MATH6222 Week 9 Lecture Notes

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2017-05-01

1 Monday's Lecture

1.1 equivalent relations

Proposition: If R is an equivalent relation often write $x \sim y$ to mean $(x,y) \in R$.

If "~" is an equivalent relation, then

- 1. If $x \sim y, [x] = [y]$.
- 2. If $x \not\sim y$, $[x] \cap [y] = \emptyset$
- 3. The distinct equivalence classes partition S, i.e. every element of S belongs in exactly one equivalence class.

Proof:

- 1. Suppose $z \in [x]$, i.e. $z \sim x$. We are given $x \sim y$. By symmetry, $z \sim y$, i.e. $z \in [y]$. This shows $[x] \subseteq [y]$, same argument in reverse shows $[y] \subseteq [x]$. Therefore, [x] = [y].
- 2. Suppose $[x] \cap [y] \neq \emptyset$, i.e. $\exists z \in [x] \cap [y]$, i.e. $z \sim x$ and $z \sim y$. By symmetry, $x \sim z$. By transitivity $x \sim y$. Contradiction.
- 3. Given any $x \in S$. By reflexivity, $x \in [x]$. So every element of S is in at least one equivalence class. Suppose x was contained in two equivalence classes, say [y] and [z]. Then $x \in [y] \cap [z] \implies y \sim z \implies [y] = [z]$.

1.2 Probability

Definition: a (finite) probability space is a finite set S together with a function $P: \{\text{subsets of } S\} \to [0,1] \text{ satisfying:}$

- 1. P(S) = 1
- 2. If $A, B \in S$ and $A \cap B = \emptyset$, then $P(A \cup B) = P(A) + P(B)$.

We call a subset $A \subseteq S$ an "event". P(A) is the probability of the event. If $A \cap B = \emptyset$, we say the events are mutually exclusive.

Example: Suppose S any finite set, define P by $P(A) = \frac{|A|}{|S|}$.

E.g. $A = \{(i, j) : 1 \le i, j \le 6\}, (|A| = 36).$

This P defines the usual probability space for pairs of die.

Remark: Alternative definition. Finite set S together with a function $P: S \to [0,1]$ satisfying $\sum P(a) = 1, a \in S$. Define $P(A) = \sum_{a \in A} P(a)$.

1.3 Conditional probability

Pick a jar at random, pick a marble out of the jar. It's black. What is the probability that I have picked for number 3? ...

Definition: Let A, B be two events in a probability space. The "probability of A given B"

$$P(A|B) = \frac{P(A \cap B)}{P(B)}$$

Bayesian Hypothesis Testing: Suppose $B_1 \ldots, B_r$ are mutually exclusive events which partition a finite probability space. Suppose wee know $P(B_i) = b_i$. Given another event $A \subseteq S$. Suppose we know $a_i = P(A|B_i)$ for each $i = 1, \ldots, r$.

Problem: Determine $P(B_i|A)$. B_i = We've picked for i, $b_i = \frac{1}{3}$. A = Picking a black marble $P(A|B_1) = 0$, $P(A|B_2) = \frac{1}{2}$, $P(A|B_3 = 1$. I asked $P(B_i|A)$, so $P(B_1|A) = 0$, $P(B_2|A) = \frac{1}{3}$, $P(B_3|A) = \frac{2}{3}$.

2 Thursday's Lecture

$$P(A|B) = \frac{P(A \cap B)}{P(B)}$$

Suppose there are mutually exclusive events B_1, \ldots, B_r which partition S. Then suppose there's another event A. We know $b_i = P(B_i)$. Problem: Determine $P(B_i|A) = b_i^*$.

Bayes Formula:

$$b_i^* = P(B_i|A) := \frac{a_i b_i}{\sum_{j=1}^r a_j b_j}$$

Proof:

$$P(B_i|A) = \frac{P(B_i \cap A)}{P(A)}$$

$$P(B_i \cap A) = P(A \cap B_i) = P(A|B_i) \cdot P(B_i) = a_i b_i$$

$$A = (A \cap B_1) \cup (A \cap B_2) \cup \dots \cup (A \cap B_r)$$

$$P(A) = \sum_{j=1}^r P(A \cap B_j) = \sum_{j=1}^r a_j b_j$$

Medical Testing: Suppose you have a test for condition X, which is 96% accurate. Question: If you test positive, what is the probability that you actually have condition X?

Suppose throughout the general population, the probability of having condition X is 1% (99% healthy).

A: testing positive +

 B_1 : you have condition $X, b_1 = P(B_1) = 0.01$

 B_2 : you don't have condition $X, b_2 = P(B_2) = 0.99$

$$a_1 = P(A|B_1) = 0.96$$

$$a_2 = P(A|B_2) = 0.04$$

$$P(B_1|A) = \frac{P(B_1 \cap A)}{P(A)} = \frac{a_1b_1}{a_1b_1 + a_2b_2} = \frac{.96 \cdot .01}{.96 \cdot .01 + .99 \cdot .04} = 19.59\%$$

Random variables and Expectation: a random variable on a probability space S is just a function $X: S \to R$.

The **expectation** of the random variable X is $E(X) := \sum_{a \in S} X(a)P(a)$

Simple Examples: $S = \{\text{people in our class}\}, X : S \to R \text{ (person } \to \text{weight)}, |S| = N$

$$E(X) = \sum_{\text{people in class}} \frac{1}{N} \text{(height of person)} = \text{average height in the class}$$

Given a random variable X on probability space S, we can define the probability $P(X = k) = \sum_{a \in S, \text{ such that } X(a) = k} P(a) = P(\{a \in S : X(a) = k\}).$

$$E(X) = \sum_{k} k \cdot P(X = k)$$

Suppose $S = \{ \text{Sequences of length 10 flips} : HT \dots \}, (|S| = 2^{10}).$ $X = \{ \text{any length 10 sequence of H T} \rightarrow \text{number of heads} \}$

$$E(X) = \sum_{k=0}^{10} k \cdot P(X = k) = \sum_{k=0}^{10} k \cdot \binom{n}{k} \frac{1}{2^{10}} = \frac{1}{2^{10}} \sum_{k=0}^{10} k \cdot \binom{n}{k}$$

3 Friday's Lecture

 $S = \{\text{Length n sequences of } \{H, T\}\}, |S| = 2^n.$ Define X on S by taking the number of heads in any given sequences.

$$E(X) = \sum_{k=0}^{n} k \cdot P(X = k) = \sum_{k=0}^{n} k \binom{n}{k} 2^{-n} = \frac{n}{2}$$

Linearity of Expectation: Let X_1, \ldots, X_n be random variables on probability space S. Let $c_1, \ldots, c_n \in \mathbb{R}$. Let $X = c_1 X_1 + c_2 X_2 + \cdots + c_n X_n$. Then $E(X) = c_1 E(X_1) + c_2 E(X_2) + \cdots + c_n E(X_n)$.

Proof:

$$E(X) = E(c_1X_1 + \dots + c_nX_n)$$

$$= \sum_{a \in S} (c_1X_1(a) + c_2X_2(a) + \dots + c_nX_n(a)) \cdot P(a)$$

$$= \sum_{a \in S} c_1X_1(a)P(a) + \dots + c_nX_nP(a)$$

$$= \sum_{a \in S} c_1X_1(a)P(a) + \sum_{a \in S} c_2X_2(a)P(a) + \dots + \sum_{a \in S} c_nX_n(a)P(a)$$

$$= c_1\sum_{a \in S} X_1(a)P(a) + \dots + c_n\sum_{a \in S} X_n(a)P(a)$$

$$= c_1E(X_1) + \dots + c_nE(X_n)$$

$$X_i = \begin{cases} 1, & \text{if the ith toss is heads} \\ 0, & \text{if the ith toss is tails} \end{cases}$$

Note: $i = 1, ..., n, X = X_1 + \cdots + X_n, E(X) = \sum_{i=1}^n E(X_i) = \frac{n}{2}$. As $E(X_1) = 1 \cdot P(X_1 = 1) + 0 \cdot P(X_1 = 0) = \frac{1}{2}$.

Problem: So n pairs of socks thrown into a laundry machine. Machine spits out a random subset of k socks. How many complete pairs of socks do we expect come out?

Outcomes = $\binom{2n}{k}$, $\{L_1, R_1, \dots, L_n, R_n\}$

X is defined as the number of complete pairs in a particular subset.

$$X_i = \begin{cases} 1, & \text{if the } \{L_i, R_i\} \text{ comes out} \\ 0, & \text{otherwise} \end{cases}$$

We need to count subsets of size $k \subset \{L_1, R_1, L_2, R_2, \dots, L_n, R_n\}$, which is $\binom{2n-2}{k-2}$. So

$$E(X_1) = 1 \cdot P(X_1 = 1) + 0 \cdot P(X_1 = 0) = \frac{\binom{2n-2}{k-2}}{\binom{2n}{k}}$$

Then

$$E(X) = \sum_{i=1}^{n} E(X_i) = \frac{n\binom{2n-2}{k-2}}{\binom{2n}{k}}$$

Finger Game: A and B can hold up 1 or 2 fingers. If the total is odd, A wins; if the total is even, B wins. Whoever wins, the losing side has to pay the amount of money of the number of fingers.

- 1, 1, A pays B 2 dollars.
- 1, 2, B pays A 3 dollars.
- 2, 1, B pays A 3 dollars.
- 2, 2 A pays B 4 dollars.

Suppose A, B play randomly, each scenario has a probability of $\frac{1}{4}$. $E(A) = \frac{1}{4}(-2) + \frac{1}{4}(3) + \frac{1}{4}3 + \frac{1}{4}(-4) = 0$. Seems fair.

A chooses

- 1. with probability x
- 2. with probability 1-x

B choose

1. with probability y

2. with probability 1 - y

A wants to choose x to maximize expected pay-off.

Then the probabilities of the 4 scenarios above become xy, (1-x)y, x(1-y), (1-x)(1-y).

$$E(X) = -2xy + 3(1-x)y + 3(1-y)x - 4(1-x)(1-y)$$

$$= -12xy + 7x + 7y - 4$$

$$= (7-12x)y + (7x - 4)$$

$$(x = \frac{7}{12}) = 7 \cdot \frac{7}{12} - 4 = \frac{1}{12} > 0$$

If $x = \frac{7}{12}$, we can make y irrelevant.

Suppose we look this from B's angle, then = (7-12y)x+(7y-4), $y=\frac{7}{12}$.

When A takes $x = \frac{7}{12}, 1 - x = \frac{5}{12}$. B holding 1 all the time! Then

$$E(A) = \frac{7}{12}(-2) + \frac{5}{12}(3) = \frac{1}{12}$$

If B holding 2 all the time!

$$E(A) = \frac{7}{12}(3) + \frac{5}{12}(-4) = \frac{1}{12}$$

No matter what B does here, A's expectation should always be $\frac{1}{12}$.

If
$$x < \frac{7}{12}$$
, B chooses $y = 0$, $(7x - 4) < \frac{1}{12}$.
If $x > \frac{7}{12}$, B chooses $y = 1$, $7 - 12x + 7x - 4 = 3 - 5x < \frac{1}{12}$.