

$$8.4) \quad \left. \begin{aligned} [A(X-c)]^2 &= [A(X) - A(c)]^2 \\ &= [A(X) - c]^2 \\ &= [A(X)]^2 - 2cA(X) + c^2 \end{aligned} \right\} \begin{array}{l} \text{ii and iii pg 7} \\ \text{combined} \end{array}$$

$$8.5) \quad \text{Let } \mathcal{C} = \{k \text{ such that } p_k > 0\} \quad X(\omega_k) = c, k \in \mathcal{C}$$

" \Rightarrow "

$$A(X) = \sum_k p_k X(\omega_k) = \sum_{k \in \mathcal{C}} p_k c = c \sum_{k \in \mathcal{C}} p_k = c$$

(Assuming $X(\omega_k)$ is real valued $\forall k$, ie $X(\omega_k) < \infty$)

$$A(X^2) = \sum_k p_k X(\omega_k)^2 = \sum_{k \in \mathcal{C}} p_k c^2 = c^2 \sum_{k \in \mathcal{C}} p_k = c^2$$

$$\text{Therefore } [A(X)]^2 = A(X^2)$$

" \Leftarrow "

$$[A(X)]^2 = A(X^2) \Rightarrow$$

$$0 = A[(X - A(X))^2] = \sum_k p_k (X(\omega_k) - \sum_k p_k X(\omega_k))^2$$

$$\text{if } p_k > 0 \text{ then } X(\omega_k) - \sum_k p_k X(\omega_k) = 0$$

$$\text{that is, } X(\omega_k) = \underbrace{\sum_k p_k X(\omega_k)}_{\text{constant}} = c, k \in \mathcal{C}$$

$$16.4) \quad |x_1 + x_2| \leq |x_1| + |x_2| \quad (\text{triangle inequality})$$

$$\Rightarrow |x_1| + |x_2| - |x_1 + x_2| \geq 0$$

$$\Rightarrow E(|x_1| + |x_2| - |x_1 + x_2|) \geq 0 \quad \text{Axiom 1 pg 15}$$

$$\Rightarrow E(|x_1|) + E(|x_2|) \geq E(|x_1 + x_2|) \quad \text{Axiom 3 pg 15}$$

□

$$16.5) \quad |x_n - x| \leq Y_n$$

$$\Rightarrow -Y_n \leq x_n - x \leq Y_n$$

$$\Rightarrow -E(Y_n) \leq E(x_n) - E(x) \leq E(Y_n) \quad (\text{positive linear operator})$$

$$\Rightarrow |E(x_n) - E(x)| \leq E(Y_n) \quad \forall n \quad (*)$$

Since $E(Y_n) \rightarrow 0$ as n increases, we have $\forall n, \exists \varepsilon > 0$ such that $|E(Y_n) - 0| < \varepsilon$

Therefore from (*), $\forall n \exists \varepsilon > 0$ such that

$$|E(x_n) - E(x)| \leq E(Y_n) \leq |E(Y_n)| < \varepsilon$$

which means $E(x_n) \rightarrow E(x)$ with n increasing

20.1) Let us show that $f(t) \geq 0$:

Assume $f(t) < 0$, we can take $H(t) > 0$, $t > 0$

and $H(0) = 0$ ($H(t) \geq 0$)

Then $E[H(\tau)] = \int_0^{\infty} \underbrace{H(t)}_{\geq 0} \underbrace{f(t)}_{< 0} dt < 0$ (Conflict with axiom 1)
pg 15

Let us show $p \geq 0$

Take $H(t) = \begin{cases} 1 & t = 0 \\ 0 & \text{otherwise} \end{cases}$ (Note $H(t) \geq 0$)

Thus $E[H(\tau)] = p \geq 0$ (Axiom 1)

$$\boxed{p + \int_0^{\infty} f(t) dt = 1}$$

$$E(1) = p \cdot 1 + \int_0^{\infty} f(t) dt = 1 \quad (\text{Axiom 4})$$

$$21.3) \quad E(X) = \lim_{D \rightarrow \infty} \frac{1}{2D} \int_{-D}^D X(\omega) d\omega$$

Axiom 1: $X \geq 0 \Rightarrow$

$$0 = \int_{-D}^D 0 d\omega \leq \int_{-D}^D X(\omega) d\omega$$

$$\Rightarrow E(X) = \lim_{D \rightarrow \infty} \frac{1}{2D} \int_{-D}^D X(\omega) d\omega \geq 0$$

Axiom 2:

$$\begin{aligned} E(cX) &= \lim_{D \rightarrow \infty} \frac{1}{2D} \int_{-D}^D cX(\omega) d\omega \\ &= c \lim_{D \rightarrow \infty} \frac{1}{2D} \int_{-D}^D X(\omega) d\omega = c E(X) \end{aligned}$$

Axiom 3:

$$\begin{aligned} E(X_1 + X_2) &= \lim_{D \rightarrow \infty} \frac{1}{2D} \int_{-D}^D X_1(\omega) + X_2(\omega) d\omega \\ &= \lim_{D \rightarrow \infty} \frac{1}{2D} \left[\int_{-D}^D X_1(\omega) d\omega + \int_{-D}^D X_2(\omega) d\omega \right] \\ &= \lim_{D \rightarrow \infty} \frac{1}{2D} \int_{-D}^D X_1(\omega) d\omega + \lim_{D \rightarrow \infty} \frac{1}{2D} \int_{-D}^D X_2(\omega) d\omega = E(X_1) + E(X_2) \end{aligned}$$

Axiom 4:

$$\begin{aligned} E(1) &= \lim_{D \rightarrow \infty} \frac{1}{2D} \int_{-D}^D 1 d\omega = \lim_{D \rightarrow \infty} \frac{1}{2D} \omega \Big|_{-D}^D \\ &= \lim_{D \rightarrow \infty} \frac{1}{2D} [D - (-D)] = \lim_{D \rightarrow \infty} \frac{2D}{2D} = \lim_{D \rightarrow \infty} 1 = 1 \end{aligned}$$

Axiom 5: Using $X_n(\omega) = \begin{cases} 1 & |\omega| \leq n \\ 0 & \text{otherwise} \end{cases}$

We have $X_1 \leq X_2 \leq X_3 \leq \dots$

and $\lim_{n \rightarrow \infty} X_n(\omega) = X = 1$ (increases monotonically)

But when $D > n$ we have

$$\lim_{n \rightarrow \infty} E(X_n(\omega)) = \lim_{n \rightarrow \infty} \lim_{D \rightarrow \infty} \frac{1}{2D} \int_{-D}^D X_n(\omega) d\omega$$

$$= \lim_{n \rightarrow \infty} \lim_{D \rightarrow \infty} \frac{1}{2D} \omega \Big|_{-n}^n$$

$$= \lim_{n \rightarrow \infty} \lim_{D \rightarrow \infty} \frac{n}{D}$$

$$= \lim_{n \rightarrow \infty} 0 = 0 \neq E(X) = 1$$

24.1) The loss function can be written as:

$$g(t) = \begin{cases} a(t-T) & \text{if } T < t \\ b(T-t) & \text{if } T > t \end{cases}$$

Or $g(t) = a(t-T) \mathbb{I}(T < t) + b(T-t) \mathbb{I}(T > t)$

The expected loss become:

$$\begin{aligned} E[g(t)] &= a(t-T) P(T < t) + b(T-t) P(T > t) \\ &= a(t-T) P(T < t) - b(t-T) P(T > t) \\ &= (t-T) [a P(T < t) - b P(T > t)] \end{aligned}$$

Taking the derivative with respect to t and set to 0 we have

$$a P(T < t) - b P(T > t) = 0$$

$$\Rightarrow a P(T < t) = b P(T > t) \quad (P(T > t) = 1 - P(T < t))$$

or $a P(T < t) = b - b P(T < t)$

$$\Rightarrow P(T < t) = \frac{b}{a+b}$$

32.7) The relation $\hat{X} - E(X) = a^T(Y - E(Y))$

become $\hat{Y}_{t+s} - \mu = \beta^s (Y_t - \mu)$

By Theorem 2.8.2 \hat{Y}_{t+s} is an LLS estimate

of Y_{t+s} if and only if a statistics the

linear relations $V_{Y_t Y_t} a = V_{Y_t Y_{t+s}} \quad \text{(we have to find)}$
 a

$$V_{Y_t Y_t} = \begin{bmatrix} \text{Cov}(Y_1, Y_1) & \dots & \text{Cov}(Y_1, Y_t) \\ \vdots & & \vdots \\ \text{Cov}(Y_t, Y_1) & \dots & \text{Cov}(Y_t, Y_t) \end{bmatrix} = \alpha \begin{bmatrix} 1 & \dots & \beta^{11-t1} \\ \vdots & & \vdots \\ \beta^{1t-11} & & 1 \end{bmatrix}$$

$$V_{Y_t Y_{t+s}} = \begin{bmatrix} \text{Cov}(Y_{t+s}, Y_1) \\ \vdots \\ \text{Cov}(Y_{t+s}, Y_t) \end{bmatrix} = \alpha \begin{bmatrix} \beta^{1t+s-11} \\ \vdots \\ \beta^{1t+s-t1} \end{bmatrix}$$

Set $a = \begin{bmatrix} 0 \\ \vdots \\ 0 \\ \beta^s \end{bmatrix}$ and we have $V_{Y_t Y_t} a = V_{Y_t Y_{t+s}}$

Thus $\hat{Y}_{s+t} = \mu + \beta^s (Y_t - \mu)$ (scalar form) is
 the L.L.S prediction

37.7) Let $X_1 = X - E(X)$ and $X_2 = Y - E(Y)$

$$X = \begin{bmatrix} X_1 \\ X_2 \end{bmatrix}$$

$$XX^T = \begin{bmatrix} X_1^2 & X_1 X_2 \\ X_1 X_2 & X_2^2 \end{bmatrix} \text{ symmetric}$$

$$\text{Then } E(XX^T) = \begin{bmatrix} \text{Var}(X) & \text{Cor}(X, Y) \\ \text{Cor}(X, Y) & \text{Var}(Y) \end{bmatrix} \text{ symmetric}$$

$$\begin{aligned} c^T E(XX^T) c &= E(c^T XX^T c) = E(c_1^2 X_1^2 + 2c_1 c_2 X_1 X_2 + c_2^2 X_2^2) \\ c &= \begin{pmatrix} c_1 \\ c_2 \end{pmatrix} \\ &= E(c_1 X_1 + c_2 X_2)^2 \\ &= E[(c^T X)^2] \geq 0 \end{aligned}$$

Therefore $E(XX^T)$ is positive definite matrix

By Theorem 2.91 $|E(X_1 X_2)|^2 \leq E(X_1^2)E(X_2^2)$ that is,

$$[\text{Cor}(X, Y)]^2 \leq \text{Var}(X)\text{Var}(Y), \text{ with equality if}$$

there is a non-trivial relation $c_1 X_1 + c_2 X_2 \stackrel{\text{m.s.}}{=} 0$, i.e.,

$$c_1 (X - E(X)) + c_2 (Y - E(Y)) \stackrel{\text{m.s.}}{=} 0 \text{ or}$$

$$c_1 X + c_2 Y \stackrel{\text{m.s.}}{=} c_1 E(X) + c_2 E(Y) = c_0$$