

9.12.) $X_1, \dots, X_n \stackrel{iid}{\sim} f(x|\theta) = \theta \exp(-\theta x)$

• Derive the LRT for:

$$H_0: \theta = \theta_0$$

$$H_1: \theta \neq \theta_0$$

$$\lambda = \frac{\sup_{\theta_0} L(\theta | \underline{x})}{\sup_{\theta} L(\theta | \underline{x})}$$

$$\begin{aligned} L(\theta | \underline{x}) &= \prod_{i=1}^n \theta \exp(-\theta x_i) \\ &= \theta^n \exp(-\theta \sum x_i) \end{aligned}$$

$$\ell(\theta | \underline{x}) = n \log(\theta) - \theta \sum x_i$$

$$\frac{\partial \ell}{\partial \theta} = \frac{n}{\theta} - \sum x_i = 0 \Rightarrow \hat{\theta} = \frac{1}{\bar{x}}$$

$$\therefore \lambda = \frac{\theta_0^n \exp(-\theta_0 \sum x_i)}{\hat{\theta}^n \exp(-\hat{\theta} \sum x_i)}$$

$$\lambda = \frac{\Theta_0^n \exp(-\Theta_0 \sum x_i)}{\left(\frac{1}{\bar{x}}\right)^n \exp\left(-\frac{n}{\bar{x}} \sum x_i\right)}$$

$$= \frac{\Theta_0^n \bar{x}^n \exp(-\Theta_0 \sum x_i)}{\exp(-n)}$$

$$= \frac{\Theta_0^n \bar{x}^n \exp(-\Theta \bar{x} n)}{\exp(-n)}$$

$$R = \{ \lambda < c \}$$

$$= \left\{ \frac{\Theta_0^n \bar{x}^n \exp(-\Theta \bar{x} n)}{\exp(-n)} < c \right\}$$

$$= \left\{ [\Theta_0 \bar{x} \exp(-\Theta \bar{x})]^n < c^* \right\}$$

$$= \left\{ \bar{x} \exp(-\bar{x}) < c^{**} \right\}$$

9.13)

$$H_0: \theta = 1$$

$$n = 10, \alpha = 0.05.$$

$$H_1: \theta \neq 1$$

$$R = \{ \bar{X} \exp(-\bar{X}) < c^{**} \}$$

a.) We want to show that Rejection region can be written in the following form:

$$R = \{ \bar{X} \leq x_0 \} \cup \{ \bar{X} \geq x_1 \}$$

* Let's differentiate the following:

$$g(a) = a \exp(-a)$$

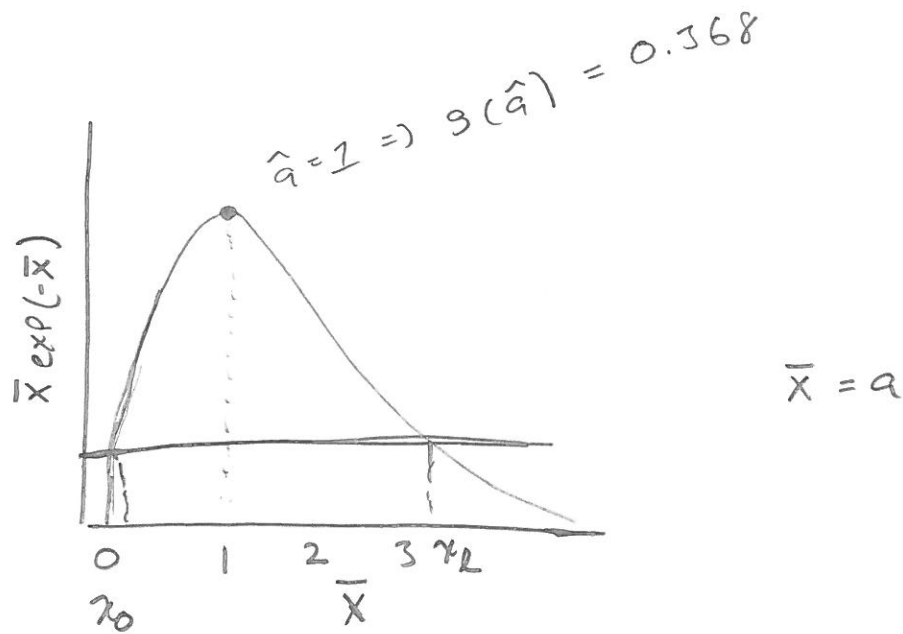
$$g'(a) = \exp(-a)(1-a)$$

$g'(a)$ is positive for $a \in (0, 1)$

$g'(a)$ is negative for $a > 1$

$\therefore g(a)$ is strictly increasing from $(0, 1)$
and strictly decreasing from $(1, \infty)$.

\therefore for $g(a)$ we can find two cut-points:



\therefore There are exactly two solutions for $g(\bar{x}) = C$.

b.) Based on N-P theory for this LRT
we determine c based on

$$R = \{ \bar{X} \exp(-\bar{X}) \leq c \}$$

$$P_{H_0}(R) = \alpha = 0.05$$

$$P_{H_0}(\bar{X} \exp(-\bar{X}) \leq c) = 0.05$$

\rightarrow under $H_0: \theta = 1$.

c.) Note $X \sim \exp(\theta)$

$$X \sim \text{gamma}(1, \theta) \stackrel{\text{Under } H_0}{\Rightarrow} X \sim \text{gamma}(1, 1)$$

$$f(x) = \frac{\beta^\alpha}{\Gamma(\alpha)} x^{\alpha-1} \exp(-x\beta)$$

$$= \frac{\beta^1}{\Gamma(1)} x^{1-1} \exp(-x\beta) = \beta \exp(-x\beta)$$

$$\Rightarrow MGF_x(t) = \left(\frac{1}{1-t/\beta} \right)^\alpha = \left(\frac{\beta}{\beta-t} \right)^\alpha$$

\uparrow
worked out in
lecture.

$$\therefore Y = \sum_{i=1}^n X_i$$

$$\begin{aligned} MGF_{\sum X_i}(t) &= \left(\frac{1}{1-t/\beta} \right)^{\alpha_1} \times \dots \times \left(\frac{1}{1-t/\beta} \right)^{\alpha_n} \\ &= \left(\frac{1}{1-t/\beta} \right)^{\alpha_1 + \dots + \alpha_n} \end{aligned}$$

$$\therefore Y \sim \text{gamma}(\alpha_1 + \dots + \alpha_n, \beta)$$

under H_0

$$\Rightarrow \text{gamma}(n, 1)$$

$$\therefore \bar{X} \sim \text{gamma}(n, n)$$

$$\Rightarrow P_{H_0}(\bar{X} \exp(-\bar{X}) \leq c) = \alpha$$

$$P_{H_0}(\bar{X} \in [0, x_0(c)] \cup [x_1(c), \infty))$$

$$= F(x_0(c)) + 1 - F(x_1(c))$$

Solve computationally:

- 1.) try a value c
- 2.) calculate x_0, x_1
- 3.) calculate $F(x_0(c)) + 1 - F(x_1(c))$
and get this close to $\alpha = 0.05$

d.) How can we solve via simulation?

$$R = \{ \bar{x} \exp(-\bar{x}) < c \}$$

1.) $x_1, \dots, x_{10} \sim \exp(\theta_0 = 1)$

2.) Calculate \bar{x}

3.) Calculate $\bar{x} \exp(-\bar{x})$

4.) We reject for small values:

$$P_{H_0}(\bar{x} \exp(-\bar{x}) < c) = 0.05$$

\therefore We use the empirical quantile for 0.05.

\rightarrow See the R portion for the solution.

$$9.24.) \quad X \sim \text{binomial}(n, p)$$

$$9.) \quad H_0: p = 0.5$$

$$H_1: p \neq 0.5$$

$$\lambda = \frac{\sup_{p_0} L(p|x)}{\sup_p L(p|x)} = \frac{\binom{n}{x} p_0^x (1-p_0)^{n-x}}{\binom{n}{x} \hat{p}^x (1-\hat{p})^{n-x}}$$

$$\bullet \text{ We can work out } \hat{p} = \bar{x}$$

$$\therefore \lambda = \frac{\left(\frac{1}{2}\right)^x \left(1 - \frac{1}{2}\right)^{n-x}}{\left(\frac{x}{n}\right)^x \left(1 - \frac{x}{n}\right)^{n-x}} = \frac{\left(\frac{1}{2}\right)^n}{\left(\frac{x}{n}\right)^x \left(1 - \frac{x}{n}\right)^{n-x}}$$

$$\frac{n^n}{n^n} = \frac{n^n}{n^x n^{n-x}}$$

$$\Rightarrow \lambda = \frac{n^n \left(\frac{1}{2}\right)^n}{n^x \left(\frac{x}{n}\right)^x n^{n-x} \left(1 - \frac{x}{n}\right)^{n-x}} = \frac{\left(n/2\right)^n}{x^x (n-x)^{n-x}}$$

$$\therefore R = \left\{ \frac{\left(n/2\right)^n}{x^x (n-x)^{n-x}} < c \right\}$$

b.) We want to show that the LRT rejects for

$$R = \{ |x - n/2| > k \}$$

• We have $\lambda = \frac{(n/2)^n}{x^x (n-x)^{n-x}}$

Let $h(x) = \frac{1}{x^x (n-x)^{n-x}}$

Also notice $h(n-x) = \frac{1}{(n-x)^{n-x} (n-(n-x))^{n-(n-x)}}$

$$= \frac{1}{(n-x)^{n-x} x^x}$$

$\therefore h(x) = h(n-x)$

• So consider $y = x - n/2 \Rightarrow x = n/2 + y$

$$\Rightarrow h(n/2 + y) = h(n/2 - y)$$

• Consider $h(n/2 + y)$ for $y \geq 0$.

$$\bullet \quad h(n/2 + y) = \binom{n/2 - y}{- (n/2 - y)} \binom{n/2 + y}{- (n/2 + y)}$$

$$\bullet \quad \text{Let } h^*(y) = \log(h(n/2 + y))$$

\bullet We want to show that $h^*(y)$ is

a non-decreasing function of y

$$\Rightarrow y \uparrow \Rightarrow x - n/2 \uparrow \Rightarrow h^*(y) \downarrow \Rightarrow \lambda \downarrow$$

↓

due to symmetry

$$x + n/2 \uparrow \Rightarrow \lambda \downarrow$$

$$h^*(y) = - (n/2 - y) \log(n/2 - y) - (n/2 + y) \log(n/2 + y)$$

$$\frac{d h^*(y)}{d y} = \log(n/2 - y) - \log(n/2 + y) \leq 0$$

as $y \geq 0$

$\therefore R = \{ |x - n/2| > K \}$ is an equivalent rejection region for some K .

$$c.) \quad R = \{ |X - n/2| > K \}$$

$$P_{H_0} (|X - n/2| > K) = P_{H_0} (X - n/2 > K)$$

$$+ P (X - n/2 < -K) = \alpha$$

$$= P_{H_0} (X > K + n/2) + P_{H_0} (X < -K + n/2)$$

$$= 1 - P_{H_0} (X \leq K + n/2) + P_{H_0} (X < -K + n/2) = \alpha$$

d.) If $n = 10$, $K = 2$ what is α :

$$\alpha = 1 - P_{H_0} (X \leq 2 + 10/2) + P_{H_0} (X < 10/2 - 2)$$

$$= 1 - P_{H_0} (X \leq 7) + P_{H_0} (X < 3)$$

$$= 1 - P_{\text{binom}}(10, 7, 1/2) + P_{\text{binom}}(10, 2, 1/2)$$

$$\approx 0.11.$$

e.) $X \overset{\text{under } H_0}{\sim} \text{Binomial} (n=100, p=1/2)$

$$\alpha = 1 - (P_{H_0}(X \leq k + n/2) + P_{H_0}(X < -k + n/2))$$

$$= 1 - P_{H_0}(X \leq 60) + P_{H_0}(X < 40)$$

Using CLT: $E(X) = 100(1/2) = 50$; $V(X) = 100(1/2)^2 = 25$

$$= 1 - P_{H_0}\left(\frac{X-50}{5} \leq \frac{60-50}{5}\right) + P_{H_0}\left(\frac{X-50}{5} \leq \frac{40-50}{5}\right)$$

$$= 1 - P_{H_0}(Z \leq 2) + P_{H_0}(Z \leq -2)$$

$$= 1 - P_{\text{norm}}(2) + P_{\text{norm}}(-2)$$

$$= 0.0455$$

$$32.) \quad \text{If } A \Rightarrow X \sim N(\mu=100, \sigma=25)$$

$$B \Rightarrow X \sim N(\mu=125, \sigma=25)$$

A datapoint is drawn: $x = 120$.

$$\begin{array}{l|l} a.) & \begin{array}{l} H_0 : \mu = 100 \\ H_1 : \mu = 125 \end{array} \end{array} \quad \left| \quad \begin{array}{l} \text{Simple vs Simple} \end{array} \right.$$

$$\begin{aligned} \lambda &= \frac{\exp\left(-\frac{1}{2(25^2)}(x-100)^2\right)}{\exp\left(-\frac{1}{2(25^2)}(x-125)^2\right)} \\ &= \exp\left(-\frac{1}{2(25^2)}\left[(x-100)^2 - (x-125)^2\right]\right) \end{aligned}$$

• when $x = 120 \Rightarrow \lambda = 0.7408$

$$\begin{aligned} b.) \quad \frac{P(A|x)}{P(B|x)} &= \frac{P(x|A)P(A)}{P(x|B)P(B)} = \frac{P(x|A)}{P(x|B)} = 0.7408 \\ &= \frac{1 - P(B|x)}{P(B|x)} = 0.7408 \end{aligned}$$

$$\therefore P(B|x) = 0.574$$

$$c.) R = \{ x > 125 \}$$

$$\begin{aligned} P_{H_0}(R) &= P_{H_0} \left(\frac{x - 100}{25} > \frac{125 - 100}{25} \right) \\ &= P_{H_0}(Z > 1) = 1 - P_{H_0}(Z \leq 1) \\ &= 0.1587 \end{aligned}$$

$$\begin{aligned} d.) \text{Power} &= P_{H_1}(R) = P_{H_1} \left(Z > \frac{125 - 125}{25} \right) \\ &= P_H(Z > 0) = 1 - P_{H_0}(Z \leq 0) \\ &= \frac{1}{2} \end{aligned}$$

e.) let's calculate the p-value:

$$\begin{aligned} P_{H_0}(X > 120) &= P_{H_0} \left(Z > \frac{120 - 100}{25} \right) \\ &= P_{H_0}(Z > 0.8) = 1 - P_{H_0}(Z \leq 0.8) \\ &= 1 - \text{Pnorm}(0.8) \\ &\approx 0.2119. \end{aligned}$$