

APPLIED STATISTICS

Logistic Regression for Two-Category Response Variables and Its Estimation

Dr Tao Zou

Research School of Finance, Actuarial Studies & Statistics
The Australian National University

Last Updated: Tue Sep 26 13:52:35 2017

Overview

- Two-Category Response Variables
- Motivating Example
- Binary Logistic Regression Model
- Estimation of Binary Logistic Regression
- Prediction of a New Observation

References

1. **F.L. Ramsey and D.W. Schafer** (2012)
Chapter 20 of *The Statistical Sleuth*
2. ANU STAT3015 Lecture Notes
3. The slides are made by **R Markdown**.
<http://rmarkdown.rstudio.com>

Two-Category Response Variables

In numerous regression applications, the response variable of interest is a categorical variable taking two values.

In such situations the response can be represented by a binary indicator variable taking on values 0 and 1. For example:

- In a study on the effectiveness of a new drug, the response might be whether a given patient survived a 5-year period.
- In a study of home ownership, the response variable is whether a given individual owns a home.

Example: Anaesthetic Data

(Taken from STAT3015 notes.)

The potency of an anaesthetic agent is measured in terms of the minimum concentration at which at least 50% of patients exhibit no response to stimulation.

Thirty patients were given a particular anaesthetic at various predetermined concentrations for 15 minutes before a stimulus was applied.

The response variable was simply an indication as to whether the patient responded to the stimulus in any way.

"Response" is 1 if the patient responded to the stimulus.

R Code

```
setwd('~\\Desktop\\Research\\AppliedStat2017\\L9')  
a=read.csv('anaesthetic.csv');a
```

##	Concentration	Response
----	---------------	----------

## 1	0.8	1
------	-----	---

## 2	0.8	1
------	-----	---

## 3	0.8	1
------	-----	---

## 4	0.8	1
------	-----	---

## 5	0.8	1
------	-----	---

## 6	0.8	1
------	-----	---

## 7	0.8	0
------	-----	---

## 8	1.0	1
------	-----	---

## 9	1.0	1
------	-----	---

## 10	1.0	1
-------	-----	---

## 11	1.0	1
-------	-----	---

## 12	1.0	0
-------	-----	---

## 13	1.2	1
-------	-----	---

## 14	1.2	1
-------	-----	---

## 15	1.2	0
-------	-----	---

## 16	1.2	0
-------	-----	---

## 17	1.2	0
-------	-----	---

## 18	1.2	0
-------	-----	---

## 19	1.4	1
-------	-----	---

## 20	1.4	1
-------	-----	---

## 21	1.4	0
-------	-----	---

## 22	1.4	0
-------	-----	---

## 23	1.4	0
-------	-----	---

## 24	1.4	0
-------	-----	---

## 25	1.6	0
-------	-----	---

## 26	1.6	0
-------	-----	---

## 27	1.6	0
-------	-----	---

## 28	1.6	0
-------	-----	---

## 29	2.5	0
-------	-----	---

## 30	2.5	0
-------	-----	---

the number of response = 1

$$\frac{(1+1+1+1+1)}{7} = \frac{6}{7}$$

sample mean of responses at $X = 0.8$

var

the proportion of response = 1



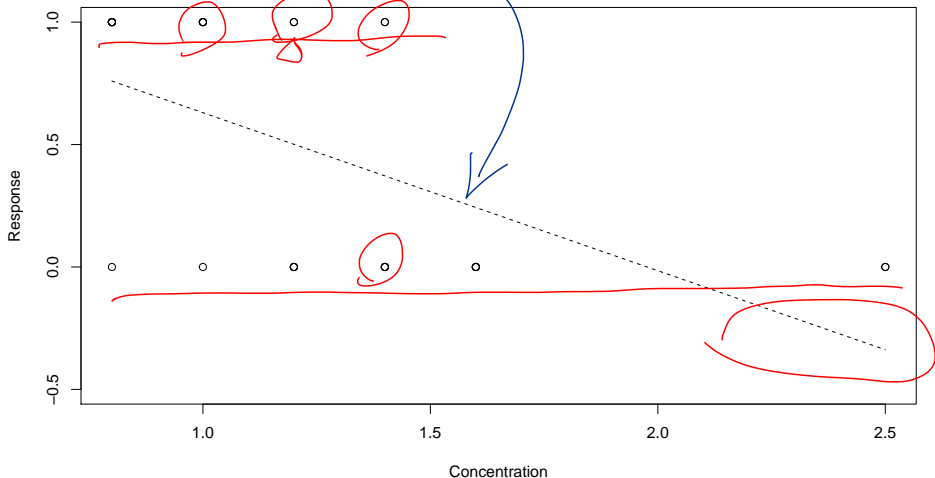
$[0, 1]$

should be in $[0, 1]$.

/L22

R Code (Con'd)

```
attach(a)
plot(Concentration, Response, ylim=c(-0.5,1))
fit=lm(Response-Concentration)
lines(Concentration, fit$fitted, lty=2)
```



On this scale, a linear regression does not seem appropriate.

Violation of Linear Regression Assumptions

Y: Response; X: Concentration.

1. Y not conform normality assumption, since Y only takes values of 0 and 1.

2. *target*
→ returns the means of target by
`tapply(Response, Concentration, mean)`
Concentration

```
##      0.8      1      1.2      1.4      1.6      2.5  
## 0.8571429 0.8000000 0.3333333 0.3333333 0.0000000 0.0000000
```

→ sample mean by different

Given $X = 0.8$, the sample mean of Y is 0.857;

given $X = 1.0$, the sample mean of Y is 0.800;

given $X = 1.2$, the sample mean of Y is 0.333;

given $X = 1.4$, the sample mean of Y is 0.333;

given $X = 1.6$, the sample mean of Y is 0.000;

given $X = 2.5$, the sample mean of Y is 0.000.

$\frac{6}{7}$

Based on data, the sample mean is actually the proportion that $Y = 1$ given $X = x$, and hence should be in the interval $[0, 1]$.

This indicates that the mean of Y given $X = x$ ($\mu\{Y|X = x\}$) should be in $[0, 1]$.
But in linear regression, $\mu\{Y|X = x\} = \beta_0 + \beta_1 x$ can take values outside of $[0, 1]$.

Violation of Linear Regression Assumptions (Con'd)

3.

```
tapply(Response, Concentration, var)
```

```
##      0.8      1      1.2      1.4      1.6      2.5  
## 0.1428571 0.2000000 0.2666667 0.2666667 0.0000000 0.0000000
```

Given $X = 0.8$, the sample variance of Y is 0.143;

given $X = 1.0$, the sample variance of Y is 0.200;

given $X = 1.2$, the sample variance of Y is 0.267;

given $X = 1.4$, the sample variance of Y is 0.267;

given $X = 1.6$, the sample variance of Y is 0.000;

given $X = 2.5$, the sample variance of Y is 0.000.

The constant variance assumption is violated, $\sigma\{Y|X = x\}$ are not constant.

Problem 3 could be fixed using weighted regression. Problem 1 may not be a problem since LS estimates are robust to some non-normal distributions.

Problem 2 is more problematic.

Generalised Linear Model (GLM)

The above example indicates that the mean of Y given $X = x$ (i.e., $\mu\{Y|X = x\}$) should be in the interval $[0, 1]$ for a binary response Y .

But in the linear regression, $\mu\{Y|X = x\} = \beta_0 + \beta_1 x$ can take values outside of $[0, 1]$.

So how about we find some transformation $h(\cdot)$ such that

$$\mu\{Y|X = x\} = h(\beta_0 + \beta_1 x) \in [0, 1] \text{ for sure?}$$

eg,

$$h(v) = \frac{e^v}{1+e^v}$$

$$h^{-1}(u) = \log \frac{u}{1-u} =: g(u)$$

Usually we consider the function $h(\cdot)$ to force that

$$u = h(v) \Rightarrow v = h^{-1}(u) = \underline{g}(u), \text{ say,}$$

definition
of inverse
function

namely g is the inverse function of h . Also h is the inverse function of g , i.e., $h(v) = g^{-1}(v)$.

link function

Then

$$\beta_0 + \beta_1 x = h^{-1}(\mu\{Y|X = x\}) = \underline{g}(\mu\{Y|X = x\}).$$

Generalised Linear Model (Con'd)

A generalised linear model (GLM) is a model where the mean of the response is related to the explanatory variables via the following relationship:

$$g(\mu\{Y|X_1, \dots, X_k\}) = \beta_0 + \beta_1 X_1 + \dots + \beta_k X_k.$$

This relationship is linear in the parameters. The function $g(\cdot)$ is called the link function.

The choice of link function $g(\cdot)$ depends on the type of the response variable, and is not limited to a binary response Y .

In this lecture we introduce the link function for two-category response Y (in such situations the response can be represented by a binary indicator variable taking on values 0 and 1).

We call this proposed model with a specific link for two-category response: **binary logistic regression** model.

Other link functions lead to other GLMs, where in these cases the response is not necessarily binary.

Overview of This Course

	Continuous X + Categorical X
Continuous Y	MLR + Indicator Variables
Two-Category Y	Binary Logistic Regression + Indicator Variables

Binary Logistic Regression Model Assumptions

1. **Bernoulli distribution:** There is a Bernoulli distributed (sub)population of responses for given values of the explanatory variables ($X_1 = x_1, \dots, X_k = x_k$). That means if we let $X = (X_1, \dots, X_k)$, the probability that $Y = 1$ given X is

$$\begin{aligned} P(Y = 1|X) &= \pi(X) \in [0, 1], \text{ and} \\ P(Y = 0|X) &= 1 - P(Y = 1|X) = 1 - \pi(X). \\ \mu\{Y|X\} &= 1 \times P(Y = 1|X) + 0 \times P(Y = 0|X) = \pi(X) \in [0, 1]. \end{aligned}$$

2. **Generalised Linearity:** The transformation of the mean of response falls on a linear function of the explanatory variables

$$\Rightarrow \mu\{Y|X\} = g^{-1}(\beta_0 + \beta_1 X_1 + \dots + \beta_k X_k)$$

$$g(\mu\{Y|X\}) = \beta_0 + \beta_1 X_1 + \dots + \beta_k X_k, \text{ for } X = (X_1, \dots, X_k),$$

where $g(u) = \log\{u/(1-u)\}$, which is called logit link function.

Binary Logistic Regression Model Assumptions (Con'd)

Remark: the inverse function of the logit link function is

$\pi(X)$
 $P(Y=1|X)$

$$g^{-1}(v) = \frac{e^v}{1 + e^v} \in [0, 1]. \rightarrow \text{logistic function}$$

Then

$$\mu\{Y|X\} = g^{-1}(\beta_0 + \beta_1 X_1 + \cdots + \beta_k X_k) \in [0, 1],$$

which is consistent with the range $\mu\{Y|X\} = \pi(X) \in [0, 1]$.

3. Independence: Observations

$$(X_{1,1}, \cdots X_{k,1}, Y_1),$$

$$\vdots$$

$$(X_{1,n}, \cdots X_{k,n}, Y_n),$$

are independent, where n is the sample size.

Binary Logistic Regression and Interpretation

Based on the above assumptions,

$$\begin{aligned} P(Y = 1|X) &= \mu\{Y|X\} = g^{-1}(\beta_0 + \beta_1 X_1 + \dots + \beta_k X_k) \\ &= \frac{e^{\beta_0 + \beta_1 X_1 + \dots + \beta_k X_k}}{1 + e^{\beta_0 + \beta_1 X_1 + \dots + \beta_k X_k}} \end{aligned}$$

Then we compute

$$\frac{P(Y = 1|X)}{1 - P(Y = 1|X)} = e^{\beta_0 + \beta_1 X_1 + \dots + \beta_k X_k}$$

which is called odds that $Y = 1$ given X .

$$\beta_0 + \beta_1 X_1 + \dots + \beta_k X_k = 0$$

- odds = 1 means there is a 50% chance that $Y = 1$ will occur
 $P(Y = 1|X) = 0.5$.
- odds > 1 means there is a better than 50% chance that $Y = 1$ will occur $P(Y = 1|X) > 0.5$.
- odds < 1 means there is less than 50% chance ~~change~~ that $Y = 1$ will occur $P(Y = 1|X) < 0.5$.

Hence, odds is another way to describe probability.

Binary Logistic Regression and Interpretation (Con'd)

Then

hold other constant

at

odds

$$= \frac{P(Y=1|X_1=x_1+1, X_2=x_2, \dots, X_k=x_k)}{1-P(Y=1|X_1=x_1+1, X_2=x_2, \dots, X_k=x_k)} = e^{\beta_0 + \beta_1(x_1+1) + \dots + \beta_k x_k} = e^{\beta_0 + \beta_1 x_1 + \dots + \beta_k x_k} e^{\beta_1} \text{ and}$$

at

odds

$$= \frac{P(Y=1|X_1=x_1, X_2=x_2, \dots, X_k=x_k)}{1-P(Y=1|X_1=x_1, X_2=x_2, \dots, X_k=x_k)} = e^{\beta_0 + \beta_1 x_1 + \dots + \beta_k x_k}$$

With the other variables held constant, if X_1 is increased by 1 unit, the odds that $Y = 1$ will change by a multiplicative factor of e^{β_1} .

$\Delta \rightarrow \beta_1$

Estimation of Binary Logistic Regression Parameters

For all generalised linear models, the method of least squares is replaced by the method of ~~maximum likelihood estimation (MLE)~~.

Consider the response $Y = y$,

$$y = 1 \Rightarrow$$

$$P(Y = 1|X) = \pi(X) = \{\pi(X)\}^1 \{1 - \pi(X)\}^{1-1} = \{\pi(X)\}^y \{1 - \pi(X)\}^{1-y},$$

$$y = 0 \Rightarrow$$

$$P(Y = 0|X) = 1 - \pi(X) = \{\pi(X)\}^0 \{1 - \pi(X)\}^{1-0} = \{\pi(X)\}^y \{1 - \pi(X)\}^{1-y}$$

Hence,

$$P(Y = y|X) = \{\pi(X)\}^y \{1 - \pi(X)\}^{1-y}.$$

y can be both 0, 1.

It is worth noting that

→ logistic function

$$\pi(X) = \mu\{Y|X\} = g^{-1}(\beta_0 + \beta_1 X_1 + \dots + \beta_k X_k) =: p(\beta_0, \dots, \beta_k), \text{ say.}$$

Since $X = (X_1, \dots, X_k)$, we have

$$\begin{aligned} P(Y = y|X_1, \dots, X_k) &= \{\pi(X_1, \dots, X_k)\}^y \{1 - \pi(X_1, \dots, X_k)\}^{1-y} \\ &= \{p(\beta_0, \dots, \beta_k)\}^y \{1 - p(\beta_0, \dots, \beta_k)\}^{1-y}. \end{aligned}$$

Estimation of Binary Logistic Regression Parameters (Con'd)

Given the independent observations

$$(\underline{X}_{1,1}, \dots, \underline{X}_{k,1}, \underline{Y}_1 = \underline{y}_1),$$

\vdots

$$(\underline{X}_{1,n}, \dots, \underline{X}_{k,n}, \underline{Y}_n = \underline{y}_n),$$

$$\begin{aligned} \underline{P(Y_i = y_i | X_{1,i}, \dots, X_{k,i})} &= \{\pi(X_{1,i}, \dots, X_{k,i})\}^{y_i} \{1 - \pi(X_{1,i}, \dots, X_{k,i})\}^{1-y_i} \\ &= \{\underline{p_i}(\beta_0, \dots, \beta_k)\}^{y_i} \{1 - \underline{p_i}(\beta_0, \dots, \beta_k)\}^{1-y_i}, \text{ say.} \end{aligned}$$

The likelihood is defined by

$$\begin{aligned} \mathcal{L}(\beta_0, \dots, \beta_k) &= P(\underline{Y_1 = y_1, \dots, Y_n = y_n} \mid \text{given all } X\text{s}) \\ &= \prod_{i=1}^n P(Y_i = y_i | X_{1,i}, \dots, X_{k,i}) \\ &= \prod_{i=1}^n \{p_i(\beta_0, \dots, \beta_k)\}^{y_i} \{1 - p_i(\beta_0, \dots, \beta_k)\}^{1-y_i}, \end{aligned}$$

which is the probability that we observe $\underline{Y_1 = y_1, \dots, Y_n = y_n}$ given all X s.

Estimation of Binary Logistic Regression Parameters (Con'd)

The maximum likelihood estimation (MLE) takes the “logic” that since we observe $Y_1 = y_1, \dots, Y_n = y_n$, there should be a pretty good chance that the observed outcome happens. Otherwise, we should not observe it.

Hence, the probability that we observe $Y_1 = y_1, \dots, Y_n = y_n$ given all X s, namely the likelihood $\mathcal{L}(\beta_0, \dots, \beta_k)$, should be very large.

We choose MLE $\hat{\beta}_0, \dots, \hat{\beta}_k$ numerically to maximize the probability $\mathcal{L}(\beta_0, \dots, \beta_k)$.

Different from the least squares estimation, we do not have a formula for MLE $\hat{\beta}_0, \dots, \hat{\beta}_k$. The MLE can only be obtained numerically.

Fitted Probabilities

Using MLE $\hat{\beta}_0, \dots, \hat{\beta}_k$, the estimated mean function is given by:

$$\hat{\mu}\{Y|X\} = g^{-1}(\hat{\beta}_0 + \hat{\beta}_1 X_1 + \dots + \hat{\beta}_k X_k) \text{ (plug-in idea).}$$

Fitted

The ~~fitting~~ probabilities are given by

$$\begin{aligned} \hat{\pi}(X) &= \hat{\mu}\{Y|X\} \\ &= g^{-1}(\hat{\beta}_0 + \hat{\beta}_1 X_1 + \dots + \hat{\beta}_k X_k) \\ &= \frac{e^{\hat{\beta}_0 + \hat{\beta}_1 X_1 + \dots + \hat{\beta}_k X_k}}{1 + e^{\hat{\beta}_0 + \hat{\beta}_1 X_1 + \dots + \hat{\beta}_k X_k}}. \end{aligned}$$

When we talk about fitted probabilities, X is usually from the training dataset (see Lecture Notes 8).

When X_{new} is from the new dataset or the test dataset, we actually talk about prediction.

Prediction of a New Observation

The forecast of probability is given by

$$\begin{aligned}\hat{\pi}(X_{\text{new}}) &= \hat{\mu}\{Y|X_{\text{new}}\} \\ &= g^{-1}(\hat{\beta}_0 + \hat{\beta}_1 X_{1,\text{new}} + \cdots + \hat{\beta}_k X_{k,\text{new}}) \\ &= \frac{e^{\hat{\beta}_0 + \hat{\beta}_1 X_{1,\text{new}} + \cdots + \hat{\beta}_k X_{k,\text{new}}}}{1 + e^{\hat{\beta}_0 + \hat{\beta}_1 X_{1,\text{new}} + \cdots + \hat{\beta}_k X_{k,\text{new}}}}.\end{aligned}$$

Recall that $P(Y = 1|X) = \pi(X)$ and $P(Y = 0|X) = 1 - P(Y = 1|X) = 1 - \pi(X)$.

Thus, if $P(Y = 1|X) > P(Y = 0|X)$ namely $\pi(X) > 0.5$, there is a better chance that $Y = 1$ will occur.

Hence, 0.5 is a commonly used threshold for predicting the response.

In conclusion, the prediction for the response Y_{new} at X_{new} is

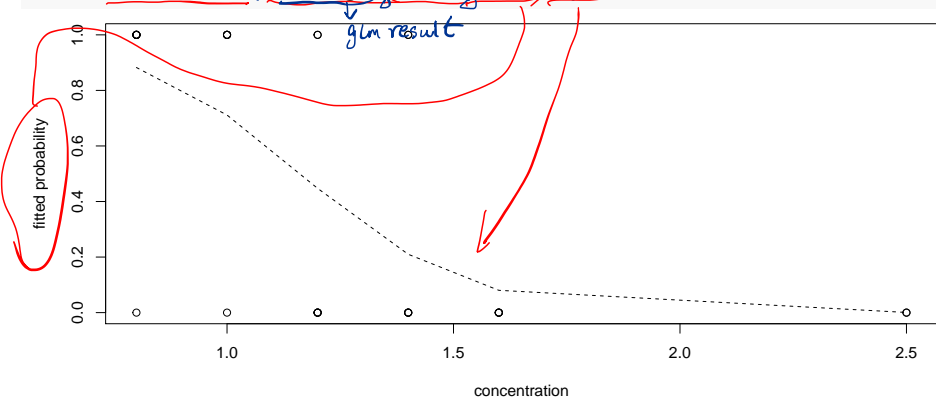
$$\hat{Y}_{\text{new}} = 1 \text{ if } \hat{\pi}(X_{\text{new}}) > 0.5; \hat{Y}_{\text{new}} = 0 \text{ otherwise.}$$

Or equivalently, \hat{Y}_{new} is the category that has the larger forecast of probability.

Example: Anaesthetic Data (Con'd)

```
##?glm
#fitting the logistic regression
ansth.logit=glm(Response~Concentration,family=binomial(link=logit))
plot(Concentration,Response,xlab="concentration",ylab="fitted probability")
lines(Concentration,ansth.logit$fitted.values,lty=2)
```

Bernoulli is a special case of



```
detach(a)
```

All the fitted probabilities are between zero and one.

Example: Anaesthetic Data (Con'd)

By using this example, we might be interested in predicting whether a patient will respond to the stimulus if an anaesthetic at a new concentration of 1.5 is given. The forecast of probability is:

```
Xnew=data.frame(Concentration=1.5)
predict(ansth.logit,Xnew,type='response')
```

1
0.1322204

Handwritten notes:
↓ result of g(m())
→ prediction of probability that $Y=1$

For this patient we predict a response of 0, i.e., we predict that the patient will not respond to the stimulus.