Multivariate Analysis of Variance (MANOVA) Model: Xij=Ni+ Eij i=1,..., k (treatments) ξή~ Nρ(0, C) => Xii~ Np(M,C) Test Ho: $\mu_1 = \mu_2$ $S_T = \sum_{i=1}^{k} \sum_{j=1}^{n} (\chi_{ij} - \overline{\chi} \chi \chi_{ij} - \overline{\chi})^T$ $=\sum_{i=1}^{k} n_i (\overline{x_i} - \overline{x}) (\overline{x_i} - \overline{x})^{\mathsf{T}} + \sum_{i=1}^{k} \sum_{j=1}^{n} (\overline{x_j} - \overline{x_j})^{\mathsf{T}}$ Idea: If the ,..., the are not all equal (i.e. Ho is false) then SB should be much bigger than Sw. Sol'n: Look at eigenvalues of Sulse i.e. Test statistics should be based on these eigenvalues. Why? Model $X_{ij} = \mu_i + \xi_{ij}$ $Y_{ij} = A_{X_{ij}} + b = A_{\mu_i} + b + A_{\xi_{ij}} = \mu_i^* + \xi_{ij}^*$ $\psi_{invertible}$ $\psi_{i} = \psi_{k} \iff \psi_{i}^* = \dots = \psi_{k}^*$ Look at $S_B \notin S_W$ based on $[X_{ij}] \notin [Y_{ij}]$ $S_B^* = \sum_{i=1}^k n(\overline{X}_i - \overline{X})(\overline{X}_i - \overline{X})^T$ S&=ASBAT $S_{u}^{\mathbf{Y}} = A S_{u}^{\mathbf{X}} A^{\mathsf{T}}$ Therefore, $(S_w^y)^T S_b^y = (A S_w^x A^T)^T (A S_b^x A^T) = (A^T)^T (S_w^x)^T A^T A S_b^x A^T = (A^T)^T (S_w^x)^T S_b^x A^T vs. (S_w^x)^T S_b^x$ Ideally, conclusions based on (SW) SB should be the same as those based on non-trivial Need to build function ϕ such that $\phi((S_w^y)^T S_b^y) = \phi((S_w^y)^T S_b^y)$ for any invertible matrix A Look at eigenvalues (AT)-1(SW)-1SBATX= XV

 $(S_{w}^{x})^{-1}S_{B}^{x} \underbrace{A^{T}x = \lambda A^{T}x}_{Y^{*}} \underbrace{A^{T}x = \lambda A^{T}x}_{X^{*}} \underbrace{(S_{w}^{x})^{-1}S_{B}^{x}}_{X^{*}} \underbrace{(S_{w}^{x})^{-1}S_{B}^{x}}_{X^{*}} \underbrace{A^{T}x = \lambda A^{T}x}_{X^{*}} \underbrace$

Thus the eigenvalues of (Sw) - SB are the same as those of (Sw) - SB for any A.

Therefore to test $H_0: \mathcal{K} = \dots = \mathcal{K}_k$, look at lest statistics based on eigenalus $\lambda_1, \dots, \lambda_p$ of $S_{W} S_{B}$. ($\lambda_1 \ge \dots \ge \lambda_p \ge 0$)

```
Need to be very careful!
Test statistics
- reject A for A≤K.
2 Pillai's Trace
Tp = ∑i=1 1+ li (reject Ho for Tp > tx)
3 Lawley- Hotelling Trace
TH= > P A; (reject Ho for TH > Kx)
(1) Roy's maximal root
R = \max(\lambda_1, \dots, \lambda_p) = \lambda_i \text{ (reject for } R > k_N)
What are the null distins of these statistics?
      - in general, difficult to derive exact dist'n.
      -approximations SForx2 distin
                              Normal approximations
     - in certain cases, A can be transformed to an exact = distin.
      P K Transformation degree of freedom
      any 2 \left(\frac{1-\Lambda}{\Lambda}\right)\left(\frac{n-p-1}{p}\right) p_n-p-1
      any 3 (\frac{1-\Lambda^{\frac{1}{2}}}{\Lambda^{\frac{1}{2}}})(\frac{n-p-2}{p}) 2p,2(n-p-2)
        2 any (\frac{1-\lambda^2}{\lambda^2})(\frac{n-k-1}{k-1}) 2(k-1), 2(n-k-1)
 Example: Group: male & female painted turtles (k=2)

p=3 variables

· length

width n_1=n_2=24

height
 S_{w}^{-1}S_{B} = \begin{pmatrix} -1.61 & -1.02 & -0.81 \\ -0.81 & -0.51 & -0.40 \\ 7.39 & 4.66 & 3.70 \end{pmatrix} -Sw<sup>-1</sup>S<sub>B</sub> has only 1 non-zero eigenvalue \lambda_{1} = 1.574 \lambda_{2} = \lambda_{3} = 0
R = 1.574 \quad T_{LH} = \lambda_1 + \lambda_2 + \lambda_3 = 1.574
T_P = \frac{\lambda_1}{1 + \lambda_1} = 0.611 \qquad \Lambda = \frac{1}{1 + \lambda_1} = 0.389
Conversion to F-statistic
Hostellings P - \text{value} = P[F(3.44) \ge 23.07] = 4.06 \times 10^{-9}

T^2 statistic (K = 2)
```