04	-8.99	0.00290	**
poly(Hight, 3, raw = TRUE)3	5.48e-07	8.33e-08	6.58	0.00715	**

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 4.01 on 3 degrees of freedom

Multiple R-squared: 0.999, Adjusted R-squared: 0.999

F-statistic: 1.6e+03 on 3 and 3 DF, p-value: 2.66e-05

Longnose Dace Abundance

From <http://udel.edu/~mcdonald/statmultreg.html>:

*I extracted some data from the Maryland Biological Stream Survey. ... The dependent variable is the number of Longnose Dace (*Rhinichthys cataractae*) per 75-meter section of [a] stream. The independent variables are the area (in acres) drained by the stream; the dissolved oxygen (in mg/liter); the maximum depth (in cm) of the 75-meter segment of stream; nitrate concentration (mg/liter); sulfate concentration (mg/liter); and the water temperature on the sampling date (in degrees C).*

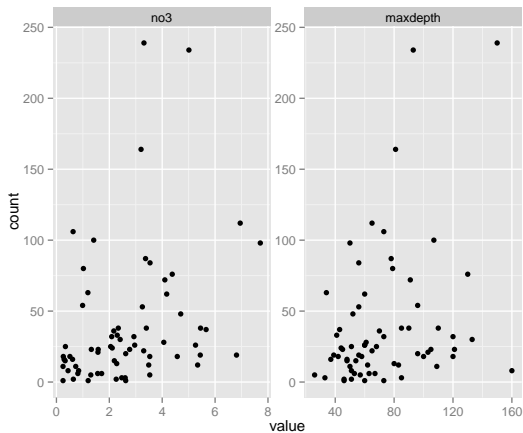
Consider the model

$$Y_i \stackrel{ind}{\sim} N(\beta_0 + \beta_1 X_{i,1} + \beta_2 X_{i,2}, \sigma^2)$$

where

- Y_i : count of Longnose Dace in stream i
- $X_{i,1}$: maximum depth (in cm) of stream i
- $X_{i,2}$: nitrate concentration (mg/liter) of stream i

Exploratory



```

DATA dace;
  INFILE 'Longnose Dace.csv' DSD FIRSTOBS=2;
  INPUT stream $ count acreage do2 maxdepth no3 so4 temp;

PROC REG DATA=dace;
  MODEL count = maxdepth no3;
  RUN;

```

The REG Procedure

Model: MODEL1

Dependent Variable: count

Number of Observations Read	67
Number of Observations Used	67

Analysis of Variance

Source	DF	Sum of Squares	Mean Square	F Value	Pr > F
Model	2	28930	14465	7.68	0.0010
Error	64	120503	1882.85220		
Corrected Total	66	149432			

Root MSE	43.39184	R-Square	0.1936
Dependent Mean	39.10448	Adj R-Sq	0.1684
Coeff Var	110.96388		

Parameter Estimates

Variable	DF	Parameter Estimate	Standard Error	t Value	Pr > t
Intercept	1	-17.55503	15.95865	-1.10	0.2754
maxdepth	1	0.48106	0.18111	2.66	0.0100
no3	1	8.28473	2.95659	2.80	0.0067

R code and output

```
d = read.csv("longnosedace.csv")
m = lm(count~no3+maxdepth,d)
summary(m)
```

```
Call:
lm(formula = count ~ no3 + maxdepth, data = d)
```

Residuals:

Min	1Q	Median	3Q	Max
-55.06	-27.70	-8.68	11.79	165.31

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	-17.555	15.959	-1.10	0.2754
no3	8.285	2.957	2.80	0.0067 **
maxdepth	0.481	0.181	2.66	0.0100 **

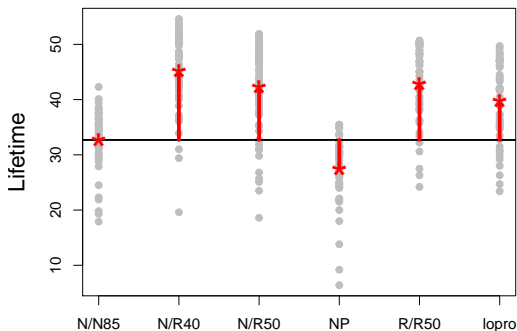
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 43.4 on 64 degrees of freedom
 Multiple R-squared: 0.194, Adjusted R-squared: 0.168
 F-statistic: 7.68 on 2 and 64 DF, p-value: 0.00102

Interpretation

- Intercept (β_0): The expected count of Longnose Dace when maximum depth and nitrate concentration are both zero is -18.
- Coefficient for maxdepth (β_1): Holding nitrate concentration constant, each cm increase in maximum depth is associated with an additional 0.48 Longnose Dace counted on average.
- Coefficient for no3 (β_2): Holding maximum depth constant, each mg/liter increase in nitrate concentration is associated with an addition 8.3 Longnose Dace counted on average.
- Coefficient of determination (R^2): The model explains 19% of the variability in the count of Longnose Dace.

Using a categorical variable as an explanatory variable.



Regression with a categorical variable

- Choose one of the levels as the **reference** level, e.g. N/N85
- Construct dummy variables using indicator functions for the other levels, e.g.

$$X_{i,1} = I(\text{diet for observation } i \text{ is NP})$$

$$X_{i,2} = I(\text{diet for observation } i \text{ is N/R50 lopro})$$

$$X_{i,3} = I(\text{diet for observation } i \text{ is N/R50})$$

$$X_{i,4} = I(\text{diet for observation } i \text{ is R/R50})$$

$$X_{i,5} = I(\text{diet for observation } i \text{ is N/R40})$$

- Run a multiple linear regression using these dummy variables.

An indicator function is

$$I(A) = \begin{cases} 1 & A \text{ is TRUE} \\ 0 & A \text{ is FALSE} \end{cases}$$

```
DATA case0501;
  INFILE 'case0501.csv' DSD FIRSTOBS=2;
  INPUT lifetime diet $;
  IF diet = 'NP' THEN x1=1; ELSE x1=0;
  IF diet = 'lopro' THEN x2=1; ELSE x2=0;
  IF diet = 'N/R50' THEN x3=1; ELSE x3=0;
  IF diet = 'R/R50' THEN x4=1; ELSE x4=0;
  IF diet = 'N/R40' THEN x5=1; ELSE x5=0;
  RUN;

PROC GLM DATA=case0501;
  MODEL lifetime = x1 x2 x3 x4 x5;
  RUN;
  ods listing close;
```

The GLM Procedure

Dependent Variable: lifetime

Source	DF	Sum of Squares	Mean Square	F Value	Pr > F
Model	5	12733.94181	2546.78836	57.10	<.0001
Error	343	15297.41532	44.59888		
Corrected Total	348	28031.35713			

R-Square	Coeff Var	Root MSE	lifetime Mean
0.454275	17.21323	6.678239	38.79713

Parameter	Estimate	Standard Error	t Value	Pr > t
Intercept	32.69122807	0.88455439	36.96	<.0001
x1	-5.28918725	1.30100640	-4.07	<.0001
x2	6.99448622	1.25652099	5.57	<.0001
x3	9.60595503	1.18768248	8.09	<.0001
x4	10.19448622	1.25652099	8.11	<.0001
x5	12.42543860	1.23521298	10.06	<.0001

SAS code and output

```
DATA case0501;
  INFILE 'case0501.csv' DSD FIRSTOBS=2;
  INPUT lifetime diet $;

PROC GLM DATA=case0501;
  CLASS diet(REF='N/N85'); /* by default, SAS uses the alphabetically last group as the reference level */
  MODEL lifetime=diet / SOLUTION;
RUN;
```

The GLM Procedure

Dependent Variable: lifetime

Source	DF	Sum of Squares	Mean Square	F Value	Pr > F
Model	5	12733.94181	2546.78836	57.10	<.0001
Error	343	15297.41532	44.59888		
Corrected Total	348	28031.35713			

R-Square	Coeff Var	Root MSE	lifetime Mean
0.454275	17.21323	6.678239	38.79713

Parameter	Estimate	Standard Error	t Value	Pr > t
Intercept	32.69122807 B	0.88455439	36.96	<.0001
diet N/R40	12.42543860 B	1.23521298	10.06	<.0001
diet N/R50	9.60595503 B	1.18768248	8.09	<.0001
diet NP	-5.28918725 B	1.30100640	-4.07	<.0001
diet R/R50	10.19448622 B	1.25652099	8.11	<.0001
diet lopro	6.99448622 B	1.25652099	5.57	<.0001
diet N/N85	0.00000000 B	.	.	.

NOTE: The X'X matrix has been found to be singular, and a generalized inverse was used to solve the

R code and output

```
# by default, R uses the alphabetically first group as the reference level
case0501$Diet = relevel(case0501$Diet, ref='N/N85')
```

```
m = lm(Lifetime~Diet, case0501)
summary(m)
```

Call:

```
lm(formula = Lifetime ~ Diet, data = case0501)
```

Residuals:

	Min	1Q	Median	3Q	Max
	-25.517	-3.386	0.814	5.183	10.014

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	32.691	0.885	36.96	< 2e-16 ***
DietN/R40	12.425	1.235	10.06	< 2e-16 ***
DietN/R50	9.606	1.188	8.09	1.1e-14 ***
DietNP	-5.289	1.301	-4.07	5.9e-05 ***
DietR/R50	10.194	1.257	8.11	8.9e-15 ***
Dietlopro	6.994	1.257	5.57	5.2e-08 ***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 6.68 on 343 degrees of freedom

Multiple R-squared: 0.454, Adjusted R-squared: 0.446

F-statistic: 57.1 on 5 and 343 DF, p-value: <2e-16

Interpretation

- $\beta_0 = E[Y_i | \text{reference level}]$, i.e. expected response for the reference level

Note: the only way $X_{i,1} = \dots = X_{i,p} = 0$ is if all indicators are zero, i.e. at the reference level.

- $\beta_p, p > 0$: expected change in the response moving from the reference level to the level associated with the p^{th} dummy variable

Note: the only way for $X_{i,p}$ to increase by one and all other indicators to stay constant is if initially $X_{i,1} = \dots = X_{i,p} = 0$ and now $X_{i,p} = 1$

For example,

- The expected lifetime for mice on the N/N85 diet is 37 weeks.
- The expected increase in lifetime for mice on the N/R40 diet compared to the N/N85 diet is 12.4 weeks.
- The model explains 45% of the variability in mice lifetimes.

Why an interaction?

*Two explanatory variables are said to **interact** if the effect that one of them has on the mean response depends on the value of the other.*

For example,

- Longnose dace: The effect of no3 on longnose dace count depends on the maxdepth. (Continuous-continuous)
- Case1002: The effect of mass on energy depends on the species type. (Continuous-categorical)
- Yield: the effect of fertilizer depends on the block (Categorical-categorical)

Continuous-continuous interaction

For observation i , let

- Y_i be the response
- $X_{i,1}$ be the first explanatory variable and
- $X_{i,2}$ be the second explanatory variable.

The mean containing only **main effects** is

$$\mu_i = \beta_0 + \beta_1 X_{i,1} + \beta_2 X_{i,2}.$$

The mean with the **interaction** is

$$\mu_i = \beta_0 + \beta_1 X_{i,1} + \beta_2 X_{i,2} + \beta_3 X_{i,1} X_{i,2}.$$

Interpretation

- β_0 : expected response when explanatory variables are zero
- β_1 : expected change in the response for each change in $X_{i,1}$ when $X_{i,2}$ is zero
- β_2 : expected change in the response for each change in $X_{i,2}$ when $X_{i,1}$ is zero
- $\beta_3 X_{i,2}$: expected change in the effect of $X_{i,1}$ on the response when $X_{i,2}$ is not zero
- $\beta_3 X_{i,1}$: expected change in the effect of $X_{i,2}$ on the response when $X_{i,1}$ is not zero

For example,

$$\begin{aligned}
 E[Y_i | X_{i,1} = x_1 + 1, X_{i,2} = x_2] &= \beta_0 + \beta_1(x_1 + 1) + \beta_2 x_2 + \beta_3(x_1 + 1)x_2 \\
 E[Y_i | X_{i,1} = x_1, X_{i,2} = x_2] &= \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \beta_3 x_1 x_2 \\
 \hline
 &= \beta_1 + \beta_3 x_2
 \end{aligned}$$

R code and output

```
d = read.csv("longnosedace.csv")
m = lm(count ~ maxdepth + no3 + maxdepth:no3, d)
summary(m)
```

Call:

```
lm(formula = count ~ maxdepth + no3 + maxdepth:no3, data = d)
```

Residuals:

Min	1Q	Median	3Q	Max
-65.11	-21.40	-9.56	5.95	151.07

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	13.32104	23.45571	0.57	0.572
maxdepth	-0.00934	0.32918	-0.03	0.977
no3	-4.64627	7.85693	-0.59	0.556
maxdepth:no3	0.20122	0.11358	1.77	0.081

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 42.7 on 63 degrees of freedom

Multiple R-squared: 0.232, Adjusted R-squared: 0.195

F-statistic: 6.34 on 3 and 63 DF, p-value: 0.000797

Continuous-categorical interaction

Let category A be the reference level. Then observation i , let

- Y_i be the response
- $X_{i,1}$ be the continuous explanatory variable,
- B_i be a dummy variable for category B, and
- C_i be a dummy variable for category C.

The mean containing only **main effects** is

$$\mu_i = \beta_0 + \beta_1 X_{i,1} + \beta_2 B_i + \beta_3 C_i.$$

The mean with the **interaction** is

$$\mu_i = \beta_0 + \beta_1 X_{i,1} + \beta_2 B_i + \beta_3 C_i + \beta_4 X_{i,1} B_i + \beta_5 X_{i,1} C_i.$$

Interpretation for the main effect model

The mean containing only **main effects** is

$$\mu_i = \beta_0 + \beta_1 X_{i,1} + \beta_2 B_i + \beta_3 C_i.$$

- β_1 is the change in the response when the continuous explanatory variable increases by one for any category
- when the continuous explanatory variable is zero
 - β_0 is the expected response for category A
 - $\beta_0 + \beta_2$ is the expected response for category B
 - $\beta_0 + \beta_3$ is the expected response for category C

Interpretation for the model with an interaction

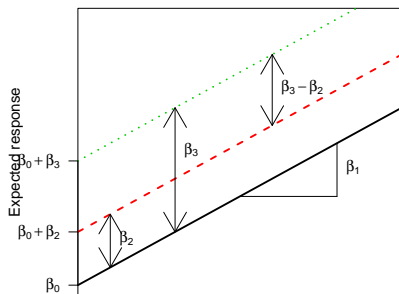
The model with an **interaction** is

$$\mu_i = \beta_0 + \beta_1 X_{i,1} + \beta_2 B_i + \beta_3 C_i + \beta_4 X_{i,1} B_i + \beta_5 X_{i,1} C_i$$

- β_0 is the expected response when the continuous explanatory variable is zero and we are in category A
- β_1 is the change in the response when the continuous explanatory variable increases by one for category A
- When the continuous explanatory variable ($X_{i,1}$) is zero,
 - β_2 is the change in the expected response when moving from category A to category B
 - β_3 is the change in the expected response when moving from category A to category C
- $\beta_1 + \beta_4$ is the change in the response when the continuous explanatory variable increases by one for category B
- $\beta_1 + \beta_5$ is the change in the response when the continuous explanatory variable increases by one for category C

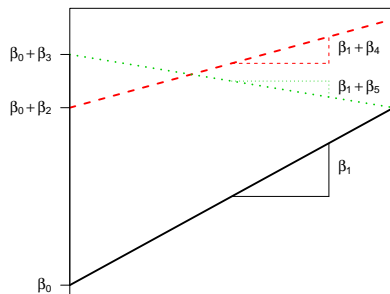
Visualizing the models

Main effects only



Continuous explanatory variable

With interactions



Continuous explanatory variable

R code and output - Main effects only

```
summary(lm(log(Energy)~log(Mass)+Type, case1002))
```

Call:

```
lm(formula = log(Energy) ~ log(Mass) + Type, data = case1002)
```

Residuals:

	Min	1Q	Median	3Q	Max
	-0.2322	-0.1220	-0.0364	0.1257	0.3446

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)	
(Intercept)	-1.4977	0.1499	-9.99	2.8e-08	***
log(Mass)	0.8150	0.0445	18.30	3.8e-12	***
Type _{non-echolocating bats}	-0.0787	0.2027	-0.39	0.70	
Type _{non-echolocating birds}	0.0236	0.1576	0.15	0.88	

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 0.186 on 16 degrees of freedom

Multiple R-squared: 0.982, Adjusted R-squared: 0.978

F-statistic: 284 on 3 and 16 DF, p-value: 4.46e-14

R code and output - with interaction

```
summary(lm(log(Energy)~log(Mass)*Type, case1002))
```

Call:

```
lm(formula = log(Energy) ~ log(Mass) * Type, data = case1002)
```

Residuals:

	Min	1Q	Median	3Q	Max
	-0.2515	-0.1264	-0.0095	0.0812	0.3284

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	-1.4705	0.2477	-5.94	3.6e-05 ***
log(Mass)	0.8047	0.0867	9.28	2.3e-07 ***
Typenon-echolocating bats	1.2681	1.2854	0.99	0.34
Typenon-echolocating birds	-0.1103	0.3847	-0.29	0.78
log(Mass):Typenon-echolocating bats	-0.2149	0.2236	-0.96	0.35
log(Mass):Typenon-echolocating birds	0.0307	0.1028	0.30	0.77

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 0.19 on 14 degrees of freedom

Multiple R-squared: 0.983, Adjusted R-squared: 0.977

F-statistic: 163 on 5 and 14 DF, p-value: 6.7e-12

Categorical-categorical

Let category A and type 0 be the reference level. For observation i , let

- Y_i be the response,
- 1_i be a dummy variable for type 1,
- B_i be a dummy variable for category B, and
- C_i be a dummy variable for category C.

The mean containing only main effects is

$$\mu_i = \beta_0 + \beta_1 1_i + \beta_2 B_i + \beta_3 C_i.$$

The mean with an interaction is

$$\mu_i = \beta_0 + \beta_1 1_i + \beta_2 B_i + \beta_3 C_i + \beta_4 1_i B_i + \beta_5 1_i C_i.$$

Interpretation for the main effects model

The mean containing only main effects is

$$\mu_i = \beta_0 + \beta_1 1_i + \beta_2 B_i + \beta_3 C_i.$$

- β_0 is the expected response for category A and type 0
- β_1 is the change in response for moving from type 0 to type 1
- β_2 is the change in response for moving from category A to category B
- β_3 is the change in response for moving from category A to category C

The means are then

Type	Category		
	A	B	C
0	β_0	$\beta_0 + \beta_2$	$\beta_0 + \beta_3$
1	$\beta_0 + \beta_1$	$\beta_0 + \beta_1 + \beta_2$	$\beta_0 + \beta_1 + \beta_3$

Interpretation for the model with an interaction

The mean with an interaction is

$$\mu_i = \beta_0 + \beta_1 1_i + \beta_2 B_i + \beta_3 C_i + \beta_4 1_i B_i + \beta_5 1_i C_i.$$

- β_0 is the expected response for category A and type 0
- β_1 is the change in response for moving from type 0 to type 1 for category A
- β_2 is the change in response for moving from category A to category B for type 0
- β_3 is the change in response for moving from category A to category C for type 0
- β_4 is the difference in change in response for moving from category A to category B for type 1 compared to type 0
- β_5 is the difference in change in response for moving from category A to category C for type 1 compared to type 0

The means are then

Type	Category				
	A	B		C	
0	β_0	$\beta_0 + \beta_2$		$\beta_0 + \beta_3$	
1	$\beta_0 + \beta_1$	$\beta_0 + \beta_1 + \beta_2 + \beta_4$		$\beta_0 + \beta_1 + \beta_3 + \beta_5$	

This is referred to as the **cell-means model**.

R code and output - main effects only

```
summary(lm(Cover~Block+Treat, case1301, subset=Block %in% c("B1","B2") & Treat %in% c("L","Lf","LfF")))
```

Call:

```
lm(formula = Cover ~ Block + Treat, data = case1301, subset = Block %in%  
  c("B1", "B2") & Treat %in% c("L", "Lf", "LfF"))
```

Residuals:

	Min	1Q	Median	3Q	Max
	-2.333	-0.667	0.000	0.792	1.833

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	4.667	0.768	6.07	0.0003 ***
BlockB2	2.167	0.768	2.82	0.0225 *
TreatLf	-1.500	0.941	-1.59	0.1496
TreatLfF	-3.000	0.941	-3.19	0.0128 *

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 1.33 on 8 degrees of freedom

Multiple R-squared: 0.694, Adjusted R-squared: 0.579

F-statistic: 6.04 on 3 and 8 DF, p-value: 0.0188

R code and output - with interaction

```
summary(lm(Cover~Block*Treat, case1301, subset=Block %in% c("B1","B2") & Treat %in% c("L","Lf","LfF")))
```

Call:

```
lm(formula = Cover ~ Block * Treat, data = case1301, subset = Block %in%
    c("B1", "B2") & Treat %in% c("L", "Lf", "LfF"))
```

Residuals:

Min	1Q	Median	3Q	Max
-1.500	-0.625	0.000	0.625	1.500

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	4.00e+00	8.90e-01	4.50	0.0041 **
BlockB2	3.50e+00	1.26e+00	2.78	0.0319 *
TreatLf	-2.72e-16	1.26e+00	0.00	1.0000
TreatLfF	-2.50e+00	1.26e+00	-1.99	0.0941 .
BlockB2:TreatLf	-3.00e+00	1.78e+00	-1.69	0.1428
BlockB2:TreatLfF	-1.00e+00	1.78e+00	-0.56	0.5945

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 1.26 on 6 degrees of freedom

Multiple R-squared: 0.795, Adjusted R-squared: 0.623

F-statistic: 4.64 on 5 and 6 DF, p-value: 0.0443

When to include interaction terms

From _The Statistical Sleuth (3rd ed)_ page 250:

- when a question of interest pertains to an interaction
- when good reason exists to suspect interaction or
- when interactions are proposed as a more general model for the purpose of examining the goodness of fit of a model without interaction