Homework 6

Qianbo Wang uni: qw2180

Consider the ChickWeight data in R. The body weights of the chicks were measured at birth (i.e., time=0) and every second day thereafter until day 20. They were also measured on day 21. There were four groups of chicks on different protein diets.

Problem 1

Determine whether there is a significant difference in the mean weights of the four groups on Day 18.

• Without adjusting for Birth Weight.

Since without adjusting for Birth Weight, the model is just simple one-way anova. The anova table is as follows:

Anova without adjusting for Birth Weight

```
Call:
```

```
aov(formula = weight ~ Diet, data = sub_day18)
```

Terms:

Diet Residuals Sum of Squares 36690.44 114840.84 Deg. of Freedom 3 43

Residual standard error: 51.67898 Estimated effects may be unbalanced

Df Sum Sq Mean Sq F value Pr(>F)
Diet 3 36690 12230 4.579 0.0072 **

Residuals 43 114841 2671

Signif. codes: 0 *** 0.001 ** 0.01 * 0.05 . 0.1 1

Since for F-test, the null and alternative hypothesis are:

$$H_0: \mu_1 = \mu_2 = \mu_3 = \mu_4$$

 $H_1: \mu_i \neq \mu_j$ for at least one pair (i, j), $i \neq j$

And the p-value=0.0072<0.05. Then we should reject the null hypothesis, i.e. there is a significant difference in the mean weights of the four groups on Day 18.

• Adjusting for Birth Weight. Give the LS Means (i.e., adjusted for Birth Weight).

Since we need do anova analysis based on adjusting for Birth Weight, then we simply add a new variable as Birth Weight, i.e. Give the LS means. Then this time the model is anovca. The anovca table is as follows:

Anova with adjusting for Birth Weight

Call:

aov(formula = weight ~ birthweight + Diet, data = sub_day18)

Terms:

birthweight Diet Residuals 15728.23 27190.80 108612.25 Sum of Squares Deg. of Freedom 3

Residual standard error: 50.85279 Estimated effects may be unbalanced

Df Sum Sq Mean Sq F value Pr(>F) birthweight 1 15728 15728 6.082 0.0178 * 27191 3.505 0.0234 * 3 9064

42 108612 2586 Residuals

Diet

Signif. codes: 0 *** 0.001 ** 0.01 * 0.05 . 0.1

And the Birth Weight means of the four groups is as follows: Since this time for F-test, the null and

Table 1: Birth Weight LS Means

Diet	Birth Weight mean	LS mean
1	41.58824	164.2240
2	40.70000	183.2445
3	40.70000	229.7409
4	41.00000	201.7337

alternative hypothesis are:

$$H_0: \hat{\mu_1} = \hat{\mu_2} = \hat{\mu_3} = \hat{\mu_4}$$

 $H_1: \hat{\mu_i} \neq \hat{\mu_j}$ for at least one pair (i, j), $i \neq j$

And the p-value for Birth Weight is p-value = 0.0178 < 0.05 and p-value for Diet is p-value = 0.0178 < 0.050.0234 < 0.05. Then we should reject the null hypothesis, i.e. there is a significant difference in the mean weights of the four groups on Day 18 regarding the Birth Weight differences.

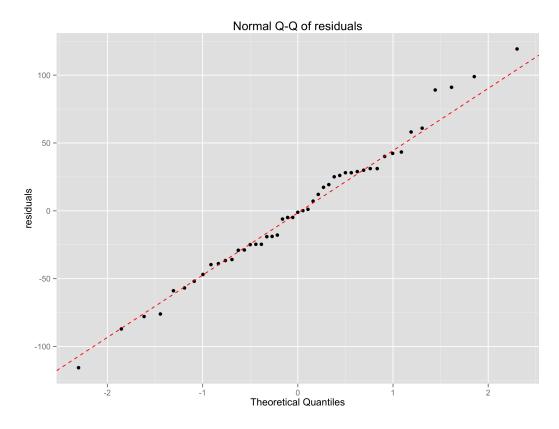
• Check the validity of your assumptions, including parallelism. Suggest measures that you would take if the assumptions are not satisfied.

Since the assumptions are:

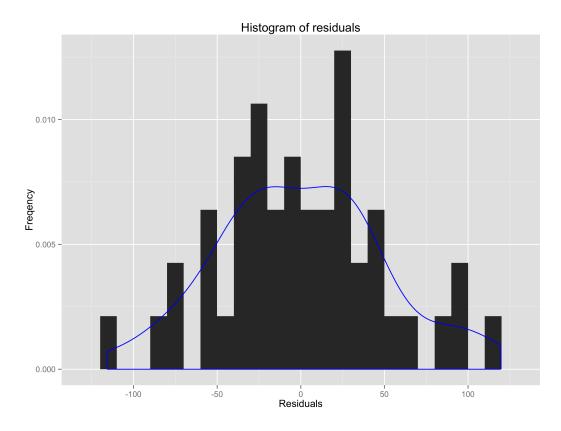
- i.i.d. normal
- constant variance

1. Check i.i.d. Normality

First, plot Q-Q plot on residuals, the plot is as follows:



Then, plot histogram with density on residuals, the plot is as follows:



Since from the plot, the residuals seem to satisfy normal distribution. Then do Shapiro-Wilk test on residuals, The null and alternative hypothesis of Shapiro-Wilk test are:

$$H_0: \epsilon_i \sim \text{Normal}$$

 $H_1: \epsilon_i$ not from Normal

And the Shapiro-Wilk test result is as follows:

Shapiro-Wilk Test on residuals

Shapiro-Wilk normality test

data: residual\$residuals
W = 0.9886, p-value = 0.9242

Since p-value=0.9242, then we should not reject the null, i.e. the residuals satisfy normal distribution.

Then, check the i.i.d. assumption, use nonparametric Kruskal-Wallis test on raw data. The null and alternative hypothesis of Kruskal-Wallis test are:

$$H_0: w_1, w_2, w_3, w_4$$
 i.i.d.

 $H_1: w_1, w_2, w_3, w_4$ not i.i.d.

The result is as follows:

Kruskal-Wallis test on raw data

Kruskal-Wallis rank sum test

data: weight by Diet

Kruskal-Wallis chi-squared = 10.6234, df = 3, p-value = 0.01395

Since p-value=0.01395, then we should reject the null, i.e. the weight data in the four Diet group are not from the same distribution.

2. Check Constant Variance

First, use Bartlett's on raw data to check the constant variance. The null and alternative hypothesis of Bartlett's test are:

$$H_0: {\sigma_1}^2 = \dots = {\sigma_4}^2$$

$$H_1: \sigma_i^2 \neq \sigma_j^2$$
 for at least one pair (i, j), $i \neq j$

And the Bartlett's test is highly dependent on the normal assumption, since we have check the validation of normality, then we can use Bartlett's test. The result is as follows:

Bartlett's test on raw data

Bartlett test of homogeneity of variances

data: sub_day18\$weight and sub_day18\$Diet
Bartlett's K-squared = 3.576, df = 3, p-value = 0.311

Since p-value=0.3111>0.05, then we should not reject the null, i.e. the constant variance assumption satisfies.

Then use Levene's Test on raw data to check the constant variance. The null and alternative hypothesis of Levene's test are:

$$H_0:{\sigma_1}^2=\ldots={\sigma_4}^2$$

$$H_1: \sigma_i^2 \neq \sigma_j^2$$
 for at least one pair (i, j), $i \neq j$

The result is as follows:

Levene's test on raw data

```
Levene's Test for Homogeneity of Variance (center = median)

Df F value Pr(>F)
group 3 1.059 0.3764

43
```

Since p-value=0.3764>0.05, then we should not reject the null, i.e. the constant variance assumption satisfies.

2. Check parallelism

Use two-way anova to check the parallelism, the null and alternative hypothesis are:

 H_0 : there is joint effect with Diet and Birth Weight

 H_1 : there is no joint effect with Diet and Birth Weight

The result is as follows:

Two-way Anova for parallelism check

```
Df Sum Sq Mean Sq F value Pr(>F)
birthweight
                     15728
                              15728
                                      6.134 0.0177 *
                  1
Diet
                     27191
                               9064
                                      3.535 0.0234 *
                  3
                  3
                      8610
                               2870
                                      1.119 0.3530
birthweight:Diet
                 39 100002
Residuals
                               2564
Signif. codes: 0 *** 0.001 ** 0.01 * 0.05 . 0.1
```

And the p-value = 0.3530 > 0.05. Then we should not reject the null hypothesis, i.e. there is no interaction effect between Diet group and the Birth Weight. So we do not need do marginal inference on BirthWeight and Diet.

Problem 2

Perform an appropriate repeated measures ANOVA to determine whether there is a significant difference in the mean weights of the four groups using the measurements on Days 10, 18, and 21.

• Do the analyses assuming compound symmetry and unstructured covariance structures and compare the results.

1. Compound Symmetry

For this part of the problem, I simply use SAS to do the repeat measures ANOVA because SAS has a easily-used implementation of repeated measures ANOVA based on different structure assumptions on covariance. Since for the assumption, it assumes that the covariance between the data is compound symmetry or unstructured. Then the result is in the following page.

2. Unstructured Covariance

Since the assumption is unstructured covariance, then the result is in the following page.

Repeated Measures on Chick Weight compound symmetry assumption

23:48 Wednesday, October 21, 2015 **1**

The Mixed Procedure

Model Information		
Data Set	REPEAT.REPEATDATA	
Dependent Variable	weight	
Covariance Structure	Compound Symmetry	
Subject Effect	Chick(Diet)	
Estimation Method	REML	
Residual Variance Method	Profile	
Fixed Effects SE Method	Model-Based	
Degrees of Freedom Method	Between-Within	

Class Level Information				
Class	Levels	Values		
Diet	4	1234		
Chick	49	1 10 11 12 13 14 15 16 17 19 2 20 21 22 23 24 25 26 27 28 29 3 30 31 32 33 34 35 36 37 38 39 4 40 41 42 43 44 45 46 47 48 49 5 50 6 7 8 9		
Time	3	10 18 21		

Dimensions		
Covariance Parameters	2	
Columns in X	20	
Columns in Z	0	
Subjects	49	
Max Obs Per Subject	3	

Number of Observations		
Number of Observations Read 141		
Number of Observations Used	141	
Number of Observations Not Used	0	

Iteration History				
Iteration	Evaluations	-2 Res Log Like	Criterion	
0	1	1396.07743029		
1	2	1338.99956090	0.00000228	
2	1	1338.99828947	0.00000000	

Convergence criteria met.

Repeated Measures on Chick Weight compound symmetry assumption

23:48 Wednesday, October 21, 2015 **2**

Estimated R Correlation Matrix for Chick(Diet) 1 1			
Row	Col1	Col2	Col3
1	1.0000	0.6626	0.6626
2	0.6626	1.0000	0.6626
3	0.6626	0.6626	1.0000

Covariance Parameter Estimates			
Cov Parm Subject Estimate			
cs	Chick(Diet)	1534.18	
Residual		781.13	

Fit Statistics		
-2 Res Log Likelihood	1339.0	
AIC (smaller is better)	1343.0	
AICC (smaller is better)	1343.1	
BIC (smaller is better)	1346.8	

Null Model Likelihood Ratio Test		
DF Chi-Square Pr > Chi		Pr > ChiSq
1	57.08	<.0001

Type 3 Tests of Fixed Effects				
Effect	Num DF	Den DF	F Value	Pr > F
Diet	3	45	5.95	0.0016
Time	2	84	192.58	<.0001
Diet*Time	6	84	3.76	0.0023

Repeated Measures on Chick Weight unstructured correlation assumption

23:48 Wednesday, October 21, 2015 **1**

Model Information		
Data Set	REPEAT.REPEATDATA	
Dependent Variable	weight	
Covariance Structure	Unstructured	
Subject Effect	Chick(Diet)	
Estimation Method	REML	
Residual Variance Method	None	
Fixed Effects SE Method	Model-Based	
Degrees of Freedom Method	Between-Within	

	Class Level Information				
Class	Levels	Values			
Diet	4	1234			
Chick	49	1 10 11 12 13 14 15 16 17 19 2 20 21 22 23 24 25 26 27 28 29 3 30 31 32 33 34 35 36 37 38 39 4 40 41 42 43 44 45 46 47 48 49 5 50 6 7 8 9			
Time	3	10 18 21			

Dimensions			
Covariance Parameters	6		
Columns in X	20		
Columns in Z	0		
Subjects	49		
Max Obs Per Subject	3		

Number of Observations		
Number of Observations Read	141	
Number of Observations Used	141	
Number of Observations Not Used	0	

Iteration History						
Iteration Evaluations -2 Res Log Like C						
0	1	1396.07743029				
1	3	1177.78560780	0.00103860			
2 1		1177.29205471	0.00004402			
3 1		1177.27077572	0.0000009			
4	1	1177.27073492	0.00000000			

Repeated Measures on Chick Weight unstructured correlation assumption

23:48 Wednesday, October 21, 2015 **2**

The Mixed Procedure

Convergence criteria met.

Estimated R Correlation Matrix for Chick(Diet) 1 1							
Row	Row Col1 Col2 Col3						
1	1.0000	0.7988	0.6615				
2	0.7988	1.0000	0.9632				
3	0.6615	0.9632	1.0000				

Covariance Parameter Estimates					
Cov Parm Subject Estim					
UN(1,1)	Chick(Diet)	428.98			
UN(2,1)	Chick(Diet)	882.94			
UN(2,2)	Chick(Diet)	2848.34			
UN(3,1)	Chick(Diet)	895.29			
UN(3,2)	Chick(Diet)	3359.26			
UN(3,3)	Chick(Diet)	4270.54			

Fit Statistics			
-2 Res Log Likelihood	1177.3		
AIC (smaller is better)	1189.3		
AICC (smaller is better)	1190.0		
BIC (smaller is better)	1200.6		

Null Model Likelihood Ratio Test					
DF	Pr > ChiSq				
5	218.81	<.0001			

Type 3 Tests of Fixed Effects					
Effect	Num DF	Den DF	F Value	Pr > F	
Diet	3	45	6.03	0.0015	
Time	2	45	100.88	<.0001	
Diet*Time	6	45	2.47	0.0375	

Compare the result:

Since both under compound symmetry covariance and unstructured covariance assumption, all of the p-values are significant, then we should conclude that the chicken have significant different of weight on at least two Diet group and between Day 10, Day18 and Day 21. And the differences between the two method on this test is just the different estimation on the covariance structure.

• Check the validity of your assumptions.

1. Normality

Since the assumptions for repeated measures are normality and constant variance. Then under the compound symmetry assumption and unstructured covariance assumption, the plots of the residuals given by SAS are as follows: Since from the plots, we see that the data is approximately normally distributed. Then do normal test on residuals, the result is also on the following pages.

Since under compound symmetry covariance assumption, for Shapiro-Wilk test on residuals gives the p-value=0.0212<0.05, then we should not reject the null, i.e. the normality assumption doesn't satisfy, i.e. it is not normal.

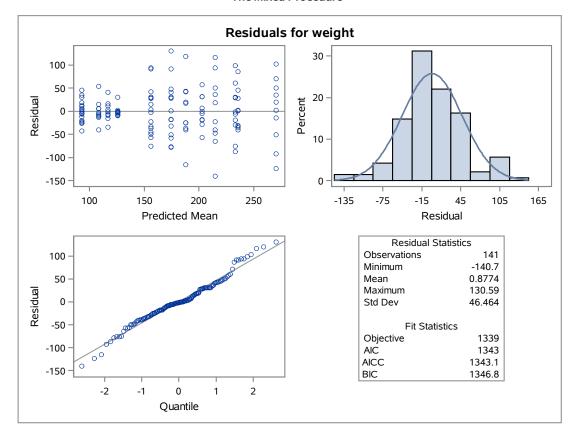
And under unstructured covariance assumption, for Shapiro-Wilk test on residuals gives the p-value=0.0119>0.05, then we should not reject the null, i.e. the normality assumption doesn't satisfy, i.e. it is not normal.

2. Constant Variance

Since normality assumption satisfies then use Bartlett's test on raw data gives the p-value = 0.0561 > 0.05, then we should not reject the null, i.e. the Constant Variance assumption satisfies.

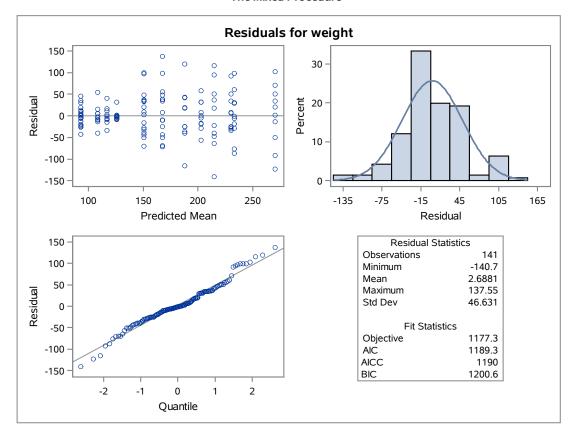
Repeated Measures on Chick Weight compound symmetry assumption

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Repeated Measures on Chick Weight unstructured correlation assumption

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Noramlity test on residuals compound symmetry assumption

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The UNIVARIATE Procedure Variable: Resid (Residual)

Moments					
N	141	Sum Weights	141		
Mean	0.87741402	Sum Observations	123.715377		
Std Deviation	46.4638738	Variance	2158.89157		
Skewness	0.09585474	Kurtosis	0.97181334		
Uncorrected SS	302353.37	Corrected SS	302244.82		
Coeff Variation	5295.5472	Std Error Mean	3.91296406		

Basic Statistical Measures						
Loc	ation	Variability				
Mean 0.8774		Std Deviation	46.46387			
Median -1.1000		Variance	2159			
Mode	-17.4135	Range	271.28652			
		Interquartile Range	54.60000			

Tests for Location: Mu0=0						
Test Statistic p Value						
Student's t	t	0.224233	Pr > t	0.8229		
Sign	М	-4.5	Pr >= M	0.5006		
Signed Rank	s	-51.5	Pr >= S	0.9160		

Topts for Normality					
Tests for Normality					
Test	Statistic p Value				
Shapiro-Wilk	w	0.977752	Pr < W	0.0212	
Kolmogorov-Smirnov	D	0.080669	Pr > D	0.0235	
Cramer-von Mises	W-Sq	0.187511	Pr > W-Sq	0.0078	
Anderson-Darling	A-Sq	1.136853	Pr > A-Sq	0.0056	

Quantiles (Definition 5)				
Quantile	Estimate			
100% Max	130.5865			
99%	119.3000			
95%	91.6745			
90%	50.7000			
75% Q3	29.9000			
50% Median	-1.1000			

Noramlity test on residuals unstructured correlation assumption

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The UNIVARIATE Procedure Variable: Resid (Residual)

Moments						
N	141					
Mean	2.68808255	Sum Observations	379.019639			
Std Deviation	46.6314288	Variance	2174.49016			
Skewness	0.11916932	Kurtosis	1.05114803			
Uncorrected SS	305447.458	Corrected SS	304428.622			
Coeff Variation	1734.74691	Std Error Mean	3.92707474			

Basic Statistical Measures					
Location Variability					
Mean	2.6881	Std Deviation	46.63143		
Median	-1.1000	Variance	2174		
Mode	-10.4541	Range	278.24589		
		Interquartile Range	55.80000		

Tests for Location: Mu0=0						
Test	Statistic p Value					
Student's t	t	0.6845	Pr > t	0.4948		
Sign	м	-4.5	Pr >= M	0.5006		
Signed Rank	s	183	Pr >= S	0.7079		

Tests for Normality						
Test Statistic p Value						
Shapiro-Wilk	w	0.975347	Pr < W	0.0119		
Kolmogorov-Smirnov	D	0.079551	Pr > D	0.0273		
Cramer-von Mises	W-Sq	0.203494	Pr > W-Sq	<0.0050		
Anderson-Darling	A-Sq	1.25304	Pr > A-Sq	<0.0050		

Quantiles (Definition 5)					
Quantile	Estimate				
100% Max	137.5459				
99%	119.3000				
95%	97.1833				
90%	54.5000				
75% Q3	31.1000				
50% Median	-1.1000				

Constant Variance test on raw data

23:48 Wednesday, October 21, 2015 **3**

The GLM Procedure

Bartlett's Test for Homogeneity of weight Variance					
Source	Pr > ChiSq				
Diet	3	7.5574	0.0561		

Constant Variance test on raw data

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The GLM Procedure

Levene's Test for Homogeneity of weight Variance ANOVA of Squared Deviations from Group Means						
Source DF Squares Square F Value Pr >						
Diet	3	3.2505E8	1.0835E8	3.90	0.0103	
Error	137	3.804E9	27766281			

R Code:

```
rm(list=ls())
data('ChickWeight')
#avova on day 18
sub_day18<-subset(ChickWeight,Time==18)
anova_day18<-aov(data = sub_day18, weight ~ Diet)
summary(anova_day18)
sink('/Users/raymond/Drive/STAT W4201/HW6/anova.txt')
anova_day18
summary(anova_day18)
sink()
#anova adjust by LS mean on day 18
sub_day0<-subset(ChickWeight,Time==0)</pre>
sub_day18[,'birthweight']<-sub_day0$weight[match(sub_day18$Chick,sub_day0$Chick)]
group_mean<-aggregate(sub_day18$birthweight,list(sub_day18$Diet),mean)
colnames(group_mean)<-c('Diet','birthweight_mean')</pre>
sub_day18<-merge(sub_day18,group_mean,by.y = 'Diet')</pre>
anova_adjust_day18<-aov(data=sub_day18, weight ~ birthweight + Diet)
sink(', /Users/raymond/Drive/STAT W4201/HW6/anova_adjust.txt')
anova_adjust_day18
summary(anova_adjust_day18)
sink()
coef<-anova_adjust_day18$coefficients[2]
v_hat<-anova_adjust_day18$fitted.values
y_frame<-data.frame(yhat=y_hat,Diet=sub_day18$Diet)
yhat_mean<-aggregate(y_frame$yhat, list(y_frame$Diet),mean)
colnames(yhat_mean)<-c('Diet','yhat')</pre>
mu<-yhat_mean$yhat-coef*(group_mean$birthweight_mean-mean(sub_day18$birthweight))
lsmean_result<-cbind(group_mean,lsmean=mu)
sink('/Users/raymond/Drive/STAT W4201/HW6/LSmeans.txt')
lsmean result
sink()
#Or simply use Ismeans function in library Ismeans
library (lsmeans)
lsmean<-lsmeans(anova_adjust_day18,'Diet')</pre>
sink('/Users/raymond/Drive/STAT W4201/HW6/LSmeans1.txt')
lsmean
sink()
library (ggplot2)
#check assumptions
#Normality
#Q-Q plot for residuals
residual < -data.frame(anova_day18['residuals'],sub_day18['Diet'])
normal_plot<-ggplot(data=residual,aes(sample=residuals))+
  stat_qq()+ggtitle('Normal Q-Q of residuals')+
 ylab('residuals')+xlab('Theoretical Quantiles')
resid_quantile < -quantile (residual residual constant (0.25, 0.75))
norm_quantile < -qnorm(c(0.25, 0.75))
slope < - diff(resid_quantile) / diff(norm_quantile)
inter < -resid\_quantile[1] - slope*norm\_quantile[1]
normal_plot+geom_abline(slope=slope,intercept=inter,linetype=2,color='red')
ggsave(filename = '/Users/raymond/Drive/STAT W4201/HW6/qqplot.png')
#histogram for residuals
histogram_plot<-ggplot(data=residual,aes(residuals))+
 geom_bar(binwidth=10,aes(y=..density..))+ggtitle("Histogram of residuals")+
 ylab('Frequency')+xlab('Residuals')
histogram_plot+geom_density(color='blue')
ggsave(filename = '/Users/raymond/Drive/STAT W4201/HW6/histogram.png')
#shapiro test on residuals
```

```
shapiro. test (residual residuals)
sink('/Users/raymond/Drive/STAT W4201/HW6/normalcheck.txt')
shapiro. test (residual $ residuals )
sink()
#check constant variance
#bartlett test on data
bartlett . test (x=sub_day18$weight,g=sub_day18$Diet)
sink('/Users/raymond/Drive/STAT W4201/HW6/constantcheck.txt')
bartlett.test(x=\!\!sub\_day18\$weight,g=\!\!sub\_day18\$Diet)
sink()
#levene test on data
library (car)
leveneTest(y=sub_day18$weight,group=as.factor(sub_day18$Diet))
#nonparametric check iid assumptions
#kruskal test on data
kruskal.test(weight~Diet,data=sub_day18)
sink('/Users/raymond/Drive/STAT W4201/HW6/identicalcheck.txt')
kruskal.test(weight~Diet,data=sub_day18)
sink()
#check for parallelism
#two-way anova
anova_inter<-aov(weight ~ birthweight*Diet,data = sub_day18)
summary(anova_inter)
sink('/Users/raymond/Drive/STAT W4201/HW6/paracheck.txt')
summary(anova_inter)
sink()
#repeated measures
#anova on day 10, 18, 21
sub_repeat<-subset(ChickWeight,Time %in% c(10,18,21))
repeated <- aov (weight ~ Diet*Time+Error(Chick), data=sub_repeat)
write.table(sub_repeat, file = '/Users/raymond/Drive/STAT W4201/HW6/ChickWeight.csv',col.names=TRUE,row.names=
        FALSE, sep=",")
```

SAS code:

```
libname repeat 'C:\Users\QIANBO\Desktop\HW6';
*import data from csv;
proc import out=repeat.repeatdata
    Datafile = 'C:\Users\QIANBO\Desktop\HW6\ChickWeight.csv'
   Dbms =csv replace;
run:
proc print data = repeat.repeatdata;
run;
ods listing close;
ods graphics on;
options papersize=letter;
ods pdf file = 'C:\Users\QIANBO\Desktop\HW6\compoundsy.pdf';
title 'Repeated Measures on Chick Weight';
title2 'compound symmetry assumption';
*correlation compound symmetry assumption;
proc mixed data = repeat.repeatdata;
class Diet Chick Time;
model weight= Diet Time Diet*Time / outp=repeat.csparam residual;
repeated/type=cs sub=Chick(Diet) rcorr;
run:
ods pdf close;
ods pdf file = 'C:\Users\QIANBO\Desktop\HW6\unstructured.pdf';
title 'Repeated Measures on Chick Weight';
title2 'unstructured correlation assumption';
*correlation unstructured assumption:
proc mixed data = repeat.repeatdata;
class Diet Chick Time;
model weight = Diet Time Diet*Time / outp=repeat.unparam residual;
repeated/type=un sub=Chick(Diet) rcorr;
run;
ods pdf close;
ods graphics off;
ods pdf file = 'C:\Users\QIANBO\Desktop\HW6\normalcs.pdf';
title 'Noramlity test on residuals';
title2 'compound symmetry assumption';
*test normality;
proc univariate data=repeat.csparam normal;
var resid;
run;
ods pdf close;
ods pdf file = 'C:\Users\QIANBO\Desktop\HW6\normalun.pdf';
title 'Noramlity test on residuals';
title2 'unstructured correlation assumption';
*test normality;
proc univariate data=repeat.unparam normal;
var resid;
run;
ods pdf close;
ods pdf file = 'C:\Users\QIANBO\Desktop\HW6\homo.pdf';
title 'Constant Variance test on raw data';
title2;
proc glm data=repeat.repeatdata;
class Diet Chick Time;
model weight = Diet;
means Diet / HOVTEST=BARTLETT;
means Diet / HOVTEST=LEVENE;
ods pdf close;
ods listing;
title;
```