

Multiparameter models

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Outline

- Independent beta-binomial
 - Independent posteriors
 - Comparison of parameters
 - JAGS
- Probability theory results
 - Scaled Inv- χ^2 distribution
 - t -distribution
 - Normal-Inv χ^2 distribution
- Normal model with unknown mean and variance
 - Jeffreys prior
 - Natural conjugate prior

Motivating example

Is Andre Dawkins 3-point percentage higher in 2013-2014 than past years?

Season	Made	Attempts
2009-2010	36	95
2010-2011	64	150
2011-2012	67	171
2013-2014	64	152

Binomial model

Assume an independent binomial model,

$$Y_s \stackrel{ind}{\sim} \text{Bin}(n_s, \theta_s), \text{ i.e. }, p(y|\theta) = \prod_{s=1}^S p(y_s|\theta_s) = \prod_{s=1}^S \binom{n_s}{y_s} \theta_s^{y_s} (1-\theta_s)^{n_s-y_s}$$

where

- y_s is the number of 3-pointers made in season s
- n_s is the number of 3-pointers attempted in season s
- θ_s is the unknown 3-pointer success probability in season s
- S is the number of seasons
- $\theta = (\theta_1, \theta_2, \theta_3, \theta_4)'$ and $y = (y_1, y_2, y_3, y_4)$

and assume independent beta priors distribution:

$$p(\theta) = \prod_{s=1}^S p(\theta_s) = \prod_{s=1}^S \frac{\theta_s^{a_s-1} (1-\theta_s)^{b_s-1}}{\text{Beta}(a_s, b_s)} \mathbf{I}(0 < \theta_s < 1).$$

Joint posterior

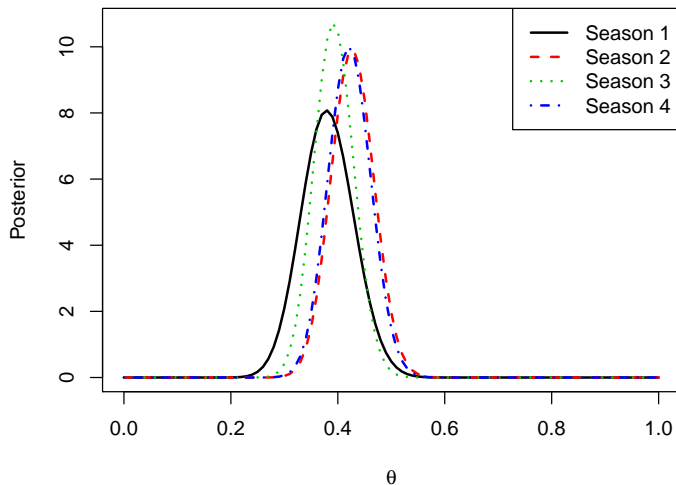
Derive the posterior according to Bayes rule:

$$\begin{aligned} p(\theta|y) &\propto p(y|\theta)p(\theta) \\ &= \prod_{s=1}^S p(y_s|\theta_s) \prod_{s=1}^S p(\theta_s) \\ &= \prod_{s=1}^S p(y_s|\theta_s)p(\theta_s) \\ &\propto \prod_{s=1}^S \text{Beta}(\theta_s|a_s + y_s, b_s + n_s - y_s) \end{aligned}$$

So the posterior for each θ_s is exactly the same as if we treated each season independently.

Joint posterior

Andre Dawkins's 3-point percentage



Monte Carlo estimates

Estimated means, medians, and quantiles.

	year	mean	median	ciL	ciU	hpdL	hpdU
1	1	0.3828454	0.3816672	0.2893217	0.4821211	0.2851402	0.4803823
2	2	0.4283304	0.4297132	0.3498912	0.5050538	0.3509289	0.5054018
3	3	0.3951943	0.3958465	0.3235839	0.4694850	0.3208512	0.4662410
4	4	0.4228666	0.4235223	0.3464835	0.4982144	0.3465337	0.4981711

Comparing probabilities across years

The scientific question of interest here is whether Dawkins's 3-point percentage is higher in 2013-2014 than previously. In probability notation this is

$$P(\theta_4 > \theta_s | y) \text{ for } s = 1, 2, 3.$$

which can be approximated via Monte Carlo as

$$P(\theta_4 > \theta_s | y) = E_{\theta|y}[\mathbf{I}(\theta_4 > \theta_s)] \approx \frac{1}{J} \sum_{j=1}^J \mathbf{I}(\theta_4^{(j)} > \theta_s^{(j)})$$

where

- $\theta_s^{(j)} \stackrel{\text{ind}}{\sim} \text{Be}(a_s + y_s, b_s + n_s - y_s)$
- $\mathbf{I}(A)$ is indicator function that is 1 if A is true and zero otherwise.

Estimated probabilities

```
# Should be able to use dcast
d = data.frame(theta_1 = sim$theta[sim$year==1],
               theta_2 = sim$theta[sim$year==2],
               theta_3 = sim$theta[sim$year==3],
               theta_4 = sim$theta[sim$year==4])

# Probabilities that season 4 percentage is higher than other seasons
mean(d$theta_4 > d$theta_1)

[1] 0.758

mean(d$theta_4 > d$theta_2)

[1] 0.454

mean(d$theta_4 > d$theta_3)

[1] 0.697
```

Using JAGS

```
library(rjags)
independent_binomials = "
model {
  for (i in 1:N) {
    y[i] ~ dbin(theta[i],n[i])
    theta[i] ~ dbeta(1,1)
  }
}
"

d = list(y=c(36,64,67,64), n=c(95,150,171,152), N=4)
m = jags.model(textConnection(independent_binomials), d)
```

```
Compiling model graph
  Resolving undeclared variables
  Allocating nodes
Graph information:
  Observed stochastic nodes: 4
  Unobserved stochastic nodes: 4
  Total graph size: 21
```

```
Initializing model
```

```
res = coda.samples(m, "theta", 1000)
```

```
summary(res)
```

```
Iterations = 1:1000
Thinning interval = 1
Number of chains = 1
Sample size per chain = 1000
```

1. Empirical mean and standard deviation for each variable,
plus standard error of the mean:

	Mean	SD	Naive SE	Time-series SE
theta[1]	0.3803	0.04728	0.001495	0.001495
theta[2]	0.4287	0.04035	0.001276	0.001276
theta[3]	0.3926	0.03767	0.001191	0.001191
theta[4]	0.4198	0.04086	0.001292	0.001292

2. Quantiles for each variable:

	2.5%	25%	50%	75%	97.5%
theta[1]	0.2893	0.3487	0.3812	0.4117	0.4741
theta[2]	0.3550	0.4011	0.4285	0.4549	0.5107
theta[3]	0.3202	0.3676	0.3921	0.4194	0.4654
theta[4]	0.3422	0.3905	0.4195	0.4475	0.5061

```
# Extract sampled theta values
theta = as.matrix(res[[1]]) # with only 1 chain, all values are in the first list element

# Calculate probabilities that season 4 percentage is higher than other seasons
mean(theta[,4] > theta[,1])

[1] 0.745

mean(theta[,4] > theta[,2])

[1] 0.448

mean(theta[,4] > theta[,3])

[1] 0.667
```

Background probability theory

- Scaled $\text{Inv-}\chi^2$ distribution
- Location-scale t -distribution
- Normal- $\text{Inv-}\chi^2$ distribution

Scaled-inverse χ^2 -distribution

If $\sigma^2 \sim IG(a, b)$, then $\sigma^2 \sim \text{Inv-}\chi^2(v, s^2)$ with

- $a = v/2$ and $b = vs^2/2$, or, equivalently,
- $v = 2a$ and $s^2 = b/a$.

Deriving from the inverse gamma, the scaled-inverse χ^2 has

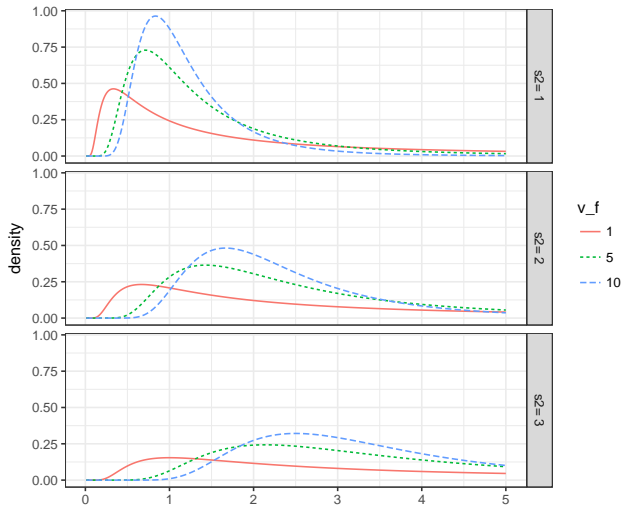
- Mean: $vs^2/(v-2)$ for $v > 2$
- Mode: $vs^2/(v+2)$
- Variance: $2v^2(s^2)^2/[(v-2)^2(v-4)]$ for $v > 4$

So s^2 is a point estimate and $v \rightarrow \infty$ means the variance decreases, since, for large v ,

$$\frac{2v^2(s^2)^2}{(v-2)^2(v-4)} \approx \frac{2v^2(s^2)^2}{v^3} = \frac{2(s^2)^2}{v}.$$

Scaled-inverse χ^2 -distribution

```
dinvgamma = function(x, a, b, ...) dgamma(1/x, a, b,...)/x^2
dsichisq = function(x, v, s2, ...) dinvgamma(x, v/2, v*s2/2, ...)
```



Location-scale *t*-distribution

The *t*-distribution is a location-scale family (Casella & Berger Thm 3.5.6), i.e. if T_v has a standard *t*-distribution with v degrees of freedom and pdf

$$f_t(t) = \frac{\Gamma([v+1]/2)}{\Gamma(v/2)\sqrt{v\pi}} (1 + t^2/v)^{-(v+1)/2},$$

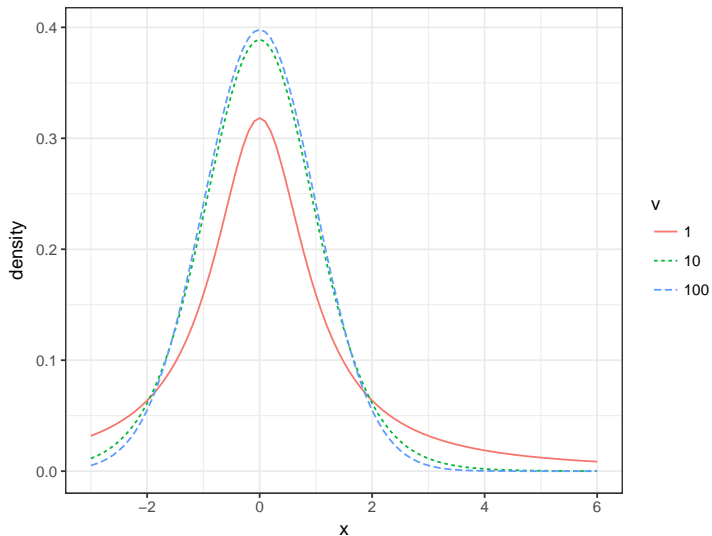
then $X = m + sT_v$ has pdf

$$f_X(x) = f_t([x-m]/s)/s = \frac{\Gamma([v+1]/2)}{\Gamma(v/2)\sqrt{v\pi}s} \left(1 + \frac{1}{v} \left[\frac{x-m}{s}\right]^2\right)^{-(v+1)/2}.$$

This is referred to as a *t* distribution with v degrees of freedom, location m , and scale s ; it is written as $t_v(m, s^2)$. Also,

$$t_v(m, s^2) \xrightarrow{v \rightarrow \infty} N(m, s^2).$$

t distribution as v changes



Normal-Inv- χ^2 distribution

Let $\mu|\sigma^2 \sim N(m, \sigma^2/k)$ and $\sigma^2 \sim \text{Inv-}\chi^2(v, s^2)$, then the kernel of this joint density is

$$\begin{aligned} p(\mu, \sigma^2) &= p(\mu|\sigma^2)p(\sigma^2) \\ &\propto (\sigma^2)^{-1/2} e^{-\frac{1}{2\sigma^2/k}(\mu-m)^2} (\sigma^2)^{-\frac{v}{2}-1} e^{-\frac{vs^2}{2\sigma^2}} \\ &= (\sigma^2)^{-(v+3)/2} e^{-\frac{1}{2\sigma^2}[k(\mu-m)^2 + vs^2]} \end{aligned}$$

In addition, the marginal distribution for μ is

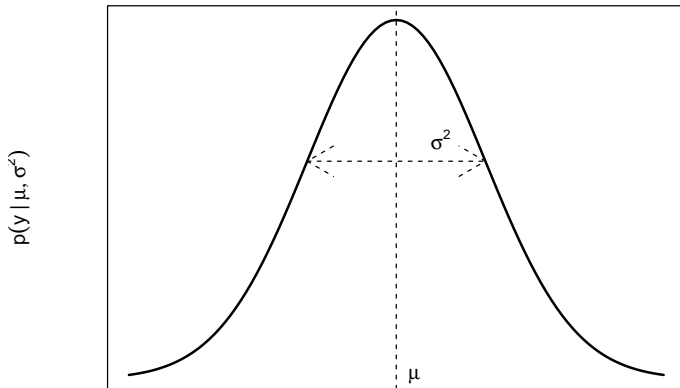
$$\begin{aligned} p(\mu) &= \int p(\mu|\sigma^2)p(\sigma^2)d\sigma^2 \\ &= \frac{\Gamma([v+1]/2)}{\Gamma(v/2)\sqrt{v\pi}s/\sqrt{k}} \left(1 + \frac{1}{v} \left[\frac{\mu-m}{s/\sqrt{k}}\right]^2\right)^{-(v+1)/2}. \end{aligned}$$

Thus $\mu \sim t_v(m, s^2/k)$.

Univariate normal model

Suppose $Y_i \stackrel{ind}{\sim} N(\mu, \sigma^2)$.

Normal model



Confidence interval for μ

Let

$$\bar{y} = \frac{1}{n} \sum_{i=1}^n y_i \quad \text{and} \quad S^2 = \frac{1}{n-1} \sum_{i=1}^n (y_i - \bar{y})^2.$$

Then,

$$T = \frac{\bar{Y} - \mu}{S/\sqrt{n}} \sim t_{n-1}$$

and an equal-tail $100(1 - \alpha)\%$ confidence interval can be constructed via

$$1 - \alpha = P(-c \leq T \leq c) = P\left(\bar{Y} - \frac{cS}{\sqrt{n}} \leq \mu \leq \bar{Y} + \frac{cS}{\sqrt{n}}\right)$$

and thus $\bar{y} \pm cS/\sqrt{n}$ is an equal-tail 95% confidence interval where $c = t_{n-1, 1-\alpha/2}$ is the t-critical value.

Default priors

Jeffreys prior can be shown to be $p(\mu, \sigma^2) \propto (1/\sigma^2)^{3/2}$. But alternative methods, e.g. reference prior, find that $p(\mu, \sigma^2) \propto 1/\sigma^2$ is a more appropriate prior.

The posterior under the reference prior is

$$\begin{aligned} p(\mu, \sigma^2 | y) &\propto (\sigma^2)^{-n/2} \exp\left(-\frac{1}{2\sigma^2} \sum_{i=1}^n (y_i - \mu)^2\right) \times \frac{1}{\sigma^2} \\ &= (\sigma^2)^{-n/2} \exp\left(-\frac{1}{2\sigma^2} \sum_{i=1}^n (y_i - \bar{y} + \bar{y} - \mu)^2\right) \times \frac{1}{\sigma^2} \\ &\quad \vdots \\ &= (\sigma^2)^{-(n-1+3)/2} \exp\left(-\frac{1}{2\sigma^2} [n(\mu - \bar{y})^2 + (n-1)S^2]\right) \end{aligned}$$

Thus

$$\mu | \sigma^2, y \sim N(\bar{y}, \sigma^2/n) \quad \sigma^2 | y \sim \text{Inv-}\chi^2(n-1, S^2).$$

Marginal posterior for μ

The marginal posterior for μ is

$$\mu|y \sim t_{n-1}(\bar{y}, S^2/n).$$

An equal-tailed $100(1 - \alpha)\%$ credible interval can be obtained via

$$\bar{y} \pm cS/\sqrt{n}$$

where $c = t_{n-1, 1-\alpha/2}$ is the t-critical value from before. This is exactly the same as the $100(1 - \alpha/2)\%$ confidence interval.

Predictive distribution

Let $\tilde{y} \sim N(\mu, \sigma^2)$. The predictive distribution is

$$\int \int p(\tilde{y}|\mu, \sigma^2)p(\mu|\sigma^2, y)p(\sigma^2|y)d\mu d\sigma^2$$

The easiest way to derive this is to write $\tilde{y} = \mu + \epsilon$ with

$$\mu|\sigma^2, y \sim N(\bar{y}, \sigma^2/n) \quad \epsilon|\sigma^2, y \sim N(0, \sigma^2)$$

independent of each other. Thus

$$\tilde{y}|\sigma^2, y \sim N(\bar{y}, \sigma^2[1 + 1/n]).$$

with $\sigma^2|y \sim \text{Inv-}\chi^2(n-1, S^2)$. Now, we can use the Normal-Inv- χ^2 theory, to find that

$$\tilde{y}|y \sim t_{n-1}(\bar{y}, S^2[1 + 1/n]).$$

Conjugate prior for μ and σ^2

The joint conjugate prior for μ and σ^2 is

$$\mu|\sigma^2 \sim N(m, \sigma^2/k) \quad \sigma^2 \sim \text{Inv-}\chi^2(v, s^2)$$

where s^2 serves as a prior guess about σ^2 and v controls how certain we are about that guess.

The posterior under this prior is

$$\mu|\sigma^2, y \sim N(m', \sigma^2/k') \quad \sigma^2|y \sim \text{Inv-}\chi^2(v', (s')^2)$$

where

$$\begin{aligned} k' &= k + n \\ m' &= [km + n\bar{y}]/k' \\ v' &= v + n \\ v'(s')^2 &= v s^2 + (n - 1)S^2 + \frac{kn}{k'}(\bar{y} - m)^2 \end{aligned}$$

Marginal posterior for μ

The marginal posterior for μ is

$$\mu|y \sim t_{v'}(m', (s')^2/k').$$

An equal-tailed $100(1 - \alpha)\%$ credible interval can be obtained via

$$m' \pm cs'/\sqrt{k'}$$

where $c = t_{v', 1-\alpha/2}$ is the t-critical value from before.

Marginal posterior via simulation

An alternative to deriving the closed form posterior for μ is to simulate from the distribution. Recall that

$$\mu|\sigma^2, y \sim N(m', \sigma^2/k') \quad \sigma^2|y \sim \text{Inv-}\chi^2(v', (s')^2)$$

To obtain a simulation from the posterior distribution $p(\mu, \sigma^2|y)$, calculate m', k', v' , and s' and then

1. simulate $\sigma^2 \sim \text{Inv-}\chi^2(v', (s')^2)$ and
2. using the simulated σ^2 , simulate $\mu \sim N(m', \sigma^2/k')$.

Not only does this provide a sample from the joint distribution for μ, σ but it also (therefore) provides a sample from the marginal distribution for μ .

The integral was suggestive:

$$p(\mu|y) = \int p(\mu|\sigma^2, y)p(\sigma^2|y)d\sigma^2$$

Predictive distribution via simulation

Similarly, we can obtain the predictive distribution via simulation. Recall that

$$p(\tilde{y}|y) = \int \int p(\tilde{y}|\mu, \sigma^2)p(\mu|\sigma^2, y)p(\sigma^2|y)d\mu d\sigma^2$$

To obtain a simulation from the predictive distribution $p(\tilde{y}|y)$, calculate m', k', v' , and s'

1. simulate $\sigma^2 \sim \text{Inv-}\chi^2(v', (s')^2)$,
2. using this σ^2 , simulate $\mu \sim N(m', \sigma^2/k')$, and
3. using these μ and σ^2 , simulate $\tilde{y} \sim N(\mu, \sigma^2)$.

Summary of normal inference

- Default analysis

- Prior: (think $\mu \sim N(0, \infty)$ and $\sigma^2 \sim \text{Inv-}\chi^2(0, 0)$)

$$p(\mu, \sigma^2) \propto 1/\sigma^2$$

- Posterior:

$$\mu|\sigma^2, y \sim N(\bar{y}, \sigma^2/n), \sigma^2|y \sim \text{Inv-}\chi^2(n-1, S^2), \mu|y \sim t_{n-1}(\bar{y}, S^2/n)$$

- Conjugate analysis

- Prior:

$$\mu|\sigma^2 \sim N(m, \sigma^2/k), \sigma^2 \sim \text{Inv-}\chi^2(v, s^2), \mu \sim t_v(m, s^2/k)$$

- Posterior:

$$\mu|\sigma^2, y \sim N(m', \sigma^2/k'), \sigma^2|y \sim \text{Inv-}\chi^2(v', (s')^2), \mu|y \sim t_{v'}(m', (s')^2/k')$$

with

$$k' = k + n, m' = [km + n\bar{y}]/k', v' = v + n,$$

$$v'(s')^2 = vs^2 + (n-1)S^2 + \frac{kn}{k'}(\bar{y} - m)^2$$