R04 - Categorical explanatory variables

STAT 401 (Engineering) - Iowa State University

April 10, 2018

Binary explanatory variable

Recall the simple linear regression model

$$Y_i \stackrel{ind}{\sim} N(\beta_0 + \beta_1 X_i, \sigma^2).$$

If we have a binary explanatory variable, i.e. the explanatory variable only has two levels say level 0 and level 1, we can code it as

$$X_i = \left\{ \begin{array}{ll} 1 & \text{obseration } i \text{ is level } 1 \\ 0 & \text{observation } i \text{ is level } 0 \end{array} \right.$$

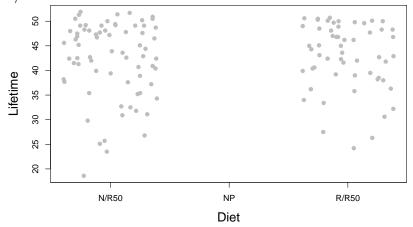
= I(observation i is level 1)

where ${\rm I}(A)$ is an indicator function that is 1 when A is true and 0 otherwise. Then

- β_0 is the expected response for level 0,
- $\beta_0 + \beta_1$ is the expected response for level 1, and therefore
- β_1 is the expected difference in response (level 1 minus level 0).

Mice lifetimes

Reconsider the mice lifetime data set but only consider the diets R/R50 and N/R50:



Regression model for mice lifetimes

Considering only the R/R50 and N/R50 diets. Let

$$Y_i \stackrel{ind}{\sim} N(\beta_0 + \beta_1 X_i, \sigma^2)$$

where Y_i is the lifetime of the ith mouse and

$$X_i = I(Diet_i = N/R50) = \begin{cases} 1 & i \text{th mouse diet is } N/R50 \\ 0 & i \text{th mouse diet is } R/R50 \end{cases}$$

then

$$\begin{array}{ll} E[\mathsf{Lifetime}|\mathsf{R}/\mathsf{R50}] &= E[Y_i|X_i = 0] &= \beta_0 \\ E[\mathsf{Lifetime}|\mathsf{N}/\mathsf{R50}] &= E[Y_i|X_i = 1] &= \beta_0 + \beta_1 \end{array}$$

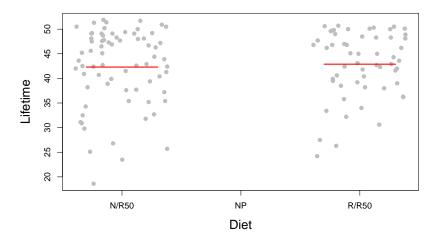
and

$$\begin{split} E[\text{Lifetime difference}] &= E[\text{Lifetime}|\text{N/R50}] - E[\text{Lifetime}|\text{R/R50}] \\ &= (\beta_0 + \beta_1) - \beta_0 = \beta_1. \end{split}$$

R code

```
case0501$X <- ifelse(case0501$Diet == "N/R50", 1, 0)
unique(case0501$X)
[1] 0 1
(m <- lm(Lifetime ~ X, data = case0501, subset = Diet %in% c("R/R50","N/R50")))
Call:
lm(formula = Lifetime ~ X, data = case0501, subset = Diet %in%
   c("R/R50", "N/R50"))
Coefficients:
(Intercept) X
42.8857 -0.5885
confint(m)
               2.5 % 97.5 %
(Intercept) 40.952257 44.819172
X -3.174405 1.997342
predict(m, data.frame(X=1), interval = "confidence") # Expected lifetime on N/R50
      fit
           lwr
                       upr
1 42 29718 40 58007 44 0143
```

Mice lifetimes



Equivalence to model for two-sample t-test

Recall that our two-sample t-test had the model

$$Y_{ij} \stackrel{ind}{\sim} N(\mu_j, \sigma^2)$$

for groups j=0,1. This is equivalent to our current regression model where

$$\mu_0 = \beta_0$$

$$\mu_1 = \beta_0 + \beta_1$$

assuming

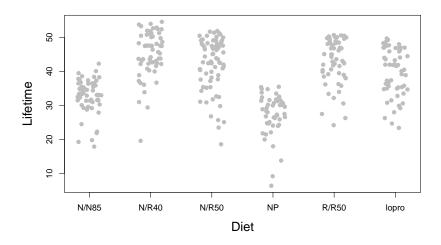
- \bullet μ_0 represents the mean for the R/R50 group and
- μ_1 represents the mean for N/R50 group.

When the models are effectively the same, but have different parameters we call it a reparameterization.

Equivalence

```
summary(m)$coefficients[2,4] # p-value
[1] 0.6531748
confint(m)
               2.5 % 97.5 %
(Intercept) 40.952257 44.819172
           -3 174405 1 997342
t.test(Lifetime ~ Diet, data = case0501, subset = Diet %in% c("R/R50","N/R50"), var.equal=TRUE)
Two Sample t-test
data: Lifetime by Diet
t = -0.45044, df = 125, p-value = 0.6532
alternative hypothesis: true difference in means is not equal to 0
95 percent confidence interval:
-3 174405 1 997342
sample estimates:
mean in group N/R50 mean in group R/R50
           42 29718
                               42.88571
```

Using a categorical variable as an explanatory variable.



Regression with a categorical variable

- \bullet Choose one of the levels as the reference level, e.g. N/N85
- Construct dummy variables using indicator functions, i.e.

$$I(A) = \begin{cases} 1 & A \text{ is TRUE} \\ 0 & A \text{ is FALSE} \end{cases}$$

for the other levels, e.g.

 $X_{i,1} = I(\text{diet for observation } i \text{ is N/R40})$ $X_{i,2} = I(\text{diet for observation } i \text{ is N/R50})$ $X_{i,3} = I(\text{diet for observation } i \text{ is NP})$ $X_{i,4} = I(\text{diet for observation } i \text{ is R/R50})$ $X_{i,5} = I(\text{diet for observation } i \text{ is lopro})$

• Estimate the parameters of a multiple regression model using these dummy variables.

Regression model

Our regression model becomes

$$Y_i \stackrel{ind}{\sim} N(\beta_0 + \beta_1 X_{i,1} + \beta_2 X_{i,2} + \beta_3 X_{i,3} + \beta_4 X_{i,4} + \beta_5 X_{i,5}, \sigma^2)$$

where

- β_0 is the expected lifetime for the N/N85 group
- $\beta_0 + \beta_1$ is the expected lifetime for the N/R40 group
- $\beta_0 + \beta_2$ is the expected lifetime for the N/R50 group
- $\beta_0 + \beta_3$ is the expected lifetime for the NP group
- $\beta_0 + \beta_4$ is the expected lifetime for the R/R50 group
- $\beta_0 + \beta_5$ is the expected lifetime for the lopro group and thus β_p for p > 0 is the difference in expected lifetimes.

R code

```
case0501 <- case0501 %>%
  mutate(X1 = Diet == "N/R40",
        X2 = Diet == "N/R50",
        X3 = Diet == "NP",
        X4 = Diet == "R/R50".
        X5 = Diet == "lopro")
m <- lm(Lifetime ~ X1 + X2 + X3 + X4 + X5, data= case0501)
m
Call:
lm(formula = Lifetime ~ X1 + X2 + X3 + X4 + X5, data = case0501)
Coefficients:
(Intercept)
                 X1TRUE
                               X2TRUE
                                            X3TRUE
                                                         X4TRUE
                                                                      X5TRUE.
     32,691
                 12.425
                               9.606
                                            -5.289
                                                         10.194
                                                                       6.994
confint(m)
                2.5 % 97.5 %
(Intercept) 30.951394 34.431062
X1TRUE
             9.995893 14.854984
X2TRUE
           7.269897 11.942013
            -7.848142 -2.730232
X3TRUE
X4TRUE
           7.723030 12.665943
X5TRUE.
            4.523030 9.465943
```

R code (cont.)

```
summary(m)
Call:
lm(formula = Lifetime ~ X1 + X2 + X3 + X4 + X5, data = case0501)
Residuals:
    Min
             10 Median
                                      Max
                               30
-25.5167 -3.3857 0.8143 5.1833 10.0143
Coefficients:
           Estimate Std. Error t value Pr(>|t|)
                       0.8846 36.958 < 2e-16 ***
(Intercept) 32.6912
X1TRUE
           12.4254
                   1.2352 10.059 < 2e-16 ***
X2TRIIE
           9.6060
                      1.1877 8.088 1.06e-14 ***
           -5.2892
                      1.3010 -4.065 5.95e-05 ***
X3TRUE
X4TRUE
           10.1945
                      1.2565 8.113 8.88e-15 ***
X5TRUE.
           6.9945
                      1.2565 5.567 5.25e-08 ***
Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' 1
Residual standard error: 6.678 on 343 degrees of freedom
Multiple R-squared: 0.4543, Adjusted R-squared: 0.4463
F-statistic: 57.1 on 5 and 343 DF, p-value: < 2.2e-16
```

Interpretation

 $eta_0 = E[Y_i | ext{reference level}], ext{ i.e. expected response for the reference level}$

Note: the only way $X_{i,1} = \cdots = X_{i,p} = 0$ is if all indicators are zero, i.e. at the reference level.

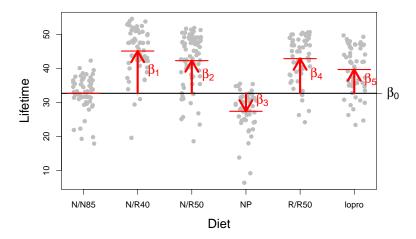
• $\beta_p, p > 0$: expected change in the response moving from the reference level to the level associated with the p^{th} dummy variable Note: the only way for $X_{i,p}$ to increase by one is if initially

$$X_{i,1} = \cdots = X_{i,p} = 0$$
 and now $X_{i,p} = 1$

For example,

- The expected lifetime for mice on the N/N85 diet is 32.7 (31.0,34.4) weeks.
- The expected increase in lifetime for mice on the N/R40 diet compared to the N/N85 diet is 12.4 (10.0,14.9) weeks.
- The model explains 45% of the variability in mice lifetimes.

Using a categorical variable as an explanatory variable.



Equivalence to multiple group model

Recall that we had a multiple group model

$$Y_{ij} \stackrel{ind}{\sim} N(\mu_j, \sigma^2)$$

for groups $j=0,1,2,\ldots,5$. This is equivalent to our current regression model where

$$N/N85$$
: $\mu_0 = \beta_0$
 $N/R40$: $\mu_1 = \beta_0 + \beta_1$
 $N/R50$: $\mu_2 = \beta_0 + \beta_2$
 NP : $\mu_3 = \beta_0 + \beta_3$
 $R/R50$: $\mu_4 = \beta_0 + \beta_4$
 $lopro$: $\mu_5 = \beta_0 + \beta_5$

assuming the groups are labeled appropriately.