# STA304/1003 H1 F - Summer 2014: Surveys, Sampling, and Observational Data

Lecture 9 - Part I: Two-Stage Cluster Sampling (con'd)

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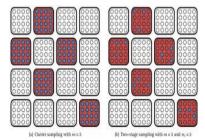
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# Two-Stage Cluster Sampling

In One-Stage Cluster sampling, all ssus in the selected psus are selected. In Two-Stage Cluster sampling:

- 1. Select an SRS S of n psus from the population of N psus.
- 2. Select an SRS of  $m_i$  ssus from each sampled psu i
- $\rightarrow$  2 sources of variability: from selecting psus and selecting ssus (both stages)

Diagram: One-Stage vs. Two-Stage Cluster Samples:



## **Review of Notation**

### Population Quantities at psu level:

- ▶ *N* = number of psus in the population
- ▶  $M_i$  = number of ssus in psu i, i = 1, 2, ..., N
- $M = \sum_{i=1}^{N} M_i$  = total number of ssus in the population
- $\overline{M} = \overline{M}/N$  = average cluster size for the population
- $ightharpoonup y_{ii} = \text{measurement for } j \text{th element in psu } i$
- $au_i = \sum_{i=1}^{M_i} y_{ij} = \text{total in psu } i$
- $au = \sum_{i=1}^{\hat{N}} \tau_i = \sum_{i=1}^{N} \sum_{j=1}^{M_i} y_{ij} = \text{population total}$
- $S_t^2 = \frac{1}{N-1} \sum_{i=1}^{N} (\tau_i \frac{\tau}{N})^2$  = population variance of the psu totals

#### Population Quantities at ssu level:

- $\bar{y}_U = \frac{1}{M} \sum_{i=1}^{N} \sum_{j=1}^{M_i} y_{ij} = \text{population mean}$
- $ightharpoonup ar{y}_{iU} = rac{1}{M_i} \sum_{j=1}^{M_i} y_{ij} = rac{\tau_i}{M_i}$  population mean in psu i
- $S^2 = \frac{1}{M-1} \sum_{i=1}^N \sum_{j=1}^{M_i} (y_{ij} \bar{y}_U)^2 = \text{population variance (per ssu)}$
- ►  $S_i^2 = \frac{1}{M_i 1} \sum_{j=1}^{M_i} (y_{ij} \bar{y}_{iU})^2$  = population variance within psu i

## Sample Quantities

- ightharpoonup n = number of psus in the sample
- $ightharpoonup m_i = \text{number of ssus in the sample from psu } i$
- ▶ S: sample of psus
- $\triangleright$   $S_i$ : sample of  $m_i$  ssus from ith psu
- $ightharpoonup ar{y}_i = rac{1}{m_i} \sum_{j \in \mathcal{S}_i} y_{ij} = \text{sample mean for psu } i$
- $\hat{\tau}_i = \sum_{i \in S_i} \frac{M_i}{m_i} y_{ij} = M_i \bar{y}_i = \text{estimated total for psu } i$
- $s_i^2 = \frac{1}{m_i 1} \sum_{j \in S_i} (y_{ij} \bar{y}_i)^2 = \text{sample variance within psu } i$

# Estimating the Population Mean

## 1. *M* is known:

$$\hat{\bar{y}}_{unb} = \frac{N}{M} \sum_{i \in \mathcal{S}} \frac{M_i \bar{y}_i}{n} = \frac{\hat{\tau}_{unb}}{M}$$
 is an unbiased estimator of the population mean

$$\blacktriangleright E(\hat{\bar{y}}_{unb}) = \bar{y}_U$$

$$\hat{V}(\hat{\bar{y}}_{unb}) = \frac{1}{n\overline{M}^2} \left( 1 - \frac{n}{N} \right) s_b^2 + \frac{1}{nN\overline{M}^2} \sum_{i \in \mathcal{S}} \left( 1 - \frac{m_i}{M_i} \right) M_i^2 \frac{s_i^2}{m_i} ;$$
 where 
$$s_b^2 = \frac{1}{n-1} \sum_{i \in \mathcal{S}} (M_i \bar{y}_i - \overline{M} \hat{\bar{y}}_{unb})^2 \text{ is the sample variance}$$

2. *M* is unknown. Use Ratio Estimation:

$$\hat{\bar{y}}_r = \frac{\sum_{i \in \mathcal{S}} \hat{\tau}_i}{\sum_{i \in \mathcal{S}} M_i} = \frac{\sum_{i \in \mathcal{S}} M_i \bar{y}_i}{\sum_{i \in \mathcal{S}} M_i}$$

among the  $M_i \bar{\nu}_i$  terms.

$$\hat{V}(\hat{\bar{y}}_r) = \frac{1}{n\overline{M}^2} \left(1 - \frac{n}{N}\right) s_r^2 + \frac{1}{nN\overline{M}^2} \sum_{i \in \mathcal{S}} M_i^2 \left(1 - \frac{m_i}{M_i}\right) \frac{s_i^2}{m_i}$$

When N is large, the second term is negligible compared to first.

Recall: 
$$s_r^2 = \frac{1}{n-1} \sum_{i \in S} (M_i \bar{y}_i - M_i \hat{y}_r)^2$$

# **Estimating the Population Total**

## **Unbiased Estimation:**

$$\hat{\tau}_{unb} = \frac{N}{n} \sum_{i \in \mathcal{S}} \hat{\tau}_i = \frac{N}{n} \sum_{i \in \mathcal{S}} M_i \bar{y}_i = \frac{N}{n} \sum_{i \in \mathcal{S}} \sum_{j \in \mathcal{S}_i} \frac{M_i}{m_i} y_{ij}$$
 is an unbiased estimator of population total

 $\hat{\tau}_i$ 's are random variables so  $\hat{\tau}_{unb}$  has 2 sources of variability:

- (1) variability between psus
- (2) variability of ssus within psus

## Properties of $\hat{\tau}_{unb}$ :

- $ightharpoonup E(\hat{ au}_{unb}) = au$
- $\hat{V}(\hat{\tau}_{unb}) = \frac{N^2}{n} \left(1 \frac{n}{N}\right) s_b^2 + \frac{N}{n} \sum_{i \in \mathcal{S}} \left(1 \frac{m_i}{M_i}\right) M_i^2 \frac{s_i^2}{m_i}$   $\hookrightarrow$  Variance from one-stage cluster + additional variance due to selection of ssus within psus

# **Design Issues**

#### 1. Precision Needed:

▶ Determine ME, e

#### 2. Choosing the psu size:

- Mostly natural like clutches of eggs, classes with students, etc. Sometimes have choice such as area of forest, time interval between costumers.
- More area ⇒ more variability within psus ⇒ ICC smaller

#### 3. Choosing subsampling sizes (how many ssus to sample in each psu):

Assuming equal cluster sizes,  $\overline{M}$  and take equal sample sizes m - minimize variance for fixed cost

$$V(\hat{\bar{y}}_{unb}) = \left(1 - \frac{n}{N}\right) \frac{MSB}{n\overline{M}} + \left(1 - \frac{m}{\overline{M}}\right) \frac{MSW}{nm}$$
:

If MSW=0,  $R_a^2=1$ : choose m=1. For other values, depends on relative costs.

• total cost = 
$$C = c_1 n + c_2 nm$$
:

$$lacksquare$$
  $n_{opt}=rac{C}{c_1+c_2m_{opt}}$  and  $m_{opt}=\sqrt{rac{c_1M(N-1)(1-R_a^2)}{c_2(NM-1)R_a^2}}$ :

Estimate  $R_a^2$  from pilot survey:  $\hat{R}_a^2 = 1 - \frac{\widehat{MSW}}{\hat{S}^2}$  and for large populations

$$m_{opt} = \sqrt{c_1(1-\hat{R}_a^2)/c_2\hat{R}_a^2}$$

For unequal cluster size use  $\bar{M}$  instead of M to determine  $\bar{m}$ : sample  $\bar{m}$  in each psu or allocate so that  $\frac{m}{M}$  is constant

## 4. Choosing the Sample Size (number of psus, *n*):

- Determine psu size and subsampling fraction. Decide on desired ME, e
- For equal-sized clusters:

$$V(\hat{y}) \leq \frac{1}{n} \left[ \frac{MSB}{\overline{M}} + \left( 1 - \frac{m}{\overline{M}} \right) \frac{MSW}{m} \right] = \frac{v}{n}$$

- $n = z_{\alpha/2}^2 v/e^2$
- ► Estimate  $v = \left[\frac{MSB}{\overline{M}} + \left(1 \frac{m}{\overline{M}}\right) \frac{MSW}{m}\right]$  from previous survey or prior knowledge

## 5. Iterate:

- Above gives the n for required ME
- Modify survey design (add stratification, auxiliary variables, etc.) until cost is within budget.

# **Example: Creamed Corn**

cases?

Dou=case ssu =can N = 580 cases Mi=24 for all i, M=24 > N Mi=580(24)=total cans in truck=13920

An inspector samples cans from a truckload of canned creamed corn to estimate the average number of worm fragments per can. The truck has 580 cases; each case contains 24 cans. It takes 20 minutes to locate and open a case, and 8 minutes to locate and examine each specified can within a case. Assume your budget is 120 minutes. A preliminary study of 12 cases at NSW = NS residuals = 4.53 your budget is 120 minutes. A preliminary study of 1  $S^2 = \frac{SSTO}{NM-1} = \frac{(N-1)NSP}{NM-1} + N(M-1)NSW}$  NM-1 OT: 1.57 OT: 1.57 OT: 1.57 OT: 1.57

 $= \frac{579(13.60) + 580(23)(453)}{13.919} = 4.91$   $R_a^2 = [-\frac{4.53}{4.91} = 0.0774$  $C_1=20$ C2=8

$$= \frac{579(13.60) + 580(23)(4.53)}{13.919} = 4.91 \qquad C2: 424 \qquad C8: 302 \qquad = 545 - 6 cans \\ | C_1 - M_{SL}| = | -4.53| = 0.0774 \qquad C3: 012 \qquad C9: 735 \qquad | C_1 - C_1 - C_2 - C_2 - C_3 - C_$$

total cost =  $(2 \text{ cases } \times 20) + 8 \times 6 \times 2) = |36$ "over budget" mins So sample 2 cases x 5 cans

# Using 'R' to get ANOVA Table:

```
> case=rep(seq(1,12,1),each=3)
> case
[1] 1 1 1 2 2 2 3 3 3 4 4 4 5 5 5 6 6 6
     7 7 7 8 8 8 9 9 9 10 10 10 11 11 11 12 12 12
> case=factor(case)
> case
[1] 1 1 1 2 2 2 3 3 3 4 4 4 5 5 5 6 6 6
    7 7 7 8 8 8 9 9 9 10 10 10 11 11 11 12 12 12
Levels: 1 2 3 4 5 6 7 8 9 10 11 12
> frag=c(1,5,7,4,2,4,0,1,2,3,6,6,4,9,8,0,7,3,5,5,1,3,0,2,7,3,5,3,1,4,4,7,9,0,0,0)
> frag
[1] 1 5 7 4 2 4 0 1 2 3 6 6 4 9 8 0 7 3 5 5 1 3 0 2 7 3 5 3 1 4 4 7 9 0 0 0
> model <- lm(frag ~ case)
> anova(model)
Analysis of Variance Table
Response: frag
         Df Sum Sq Mean Sq F value Pr(>F)
         11 149.64 13.6035 3.0045 0.01172 *
case
Residuals 24 108.67 4.5278
Signif. codes: 0 *** 0.001 ** 0.01 * 0.05 . 0.1 1
```

# Summary and Advantages/Disadvantages of Cluster Sampling

- Cluster sampling used commonly in large surveys
- Convenient, easy to access elements by clusters since clusters occur naturally together
- In cluster sampling, want elements to be heterogenous within groups; In STRS, want elements to be homogeneous within groups (opposites)
- If elements within clusters are homogenous, two-stage cluster sampling is better
- ► One-Stage is a special case of the general Two-Stage Cluster sample (using  $M_i = m_i$ )
- Cluster sampling usually has larger variance than using SRS for the same sample size
- Cluster sampling can give more precision per dollar if measuring individual elements is much more costly than sampling clusters
- ► Two types of estimation for population parameters: Unbiased and Ratio estimation
  - If cluster sizes vary greatly, ratio estimation is better to use (smaller variance) and may be an advantage to sample with probabilities proportional to cluster size
  - ► For equal cluster sizes, both types of estimates are equivalent