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Chapter I: Combinatorial Analysis

Counting

Basic Counting Principles

Theorem 1.1 (Pigeonhole principle. Dirichlet, 1834) if n items are to be put into m containers, with $n > m$, then at least one container must contain more than one item.

Corollary 1.2: if n items are to be put into m containers, with $n > m$, then at least one container must contain at least $\lceil \frac{n}{m} \rceil$ item.

example: you go to a restaurant. That day, they propose 4 starters, 4 courses and 3 desserts. You choose one starter, one course and one dessert to form into your meal set. How many different sets are possible?

answer: $4 \times 4 \times 3 = 48$.

thus, we can get the following Proposition.

Proposition 1.3: If r experiments are to be performed sequentially (按顺序) and the first experiment can be performed in n_1 ways, \dots , the r th experiment in n_r ways, then there are $\prod_{i=1}^n n_i$ ways to perform the r experiments.

example bis: Still for the meal selection, you can either choose one starter and one course, or one course and one dessert. That is, you cannot take a starter, a course and a dessert. How many different sets are

possible? **answer:** $4 \times 4 + 4 \times 3 = 28$.

Permutations

let's begin with an example:How many different ranking orders are possible for 10 tennis players? A_{10}^{10} .

Definition 1.4 (Permutation): An ordered ranking of $n \in N^*$ distinct elements is called a **permutation**.

Proposition 1.5: There are $n(n - 1) \cdots 1$ permutations of $n \in N^*$ distinct elements.

$$P_1E_1P_2P_3E_2R = \frac{A_6^6}{A_2^2 \times A_3^3}$$

组合:

Combinations

Multinomials

consider:

$$(x_1 + x_2 + \cdots + x_r)^n$$

for $x_1^{n_1}x_2^{n_2} \cdots x_r^{n_r}$, we choose n_1 's x_1 in n item, and n_2 's x_2 in $n - n_1$ item.

Prop: $n_1 + n_2 + n_3 + \cdots + n_r = n$ have $\binom{n+r-1}{r-1}$ distinct nonnegative integer_valued vectors (n_1, n_2, \cdots , n_r)

Proof: We can see this problem as the form following:

$$y_1 + y_2 + \cdots + y_r = n + r$$

where $y_i = n_i + 1, y_i \geq 1$.Specifically,what we need to anwser is how many methods of inserting $r - 1$ spacers in $n + r - 1$ gaps. **Ex:** $n = 1, r = 3$

$$\bigcirc \bigcirc | \bigcirc \bigcirc : (0, 0, 1)$$

$$\bigcirc | \bigcirc \bigcirc | \bigcirc : (0, 1, 0)$$

$$\bigcirc \bigcirc | \bigcirc | \bigcirc : (1, 0, 0)$$

i.e. $\binom{1+3-1}{3-1} = \binom{3}{2} = 3$

Chapter.II Axioms of Probability

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Set

Definition 2.1

1. random experiment
2. sample space (is denoted by S)
3. sample point

Remark 2.2 S can be finite or **infinite** (countable or uncountable)

Definition 2.2 subset/superset

Definition 2.3 event

Remark 2.3 Sure set: S , impossible set : \emptyset

Relation

1. Union
2. Intersection
3. Countable union/intersection
4. Mutually exclusive: $E \cap F = \emptyset$
5. Complement: E^c
6. Difference: $E \setminus F$
7. Symmetric difference: $E \Delta F = \{\omega | \omega \in E \setminus F \text{ or } \omega \in F \setminus E\}$

Proposition 2.5 (De Morgan's Law)

$$\begin{aligned} \left(\bigcup_{i=1}^n E_i \right)^c &= \bigcap_{i=1}^n E_i^c \\ \left(\bigcap_{i=1}^n E_i \right)^c &= \bigcup_{i=1}^n E_i^c \end{aligned}$$

Axioms of Probability

Some Properties

$$P(\emptyset) = 0$$

Proof If we consider a sequence $\{E_i\}$ where $2E_1 = S, E_i = \emptyset$, for $i \geq 2$, then $S = \bigcup_{i=1}^{\infty} E_i$. Hence,

$$P(S) = P\left(\bigcup_{i=1}^{\infty} E_i\right) = \sum_{i=1}^{\infty} P(E_i) = P(S) + \sum_{i=2}^{\infty} P(E_i) \implies P(\emptyset) = 0$$

$$P\left(\bigcup_{i=1}^n E_i\right) = \sum_{i=1}^n P(E_i)$$

where $E_i, E_j, i \neq j$ are mutually exclusive

Proof Similar to **Proof.1**

$$P(E) \leq P(F)$$

where $E \subseteq F \subseteq A$

Proof $P(F) = P(E + F \setminus E) = P(E) + P(F \setminus E)$

4. (inclusion and exclusion identity) For any two events E, F

$$P(E \cup F) = P(E) + P(F) - P(E \cap F)$$

$$P\left(\bigcup_{i=1}^n E_i\right) = \sum_{k=1}^n (-1)^{k-1} \sum_{1 \leq i_1 \leq i_2 \leq \dots \leq i_k \leq n} P(E_{i_1} \cap E_{i_2} \dots \cap E_{i_k}) \quad (2.1)$$

Proof $P(E \cup F) = P(E \cup E^c F) = P(E) + P(E^c F),$

consider $EF + E^c F = F, EF \cap E^c F = \emptyset \implies P(E) + P(E^c F) = P(E) + P(F) - P(EF)$

Ex $n = 4$:

$$\begin{aligned} & P\left(\bigcup_{i=1}^4 E_i\right) \\ &= P(E_1) + P(E_2) + P(E_3) + P(E_4) \\ &\quad - P(E_1 E_2) - P(E_2 E_3) - P(E_2 E_4) - P(E_1 E_4) - P(E_1 E_3) - P(E_3 E_4) \\ &\quad + P(E_1 E_2 E_3) + P(E_1 E_3 E_4) + P(E_1 E_2 E_4) + P(E_2 E_3 E_4) \\ &\quad - P(E_1 E_2 E_3 E_4) \end{aligned}$$

$$P(E \cup F) \leq P(E) + P(F)$$

(A generalization) For a finite sequence of events E_1, E_2, \dots, E_n

$$P\left(\bigcup_{i=1}^n E_i\right) \leq P(E_1) + P(E_2) + \dots + P(E_n) \quad (2.2)$$

(Infinite) *Boole's inequality* : For a countably infinite sequence of events $\{E_i\}_{i \geq 1}$

$$P\left(\bigcup_{i=1}^{\infty} E_i\right) \leq \sum_{i=1}^{\infty} P(E_i)$$

Proof (2.2) Note the identity:

$$\bigcup_{i=1}^n E_i = E_1 + E_1^c E_2 + \dots + E_1^c E_2^c \dots E_{n-1}^c E_n$$

or this form:

$$\begin{aligned} F_1 &= E_1 \\ F_2 &= E_2 \setminus E_1 \\ &\vdots \\ F_k &= E_k \setminus \bigcup_{i=1}^k E_i \end{aligned}$$

$$P\left(\bigcup_{i=2}^n E_i\right) = P(E_1) + \sum_{i=1}^n P(E_1^c E_2^c \cdots E_{i-1}^c E_i)$$

denote $E_1^c E_2^c \cdots E_{i-1}^c E_i = B_i$, where $P(B_i) \leq P(E_i)$. Thus,

$$P\left(\bigcup_{i=1}^n E_i\right) \leq P(E_1) + P(E_2) + \cdots P(E_n)$$

Ex 2.1

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Continue Properties

Definition Increasing/Decreasing sequence $\{E_n\}$, we **define a new event** $\lim_{n \rightarrow \infty} E_n$ by $\lim_{n \rightarrow \infty} E_n = \bigcup_{n=1}^{\infty} E_n$

5. For decreasing/increasing sequence E_n

$$\lim_{n \rightarrow \infty} P(E_n) = P(\lim_{n \rightarrow \infty} E_n)$$

We prove the case for increasing sequence $\{E_n\}$

$$RHS = P\left(\bigcup_{n=1}^{\infty} E_n\right) = P\left(\bigcup_{n=1}^{\infty} F_n\right) = \sum_{n=1}^{\infty} P(F_n) = \lim_{n \rightarrow \infty} \sum_{k=1}^n P(F_n) = \lim_{n \rightarrow \infty} P(E_n) = LHS$$

Ex 2.2

Suppose we have an infinitely large urn, and an infinite collection of balls labeled as number 1,2,3,4,...

- At 1 min to 12p.m., balls numbered 1 through 10 are placed in the urn and a ball is randomly selected and withdrawn.
- At $\frac{1}{2}$ min to 12p.m., balls numbered 11 through 20 are placed in the urn and a ball is randomly selected and withdrawn.
- At $\frac{1}{4}$ min to 12p.m., balls numbered 21 through 30 are placed in the urn and a ball is randomly selected and withdrawn.

and so on. How many balls are there in the urn at 12pm.

Proof Consider 1^{th} ball, denote event $\{1^{th} \text{ ball is still in the urn at } \frac{1}{2^k} \text{ min to 12pm}\}$ as E_k , and event $\{1^{th} \text{ ball is in the urn at 12pm}\}$ as E apparently,

$$E_n \subseteq \cdots E_2 \subseteq E_1$$

$$\begin{aligned}
P(E) &= \lim_{n \rightarrow \infty} P(\bigcup_{k=1}^n E_k) \\
&= \lim_{n \rightarrow \infty} P(E_n) \\
&= \lim_{k \rightarrow \infty} \frac{9}{10} \frac{18}{19} \frac{27}{28} \cdots \frac{9k}{9k+1} \\
&= \exp\left(\sum_{k=1}^{\infty} \ln\left(1 - \frac{1}{9k+1}\right)\right) \rightarrow 0
\end{aligned}$$

Similarly, the event $\{i^{th} \text{ is in the urn at 12pm}\}$, denoted by F_i , $P(F_i) = 0$.

$\{\text{the urn is not empty}\} \Leftrightarrow \{\text{there is at least one ball in the urn}\}$. Finally, the urn is empty.

Uniform Probability measure on finite sample

Poker Problem

52 cards to 4 people. What is the probability that

- one of the players receives all 13 spades. (E_1)
- each player receives 1 ace. (E_2)

(1)

$$P(E_1) = \frac{|E|}{|S|} = \frac{\binom{39}{13 \ 13 \ 13}}{\binom{52}{13 \ 13 \ 13 \ 13}}$$

(2)

$$P(E_2) = \frac{\binom{4}{1 \ 1 \ 1 \ 1} \binom{48}{12 \ 12 \ 12 \ 12}}{\binom{52}{13 \ 13 \ 13 \ 13}}$$

Birthday Problem

The Matching Problem

Each of N men throw his hat into the center of room, and the hats are first mixed up. Then each man randomly selects a hat. What is the probability that none of the men selects his own hat?

Solution Denote the event 'the i^{th} man gets his hat' as E_i . According to *Inclusive&Exclusive Theorem*

$$P\left(\bigcup_{i=1}^N E_i\right) = \sum_{i=1}^N P(E_i) - \sum_{1 \leq i_1, i_2 \leq N} P(E_{i_1} E_{i_2}) + \cdots + (-1)^{k-1} \sum_{1 \leq i_1, i_2, \dots, i_k \leq N} P(E_{i_1} E_{i_2} \cdots E_{i_k}) + \cdots + (-1)^{N-1} P(E_1 E_2 \cdots E_N)$$

note that

$$P(E_{i_1} E_{i_2} \cdots E_{i_k}) = \frac{(N-k)!}{N!}, \binom{N}{k} \frac{(N-k)!}{N!} = \frac{1}{k!}$$

then

$$P(\bigcup_{i=1}^N E_i) = 1 - \frac{1}{2!} + \frac{1}{3!} - \dots + (-1)^{N-1} \frac{1}{N!}$$

The event "none of men selects his own hat" and E is complementary.

Chapter III

Conditional Probability
