

Time Series

4. Estimation, validation and forecasting

Josep A. Sanchez-Espigares

Department of Statistics and Operations Research
Universitat Politècnica de Catalunya
Barcelona, Spain



UNIVERSITAT POLITÈCNICA
DE CATALUNYA
BARCELONATECH

- Estimation of parameters
- Validation of the model
- Forecasting with ARIMA models

Let $X = \{x_t\}_{t=0}^N$ a time series with conditional gaussian distribution

Likelihood of the model:

$$L(\phi, \theta, \sigma_z^2; X) = f_{(x_1, \dots, x_T)}(x_1, \dots, x_T) = f_1(x_1) \sum_{i=2}^T f(x_i | x_{i-1}, \dots, x_1)$$

For example, for an AR(1) model:

$$X_i = \phi X_{i-1} + Z_t \quad Z_t \sim N(0, \sigma_Z^2)$$

$$E(X_i | X_{i-1}, \dots, X_1) = \phi X_{i-1}$$

$$V(X_i | X_{i-1}, \dots, X_1) = \sigma_Z^2$$

$$L(\phi, \sigma_z^2; X) = f_1(x_1) \sum_{i=2}^N f\left(\frac{x_i - \phi x_{i-1}}{\sigma_Z}\right)$$

Conditional density: it depends on the distribution considered for the first observation (f_1)

Maximum likelihood estimation

$$\hat{\Lambda}_{ML} = (\hat{\phi}, \hat{\theta}, \hat{\sigma}_z^2) = \operatorname{argmax} (L(\phi, \theta, \sigma_z^2; X))$$

Maximum likelihood estimators are **consistent** and **asymptotically efficient and gaussian**

$$\hat{\Lambda}_{ML} \approx N(\Lambda, I_{\Lambda}^{-1})$$

where $I_{\Lambda} = E \left(-\frac{\partial^2 \log L(\Theta; X)}{\partial \Theta^2} \right)$. This is the Hessian of the objective function ($H(L)$)

In particular, $se(\hat{\Lambda}_{ML}) = \sqrt{\operatorname{diag}[H(L)^{-1}]}$

Some details of the implementation in R:

- 1 Expression of the model in State-Space form
- 2 Start at some initial values for the parameters
- 3 Calculation of the logLikelihood function for these values
- 4 Non-linear optimization method (`optim`) to find the Maximum Likelihood estimates
- 5 The `optim` method gives the values of the Hessian of the function. Its diagonal is used to calculate the standard deviation of the estimates (standard errors)

Significance test for the parameters of the model:

Exemple for a coefficient of the AR-part:

$$\begin{cases} H_0 : \phi_i = 0 \\ H_1 : \phi_i \neq 0 \end{cases}$$

Test statistic:

$$\hat{\phi}_i \approx N(\phi_i, \sigma_{\phi_i}) \Rightarrow \hat{t} = \frac{\hat{\phi}_i}{se(\hat{\phi}_i)} \approx t - Student_{T-k}$$

where $se(\hat{\phi}_i)$ is the estimate of σ_{ϕ_i} based on the Hessian of the function, and k is the number of total parameters of the model.

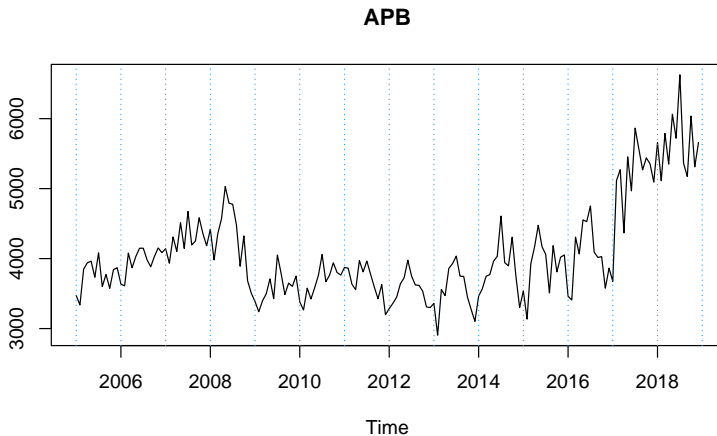
For practical purposes: if $|\hat{t}| > 2$, then the parameter is significant.

APB: Port traffic in the port of Barcelona (millions of tones)

Source: Ministry of Public Works of Spain (<http://www.fomento.gob.es/BE/?nivel=2&orden=04000000>)

##		Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
##	2005	3473	3338	3850	3940	3963	3733	4082	3600	3773	3575	3842	3870
##	2006	3636	3611	4079	3869	4030	4148	4149	3984	3884	4032	4152	4086
##	2007	4141	3934	4309	4102	4512	4145	4674	4195	4250	4584	4359	4182
##	2008	4418	3981	4357	4570	5030	4792	4776	4482	3891	4324	3682	3504
##	2009	3385	3240	3397	3497	3710	3428	4050	3785	3486	3649	3603	3750
##	2010	3382	3265	3578	3423	3583	3760	4058	3669	3763	3940	3802	3764
##	2011	3874	3862	3633	3556	3971	3810	3965	3776	3593	3425	3628	3197
##	2012	3286	3362	3449	3639	3726	3977	3750	3623	3618	3532	3309	3300
##	2013	3359	2906	3559	3471	3859	3929	4036	3749	3744	3449	3271	3103
##	2014	3464	3569	3747	3772	3962	4033	4607	3942	3896	4305	3749	3300
##	2015	3536	3135	3927	4166	4477	4171	4061	3510	4185	3810	4021	4052
##	2016	3460	3410	4306	4067	4553	4528	4752	4092	4016	4028	3577	3863
##	2017	3674	5116	5270	4368	5454	4969	5865	5564	5268	5440	5353	5094
##	2018	5659	5112	5789	5350	6063	5720	6627	5360	5173	6037	5311	5659

APB: Port traffic in the port of Barcelona (millions of tones)



Data: APB

Proposed transformations: $W_t = (1 - B)(1 - B^{12})\log(X_t)$

Proposed model: $ARIMA(0, 1, 4)(0, 1, 1)_{12}$

```
(model=arima(serie,order=c(0,1,4),seasonal=list(order=c(0,1,1),period=12)))
```

```
##
## Call:
## arima(x = serie, order = c(0, 1, 4), seasonal = list(order = c(0, 1, 1), period = 12))
##
## Coefficients:
##          ma1          ma2          ma3          ma4          sma1
##      -0.4570   -0.0081    0.0405   -0.1574   -0.9227
## s.e.    0.0823    0.0946    0.0887    0.0861    0.1530
##
## sigma^2 estimated as 0.004536:  log likelihood = 187.27,  aic = -362.54
cat("\nT-ratios:",round(model$coef/sqrt(diag(model$var.coef)),2))
```

```
##
## T-ratios: -5.55 -0.09 0.46 -1.83 -6.03
cat("\nSignificant?:",abs(model$coef/sqrt(diag(model$var.coef)))>2)
```

```
##
## Significant?: TRUE FALSE FALSE FALSE TRUE
```

For all $ARIMA(p, d, q)(P, D, Q)_s$ model, the general expression is:

$$\phi_p(B)\Phi_P(B^S) \left((1-B)^d(1-B^S)^D X_t - \mu \right) = \theta_q(B)\Theta_Q(B^S)Z_t$$

The random part of the model is $Z_t \sim N(0, \sigma_Z^2)$

- **Residual Analysis** for Z_t . Premises:
 - Homogeneity of variance (σ_Z^2 constant)
 - Normality ($Z_t \sim \text{Normal}$)
 - Independence ($\rho(k) = 0 \quad \forall k$)

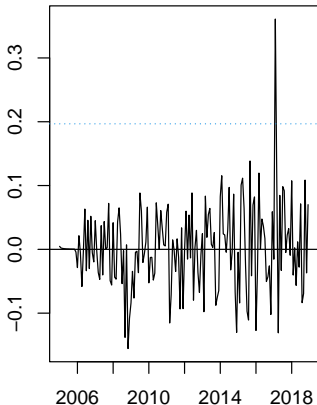
Homogeneity of variance. Tools

- Residuals plot
- Square root of absolute values of the residuals with smooth fit
- ACF/PACF of square of residuals

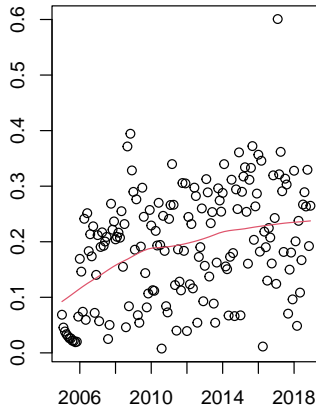
Validation

```
resid=model$residuals
par(mfrow=c(1,2),mar=c(3,3,3,3))
#Residuals plot
plot(resid,main="Residuals")
abline(h=0)
abline(h=c(-3*sd(resid),3*sd(resid)),lty=3,col=4)
#Square Root of absolute values of residuals (Homocedasticity)
scatter.smooth(sqrt(abs(resid)),main="Square Root of Absolute residuals",
               lpars=list(col=2))
```

Residuals



Square Root of Absolute residuals



- Outlier observations: Outlier detection and treatment
- Volatility: Models for the variance (GARCH)

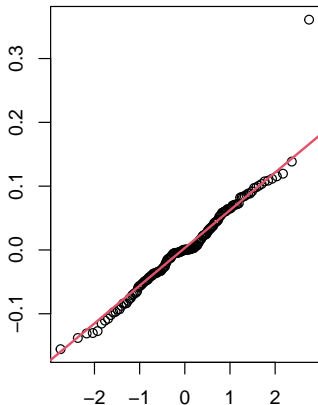
Normality. Tools

- Quantile-Quantile plot
- Histogram with theoretical density overlapped
- Shapiro-Wilks test

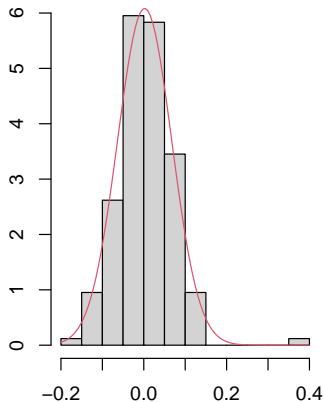
Validation

```
par(mfrow=c(1,2),mar=c(3,3,3,3))  
#Normal plot of residuals  
qqnorm(resid)  
qqline(resid,col=2,lwd=2)  
##Histogram of residuals with normal curve  
hist(resid,breaks=10,freq=F)  
curve(dnorm(x,mean=mean(resid),sd=sd(resid)),col=2,add=T)
```

Normal Q-Q Plot



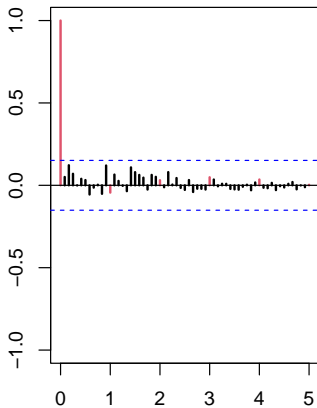
Histogram of resid



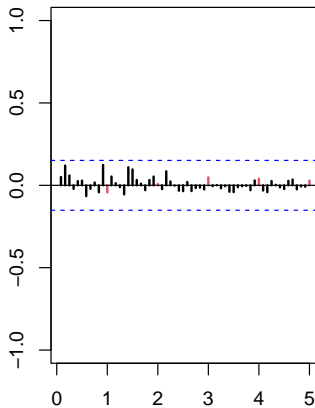
Validation

```
par(mfrow=c(1,2),mar=c(3,3,3,3))
s=12
#ACF & PACF of residuals
par(mfrow=c(1,2))
acf(resid^2,ylim=c(-1,1),lag.max=60,col=c(2,rep(1,s-1)),lwd=2)
pacf(resid^2,ylim=c(-1,1),lag.max=60,col=c(rep(1,s-1),2),lwd=2)
```

Series resid^2



Series resid^2



```
par(mfrow=c(1,1))
```


- Outlier observations: Outlier detection and treatment
- Asymmetry or bi-modality: Transformation, outliers or change residuals distribution
- Heavy tails (excess of kurtosi): Volatility models or t-distributions for the residuals

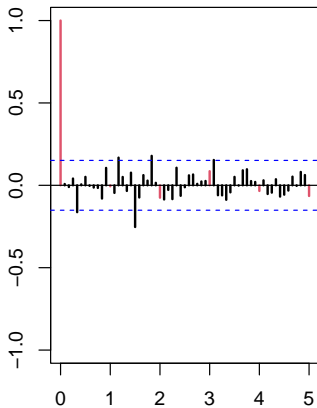
Independence. Tools

- ACF/PACF of residuals
- LJung-Box test
- Durbin-Watson test

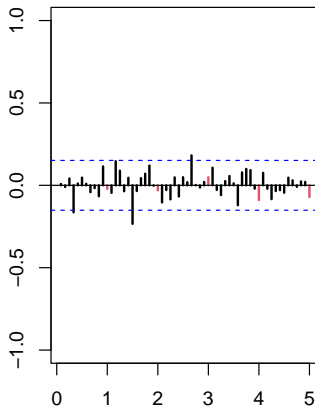
Validation

```
par(mfrow=c(1,2),mar=c(3,3,3))
s=12
#ACF & PACF of residuals
par(mfrow=c(1,2))
acf(resid,ylim=c(-1,1),lag.max=60,col=c(2,rep(1,s-1)),lwd=2)
pacf(resid,ylim=c(-1,1),lag.max=60,col=c(rep(1,s-1),2),lwd=2)
```

Series resid



Series resid



```
par(mfrow=c(1,1))
```

Checking the significance of individual autocorrelation might ignore that the configuration of all (or a subset) of the lags may be jointly significant.

Ljung-Box Q-statistic: For a lag k test the joint hypothesis that the first k autocorrelations of the residuals are jointly zero:

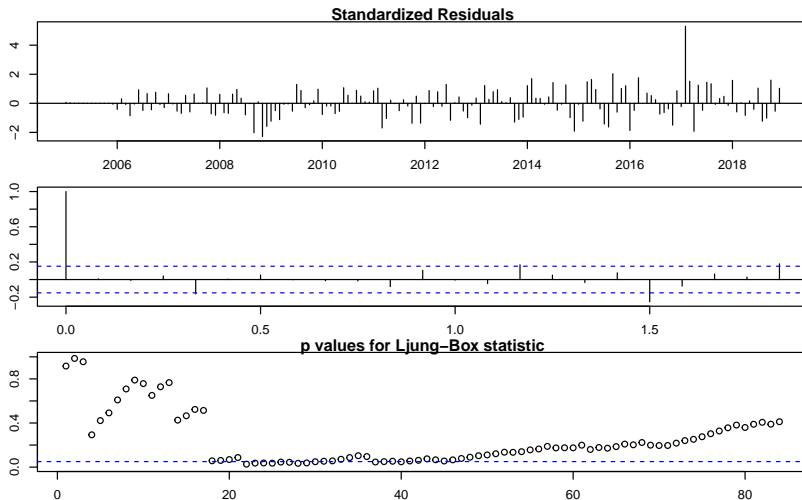
$$H_0 : \rho_Z(1) = \cdots = \rho_Z(k) = 0$$

Test Statistic: $Q = T \sum_{i=1}^K \hat{\rho}_Z^2(i)$

Asymptotically, $Q \rightarrow \chi_k^2$

Validation

```
par(mfrow=c(1,1),mar=c(3,3,3,3))  
#Ljung-Box p-values  
par(mar=c(2,2,1,1))  
tsdiag(model,gof.lag=7*12)
```



Independence. Problems

- Significant lags: Re-identify or add parameters to the model

- Information Criterion

- Balance between goodness-of-fit and simplicity of the model
- Component for the goodness-of-fit: $-2 \log L(\phi, \theta, \sigma^2 | X)$
- Component for the simplicity: $K * par$

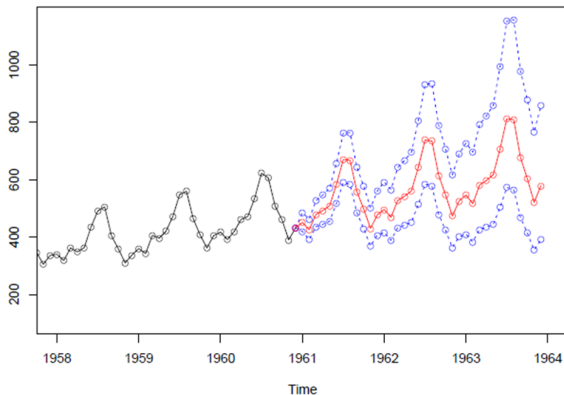
$$AIC = -2 \log L(\phi, \theta, \sigma^2 | X) + 2par$$

$$BIC = -2 \log L(\phi, \theta, \sigma^2 | X) + \log(N)par$$

$$AICc = -2 \log L(\phi, \theta, \sigma^2 | X) + 2N(p + q + 1)/(N - (p + q) - 2)$$

$ARIMA(0,1,1)(0,1,1)_s$; period $s = 12$; $n.ahead = 36$

3-years ahead prediction for AirPassengers



Remember: \A times series $\{X_t; t = 0, 1, \dots\}$ is an **ARMA(p,q)** model if it is **stationary/causal** and **invertible**

$$X_t = \phi_1 X_{t-1} + \dots + \phi_p X_{t-p} + Z_t + \theta_1 Z_{t-1} + \dots + \theta_q Z_{t-q} \quad Z_t \sim WN(0, \sigma_Z^2)$$

with $\phi_p \neq 0$; $\theta_q \neq 0$ and $\sigma_Z^2 > 0$

ARMA(p,q) model can be written in concise form as

$$\Phi_p(B)X_t = \Theta_q(B)Z_t$$

Objective: To find the best forecast for X_{t+h} at time t using the **h-step-ahead minimum mean square error predictor**, denoted by $\tilde{X}_{t+h|t}$

Questions to help the construction process:

- Which is the **forecasting function**?
i.e., Which are the **values for** $\tilde{X}_{t+h|t}$ as a function of **h**?
- Which is the **memory of the model**? How are the forecast expressed as a (linear) **function of the previous** (known) observations?
- Which is the **forecasting confidence interval**?

Three equivalent ARMA(p,q) equations:

- **Previous observations and random noises**

$$X_t = \phi_1 X_{t-1} + \cdots + \phi_p X_{t-p} + Z_t + \theta_1 Z_{t-1} + \cdots + \theta_q Z_{t-q}$$

- **AR process of infinite order**

$$X_t = -\pi_1 X_{t-1} - \pi_2 X_{t-2} - \cdots + Z_t$$

- **MA process of infinite order**

$$X_t = Z_t + \psi_1 Z_{t-1} + \psi_2 Z_{t-2} + \cdots$$

h-step-ahead-forecast

- **Previous observations and random noises**

$$X_{t+h} = \phi_1 X_{t+h-1} + \cdots + \phi_p X_{t+h-p} + Z_{t+h} + \theta_1 Z_{t+h-1} + \cdots + \theta_q Z_{t+h-q}$$

- **AR process of infinite order**

$$X_{t+h} = -\pi_1 X_{t+h-1} - \pi_2 X_{t+h-2} - \cdots + Z_{t+h}$$

- **MA process of infinite order**

$$X_{t+h} = Z_{t+h} + \psi_1 Z_{t+h-1} + \psi_2 Z_{t+h-2} + \cdots$$

Given the model

$$X_t = \phi_1 X_{t-1} + \phi_2 X_{t-2} + Z_t + \theta_1 Z_{t-1} + \theta_2 Z_{t-2}$$

Suppose that **values of past observations are known**, then $Z_t = X_t - \tilde{X}_{t|t-1}$;
 $Z_{t-1} = X_{t-1} - \tilde{X}_{t-1|t-2}$

Substituting

$$\tilde{X}_{t+1|t} = \phi_1 X_t + \phi_2 X_{t-1} + E(Z_{t+1}) + \theta_1 Z_t + \theta_2 Z_{t-1}$$

$$\tilde{X}_{t+2|t} = \phi_1 \tilde{X}_{t+1|t} + \phi_2 X_t + E(Z_{t+2}) + \theta_1 E(Z_{t+1}) + \theta_2 Z_t$$

$$\tilde{X}_{t+3|t} = \phi_1 \tilde{X}_{t+2|t} + \phi_2 \tilde{X}_{t+1|t} + E(Z_{t+3}) + \theta_1 E(Z_{t+2}) + \theta_2 E(Z_{t+1})$$

...

$$\tilde{X}_{t+h-3|t} = \phi_1 \tilde{X}_{t+h-1|t} + \phi_2 \tilde{X}_{t+h-2|t}, \quad h > 2$$

In fact,

Given a general **ARMA(p,q)** model, for $h > q$:

h-step-ahead predictor is determined by the difference equations for the autocorrelations (which only depend on the autocorrelation characteristic polynomial)

Model:

$$X_t = X_{t-1} + Z_t + \theta Z_{t-1}$$

also defined by

$$(1 - B)X_t = (1 - \theta B)Z_t$$

X_t can be expressed as

$$\pi(B) = \frac{1 - B}{1 + \theta B} = 1 - (\theta + 1)B + \theta(\theta + 1)B^2 - \theta^2(\theta + 1)B^3 + \dots$$

Model:

$$\tilde{X}_{t+1|t} = (\theta + 1)X_t - \theta(\theta + 1)X_{t-1} + \theta^2(\theta + 1)X_{t-2} - \dots$$

moving θ

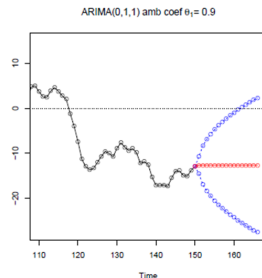
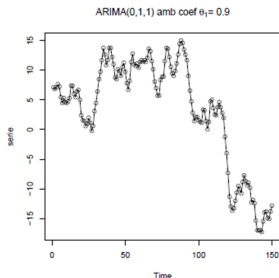
$$\tilde{X}_{t+1|t} = (\theta + 1)X_t - \theta\tilde{X}_{t|t-1}$$

Using the parameter $\lambda = 1 + \theta$. Clearly $\theta = \lambda - 1 = -(1 - \lambda)$

$$\tilde{X}_{t+1|t} = \lambda X_t - (1 - \lambda)\lambda X_{t-1} + (1 - \lambda)^2 \lambda X_{t-2} - \dots = \lambda X_t + (1 - \lambda)\tilde{X}_{t|t-1}$$

Non Stationary ARIMA process Forecast. ARIMA(0,1,1) (3/3)

Hence, the **forecast is the a linear combination** of the last observation and the forecast obtained in the previous step.



Variance of the 1-step-ahead forecasting error

$$\text{Var}(e_t(1)) = E((X_{t+1} - \tilde{X}_{t+1|t})^2) = E(Z_{t+1}^2) = \sigma_Z^2$$

H step forecast for h 2 is

$$\text{Var}(e_t(h)) = \sigma_Z^2(1 + \psi_1^2 + \psi_2^2 + \cdots \psi_{h-1}^2)$$

Model:

$$(1 - B)(1 - B^{12})X_t = (1 + \theta B)(1 + \theta_{12}B^{12})Z_t$$

also defined by

$$(X_t - X_{t-12}) - (X_{t-1} - X_{t-13}) = Z_t + \theta_1 Z_{t-1} + \theta_{12} Z_{t-12} + \theta_1 \theta_{12} Z_{t-13}$$

Hence,

$$X_{t+1} = X_t + (X_{t-11} - X_{t-12} + Z_{t+1} + \theta_1 Z_{t-11} + \theta_1 \theta_{12} Z_{t-12})$$

The forecasting function:

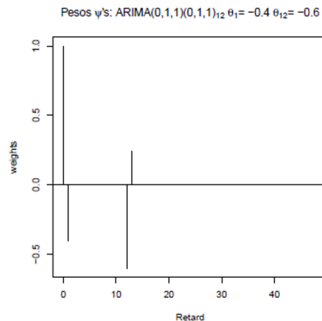
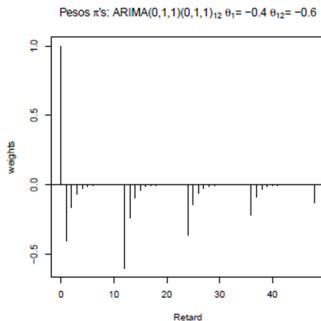
$$\tilde{X}_{t+1|t} = X_t + (X_{t-11} - X_{t-12}) + \theta_1 Z_t + \theta_1 \theta_{12} Z_{t-11} + \theta_1 \theta_{12} Z_{t-12}$$

$$\tilde{X}_{t+2|t} = \tilde{X}_{t+1|t} + (X_{t-10} - X_{t-11}) + \theta_{12} Z_{t-10} + \theta_1 \theta_{12} Z_{t-11}$$

...

$$\tilde{X}_{t+h|t} = \tilde{X}_{t+h-1|t} + (\tilde{X}_{t+h-12|t} - \tilde{X}_{t+h-13|t}), \quad h > 13$$

Seasonal Processes Forecasting. ARIMA(0,1,1)(0,1,1)₁₂² (3/3)



Weights π and ψ for the model ARIMA(0,1,1)(0,1,1)₁₂ for the parameters above

