

# STAT 401A - Statistical Methods for Research Workers

## Regression diagnostics

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*This isn’t just true in statistics! Maps are a type of model; they are wrong. But good maps are very useful.*

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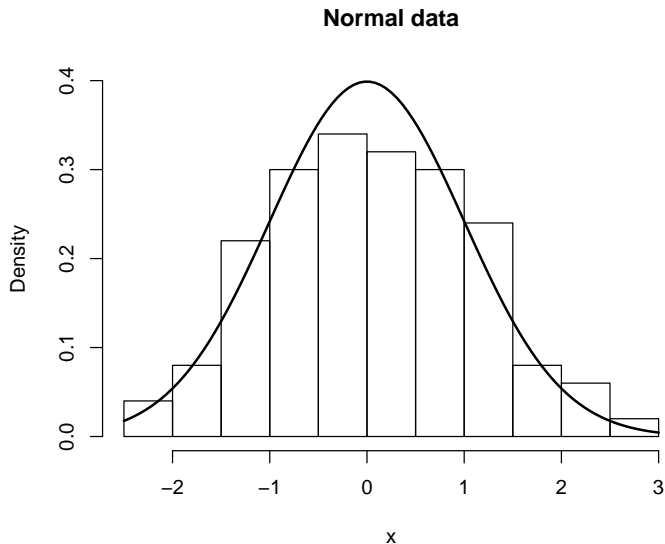
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- Linearity between mean response and explanatory variable

# Histograms with best fitting bell curves



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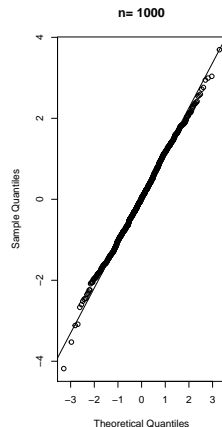
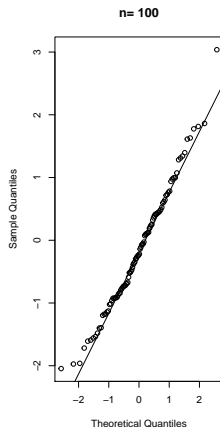
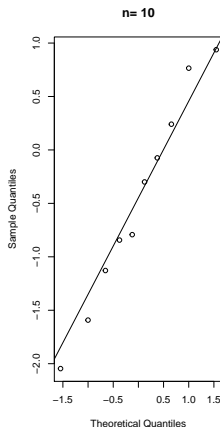
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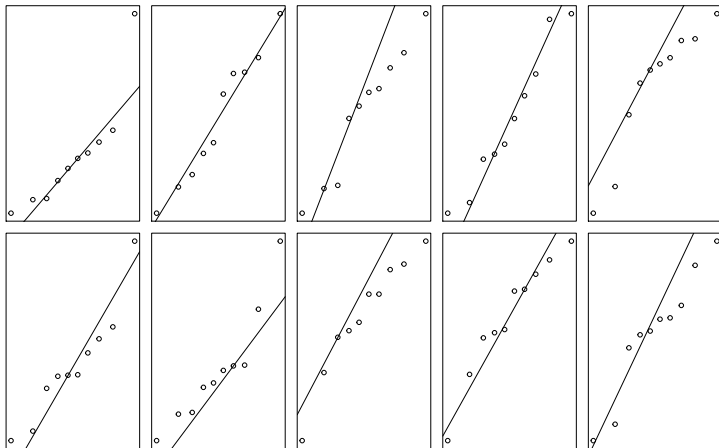
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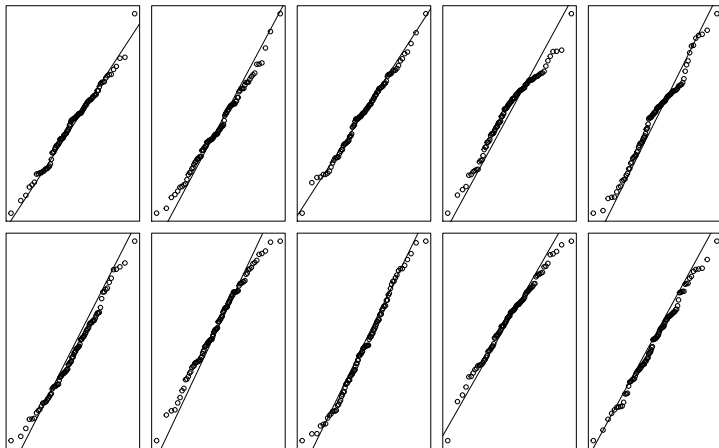
**Remark** The bottom line is that, if the distribution assumption is satisfied, the points should fall roughly along the line. Systematic variation from this line indicates skewness,

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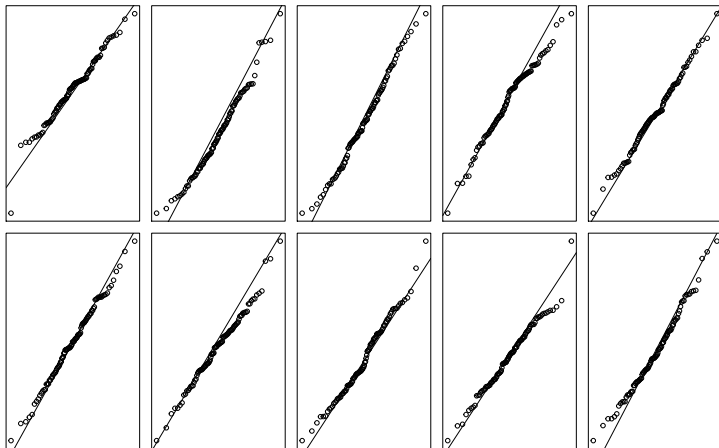


# Normal ( $n=10$ )

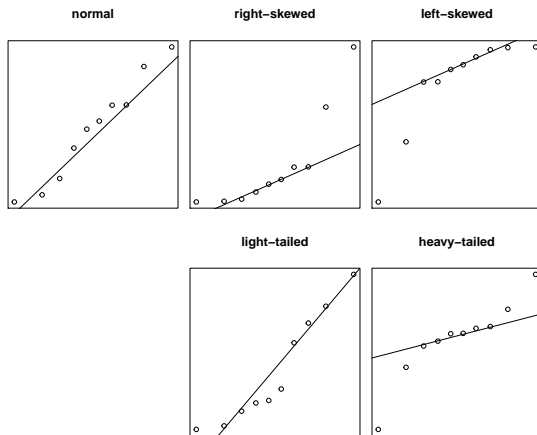


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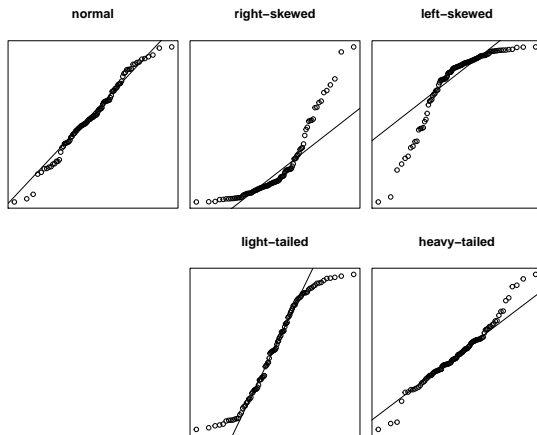
# Normal ( $n=1000$ )



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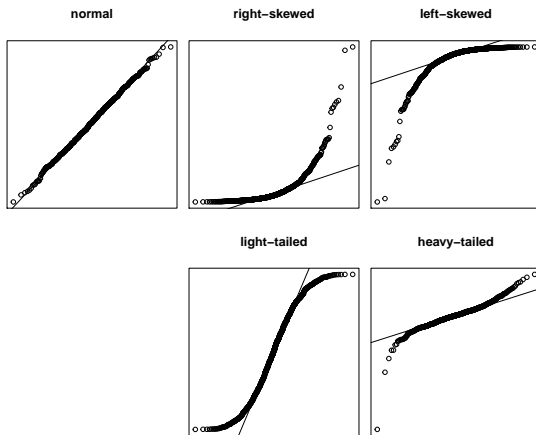


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Other patterns are certainly possible, but these are the most common.

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Recall the model

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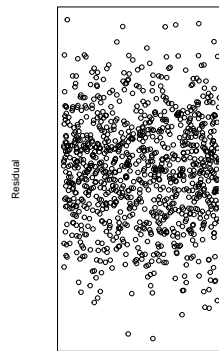
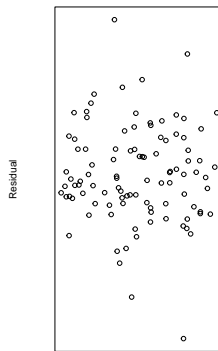
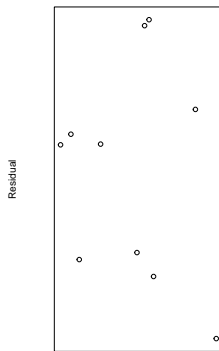
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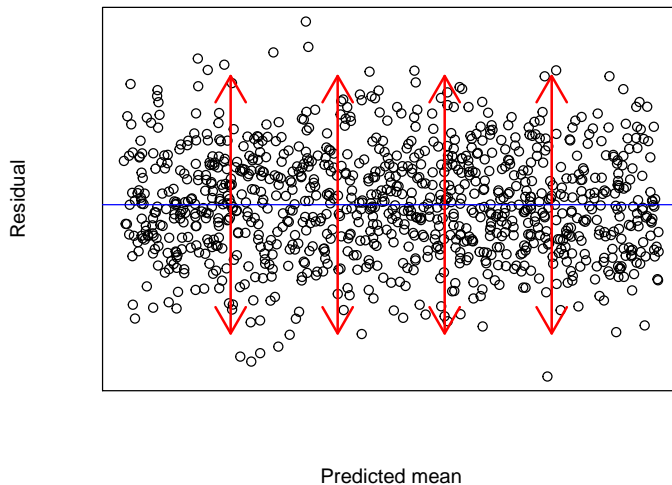
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The most common way this assumption is violated is by having increasing variance with increasing mean, thus we often look at a residuals vs predicted (fitted) mean plot.

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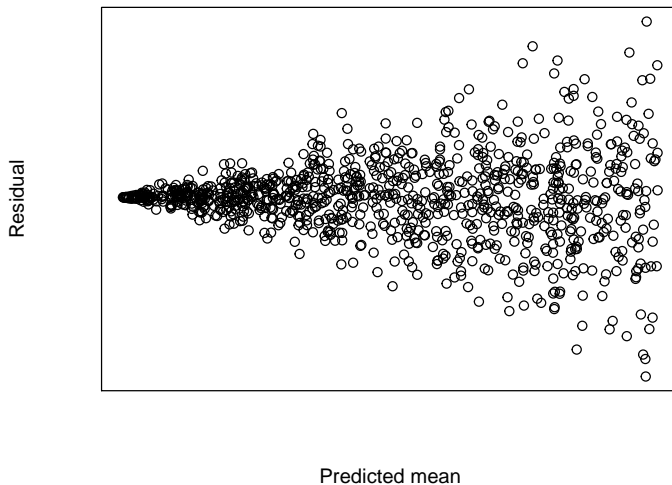


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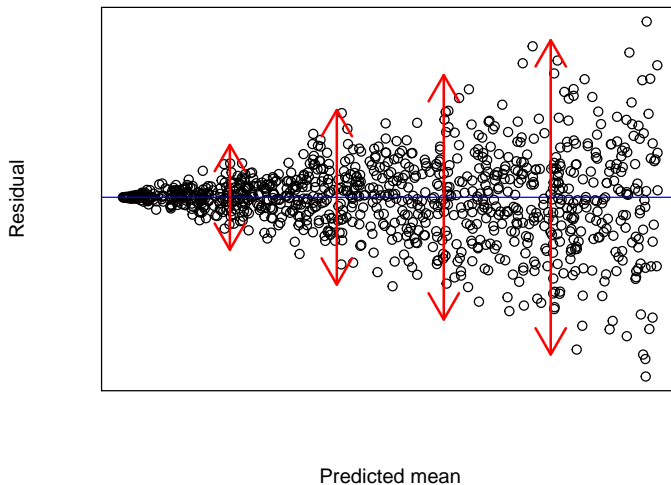




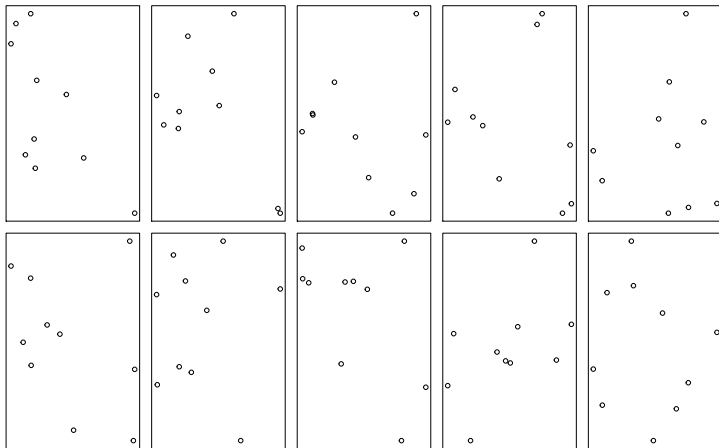
# Extreme non-constant variance (funnel)



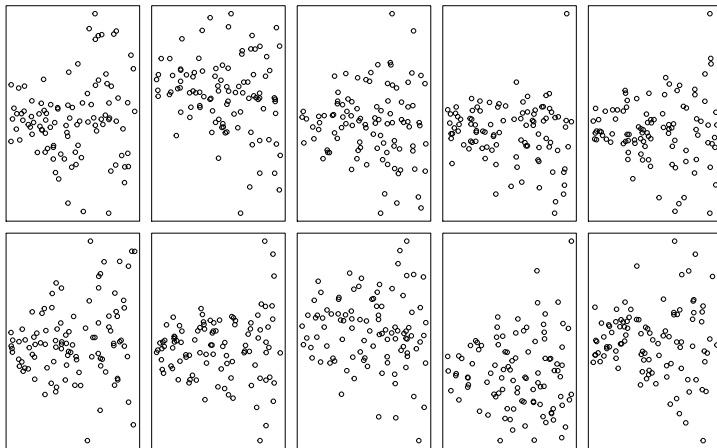
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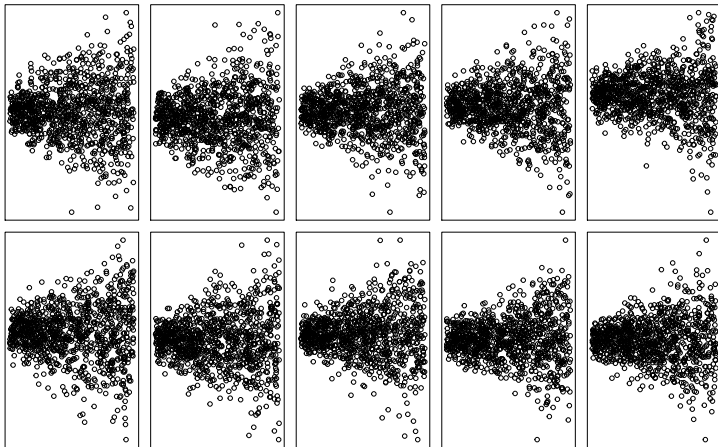
# Non-constant variance ( $n=10$ , $\sigma_2/\sigma_1 = 4$ )



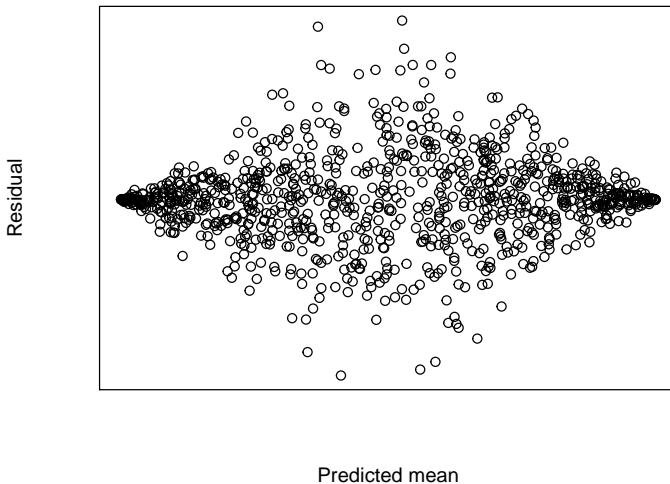
# Non-constant variance ( $n=100$ , $\sigma_2/\sigma_1 = 4$ )



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- Serial correlation
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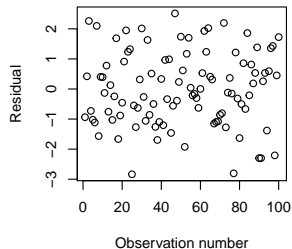
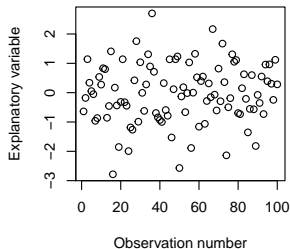
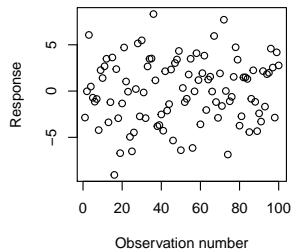
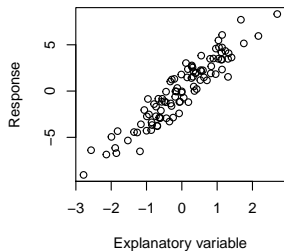
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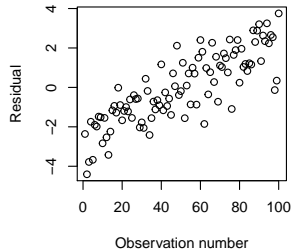
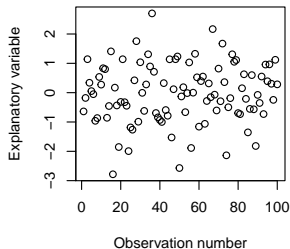
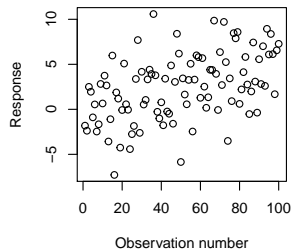
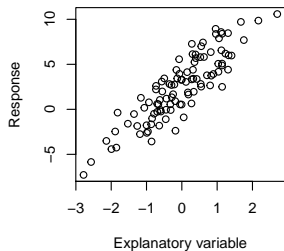
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# No evidence for lack of independence

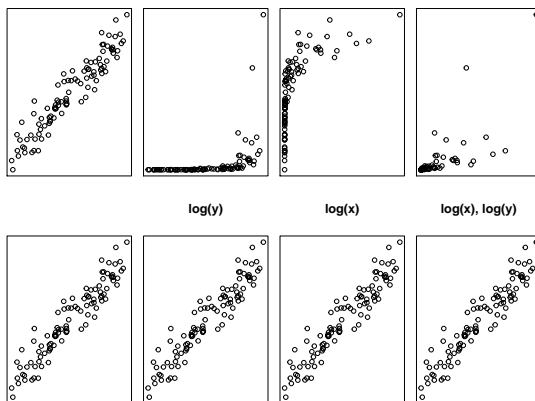


# Evidence for lack of independence



# Linearity

Assess using scatterplots of (transformed) response vs (transformed) explanatory variable:



# Testing Composite hypotheses

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Do the following

1. Calculate extra sum of squares.
2. Calculate extra degrees of freedom
3. Calculate

$$\text{F-statistic} = \frac{\text{Extra sum of squares} / \text{Extra degrees of freedom}}{\hat{\sigma}_{full}^2}$$

4. Compare this to an F-distribution with

- numerator degrees of freedom = extra degrees of freedom
- denominator degrees of freedom = degrees of freedom in estimating  $\hat{\sigma}_{full}^2$



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- Small pvalues indicate a lack-of-fit, i.e. the reduced model is not adequate.

# Lack-of-fit F-test

Let  $Y_{ij}$  be the  $i^{th}$  observation from the  $j^{th}$  group where the group is defined by those observations having the same explanatory variable value ( $X_j$ ).

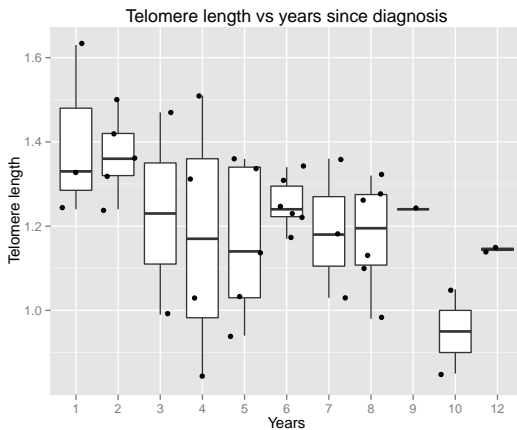
Two models:

$$\text{ANOVA: } Y_{ij} \stackrel{\text{ind}}{\sim} N(\mu_j, \sigma^2) \quad (\text{full})$$

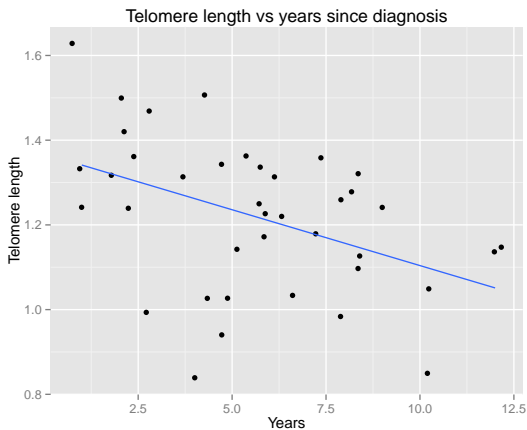
$$\text{Regression: } Y_{ij} \stackrel{\text{ind}}{\sim} N(\beta_0 + \beta_1 X_j, \sigma^2) \quad (\text{reduced})$$

- Regression model is reduced:
  - ANOVA has  $J$  parameters for the mean
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- Lack-of-fit F-test requires multiple observations at a few  $X_j$  values!

# Telomere length



# Telomere length



# SAS code

```
DATA t;
  INFILE 'telomeres.csv' DSD FIRSTOBS=2;
  INPUT years length;

PROC REG DATA=t;
  MODEL length = years / CLB LACKFIT;
  RUN;
```

The REG Procedure  
 Model: MODEL1  
 Dependent Variable: length

Number of Observations Read	39
Number of Observations Used	39

## Analysis of Variance

Source	DF	Sum of Squares	Mean Square	F Value	Pr > F
Model	1	0.22777	0.22777	8.42	0.0062
Error	37	1.00033	0.02704		
Lack of Fit	9	0.18223	0.02025	0.69	0.7093
Pure Error	28	0.81810	0.02922		
Corrected Total	38	1.22810			

```
# Use as.factor to turn a continuous variable into a categorical variable
m_anova = lm(telomere.length ~ as.factor(years), Telomeres)
m_reg   = lm(telomere.length ~ years, Telomeres)
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#### Analysis of Variance Table

Model 1: telomere.length ~ years

Model 2: telomere.length ~ as.factor(years)

	Res.Df	RSS	Df	Sum of Sq	F	Pr(>F)
1	37	1.000				
2	28	0.818	9	0.182	0.69	0.71

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No evidence of a lack of fit.

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  - Transform explanatory variable
  - Add other explanatory variable(s)

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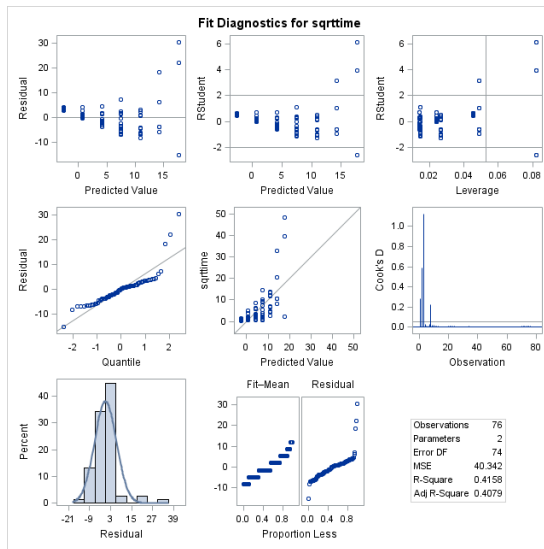
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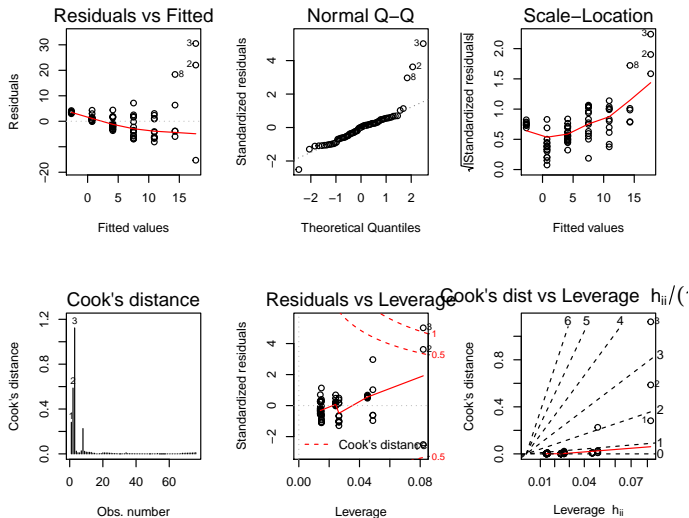
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# Default diagnostics in SAS





# Default diagnostics in R



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  - $\beta_1$  will affect the change in the response when there is a multiplicative change in  $X$

# Summary of interpretations when using logarithms

- When using the log of the response,
  - $\beta_0$  will affect the median response
  - $\beta_1$  will affect the multiplicative change in the median response
- When using the log of the explanatory variable ( $X$ ),
  - $\beta_0$  will affect the response when  $X = 1$
  - $\beta_1$  will affect the change in the response when there is a multiplicative change in  $X$



# Summary of interpretations when using logarithms

- When using the log of the response,
  - $\beta_0$  will affect the median response
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  - $\beta_0$  will affect the response when  $X = 1$
  - $\beta_1$  will affect the change in the response when there is a multiplicative change in  $X$

To construct confidence intervals for  $e^\beta$ , find a confidence interval for  $\beta$  and exponentiate the endpoints, i.e. if  $(L, U)$  is a confidence interval for  $\beta$ , then  $(e^L, e^U)$  is a confidence interval for  $e^\beta$ .