

## Classification

Observe  $(g_1, x_1), \dots, (g_n, x_n)$

$g_i$  takes values in  $\{1, \dots, k\}$

Model: Dist'n of  $(G, X)$ :  $P(G=j) = \lambda_j$

Conditional density of  $X$  given  $G=j$ :  $f_j(x)$

Marginal density of  $X$  is  $f(x) = \lambda_1 f_1(x) + \lambda_2 f_2(x) + \dots + \lambda_k f_k(x)$

Optimal classification: 2 approaches

① Minimize  $P(\text{error})$

Given  $x$ , classify as  $\hat{G}(x) = j$  if  $\lambda_j f_j(x) > \lambda_i f_i(x)$  for all  $i \neq j$

i.e.  $R_j = \{x: \lambda_j f_j(x) > \lambda_i f_i(x) \text{ for all } i \neq j\}$

② Look at conditional dist'n of  $G$  given  $X=x$

$$\text{Bayes Thm: } P(G=j | X=x) = \frac{P(G=j, X=x)}{P(X=x)}$$

$$= \frac{P(G=j)P(X=x | G=j)}{P(X=x)}$$

$$= \frac{\lambda_j f_j(x)}{\lambda_1 f_1(x) + \dots + \lambda_k f_k(x)}$$

$\Rightarrow$  if we choose  $\hat{G}(x)$  to maximize  $P(G=j | X=x)$

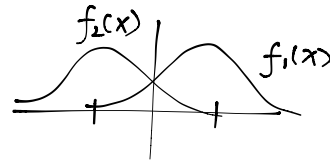
we get the classification rule in ①

But  $P(G=j | X=x)$  provides more information.

Example  $k=2, p=1, \lambda_1 = \lambda_2 = \frac{1}{2}$

$$f_1(x) = \frac{1}{\sqrt{2\pi}} \exp(-\frac{1}{2}(x-1)^2) \quad (N(1, 1))$$

$$f_2(x) = \frac{1}{\sqrt{2\pi}} \exp(-\frac{1}{2}(x+1)^2) \quad (N(-1, 1))$$



① gives  $\hat{G}(x) = 1$  if  $x > 0$

$\hat{G}(x) = 2$  if  $x < 0$

$$\text{② } P(G=1 | X=x) = \frac{f_1(x)}{f_1(x) + f_2(x)}$$

$x$	$P(G=1   X=x)$
-1	0.119
-0.5	0.269
-0.25	0.378
0	0.500
0.25	0.622
0.5	0.731
1	0.881

## Linear discriminant analysis (LDA)

$f_1(x), \dots, f_k(x)$  multivariate normal with distinct means  $\mu_1, \dots, \mu_k$  and common covariance matrix  $C$ .

If everything is known then  $\hat{G}(x) = j$  if  $d_j(x) > d_i(x)$  for all  $i \neq j$  where

$$d_j(x) = x^T C^{-1} \mu_j - \frac{1}{2} \mu_j^T C^{-1} \mu_j + \ln(\lambda_j)$$

Given data  $(g_1, x_1), \dots, (g_n, x_n)$  we have 
$$\hat{\mu}_j = \frac{\sum_{i=1}^n x_i I(g_i=j)}{\sum_{i=1}^n I(g_i=j)} \quad (j=1, \dots, k)$$

$$\hat{C} = \frac{1}{n-k} \sum_{i=1}^n (x_i - \hat{\mu}_j)(x_i - \hat{\mu}_j)^T$$

Example: Fisher's iris data (see Blackboard)

3 species of irises  $\left\{ \begin{array}{l} \text{virginica} \\ \text{setosa} \\ \text{versicolour} \end{array} \right.$

4 variables  $\left\{ \begin{array}{l} \text{sepal length} \\ \text{sepal width} \\ \text{petal length} \\ \text{petal width} \end{array} \right.$

- LDA works very well here.
- But, could prob do as well from pairwise scatterplots!

How to estimate misclassification rate?

(as well as posterior dist'n)

- resubstitution estimate is typically biased downwards.
- bias increases often severely, as model complexity increases.

**Solution:** Do some sort of **cross-validation**

- divide data into 2 sets: training set ( $n-m$  obs) and test set ( $m$  obs)
- estimate classification rate based on training data and use test data to estimate misclassification rate.

For example, **10-fold cross-validation**

- divide data into 10 sets
- successively leave out one set  $\rightarrow$  test data  
use other 9 as training data

- **leave-one-out CV**: for each "observation"  $x_i$ .

we compare  $\hat{G}_i(x_i)$  where  $\hat{G}_i$  is the classification rate using all data except  $x_i$ .

$\Rightarrow$  used in R function: `lda`, `qda`

For the iris data: est'd misclassification rate =  $\frac{3}{150}$  (assume  $\lambda_1 = \lambda_2 = \lambda_3 = \frac{1}{3}$ )

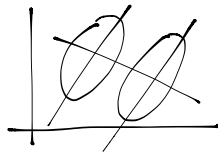
$$\text{Est'd misclassification rate} = \frac{\sum_{i=1}^n \lambda_{g_i} I(\hat{G}_i(x_i) \neq g_i)}{\sum_{i=1}^n \lambda_{g_i}} = 2\%$$

Can also estimate  $P(G=j|x_i)$  use **LOO CV**

- assume model is approx correct.

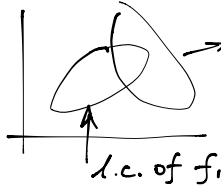
$\rightarrow 1 - 0 - 0$

### - Quadratic Discriminant Analysis (QDA)



$$k=2, p=2, G=G_2=C$$

level curves of  $f_1(\underline{x})$  &  $f_2(\underline{x})$  have same orientation



→ l.c. of  $f_2$

⇒ no linear boundaries for classification rule.

l.c. of  $f_1$

- general form of classification rule is the same but regions more complicated.