

MULTIVARIATE PROBABILITY DISTRIBUTIONS (Chapter 5)

We'll first look at the discrete case; the continuous case will be discussed later.

The discrete case

(3 mixed case)

Example 1 A die is rolled. Let X = no. of 6's and Y = no. of even numbers.
Find the joint probability distribution of X and Y .

Number on die	1	2	3	4	5	6
Value of X	0	0	0	0	0	1
Value of Y	0	1	0	1	0	1

↓ ↓ ↓ ↓ ↓

same

$P(X=1, Y=1) = P(6) = 1/6$

$P(X=0, Y=1) = P(2 \text{ or } 4) = 2/6 = 1/3$

$P(X=0, Y=0) = P(1 \text{ or } 3 \text{ or } 5) = 3/6 = 1/2$

We say that X and Y have a joint probability distribution.

(pmf)
The joint pdf of X and Y is

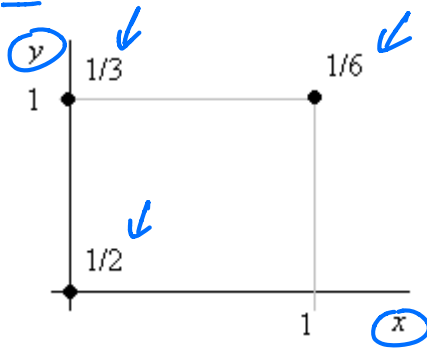
$$p(x, y) = P(X=x, Y=y) = \begin{cases} 1/2, & x=y=0 \\ 1/3, & x=0, y=1 \\ 1/6, & x=y=1 \\ 0, & \text{otherwise} \end{cases}$$

The joint probability distribution of X and Y can also be presented in other ways.

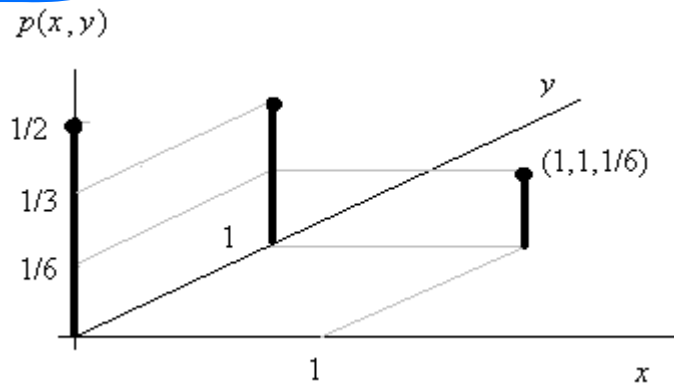
Table of $p(x, y)$:

		y	
		0	1
x	0	1/2	1/3
	1		1/6

Graph (two-dimensional top view):



Three-dimensional graph:



Two properties of discrete joint pdfs:

1. $0 \leq p(x,y) \leq 1$ for all x and y
2. $\sum_{x,y} p(x,y) = 1$. (Or, equivalently, $\sum_x \sum_y p(x,y) = 1$.)

In Example 1: $1/2$, $1/3$ and $1/6$ are all in the interval $[0,1]$; and $1/2 + 1/3 + 1/6 = 1$.

The joint cdf of X and Y is

$$F(x,y) = P(X \leq x, Y \leq y).$$

In Example 1 observe that:

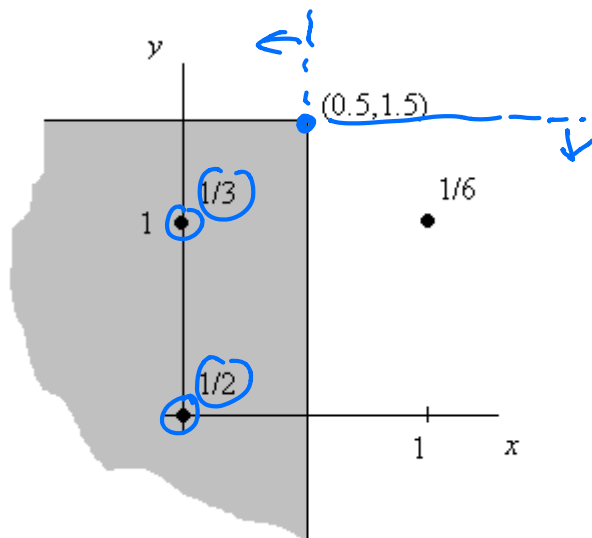
$$F(0,0) = P(X \leq 0, Y \leq 0) = p(0,0) = 1/2$$

$$F(0,0.5) = P(X \leq 0, Y \leq 0.5) = p(0,0) = 1/2$$

$$F(0,1) = P(X \leq 0, Y \leq 1) = p(0,0) + p(0,1) = 1/2 + 1/3 = 5/6$$

$$F(0.5,1.5) = P(X \leq 0.5, Y \leq 1.5) = p(0,0) + p(0,1) = 1/2 + 1/3 = 5/6 \text{ etc.}$$

The following figure illustrates the working for $F(0.5,1.5)$. The region to the left of and below $(0.5,1.5)$ is shaded, and we see that this region contains $(0,0)$ and $(0,1)$. So we sum the joint pdf $f(x,y)$ over those points to get the joint cdf $F(0.5,1.5)$. The region includes its boundary lines. If there were any points (x,y) with positive $f(x,y)$ on those boundaries, the values of $f(x,y)$ at those points would also contribute to $F(0.5,1.5)$.

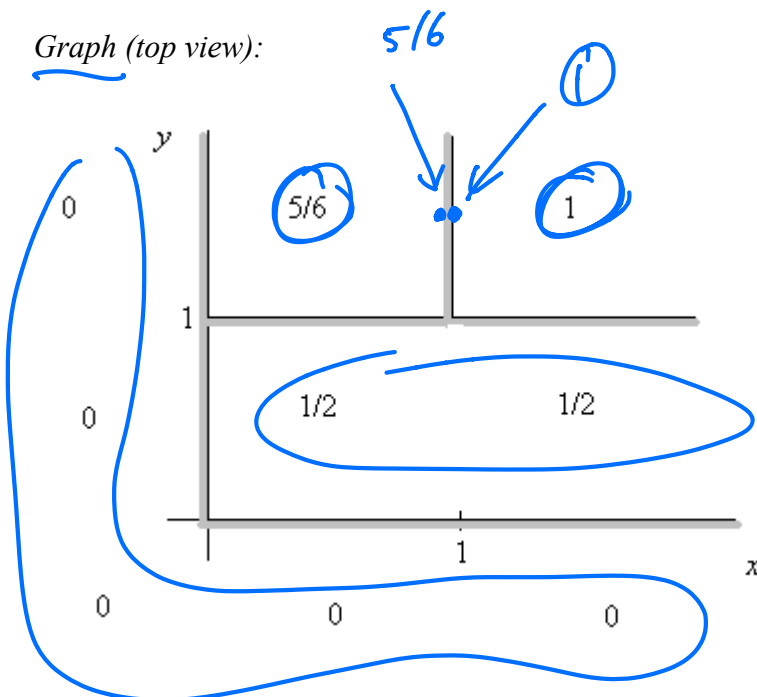


$(x < \neq y)$

We find that X and Y have joint cdf

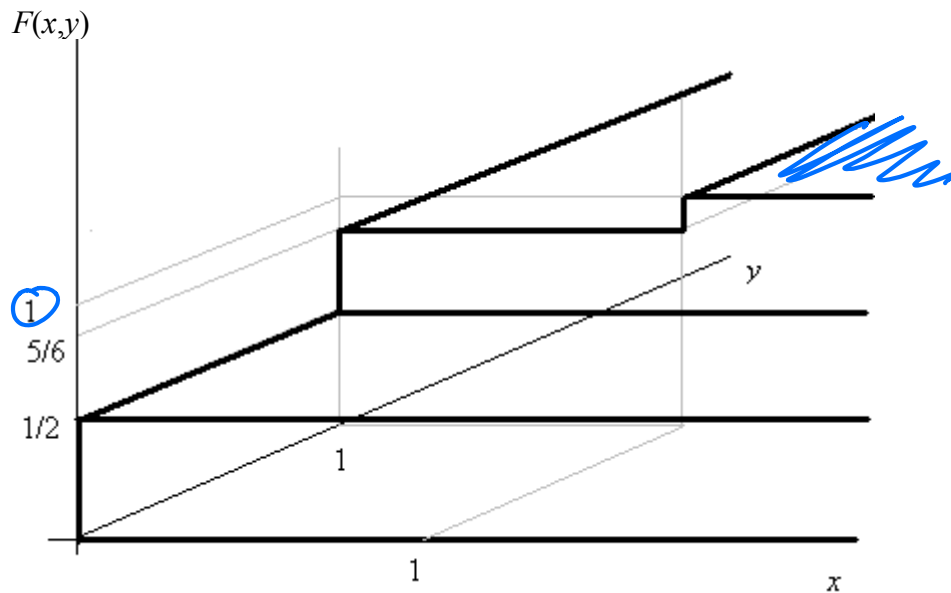
$$F(x, y) = \begin{cases} 0, & x < 0 \text{ or } y < 0 \text{ (or both)} \\ 1/2, & x \geq 0, 0 \leq y < 1 \\ 5/6, & 0 \leq x < 1, y \geq 1 \\ 1, & x, y \geq 1 \end{cases}$$

Graph (top view):



The shading here indicates that, for example, $F(1, 1.5) = 1$, $F(0.999, 1.5) = 5/6$.

3-d graph (non-assessable):



Some properties of all joint cdf's:

1. $F(x, y) \rightarrow 0$ as $x \rightarrow -\infty$ or $y \rightarrow -\infty$ (or both).
2. $F(x, y) \rightarrow 1$ as $x \rightarrow \infty$ and $y \rightarrow \infty$.
3. $F(x, y)$ is nondecreasing in both x and y directions.
4. $F(x, y)$ is right-continuous in both x and y directions.

Note that with these properties in mind, the joint cdf of X and Y in our example could be written more simply as

$$F(x, y) = \begin{cases} 1/2, & x > 0, 0 < y < 1 \\ 5/6, & 0 < x < 1, y > 1 \end{cases}$$

But, for clarity, it is best to write *joint* cdf's in full detail.

Now for some more definitions.

The *marginal pdf* of X is

$$P(X=x) = f(x) = \sum_y f(x, y) = \sum_y P(X=x, Y=y)$$

This pdf defines the *marginal probability distribution* of X .

We may also write $f(x)$ as $f_X(x)$.

In our example:

$$f_X(0) = \sum_y p(0,y) = p(0,0) + p(0,1) = \frac{1}{2} + \frac{1}{3} = \frac{5}{6}$$

$$f_X(1) = \sum_y p(1,y) = p(1,1) = \frac{1}{6}$$

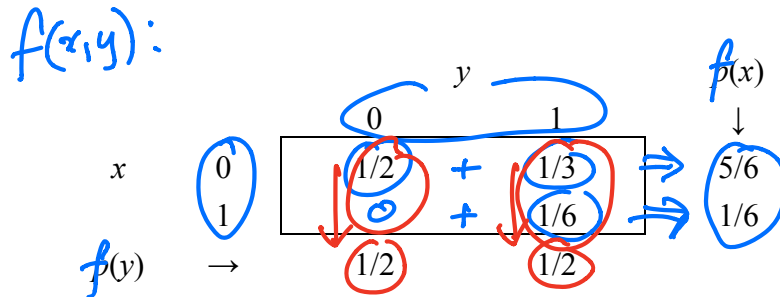
Thus $f_X(x) = \begin{cases} 5/6, & x=0 \\ 1/6, & x=1 \end{cases}$

In words we may say that X 's marginal probability distribution is Bernoulli with parameter $1/6$. That is, $X \sim \text{Bern}(1/6)$.

(This makes sense: X is the number of 6's on one roll of a die.)

Similarly, we find that $Y \sim \text{Bern}(1/2)$.

Note that what we have done is equivalent to computing column and row totals:

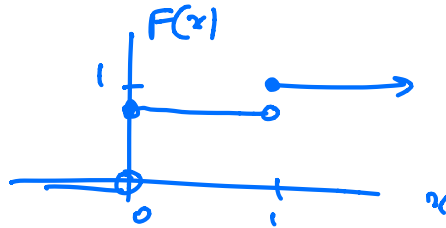


The marginal cdf of X is

$$F(x) = P(X \leq x) = \sum_{t \leq x} f_X(t) \quad (\text{Ch 3})$$

This is just the ordinary cdf of X , and can be computed in the usual way.

For example, $F(x) = \begin{cases} 0, & x < 0 \\ 5/6, & 0 \leq x < 1 \\ 1, & x \geq 1 \end{cases}$



The conditional pdf of X given that $Y=y$ is

$$p(x|y) = \frac{p(x,y)}{p(y)} \quad (\text{Or equivalently, } p(x|y) = \frac{P(X=x, Y=y)}{P(Y=y)}.)$$

$P(X=x|Y=y) \rightarrow$

This density defines the conditional probability distribution of X given that $Y=y$.

In our example, what's the conditional probability distribution of X given that $Y = 1$?

$$p(x|1) = \frac{p(x,1)}{p_Y(1)} \text{ for } x = 0, 1.$$

Explicitly:

$$p_{X|Y}(0|1) = \frac{p_{X,Y}(0,1)}{p_Y(1)} = \frac{1/3}{1/2} = \frac{2}{3}$$

$$p_{X|Y}(1|1) = \frac{p_{X,Y}(1,1)}{p_Y(1)} = \frac{1/6}{1/2} = \frac{1}{3}$$

So $p(x|1) = \begin{cases} 2/3, & x = 0 \\ 1/3, & x = 1 \end{cases}$

Thus $(X|Y=1) \sim \text{Bern}(1/3)$.

2, 4 or 6

(This makes sense: If an even number comes up (2, 4 or 6) then there is obviously a one-in-three chance of that number being 6. So $P(X=1|Y=1) = 1/3$, etc.)

What is the dsn of X given that $Y = 0$?

$Y = 0$ implies that a 1, 3 or 5 comes up, meaning that a 6 definitely does *not* come up.

So $P(X=1|Y=0) = 0$ and $P(X=0|Y=0) = 1$.

(If $Y = 0$ then $X = 0$ with probability one.)

w.p. 1

Thus $(X|Y=0) \sim \text{Bern}(0)$, and $p(x|0) = I(x=0) = \begin{cases} 1, & x = 0 \\ 0, & x = 1 \end{cases}$

This is an example of a degenerate dsn (a discrete dsn with only one possible value).

The conditional cdf of $(X|Y=y)$ is

$$F(x|y) = P(X \leq x | Y = y).$$

not $P(X \leq x | Y \leq y)$
(x)

This function can be computed in the same way as the marginal cdf of X , but using $p(x|y)$ instead of $p(x)$.

$$\text{For example, } F(x|1) = \begin{cases} 0, & x < 0 \\ 2/3, & 0 \leq x < 1 \\ 1, & x \geq 1 \end{cases}$$

Note: For clarity, we could also write this as $F_{X|Y}(x|1)$.

$$= P(X \leq x | Y = 1)$$

$F(1|0) = ?$ confusing
 $= F_{X|Y}(1,1) \text{ or } F_Y(x|1,1)$

Independence of random variables

Recall that two events A and B are independent if $P(AB) = P(A)P(B)$. Similarly...

Two random variables X and Y are *independent* if

$$p(x,y) = p(x)p(y) \text{ for all } x \text{ and } y. \quad (*)$$

We then write $X \perp Y$.

If $(*)$ is false for some x and y , then X and Y are *dependent*, and we write $X \not\perp Y$.

In our example, are X and Y independent?

Recall that $p_{X,Y}(1,1) = 1/6$, $p_X(1) = 1/6$, $p_Y(1) = 1/2$.

Thus $p_{X,Y}(1,1) \neq p_X(1)p_Y(1)$.

Therefore X and Y are not independent.

NB: If $p(x|y) = p(x)$ or $p(y|x) = p(y)$ then $X \perp Y$.

If $p(x|y) \neq p(x)$ or $p(y|x) \neq p(y)$ then $X \not\perp Y$.

In our example, $p_{X|Y}(1|1) = 1/3$ and $p_X(1) = 1/6$.

These are not the same, and therefore $X \not\perp Y$.

Multivariate expectation

$$\text{ch 3: } E g(X) = \sum_x g(x) p(x)$$

$$Eg(X,Y) = \sum_{x,y} g(x,y) p(x,y) = \sum_x \sum_y g(x,y) p(x,y)$$

In our example, what is the expected value of XY ?

$$E(XY) = \sum_{x,y} xy p(x,y) = 0(0)p(0,0) + 0(1)p(0,1) + 1(1)p(1,1) \\ = 0 + 0 + p(1,1) = 1/6.$$

Also, what is the expected value of $(X+1)^Y$?

$$E\{(X+1)^Y\} = \sum_{x,y} (x+1)^y p(x,y) = (0+1)^0 p(0,0) + (0+1)^1 p(0,1) + (1+1)^1 p(1,1) \\ = (0+1)^0 (1/2) + (0+1)^1 (1/3) + (1+1)^1 (1/6) = 7/6.$$

0	1
0	1
0	1

(For each example here, it may help to draw a matrix showing a cell for each value of the pair (x,y) . In each cell write the value of the pdf and the value of the function.)

Covariance and correlation

The covariance between X and Y is

$$\text{Cov}(X, Y) = E\{(X - EX)(Y - EY)\}.$$

What's the covariance between X and Y in our example?

Recall that $X \sim \text{Bern}(1/6)$ and $Y \sim \text{Bern}(1/2)$.

Therefore $EX = 1/6$ and $EY = 1/2$.

It follows that

$$\begin{aligned} \text{Cov}(X, Y) &= \sum_{x,y} \left(x - \frac{1}{6}\right) \left(y - \frac{1}{2}\right) p(x, y) \\ &= \left(0 - \frac{1}{6}\right) \left(0 - \frac{1}{2}\right) p(0, 0) + \left(0 - \frac{1}{6}\right) \left(1 - \frac{1}{2}\right) p(0, 1) + \left(1 - \frac{1}{6}\right) \left(1 - \frac{1}{2}\right) p(1, 1) \\ &= \left(\frac{1}{12}\right) \frac{1}{2} + \left(-\frac{1}{12}\right) \frac{1}{3} + \left(\frac{5}{12}\right) \frac{1}{6} = \frac{1}{12}. \end{aligned}$$

A useful result: $\text{Cov}(X, Y) = E(XY) - (EX)EY$.

$$\begin{aligned} \text{Proof: LHS} &= E\{(X - \mu_X)(Y - \mu_Y)\} = E(XY - X\mu_Y - Y\mu_X + \mu_X\mu_Y) \\ &= E(XY) - \mu_Y EX - \mu_X EY + \mu_X\mu_Y \\ &= E(XY) - \mu_Y\mu_X - \mu_X\mu_Y + \mu_X\mu_Y = \text{RHS.} \end{aligned}$$

Let's illustrate this result by using it to check $\text{Cov}(X, Y)$ in our example.

Recall that $E(XY) = 1/6$.

It follows that $\text{Cov}(X, Y) = \frac{1}{6} - \frac{1}{6} \left(\frac{1}{2}\right) = \frac{1}{12}$, as before.

The correlation between X and Y is

$$\rho = \text{Corr}(X, Y) = \frac{\text{Cov}(X, Y)}{SD(X)SD(Y)}.$$

ρ
rho

$SD(X)$
 $SD(Y)$

What's the correlation between X and Y in our example?

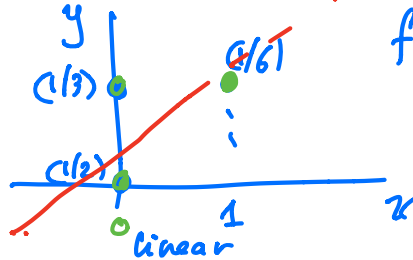
$$X \sim \text{Bern}(1/6) \Rightarrow \text{Var}X = \frac{1}{6} \left(1 - \frac{1}{6}\right) = \frac{5}{36} \Rightarrow SD(X) = \frac{\sqrt{5}}{6}.$$

$$Y \sim \text{Bern}(1/2) \Rightarrow \text{Var}Y = \frac{1}{2} \left(1 - \frac{1}{2}\right) = \frac{1}{4} \Rightarrow SD(Y) = \frac{1}{2}.$$

$$\text{So } \rho = \frac{1/12}{\frac{\sqrt{5}}{6} \times \frac{1}{2}} = 0.4472.$$

positive
linear
relationship

$f(x, y)$:



Notes

1. ρ provides information about the relationship between X and Y .
If $\rho > 0$ then high values of X are associated with high values of Y (eg $\rho = 0.4472$ above).

If $\rho < 0$ then high values of X are associated with low values of Y .

2. $-1 \leq \rho \leq 1$.

$\text{corr}(X, Y)$
 $\rho =$

(By contrast, $\text{Cov}(X, Y)$ can be anything from minus infinity to infinity.
So ρ is easier to interpret.)

3. $X \perp Y \Rightarrow \rho = 0$. (Prove this as an exercise.)

4. $\rho \neq 0 \Rightarrow X \not\perp Y$.

(In logical parlance, this follows from Note 3 by the principle of contraposition. The contrapositive of $P \Rightarrow Q$ is $\text{not}Q \Rightarrow \text{not}P$.

For example, since it is true that all dogs are animals, it follows by contraposition that if something is not an animal, it is also not a dog.)

5. $\rho = 0 \not\Rightarrow X \perp Y$.

(A proof of this fact is provided by Example 5.24 in the text.)