

Notebook 2 One Way ANOVA Example (Corn Yields)

The following R analysis of the corn data described on pages 6 and 7 of the brick, is an expanded version of the code shown on pages 7 to 11 of the brick.

There are some differences between this and the S-Plus version shown in the brick, which I will note and discuss below.

Read and attach the data and examine the contents. The advantage of attaching the data is that we can refer to the elements of the data as if they were vectors, without having to specify the data= option in functions such as lm():

```
corn <- read.table("Corn.txt",header=T)
corn
```

```
##      yield      fert
## 1      99 Control
## 2      40 Control
## 3      61 Control
## 4      72 Control
## 5      76 Control
## 6      84 Control
## 7      96  K20+N
## 8      84  K20+N
## 9      82  K20+N
## 10     104  K20+N
## 11      99  K20+N
## 12     105  K20+N
## 13      63 K20+P205
## 14      57 K20+P205
## 15      81 K20+P205
## 16      59 K20+P205
## 17      64 K20+P205
## 18      72 K20+P205
## 19      79  N+P205
## 20      92  N+P205
## 21      91  N+P205
## 22      87  N+P205
## 23      78  N+P205
## 24      71  N+P205
```

```
attach(corn)
names(corn)
```

```
## [1] "yield" "fert"
```

```
yield
```

```
## [1] 99 40 61 72 76 84 96 84 82 104 99 105 63 57 81 59 64
## [18] 72 79 92 91 87 78 71
```

```
fert
```

```
## [1] Control Control Control Control Control Control K20+N
## [8] K20+N K20+N K20+N K20+N K20+N K20+P205 K20+P205
```

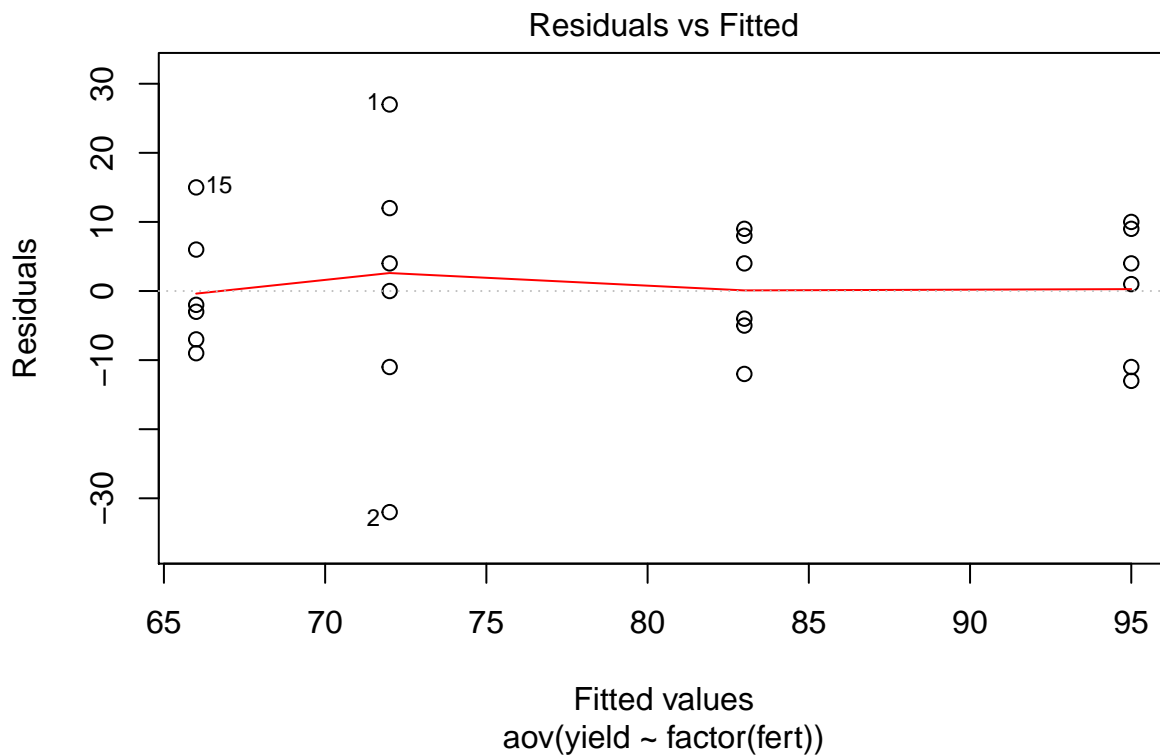
```
## [15] K20+P205 K20+P205 K20+P205 K20+P205 N+P205    N+P205    N+P205
## [22] N+P205    N+P205    N+P205
## Levels: Control K20+N K20+P205 N+P205

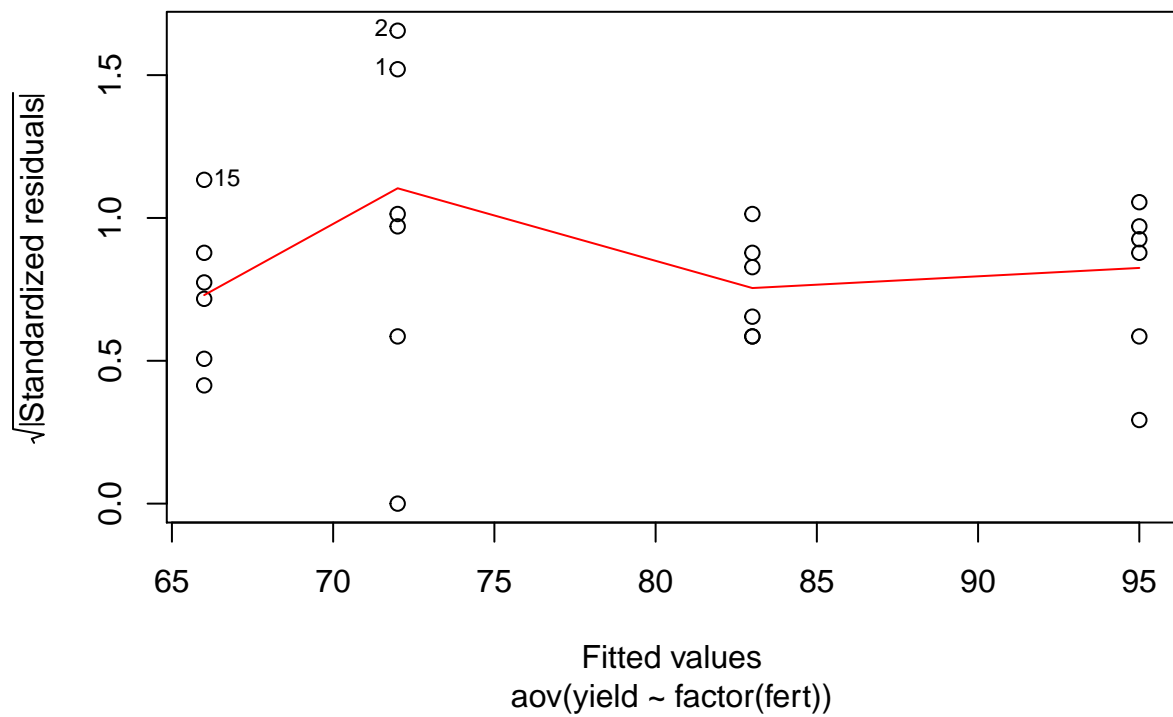
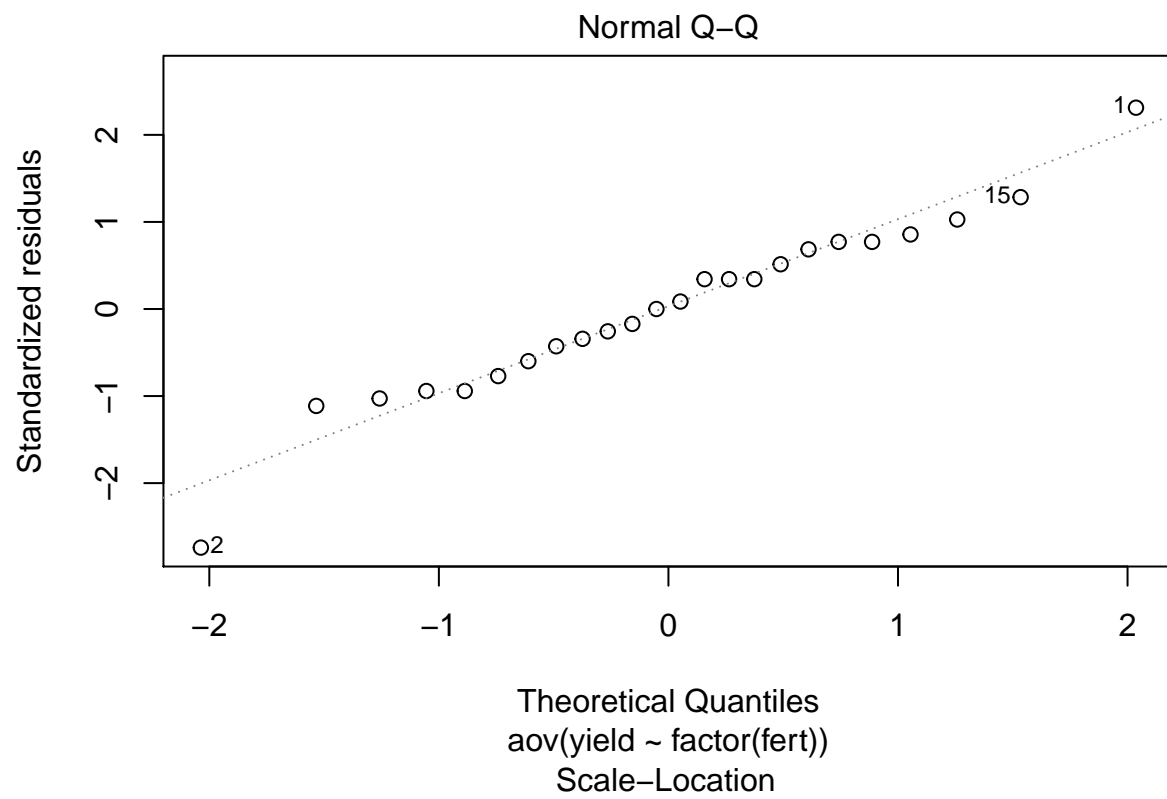
# The aov command in R (or S-Plus) is used to fit the one-way ANOVA model described in the brick.
# The factor command is used to indicate that fert is a factor variable:

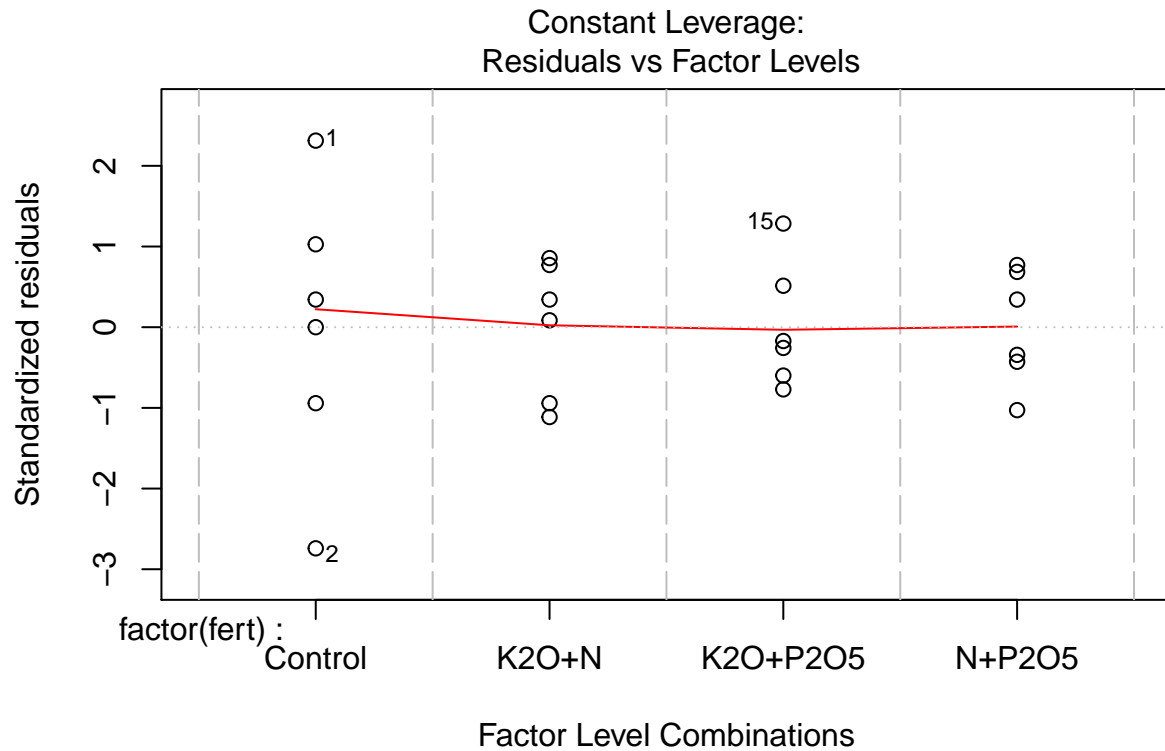
corn.aov <- aov(yield ~ factor(fert))

# We can use the plot command to assess whether this is an appropriate model for the data.

plot(corn.aov)
```







*# Here is a quick description of what the generic plot() function produces in R for a
lm() or aov() object. For more details, see:*

`help(plot.lm)`

The four plots produced by default for a stored linear model are:

*# Plot 1 - a plot of the residuals vs fitted values which is very important for assessing the
underlying assumptions of independence and constant variance*

Plot 2 - a normal quantile plot of the rstandard() residuals

*# Plot 3 - another plot of the (transformed) residuals vs the fitted values, this is an another
attempt to graphically examine the assumption of constant variance*

*# Plot 4 - a leverage plot - for a lm() object, this will typically be a plot of the standardised
residuals against the leverage values. On such a plot, arbitrary limits can be drawn for the
Cook's D values for each of the data points, as Cook's D is a function of the
standardised residuals and the leverage values.*

*# For an aov() object, especially for a balanced experimental design where all the observations
have the same (constant leverage), plot 4 becomes a plot of the standardised residuals against
the treatments (the factor level combinations).*

*# These plots in R suggest a problem - there appear to be two possible outliers in one of the
groups - observations 1 and 2 - more about this problem later.*

*# We can examine the ANOVA table for an object created using the aov() function by applying the
summary function:*

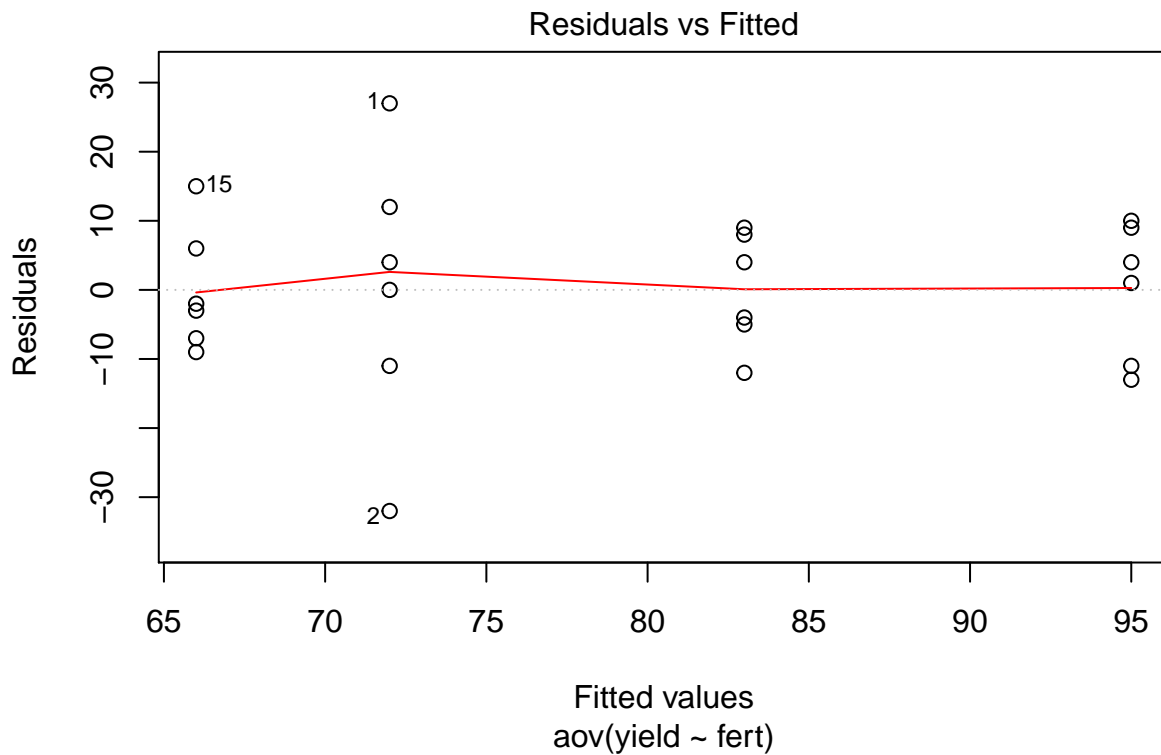
`summary(corn.aov)`

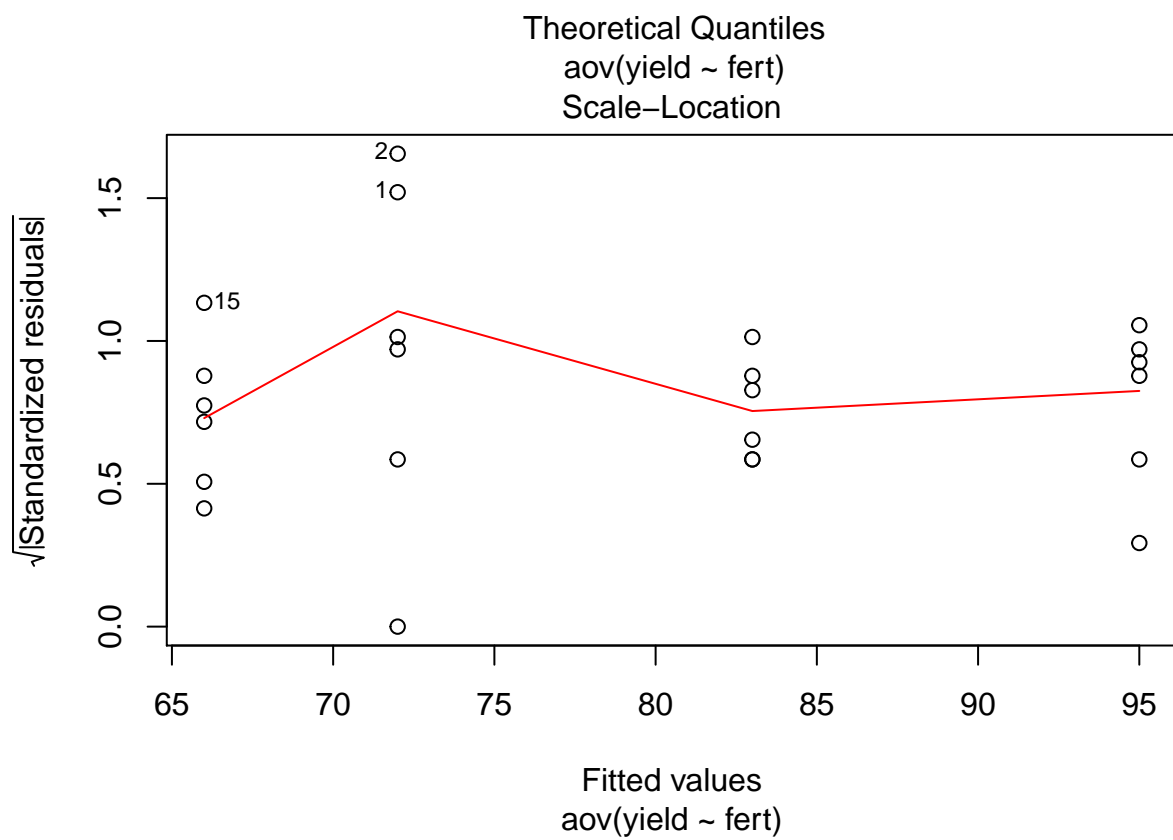
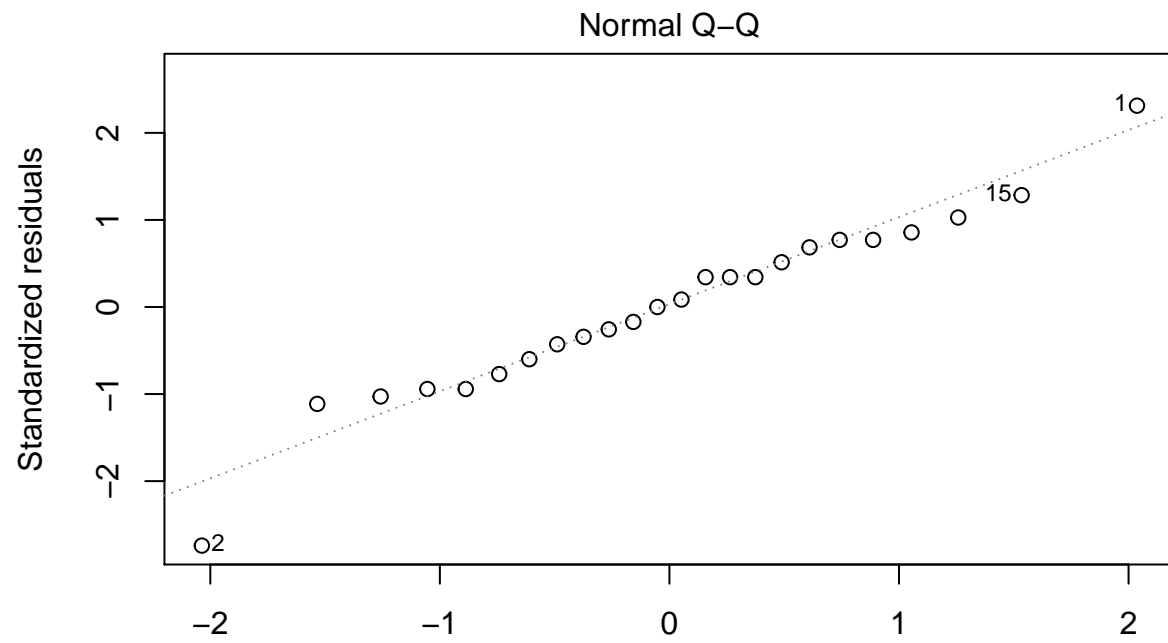
```
##           Df Sum Sq Mean Sq F value Pr(>F)
## factor(fert) 3   2940    980.0    5.99 0.00439 **
## Residuals   20   3272    163.6
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

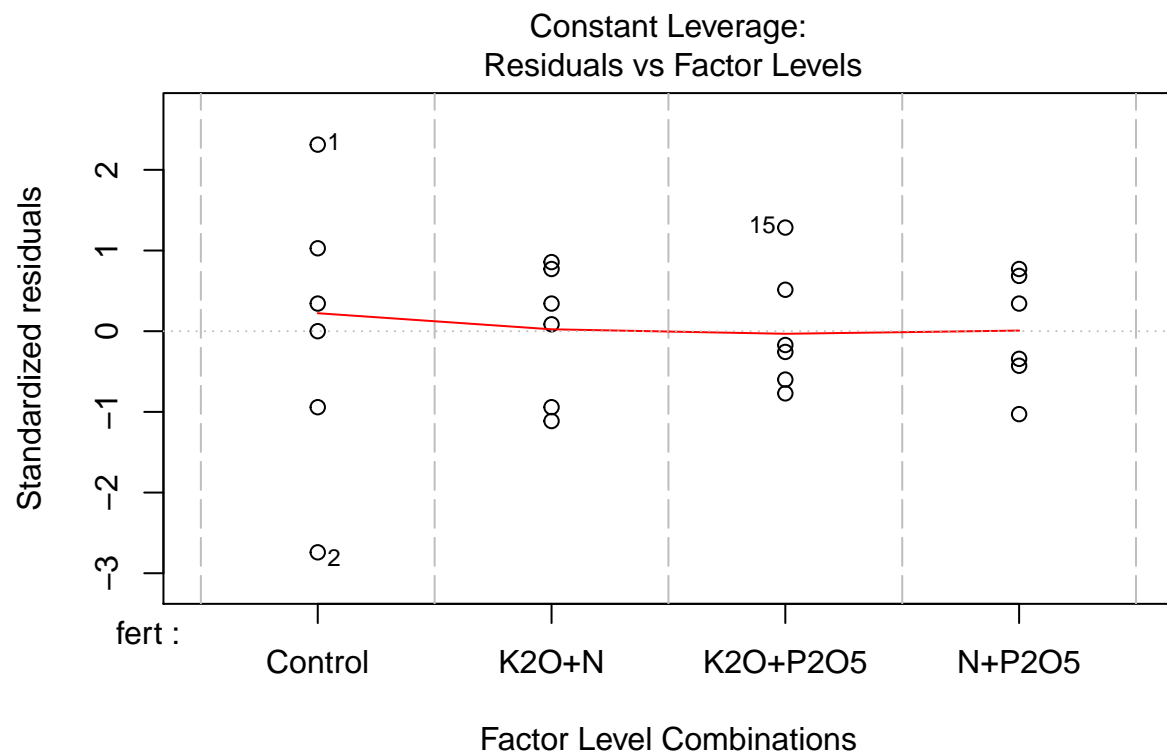
# Note that R adopts the dubious practice of starring the results at different level of
# significance - dubious because the choice of significance level should be chosen a priori
# (before the experiment), rather than adjusted to suit, after the analysis (a posteriori).

# Note the factor command is only really necessary when the categorical variable is coded using
# numbers such as 1,2,3,4. Here fert is a series of qualitative labels, which R would have
# correctly interpreted as a factor, not a continuous variable - as can be seen if we refit
# the model:

corn.aov2 <- aov(yield ~ fert)
plot(corn.aov2)
```







```
summary(corn.aov2)
```

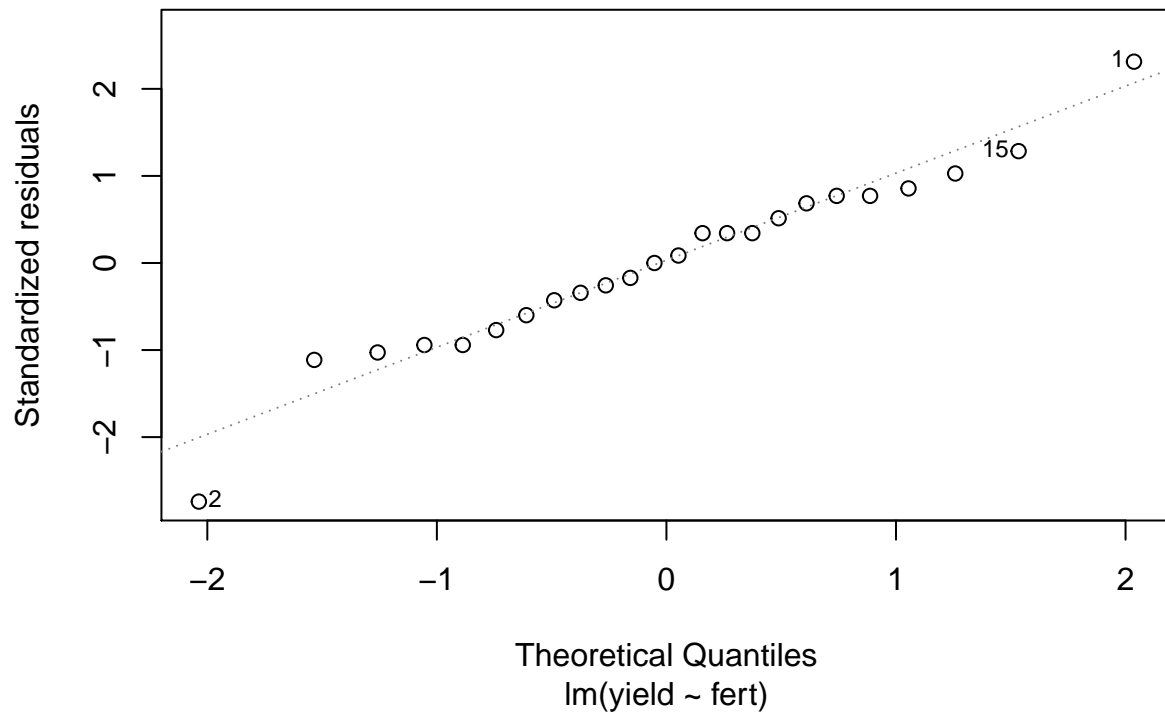
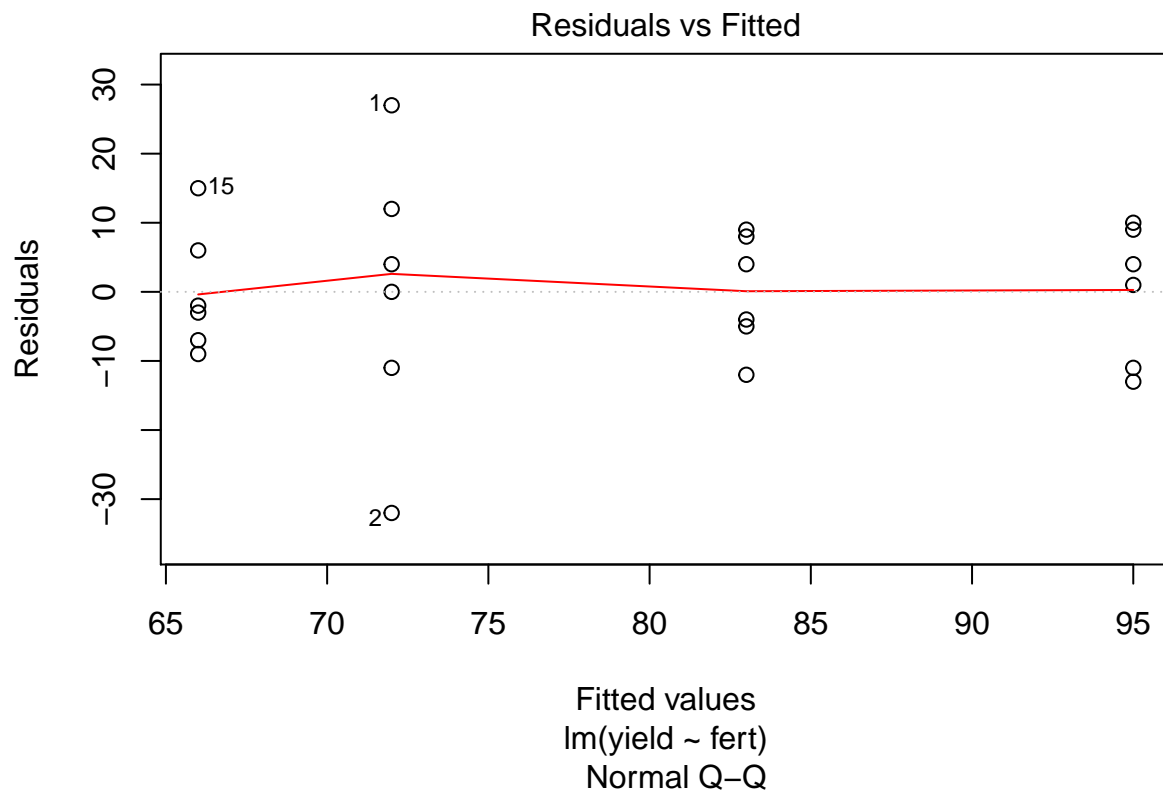
```
##           Df Sum Sq Mean Sq F value    Pr(>F)
## fert         3   2940    980.0     5.99 0.00439 **
## Residuals    20   3272    163.6
```

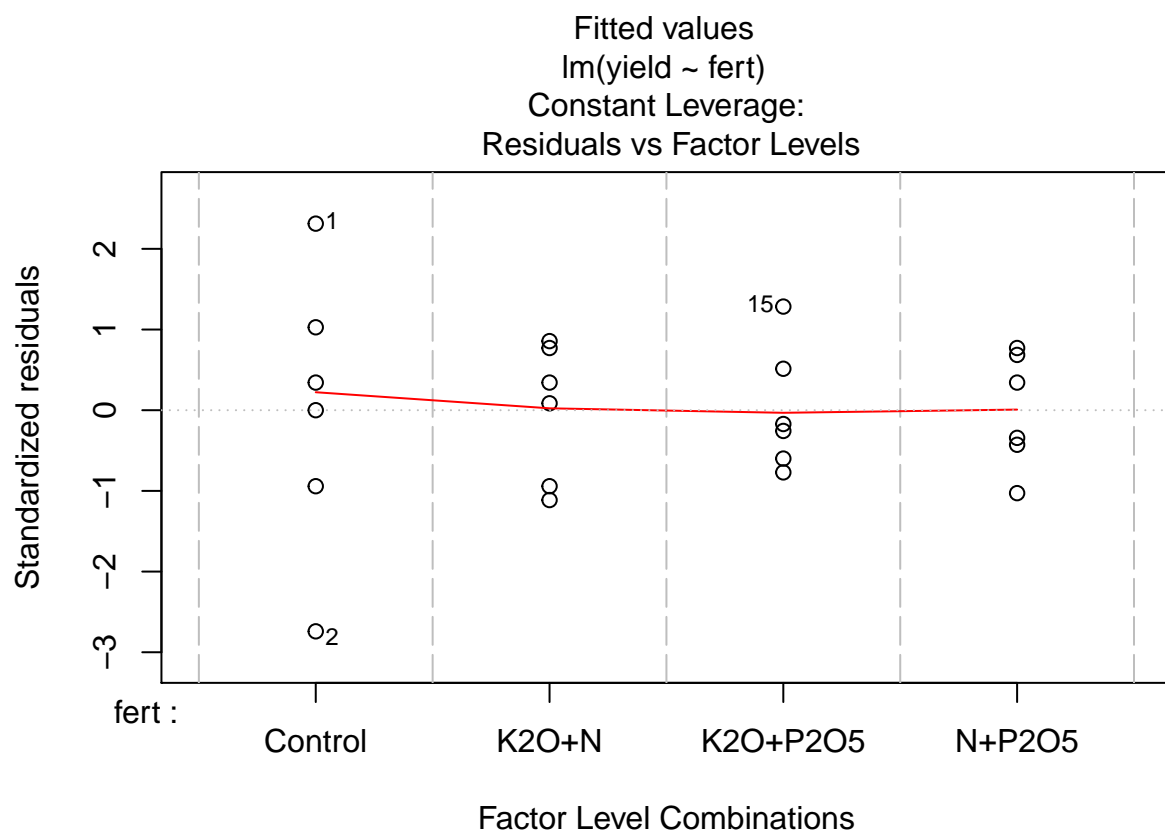
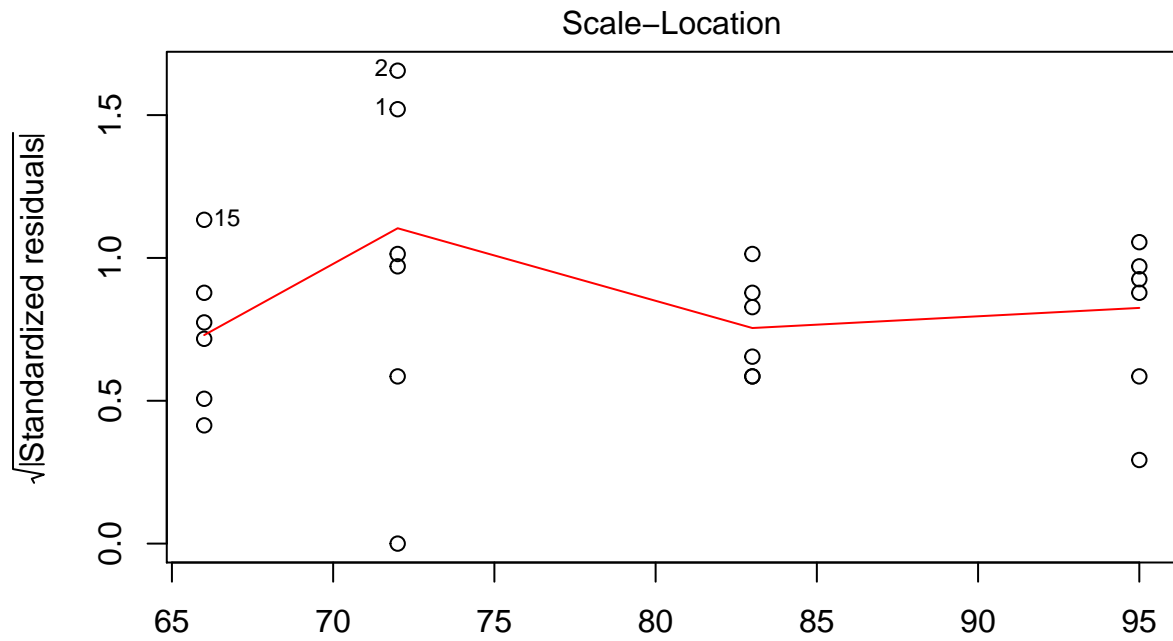
```
## ---
```

```
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
# We could also use the lm() function used to fit regression models in STAT2008 to fit the same
# model:
```

```
corn.lm <- lm(yield ~ fert)
plot(corn.lm)
```





```
anova(corn.lm)
```

```
## Analysis of Variance Table
##
## Response: yield
##          Df Sum Sq Mean Sq F value    Pr(>F)
## fert      3   2940    980.0    5.9902 0.004387 **
```

```
## Residuals 20    3272    163.6
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
summary(corn.lm)
```

```
##
## Call:
## lm(formula = yield ~ fert)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -32.00  -7.50   0.50   8.25  27.00
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)    72.000     5.222  13.788 1.13e-11 ***
## fertK20+N      23.000     7.385   3.115  0.00546 **
## fertK20+P205   -6.000     7.385  -0.812  0.42607
## fertN+P205     11.000     7.385   1.490  0.15194
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 12.79 on 20 degrees of freedom
## Multiple R-squared:  0.4733, Adjusted R-squared:  0.3943
## F-statistic:  5.99 on 3 and 20 DF,  p-value: 0.004387
```

*# The main difference between an lm() and an aov() object is that summary gives the table of coefficients for the lm() object and you need to use anova() to see the ANOVA table.
To get the coefficients for an aov() object you need to use:*

```
summary.lm(corn.aov2)
```

```
##
## Call:
## aov(formula = yield ~ fert)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -32.00  -7.50   0.50   8.25  27.00
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)    72.000     5.222  13.788 1.13e-11 ***
## fertK20+N      23.000     7.385   3.115  0.00546 **
## fertK20+P205   -6.000     7.385  -0.812  0.42607
## fertN+P205     11.000     7.385   1.490  0.15194
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 12.79 on 20 degrees of freedom
## Multiple R-squared:  0.4733, Adjusted R-squared:  0.3943
## F-statistic:  5.99 on 3 and 20 DF,  p-value: 0.004387
```

How do we interpret these coefficients? Unlike version 7 and earlier versions of S-Plus, # (where the rather strange choice of defaults were Helmert contrasts, which are an attempt

```
# to deal with highly unbalanced or non-orthogonal experimental designs) the default
# parameterisation for factor variables in R and in S-Plus from version 8 onwards are treatment
# contrasts:
```

```
contrasts(fert)
```

```
##           K20+N K20+P205 N+P205
## Control      0         0       0
## K20+N        1         0       0
## K20+P205     0         1       0
## N+P205       0         0       1
```

```
# Under treatment contrasts, the model parameters ARE closely related to the mean yields
# for the four fertilizers, which can be calculated as follows:
```

```
mean(yield)
```

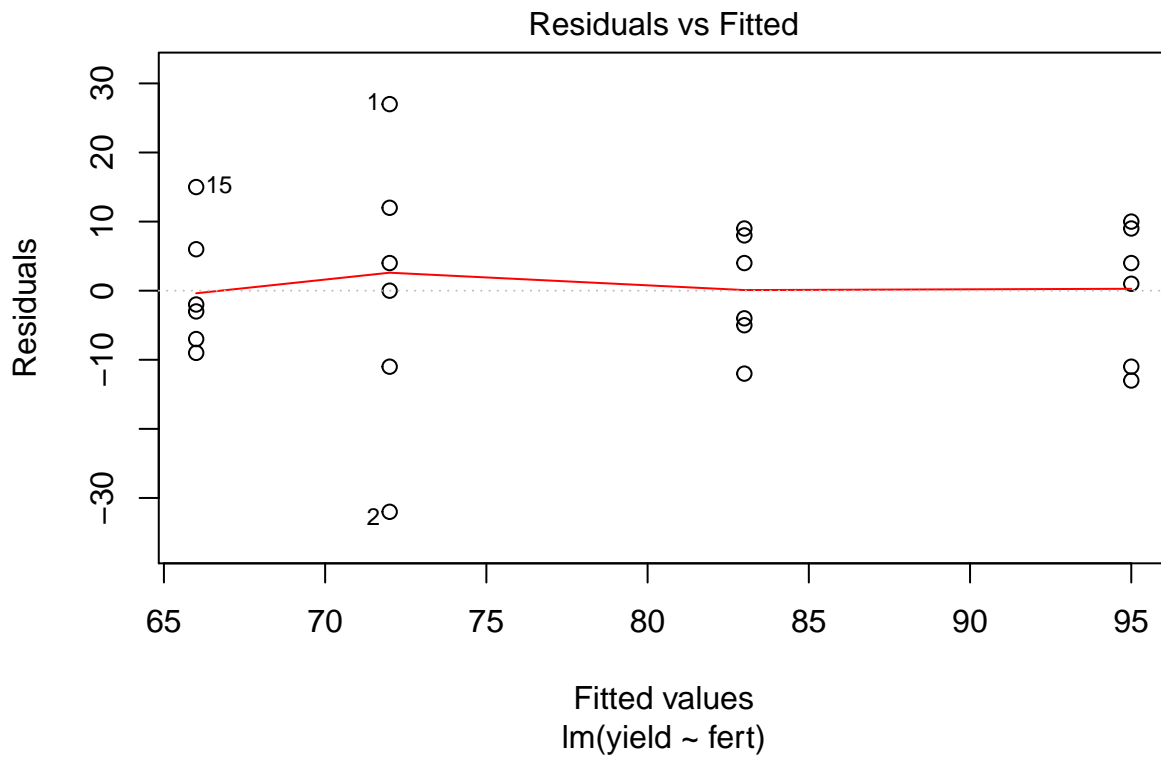
```
## [1] 79
```

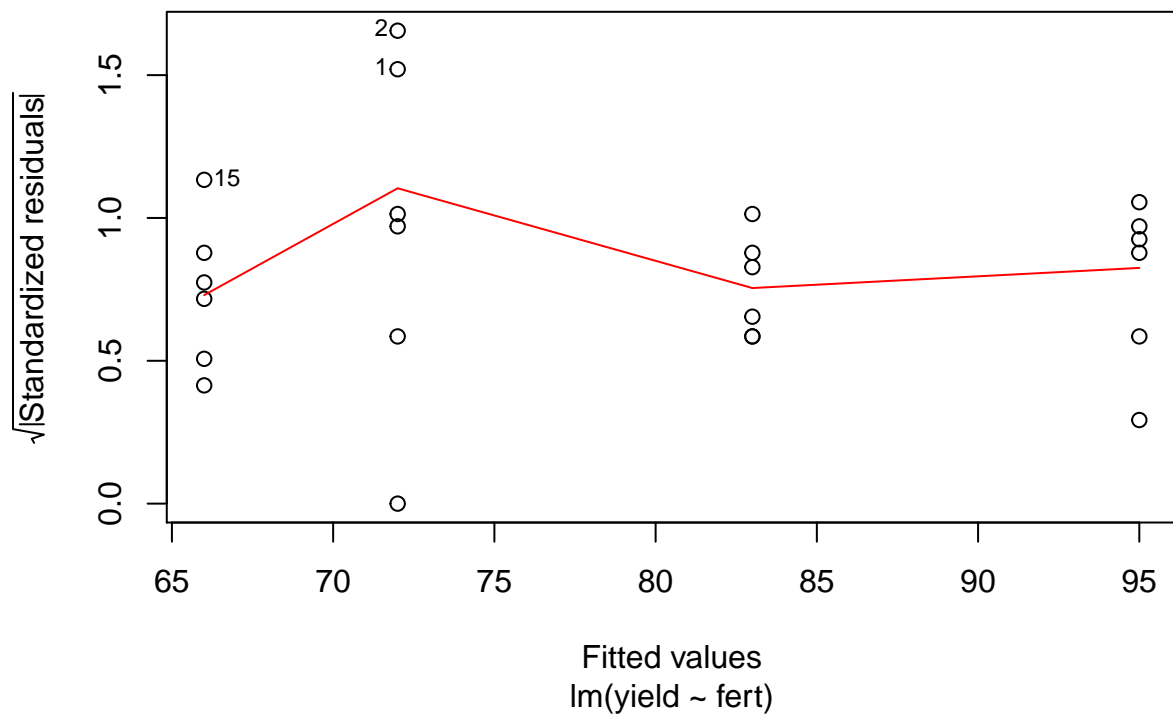
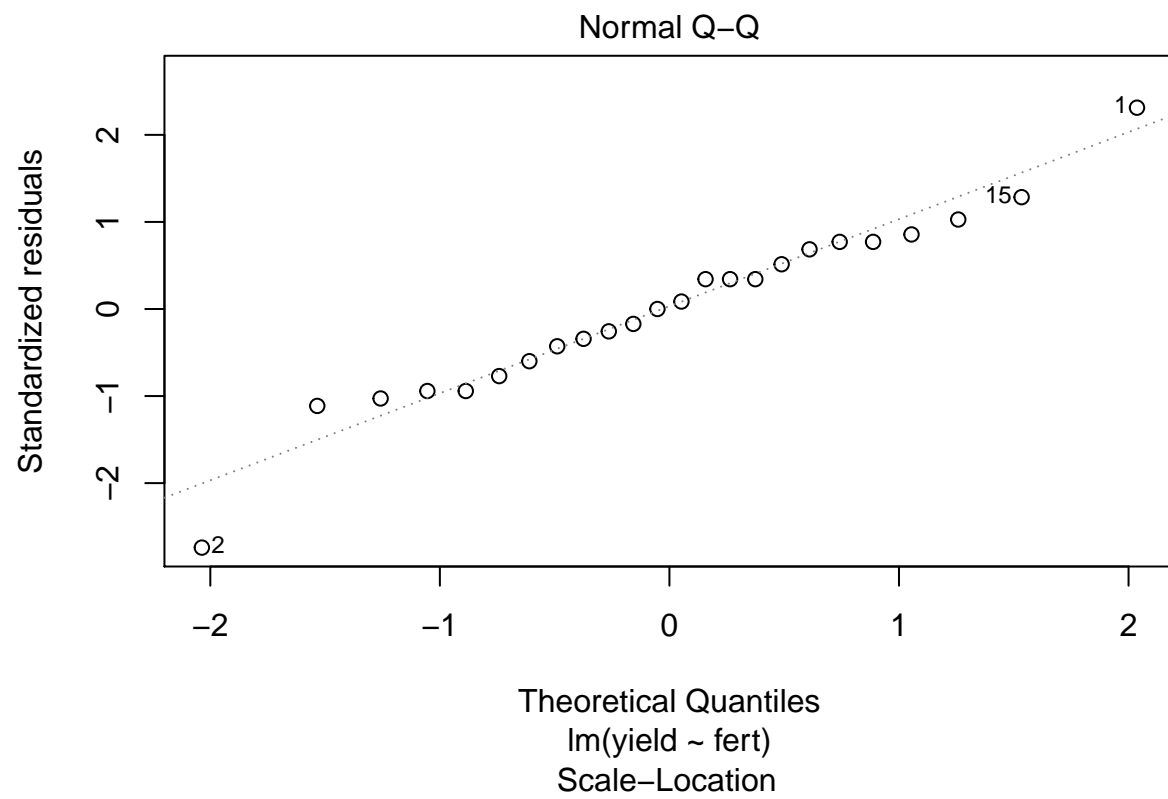
```
lvl.mns <- tapply(yield, fert, mean)
lvl.mns
```

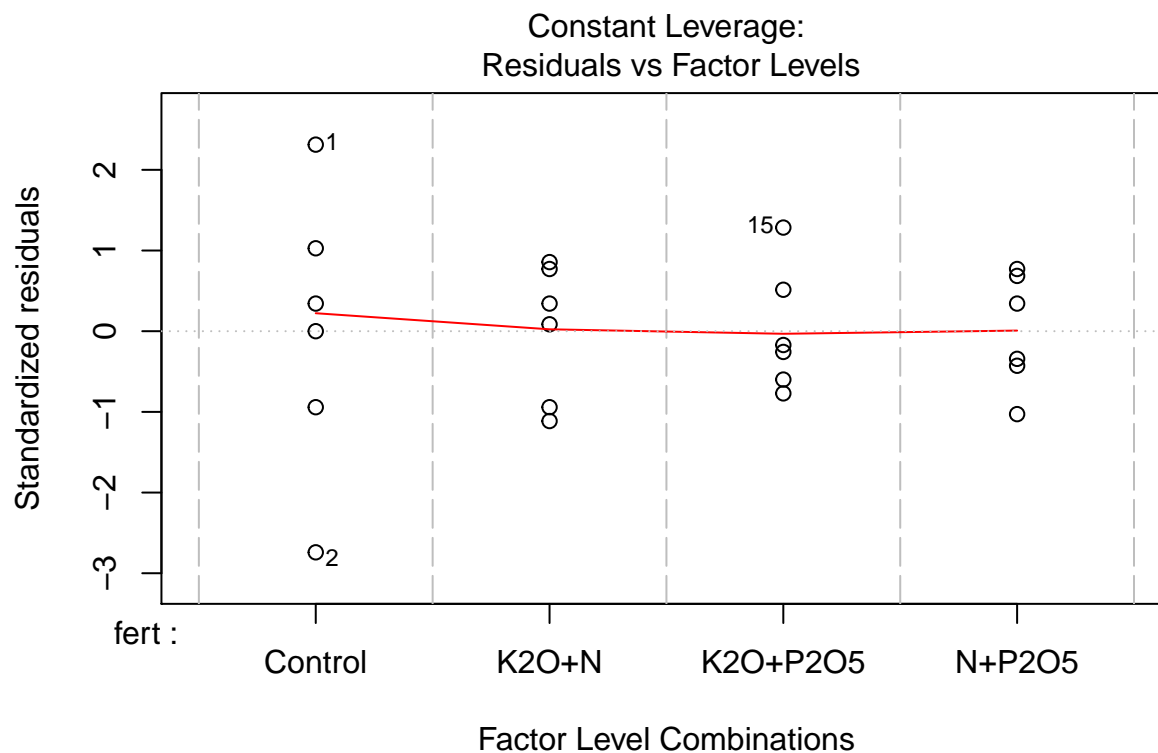
```
## Control      K20+N K20+P205  N+P205
##       72       95       66      83
```

```
# The following approach will also fit the 0-1 dummy or indicator variables described on page 10 of the
# In R or S-Plus version 8, this is the default version, so will give identical results to the previous
```

```
corn.lm2 <- lm(yield ~ fert, contrasts=list(fert=contr.treatment))
plot(corn.lm2)
```







```
anova(corn.lm2)
```

```
## Analysis of Variance Table
##
## Response: yield
##          Df Sum Sq Mean Sq F value    Pr(>F)
## fert      3  2940   980.0    5.9902 0.004387 **
## Residuals 20  3272   163.6
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
summary(corn.lm2)
```

```
##
## Call:
## lm(formula = yield ~ fert, contrasts = list(fert = contr.treatment))
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -32.00  -7.50   0.50   8.25  27.00
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)    72.000     5.222  13.788 1.13e-11 ***
## fert2          23.000     7.385   3.115 0.00546 **
## fert3         -6.000     7.385  -0.812 0.42607
## fert4          11.000     7.385   1.490 0.15194
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 12.79 on 20 degrees of freedom
```

```
## Multiple R-squared:  0.4733, Adjusted R-squared:  0.3943
## F-statistic:  5.99 on 3 and 20 DF,  p-value: 0.004387

# In this "treatment" parameterisation, R has chosen the first or Control group as the reference
# group by creating the following 0-1 indicator variables:

contr.treatment(4)

##      2 3 4
## 1 0 0 0
## 2 1 0 0
## 3 0 1 0
## 4 0 0 1

# If we are not happy with R's default choice of reference group, we could do our own manual
# coding. Here is the manually coded equivalent of what R has done above:

fert1 <- ifelse(fert=="Control",1,0)
fert2 <- ifelse(fert=="K20+N",1,0)
fert3 <- ifelse(fert=="K20+P205",1,0)
fert4 <- ifelse(fert=="N+P205",1,0)
ferts <- cbind(fert1, fert2, fert3, fert4)
ferts

##      fert1 fert2 fert3 fert4
## [1,]      1      0      0      0
## [2,]      1      0      0      0
## [3,]      1      0      0      0
## [4,]      1      0      0      0
## [5,]      1      0      0      0
## [6,]      1      0      0      0
## [7,]      0      1      0      0
## [8,]      0      1      0      0
## [9,]      0      1      0      0
## [10,]     0      1      0      0
## [11,]     0      1      0      0
## [12,]     0      1      0      0
## [13,]     0      0      1      0
## [14,]     0      0      1      0
## [15,]     0      0      1      0
## [16,]     0      0      1      0
## [17,]     0      0      1      0
## [18,]     0      0      1      0
## [19,]     0      0      0      1
## [20,]     0      0      0      1
## [21,]     0      0      0      1
## [22,]     0      0      0      1
## [23,]     0      0      0      1
## [24,]     0      0      0      1

corn.lm2a <- lm(yield ~ ferts)

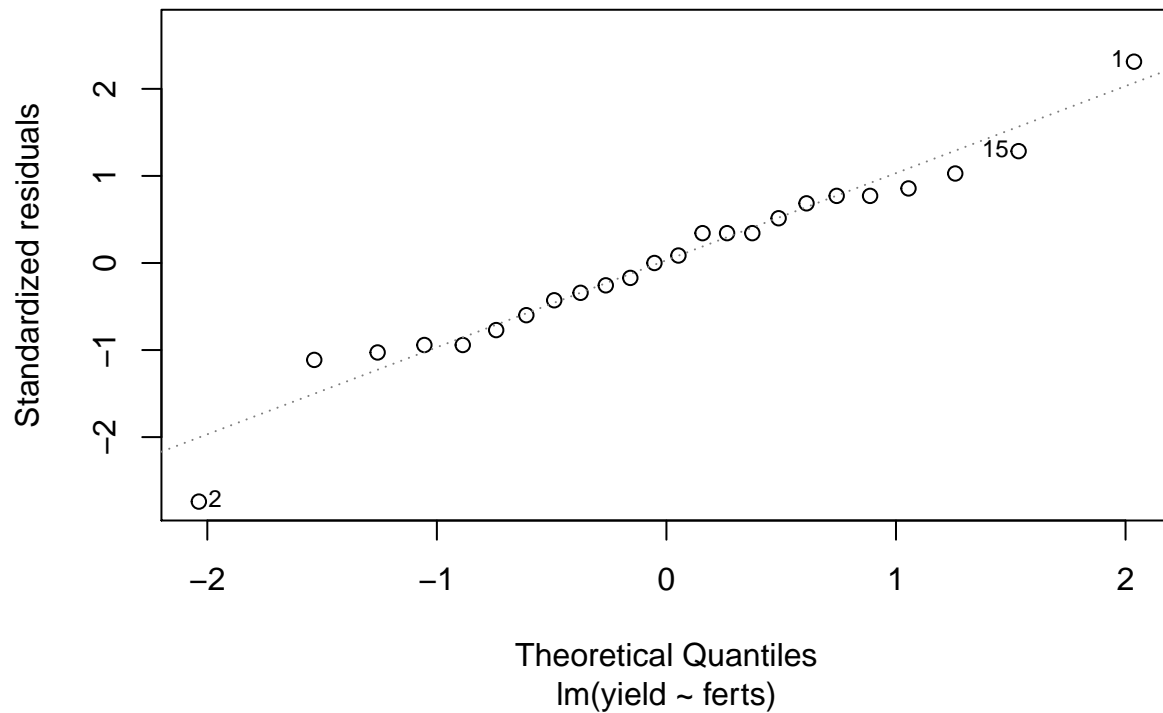
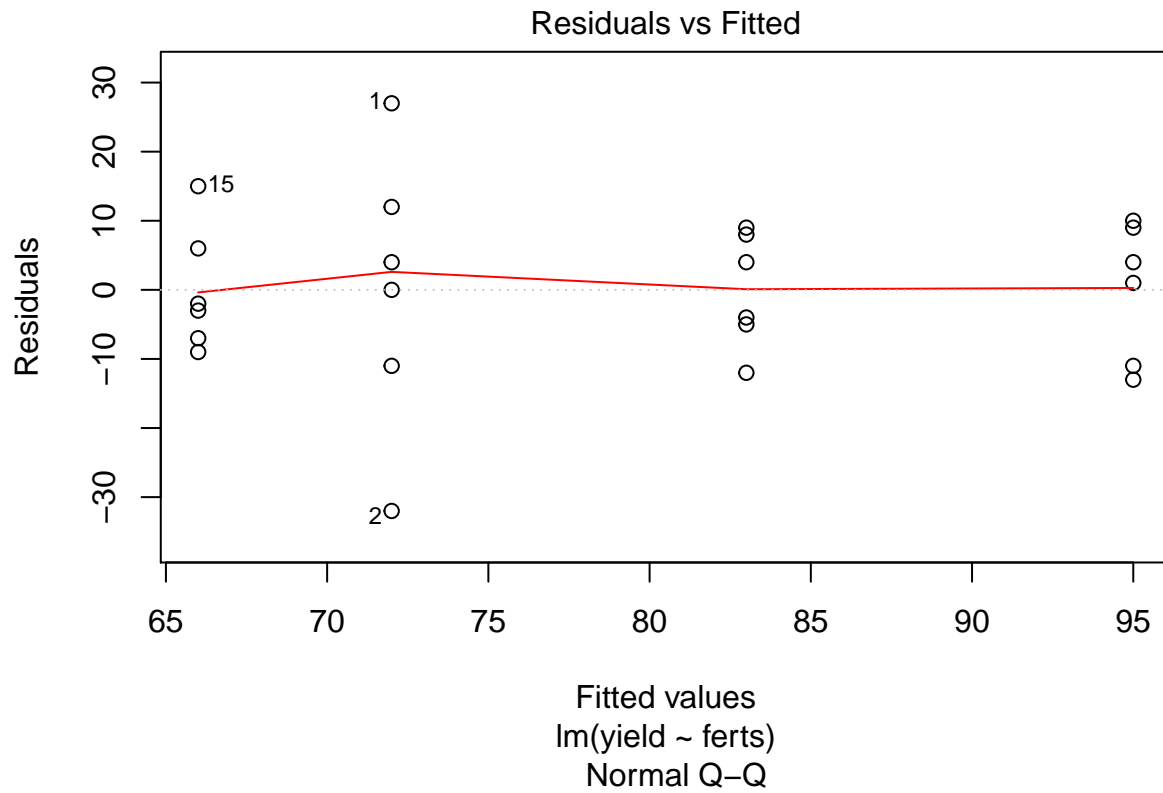
# Note that unlike S-Plus, R does not give an error message to indicate that this model is
# over-parameterised, but if you examine the coefficients for this model, you find that R has
# simply decided not to fit the last of the parameters (fert4):

```

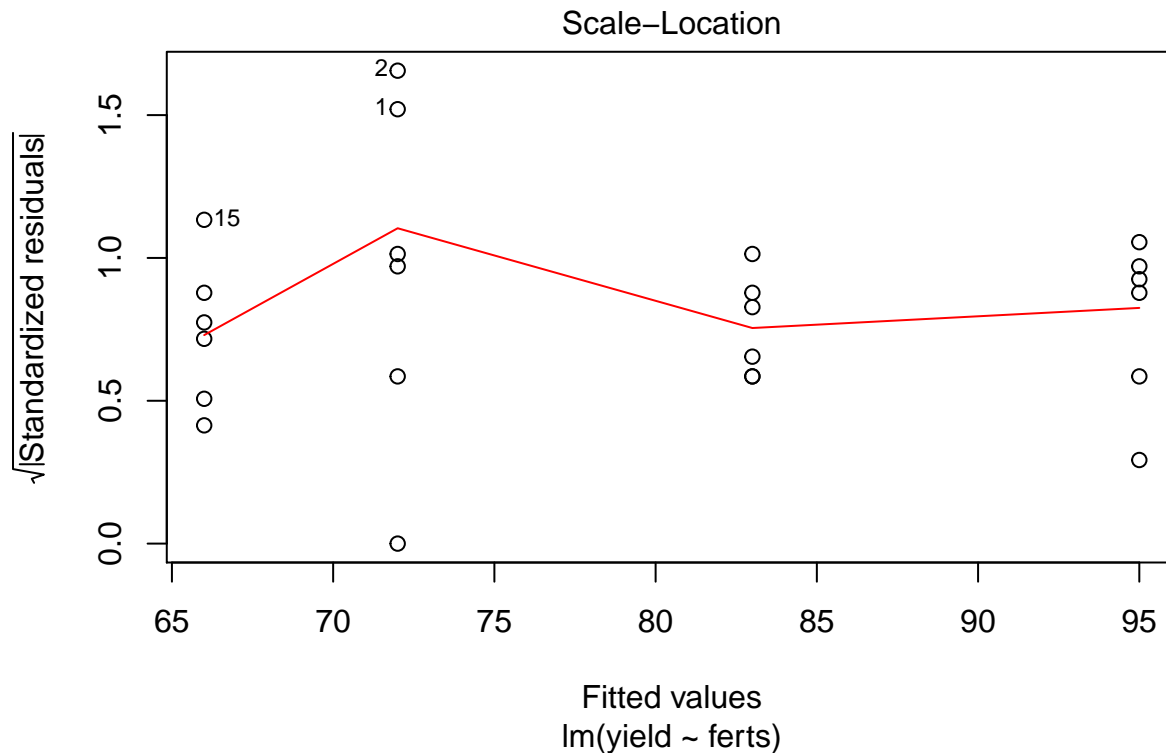
```
summary(corn.lm2a)
```

```
##
## Call:
## lm(formula = yield ~ ferts)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -32.00  -7.50   0.50   8.25  27.00
##
## Coefficients: (1 not defined because of singularities)
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)   83.000      5.222  15.895 8.23e-13 ***
## fertsfert1   -11.000      7.385  -1.490  0.1519
## fertsfert2    12.000      7.385   1.625  0.1198
## fertsfert3   -17.000      7.385  -2.302  0.0322 *
## fertsfert4         NA          NA      NA      NA
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 12.79 on 20 degrees of freedom
## Multiple R-squared:  0.4733, Adjusted R-squared:  0.3943
## F-statistic:  5.99 on 3 and 20 DF,  p-value: 0.004387
# Using treatment coding, the intercept becomes the mean for the group for which we don't fit
# an indicator variable - this is called the reference group. In this instance, a better choice
# for the reference group is the Control group, so we should delete fert1 rather than fert4:

ferts <- cbind(fert2, fert3, fert4)
corn.lm2a <- lm(yield ~ ferts)
plot(corn.lm2a)
```



```
## hat values (leverages) are all = 0.1666667
## and there are no factor predictors; no plot no. 5
```

```
anova(corn.lm2a)
```

```
## Analysis of Variance Table
##
## Response: yield
##          Df Sum Sq Mean Sq F value    Pr(>F)
## ferts      3  2940   980.0    5.9902 0.004387 **
## Residuals 20  3272   163.6
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
summary(corn.lm2a)
```

```
##
## Call:
## lm(formula = yield ~ ferts)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -32.00  -7.50   0.50   8.25  27.00
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)    72.000     5.222  13.788 1.13e-11 ***
## fertsfert2     23.000     7.385   3.115  0.00546 **
## fertsfert3     -6.000     7.385  -0.812  0.42607
## fertsfert4     11.000     7.385   1.490  0.15194
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 12.79 on 20 degrees of freedom
```

```
## Multiple R-squared:  0.4733, Adjusted R-squared:  0.3943
## F-statistic:  5.99 on 3 and 20 DF,  p-value: 0.004387
```

```
# If we remember the level means calculated earlier, we can see that under the treatment
# parameterisation, the intercept is the mean for the "Control" group and the other parameters
# are the deviations away from this mean to get to the other group means:
```

```
lvl.mns
```

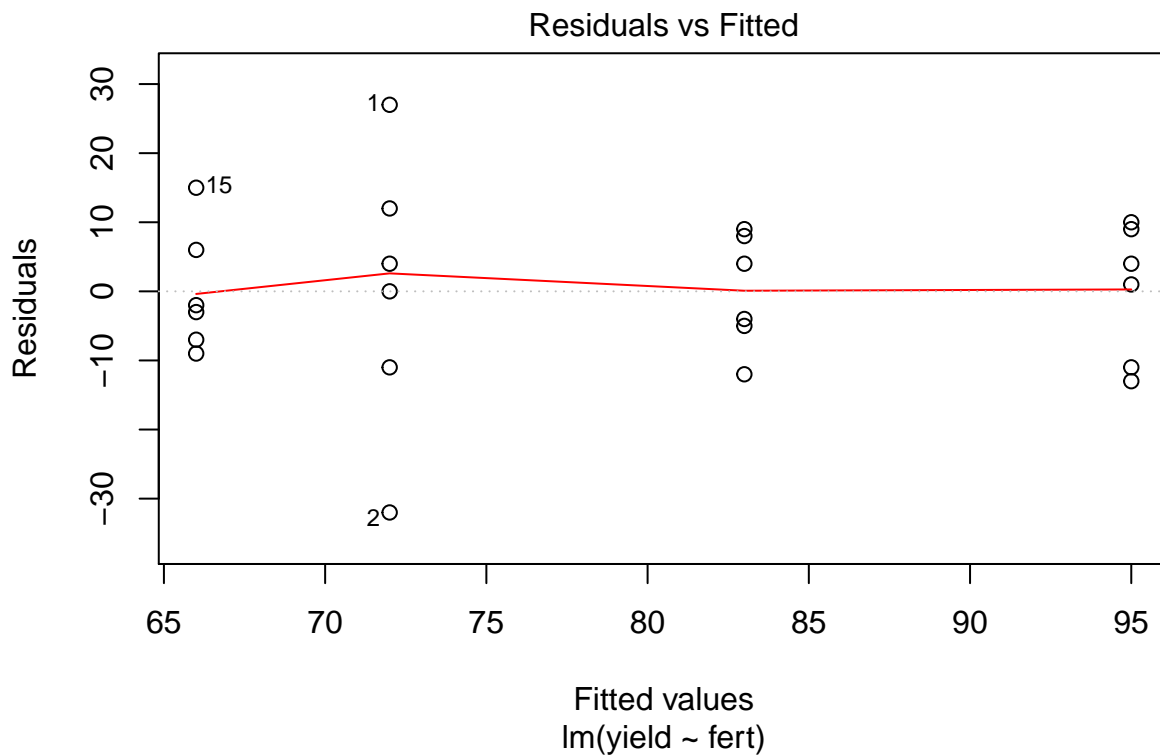
```
## Control    K20+N K20+P205    N+P205
##      72      95      66      83
```

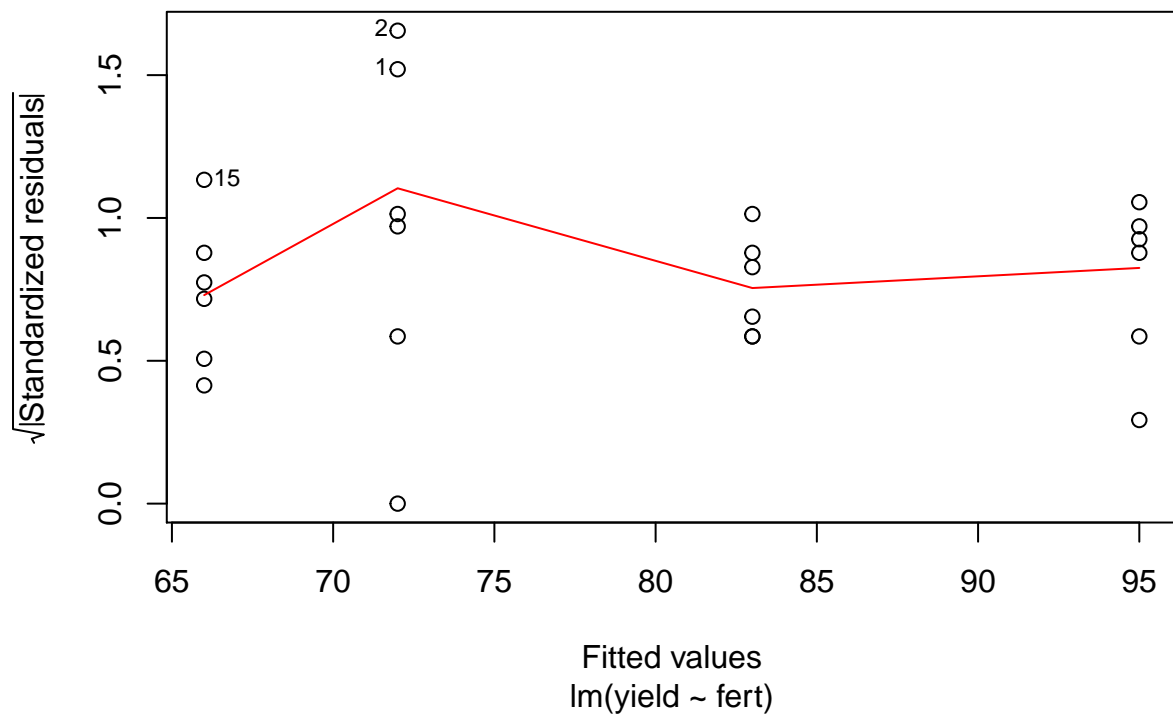
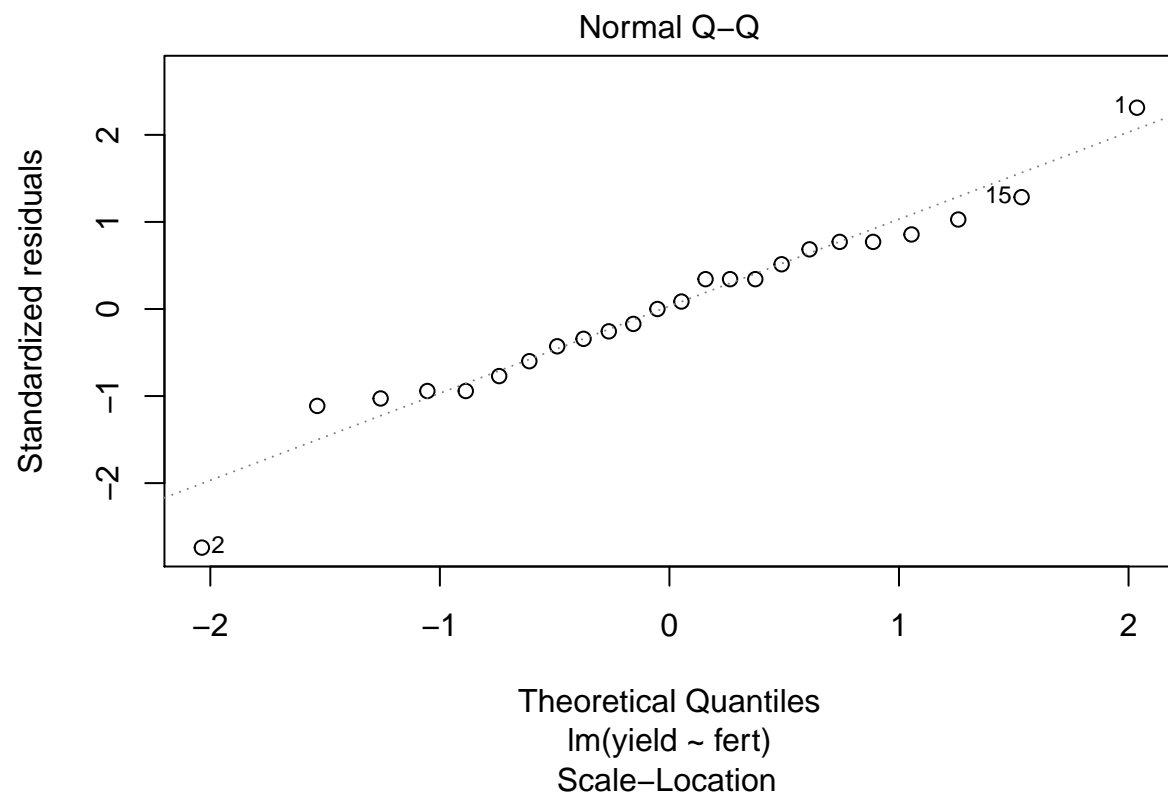
```
lvl.mns - lvl.mns["Control"]
```

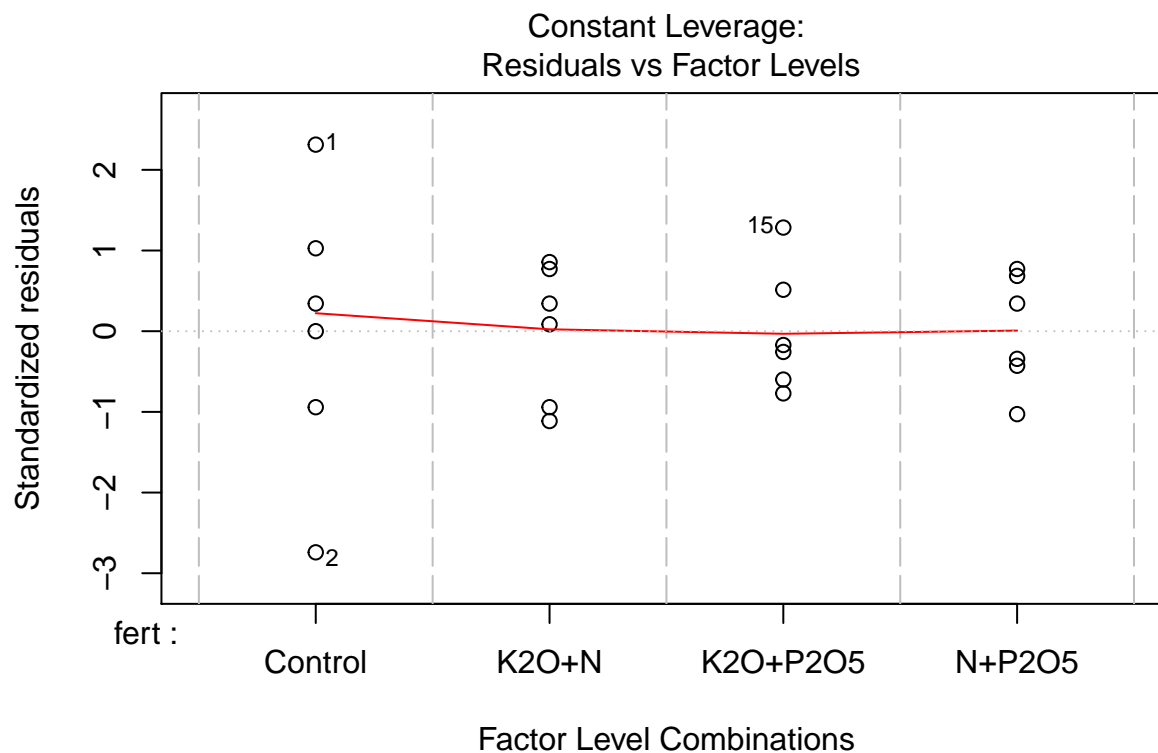
```
## Control    K20+N K20+P205    N+P205
##      0      23      -6      11
```

```
# Another approach is the "sum" parameterisation, which also has a sensible interpretation:
```

```
corn.lm3 <- lm(yield ~ fert, contrasts=list(fert=contr.sum))
plot(corn.lm3)
```







```
anova(corn.lm3)
```

```
## Analysis of Variance Table
##
## Response: yield
##          Df Sum Sq Mean Sq F value    Pr(>F)
## fert      3   2940   980.0    5.9902 0.004387 **
## Residuals 20   3272   163.6
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
summary(corn.lm3)
```

```
##
## Call:
## lm(formula = yield ~ fert, contrasts = list(fert = contr.sum))
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -32.00  -7.50   0.50   8.25  27.00
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)    79.000     2.611  30.258 < 2e-16 ***
## fert1         -7.000     4.522  -1.548  0.13732
## fert2         16.000     4.522   3.538  0.00206 **
## fert3        -13.000     4.522  -2.875  0.00937 **
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 12.79 on 20 degrees of freedom
```

```
## Multiple R-squared:  0.4733, Adjusted R-squared:  0.3943
## F-statistic:  5.99 on 3 and 20 DF,  p-value: 0.004387

# Here the intercept is the overall mean and the other parameters are deviations away from the
# overall mean to the first three level means (in alphabetical order). The "sum"
# parameterisation is equivalent to applying the constraint that the (weighted) deviations for
# the four groups sum to 0, so the deviation for the fourth group can be found by subtraction:

mean(yield)

## [1] 79

lvl.mns - mean(yield)

## Control      K20+N K20+P205      N+P205
##      -7         16       -13         4

coef(corn.lm3)

## (Intercept)      fert1      fert2      fert3
##          79         -7         16        -13

coef(corn.lm3)[2:4]

## fert1 fert2 fert3
##    -7    16  -13

-sum(coef(corn.lm3)[2:4])

## [1] 4

# Note that this constraint is  $\text{sum}(\text{tau}_i) = 0$  not  $\text{sum}(n_i * \text{tau}_i) = 0$  from the notes. This will
# not make a difference if the design is balanced, although if the design is unbalanced the
# contr.sum function cannot be used to give the grand mean interpretation. In this case the
# coefficients for the fourth group should be:
#  $(-n_i[1]/n_i[4], -n_i[2]/n_i[4], -n_i[3]/n_i[4])$  and not  $(-1, -1, -1)$ .

contr.sum(4)

##      [,1] [,2] [,3]
## 1      1      0      0
## 2      0      1      0
## 3      0      0      1
## 4     -1     -1     -1

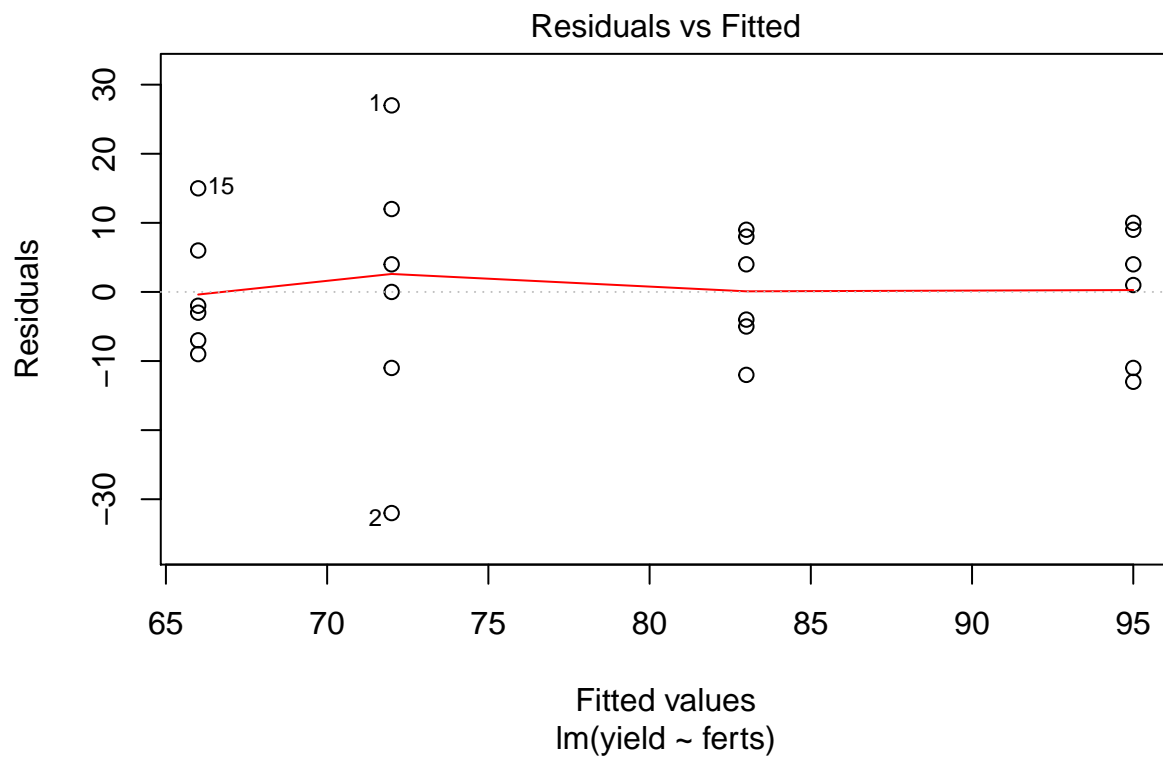
# We could also do a manual equivalent of this "sum" parameterisation:

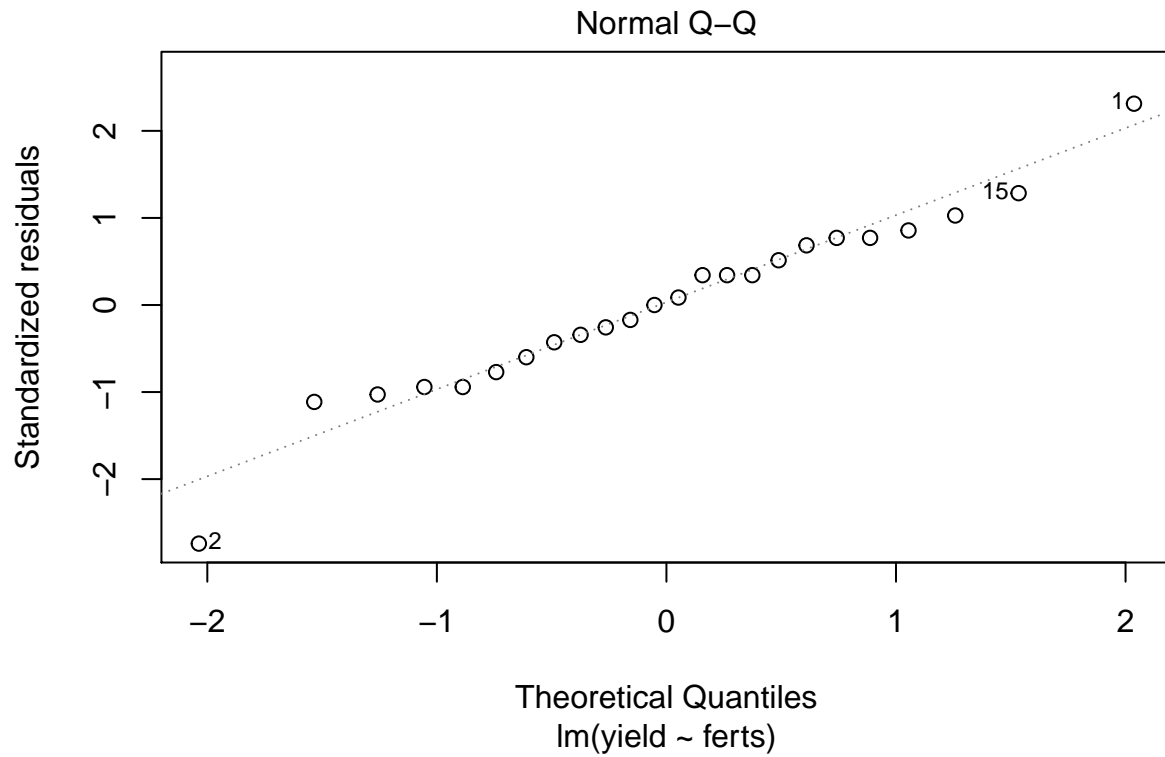
fert1 <- c(rep(1,6), rep(0,12), rep(-1,6))
fert2 <- c(rep(0,6), rep(1,6), rep(0,6), rep(-1,6))
fert3 <- c(rep(0,12), rep(1,6), rep(-1,6))
ferts <- cbind(fert1, fert2, fert3)
ferts

##      fert1 fert2 fert3
## [1,]      1      0      0
## [2,]      1      0      0
## [3,]      1      0      0
## [4,]      1      0      0
## [5,]      1      0      0
## [6,]      1      0      0
```

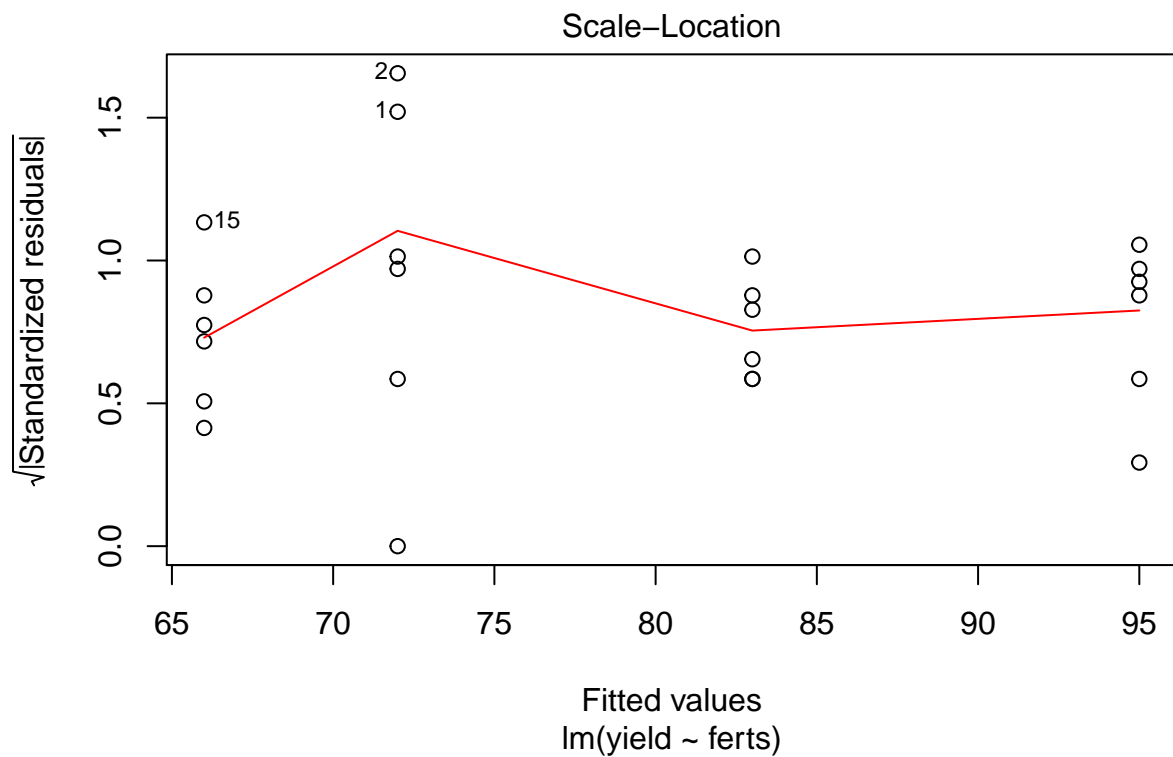
```
## [7,] 0 1 0
## [8,] 0 1 0
## [9,] 0 1 0
## [10,] 0 1 0
## [11,] 0 1 0
## [12,] 0 1 0
## [13,] 0 0 1
## [14,] 0 0 1
## [15,] 0 0 1
## [16,] 0 0 1
## [17,] 0 0 1
## [18,] 0 0 1
## [19,] -1 -1 -1
## [20,] -1 -1 -1
## [21,] -1 -1 -1
## [22,] -1 -1 -1
## [23,] -1 -1 -1
## [24,] -1 -1 -1
```

```
corn.lm3a <- lm(yield ~ ferts)
plot(corn.lm3a)
```





```
## hat values (leverages) are all = 0.1666667
## and there are no factor predictors; no plot no. 5
```



```
anova(corn.lm3a)
```

```
## Analysis of Variance Table
```

```
##
## Response: yield
##           Df Sum Sq Mean Sq F value    Pr(>F)
## ferts      3   2940    980.0   5.9902 0.004387 **
## Residuals 20   3272    163.6
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
summary(corn.lm3a)
```

```
##
## Call:
## lm(formula = yield ~ ferts)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -32.00  -7.50   0.50   8.25  27.00
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)   79.000      2.611  30.258 < 2e-16 ***
## fertsfert1    -7.000      4.522  -1.548  0.13732
## fertsfert2    16.000      4.522   3.538  0.00206 **
## fertsfert3   -13.000      4.522  -2.875  0.00937 **
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 12.79 on 20 degrees of freedom
## Multiple R-squared:  0.4733, Adjusted R-squared:  0.3943
## F-statistic:  5.99 on 3 and 20 DF,  p-value: 0.004387
```

By default, R gives explicit parameters for all but the last of the four groups. If we are not happy with this approach, we can force R to use a different group as the reference group:

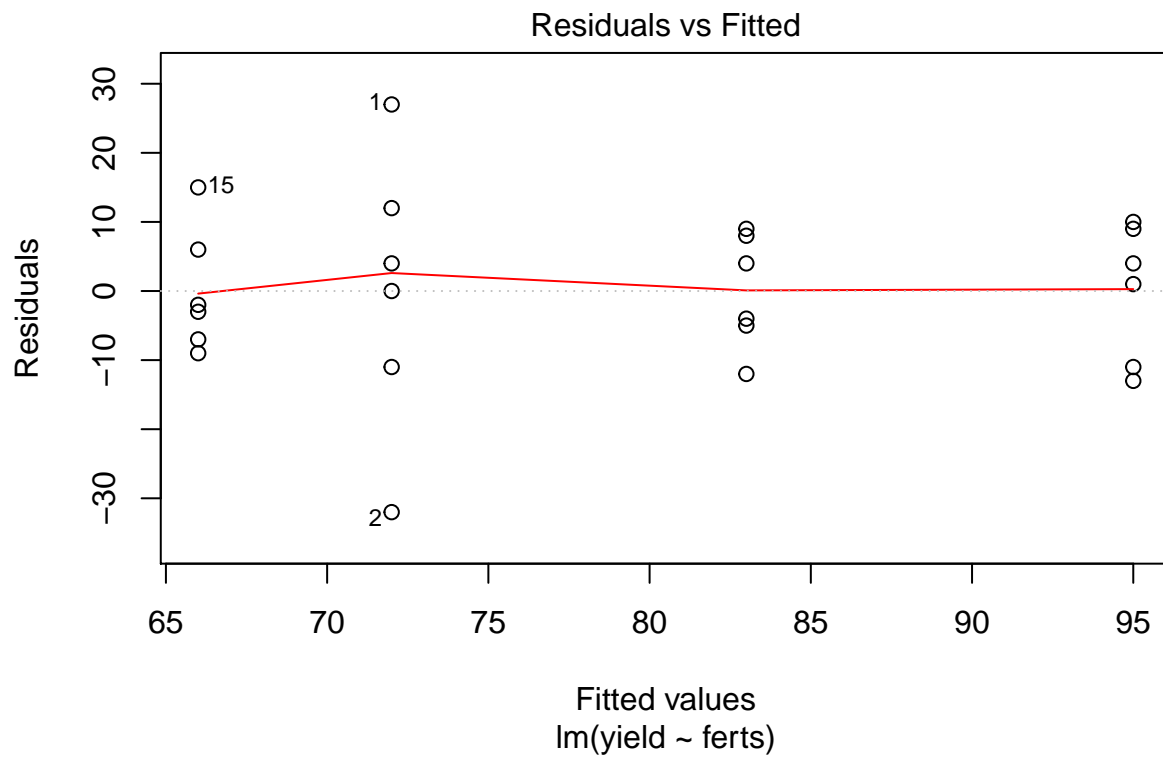
```
fert2 <- c(rep(-1,6), rep(1,6), rep(0,12))
fert3 <- c(rep(-1,6), rep(0,6), rep(1,6), rep(0,6))
fert4 <- c(rep(-1,6), rep(0,12), rep(1,6))
ferts <- cbind(fert2, fert3, fert4)
ferts
```

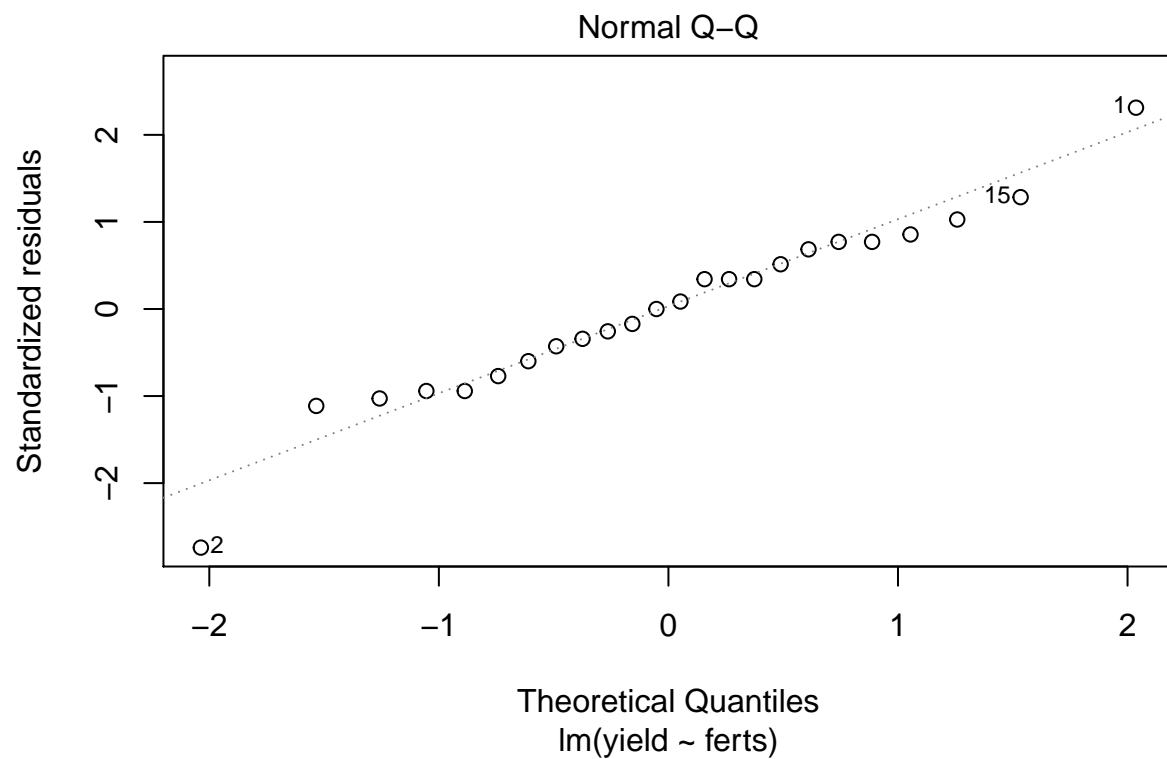
```
##      fert2 fert3 fert4
## [1,]    -1    -1    -1
## [2,]    -1    -1    -1
## [3,]    -1    -1    -1
## [4,]    -1    -1    -1
## [5,]    -1    -1    -1
## [6,]    -1    -1    -1
## [7,]     1     0     0
## [8,]     1     0     0
## [9,]     1     0     0
## [10,]    1     0     0
## [11,]    1     0     0
## [12,]    1     0     0
## [13,]     0     1     0
## [14,]     0     1     0
## [15,]     0     1     0
```



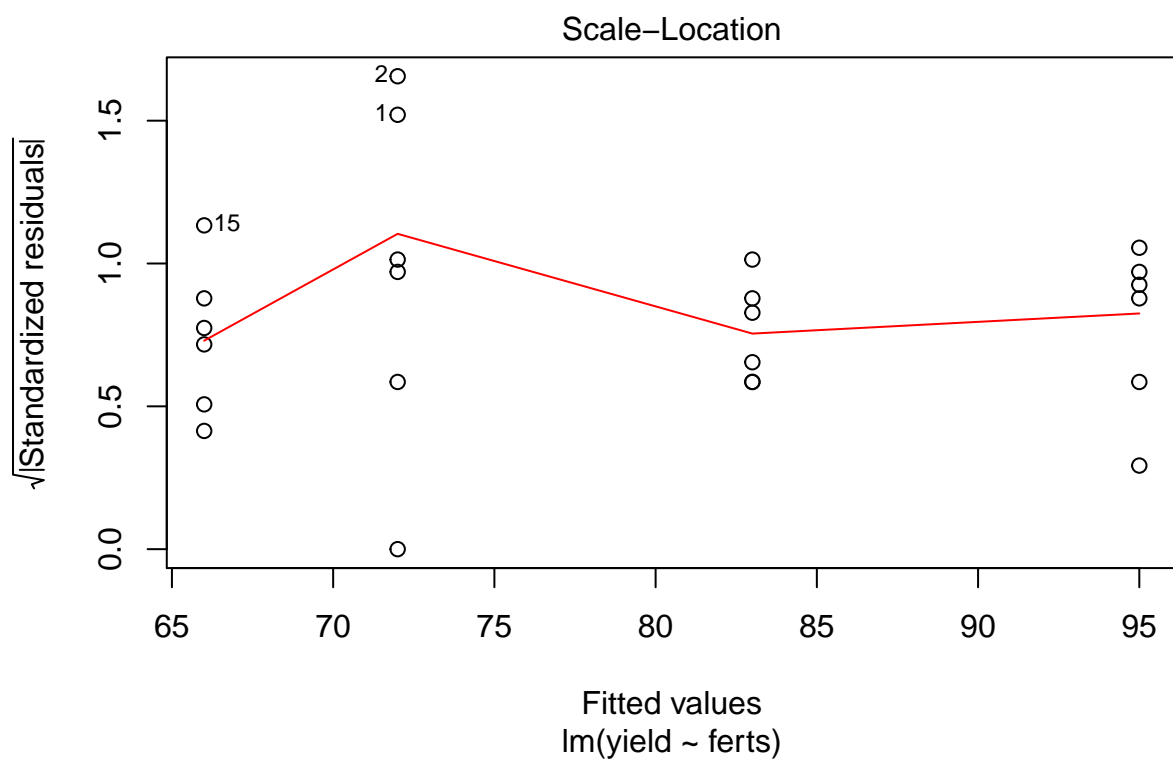
```
## [16,]    0    1    0
## [17,]    0    1    0
## [18,]    0    1    0
## [19,]    0    0    1
## [20,]    0    0    1
## [21,]    0    0    1
## [22,]    0    0    1
## [23,]    0    0    1
## [24,]    0    0    1
```

```
corn.lm3b <- lm(yield ~ ferts)
plot(corn.lm3b)
```





```
## hat values (leverages) are all = 0.1666667
## and there are no factor predictors; no plot no. 5
```



```
anova(corn.lm3b)
```

```
## Analysis of Variance Table
##
## Response: yield
##           Df Sum Sq Mean Sq F value    Pr(>F)
## ferts      3   2940    980.0   5.9902 0.004387 **
## Residuals 20   3272    163.6
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
summary(corn.lm3b)
```

```
##
## Call:
## lm(formula = yield ~ ferts)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -32.00  -7.50   0.50   8.25  27.00
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)    79.000     2.611  30.258 < 2e-16 ***
## fertsfert2     16.000     4.522   3.538 0.00206 **
## fertsfert3    -13.000     4.522  -2.875 0.00937 **
## fertsfert4      4.000     4.522   0.885 0.38692
```

```
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 12.79 on 20 degrees of freedom
## Multiple R-squared:  0.4733, Adjusted R-squared:  0.3943
## F-statistic:  5.99 on 3 and 20 DF,  p-value: 0.004387

# Note that all the above models produce identical plots and ANOVA tables - they are basically
# all the same model, all of which use essentially the same information coded in different ways.
# The differences between the models lie in how we use the different parameters to recover
# information about the level means.

# Is this model really appropriate? The outliers in the control group appear to be inflating
# the variance of that group, which may cause problems as we are assuming that the variance is
# constant, ie. the same for all four groups. The rule of thumb for assessing this is on page 4
# of the brick, the ratio of the largest to smallest group variance should not be greater than
# the number of groups:

lvl.vars <- tapply(yield, fert, var)
lvl.vars

## Control      K20+N K20+P205    N+P205
##    406.8      97.6     80.8     69.2

max(lvl.vars)

## [1] 406.8

min(lvl.vars)

## [1] 69.2

max(lvl.vars)/min(lvl.vars)

## [1] 5.878613

# There are only 4 groups, so we have a problem (one which is simply ignored in the analysis in
# the brick). If we do decide to include these outliers, we should certainly qualify any
# conclusions we make based on this analysis, as an important underlying assumption has been
# violated. Here the presence of two outliers in different directions may just inflate the
# overall estimate of the error variance, which would inflate the standard errors and therefore
# reduce the precision of any contrasts between the treatment groups, making our inference
# more conservative.

anova(corn.lm2)

## Analysis of Variance Table
##
## Response: yield
##          Df Sum Sq Mean Sq F value    Pr(>F)
## fert      3   2940    980.0   5.9902 0.004387 **
## Residuals 20   3272    163.6
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

# The ANOVA table indicates a strong difference between the groups, but not which fertilizer
# treatments produce the best yields. Here are some alternative ways of doing the analysis
# shown at the top of page 8 of the brick:
```

```
summary(corn.lm2)
```

```
##
## Call:
## lm(formula = yield ~ fert, contrasts = list(fert = contr.treatment))
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -32.00  -7.50   0.50   8.25  27.00
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)    72.000     5.222  13.788 1.13e-11 ***
## fert2          23.000     7.385   3.115  0.00546 **
## fert3          -6.000     7.385  -0.812  0.42607
## fert4          11.000     7.385   1.490  0.15194
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 12.79 on 20 degrees of freedom
## Multiple R-squared:  0.4733, Adjusted R-squared:  0.3943
## F-statistic:  5.99 on 3 and 20 DF,  p-value: 0.004387
```

```
# The basic results appear to be that fertilizer 1 "K20 + N", produces significantly
# higher corn yields than the "Control" group (which is presumably some default
# fertilizer regime or possibly no added fertilizer at all), whilst the other two
# groups do not produce significantly different yields from the "Control" group.
```

```
# These conclusions do not change, even if we apply an appropriate Bonferroni correction
# to account for the fact that are performing a series of 3 comparisons (i.e. we have
# to somehow overcome the problem of multiple comparisons):
```

```
names(summary(corn.lm2))
```

```
## [1] "call"          "terms"          "residuals"      "coefficients"
## [5] "aliases"        "sigma"           "df"              "r.squared"
## [9] "adj.r.squared" "fstatistic"      "cov.unscaled"
```

```
summary(corn.lm2)$sigma
```

```
## [1] 12.79062
```

```
se <- summary(corn.lm2)$sigma*sqrt(1/6 + 1/6)
se
```

```
## [1] 7.384669
```

```
est <- coef(corn.lm2)[2:4]
est
```

```
## fert2 fert3 fert4
##    23    -6    11
```

```
lower <- est - qt(1 - 0.05/(2*3), 20)*se
lower
```

```
##      fert2      fert3      fert4
##  3.706922 -25.293078 -8.293078
```

```

upper <- est + qt(1 - 0.05/(2*3), 20)*se
upper

##      fert2      fert3      fert4
## 42.29308 13.29308 30.29308

cbind(lower, est, upper)

##           lower est      upper
## fert2    3.706922 23 42.29308
## fert3   -25.293078 -6 13.29308
## fert4    -8.293078 11 30.29308

# Different research questions would call for different analyses and therefore different
# interpretations of the results. Here is the analysis at the bottom of page 7 of the brick:

ni <- tapply(yield, fert, length)
ni

## Control      K20+N K20+P205      N+P205
##         6         6         6         6

h <- c(-1, 1/3, 1/3, 1/3)
h

## [1] -1.0000000  0.3333333  0.3333333  0.3333333

est <- t(h) %*% lvl.mns
est

##           [,1]
## [1,] 9.333333

MSE <- sum((yield-fitted(corn.lm2))^2)/corn.lm2$df.residual
MSE

## [1] 163.6

se <- sqrt(MSE)*sqrt(sum((h^2)/ni))
lower <- est - qt(0.975, corn.lm2$df.residual)*se
upper <- est + qt(0.975, corn.lm2$df.residual)*se
c(lower, est, upper)

## [1] -3.244102  9.333333 21.910768

# Finally, here's the analysis from the bottom of page 8. Please read the discussion of
# these results in the brick.

h <- c(0, 0.5, -1, 0.5)
est <- t(h) %*% lvl.mns
se <- sqrt(MSE)*sqrt(sum((h^2)/ni))
lower <- est - qt(0.975, corn.lm2$df.residual)*se
upper <- est + qt(0.975, corn.lm2$df.residual)*se
c(lower, est, upper)

## [1]  9.659615 23.000000 36.340385

```