

## Set S03 - Random effects

STAT 401 (Engineering) - Iowa State University

April 21, 2017

# Regression models

For continuous  $Y_i$ , we have linear regression

$$Y_i \stackrel{ind}{\sim} N(\mu_i, \sigma^2), \quad \mu_i = \beta_0 + \beta_1 X_{i,1} + \cdots + \beta_p X_{i,p}$$

For binary or count with an upper maximum  $Y_i$ , we have logistic regression

$$Y_i \stackrel{ind}{\sim} \text{Bin}(n_i, \theta_i), \quad \log\left(\frac{\theta_i}{1 - \theta_i}\right) = \beta_0 + \beta_1 X_{i,1} + \cdots + \beta_p X_{i,p}$$

For count data with no upper maximum, we have Poisson regression

$$Y_i \stackrel{ind}{\sim} \text{Po}(\lambda_i), \quad \log(\lambda_i) = \beta_0 + \beta_1 X_{i,1} + \cdots + \beta_p X_{i,p}$$

But what if our observations cannot reasonably be assumed to be independent given these explanatory variables?

# Random effect model

Suppose we have continuous observations  $Y_{ij}$  for individual  $i$  from group  $j$ . A random effects model (with a common variance) assumes

$$Y_{ij} = \mu + \alpha_j + \epsilon_{ij} \quad \epsilon_{ij} \stackrel{\text{ind}}{\sim} N(0, \sigma_\epsilon^2)$$

and, to make the  $\alpha_i$  random effects, independent of  $\epsilon_{ij}$

$$\alpha_j \stackrel{\text{ind}}{\sim} N(0, \sigma_\alpha^2).$$

This makes observations within the group correlated since

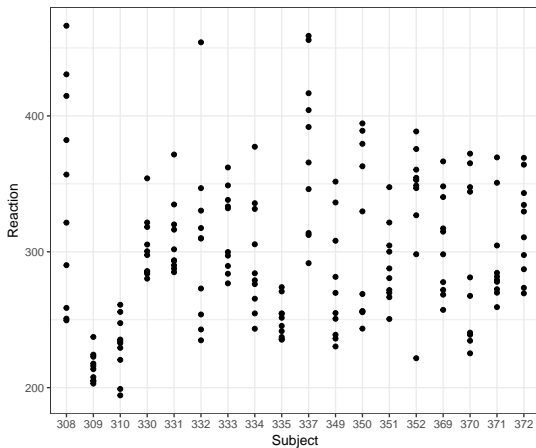
$$\begin{aligned} \text{Cov}[Y_{ij}, Y_{i'j}] &= \text{Cov}[\alpha_j + \epsilon_{ij}, \alpha_j + \epsilon_{i'j}] \\ &= \text{Var}[\alpha_j] = \sigma_\alpha^2 \end{aligned}$$

and

$$\text{Cor}[Y_{ij}, Y_{i'j}] = \frac{\text{Cov}[Y_{ij}, Y_{i'j}]}{\sqrt{\text{Var}[Y_{ij}] \text{Var}[Y_{i'j}]}} = \frac{\sigma_\alpha^2}{\sigma_\alpha^2 + \sigma_\epsilon^2}.$$

# Sleep study example

```
ggplot(sleepstudy, aes(Subject, Reaction)) + geom_point() + theme_bw()
```



# Sleep study example

```
summary(me <- lmer(Reaction ~ (1|Subject), sleepstudy))
```

Linear mixed model fit by REML ['lmerMod']

Formula: Reaction ~ (1 | Subject)

Data: sleepstudy

REML criterion at convergence: 1904.3

Scaled residuals:

Min	1Q	Median	3Q	Max
-2.4983	-0.5501	-0.1476	0.5123	3.3446

Random effects:

Groups	Name	Variance	Std.Dev.
Subject	(Intercept)	1278	35.75
Residual		1959	44.26

Number of obs: 180, groups: Subject, 18

Fixed effects:

	Estimate	Std. Error	t value
(Intercept)	298.51	9.05	32.98

# Mixed effect model

Suppose we have continuous observations  $Y_{ij}$  for individual  $i$  from group  $j$  and an associated explanatory variable  $X_{ij}$ . A mixed effect model assumes

$$Y_{ij} = \beta_0 + \beta_1 X_{ij} + \alpha_j + \epsilon_{ij} \quad \epsilon_{ij} \stackrel{ind}{\sim} N(0, \sigma_\epsilon^2)$$

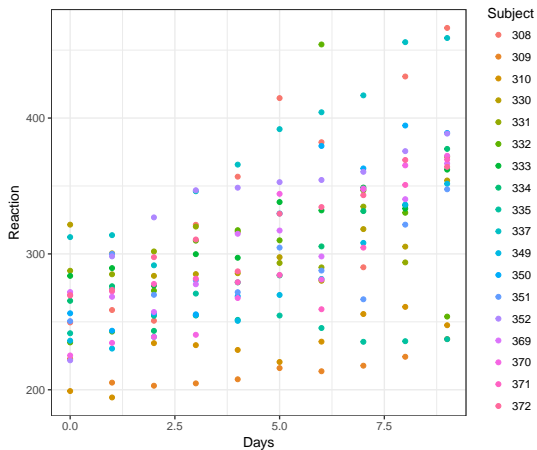
and, to make the  $\alpha_i$  random effects, independent of  $\epsilon_{ij}$

$$\alpha_j \stackrel{ind}{\sim} N(0, \sigma_\alpha^2).$$

Again, this enforces a correlation between the observations within a group. This model is often referred to as a **random intercept model** because each group has its own intercept ( $\beta_0 + \alpha_j$ ) and these are *random* since  $\alpha_j$  has a distribution. Thus this model is related to a model that includes a fixed effect for each subject. But here those subject specific effects are shrunk toward an overall mean ( $\beta_0$ ).

# Sleep study example

```
ggplot(sleepstudy, aes(Days, Reaction, color = Subject)) +  
  geom_point() + theme_bw()
```



# Sleep study example

```
summary(me <- lmer(Reaction ~ Days + (1|Subject), sleepstudy))
```

Linear mixed model fit by REML ['lmerMod']

Formula: Reaction ~ Days + (1 | Subject)

Data: sleepstudy

REML criterion at convergence: 1786.5

Scaled residuals:

	Min	1Q	Median	3Q	Max
	-3.2257	-0.5529	0.0109	0.5188	4.2506

Random effects:

Groups	Name	Variance	Std.Dev.
Subject	(Intercept)	1378.2	37.12
Residual		960.5	30.99

Number of obs: 180, groups: Subject, 18

Fixed effects:

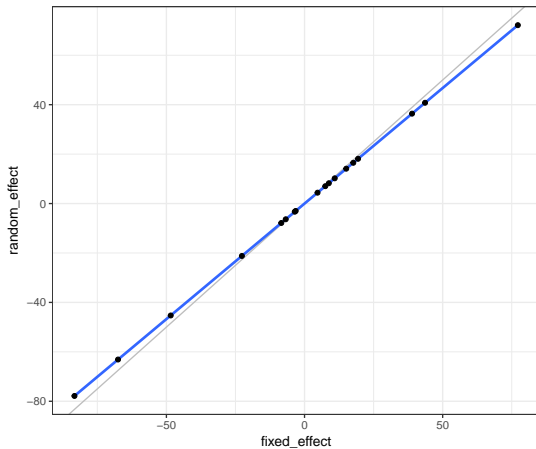
	Estimate	Std. Error	t value
(Intercept)	251.4051	9.7467	25.79
Days	10.4673	0.8042	13.02

Correlation of Fixed Effects:

	(Intr)
Days	-0.371

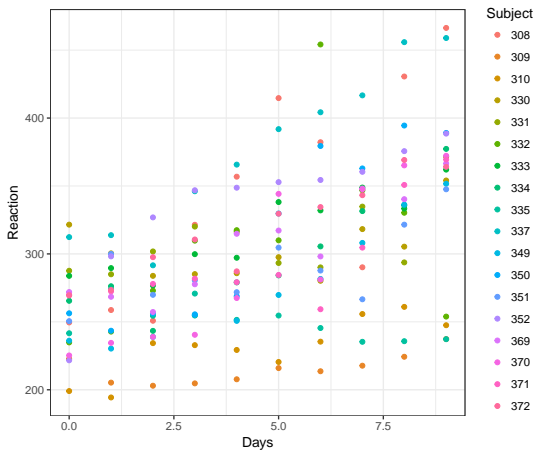


# Shrinkage



# Sleep study example

```
ggplot(sleepstudy, aes(Days, Reaction, color = Subject)) +  
  geom_point() + theme_bw()
```



## Random slope model

Suppose we have continuous observations  $Y_{ij}$  for individual  $i$  from group  $j$ . A mixed effect model with group specific slopes assumes

$$Y_{ij} = \beta_0 + \beta_1 X_{ij} + \alpha_{0j} + \alpha_{1j} X_{ij} + \epsilon_{ij} \quad \epsilon_{ij} \stackrel{ind}{\sim} N(0, \sigma_\epsilon^2)$$

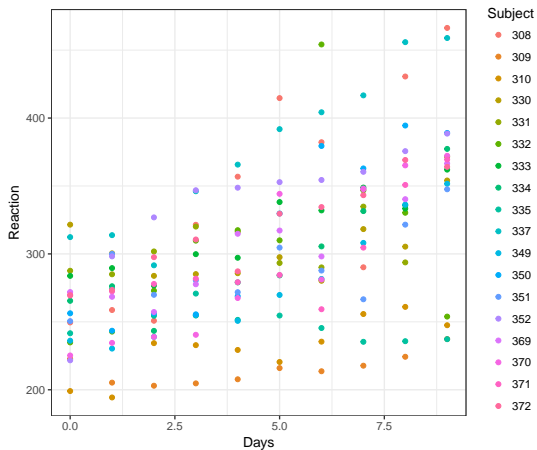
and, independent of  $\epsilon_{ij}$ ,

$$\begin{pmatrix} \alpha_{0j} \\ \alpha_{1j} \end{pmatrix} \stackrel{ind}{\sim} N(0, \Sigma_\alpha)$$

$N(0, \Sigma_\alpha)$  represents a bivariate normal with mean 0 and covariance matrix  $\Sigma_\alpha$ . This model is often referred to as a **random slope model** because each group has its own slope ( $\beta_1 + \alpha_{1j}$ ) and these are *random* since  $\alpha_{1j}$  has a distribution. Thus this model is related to a model that includes an interaction between the group and the explanatory variable, but here those subject specific slopes are shrunk toward an overall slope ( $\beta_1$ ).

# Sleep study example

```
ggplot(sleepstudy, aes(Days, Reaction, color = Subject)) +  
  geom_point() + theme_bw()
```



# Sleep study example

```
summary(me <- lmer(Reaction ~ Days + (Days|Subject), sleepstudy))
```

Linear mixed model fit by REML ['lmerMod']  
 Formula: Reaction ~ Days + (Days | Subject)  
 Data: sleepstudy

REML criterion at convergence: 1743.6

Scaled residuals:

Min	1Q	Median	3Q	Max
-3.9536	-0.4634	0.0231	0.4634	5.1793

Random effects:

Groups	Name	Variance	Std.Dev.	Corr
Subject	(Intercept)	612.09	24.740	
	Days	35.07	5.922	0.07
Residual		654.94	25.592	

Number of obs: 180, groups: Subject, 18

Fixed effects:

	Estimate	Std. Error	t value
(Intercept)	251.405	6.825	36.84
Days	10.467	1.546	6.77

Correlation of Fixed Effects:

	(Intr)
Days	-0.138