

Discrete analysis_HW4

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Problem 1

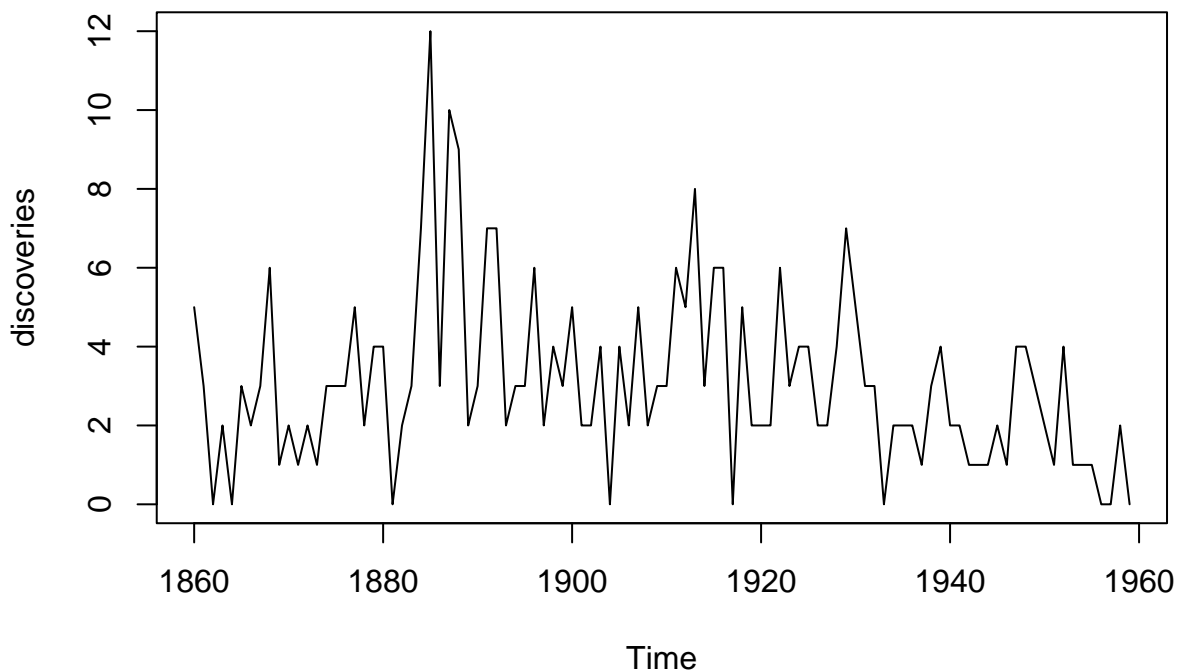
Question 1.

The dataset `discoveries` found in the R base package (use `"data(discoveries)"` to load the dataset in R) lists the numbers of "great" inventions and scientific discoveries in each year from 1860 to 1959 (use `"help(discoveries)"` to get more information about the data set). Has the discovery rate remained constant over time?

Sol.

First, we observe the time series plot of 'discoveries' data by 'ts.plot' function in R:

```
data(discoveries)
ts.plot(discoveries)
```



Averagely, there seems to be a tendency for discovery counts (y_t) to decrease as time increases. Then, with a significant level 0.05, do the hypothesis test to compare two models:

$$H_0 : \text{null model (only intercept)} \text{ v.s. } H_1 : \log(y_t) = \beta_0 + \beta_1 \times \text{time}$$

```
time = 1860:1959
model1 = glm( discoveries ~ 1+ time, data= discoveries, family ="poisson")
summary(model1)
```

Call:

```
glm(formula = discoveries ~ 1 + time, family = "poisson", data = discoveries)
```

Coefficients:

	Estimate	Std. Error	z value	Pr(> z)
(Intercept)	11.354807	3.775677	3.007	0.00264 **
time	-0.005360	0.001982	-2.705	0.00683 **

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

(Dispersion parameter for poisson family taken to be 1)

Null deviance: 164.68 on 99 degrees of freedom
Residual deviance: 157.32 on 98 degrees of freedom
AIC: 430.32

Number of Fisher Scoring iterations: 5

The test statistic, $D_{H_0} - D_{H_1} = 164.68 - 157.32 = 7.36 \geq \chi_{0.05,1}^2 = 3.841459$, so we reject H_0 at level 0.05. That is, it is confident to reject 'discoveries rate is not constant over time' at level 0.05.

Problem 2

Question 2.

The [data](#) was collected in a salmonella reverse mutagenicity assay. The predictor is the dose level of quinoline and the response is the numbers of revertant colonies of TA98 Salmonella observed on each of three replicate plates. Show that a Poisson GLM is inadequate and that some overdispersion must be allowed for. Do not forget to check out other reasons for a high deviance.

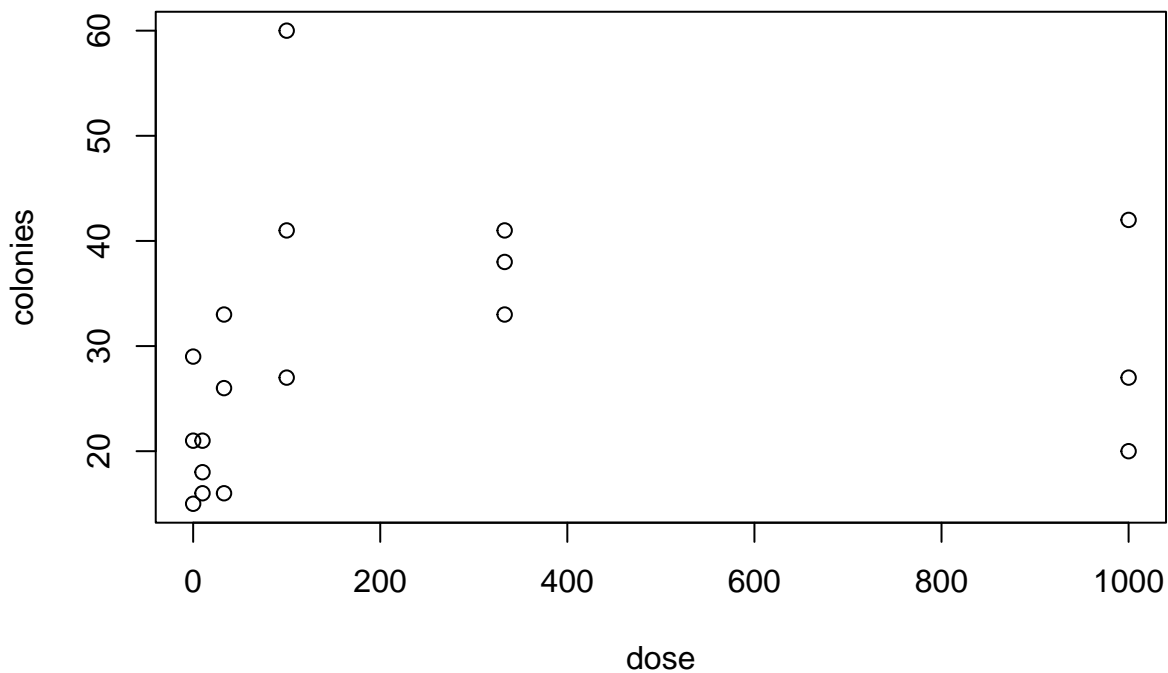
Sol.

```
data = read.table("salmonella.txt")
dim(data) ; unique(data$dose)
```

```
[1] 18 2
```

```
[1] 0 10 33 100 333 1000
```

```
plot(x = data$dose, y = data$colonies, xlab= "dose", ylab = "colonies")
```



First, We should convert dose variable into an ordered factor variable:

```
data$dose = as.ordered(factor(data$dose))
unique(data$dose )
```

```
[1] 0    10   33   100  333  1000
Levels: 0 < 10 < 33 < 100 < 333 < 1000
```

Then, fit the Poisson model with significant level 0.05:

```
model = glm(formula = colonies ~ ., data = data, family = "poisson")
summary(model)
```

Call:

```
glm(formula = colonies ~ ., family = "poisson", data = data)
```

Coefficients:

	Estimate	Std. Error	z value	Pr(> z)
(Intercept)	3.32778	0.04561	72.968	< 2e-16 ***
dose.L	0.50669	0.11530	4.394	1.11e-05 ***
dose.Q	-0.22791	0.11078	-2.057	0.039668 *
dose.C	-0.41331	0.11383	-3.631	0.000282 ***
dose^4	0.15583	0.11263	1.384	0.166501
dose^5	0.13253	0.10577	1.253	0.210187

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

(Dispersion parameter for poisson family taken to be 1)

```
Null deviance: 78.358  on 17  degrees of freedom
Residual deviance: 33.496  on 12  degrees of freedom
AIC: 138.03
```

```
Number of Fisher Scoring iterations: 4
```

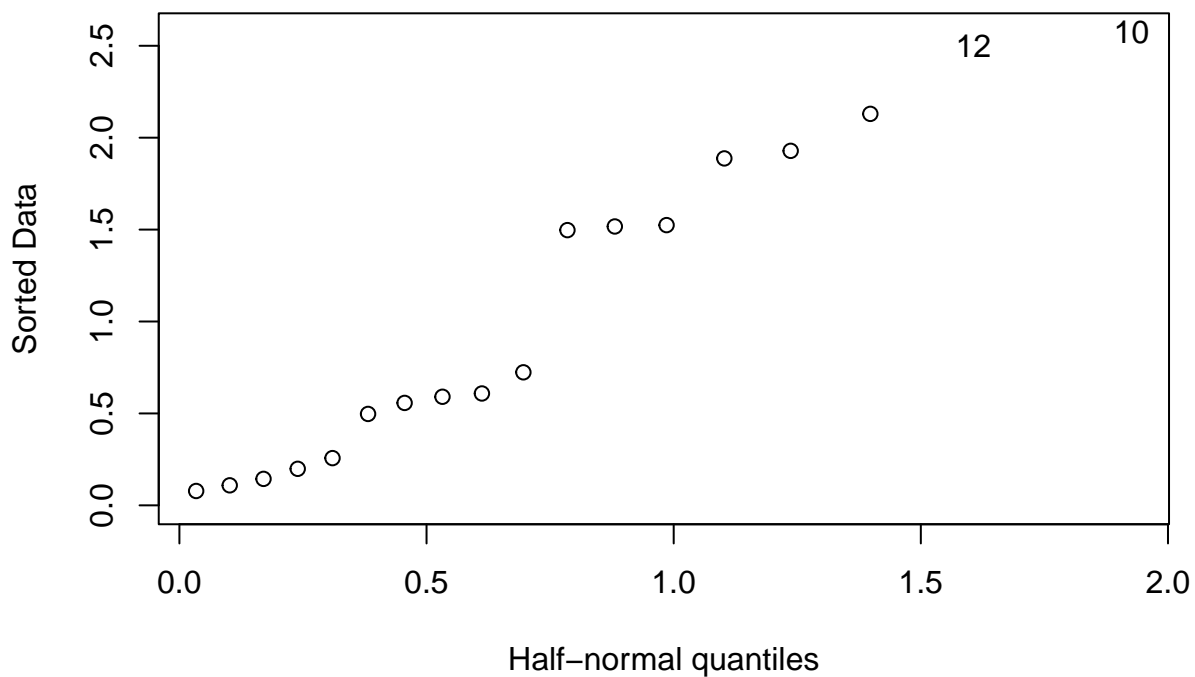
```
model$deviance > qchisq(p = 0.95,df = 12)
```

```
[1] TRUE
```

The deviance is too larger, then we should check whether this data has outliers or may be overdispersed:

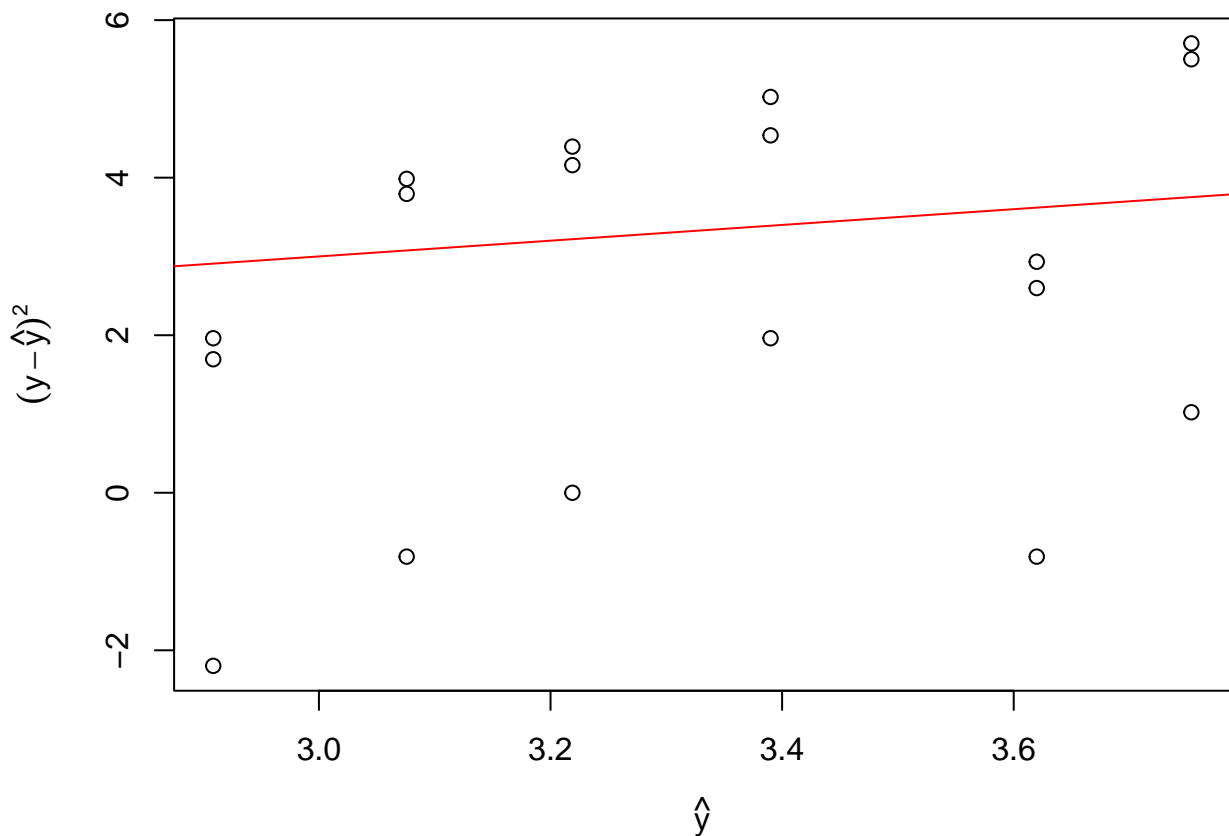
- For detecting outlier (using half-normal plot):

```
library(faraway)
halfnorm( residuals(model))
```



- For detecting overdispersion (comparing mean and variance of fitted model):

```
pois.fit.mean = log( fitted(model) )
pois.fit.var = log( (data$colonies - exp(pois.fit.mean))^2 )
par(mar=c(4.5,4.5,0.5,0.5))
plot( x= pois.fit.mean, y= pois.fit.var,
      xlab = expression( hat(y)),
      ylab = expression( (y- hat(y))^2 )
      abline(0,1, cex=1.5, col="red")
```



Obviously, there are some outliers and there is a clear difference between mean and variance. Therefore, we refit this model after adding dispersion parameter:

```
sigma.hat = sum( residuals( model, type = "pearson")^2) / model$df.residual
summary(model, dispersion = sigma.hat)
```

Call:

```
glm(formula = colonies ~ ., family = "poisson", data = data)
```

Coefficients:

	Estimate	Std. Error	z value	Pr(> z)
(Intercept)	3.32778	0.07601	43.779	< 2e-16 ***
dose.L	0.50669	0.19218	2.637	0.00838 **
dose.Q	-0.22791	0.18465	-1.234	0.21710
dose.C	-0.41331	0.18972	-2.178	0.02937 *
dose^4	0.15583	0.18772	0.830	0.40649
dose^5	0.13253	0.17628	0.752	0.45217

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

(Dispersion parameter for poisson family taken to be 2.777999)

Null deviance: 78.358 on 17 degrees of freedom
 Residual deviance: 33.496 on 12 degrees of freedom
 AIC: 138.03

Number of Fisher Scoring iterations: 4

Finally, we can find that the $s.e.(\hat{\beta})$'s are corrected, and the significant effects are different compared with the

model without adding dispersion parameter.

Problem 3

Question 3.

The `data` was collected from subjects in Denmark. Build a Poisson model to describe how marital status (i.e., single, married, or divorced) changes with age, taking care to give an understandable interpretation of the important parameter estimates. Under your model, predict the probability that a Dane, aged 55, is divorced.

Sol.

```
data = read.table("maritaldane.txt", header = T)
data$Age = as.ordered(data$Age)
head(data)
```

	Age	Single	Married	Divorced	Total
1	17-21	17	1	0	18
2	21-25	16	8	0	24
3	25-30	8	17	1	26
4	30-40	6	22	4	32
5	40-50	5	21	6	32
6	50-60	3	17	8	28

For Poisson model analysis, convert this data into long format:

```
library(tidyverse)
long.data = data %>%
  gather(status, count, -Total, -Age)
long.data$status = factor(long.data$status, levels = c("Single", "Married", "Divorced"))
head(long.data)
```

	Age	Total	status	count
1	17-21	18	Single	17
2	21-25	24	Single	16
3	25-30	26	Single	8
4	30-40	32	Single	6
5	40-50	32	Single	5
6	50-60	28	Single	3

It is observed that this data is collected under the different value of size variable (Total), so use Poisson rate model to analyse it:

```
model = glm(count~offset(log(Total))+(. - Total)^2, data = long.data, family = "poisson")
summary(model)
```

Call:

```
glm(formula = count ~ offset(log(Total)) + (. - Total)^2, family = "poisson",
    data = long.data)
```

Coefficients:

	Estimate	Std. Error	z value	Pr(> z)
(Intercept)	-1.460e+00	1.955e-01	-7.469	8.09e-14 ***
Age.L	-2.060e+00	6.475e-01	-3.181	0.001467 **

```

Age.Q          7.424e-01  6.255e-01  1.187 0.235298
Age.C          8.130e-02  5.894e-01  0.138 0.890287
Age^4         -2.020e-01  5.417e-01 -0.373 0.709281
Age^5          6.103e-02  5.084e-01  0.120 0.904437
Age^6         -1.778e-01  4.823e-01 -0.369 0.712415
Age^7         -1.057e-01  4.451e-01 -0.238 0.812215
statusMarried  5.227e-01  2.574e-01  2.031 0.042304 *
statusDivorced -6.353e+00  1.231e+04 -0.001 0.999588
Age.L:statusMarried  3.163e+00  9.232e-01  3.426 0.000612 ***
Age.Q:statusMarried -2.514e+00  8.954e-01 -2.808 0.004991 **
Age.C:statusMarried  6.064e-01  8.022e-01  0.756 0.449716
Age^4:statusMarried -1.615e-01  6.943e-01 -0.233 0.816040
Age^5:statusMarried -3.409e-02  6.116e-01 -0.056 0.955546
Age^6:statusMarried  2.168e-01  5.534e-01  0.392 0.695190
Age^7:statusMarried  8.911e-02  5.004e-01  0.178 0.858671
Age.L:statusDivorced  2.622e+01  4.623e+04  0.001 0.999547
Age.Q:statusDivorced -1.483e+01  3.800e+04  0.000 0.999689
Age.C:statusDivorced  2.164e+00  3.688e+04  0.000 0.999953
Age^4:statusDivorced  6.217e+00  4.144e+04  0.000 0.999880
Age^5:statusDivorced -7.916e+00  3.583e+04  0.000 0.999824
Age^6:statusDivorced  5.637e+00  2.186e+04  0.000 0.999794
Age^7:statusDivorced -2.028e+00  8.407e+03  0.000 0.999808
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

```

(Dispersion parameter for poisson family taken to be 1)

```

Null deviance: 1.2038e+02 on 23 degrees of freedom
Residual deviance: 3.0331e-10 on 0 degrees of freedom
AIC: 128.04

```

Number of Fisher Scoring iterations: 21

Since most of the effects were not significant, model selection was performed:

```

while (prod(coef(summary(model))[, "Pr(>|z|)"] <= 0.05) != 1) {
  model_mat = model.matrix(model)[, -1]
  minaic = 1000
  for (i in 1:dim(model_mat)[2]) {
    t_dat = data.frame(count = long.data$count, model_mat[, -i],
                      Total = long.data$Total)
    fit = glm(count ~ offset(log(Total)) + (.-Total), data = t_dat, family = "poisson")
    if (fit$AIC < minaic) {
      minaic = fit$AIC
      selectX = i
    }
  }
  t_dat = data.frame(count = long.data$count, model_mat[, -selectX], Total = data$Total)
  model = glm(count ~ offset(log(Total)) + (.-Total), data = t_dat, family = "poisson")
}
summary(model)

```

```
Call:
glm(formula = count ~ offset(log(Total)) + (. - Total), family = "poisson",
    data = t_dat)
```

Coefficients:

	Estimate	Std. Error	z value	Pr(> z)
(Intercept)	-1.8441	0.1486	-12.410	< 2e-16 ***
Age.L	-3.2406	0.4711	-6.879	6.05e-12 ***
statusMarried	0.9727	0.2025	4.804	1.56e-06 ***
Age.L.statusMarried	3.7454	0.6327	5.920	3.22e-09 ***
Age.Q.statusMarried	-1.3857	0.4275	-3.242	0.00119 **
Age.L.statusDivorced	5.9468	0.8390	7.088	1.36e-12 ***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

(Dispersion parameter for poisson family taken to be 1)

Null deviance: 120.384 on 23 degrees of freedom
Residual deviance: 12.371 on 18 degrees of freedom
AIC: 104.41

Number of Fisher Scoring iterations: 5

As seen from the above summary, the interaction effects between Age and marital status are significant, which means that the number of people in different marital statuses will change with different age groups. Finally, predict the probability of (Age=55 \in (50,60), status = Divorced):

```
which1 = which(long.data$Age == "50-60" & long.data$status == "Divorced")
which2 = which(long.data$Age == "50-60")
cat("P(Divoced, age=55) = ", fitted(model)[which1]/ (sum(fitted(model)[which2])))
```

P(Divoced, age=55) = 0.2904925