

1 Gaussian distribution is normalized

(1) Σ symmetric $\rightarrow \Sigma^{-1}$ symmetric

Proof:

$$\begin{aligned}\Sigma \Sigma^{-1} &= I \\ (\Sigma \Sigma^{-1})^T &= I \\ (\Sigma^{-1})^T \Sigma^T &= I \\ \Sigma = \Sigma^T \rightarrow \Sigma^{-1})^T \Sigma &= I \\ \Sigma^{-1} \Sigma &= I \rightarrow \Sigma^{-1} = (\Sigma^{-1})^T\end{aligned}$$

Gaussian distribution:

$$N(x|\mu, \Sigma) = \frac{1}{(2\pi)^{D/2} |\Sigma|^{1/2}} (x - \mu)^T \Sigma^{-1} (x - \mu)$$

$$\begin{aligned}\Delta^2 &= (x - \mu)^T \Sigma^{-1} (x - \mu) \\ &= x^T \Sigma^{-1} x - \mu^T \Sigma^{-1} x - x^T \Sigma^{-1} \mu + \mu^T \Sigma^{-1} \mu\end{aligned}$$

$$\begin{aligned}\mu^T \Sigma^{-1} x &= a \\ (\mu^T \Sigma^{-1} x)^T &= a^T \\ x^T \Sigma^{-1} \mu &= a \\ \mu^T \Sigma^{-1} x &= x^T \Sigma^{-1} \mu\end{aligned}$$

$$\rightarrow \Delta^2 = x^T \Sigma^{-1} x - 2x^T \Sigma^{-1} \mu + \text{const}$$

(2) Σ symmetric \rightarrow eigenvectors orthogonal

proof :

$$\lambda_i u_i = \Sigma u_i \rightarrow \lambda_i u_j^T u_i = u_j^T \Sigma u_i = (u_j \Sigma)^T u_i = (u_j \lambda_j)^T u_i = \lambda_j u_j^T u_i \rightarrow u_i \text{ and } u_j \text{ orthogonal}$$

$$\begin{aligned}\Sigma &= \sum_{i=1}^D \lambda_i u_i u_i^T = U \Lambda U^T \\ \Sigma^{-1} &= (U \Lambda U^T)^{-1} \\ \Sigma^{-1} &= U \Lambda^{-1} U^T \\ \Sigma^{-1} &= \sum_{i=1}^D \frac{1}{\lambda_i} u_i u_i^T\end{aligned}$$

$$\begin{aligned}
\Delta^2 &= (x - \mu)^T \Sigma^{-1} (x - \mu) \\
&= (x - \mu)^T \sum_{i=1}^D \frac{1}{\lambda_i} u_i u_i^T (x - \mu) \\
&= \sum_{i=1}^D \frac{1}{\lambda_i} (x - \mu)^T u_i u_i^T (x - \mu)
\end{aligned}$$

Let $y_i = u_i^T (x - \mu) \rightarrow \Delta^2 = \sum_{i=1}^D \frac{y_i^2}{\lambda_i}$

We have: $|\Sigma|^{1/2} = \prod_{i=1}^D \lambda_i^{1/2} \rightarrow p(y) = \prod_{i=1}^D \frac{1}{(2\pi\lambda_i)^{1/2}} e^{-\frac{y_i^2}{2\lambda_i}}$

$$\begin{aligned}
\int_{-\infty}^{\infty} p(y) dy &= \int_{-\infty}^{\infty} \prod_{i=1}^D \frac{1}{(2\pi\lambda_i)^{1/2}} e^{-\frac{y_i^2}{2\lambda_i}} dy \\
&= \prod_{i=1}^D \frac{1}{(2\pi\lambda_i)^{1/2}} \int_{-\infty}^{\infty} e^{-\frac{y_i^2}{2\lambda_i}} dy \\
&= 1
\end{aligned}$$

\rightarrow Gaussian distribution is normalized

2 Conditional Gaussian distribution

$$x = \begin{bmatrix} x_a \\ x_b \end{bmatrix} \quad \mu = \begin{bmatrix} \mu_a \\ \mu_b \end{bmatrix} \quad \Sigma = \begin{bmatrix} \Sigma_{aa} & \Sigma_{ab} \\ \Sigma_{ba} & \Sigma_{bb} \end{bmatrix}$$

Let $\Lambda \equiv \Sigma^{-1} \rightarrow \Lambda = \begin{bmatrix} \Lambda_{aa} & \Lambda_{ab} \\ \Lambda_{ba} & \Lambda_{bb} \end{bmatrix}$

We have

$$\begin{aligned}
-\frac{1}{2}(x - \mu)^T \Sigma^{-1} (x - \mu) &= -\frac{1}{2}(x - \mu)^T \Lambda (x - \mu) \\
&= -\frac{1}{2}(x_a - \mu_a)^T \Lambda_{aa} (x_a - \mu_a) - \frac{1}{2}(x_a - \mu_a)^T \Lambda_{ab} (x_b - \mu_b) \\
&\quad - \frac{1}{2}(x_b - \mu_b)^T \Lambda_{ba} (x_a - \mu_a) - \frac{1}{2}(x_b - \mu_b)^T \Lambda_{bb} (x_b - \mu_b)
\end{aligned}$$

We calculate $p(x_a | x_b) \rightarrow x_b$ is regarded as a constant

$$-\frac{1}{2}(x - \mu)^T \Sigma^{-1} (x - \mu) = -\frac{1}{2} x_a^T \Lambda_{aa} x_a + x_a^T (\Lambda_{aa} \mu_a - \Lambda_{ab} (x_b - \mu_b)) + const$$

It is quadratic form of x_a hence conditional distribution $p(x_a|x_b)$ will be Gaussian, because this distribution is characterized by its mean and its variance. Compare with Gaussian distribution

$$-\frac{1}{2}\Delta^2 = -\frac{1}{2}x^T\Sigma^{-1}x + x^T\Sigma^{-1}\mu + \text{const}$$

$$\rightarrow \Sigma_{a|b} = \Lambda_{aa}^{-1}$$

$$\rightarrow \Sigma_{a|b}^{-1}\mu_{a|b} = \Lambda_{aa}\mu_a - \Lambda_{ab}(x_b - \mu_b)$$

$$\rightarrow \mu_{a|b} = \Sigma_{a|b}(\Lambda_{aa}\mu_a - \Lambda_{ab}(x_b - \mu_b)) = \mu_a - \Lambda_{aa}^{-1}\Lambda_{ab}(x_b - \mu_b)$$

By using Schur complement, with $M = (A - BD^{-1}C)^{-1}$

$$\begin{bmatrix} A & B \\ C & D \end{bmatrix}^{-1} = \begin{bmatrix} M & -MBD^{-1} \\ -D^{-1}CMD^{-1} & -D^{-1}CMBD^{-1} \end{bmatrix}$$

$$\rightarrow \Lambda_{aa} = (\Sigma_{aa}^{-1} - \Sigma_{aa}^{-1}\Sigma_{ab}\Sigma_{bb}^{-1}\Sigma_{ba})^{-1}$$

$$\Lambda_{ab} = -(\Sigma_{aa} - \Sigma_{ab}\Sigma_{bb}^{-1}\Sigma_{ba})^{-1}\Sigma_{ab}\Sigma_{bb}^{-1}$$

As a result

$$\begin{aligned} \mu_{a|b} &= \mu_a + \Sigma_{ab}\Sigma_{bb}^{-1}(x_b - \mu_b) \\ \Sigma_{a|b} &= \Sigma_{aa} - \Sigma_{ab}\Sigma_{bb}^{-1}\Sigma_{ba} \\ p(a|b) &= N(x_a|\mu_{a|b}, \Sigma_{a|b}) \end{aligned}$$

3 Marginal Gaussian distribution

$$p(x_a) = \int p(x_a, x_b)dx_b$$

We need to integrate out x_b by looking the quadratic form related to x_b with $m = \Lambda_{bb}\mu_b - \Lambda_{ba}(x_a - \mu_a)$

$$\begin{aligned} -\frac{1}{2}(x - \mu)^T\Lambda(x - \mu) &= -\frac{1}{2}x_b^T\Lambda_{bb}x_b + x_b^Tm - \frac{1}{2}x_a^T\Lambda_{aa}x_a + x_a^T(\Lambda_{aa}\mu_a + \Lambda_{ab}\mu_b) + \text{const} \\ -\frac{1}{2}x_b^T\Lambda_{bb}x_b + x_b^Tm &= -\frac{1}{2}(x_b - \Lambda_{bb}^{-1}m)^T\Lambda_{bb}(x_b - \Lambda_{bb}^{-1}m) + \frac{1}{2}m^T\Lambda_{bb}^{-1}m \\ \rightarrow -\frac{1}{2}(x - \mu)^T\Lambda(x - \mu) &= -\frac{1}{2}(x_b - \Lambda_{bb}^{-1}m)^T\Lambda_{bb}(x_b - \Lambda_{bb}^{-1}m) \\ &\quad + \frac{1}{2}m^T\Lambda_{bb}^{-1}m - \frac{1}{2}x_a^T\Lambda_{aa}x_a + x_a^T(\Lambda_{aa}\mu_a + \Lambda_{ab}\mu_b) + \text{const} \\ &= -\frac{1}{2}(x_b - \Lambda_{bb}^{-1}m)^T\Lambda_{bb}(x_b - \Lambda_{bb}^{-1}m) \\ &\quad - \frac{1}{2}x_a^T(\Lambda_{aa} - \Lambda_{ab}\Lambda_{bb}^{-1}\Lambda_{ba})x_a + x_a^T(\Lambda_{aa} - \Lambda_{ab}\Lambda_{bb}^{-1}\Lambda_{ba})^{-1}\mu_a + \text{const} \end{aligned}$$

We can integrate over unnormalized Gaussian:

$$\int e^{-\frac{1}{2}(x_b - \Lambda_{bb}^{-1}m)^T \Lambda_{bb}(x_b - \Lambda_{bb}^{-1}m)} dx_b$$

The remaining term:

$$-\frac{1}{2}x_a^T(\Lambda_{aa} - \Lambda_{ab}\Lambda_{bb}^{-1}\Lambda_{ba})x_a + x_a^T(\Lambda_{aa} - \Lambda_{ab}\Lambda_{bb}^{-1}\Lambda_{ba})^{-1}\mu_a + const$$

Compare with $-\frac{1}{2}\Delta^2 = -\frac{1}{2}x^T\Sigma^{-1}x + x^T\Sigma^{-1}\mu + const$

We have:

$$\begin{aligned} E[x_a] &= \mu_a \\ cov[x_a] &= \Sigma_{aa} \\ p(x_a) &= N(x_a|\mu_a, \Sigma_{aa}) \end{aligned}$$