

EE6106 (Lecture 5) (Date: 30 Jan 2024)

Convergence of random sequences

$\{X_n\}_{n \geq 1} \sim$ seq of random variables

• $X_n \rightarrow X$ almost surely (w.p. 1)

($X_n \xrightarrow{a.s.} X$) if

$$P(\lim_{n \rightarrow \infty} X_n(\omega) = X(\omega)) = 1$$

• $X_n \rightarrow X$ in probability

($X_n \xrightarrow{P} X$) if

$$\lim_{n \rightarrow \infty} P(|X_n - X| > \varepsilon) = 0 \quad \forall \varepsilon > 0$$



this is an "event"

since it is the difference of
two random variables

• $X_n \rightarrow X$ in distribution

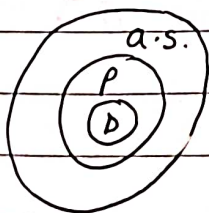
(weak convergence)

($X_n \xrightarrow{D} X$) if

$$\lim_{n \rightarrow \infty} P(X_n \leq x) = P(X \leq x)$$

at all points of continuousness x of $P(X \leq x)$

Note: $X_n \xrightarrow{a.s.} X \Rightarrow X_n \xrightarrow{P} X \Rightarrow X_n \xrightarrow{D} X$



$$A_n(\varepsilon) = \{ |X_n - X| \geq \varepsilon \}$$

$$X_n \xrightarrow{P} X \iff \lim_{n \rightarrow \infty} P(A_n(\varepsilon)) = 0 \quad \forall \varepsilon > 0$$

$$X_n \xrightarrow{a.s.} X \iff \lim_{n \rightarrow \infty} P\left(\bigcup_{m=n}^{\infty} A_m(\varepsilon)\right) = 0$$

The probability of ever
having an ε -deviation in the future
goes to zero.

$\forall \varepsilon > 0$

Limit theorems

$$S_0 = 0$$
$$S_n = \sum_{i=1}^n X_i \quad (n \geq 1)$$

$\left\{ S_n \right\}$ is thus the random sum

Strong Law of Large Numbers (SLLN)

If $\{X_i\}$ are iid with finite mean μ ,

$$\frac{S_n}{n} \xrightarrow{\text{a.s.}} \mu$$

$$\text{i.e. } \mathbb{P}\left(\lim_{n \rightarrow \infty} \frac{S_n(\omega)}{n} = \mu(\omega)\right) = 1$$

Weak Law of Large Numbers (WLLN)

$$\frac{S_n}{n} \xrightarrow{P} \mu$$

$$\text{i.e. } \lim_{n \rightarrow \infty} \mathbb{P}\left(\left|\frac{S_n}{n} - \mu\right| > \varepsilon\right) = 0$$

SLLN: running average will become very close to the expected average
 S_n grows linearly with n (generally), with a slope equal to μ

$$S_n \approx n\mu + O(n)$$



we do not know
the "nature" of sublinearity?



we do, under certain assumptions $\Rightarrow \sqrt{n}$

Central Limit Theorem (CLT)

If X_i are i.i.d. with finite mean μ , finite variance σ^2 ,

$$\frac{S_n - n\mu}{\sqrt{n}\sigma} \xrightarrow{P} Z$$

↓
standard Gaussian

$$\sim N(0,1)$$

Moral : $S_n \approx n\mu + \sqrt{n}Z + o(\sqrt{n})$



$N(0, \sigma^2) \rightarrow$ here the variance is σ^2 instead of 1

Thus, the typical deviations between S_n & $n\mu$ (running sum) & $n\mu$ are of the order $O(\sqrt{n})$.

Suppose we ask,

what is $\lim_{n \rightarrow \infty} P(0 \leq S_n - n\mu \leq \sqrt{n})$ or otherwise

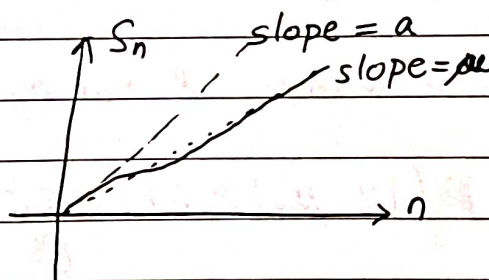
$\lim_{n \rightarrow \infty} P(a\sqrt{n} \leq S_n - n\mu \leq b\sqrt{n})$?

A: This is simply $P(a \leq Z \leq b)$, where Z is $N(0, \sigma^2)$

*The CLT thus says that "typical" deviations of S_n from $n\mu$ are $O(\sqrt{n})$.

Q: $P(S_n > na)$

$\rightarrow a > \mu$



Order of "n" deviation
What is the probability
(large deviations)

Claim 1 : Goes to 0 as $n \rightarrow \infty$

Trivial consequence of the WLLN

$P(S_n > na) = P(\frac{S_n}{n} > a) = P(\frac{S_n - n\mu}{n} > \underbrace{a - \mu}_0) = 0$ (by WLLN)

CONCENTRATION INEQUALITIES

Markov inequality

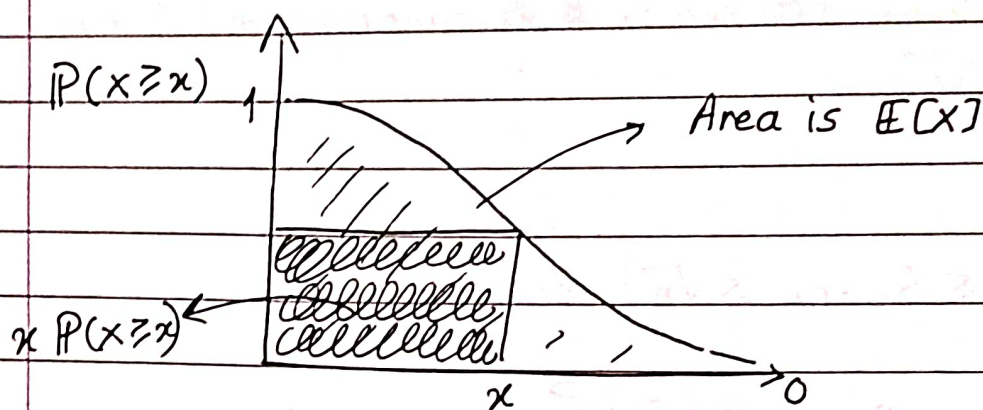
If X is a non-negative RV, $x > 0$

$P(X \geq x) \leq \frac{E[X]}{x}$

Inequality becomes useful only when $x \gg E[X]$ by a large margin

For a non-negative RV,

$$E[X] = \int_0^{\infty} P(X \geq t) dt$$



Then $x P(X \geq x) \leq E[X]$

Typically, we apply a monotone transformation to X & then apply Markov's inequality to both sides.

Corollary $Y \sim \text{mean } \mu_Y \text{ \& var } \sigma_Y^2, t > 0$

$$P(|Y - \mu_Y| \geq t) \leq \frac{\sigma_Y^2}{t^2} \quad \dots \text{CHEBYSHEV INEQUALITY}$$

Proof

$$P(|Y - \mu_Y| \geq t) = P(|Y - \mu_Y|^2 \geq t^2) \underset{\text{M.I.}}{\leq} \frac{E[|Y - \mu_Y|^2]}{t^2}$$

Markov: true when mean is finite but get $1/t$ bound on tail

Chebyshev: variance must also be finite but get $1/t^2$ bound on tail

$$P(Y > y) = P(e^{sY} > e^{sy}) \leq E[e^{sY}] e^{-sy}$$

↓
 for $s > 0$
 to ensure monotonic
 increase of e^x

↓
 ... CHERNOFF INEQUALITY
 Moment
 Generating
 Function

For a r.v. Y ,

$$MGF_Y(s) = M_Y(s) = E[e^{sY}]$$

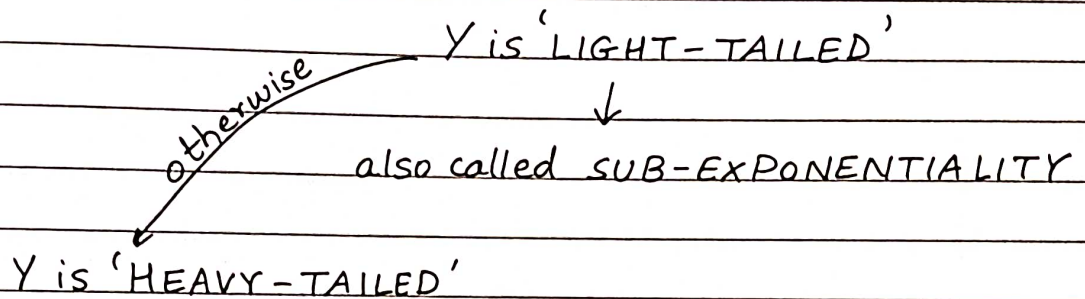
Chernoff Bound:

$$P(Y > y) \leq \inf_{s > 0} [M_Y(s)e^{-sy}]$$

it might also be possible to find the explicit value of s as well.

The Chernoff Bound can also give an exponential bound

For Chernoff bound to hold, we need $M_Y(s) < \infty$ for some $s > 0$



Q: Can we use the Chernoff bound to talk about $P(S_n > na)$

Note: Using Markov inequality yields $P(S_n > na) \leq \frac{\mu}{a}$

Assume $\{X_n\}_{n \geq 1} \sim \text{iid}$, mean μ

$$P(S_n > na) \stackrel{(a > \mu)}{\leq} \inf_{(s > 0)} M_{S_n}(s) e^{-sna} \quad (\text{Chernoff bound})$$

$$M_{S_n}(s) = E[e^{s(X_1 + \dots + X_n)}] = (M_X(s))^n \quad \dots (\text{iid})$$

$$\therefore P(S_n > na) \leq \inf_{s > 0} [M_X(s)]^n e^{-sna}$$

Define $\Lambda_X(s) = \log M_X(s)$
 \downarrow
 log MGF

$$P(S_n > na) \leq e^{n \Lambda_X(s) - sna} = e^{-n[sa - \Lambda_X(s)]}$$

$$\Rightarrow P(S_n > na) \leq e^{-n \sup_{s > 0} (sa - \Lambda_X(s))}$$

as long as $\sup_{s > 0} (sa - \Lambda_X(s))$ is positive
 we can get an exponentially decaying term

* For light tailed x , ^{with $a > \mu$} it turns out that

$$\sup_{s > 0} [sa - \Lambda_x(s)] > 0$$

It also turns out that this is the tightest bound possible.

Thus the upper bound is asymptotically tight (Cramer's Theorem)

Depending on the distribution of x , we can get the "customized" bounds. However, this is a downside when it comes to learning algorithms because now the ^{distribution, and so,} bounds are not available.