STA305/1004-Class19

Today's Class

- ► Sample size for ANOVA
- ► Factorial designs at two levels
- ► Cube plots
- ► Calculation of factorial effects

Sample size for ANOVA - Designing a study to compare more than two treatments

Compare more than

2 means e.g. Ho:
$$\mu_1 = \mu_2 = \mu_3$$

H1: $\mu_i \neq \mu_j$

- Consider the hypothesis that k means are equal vs. the alternative that at least two differ
- ▶ What is the probability that the test rejects if at least two means differ?
- ▶ Power = 1 P(Type II error) is this probability.

Sample size for ANOVA - Designing a study to compare more than two treatments

The null and alternative hypotheses are:

$$H_0: \mu_1 = \mu_2 = \cdots = \mu_k \text{ vs. } H_1: \mu_i \neq \mu_i.$$

The test rejects at level $\boldsymbol{\alpha}$ if

$$MS_{Treat}/MS_E \geq F_{k-1,N-K,\alpha}$$
.

The power of the test is

$$1 - \beta = P\left(MS_{Treat}/MS_{E} \ge F_{k-1,N-K,\alpha}\right),$$

when H_0 is false.

Sample size for ANOVA - Designing a study to compare more than two treatments

When H_0 is false it can be shown that:

- MS_{Treat}/σ^2 has a non-central Chi-square distribution with k-1 degrees of freedom and non-centrality parameter δ .
- MS_{Treat}/MS_E has a non-central F distribution with the numerator and denominator degrees of freedom k-1 and N-k respectively, and non-centrality parameter

 $\delta = \frac{\sum_{i=1}^{k} n_i (\mu_i - \bar{\mu})^2}{\sigma^2},$ between variance

With

where n_i is the number of observations in group i, $\bar{\mu}=\sum_{i=1}^k \mu_i/k$, and σ^2 is the within group error variance .

This is dentoted by $F_{k-1,N-k}(\delta)$.

Direct calculation of Power

► The power of the test is

$$P\left(F_{k-1,N-k}(\delta) > F_{k-1,N-K,\alpha}\right)$$

 $P(F_{k-1,N-k}(\delta) > F_{k-1,N-K,\alpha})$. as G gets larger then power

- ▶ The power is an increasing function δ •
- ▶ The power depends on the true values of the treatment means μ_i , the error variance σ^2 , and sample size n_i .
- ▶ If the experimentor has some prior idea about the treament means and error variance the sample size (number of replications) that will guaruntee a pre-assigned power of the test.

Blood coagulation example - sample size

anova(lm.diets)

Suppose that an investigator would like to replicate the blood coagulation study with only 3 animals per diet. In this case k=4, $n_i=3$. The treatment means from the initial study are:

Average 61 66 68				
	Average	61	66	68

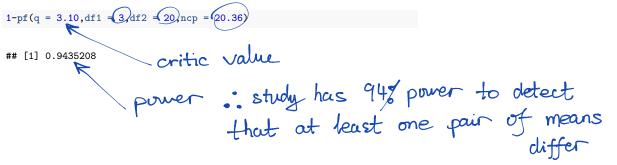
- $\mu_1 = 61, \ \mu_2 = 66, \ \mu_3 = 68, \ \mu_4 = 61.$
- ▶ The error variance σ^2 was estimated as $MS_E = 5.6$.
- ▶ Assuming that the estimated values are the true values of the parameters, the non-centrality parameter of the *F* distribution is

$$\delta = 3 \times \left((61 - 64)^2 + (66 - 64)^2 + (68 - 64)^2 + (61 - 64)^2 \right) / 5.6 = 20.35714$$

Blood coagulation example - sample size

If we choose $\alpha=0.05$ as the significance level then $F_{3,20,0.05}=3.0983912$. The power of the test is then $P(F_{3,20}(20.36)>3.10)=0.94.$

This was calculated using the CDF for the F distribution in R pf().



Calculating power and sample size using the pwr library

- ► There are several libraries in R which can calculate power and sample size for statistical tests. The library pwr() has a function
- pwr.anova.test(k = NULL, n = NULL, f = NULL, sig.level = 0.05, power = NULL)

for computing power and sample size.

- k Number of groups
- n Number of observations (per group)
- f Effect size
- ► The effect size is the square root of the non-centrality parameter of the non-central *F* distribution.

$$f = \sqrt{\frac{\sum_{i=1}^{k} n_i (\mu_i - \bar{\mu})^2}{\sigma^2}}.$$

where n_i is the number of observations in group i, $\bar{\mu} = \sum_{i=1}^k \mu_i/k$, and σ^2 is the within group error variance.

Calculating power and sample size using the pwr library

all the properties of as d power 1 power are still true! In the previous example $\delta = 20.35714$ so $f = \sqrt{20.35714} = 4.5118887$.

when Ho is true f=0 size is 4x3=12 then power = 0.05

When Ho is true

Ho:
$$W_1 = W_2 = W_3 = W_4$$
 is true
$$S = \frac{\sum_{i=1}^4 n_i (|w_i - \overline{w}|^2)}{\sigma^2}$$

$$= 0$$

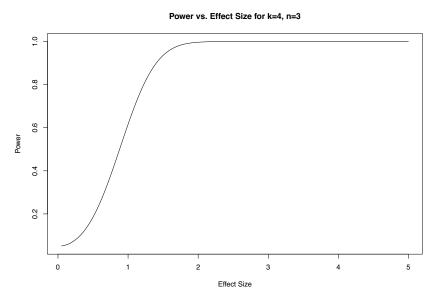
$$P(F_{3,20}(0) > F_{3,20}) = 0.05$$

= type I emor rate

- . Which ANOVA study should have a greater type I error rate (assume both studies with use the same significance
 - A study that is designed to detect an effect size of 0.5 - - 0.75Study 1: $1-\beta_1$ $1-\beta_2 > 1-\beta_1$ $5 + \beta_2 > 1-\beta_2 = \beta_1 > \beta_2$

Study 1:
$$1-\beta_1$$
 $1-\beta_2 > 1-\beta_1$
Study 2: $1-\beta_2$ $\Longrightarrow \beta_1 > \beta_2$

Calculating power and sample size using the pwr library



Calculating power using simulation

The general procedure for simulating power is:

- Use the underlying model to generate random data with (a) specified sample sizes,
 (b) parameter values that one is trying to detect with the hypothesis test, and (c) nuisance parameters such as variances.
- Run the estimation program (e.g., t.test(),lm()) on these randomly generated data.
- 3. Calculate the test statistic and p-value.
- 4. Do Steps 1–3 many times, say, N, and save the p-values. The estimated power for a level alpha test is the proportion of observations (out of N) for which the p-value is less than alpha.

Calculating power using simulation

One of the advantages of calculating power via simulation is that we can investigate what happens to power if, say, some of the assumptions behind one-way ANOVA are violated.

pur anova test essumes that data is normally distributed

Calculating power using simulation - R program

```
#Simulate power of ANOVA for three groups
NSIM <- 1000 # number of simulations
res <- numeric(NSIM) # store p-values in res
mu1 <- 2; mu2 <- 2.5; mu3 <- 2 # true mean values of treatment groups
sigma1 <- 1; sigma2 <- 1; sigma3 <- 1 #variances in each group
n1 <- 40; n2 <- 40; n3 <- 40 #sample size in each group
for (i in 1:NSIM) # do the calculations below N times
# generate sample of size n1 from N(mu1, sigma1^2)
v1 \leftarrow rnorm(n = n1, mean = mu1, sd = sigma1)
# generate sample of size n2 from N(mu2, sigma2^2)
y2 \leftarrow rnorm(n = n2, mean = mu2, sd = sigma2)
# generate sample of size n3 from N(mu3, sigma3~2)
v3 \leftarrow rnorm(n = n3, mean = mu3, sd = sigma3)
y \leftarrow c(y1,y2,y3) # store all the values from the groups
# generate the treatment assignment for each group
trt \leftarrow as.factor(c(rep(1,n1),rep(2,n2),rep(3,n3)))
m <- lm(y~trt) # calculate the ANOVA
res[i] \leftarrow anova(m)[1,5] # p-value of F test
sum(res<=0.05)/NSIM # calculate p-value</pre>
```

Example of a factorial design

Suppose that an investigator is interested in examining three components of a weight loss intervention. The three components are:

- 1. Keeping a food diary (yes/no)
- 2. Increasing activity (yes/no)
- 3. Home visit (yes/no)

Factorial designs

The investigator plans to investigate all $2x2x2=2^3=8$ combinations of experimental conditions.

The experimental conditions will be.

Expt condition	Keep food diary	Increase physical activity	Home visit	weight loss
1	No	No	No	
2	No	No	Yes	<i>y</i> 2
3	No	Yes	No	<i>y</i> 3
4	No	Yes	Yes	<i>y</i> 4
5	Yes	No	No	<i>y</i> 5
6	Yes	No	Yes	<i>y</i> 6
7	Yes	Yes	No	У 7
8	Yes	Yes	Yes	<i>y</i> 8

Factorial designs at two levels

► To perform a factorial design, you select a fixed number of levels of each of a number of factors (variables) and then run experiments in all possible combinations.

Factorial designs at two levels

- ▶ The factors can be quantitative or qualitative.
- ► Two levels of a quantitative variable could be two different temperatures or two different concentrations.
- Qualitative factors might be two types of catalysts or the presence and absence of some entity.

Factorial designs at two levels Factor 1 Factor 2 Is a study where patients vordomly assigned to one of four Beta -carotere Aspirin N (4) DAspirin only (beta-carotene placebo) N (I) # factors Y Y (2) 2) Beta carotene only Caspirin placebo) Y (3) 3 Aspirin and beta carotene (Neither (both aspirin & beta canotene placebos) a possible expt: JES factorial design. Conditions # Jarels -: the study involves all factor-level combinations it's a factorial design

Factorial design

The notation 2^3 identifies: - the number of factors (3) - the number of levels of each factor (2) - how many experimental conditions are in the design ($2^3 = 8$)

Factorial experiments can involve factors with different numbers of levels.

Factorial design

- 1. How many factors?
- 2. How many levels of each factor?
- 3. How many experimental conditions (runs)?
- > 2 factors at 4 levels
 2 factors at 3 levels
 1 factor at 2 levels
 -> 16×9×2=288 expt. conditions



Difference between ANOVA and Factorial Designs

In ANOVA the objective is to compare the individual experimental conditions with each other. In a factorial experiment the objective is generally to compare combinations of experimental conditions.

Let's consider the food diary study above. What is the effect of keeping a food diary?

Expt condition	Keep food diary	Increase physical activity	Home visit	weight loss
1	No	No	No	<i>y</i> ₁
2	No	No	Yes	<i>y</i> ₂
3	No	Yes	No	<i>y</i> 3
4	No	Yes	Yes	<i>y</i> 4
5	Yes	No	No	<i>y</i> ₅
6	Yes	No	Yes	<i>y</i> ₆
7	Yes	Yes	No	<i>y</i> ₇
8	Yes	Yes	Yes	<i>y</i> 8

We can estimate the effect of food diary by comparing the mean of all conditions where food diary is set to NO (conditions 1-4) and mean of all conditions where food diary set to YES (conditions 5-8). This is also called the **main effect** of food diary, the adjective *main* being a reminder that this average is taken over the levels of the other factors.

Difference between ANOVA and Factorial Designs

Expt condition	Keep food diary	Increase physical activity	Home visit	weight loss
1	No	No	No	<i>y</i> ₁
2	No	No	Yes	<i>y</i> ₂
3	No	Yes	No	<i>y</i> 3
4	No	Yes	Yes	<i>y</i> 4
5	Yes	No	No	<i>y</i> 5
6	Yes	No	Yes	<i>y</i> 6
7	Yes	Yes	No	У 7
8	Yes	Yes	Yes	<i>y</i> ₈

The main effect of food diary is:

$$\frac{y_1+y_2+y_3+y_4}{4}-\frac{y_5+y_6+y_7+y_8}{4}.$$

The main effect of physical activity is:

$$\frac{y_1+y_2+y_5+y_6}{4}-\frac{y_3+y_4+y_7+y_8}{4}.$$

The main effect of home visit is:

$$\frac{y_1 + y_3 + y_5 + y_7}{4} - \frac{y_2 + y_4 + y_6 + y_8}{4}.$$

Factorial designs at two levels

To perform a factorial design:

- 1. Select a fixed number of levels of each factor.
- $\begin{tabular}{ll} 2. & Run \ experiments \ in \ all \ possible \ combinations. \end{tabular}$

Factorial designs at two levels

- ▶ We will discuss designs where there are just two levels for each factor.
- ▶ Factors can be quantitative or qualitative.
- Two levels of quantitative variable could be two different temperatures or concentrations.
- Two levels of a quantitative variable could be two different types of catalysts or presence/absence of some entity.

Pilot plant investigation - example of factorial design

A pilot plant invsetiagtion employed a 2^3 factorial design (Box, Hunter, and Hunter (2005)) with

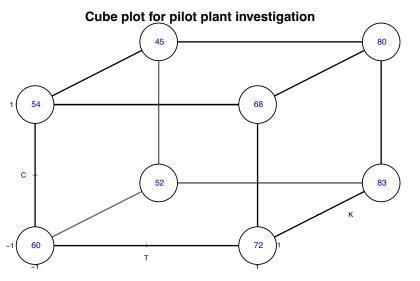
Factors	level 1	level 2
Temperature	160C°(-1)	180C°(+1)
Concentration	20% (-1)	40% (+1)
Catalyst	A (-1)	B(+1)

un	Т	С	K	у
1	-1	-1	-1	60
2	1	-1	-1	72
3	-1	1	-1	54
4	1	1	-1	68
5	-1	-1	1	52
6	1	-1	1	83
7	-1	1	1	45
8	1	1	1	80

► Each data value recorded is for the response yield *y* averaged over two duplicate runs.

Cube plots

```
library("FrF2")
bhh54 <- lm(y~T*C*K,data=tab0502)
cubePlot(bhh54,"T","K","C",main="Cube plot for pilot plant investigation")</pre>
```



Cube plots

- ▶ 8 run design produces 12 comparisons
- ▶ Each edge of cube only one factor changed while other 2 held constant.
- ► Therefore experimenter that believes in only changing one factor at a time is satisfied.