

CSE 527 Problem Set 1

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1.

a)

The chances that you actually have the disease is denoted by, (D=Disease, P=Positive)

$$\begin{aligned} P(D|P) &= \frac{P(PD)}{P(P)} = \frac{P(P|D) \times P(D)}{P(\bar{D}) \cdot P(P|\bar{D}) + P(D) \cdot P(P|D)} = \frac{P(P|D) \times P(D)}{(1 - P(D)) \cdot (1 - P(\bar{P}|\bar{D})) + P(D) \cdot P(P|D)} \\ &= \frac{l \cdot P(D)}{k + m \cdot P(D)} \end{aligned}$$

Where, $l = P(P|D)$, $k = 1 - P(\bar{P}|\bar{D})$, $m = P(P|D) + P(\bar{P}|\bar{D}) - 1$. Thus, k, l, m are all positive numbers.

Take derivative,

$$\frac{dP(D|P)}{dP(D)} = \frac{l \cdot (k + m \cdot P(D)) - m \cdot l \cdot P(D)}{(k + m \cdot P(D))^2} = \frac{k \cdot l}{(k + m \cdot P(D))^2}$$

The derivative is poitive, so as $P(D)$ decrease, $P(D|P)$ decrease, Thus the disease is rare is a good news since $P(D|P)$ is lower.

Where, $P(P|D) = 0.99$, $P(\bar{P}|\bar{D}) = 0.99$, $P(D) = 1/10,000$

$$P(D|P) = 0.0098$$

b)

According to the Bayes' rule:

$$\begin{aligned} P(A|B, E)P(B|E) &= \frac{P(A, B, E)}{P(B, E)} \times \frac{P(B, E)}{P(E)} \\ &= \frac{P(A, B, E)}{P(E)} \\ &= P(A, B|E) \end{aligned}$$

$$\begin{aligned} \frac{P(B|A, E)P(A|E)}{P(B|E)} &= \frac{\frac{P(A, B, E)}{P(A, E)} \frac{P(A, E)}{P(E)}}{\frac{P(B, E)}{P(E)}} \\ &= \frac{P(A, B, E)}{P(B, E)} \\ &= P(A|B, E) \end{aligned}$$

2.

a)

Denote 1 as up, 0 as down.

x2, x3	x1 up	x1 down
1, 1	$\theta_{x1 1,1}$	$1 - \theta_{x1 1,1}$
1, 0	$\theta_{x1 1,0}$	$1 - \theta_{x1 1,0}$
0, 1	$\theta_{x1 0,1}$	$1 - \theta_{x1 0,1}$
0, 0	$\theta_{x1 0,0}$	$1 - \theta_{x1 0,0}$

2.

b).

$$\begin{aligned} \mathcal{L}(\theta; D) &= P(D|\theta) \\ &= \prod_i P(D_i|\theta) \\ &= \prod_i P(X_{1i}, X_{2i}, X_{3i}|\theta) \\ &= \prod_i \left(P(X_{2i}|\theta) \right. \\ &\quad \left. P(X_{3i}|\theta) \right. \\ &\quad \left. P(X_{1i}|X_{2i}, X_{3i}|\theta) \right) \\ &= \prod_i P(X_{1i}|X_{2i}, X_{3i}) \prod_i P(X_{2i}) \prod_i P(X_{3i}) \end{aligned}$$

$$c). \arg\max_{\theta} \mathcal{L}(\theta; D) = \arg\max_{\theta} \ln(\mathcal{L}(\theta; D)) = \arg\max_{\theta} l(\theta).$$

$$\text{Since } X_k \text{ is binary, } \therefore P(X_k) = \begin{cases} 1 & \theta_{X_k|P_k} \\ 0 & 1 - \theta_{X_k|P_k} \end{cases}$$

$$\therefore l(\theta) = \sum_k [M_{X_k|P_k} \log \theta_{X_k|P_k} + (M_{P_k} - M_{X_k|P_k}) \log (1 - \theta_{X_k|P_k})] \quad (k=1, 2, \dots, N)$$

In order to $\arg\max l(\theta)$, we take derivative w.r.t. each $\theta_{X_k|P_k}$ and set to zero.

$$\frac{dl(\theta)}{d\theta_{X_k|P_k}} = 0, \text{ we have, } \frac{M_{X_k|P_k}}{\theta_{X_k|P_k}} - \frac{M_{P_k} - M_{X_k|P_k}}{1 - \theta_{X_k|P_k}} = 0$$

$$\text{So, } \hat{\theta}_{X_k|P_k} = \frac{M_{X_k|P_k}}{M_{P_k}}.$$

3.

a). Denote $G_{A|B} = A$, $G_{A|C} = B$, $G_{B|C} = C$,

In model 1: $\hat{\theta}_A, \hat{\theta}_{B|A}=1, \hat{\theta}_{C|A}=0, \hat{\theta}_{C|B}=1, \hat{\theta}_{C|B}=0$.

In model 2: $\hat{\theta}_A, \hat{\theta}_B, \hat{\theta}_{C|A}=1, \hat{\theta}_{C|B}=0, \hat{\theta}_{C|B}=1, \hat{\theta}_{C|B}=0$.

b). For model 1:

$$\begin{aligned} \mathcal{L}(\theta; D) &= \prod_i P(A_i) P(B_i|A_i) P(C_i|B_i) \\ &= \theta_A^{M_A} (1 - \theta_A)^{M - M_A} \cdot \theta_{B|A=1}^{M_{B|A=1}} (1 - \theta_{B|A=1})^{M_{A=1} - M_{B|A=1}} \cdot \theta_{B|A=0}^{M_{B|A=0}} (1 - \theta_{B|A=0})^{M_{A=0} - M_{B|A=0}} \\ &\quad \cdot \theta_{C|B=1}^{M_{C|B=1}} (1 - \theta_{C|B=1})^{M_{B=1} - M_{C|B=1}} \cdot \theta_{C|B=0}^{M_{C|B=0}} (1 - \theta_{C|B=0})^{M_{B=0} - M_{C|B=0}} \end{aligned}$$

For model 2:

$$\begin{aligned} \mathcal{L}(\theta; D) &= \prod_i P(A_i) P(B_i) P(C_i|A_i|B_i) \\ &= \theta_A^{M_A} (1 - \theta_A)^{M - M_A} \cdot \theta_B^{M_B} (1 - \theta_B)^{M - M_B} \\ &\quad \times \theta_{C|A=1, B=1}^{M_{C|A=1, B=1}} (1 - \theta_{C|A=1, B=1})^{M_{A=1, B=1} - M_{C|A=1, B=1}} \times \theta_{C|A=1, B=0}^{M_{C|A=1, B=0}} (1 - \theta_{C|A=1, B=0})^{M_{A=1, B=0} - M_{C|A=1, B=0}} \\ &\quad \times \theta_{C|A=0, B=1}^{M_{C|A=0, B=1}} (1 - \theta_{C|A=0, B=1})^{M_{A=0, B=1} - M_{C|A=0, B=1}} \times \theta_{C|A=0, B=0}^{M_{C|A=0, B=0}} (1 - \theta_{C|A=0, B=0})^{M_{A=0, B=0} - M_{C|A=0, B=0}} \end{aligned}$$

c. Mathematically we should take derivative w.r.t. each θ and set to zero. Computationally, we can compute the cpd for each node to get all the θ .

d.

Codes are attached in the end.

CPDs for Model 1

Gal80(0)	0.5179
Gal80(1)	0.4821

Gal80	0	1
Gal4(0)	0.3103	0.6667
Gal4(1)	0.6897	0.3333

Gal4	0	1
Gal2(0)	0.6296	0.2931
Gal2(11)	0.3704	0.7069

CPDs for Model 2

Gal80(0)	0.5179
Gal80(1)	0.4821

Gal4(0)	0.4821
Gal4(1)	0.5179

Gal4	0	0	1	1
Gal80	0	1	0	1
Gal2(0)	0.6667	0.6111	0.2	0.5
Gal2(1)	0.3333	0.3889	0.8	0.5

Likelihood for two models.

Model	Likelihood Score
1	-218.535478586
2	-223.129027335

e.

Model 1 has higher score, so it's selected.

Code for 3.d

```

from __future__ import division, absolute_import, print_function
import numpy as np
import pandas as pd
from pgmpy.models import BayesianModel
from pgmpy.factors.discrete import TabularCPD
from pgmpy.estimators import K2Score
###
#Method 1

# Read data
data = pd.read_table('disc-gal80-gal4-gal2.txt', index_col=0).values

# Get params for Model 1
M = len(data[0])
Ma1 = np.count_nonzero(data[0])
Ma0 = M - Ma1
Mb1 = np.count_nonzero(data[1])
Mb0 = M - Mb1
Mb_a1 = np.count_nonzero(data[1][np.where(data[0]==1)])
Mb_a0 = np.count_nonzero(data[1][np.where(data[0]==0)])
Mc_b1 = np.count_nonzero(data[2][np.where(data[1]==1)])
Mc_b0 = np.count_nonzero(data[2][np.where(data[1]==0)])

theta_a = Ma1 / M
theta_b_1 = Mb_a1 / Ma1
theta_b_0 = Mb_a0 / Ma0
theta_c_1 = Mc_b1 / Mb1
theta_c_0 = Mc_b0 / Mb0

#Compute log Likelihood
def local_l(theta, M, M_):

    return M_ * np.log(theta) + (M - M_) * np.log(1 - theta)

L1 = np.sum(local_l(np.array([theta_a, theta_b_1, theta_b_0, theta_c_1, theta_c_0]),
                        np.array([M, Ma1, Ma0, Mb1, Mb0]),
                        np.array([Ma1, Mb_a1, Mb_a0, Mc_b1, Mc_b0])))
print('log likelilhood for model 1 is ' + str(L1))
###
# Get params for Model 2
M = len(data[0])
Ma1 = np.count_nonzero(data[0])
#Ma0 = M - Ma1
Mb1 = np.count_nonzero(data[1])
#Mb0 = M - Mb1
M11 = len(np.intersect1d(np.where(data[0]==1)[0], np.where(data[1]==1)[0]))
M10 = len(np.intersect1d(np.where(data[0]==1)[0], np.where(data[1]==0)[0]))
M01 = len(np.intersect1d(np.where(data[0]==0)[0], np.where(data[1]==1)[0]))
M00 = len(np.intersect1d(np.where(data[0]==0)[0], np.where(data[1]==0)[0]))
Mc_11 = np.count_nonzero(data[2][np.intersect1d(np.where(data[0]==1)[0], np.where(data[1]==1)
[0])])

Mc_10 = np.count_nonzero(data[2][np.intersect1d(np.where(data[0]==1)[0], np.where(data[1]==0)

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[0]))
Mc_01 = np.count_nonzero(data[2][np.intersect1d(np.where(data[0]==0)[0], np.where(data[1]==1)
[0]))
Mc_00 = np.count_nonzero(data[2][np.intersect1d(np.where(data[0]==0)[0], np.where(data[1]==0)
[0]))

theta_a = Ma1 / M
theta_b = Mb1 / M
theta_c_11 = Mc_11 / M11
theta_c_10 = Mc_10 / M10
theta_c_01 = Mc_01 / M01
theta_c_00 = Mc_00 / M00

#Compute Likelihood
L2 = np.sum(local_l(np.array([theta_a, theta_b, theta_c_11, theta_c_10, theta_c_01,
theta_c_00]),
                    np.array([M, M, M11, M10, M01, M00]),
                    np.array([Ma1, Mb1, Mc_11, Mc_10, Mc_01, Mc_00])))
print('log likelilhood for model 2 is ' + str(L2))
###
#Method 2, Using pgmpy

# Read Data
train_data = pd.DataFrame(pd.read_table('disc-gal80-gal4-gal2.txt', index_col=0).values.T,
columns=['Gal80', 'Gal4', 'Gal2'])

# Define Model
model1 = BayesianModel([('Gal80', 'Gal4'), ('Gal4', 'Gal2')])
model2 = BayesianModel([('Gal80', 'Gal2'), ('Gal4', 'Gal2')])

# Fit the data
model1.fit(train_data)
model2.fit(train_data)

# Get CPDs
print('For Model 1')
print(model1.get_cpds('Gal80'))
print(model1.get_cpds('Gal4'))
print(model1.get_cpds('Gal2'))

print('For Model 2')
print(model2.get_cpds('Gal80'))
print(model2.get_cpds('Gal4'))
print(model2.get_cpds('Gal2'))

#Calculate K2 Score
print('K2Score of Model1 and 2')
print(K2Score(train_data).score(model1))
print(K2Score(train_data).score(model2))

```