Note for Bandit Algorithms

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Theorem 0.1

Suppose that $X - \mu$ obeys σ -subgaussian distribution, we have

$$\mathbb{P}(\hat{\mu} - \mu \ge \epsilon) \le \exp(-\frac{n\epsilon^2}{2\sigma^2})$$

for any $\epsilon > 0$, where $\mu = \mathbb{E}[X]$ and $\hat{\mu} = \overline{X}$. Equivalently, for $\forall \delta \in (0,1]$,

$$\mathbb{P}(\hat{\mu} - \mu \ge \sigma \sqrt{\frac{2}{n} \log \frac{1}{\delta}}) \le \delta$$

1 Stochastic Bandits

Theorem 1.1: ETC

Suppose that $X_i - \mu_i$ obeys σ -subgaussian distribution, the regret for ETC is bounded by

$$m\sum_{i=1}^{k} \Delta_i + (n - mk)\sum_{i=1}^{k} \Delta_i \exp(-\frac{m\Delta_i^2}{4\sigma^2})$$

when $n \geq km$.

Theorem 1.2: UCB

Suppose that $X_i - \mu_i$ obeys σ -subgaussian distribution, the regret for UCB is bounded by

$$3\sum_{i=1}^{k} \Delta_i + \min\{\sum_{\Delta_i \neq 0} \frac{16\log n}{\Delta_i}, 8\sqrt{nk\log n}\}$$

when $\delta = 1/n^2$.

Theorem 1.3

Suppose that X obeys continuous distribution, then we have

$$CVaR_{\alpha}(X) := \inf_{x \in \mathbb{R}} \{ x + \frac{1}{\alpha} \mathbb{E}[(X - x)^{+}] \} = \mathbb{E}[X | X \ge VaR_{\alpha}(X)]$$

Proof. Let $G(x) = x + \frac{1}{\alpha} \mathbb{E}[(X - x)^+]$. Notice that

$$\mathbb{E}[(X-x)^{+}] = \int_{F(x)}^{1} (t-x) dF(t) = \int_{F(x)}^{1} (F^{-1}(u) - x) du$$

Then
$$\frac{dG}{dx} = 1 - \frac{1}{\alpha}(1 - F(x)) = 0 \Leftrightarrow x = F^{-1}(1 - \alpha) = \text{VaR}_{\alpha}(X)$$
 and $\frac{d^2G}{dx^2} = \frac{1}{\alpha}f(x) \ge 0$. Hence

$$\inf_{x \in \mathbb{R}} G(x) = G(\operatorname{VaR}_{\alpha}(X)) = \operatorname{VaR}_{\alpha}(X) + \underbrace{\frac{1}{\alpha} \mathbb{E}[(X - \operatorname{VaR}_{\alpha}(X))^{+}]}_{\mathbb{E}[X - \operatorname{VaR}_{\alpha}(X)|X \geq \operatorname{VaR}_{\alpha}(X)]} = \mathbb{E}[X | X \geq \operatorname{VaR}_{\alpha}(X)]$$

as we desire. \Box

Theorem 1. Suppose that $\overline{R}_n = \sum_{t=1}^n \Delta_{A_t}$

Proof. For any $(u_1, \ldots, u_k) \in \mathbb{N}_+^k$, define

$$G_i = \{\mu^* < \min_{t \in \{1, \dots, n\}} UCB^*(t, \delta)\} \cap \{\hat{\mu}_{iu_i} + \sqrt{\frac{2}{u_i} \log \frac{1}{\delta}} < \mu^*\}$$

When G_i occurs, there is $T_i(n) \leq u_i$. Hence $T_i(n) \geq u_i + 1$ implies G_i^c . Notice that

$$G_i^c \subset \{\mu^* \ge \min_{s \in \{1, \dots, n\}} \hat{\mu}^* + \sqrt{\frac{2}{s} \log \frac{1}{\delta}}\} \cup \{\hat{\mu}_{iu_i} - \mu_i \ge \Delta_i - \sqrt{\frac{2}{u_i} \log \frac{1}{\delta}}\}$$

Thus we have

$$\mathbb{P}(G_i^c) \le n\delta + \exp(-\frac{u_i(\Delta_i - \sqrt{\frac{2}{u_i}\log\frac{1}{\delta}})^2}{2})$$

There is

$$\mathbb{P}(\overline{R}_n \ge \sum_{\Delta_i \ne 0} (u_i + 1)\Delta_i) \le \mathbb{P}(\exists i \in \{1, \dots, k\}, \Delta_i \ne 0, T_i \ge u_i + 1) \le \sum_{\Delta_i \ne 0} \mathbb{P}(T_i \ge u_i + 1)$$

$$\le \sum_{\Delta_i \ne 0} \mathbb{P}(G_i^c) \le (k - 1)n\delta + \sum_{\Delta_i \ne 0} \exp(-\frac{u_i(\Delta_i - \sqrt{\frac{2}{u_i}\log\frac{1}{\delta}})^2}{2})$$

References