Machine Learning

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Abstract

The course will cover several topics, starting with an introduction to basic concepts. Learning theory will be explored, including the bias-variance tradeoff, Union and Chernoff/Hoeffding bounds, VC dimension, worst-case (online) learning, and practical advice on using learning algorithms effectively.

In supervised learning, key areas of focus include the supervised learning setup, LMS, logistic regression, perceptron, the exponential family, and kernel methods such as Radial Basis Networks, Gaussian Processes, and Support Vector Machines. Additionally, topics like model selection, feature selection, ensemble methods (e.g., bagging and boosting), and strategies for evaluating and debugging learning algorithms will be addressed.

The course will also delve into reinforcement learning and control, examining Markov Decision Processes (MDPs), Bellman equations, value iteration, policy iteration, TD, SARSA, Q-learning, value function approximation, policy search, REINFORCE, POMDPs, and the Multi-Armed Bandit problem.

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Introduction

1.1 Machine Learning

Definition (Learning). A computer program is said to learn from experience E with respect to some class of tasks T and performance measure P if it improves with experience E.

Machine Learning derives knowledge from experience and induction.

In Machine Learning, we depend on computers to make informed decisions using new, unfamiliar data. Designing a comprehensive set of meaningful rules can prove to be exceedingly difficult. Machine Learning facilitates the automatic extraction of relevant insights from historical data and effectively applies them to new datasets.

The objective is to automate the programming process for computers, acknowledging the bottleneck presented by writing software. Instead, our aim is to utilize the data itself to accomplish the required tasks.

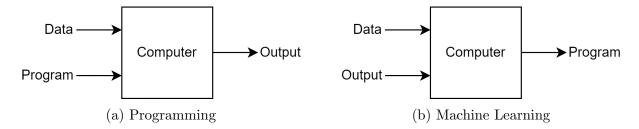


Figure 1.1: Difference between programming and Machine Learning

Machine Learning paradigms can be categorized into three main types:

- Supervised learning: involves labeled data and direct feedback, aiming to predict outcomes or future events.
- Unsupervised learning: operates without labeled data or feedback, focusing on discovering hidden structures within the data.
- Reinforcement learning: centers around a decision-making process, incorporating a reward system to learn sequences of actions.

1.2 Supervised learning

Supervised learning encompasses several distinct tasks:

- Classification: this involves assigning predefined categories or labels to data points based on their features. The model is trained on labeled data, learning patterns to predict the class labels of new data points.
- Regression: the goal here is to predict continuous numerical values based on input features, as opposed to discrete class labels in classification. The model learns a function mapping input features to output values.
- *Probability estimation*: this task predicts the likelihood of certain events or outcomes occurring, often used to gauge the confidence of model predictions.

Formally, in supervised learning, a model learns from data to map known inputs to known outputs. The training set is denoted as $\mathcal{D} = \{\langle x, t \rangle\}$, where t = f(x), with f representing the unknown function to be determined using supervised learning techniques.

Various techniques can be employed for supervised learning, including linear models, artificial neural networks, support vector machines, and decision trees.

1.3 Unsupervised learning

Unsupervised learning encompasses two main tasks:

- Clustering: in this task, the objective is to group similar data points together based on their features, without predefined labels. The goal is to uncover underlying patterns or structures within the data. Clustering algorithms segment the data into clusters or groups, where data points within the same cluster exhibit greater similarity compared to those in different clusters.
- Dimensionality reduction: this task involves reducing the number of input variables or features in a dataset while retaining essential information. This is often done to address the curse of dimensionality, enhance computational efficiency, and mitigate overfitting risks in models. Dimensionality reduction techniques aim to transform high-dimensional data into a lower-dimensional representation while preserving most relevant information.

Formally, in unsupervised learning, computers learn previously unknown patterns and efficient data representations. The training set is defined as $\mathcal{D} = \{x\}$, where the goal is to find a function f that extracts a representation or grouping of the data.

Various techniques are used for unsupervised learning, including k-means clustering, self-organizing maps, and principal component analysis.

1.4 Reinforcement learning

Reinforcement learning encompasses several key approaches:

• Markov Decision Process: a mathematical framework for modeling decision-making, involving states, actions, transition probabilities, and rewards. The goal is to find a policy that maximizes cumulative rewards while considering uncertainty.

- Partially Observable MDP: an extension of MDP where the current state is uncertain and must be inferred from observations. The objective remains the same, but the agent maintains a belief over possible states based on observations.
- Stochastic games: models for decision-making with multiple agents, where outcomes depend on actions and random factors. Players aim to optimize strategies considering other players' actions and uncertainties.

In reinforcement learning, the computer learns the optimal policy based on a training set \mathcal{D} containing tuples $\langle x, u, x', r \rangle$, where x is the input, u is the action, x' is the resulting state after the action, and r is the reward. The policy Q^* is defined to maximize $Q^*(x, u)$ over actions u for each state x in the training set.

Various techniques such as Q-learning, SARSA, and fitted Q-iteration are used to find this optimal policy.

Supervised learning

2.1 Introduction

Supervised learning stands as the predominant and well-established learning approach. Its core objective is to enable a computer, given a training set $\mathcal{D} = \{\langle x, t \rangle\}$, to approximate a function f that maps an input x to an output t. The input variables x, often referred to as features or attributes, are paired with output variables t, also known as targets or labels. The tasks undertaken in supervised learning are as follows:

- Classification: when t is discrete.
- Regression: when t is continuous.
- Probability estimation: when trepresents a probability.

Supervised learning finds application in scenarios where:

- Humans lack the capability to perform the task directly (e.g., DNA analysis).
- Humans can perform the task but lack the ability to articulate the process (e.g., medical image analysis).
- The task is subject to temporal variations (e.g., stock price prediction).
- The task demands personalization (e.g., movie recommendation).

2.1.1 Function approximation

The process of approximating a function f from a dataset \mathcal{D} involves several steps:

- 1. Define a loss function \mathcal{L} : this function calculates the discrepancy between f and h, a chosen approximation.
- 2. Select a hypothesis space \mathcal{H} : this space consists of a set of candidate functions from which to choose an approximation h.
- 3. Minimize \mathcal{L} within \mathcal{H} : the goal is to find an approximation h within the hypothesis space \mathcal{H} that minimizes the loss function \mathcal{L} .

The hypothesis space \mathcal{H} can be expanded to theoretically achieve a perfect approximation of the function f. However, a significant challenge arises because the loss function \mathcal{L} cannot be easily determined, primarily due to the absence of the actual function f.

2.1.2 Taxonomy

The taxonomy is as follows:

- Parametric or nonparametric: parametric methods are characterized by having a fixed and finite number of parameters, while nonparametric methods have a number of parameters that depend on the training set.
- Frequentist or Bayesian: frequentist approaches utilize probabilities to model the sampling process, whereas Bayesian methods use probability to represent uncertainty about the estimate.
- Empirical risk minimization or structural risk minimization: empirical risk refers to the error over the training set, while structural risk involves balancing the training error with model complexity.

The type of Machine Learning can be:

- Direct: this method involves learning an approximation of f directly from the dataset \mathcal{D} .
- Generative: in this approach, the model focuses on modeling the conditional density Pr(t|x) and then marginalizing to find the conditional mean:

$$\mathbb{E}\left[t|x\right] = \int t \Pr(t|x) dt$$

• Discriminative: this method models the joint density Pr(x,t), infers the conditional density Pr(t|x), and then marginalizes to find the conditional mean:

$$\mathbb{E}\left[t|x\right] = \int t \Pr(t|x) dt$$

2.2 Linear regression

The goal of regression is to approximate a function $f(\mathbf{x})$ that maps input \mathbf{x} to a continuous output t from a dataset \mathcal{D} :

$$\mathcal{D} = \{ \langle \mathbf{x}, t \rangle \} \implies t = f(\mathbf{x})$$

To perform regression, we assume the existence of a function capable of performing this mapping.

In linear regression, the function $f(\cdot)$ is modeled using linear functions. This choice is motivated by several factors:

- Linear models are easily interpretable, making them suitable for explanation.
- Linear regression problems can be solved analytically, allowing for efficient computation.
- Linear functions can be extended to model nonlinear relationships.

• More sophisticated methods often build upon or incorporate elements of linear regression.

The key components of constructing a linear regression problem include:

• Hypothesis space: the mapping function can be defined as:

$$y(\mathbf{x}, \mathbf{w}) = w_0 + \sum_{j=1}^{D-1} w_j x_j = w_0 1 + \sum_{j=1}^{D-1} w_j x_j = \sum_{j=0}^{D-1} w_j x_j = \mathbf{w}^T \mathbf{x}$$

The parameter $w_0 = -b$ is called bias parameter. In a two-dimensional space, our hypothesis space will be the set of all points in the plane (w_0, w_1) . The coordinates of each point will correspond to a line in the (\mathbf{x}, y) space.

• Loss function: we usually employ the Sum of Squared Errors:

$$SSE(\mathbf{w}) = \frac{1}{2} \sum_{n=1}^{N} (y(x_n, \mathbf{w}) - t_n)^2 = \frac{1}{2} \sum_{n=1}^{N} (\phi(x_n) - t_n)^2 = RSS(\mathbf{w}) = \sum_{i=1}^{N} \epsilon_i^2$$

• Optimization: a closed-form optimization of the RSS, known as least squares, begins with the matrix representation of the loss function:

$$LS(\mathbf{w}) = \frac{1}{2}RSS(\mathbf{w}) = \frac{1}{2}(\mathbf{\Phi}_{\mathbf{w}} - \mathbf{t})^{2}$$

To find the optimal \mathbf{w} , we compute the first derivative of $\mathrm{LS}(\mathbf{w})$ and set it to zero, obtaining:

$$\hat{\mathbf{w}}_{\mathrm{LS}} = \left(\mathbf{\Phi}^T \mathbf{\Phi}\right)^{-1} \mathbf{\Phi}^T \mathbf{t}$$

The inversion of the matrix can be computationally expensive, especially for large datasets, assuming the matrix is non-singular (invertible).

To mitigate this, stochastic gradient descent can be employed. The algorithm known as Least Mean Squares (LMS) uses the following update rule:

$$\mathbf{w}^{(n+1)} = \mathbf{w}^{(n)} - \alpha \left(\mathbf{w}^{(n)} \phi(\mathbf{x}_n) - t_n \right) \phi(\mathbf{x}_n)$$

The same update rule can be also applied for batches of size K:

$$\mathbf{w}^{(n+1)} = \mathbf{w}^{(n)} - \frac{\alpha}{K} \left(\mathbf{w}^{(n)} \phi(\mathbf{x}_n) - t_n \right) \phi(\mathbf{x}_n)$$

Multiple outputs If the regression problem involves multiple outputs, meaning that \mathbf{t} is not a scalar, we can solve each regression problem independently. The solution for the weight vectors for all outputs can be expressed as:

$$\hat{\mathbf{W}} = \left(\mathbf{\Phi}^T\mathbf{\Phi}
ight)^{-1}\mathbf{\Phi}^T\mathbf{T}$$

This solution can be easily decoupled for each output k:

$$\hat{\mathbf{w}}_k = \left(\mathbf{\Phi}^T\mathbf{\Phi}\right)^{-1}\mathbf{\Phi}^T\mathbf{t}_k$$

An advantage of this approach is that $(\Phi^T \Phi)^{-1}$ only needs to be computed once, regardless of the number of outputs.

2.2.1 Basis functions

While a linear combination of input variables may not always suffice to model data, we can still construct a regression model that is linear in its parameters. This can be achieved by defining a model using non-linear basis functions, expressed as:

$$y(\mathbf{x}, \mathbf{w}) = \mathbf{w}^T \boldsymbol{\phi}(\mathbf{x})$$

Here, the components of the vector $\phi(\mathbf{x})$ are referred to as features. These features allow for a more flexible representation of the input data, enabling the model to capture non-linear relationships between the input variables and the output.

Example:

Let's reconsider a set of data regarding individuals' weight and height, along with their completion times for a one-kilometer run:

Height (cm)	Weight (kg)	Completion time (s)
180	70	180
184	80	220
174	60	170

We can model this problem using a dummy variable and introduce the Body Mass Index (BMI) as a new feature:

Dummy variable	Height (cm)	Weight (kg)	$\mid \mathbf{BMI} \mid$	Completion time (s)
x_0	x_1	x_2	x_3	t
1	180	70	21	180
1	184	80	23	220
1	174	60	20	170

Here, the dummy variable x_0 is always initialized to one. Now, we have the option to retain or discard the weight and height variables, considering only the BMI values for analysis.

The most commonly used basis functions in regression are:

Basis function	Formula
Polynomial	$\phi_j(x) = x^j$
Gaussian	$\phi_j(x) = e^{-\frac{\left(x - \mu_j\right)^2}{2\sigma^2}}$
Sigmoidal	$\phi_j(x) = \frac{1}{1 + e^{\frac{\mu_j - x}{\sigma}}}$

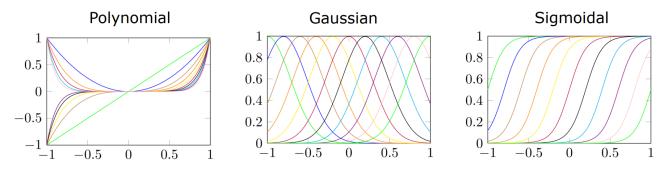


Figure 2.1: Polynomial, Gaussian, and sigmoidal basis functions

It's noteworthy that the Gaussian basis function allows for a local approximation by omitting values that are close to zero. This approach enables capturing the relationship between the input and output in a reduced input space area. As we move away from the mean, approaching zero, the values become negligible.

2.2.2 Normalization

Given a set of N samples, $\{s_1, \ldots, s_N\}$, normalization can be performed using two common methods:

• Z-score (normalization): scales the data based on the dataset's mean and standard deviation.. Given the mean \bar{s} and the variance $S^2 = \frac{1}{N-1} \sum_{n=1}^{N} (s_n - \bar{s})^2$, the normalized value of a sample s is calculated as:

$$s_{\text{z-score}} = \frac{s - \bar{s}}{S}$$

This method transforms the data into a distribution with a mean of 0 and a standard deviation of 1, making it useful when working with data that needs to be compared across different scales or distributions.

• Minmax (feature scaling): rescales the data so that all values lie between a defined range, typically $\begin{bmatrix} 0 & 1 \end{bmatrix}$. Given the minimum value s_{\min} and maximum value s_{\max} in the dataset, the normalized value of a sample s is:

$$s_{\rm Min\text{-}max} = \frac{s - s_{\rm min}}{s_{\rm max} - s_{\rm min}}$$

This method is particularly useful when the data needs to be transformed to a bounded range.

Both methods have their applications, with z-score normalization being more effective for data with outliers or differing variances, and Min-Max scaling suited for data that needs to be normalized to a specific range.

2.2.3 Regularization

A function can achieve a better approximation by increasing the degree of the polynomial used in the regression. However, increasing the polynomial degree also increases the complexity of the model parameters. To address this complexity, adjustments are needed in the loss function:

$$\mathcal{L}(\mathbf{w}) = \mathcal{L}_D(\mathbf{w}) + \lambda \mathcal{L}_W(\mathbf{w})$$

Here, $\mathcal{L}_D(\mathbf{w})$ represents the usual loss function, $\mathcal{L}_W(\mathbf{w})$ reflects model complexity (a hyper-parameter), and λ is the regularization coefficient. The model complexity loss function can be:

• Ridge, in which the loss function becomes:

$$\mathcal{L}_{\text{ridge}}(\mathbf{w}) = \frac{1}{2} \sum_{i=1}^{N} \epsilon_i^2 + \lambda \frac{1}{2} \|\mathbf{w}\|_2^2$$

This new loss function remains quadratic with respect to \mathbf{w} , allowing for closed-form optimization:

$$\hat{\mathbf{w}}_{\text{ridge}} = \left(\lambda \mathbf{I} + \mathbf{\Phi}^T \mathbf{\Phi}\right)^{-1} \mathbf{\Phi}^T \mathbf{t}$$

The term $\lambda \mathbf{I}$ is crucial in solving the singularity problem, as it transforms a non-singular matrix into a singular one with an appropriate choice of λ . In particular, the eigenvalues of the $(\lambda \mathbf{I} + \mathbf{\Phi}^T \mathbf{\Phi})$ matrix must be greater or equal than λ since $\mathbf{\Phi}^T \mathbf{\Phi}$ is positive semidefinite.

• Lasso, in which the loss function becomes:

$$\mathcal{L}_{\text{lasso}}(\mathbf{w}) = \frac{1}{2} \sum_{i=1}^{N} \epsilon_i^2 + \lambda \frac{1}{2} \|\mathbf{w}\|_1$$

In this case, closed-form optimization is not possible. However, lasso typically leads to sparse regression models: when the regularization coefficient λ is large enough, some components of $\hat{\mathbf{w}}$ become equal to zero.

2.2.4 Model evaluation

The performance of the resulting model can be assessed through various metrics and statistical tests:

- Residual Sum of Squares: measures the discrepancy between the predicted and actual target values. A lower RSS indicates a better fit of the model to the data.
- Mean Square Error: average of the squared differences between the predicted values and the actual values. It is calculated as:

$$MSE(\mathbf{w}) = \frac{RSS(\mathbf{w})}{N}$$

where N is the number of samples. MSE penalizes larger errors more heavily due to the squaring of differences.

• Root Mean Square Error: square root of the MSE, giving an error metric in the same units as the target variable:

$$RMSE(\mathbf{w}) = \sqrt{\frac{RSS(\mathbf{w})}{N}}$$

RMSE is often easier to interpret as it provides an error measure on the same scale as the original data.

• Coefficient of determination: measures how well the model explains the variance in the target variable. It is calculated as:

$$R^2 = 1 - \frac{RSS(\mathbf{w})}{TSS}$$

Here, $TSS = \sum_{n=1}^{N} (\bar{t} - t_n)^2$ is the Total Sum of Squares, and \bar{t} is the mean of the target values. An R^2 close to 1 indicates a good fit, while a value near 0 suggests the model performs poorly compared to a simple mean.

• Degrees of freedom: represent the difference between the number of samples and the number of model parameters:

$$dfe = N - M$$

Here, M is the number of parameters in the model.

• Adjusted coefficient of determination: accounts for the number of predictors in the model and adjusts for the degrees of freedom:

$$R_{\text{adj}}^2 = 1 - (1 - R^2) \frac{N - 1}{\text{dfe}}$$

This metric is useful when comparing models with different numbers of predictors, as it penalizes overfitting.

Statistical tests on coefficients To determine the statistical significance of the model's parameters, hypothesis tests can be performed:

1. Test on single coefficients: this test examines whether each estimated weight \hat{w}_j is significantly different from zero. The distribution for this test is given by:

$$t_{\rm dfe} \sim \frac{\hat{w}_j - w_j}{\sigma \hat{\sqrt{v_j}}}$$

Here, w_j is the true parameter, \hat{w}_j is the estimated parameter, v_j is the j-th diagonal element of $(\mathbf{x}^T\mathbf{x})^{-1}$, and $\hat{\sigma}^2$ is the unbiased estimate of the variance:

$$\hat{\sigma}^2 = \frac{\text{RSS}(\hat{\mathbf{w}})}{\text{dfe}}$$

If the test shows that the coefficient is significantly different from zero, the null hypothesis (that the coefficient is zero) is rejected.

2. Test on overall model significance: this test assesses the significance of the overall model by comparing it to a null model (a model with no predictors). The test uses the Fisher-Snedecor distribution:

$$F_{\mathrm{stat}} \sim \frac{\mathrm{dfe}}{M-1} \frac{\mathrm{TSS} - \mathrm{RSS}(\hat{\mathbf{w}})}{\mathrm{RSS}(\hat{\mathbf{w}})}$$

If the F-statistic is large, the null hypothesis (that all model coefficients are zero) is rejected, indicating that the model significantly improves prediction compared to a constant (mean) model.

2.2.5 Maximum Likelihood

We can approach regression in a probabilistic framework by defining a model that maps inputs to target values probabilistically. This allows us to express uncertainty in the predictions.

Given a regression model denoted by $y(x, \mathbf{w})$, where \mathbf{w} represents the unknown parameters, we assume that the observed data \mathcal{D} is generated with some inherent noise. The model provides the conditional probability of the target given the input, and we express the likelihood of the data \mathcal{D} given the parameters \mathbf{w} as $\Pr(\mathcal{D}|\mathbf{w})$.

To estimate the parameters, we seek to find the set of parameters \mathbf{w} that maximizes this likelihood. This approach is known as Maximum Likelihood Estimation (ML), and the parameters are found by solving the following optimization problem:

$$\mathbf{w}_{\mathrm{ML}} = \operatorname*{argmax}_{\mathbf{w}} \Pr(\mathcal{D}|\mathbf{w})$$

Our probabilistic regression model can be written as:

$$t = y(\mathbf{x}, \mathbf{w}) + \epsilon = \mathbf{w}^T \mathbf{\Phi}(\mathbf{x}) + \epsilon$$

Here, $y(\mathbf{x}, \mathbf{w})$ is assumed to be a linear model in terms of a set of basis functions $\Phi(\mathbf{x})$, with additive noise ϵ that follows a Gaussian distribution with zero mean and variance σ^2 .

Given a dataset \mathcal{D} of N samples with inputs \mathbf{x}_n and targets \mathbf{t}_n , we express the likelihood of the data \mathcal{D} given the model parameters \mathbf{w} as:

$$\Pr(\mathcal{D}|\mathbf{w}) = \Pr(\mathbf{t}|\mathbf{x}, \mathbf{w}, \sigma^2) = \prod_{n=1}^{N} \mathcal{N}(t_n|\mathbf{w}^T \mathbf{\Phi}(\mathbf{x}_n), \sigma^2)$$

Here, $\mathcal{N}(t_n|\mathbf{w}^T\mathbf{\Phi}(\mathbf{x}_n), \sigma^2)$ represents the Gaussian distribution for each target, with mean $\mathbf{w}^T\mathbf{\Phi}(\mathbf{x}_n)$ and variance σ^2 .

To find the maximum likelihood estimate \mathbf{w}_{ML} , we maximize the log-likelihood, which simplifies the product into a sum:

$$\mathcal{L}(\mathbf{w}) = \ln \Pr(t_n | \mathbf{x}_n, \mathbf{w}, \sigma^2) = -\frac{N}{2} \ln(2\pi\sigma^2) - \frac{1}{2\sigma^2} RSS(\mathbf{w})$$

Here, $RSS(\mathbf{w})$ is the Residual Sum of Squares.

The first term, $-\frac{N}{2}\ln(2\pi\sigma^2)$, is independent of \mathbf{w} , so we can ignore it when maximizing the log-likelihood. This leaves us with the second term, which is proportional to the residual sum of squares. Therefore, maximizing the log-likelihood is equivalent to minimizing RSS(\mathbf{w}).

To find \mathbf{w}_{ML} , we set the gradient of $\mathcal{L}(\mathbf{w})$ with respect to \mathbf{w} to zero:

$$\frac{\partial \mathcal{L}(\mathbf{w})}{\partial \mathbf{w}} = 0$$

Solving this yields the closed-form solution for the maximum likelihood estimate of w

$$\mathbf{w}_{\mathrm{ML}} = \left(\mathbf{\Phi}^T \mathbf{\Phi}\right)^{-1} \mathbf{\Phi}^T \mathbf{t}$$

This result matches the solution for the Ordinary Least Squares (OLS) method, showing that the maximum likelihood estimate under the assumption of Gaussian noise is equivalent to minimizing the squared error. The Maximum Likelihood estimate \mathbf{w}_{ML} has the smallest variance among unbiased linear estimators, according to the Gauss-Markov theorem.

2.2.6 Bayesian linear regression

Bayesian linear regression offers a probabilistic framework for modeling linear relationships by incorporating uncertainty about the model parameters, unlike traditional methods that provide only point estimates. In this approach, we treat the model parameters as random variables and update our beliefs about them as more data becomes available. The process is outlined in the following steps:

- 1. Formulation of a probabilistic model: initially, we express our prior knowledge about the model parameters probabilistically, defining a prior distribution that encapsulates assumptions about these parameters before observing any data. This prior reflects what we know or assume about the parameter values based on domain expertise or past experience.
- 2. Data observation: as we collect data, we obtain a likelihood function that measures the probability of observing the data given particular values of the model parameters.
- 3. Posterior distribution calculation: after observing the data, we use Bayes' theorem to compute the posterior distribution, which combines the prior distribution with the likelihood of the data:

$$Pr(params|data) = \frac{Pr(data|params) Pr(params)}{Pr(data)}$$

The posterior distribution provides a refined belief about the model parameters after seeing the data.

4. Prediction and decision making: to make predictions, we use the posterior distribution by averaging over all possible parameter values weighted by their posterior probabilities. This allows for uncertainty in the predictions and enables decisions that minimize expected loss.

In Bayesian linear regression, the posterior distribution is computed by combining the prior with the likelihood of the parameters given the observed data:

$$\Pr(\mathbf{w}|\mathcal{D}) = \frac{\Pr(\mathcal{D}|\mathbf{w})\Pr(\mathbf{w})}{\Pr(\mathcal{D})}$$

Here, $\Pr(\mathbf{w})$ is the prior distribution over the parameters \mathbf{w} , $\Pr(\mathcal{D}|\mathbf{w})$ is the likelihood of the data given the parameters, and $\Pr(\mathcal{D})$ is the marginal likelihood, ensuring normalization:

$$\Pr(\mathcal{D}) = \int \Pr(\mathcal{D}|\mathbf{w}) \Pr(\mathbf{w}) d\mathbf{w}$$

The mode of the posterior distribution is known as the Maximum A Posteriori (MAP) estimate, which gives the most probable parameter values given the data.

Assuming a Gaussian likelihood function allows the use of a conjugate Gaussian prior, which simplifies the Bayesian updating process. The prior is typically modeled as:

$$\Pr(\mathbf{w}) = \mathcal{N}(\mathbf{w}|\mathbf{w}_0, \mathbf{S}_0)$$

Here, \mathbf{w}_0 is the prior mean, and \mathbf{S}_0 is the prior covariance matrix. After observing the data, the posterior remains Gaussian:

$$\begin{cases} \Pr(\mathbf{w}|\mathbf{t}, \mathbf{\Phi}, \sigma^2) = \mathcal{N}(\mathbf{w}|\mathbf{w}_N, \mathbf{S}_N) \\ \mathbf{w}_N = \mathbf{S}_N \left(\mathbf{S}_0^{-1} \mathbf{w}_0 + \frac{\mathbf{\Phi}^T \mathbf{t}}{\sigma^2} \right) \\ \mathbf{S}_N^{-1} = \mathbf{S}_0^{-1} + \frac{\mathbf{\Phi}^T \mathbf{\Phi}}{\sigma^2} \end{cases}$$

Here, \mathbf{w}_N is the posterior mean, and \mathbf{S}_N is the posterior covariance matrix. The prior mean could be:

• Infinitely broad: if the prior is uninformative, the covariance matrix S_0 ends to infinity, leading to:

$$\lim_{\mathbf{S}_0 o \infty} \mathbf{w}_N = \left(\mathbf{\Phi}^T \mathbf{\Phi}\right)^{-1} \mathbf{\Phi}^T \mathbf{t} \qquad \lim_{\mathbf{S}_0 o \infty} \mathbf{S}_N^{-1} = \frac{\mathbf{\Phi}^T \mathbf{\Phi}}{\sigma^2}$$

This reduces the Bayesian solution to the ordinary least squares (OLS) solution, and the MAP estimate becomes equivalent to the Maximum Likelihood estimate. The variance σ^2 can be estimated as:

$$\sigma^2 = \frac{1}{N-M} \sum_{n=1}^{N} (t_n - \hat{\mathbf{w}}^T(\phi)(\mathbf{x}_n))^2$$

• Not infinitely broad: in cases where the prior is informative (e,g, $\mathbf{w}_0 = 0$, $\mathbf{S}_0 = \tau^2 \mathbf{I}$), the posterior can be expressed as:

$$\ln \Pr(\mathbf{w}|\mathbf{t}) = -\frac{1}{2} \sum_{i=1}^{N} (t_i - \mathbf{w}^T \boldsymbol{\phi}(\mathbf{x}_i))^2 - \frac{\sigma^2}{2\tau^2} \|\mathbf{w}\|_2^2$$

The MAP estimate coincides with the solution to ridge regression, where the regularization parameter λ is related to the prior by $\lambda = \frac{\sigma^2}{\tau^2}$.

In Bayesian linear regression, the predictive distribution for a new data point \mathbf{x}^* is given by:

$$\Pr(t|\mathbf{x}, \mathcal{D}) = \mathbb{E}\left[t^*|\mathbf{x}^*, \mathbf{w}, \mathcal{D}\right] = \int \Pr(t^*|\mathbf{x}^*, \mathbf{w}, \mathcal{D}) \Pr(\mathbf{w}|\mathcal{D}) d\mathbf{w}$$

Under Gaussian assumptions, the predictive distribution remains Gaussian with mean and variance:

$$\mu_N(\mathbf{x}) = \boldsymbol{\phi}(\mathbf{x})^T \mathbf{W}_N \qquad \sigma_N^2(\mathbf{x}) = \sigma^2 + \boldsymbol{\phi}(\mathbf{x})^T \mathbf{S}_N \boldsymbol{\phi}(\mathbf{x})$$

As the number of data points N increases, the uncertainty in the parameters (captured by the second term) diminishes, leaving only the variance of the noise σ^2 .

2.2.7 Challenges and limitations

Modeling presents challenges in ensuring our model effectively represents a wide range of plausible functions while maintaining informative priors without overly spreading out probabilities or assigning negligible values.

On the computational side, limitations arise with analytical integration, particularly in cases involving non-conjugate priors and complex models. Approaches like Gaussian approximation, Monte Carlo integration, and variational approximation become necessary for addressing these complexities and achieving accurate results.

Linear models with fixed basis functions offer several benefits:

- They permit closed-form solutions, facilitating efficient computation.
- They lend themselves to tractable Bayesian treatment, enabling principled uncertainty quantification.

• They can capture non-linear relationships by employing appropriate basis functions.

However, these models also come with several drawbacks:

- Basis functions remain static and non-adaptive to variations in the training data.
- These models are susceptible to the curse of dimensionality, particularly when dealing with high-dimensional feature spaces.

2.3 Classification

Classification involves learning an approximation of a function f(x) that maps inputs x to discrete classes C_k (with k = 1, ..., K) from a dataset \mathcal{D} :

$$\mathcal{D} = \{ \langle x, C_k \rangle \} \implies C_k = f(x)$$

Various approaches to classification include:

- Discriminant function: modeling a parametric function that directly maps inputs to classes and learning the parameters from the data.
- Probabilistic discriminative approach: designing a parametric model of $Pr(C_k|\mathbf{x})$ and learning the model parameters from the data.
- Probabilistic generative approach: modeling $Pr(\mathbf{x}|C_k)$ and class priors $Pr(C_k)$, fitting models to the data, and inferring the posterior using Bayes' rule.

In linear classification, we will use generalized linear models:

$$f(\mathbf{x}, \mathbf{w}) = f(\mathbf{x}^T \mathbf{w} + w_0)$$

Here, the function $f(\cdot)$ is not linear in **w** and partitions the input space into decision regions, with their decision boundaries. Notably, these decision boundaries are linear functions of **x** and **w**, expressed as:

$$\mathbf{x}^T \mathbf{w} + w_0 = \text{constant}$$

The labels in a classification problem can be encoded in different ways, depending on the numbers of labels:

- Two labels: we can choose between $t \in \{0,1\}$ and $t \in \{-1,1\}$ depending on the specific situation. The first encoding is useful when we need to model probabilities, the second one is preferable for certain algorithms.
- Multiple lables: in this scenario we have K labels and the typical encoding is called 1-of-K. Here, t is a vector of length K, with a 1 in the position corresponding to the encoded class.

Two-class problem The most general formulation for a discriminant linear function in a two-class linear problem is:

$$f(\mathbf{x}, \mathbf{w}) = \begin{cases} C_1 & \text{if } \mathbf{x}^T \mathbf{w} + w_0 \ge 0 \\ C_2 & \text{otherwise} \end{cases}$$

From this formulation, we can deduce the following properties:

- The decision boundary is $y(\cdot) = \mathbf{x}^T \mathbf{w} + w_0 = 0$.
- The decision boundary is orthogonal to w.
- The distance of the decision boundary from the origin is $\frac{w_0}{\|\mathbf{w}\|_2}$.
- The distance of the decision boundary from \mathbf{x} is $\frac{y(\mathbf{x})}{\|\mathbf{w}\|_2}$.

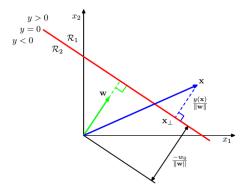


Figure 2.2: Two-class decision problem boundaries

Multiple-class problem In multiple class problems with K classes, various encoding methods can be employed:

- One versus the rest: this approach involves using K-1 binary classifiers, where each classifier distinguishes between one class and the rest of the classes. However, this method introduces ambiguity since there may be regions mapped to multiple classes.
- One versus one: this method utilizes $\frac{K(K-1)}{2}$ binary classifiers, where each classifier discriminates between pairs of classes. Similar to the one versus the rest approach, this method also suffers from ambiguity.
- Linear discriminant functions: one solution to mitigate the ambiguity in multi-class classification is to employ K linear discriminant functions:

$$y_k(\mathbf{x}) = \mathbf{x}^T \mathbf{w}_k + w_{k0} \qquad k = 1, \dots, K$$

In this approach, an input vector \mathbf{x} is assigned to class C_k if $y_k > y_j$ for all $j \neq k$. This method ensures that the decision boundaries are singly connected and convex.

Basis functions Up to this point, we have focused on models operating within the input space. However, we can enhance these models by incorporating a fixed set of basis functions $\phi(\mathbf{x})$. Essentially, this involves applying a non-linear transformation to map the input space into a feature space. Consequently, decision boundaries that are linear within the feature space would correspond to nonlinear boundaries within the input space. This extension enables the application of linear classification models to problems where samples are not linearly separable.

Ordinary Least Squares Let's consider a K-class problem using a 1-of-K encoding for the target. Each class is modeled with a linear function:

$$y_k(\mathbf{x}) = \mathbf{x}^T \mathbf{w}_k + w_{k0} \qquad k = 1, \dots, K$$

In matrix notation, this can be expressed as $\mathbf{y}(\mathbf{x}) = \tilde{\mathbf{w}}^T \tilde{\mathbf{x}}$. Given a dataset $\mathcal{D} = \{\mathbf{x}_i, \mathbf{t}_i\}$ where $i = 1, \dots, N$, we can utilize the Least Squares method to determine the optimal value of $\tilde{\mathbf{w}}$, resulting in:

$$\tilde{\mathbf{w}} = (\tilde{\mathbf{x}}^T \tilde{\mathbf{x}}) \, \tilde{\mathbf{x}}^T \tilde{\mathbf{t}}$$

The primary challenge with employing ordinary Least Squares in classification is that the resulting decision boundaries between regions can vary significantly based on the distribution of the data. This method may yield effective or suboptimal boundaries depending on the characteristics of the dataset.

2.3.1 Discriminant function approach

Percetron To address the issue of poor boundaries, one approach is to utilize a model known as the perceptron. Proposed by Rosenblatt in 1958, the perceptron is a generalized linear model designed specifically for two-class problems. The perceptron model is defined as:

$$f(\mathbf{x}, \mathbf{w}) = \begin{cases} +1 & \text{if } \mathbf{x}^T \mathbf{w} + w_0 \ge 0 \\ -1 & \text{otherwise} \end{cases}$$

The perceptron algorithm aims to determine a decision surface, also known as a separating hyperplane, by minimizing the distance of misclassified samples to the boundary. This minimization of the loss function can be achieved using stochastic gradient descent. Although simpler loss functions could theoretically be used, they are often more complex to minimize in practice.

The perceptron loss function is expressed as:

$$\mathcal{L}_P(\mathbf{w}) = -\sum_{n \in \mathcal{M}} \mathbf{w}^T \mathbf{x}_n t_n$$

Here, correctly classified samples do not contribute to L.

Minimizing \mathcal{L}_P is achieved using stochastic gradient descent:

$$\mathbf{w}^{(k+1)} = \mathbf{w}^{(k)} + \alpha \mathbf{x}_n t_n$$

Since the scale of **w** does not affect the perceptron function, the learning rate α is often set to 1. The perceptron algorithm takes a dataset $\mathcal{D} = \{\mathbf{x}_i, \mathbf{t}_i\}$ where $i = 1, \dots, N$.

Algorithm 1 Perceptron

```
1: Initialize \mathbf{w}_0

2: k = 0

3: repeat

4: k = k + 1

5: n = k \mod N

6: if \hat{t}_n \neq t_n then

7: \mathbf{w}_{k+1} = \mathbf{w}_k + \mathbf{x}_n t_n

8: end if

9: until convergence
```

Theorem 2.3.1 (Perceptron convergence). If the training dataset is linearly separable in the feature space, then the perceptron learning algorithm is guaranteed to find an exact solution in a finite number of steps.

Several steps may be necessary, making it challenging to distinguish between non-separable problems and slowly converging ones. If multiple solutions exist, the one obtained by the algorithm depends on the order of the elements in the dataset.

K-Nearest Neighbors The K-Nearest Neighbors algorithm, specifically the 1-Nearest Neighbors variant, follows a discriminative approach by using the proximity of points in the feature space to make predictions for new data points. The core idea is to find the closest neighbors to predict the target label of an unseen data point.

Given a dataset $\mathcal{D} = \{(x_n, t_n)\}_{i=1}^M$ and a new data point \mathbf{x}_q , 1NN predicts the target by finding the nearest neighbor according to the Euclidean distance:

$$i_1 \in \underset{n \in \{1, \dots, N\}}{\operatorname{argmin}} \|\mathbf{x}_q - \mathbf{x}_n\|_2$$

KNN works effectively for both regression and classification tasks, and it requires no explicit training phase; the model learns by simply querying the dataset at prediction time.

For K > 1, combining the targets from multiple neighbors depends on the type of task:

- Classification: predict the most frequent class among the K-nearest neighbors, with a tie-breaking rule if needed.
- Regression: predict the average target value among the K-nearest neighbors.

This approach can easily handle multiple classes without modification. In multi-class classification, the target is the mode class of the neighbors, while in multi-class regression, it is the average target value.

Note that this algorithm needs to have all the dataset stored in main memory and it is non parametric since we do not have explicit parameters to compute. Higher values for k reduces the variance.

2.3.2 Probabilistic discriminative approach

In a discriminative approach, we model the conditioned class probability directly:

$$\Pr(C_1|\boldsymbol{\phi}) = \frac{1}{1 + e^{-\mathbf{w}^T \boldsymbol{\phi}}} = \sigma(\mathbf{w}^T \boldsymbol{\phi})$$

This model is commonly referred to as logistic regression (generalized linear model).

Maximum Likelihood Given a dataset $\mathcal{D} = \{\mathbf{x}_i, t_i\}$, where i = 1, ..., N and $t_i \in \{0, 1\}$, we aim to maximize the likelihood. We model the likelihood of a single sample using a Bernoulli distribution, employing the logistic regression model for conditioned class probability:

$$\Pr(t_n|\mathbf{x}_n,\mathbf{w}) = y_n^{t_n}(1-y_n)^{1-t_n} \qquad y_n = \sigma(\mathbf{w}^T\boldsymbol{\phi}_n)$$

Assuming independent sampling of data in \mathcal{D} , we have:

$$\Pr(\mathbf{t}|\mathbf{x},\mathbf{w}) = \prod_{n=1}^{N} y_n^{t_n} (1 - y_n)^{(1 - t_n)} \qquad y_n = \sigma(\mathbf{w}^T \boldsymbol{\phi}_n)$$

The negative log-likelihood (also known as cross-entropy error function) serves as a convenient loss function to minimize:

$$\mathcal{L}(\mathbf{w}) = -\ln \Pr(\mathbf{t}|\mathbf{X}, \mathbf{w}) = -\sum_{n=1}^{N} (t_n \ln y_n + (1 - t_n) \ln(1 - y_n))$$

The derivative of the loss function yields the gradient of the loss function:

$$\frac{\partial \mathcal{L}(\mathbf{w})}{\partial \mathbf{w}} = \sum_{n=1}^{N} (y_n - t_n) \, \boldsymbol{\phi}_n$$

Due to the nonlinearity of the logistic regression function, a closed-form solution is not feasible. Nevertheless, the error function is convex, allowing for gradient-based optimization (online gradient descent). The convergence is asymtotically guaranteed, also in case of non linearly separable elements.

Multi class logistic regression In multi class problems, $Pr(C_k|\phi)$ is modeled by applying a softmax transformation to the output of K linear functions (one for each class):

$$\Pr(C_k|\boldsymbol{\phi}) = \frac{e^{\mathbf{w}_k^T \boldsymbol{\phi}}}{\sum_j e^{\mathbf{w}_j^T \boldsymbol{\phi}}}$$

Similar to the two-class logistic regression and assuming 1-of-K encoding for the target, we compute the likelihood as:

$$\Pr(\mathbf{t}|\mathbf{\Phi},\mathbf{w}_1,\ldots,\mathbf{w}_K) = \prod_{n=1}^N \left(\prod_{k=1}^K \Pr\left(C_k|\boldsymbol{\phi}_n\right)^{t_n k}\right) = \prod_{n=1}^N \left(\prod_{k=1}^K y_{nk}^{t_n k}\right)$$

As in the two-class problem, we minimize the cross-entropy error function:

$$L(\mathbf{w}_1, \dots, \mathbf{w}_K) = -\ln \Pr(\mathbf{t}|\mathbf{\Phi}, \mathbf{w}_1, \dots, \mathbf{w}_K) = -\sum_{n=1}^N \left(\sum_{k=1}^K t_{nk} \ln y_{nk}\right)$$

Then, we compute the gradient for each weight vector:

$$\frac{\partial \mathcal{L}_{\mathbf{w}_j}(\mathbf{w}_1, \dots, \mathbf{w}_K)}{\partial \mathbf{w}} = \sum_{n=1}^{N} (y_{nj} - t_{nj}) \, \boldsymbol{\phi}_n$$

Replacing the logistic function with a step function in logistic regression yields the same updating rule as the perceptron algorithm.

2.3.3 Probabilistic generative approach

The primary probabilistic generative model for classification is known as Naive Bayes, which relies on a simplifying assumption known as the naive or conditional independence assumption. This assumption states that, given a class label C_k , the features x_i in the input vector \mathbf{x} are conditionally independent of one another. Under this assumption, the probability of class given an input is expressed as:

$$\Pr(C_k|\mathbf{x}) = \Pr(C_k) \prod_{i=1}^{M} \Pr(x_i|C_k)$$

In Naive Bayes classification, our goal is to predict the most likely class $y(\mathbf{x})$ for a given input \mathbf{x} by maximizing the posterior probability:

$$y(\mathbf{x}) = \underset{k}{\operatorname{argmax}} \Pr(C_k) \prod_{i=1}^{M} \Pr(x_i | C_k)$$

To fit a Naive Bayes model, we typically use a logarithmic transformation of the posterior probability, resulting in a log-likelihood function that we maximize. This optimization is performed through maximum likelihood estimation (MLE), where both the class prior and feature likelihoods are directly estimated from the data.

It is important to note that, despite its name, Naive Bayes is not a Bayesian method. In Bayesian analysis, priors are updated with new evidence. Here, however, we estimate priors directly from the data without subsequent updates.

Naive Bayes, as a generative model, allows us to generate synthetic data resembling the original dataset. The process to generate a new sample is as follows:

- 1. Select a class $C_{\hat{k}}$ according to the estimated multinomial prior distribution with parameters $\hat{\Pr}(C_1), \ldots, \hat{\Pr}(C_K)$.
- 2. For each feature j, draw a sample x_j from the distribution $\mathcal{N}(\hat{\mu}_{j\hat{k}}, \hat{\sigma}_{j\hat{k}}^2)$
- 3. Repeat this process to generate additional samples as needed.

2.3.4 Model evaluation

To assess the performance of a classifier, we can use a confusion matrix. This matrix provides a summary of the number of correctly classified and misclassified samples, offering insights into how well the model distinguishes between classes. Using the values from the confusion matrix,

	Actual Class: 1	Actual Class: 0
Predicted Class: 1	TP	FP
Predicted Class: 0	FN	TN

we can calculate various performance metrics:

• Accuracy: the fraction of total samples that are correctly classified. This is a general measure of the classifier's performance over the entire dataset.

$$Accuracy = \frac{TP + TN}{N}$$

• *Precision*: the fraction of samples correctly classified as positive out of all samples predicted as positive. Precision indicates how many of the predicted positive samples are actually positive.

$$\operatorname{Precision} = \frac{\operatorname{TP}}{\operatorname{TP} + \operatorname{FP}}$$

• *Recall*: the fraction of actual positive samples that are correctly classified. Recall measures the model's ability to identify positive samples from the dataset.

$$Recall = \frac{TP}{TP + FN}$$

• F1 score: the harmonic mean of precision and recall, providing a single metric that balances both. The F1 score is particularly useful when you need to find an equilibrium between precision and recall.

$$F1 = 2 \frac{\text{Precision} \cdot \text{Recall}}{\text{Precision} + \text{Recall}}$$

Higher values for these metrics generally indicate better model performance. These measures are not symmetric; they depend on the choice of the positive class. In some applications, you may switch the positive class to obtain metrics that better reflect the classifier's predictive power in that context.

2.4 Kernel methods

Kernel methods are powerful tools for capturing nonlinear relationships in data. In nonlinear regression, the relationship between input and output variables deviates from a straightforward linear association. Similarly, in classification tasks, the class boundaries may not be linearly separable. When standard linear models fail to capture these complexities, kernel methods transform data into higher-dimensional spaces, allowing linear models to perform effectively even in nonlinear scenarios.

The transformation of data into this higher-dimensional feature space is known as feature mapping, typically represented as:

$$\Phi: x \to \phi(x)$$

However, feature mapping often encounters the curse of dimensionality, where the number of features increases exponentially with the number of input variables, making computation infeasible. Kernel methods address this by sidestepping explicit computation of the feature mapping, thus keeping the computations manageable while preserving the power of the transformation.

2.4.1 Kernel function design

The kernel function measures the similarity between two data samples and is defined as the inner product of their feature vectors:

$$k(\mathbf{x}, \mathbf{x}') = \boldsymbol{\phi}(\mathbf{x})^T \boldsymbol{\phi}(\mathbf{x})$$

This function allows us to assess similarity without explicitly calculating the high-dimensional feature vectors. Large feature vectors can make the kernel function computationally practical, as kernel methods only require pairwise similarities, avoiding direct computation in the higher-dimensional space.

Two main strategies exist for designing kernel functions:

- Direct construction: designing kernel functions from scratch.
- Composition rules: applying a set of rules to existing kernels to construct new ones.

In both approaches, ensuring that a kernel function is valid is crucial, meaning it must correspond to a scalar product in some feature space.

Design rules According to Mercer's theorem, a valid kernel function must be continuous, symmetric, and positive semi-definite.

Theorem 2.4.1 (Mercer). Any continuous, symmetric, positive semi-definite kernel function $k(\mathbf{x}, \mathbf{x}')$ can be expressed as a dot product in a high-dimensional space.

To satisfy Mercer's conditions, the Gram matrix $\mathbf{K} = \mathbf{\Phi}(\mathbf{x})^T \mathbf{\Phi}(\mathbf{x})$, formed by evaluating the kernel on data pairs, must be positive semi-definite:

$$\mathbf{x}^T \mathbf{K} \mathbf{x} > 0 \qquad \forall \mathbf{x} \neq \mathbf{0}$$

Given valid kernels $k_1(\mathbf{x}, \mathbf{x}')$ and $k_2(\mathbf{x}, \mathbf{x}')$ the following rules can be applied to design a new valid kernel:

Transformation	Equivalence	Notes
Constant	$ck_1(\mathbf{x}, \mathbf{x}')$	c > 0 is a constant.
Function	$f(\mathbf{x})k_1(\mathbf{x},\mathbf{x}')f(\mathbf{x}')$	$f(\cdot)$ is any function
Polynomial	$q(k_1(\mathbf{x}, \mathbf{x}'))$	$q(\cdot)$ is a polynomial with non-negative coefficients
Exponential	$e^{k_1(\mathbf{x},\mathbf{x}')}$	Exponential of the kernel function
Sum	$k_1(\mathbf{x}, \mathbf{x}') + k_2(\mathbf{x}, \mathbf{x}')$	Addition of two kernel functions
Product	$k_1(\mathbf{x}, \mathbf{x}')k_2(\mathbf{x}, \mathbf{x}')$	Multiplication of two kernel functions
Composite	$k_3(oldsymbol{\phi}(\mathbf{x}),oldsymbol{\phi}(\mathbf{x}'))$	$\phi(\mathbf{x})$ maps \mathbf{x} to \mathbb{R}^M k_3 is a valid kernel in \mathbb{R}^M
Matrix-based	$\mathbf{x}^T \mathbf{A} \mathbf{x}$	A is a symmetric, positive semi-definite matrix
Subspace kernels sum	$k_a(\mathbf{x}_a, \mathbf{x}_a') + k_b(\mathbf{x}_b, \mathbf{x}_b')$	Kernel sum over partitions of \mathbf{x}
Subspace kernels product	$k_a(\mathbf{x}_a, \mathbf{x}_a') k_b(\mathbf{x}_b, \mathbf{x}_b')$	Kernel product over partitions of \mathbf{x}

Kernel trick A key aspect of kernel methods is the kernel trick, where we replace terms involving $\phi(\mathbf{x})$ with equivalent expressions that rely only on $k(\mathbf{x}, \cdot)$. This allows models to compute outputs based solely on pairwise similarities between data points, making kernel methods versatile across a range of algorithms, including ridge regression, kNN regression, the perceptron, nonlinear PCA, and Support Vector Machines.

Symbolic kernels Kernel methods are not limited to real-valued vectors; they can be applied to more complex data types, such as graphs, sets, strings, and text. In such cases, the kernel function adapts to measure similarity in non-numeric data structures, enabling kernel methods to support a wide array of applications.

Generative kernels Kernel functions can also derive from probability distributions. For example, in a generative modeling context where $Pr(\mathbf{x})$ denotes a probability distribution, a kernel function could be defined as:

$$k(\mathbf{x}, \mathbf{x}') = \Pr(\mathbf{x}) \Pr(\mathbf{x}')$$

This kernel is valid because it represents an inner product in a one-dimensional feature space where each data point \mathbf{x} to $\Pr(\mathbf{x})$.

2.4.2 Kernel ridge regression

In kernel ridge regression, we aim to minimize a loss function that combines the Residual Sum of Squares with a regularization term to control the model complexity. The loss function, $\mathcal{L}(\mathbf{w})$, is defined as:

$$\mathcal{L}(\mathbf{w}) = \frac{1}{2} RSS + \frac{\lambda}{2} ||\mathbf{w}||_2^2$$

Here, λ is the regularization parameter, controlling the trade-off between the fit to the data and the penalty on the magnitude of the weights **w**.

To find the optimal weight vector \mathbf{w} , we set the gradient of $\mathcal{L}(\mathbf{w})$ with respect to \mathbf{w} to zero:

$$\lambda \mathbf{w} - \mathbf{\Phi}^T \mathbf{t} + \mathbf{\Phi}^T \mathbf{\Phi} \mathbf{w} = 0$$

By rearranging and isolating w, we can rewrite it as:

$$\mathbf{w} = \mathbf{\Phi}^T \mathbf{a}$$

Here, $\mathbf{a} = (\mathbf{K} + \lambda \mathbf{I})^{-1}\mathbf{t}$, and $\mathbf{K} = \mathbf{\Phi}\mathbf{\Phi}^T$ is the Gram matrix. The Gram matrix is an $N \times N$ matrix representing the inner products of feature vectors, indicating the pairwise similarity between training samples.

Prediction With this formulation, predictions for a new input \mathbf{x} can be computed as:

$$y(\mathbf{x}) = \mathbf{k}(\mathbf{x})^T (\mathbf{K} + \lambda \mathbf{I})^{-1} \mathbf{t}$$

Here, $\mathbf{k}(\mathbf{x})$ is a vector with elements $k_n(\mathbf{x}) = k(\mathbf{x}_n, \mathbf{x})$ for aeachll \mathbf{x}_n in the training set \mathcal{D} . This formulation expresses predictions as a linear combination of the target values of the training samples, weighted by the kernel-based similarities to the new input.

Comparison In traditional ridge regression, finding **w** requires inverting an $M \times M$ matrix $(\Phi \Phi^T + \lambda \mathbf{I}_M)$ which is efficient when M, the number of features, is relatively small.

In kernel ridge regression, we invert an $N \times N$ matrix $(\mathbf{K} + \lambda \mathbf{I}_N)$, which is preferable when N the number of samples, is relatively large. This approach also avoids explicit computation of the feature map $\mathbf{\Phi}$, as the kernel function directly captures similarities between samples, resulting in a more efficient and flexible implementation.

2.4.3 Kernel k-NN regression

The k-Nearest Neighbors algorithm algorithm can be adapted for regression by predicting the output as the average of the target values of the k nearest samples in the training data. Specifically, the prediction $\hat{f}(\mathbf{x})$ for a new input \mathbf{x} is given by:

$$\hat{f}(\mathbf{x}) = \frac{1}{k} \sum_{\mathbf{x}_i \in N_k(\mathbf{x})} t_i$$

Here, $N_k(\mathbf{x})$ represents the set of the k nearest neighbors of \mathbf{x} , and t_i is the target value associated with \mathbf{x}_i .

While k-NN regression is simple and intuitive, it often produces noisy predictions due to the abrupt changes in neighborhood averages as the test point moves across different sample regions.

Nadaraya-Watson model The Nadaraya-Watson model, also referred to as kernel regression, smooths these abrupt changes by applying a kernel function to weight each sample based on its distance to \mathbf{x} . This results in a continuous, weighted average of the target values:

$$\hat{f}(\mathbf{x}) = \frac{\sum_{i=1}^{N} k(\mathbf{x}, \mathbf{x}_i) t_i}{\sum_{i=1}^{N} k(\mathbf{x}, \mathbf{x}_i)}$$

Here, $k(\mathbf{x}, \mathbf{x}_i)$ is a kernel function that assigns larger weights to samples closer to \mathbf{x} and smaller weights to those further away. This approach reduces the discontinuity of predictions, offering a smoother and often more accurate regression output compared to traditional k-NN regression.

2.4.4 Gaussian processes

A Gaussian Process defines a probability distribution over functions $y(\mathbf{x})$ such that, for any set of input points $\{\mathbf{x}_i\}_{i=1}^N$, the function value $\{y(\mathbf{x}_i)\}_{i=1}^N$ are jointly Gaussian. This property enables GPs to serve as a powerful tool for regression tasks, where the relationship between input data and output predictions is modeled probabilistically.

In particular, a GP with inputs \mathbf{x} has a prior distribution over possible functions $y(\mathbf{x})$ as:

$$y(\mathbf{x}) = \mathcal{N}(\mathbf{0}, \mathbf{K})$$

Here we commonly assume a zero mean (i.e., $\mu = 0$) $y(\mathbf{x})$ unless prior information suggests otherwise. The covariance matrix \mathbf{K} is built from a kernel function $k(\mathbf{x}_i, \mathbf{x}_j)$, which quantifies the similarity between pairs of input points.

Output The target variable **t** is modeled as the true process output **y**, perturbed by Gaussian noise ϵ that is independent of the input:

$$\mathbf{t}_N = \mathbf{y}_N + \epsilon$$

Here, $\epsilon \sim \mathcal{N}(0, \sigma^2)$. Thus, the conditional distribution of the observed targets given the latent function values is:

$$\Pr(\mathbf{t}_n|\mathbf{y}_n) = \mathcal{N}(\mathbf{t}_N|\mathbf{y}_N, \sigma^2\mathbf{I}_N)$$

The prior distribution over the latent function values \mathbf{y}_N is a Gaussian $\mathcal{N}(\mathbf{0}, \mathbf{K}_n)$. Therefore, the marginal distribution of the targets \mathbf{t}_N is also Gaussian:

$$\Pr(\mathbf{t}_N) = \mathcal{N}(\mathbf{t}_N | \mathbf{0}, \mathbf{C}_N)$$

Here, $\mathbf{C}_N = \mathbf{K}_N + \sigma^2 \mathbf{I}_N$ represents the combined effect of the kernel-based covariance and noise.

Gaussian kernel A commonly used kernel in GPs is the Gaussian (or RBF) kernel, defined as:

 $k(\mathbf{x}, \mathbf{x}') = e^{-\frac{\|\mathbf{x} - \mathbf{x}'\|}{2\sigma^2}}$

This kernel measures similarity based on Euclidean distance, where σ^2 controls the smoothness of the function.

Additionally, the Gaussian kernel can be generalized by replacing the Euclidean distance with a more general kernel-based distance measure $k(\mathbf{x}, \mathbf{x}')$, yielding:

$$k(\mathbf{x}, \mathbf{x}') = e^{-\frac{k(\mathbf{x}, \mathbf{x}) + k(\mathbf{x}', \mathbf{x}') - 2k(\mathbf{x}, \mathbf{x}')}{2\sigma^2}}$$

Prediction To predict the target t_{N+1} at a new, unseen input \mathbf{x}_{N+1} , we use the GP framework to define the conditional predictive distribution:

$$\Pr(\mathbf{t}_{N+1}) = \mathcal{N}(\mathbf{t}_{N+1}|\mathbf{0}, \mathbf{C}_{N+1}) = \mathcal{N}(m(\mathbf{x}_{N+1}), \sigma^2(\mathbf{x}_{N+1}))$$

Here, $m(\mathbf{x}_{N+1}) = \mathbf{k}^T \mathbf{C}_N^{-1} \mathbf{t}$ is the predictive mean, and $\sigma^2(\mathbf{x}_{N+1}) = c - \mathbf{k}^T \mathbf{C}_N^{-1} \mathbf{k}$ is the predictive variance, providing a measure of uncertainty. The parameter \mathbf{k} is the covariance vector between \mathbf{x}_{N+1} and the training data points, and $c = k(\mathbf{x}_{N+1}, \mathbf{x}_{N+1})$ is the self-covariance.

Hyperparamtere estimation Although GPs are non-parametric, hyperparameters such as σ^2 (noise level) and those governing the kernel function need to be optimized or chosen. Common approaches include:

- Incorporating prior knowledge of the problem domain.
- Maximizing the log-likelihood on a validation dataset.
- Dynamically adjusting as new data becomes available.

2.4.5 Support Vector Machines

Kernel methods are a powerful tool in Machine Learning, yet they have a significant limitation: calculating the kernel function for every sample in a large training set can be computationally prohibitive. Sparse kernel methods address this by selecting a subset of the training samples, focusing only on those critical for defining decision boundaries. Two prominent examples of sparse kernel methods are Support Vector Machines (SVMs) and Relevance Vector Machines (RVMs).

Linearly separable problems To achieve optimal separation in linearly separable cases, SVMs aim to find a hyperplane that maximizes the margin. This can be formulated as:

$$\underset{\mathbf{w},b}{\operatorname{argmax}} \left\{ \min_{n} \left[\frac{t_n \left(\mathbf{w}^T \boldsymbol{\phi}(\mathbf{x}_n) + b \right)}{\|\mathbf{w}\|} \right] \right\}$$

The optimization problem is often simplified by establishing a canonical hyperplane where only solutions satisfying:

$$t_n\left(\mathbf{w}^T\boldsymbol{\phi}(\mathbf{x}_n) + b\right) = 1 \quad \forall \mathbf{x}_n \in \mathcal{S}$$

are considered. This reformulation allows us to convert the problem into a quadratic programming task:

$$\min \frac{1}{2} \|\mathbf{w}\|_2^2$$

such that $t_n \left(\mathbf{w}^T \boldsymbol{\phi}(\mathbf{x}_n) + b \right) \ge 1 \quad \forall n$

Using Lagrange multipliers, we derive the dual form of this problem:

$$\max \ \tilde{\mathcal{L}}(\boldsymbol{\alpha}) = \sum_{n=1}^{N} \alpha_n - \frac{1}{2} \sum_{n=1}^{N} \sum_{m=1}^{N} \alpha_n \alpha_m t_n t_m k(\mathbf{x}_n, \mathbf{x}_m)$$
 such that $\alpha_n \ge 0$

$$\sum_{n=1}^{N} \alpha_n t_n = 0 \quad \forall n$$

The resulting classifier is represented as:

$$y(\mathbf{x}) = \sum_{n=1}^{N} \alpha_n t_n k(\mathbf{x}, \mathbf{x}_n) + b$$

HOnly the training points with non-zero α_i values (support vectors) influence the decision boundary, and the bias term b, is computed as:

$$b = \frac{1}{|\mathcal{S}|} \sum_{\mathbf{x}_n \in \mathcal{S}} \left(t_n - \sum_{\mathbf{x}_m \in \mathcal{S}} \alpha_m t_m k(\mathbf{x}_n, \mathbf{x}_m) \right)$$

Linearly non-separable problems In real-world applications, data is often not perfectly separable. To accommodate such cases, SVMs introduce slack variables, $\xi_n \geq 0$, which allow some margin violations. This leads to the soft-margin formulation:

$$\min \frac{1}{2} \|\mathbf{w}\|_{2}^{2} + C \sum_{n=1}^{N} \xi_{n}$$

such that $t_{n} \left(\mathbf{w}^{T} \boldsymbol{\phi}(\mathbf{x}_{n}) + b \right) \ge 1 - \xi_{n} \quad \forall n$

The parameter C controls the trade-off between maximizing the margin and minimizing classification errors, with higher C values penalizing errors more heavily.

The dual form of the soft-margin problem is:

$$\max \ \tilde{\mathcal{L}}(\boldsymbol{\alpha}) = \sum_{n=1}^{N} \alpha_n - \frac{1}{2} \sum_{n=1}^{N} \sum_{m=1}^{N} \alpha_n \alpha_m t_n t_m k(\mathbf{x}_n, \mathbf{x}_m)$$
 such that $0 \le \alpha_n \le C$

$$\sum_{n=1}^{N} \alpha_n t_n = 0 \quad \forall n$$

In this setup, the support vectors are the samples for which $\alpha_n > 0$. If $\alpha_n < C$, then $\xi_n = 0$, indicating that the sample lies exactly on the margin boundary. Conversely, if $\alpha_n = C$, the sample may lie within the margin or be misclassified if $\xi_n > 1$.

Alternatively, this optimization problem can be reformulated using a parameter ν to control the fraction of margin violations and the number of support vectors. The equivalent formulation is:

$$\max \ \tilde{\mathcal{L}}(\boldsymbol{\alpha}) = -\frac{1}{2} \sum_{n=1}^{N} \sum_{m=1}^{N} \alpha_n \alpha_m t_n t_m k(\mathbf{x}_n, \mathbf{x}_m)$$
 such that $0 \le \alpha_n \le \frac{1}{N}$
$$\sum_{n=1}^{N} \alpha_n t_n = 0$$
$$\sum_{n=1}^{N} \alpha_n \ge \nu \quad \forall n$$

In this case, $0 \le \nu < 1$ is a user-defined parameter, enabling control over both the fraction of errors and the number of support vectors, ensuring that margin errors and the proportion of support vectors are limited by ν .

Training Solving the optimization problem to determine the values of α_i and b can be computationally demanding, especially for large datasets. Directly solving the SVM optimization problem typically requires $\mathcal{O}(n^3)$ operations, where n is the number of training samples. To address this, various methods have been developed to improve efficiency:

- 1. Chunking: this approach divides the problem into smaller, more manageable chunks. Each chunk, or working set, consists of the current support vectors and a subset of samples with the highest error rates (the worst set). As iterations proceed, the size of the working set may adjust dynamically. While chunking expands the working set as necessary, it converges to the optimal solution.
- 2. Osuna's methods: a variant of chunking designed specifically for SVM optimization, Osuna's method also employs an iterative approach but maintains a fixed working set size. Misclassified samples from the dataset replace samples in the working set during iterations, ensuring convergence while keeping the working set stable.
- 3. Sequential Minimal Optimization: SMO further reduces computational demands by iteratively optimizing only two variables at a time. This minimal working set size allows analytical solutions for each iteration, resulting in a faster convergence to the optimal solution.

For situations requiring online learning, incremental or chunking-based methods can be used to update the model continuously as new data arrives, bypassing the need for full retraining.

Multi-class While SVMs are naturally binary classifiers, they can be adapted to handle multi-class problems through the following techniques:

• One against all: for a k-class problem, one against all creates k binary classifiers, each distinguishing a single class from the others. During testing, the classifier with the highest margin determines the predicted class. One against all is memory-efficient but may require more training time.

• One against one: decomposes a k-class problem into $\frac{k(k-1)}{2}$ binary classifiers, each trained on a pair of classes. In testing, each classifier votes, and the label with the most votes is chosen. One against one offers higher performance and often yields better results.

• Directed Acyclic Graph SVM: like one against one, DAGSVM creates $\frac{k(k-1)}{2}$ classifiers but uses a Directed Acyclic Graph during testing, reducing the number of classifiers needed to k-1. This approach reduces test time while maintaining classification accuracy.

Among these methods, one against one is widely considered to perform best due to its robust pairwise decomposition, while DAGSVM is often preferred for faster testing due to its reduced computational complexity.

Model evaluation

3.1 Bias variance tradeoff

The bias-variance framework provides a structured approach to evaluating model performance. In this framework, we represent the data as a combination of a deterministic component and noise with zero mean and variance σ^2 :

$$t = f(\mathbf{x}) + \varepsilon$$

Known process When the underlying process $f(\mathbf{x})$ generating the data is known, the correct model $y(\mathbf{x})$ for the given process can be determined using Population Risk Minimization:

$$y^*(\mathbf{x}) = \underset{y \in \mathcal{H}}{\operatorname{argmin}} \mathbb{E}_{t,x}[(t - y(\mathbf{x}))^2] = \int \Pr(\mathbf{x})(f(\mathbf{x}) - y(\mathbf{x}))^2 d\mathbf{x}$$

Unknown process When the underlying process is unknown, we use Empirical Risk Minimization:

$$\hat{y}(\mathbf{x}) = \underset{y \in \mathcal{H}}{\operatorname{argmin}} \frac{1}{N} \sum_{n=1}^{n} (t_n - y(x_n))^2$$

Here, the random variable \hat{y} depends on the dataset.

The expected squared error can then be decomposed as follows:

$$\underbrace{\mathbb{E}_{\mathcal{D},t}\left[\left(t-\hat{y}(\mathbf{x})\right)^{2}\right]}_{\text{error}} = \underbrace{\sigma^{2}}_{\text{irreducible error}} + \underbrace{\operatorname{Var}_{\mathcal{D}}\left[\hat{y}(\mathbf{x})\right]}_{\text{variance}} + \underbrace{\mathbb{E}_{\mathcal{D}}[f(\mathbf{x})-\hat{y}(\mathbf{x})]^{2}}_{\text{squared bias}}$$

In this decomposition:

- Expected error: averaged over the training dataset \mathcal{D} and the target t.
- Irreducible error: unaffected by model choice or the number of samples.
- Variance: measures variability between models trained on different datasets, reducing as model complexity decreases or as the sample size grows. High variance leads to overfitting.

• Squared bias: measures the deviation between the true function and the expected learned function, depending on the hypothesis space \mathcal{H} . Bias generally decreases with model complexity. High bias results in underfitting.

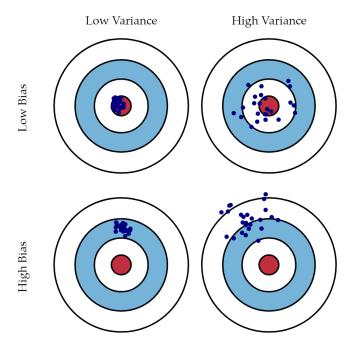


Figure 3.1: Bias-variance framework

The bias-variance decomposition shows why regularization helps reduce error on unseen data. Lasso regression tends to outperform ridge regression when only a few features contribute to the output.

3.1.1 Training error

Given a dataset $\mathcal{D} = \{\mathbf{x}_i, t_i\}$ with i = 1, ..., N, a model is chosen based on the computed loss \mathcal{L} over \mathcal{D} . For regression, the loss function is:

$$L_{train} = \frac{1}{N} \sum_{n=1}^{N} (t_n - y(\mathbf{x}_n))^2$$

The training error decreases as model complexity increases. However, training error does not provide an accurate estimate of the error on new data, known as the prediction error. For regression, the prediction error is represented by:

$$\mathcal{L}_{true} = \iint (t - y(\mathbf{x}))^2 \Pr(\mathbf{x}, t) d\mathbf{x} dt$$

Modeling the joint probability distribution $Pr(\mathbf{x}, t)$ is often infeasible.

In practice, data is typically split into a training set and a test set. Model parameters are optimized using the training set, and prediction error is estimated using the test set. As the sample size grows, training and test errors converge. Examining train and test errors helps identify issues:

• *High bias*: when both training and test errors are high and close to each other.

3.2. Model validation 30

• High variance: when training error is low, but test error gradually increases to match it.

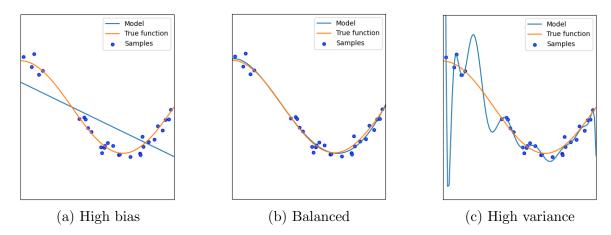


Figure 3.2: Bias-variance balancing

When data is limited, test error may appear minimal, leading to under- or over-estimation of the prediction error. Using test error for model selection can result in overfitting to the test set. An unbiased estimate of prediction error is only achievable if the test set remains separate from training and model selection processes.

3.2 Model validation

To select the optimal model and tune hyperparameters effectively, we first divide the data into three subsets:

- 1. Training set $\mathcal{D}_{\text{train}}$: used to learn the model parameters.
- 2. Validation set $\mathcal{D}_{\text{validation}}$: used to select the best model.
- 3. Test set $\mathcal{D}_{\text{test}}$: used to evaluate the final model's performance.

A typical split is 50%-25%-25% for training, validation, and test sets, respectively. For reliable validation, the validation set must be large enough to prevent overfitting to its specific samples, which could lead to suboptimal model selection.

3.2.1 Leave-One-Out Cross Validation

In leave-one-out cross-validation, the model is trained on all samples except one $\{\mathbf{x}_i, t_i\}$, and the performance is assessed on that omitted sample. The overall prediction error estimate is the average error over all samples:

$$\mathcal{L}_{LOO} = \frac{1}{N} \sum_{i=1}^{N} (t_i - y_{\mathcal{D}_i}(\mathbf{x}_i))^2$$

Here, $y_{\mathcal{D}_i}$ is the model trained on \mathcal{D} excluding $\{\mathbf{x}_i, t_i\}$.

The \mathcal{L}_{LOO} estimate is nearly unbiased, though slightly pessimistic. However, LOO-CV is computationally intensive, as it requires training N models.

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3.2.2 K-Fold Cross Validation

In K-fold cross-validation, the training data \mathcal{D} is split into k equally sized folds: $\mathcal{D}_1, \mathcal{D}_2, \dots, \mathcal{D}_k$. For each fold \mathcal{D}_i , the model is trained on \mathcal{D} excluding \mathcal{D}_i , and the error is calculated on \mathcal{D}_i :

$$\mathcal{L}_{\mathcal{D}_i} = \frac{k}{N} \sum_{(\mathbf{x}_n, t_n) \in \mathcal{D}_i} (t_n - y_{\mathcal{D} \setminus \{\mathcal{D}_i\}}(\mathbf{x}_n))^2$$

The prediction error is then estimated by averaging across all folds:

$$\mathcal{L}_{ ext{k-fold}} = rac{1}{k} \sum_{i=1}^{k} \mathcal{L}_{\mathcal{D}_i}$$

The $\mathcal{L}_{k\text{-fold}}$ estimate of prediction error is slightly biased (pessimistic) but more computationally feasible than LOO-CV. Typically, k is set to 10.

3.2.3 Adjustment tecniques

Several metrics adjust the training error based on model complexity to aid in model evaluation for complex models:

Criteria	Formula
Mallows's C_p	$C_p = \frac{1}{N} \left(\text{RSS} + 2M\sigma^2 \right)$
Akaike Information Criteria (AIC)	$AIC = -2\ln(L) + 2M$
Bayesian Information Criteria (BIC)	$BIC = -2\ln(L) + M\ln(N)$
Adjusted R^2	$A_{R^2} = 1 - \frac{\text{RSS}/(n-m-1)}{\text{TSS}/(N-1)}$

In these formulas:

• M: number of model parameters

• N: number of samples

• L: likelihood function

• σ^2 : estimate of noise variance

• RSS: residual sum of squares

• TSS: total sum of squares

AIC and BIC are often used when maximizing the log-likelihood. Compared to AIC, BIC imposes a stronger penalty for model complexity, favoring simpler models.

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3.3 Model selection

Adding additional features increases the dimensionality of the input space exponentially. This growth not only increases computational cost but also requires more data and may introduce high variance in the model. Our objective is to select a model that minimizes prediction error by reducing variance. Achieving this requires methods that balance complexity with performance, such as:

- Feature selection: selecting a subset of the most relevant features to avoid unnecessary complexity.
- Dimensionality reduction: transforming features into a lower-dimensional space while retaining essential information.
- Regularization: adding penalty terms to the loss function to discourage complex models, helping to prevent overfitting.

These approaches can be combined to improve model performance, as they address complementary aspects of the model selection process.

3.3.1 Feature selection

The simplest approach to feature selection is to evaluate all possible combinations of features. However, given M features, the number of models with exactly k features to evaluate is $\binom{M}{k}$ for each subset. This exhaustive search quickly becomes computationally infeasible as M grows.

In practice, feature selection is often adapted to the type of model being used, and there are three main methods to perform feature selection: filter, embedded, and wrapper methods.

Filter methods Filter methods evaluate each feature independently, using statistical measures to assess its relevance to the target variable. The most relevant k features are then selected based on these metrics. While filter methods are computationally efficient, they may overlook interactions between features because they assess each feature individually, independent of the model.

One example of a filter method is the Pearson correlation coefficient, which measures the linear association between each feature x_j and a target y:

$$\hat{\rho}(x_j, y) = \frac{\sum_{n=1}^{N} (x_{j,n} - \bar{x}_j)(y_n - \bar{y})}{\sqrt{\sum_{n=1}^{N} (x_{j,n} - \bar{x}_j)^2} \sqrt{\sum_{n=1}^{N} (y_n - \bar{y})^2}}$$

Here, \bar{x}_j and \bar{y} are the mean of all the x_j and y respectively. Features with higher correlation coefficients are prioritized. Filter methods typically capture only linear relationships, though there are also extensions to detect nonlinear associations.

Embedded methods Embedded methods incorporate feature selection within the model training process itself. This approach is often used with regularized models, such as Lasso regression, which automatically drives irrelevant feature weights toward zero, effectively eliminating them. Although embedded methods are computationally efficient, they are specific to the chosen model and may not generalize to other algorithms.

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Wrapper methods Wrapper methods utilize a search algorithm to find optimal feature subsets by iteratively training models on different subsets and evaluating their performance. Unlike filter methods, wrapper methods consider interactions between features. Common strategies for searching subsets include greedy algorithms such as forward selection (starting with no features and adding one at a time) and backward elimination (starting with all features and removing one at a time). Though potentially more accurate than filter methods, wrapper methods are often computationally intensive.

3.3.2 Dimensionality reduction

Dimensionality reduction seeks to reduce the number of features in the input space, but it differs from feature selection in two important ways: it utilizes all available features and projects them into a lower-dimensional space, rather than selecting a subset of the original features. Additionally, dimensionality reduction is generally an unsupervised approach, as it does not rely on labeled data.

Several popular methods for dimensionality reduction, each with distinct strengths and use cases, include Principal Component Analysis (PCA), Independent Component Analysis (ICA), self-organizing maps, autoencoders, ISOMAP, and t-SNE.

Principal Component Analysis PCA is an unsupervised dimensionality reduction technique that performs a linear transformation on the original data to extract lower-dimensional features. The core idea of PCA is to find a set of orthogonal directions, or principal components, that capture the maximum variance in the data. The principal components are ranked so that the first component accounts for the highest variance, the second component for the next highest, and so on.

The steps for performing PCA are as follows:

- 1. Translate the original data \mathbf{x} to $\tilde{\mathbf{x}}$ to ensure it has zero mean.
- 2. Compute the covariance matrix of $\tilde{\mathbf{x}}$:

$$\mathbf{C} = \tilde{\mathbf{x}}^T \tilde{\mathbf{x}}$$

- 3. The eigenvectors of C, which correspond to the principal components of the data.
- 4. The eigenvectors can be computed using Singular Value Decomposition.

There are various methods for determining the number of principal components to retain:

- Retain components until the cumulative variance reaches 90%-95%. Cumulative variance is calculated as the fraction of each eigenvalue λ_i relative to the sum of all eigenvalues.
- Retain components that individually explain more than a specified percentage (e.g., 5%) of the variance, discarding only those with very low variance.
- Identify the elbow in the cumulative variance plot, where the marginal gain in explained variance begins to diminish significantly.

PCA is commonly used for feature extraction, data compression, and visualization in lower dimensions.

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3.3.3 Regularization

Regularization is another key method for model selection, primarily aimed at reducing model complexity and preventing overfitting by penalizing larger model coefficients. Regularization techniques such as Lasso, Ridge, and Elastic Net are typically applied to linear regression models, but they can also be adapted to other types of models:

- Lasso (L1 regularization): Adds a penalty equal to the absolute value of the coefficients. Lasso performs feature selection by forcing some coefficients to zero, effectively removing certain features from the model.
- Ridge (L2 regularization): adds a penalty equal to the square of the coefficients, which helps to shrink all coefficients but does not eliminate any entirely. This method is particularly useful for multicollinearity.
- Elastic Net: combines L1 and L2 regularization, balancing feature selection and coefficient shrinkage. Elastic Net is effective in scenarios where there are many correlated features.

Regularization not only aids in model selection by balancing model complexity and predictive accuracy but also improves generalization, particularly in high-dimensional datasets where overfitting is a concern.

3.4 Ensemble

Ensemble methods aim to reduce variance or bias (or both) by combining multiple models to improve overall predictive performance. These objectives are typically achieved through two main techniques:

- Bagging: reduces variance without increasing bias by training multiple models on different subsets of the data.
- Boosting: reduces bias by sequentially combining weak learners to create a strong model.

3.4.1 Bagging

Bagging is an ensemble technique designed to reduce model variance, making it particularly useful for high-variance, low-bias models. The steps for bagging are as follows:

- 1. Generate N different datasets by applying random sampling with replacement (bootstrapping) from the original training data.
- 2. Train a separate model (learner) on each of these datasets in parallel.

For prediction, each model is applied to a new sample, and the outputs are combined:

- In classification, a majority vote across the models is used.
- In regression, predictions are averaged.

Bagging reduces variance by averaging out the noise associated with individual models. It is particularly effective for unstable learners (models highly sensitive to small changes in the data). For such learners, bagging can improve stability and performance. However, bagging is less effective for robust learners with inherently low variance, like linear models.

	Bagging
Primary Goal	Reduces variance
Works Best With	Unstable learners (sensitive to data changes)
Noise Handling	Can be applied with noisy data
Effectiveness	Usually helps, but the difference might be small
Execution	Parallel

3.4.2 Boosting

Boosting is an iterative ensemble technique that combines weak learners sequentially to build a strong model, focusing on reducing bias. The steps in boosting are as follows:

- 1. Assign equal weights to all samples in the training set initially.
- 2. Train a weak learner (often a simple model with high bias) on the weighted dataset.
- 3. Calculate the error of this model on the training set.
- 4. Increase the weights of the misclassified samples so that the next model focuses more on these difficult cases.
- 5. Repeat from step 2 until a stopping criterion is met (e.g., a maximum number of learners or a desired level of accuracy).

Once all learners are trained, they are combined by weighting each model's predictions according to its accuracy. For new samples, the ensemble prediction is a weighted sum (or weighted vote) of the predictions from all the weak learners, with more accurate learners contributing more heavily.

	Boosting
Primary Goal	Reduces bias (generally without overfitting)
Works Best With	Stable learners (less sensitive to data changes)
Noise Handling	Might have problems with noisy data
Effectiveness	Not always helpful, but can make a significant difference
Execution	Sequential

3.5 Computational learning theory

We start by considering an input space \mathcal{X} with M-dimensional features, with an output space denoted as \mathcal{Y} . We also have a joint probability $\Pr(\mathbf{x},t)$, a loss function \mathcal{L} , and the hypothesis space $\mathcal{H} \subset \{h : \mathcal{X} \times \mathcal{Y}\}$

Let's suppose L has identified a hypothesis h^* that makes no error on the training data. We need to find how many training samples from \mathcal{X} are required to ensure that L has learned a true concept.

Theorem 3.5.1 (No free lunch). Let $acc_G(L)$ represent the accuracy of learner L on samples not included in the training set. Let \mathcal{F} be the collection of all potential concepts where $y = f(\S)$. For any binary classifier L and any possible training set:

$$\frac{1}{|\mathcal{F}|} \sum_{\mathcal{F}} acc_G(L) = \frac{1}{2}$$

This means that on avereage every binary classification will behave like a random guess.

This means that there is no model in Machine Learning that is superior with respect to all other models. This means also that in Machine Learning we always operate under some assumptions (such as the assumptions that at least a good approximation of the searched function is in the hypothesis space).

Corollary 3.5.1.1. For any two learners, L_1 and L_2 , if exists $f(\cdot)$ where $acc_G(L_1) > acc_G(L_2)$ then exists $f'(\cdot)$ where $acc_G(L_2) > acc_G(L_1)$.

3.5.1 Ideal learning

Let's assume that the hypothesis space \mathcal{H} contains the real function, so the learner L can obtain a null training error.

Let be \mathcal{D} be the training data drawn from a stationary distribution and labeled (without noise) according to a concept we intend to learn. A binary classifier L produces a hypothesis $h \in H$ such that:

$$h^* = \underset{h \in H}{\operatorname{argmin}} \operatorname{error}_{\operatorname{train}}(h)$$

We determine the training error of a hypothesis as the probability of misclassifying a sample:

$$\operatorname{error}_{\mathcal{D}}(h) = \Pr_{x \in \mathcal{D}} [h(x) \neq c(x)]$$

However, our interest lies in the true error (probability of making a mistake on a sample) of:

$$\operatorname{error}_{\operatorname{true}}(h) = \Pr_{x \sim P(X)} [h(x) \neq c(x)]$$

We say that \mathcal{H} overfits the training data if $\operatorname{error}_{\operatorname{true}} > \operatorname{error}_{\mathcal{D}}$, but we cannot accurately bound $\operatorname{error}_{\operatorname{true}}$ given $\operatorname{error}_{\mathcal{D}}$ because the training data are not independent of \mathcal{H} Therefore, we require a stricter bounding of the error under additional assumptions.

Definition (Consistent hypothesis). A hypothesis \mathcal{H} is deemed consistent with a training dataset \mathcal{D} of the concept c if and only if h(x) = c(x) for each training sample in \mathcal{D} :

consistent
$$(h, \mathcal{D}) \stackrel{\text{def}}{=} h(x) = c(x) \quad \forall \langle x, c(x) \rangle \in \mathcal{D}$$

In other words, is an hypothesis with null training error.

Definition (Version space). The version space, $VS_{\mathcal{H},\mathcal{D}}$, with respect to the hypothesis space \mathcal{H} and the labeled dataset \mathcal{D} , is the subset of hypothesis in \mathcal{H} consistent with \mathcal{D} :

$$VS_{\mathcal{H},\mathcal{D}} \stackrel{\text{def}}{=} \{ h \in \mathcal{H} \mid \text{consistent}(h,\mathcal{D}) \}$$

Thus, the version space consists of all the hypothesis spaces in which we have a null training error.

From now on, we consider only consistent learners, which always output a consistent hypothesis. This means that each time we train a Machine Learning algorithm, it will always find the function with zero training error when we have a consistent hypothesis.

If we aim to bound the error_{true} of a consistent learner, we need to find a bound for all the hypothesis in $VS_{\mathcal{H},\mathcal{D}}$.

Theorem 3.5.2. If the hypothesis space \mathcal{H} is finite and \mathcal{D} is a sequence of $N \geq 1$ independent random examples of some target concept c, then for any $0 \leq \varepsilon \leq 1$, the probability that $VS_{\mathcal{H},\mathcal{D}}$ contains a hypothesis error greater than ε is less than $|\mathcal{H}| e^{\varepsilon N}$:

$$\Pr(\exists h \in \mathcal{H} \mid error_{\mathcal{D}}(h) = 0 \land error_{true}(h) \ge \varepsilon) \le |\mathcal{H}| e^{\varepsilon N}$$

Practical application Let's denote δ as the probability of having $\operatorname{error}_{\operatorname{true}} > \varepsilon$ for a consistent hypothesis:

$$|H| e^{-\varepsilon N} \le \delta$$

By using the logaritms, we can bound both N and ε .

Definition (Probably Approximately Correct learnable concept). A concept C is PAC-learnable by L using \mathcal{H} if:

$$\forall c \in C$$
, distributions $\Pr(\mathcal{X}), 0 < \varepsilon < \frac{1}{2}, 0 < \delta < \frac{1}{2}$

The learner L will with a probability at least $1 - \delta$ output a hypothesis $h \in \mathcal{H}$ such that:

$$\operatorname{error}_{\operatorname{true}}(h) \leq \varepsilon$$

In time that is polynomial in $\frac{1}{\varepsilon}$, $\frac{1}{\delta}$, M, and size(c).

A sufficient condition to prove PAC learnability is proving that a learner L requires only a polynomial number of training examples, and processing per example is polynomial.

3.5.2 Agnostic learning

Up to this point, we've operated under the assumption that the version space $VS_{\mathcal{H},\mathcal{D}}$ is not empty, and that the learner L will consistently output a hypothesis \mathcal{H} such that the error on the dataset is zero. However, in a more general scenario, an agnostic learner might output a hypothesis \mathcal{H} with $error_{\mathcal{D}}(h) > 0$. This means, that it will output a function that is similar to the real one, but not exactly the same.

Theorem 3.5.3. If the hypothesis space \mathcal{H} is finite and \mathcal{D} is a sequence of $N \geq 1$ independent and identically distributed random variables examples of some target concept c, then for any $0 \leq \varepsilon \leq 1$, and for any learned hypothesis \mathcal{H} , the probability that $error_{true}(h) - error_{\mathcal{D}}(h) > \varepsilon$ is less than $|\mathcal{H}| e^{-2N\varepsilon^2}$:

$$\Pr(\exists h \in \mathcal{H} \mid error_{true}(k) > error_{\mathcal{D}}(h) + \varepsilon) \leq |\mathcal{H}| e^{-2N\varepsilon^2}$$

Similar to previous derivations, we can establish a bound on the sample complexity:

$$N \ge \frac{1}{2\varepsilon^2} \left(\ln |\mathcal{H}| + \ln \left(\frac{1}{\delta} \right) \right)$$

Furthermore, we can also constrain the true error of the hypothesis as follows:

$$\operatorname{error}_{\operatorname{true}}(h) \leq \operatorname{error}_{\mathcal{D}}(h) + \sqrt{\frac{\ln |\mathcal{H}| + \ln \frac{1}{\delta}}{2N}}$$

Here, $\operatorname{error}_{\mathcal{D}}(h)$ describe the bias, while the other term describes the variance.

3.5.3 Vapnik-Chervonenkis dimension

The VC dimension represents the size of the subset of X for which $|\mathcal{H}|$ can ensure a zero training error, regardless of the target function.

Definition (Dichotomy). A dichotomy of a set S of instances is defined as a partition of S into two disjoint subsets.

Definition (Shattered). A set of instances S is said to be shattered by hypothesis space \mathcal{H} if and only if for every dichotomy of S, there exists some hypothesis in \mathcal{H} consistent with this dichotomy.

Definition ($VC\ dimension$). The Vapnik-Chervonenkis dimension, $VC(\mathcal{H})$, of hypothesis space \mathcal{H} over instance space X, is the largest finite subset of X shattered by \mathcal{H} .

If an arbitrarily large set of \mathcal{X} can be shattered by \mathcal{H} , then $VC(\mathcal{H}) = \infty$.

If $|\mathcal{H}| < \infty$, then $VC(\mathcal{H}) \le \log_2(|\mathcal{H}|)$. When $VC(\mathcal{H}) = d$, it implies that there are at least 2^d hypothesis in \mathcal{H} to label d instances. Consequently, $|\mathcal{H}| \ge 2^d$. With a probability of at least $(1 - \delta)$, every $h \in \mathcal{H}$ satisfies the following inequality:

$$\operatorname{error}_{\operatorname{true}}(h) \leq \operatorname{error}_{\mathcal{D}}(h) + \sqrt{\frac{\operatorname{VC}(\mathcal{H})\left(\ln\frac{2N}{\operatorname{VC}(\mathcal{H})} + 1\right) + \ln\frac{4}{\delta}}{N}}$$

Reinforcement learning

4.1 Introduction

In reinforcement learning we train a model by providing it with an evaluation of its output

Sequential decision making In the realm of sequential decision making, we navigate through a series of choices or actions aimed at achieving a specific objective. The optimal actions are contingent upon the context in which they occur, often lacking clear-cut examples of correctness. Moreover, these actions can yield long-term ramifications, and while the short-term outcomes of optimal decisions may appear unfavorable, they serve a greater purpose in the pursuit of our goals.

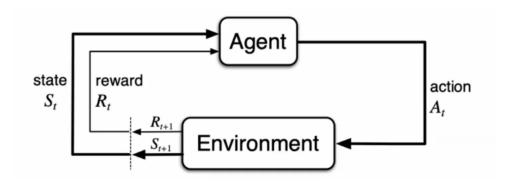


Figure 4.1: Agent-environment interface

At discrete time steps t = 0, 1, 2, K, the agent and environment interact as follows: the agent observes the state at step t, denoted as $S_t \in \mathcal{S}$, produces an action at step t, represented by $A_t \in \mathcal{A}(S_t)$, gets the resulting reward $R_{t+1} \in \mathcal{R}$, and transitions the environment to the next state $S_{t+1} \in \mathcal{S}$.

$$\cdots \underbrace{S_{t}}_{A_{t}} \underbrace{A_{t+1}}_{A_{t+1}} \underbrace{S_{t+1}}_{A_{t+1}} \underbrace{A_{t+2}}_{A_{t+2}} \underbrace{S_{t+3}}_{A_{t+2}} \underbrace{S_{t+3}}_{A_{t+3}} \cdots$$

Figure 4.2: Agent-environment interface

4.2 Markov decision process

Markov decision processes adhere to Markov property:

Property 4.2.1. The future state (s') and reward (r) only depend on current state (s) and action (a)

It is not a limiting assumption, it can be seen as a property of state

One-step dynamic In a Markov Decision Process (MDP), the one-step dynamic can be described as:

$$p(s', r \mid s, a)$$

That is defined on: $p: \mathcal{S} \times \mathcal{R} \times \mathcal{S} \times \mathcal{A} \rightarrow [0,1]$. The main property is that:

$$\sum_{s' \in \mathcal{S}} \sum_{r \in \mathcal{R}} p(s', r \mid s, a) = 1 \qquad \forall s \in \mathcal{S}, \forall a \in \mathcal{A}(s)$$

4.2.1 Finite Markov decision processes

When the Markov Property holds and both the state and action sets are finite, the problem is termed as a finite Markov Decision Process. To formally define a finite MDP, it's necessary to specify the following: sets for states and actions, and one-step dynamics:

$$p(s', r \mid s, a) = \Pr\{S_{t+1} = s', R_{t+1} = r \mid S_t = s, A_t = a\}$$

We can further deduce the distribution of the next state and the expected reward: The overarching formulation is as follows:

$$p(s' \mid s, a) \doteq \Pr\{S_{t+1} = s' \mid S_t = s, A_t = a\} = \sum_{r \in \mathcal{R}} p(s', r \mid s, a)$$
$$r(s, a) \doteq \mathbb{E}[R_{t+1} \mid S_t = s, A_t = a] = \sum_{r \in \mathcal{R}} r \sum_{s' \in \mathcal{S}} p(s', r \mid s, a)$$

Return The agent should refrain from selecting actions solely based on immediate rewards. Instead, prioritizing long-term consequences over short-term gains is crucial. Hence, it's imperative to consider the sequence of future rewards. To this end, we define the return, G_t , as a function of the sequence of future rewards.

$$G_t \doteq f(R_{t+1} + R_{t+2} + R_{t+3} + \cdots)$$

To achieve success, the agent must aim to maximize the expected return, denoted as $\mathbb{E}[G_t]$. Various definitions of return are conceivable, including: total reward, discounted reward, or average reward.

Episodic task In episodic task the agent-environment interaction naturally breaks into chunks called episodes.

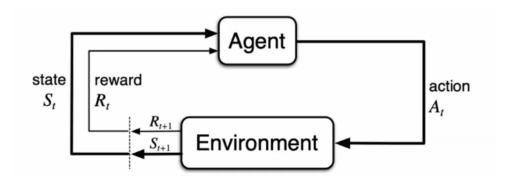


Figure 4.3: Markov decision processes episodic tasks

It is possible to maximize the expected total reward:

$$\mathbb{E}[G_t] = \mathbb{E}[R_{t+1} + R_{t+2} + R_{t+3} + \dots + R_T]$$

Continuing tasks In continuing task the agent-environment interaction goes on continually and there are no terminal state The total reward is a sum over an infinite sequence and might not be finite:

$$G_t \doteq R_{t+1} + R_{t+2} + R_{t+3} + \dots + R_{t+k} + \dots \stackrel{?}{=} \infty$$

To solve this issue we can discount the future rewards by a factor γ (0 < γ < 1):

$$G_t \doteq R_{t+1} + \gamma R_{t+2} + \gamma^2 R_{t+3} + \dots + \gamma^{k-1} R_{t+k} + \dots \stackrel{\leqslant}{\infty}$$

Thus, the expected reward to maximize will be defined as:

$$\mathbb{E}\left[G_{t}\right] = \mathbb{E}\left[\sum_{k=0}^{\infty} \gamma^{k} R_{t+k+1}\right]$$

Return general notation In episodic tasks, we number from zero the time steps for each episode We can design terminal state as absorbing states that always produce zer. reward: We can use the same definition of expected reward for episodic and continuing tasks:

$$\mathbb{E}\left[G_{t}\right] = \mathbb{E}\left[\sum_{k=0}^{\infty} \gamma^{k} R_{t+k+1}\right]$$

Where:

- $\gamma = 1$ can be used if an absorbing state is always reached.
- $\gamma = 0$, agent would only care about immediate reward.
- As $\gamma \to 1$, agent would take future rewards into account more strongly.

Goal and reward A goal should state what we want to achieve, not how we want to achieve it. This idea aligns with the Reward Hypothesis, suggesting that our goals can be seen as maximizing the expected cumulative sum of received rewards.

4.2.2 Policy

A policy, at any particular moment, determines the action that the agent selects. It entirely characterizes the behavior of an agent. Policies can vary in several dimensions:

- Markovian or non-Markovian.
- Deterministic or stochastic.
- Stationary or non-Stationary.

Deterministic policy In its simplest form, the policy can be represented as a function $(\pi: \mathcal{S} \to \mathcal{A})$:

$$\pi(s) = a$$

In this setup, the policy directly maps each state to a specific action. Such policies can be effectively depicted using a table.

Stochastic policy A more versatile approach involves modeling the policy as a function that assigns each state a probability distribution over the available actions:

$$\pi(a|s)$$

Where:

- $\sum_{a \in \mathcal{A}(s)} \pi(a|s) = 1.$
- $\pi(a|s) \ge 0$

A stochastic policy can accommodate deterministic policies as well.

4.2.3 Value functions

For a given policy π , we can compute the state-value function as:

$$V_{\pi}(s) \doteq \mathbb{E}\left[G_t|S_t = s\right] = \mathbb{E}_{\pi}\left[\sum_{k=0}^{\infty} \gamma^k R_{t+k+1}|S_t = s\right]$$

This function signifies the expected return from a specific state s, following policy π , Similarly, we can calculate the action-value function as:

$$Q_{\pi}(s, a) \doteq \mathbb{E}_{\pi} \left[G_t | S_t = s, A_t = a \right] \mathbb{E}_{\pi} \left[\sum_{k=0}^{\infty} \gamma^k R_{t+k+1} | S_t = s, A_t = a \right]$$

This function represents the expected return from a given state s when a particular action a is chosen, followed by policy π .

Bellman expectation equation The state-value function can again be decomposed into immediate reward plus discounted value of successor state:

$$V_{\pi}(s) \doteq \mathbb{E}\left[R_{t+1} + \gamma V_{\pi}(S_{t+1}) | S_t = s\right] = \sum_{a \in \mathcal{A}} \pi(a|s) \left[r(s, a) + \gamma \sum_{s' \in \mathcal{S}} p(s'|s, a) V_{\pi}(s') \right]$$

The action-value function can be similarly decomposed:

$$Q_{\pi}(s, a) = \mathbb{E}_{\pi} \left[R_{t+1} + \gamma V_{\pi} \left(S_{t+1} \right) | S_t = s, A_t = a \right]$$

$$= r(s, a) + \gamma \sum_{s' \in \mathcal{S}} p(s'|s, a) V_{\pi}(s')$$

$$= r(s, a) + \gamma \sum_{s' \in \mathcal{S}} p(s'|s, a) \sum_{a' \in \mathcal{A}} \pi(a'|s') Q_{\pi}(s', a')$$

4.2.4 Optimality

We denote that $\pi \geq \pi'$ if and only if $V_{\pi}(s) \geq V'_{\pi}(s)$ for all $s \in \mathcal{S}$. For any Markov Decision Process, there always exists at least one optimal deterministic policy π^* that is superior or equal to all others:

$$\pi^* > \pi \qquad \forall \pi$$

This occurs because we can select different policies for each interval, consistently choosing the optimal one. In Markov decision processes, we have $|\mathcal{A}|^{|\mathcal{S}|}$ deterministic policies, making brute force search computationally infeasible.

Optimal Value Function To solve this computational problem we can use the optimal value function. Given the optimality definition for the policy, we can compute optimal state-value function and optimal action-value function as:

$$V^*(s) \doteq \max_{\pi} V_{\pi}(s) \qquad \forall s \in \mathcal{S}$$
$$Q^*(s, a) \doteq \max_{\pi} Q_{\pi}(s, a) \qquad \forall s \in \mathcal{S}, \forall a \in \mathcal{A}$$

The corresponding Bellman Optimality Equation for $V^*(s)$ is:

$$V^*(s) = \sum_{a \in \mathcal{A}} \pi^*(a|s) \left(r(s,a) + \sum_{s' \in \mathcal{S}} p(s'|s,a) V^*(s') \right)$$
$$= \max_a \left\{ r(s,a) + \gamma \sum_{s' \in \mathcal{S}} p(s'|s,a) V^*(s') \right\}$$

The corresponding Bellman Optimality Equation for $Q^*(s)$ is:

$$Q^*(s, a) = r(s, a) + \gamma \sum_{s' \in S} p(s'|s, a) \sum_{a' \in A} \pi^*(a'|s') Q^*(s', a')$$
$$= r(s, a) + \gamma \sum_{s' \in S} p(s'|s, a) \max_{a'} Q^*(s', a')$$

We can make this change since the considered policy is optimal. From $V^*(s)$ and $Q^*(s,a)$ we can easily compute the optimal policy π^* as:

$$V^*(s) = \underset{a}{\operatorname{argmax}} \left\{ r(s, a) + \gamma \sum_{s' \in \mathcal{S}} p(s'|s, a) V^*(s') \right\}$$

The problem with this function is that it is not linear. To make it linear we again need the value of π^* , and since it is computationally infeasible to compute this function for every policy we need a different approach.

4.3 Dynamic programming

To resolve an MDP, locating the optimal policy is essential. However, employing a brute force method is impractical due to the necessity to solve $|\mathcal{S}|$ linear equations for each policy. Dynamic Programming (DP) offers a solution by dissecting the intricate problem into more manageable sub-problems recursively. Through the utilization of DP, we'll explore how to effectively tackle an MDP using the Bellman Equations.

By utilizing Dynamic programming we can evaluate multiple policies and compute the corresponding state-value function. Then, by using the Bellman equation we can find the optimal one again using dynamic programming.

4.3.1 Policy evaluation

We search the solution of the Bellman expectation equation:

$$V_{\pi}(s) = \sum_{a \in \mathcal{A}} \pi(a|s) \left[r(s,a) + \gamma \sum_{s' \in \mathcal{S}} p(s'|s,a) V_{\pi}(s') \right]$$

DP solves this problem through iterative application of Bellman equation:

$$V_{k+1}(s) = \sum_{a \in \mathcal{A}} \pi(a|s) \left[r(s,a) + \gamma \sum_{s' \in \mathcal{S}} p(s'|s,a) V_k(s') \right]$$

At each iteration k, the value-function V_k is updated for all state $s \in \mathcal{S}$. It can be proved that V_k converge to V_{π} as k tends to infinity for any initial value V_0 .

Algorithm 2 Iterative policy evaluation algorithm

```
1: Initialize V(s) for all s \in \mathcal{S}^+ arbitrarily

2: V(\text{terminal}) = 0

3: \mathbf{repeat}

4: \Delta = 0

5: \mathbf{for} \ \text{each} \ s \in \mathcal{S} \ \mathbf{do}

6: v = V(s)

7: V(s) = \sum_{a} \pi(a|s) \sum_{s',r} p(s',r|s,a) \left[r + \gamma V(s')\right]

8: \Delta = \max\left(\Delta, |v - V(s)|\right)

9: \mathbf{end} \ \mathbf{for}

10: \mathbf{until} \ \Delta < \theta
```

The input of this algorithm is the policy to be evaluated π . It also has a small threshold $\theta > 0$, that is a parameter used to determine the accuracy of the estimation.

4.3.2 Policy improvement

Normally, the optimal policy from optimal value functions is derived as:

$$\pi^*(s) = \operatorname*{argmax}_{a} \left\{ r(s, a) + \gamma \sum_{s' \in \mathcal{S}} p(s'|s, a) V^*(s') \right\} = \operatorname*{argmax}_{a} Q^*(s, a)$$

If we act greedy with respect to non optimal value function we have:

$$\pi'(s) = \operatorname*{argmax}_{a} \left\{ r(s, a) + \gamma \sum_{s' \in \mathcal{S}} p(s'|s, a) V_{\pi}(s') \right\} = \operatorname*{argmax}_{a} Q_{\pi}(s, a)$$

We may have two outcomes:

- $\pi' = \pi$ it means that π is already the optimal policy π^* .
- $\pi' \neq \pi$ it means that π' is better or as good as π .

The second point is guaranteed by the following theorem.

Theorem 4.3.1. For any pair deterministic policies π' and π such that:

$$Q_{\pi}(s, \pi'(s)) \ge Q_{\pi}(s, \pi(s)) \quad \forall s \in \mathcal{S}$$

Then π' is better or as good as π :

$$\pi' \geq \pi$$

Corollary 4.3.1.1. If exists $s \in \mathcal{S}$ such that $Q_{\pi}(s, \pi'(s)) > Q_{\pi}(s, \pi(s))$, then $\pi' > \pi$.

Proof. We have that:

$$V_{\pi}(s) \leq Q_{\pi}(s, \pi(s')) = \mathbb{E}_{\pi'} \left[R_{t+1} + \gamma V_{\pi}(S_{t+1}) | S_t = s \right]$$

$$\leq \mathbb{E}_{\pi'} \left[R_{t+1} + \gamma Q_{\pi}(S_{t+1}, \pi'(S_{t+1})) | S_t = s \right]$$

$$\leq \mathbb{E}_{\pi'} \left[R_{t+1} + \gamma R_{t+2} + \gamma^2 Q_{\pi}(S_{t+2}, \pi'(S_{t+2})) | S_t = s \right]$$

$$\leq \mathbb{E}_{\pi'} \left[R_{t+1} + \gamma R_{t+2} + \dots | S_t = s \right] = V_{\pi'}(s)$$

4.3.3 Policy iteration

We can exploit the policy improvement theorem to find the optimal policy:

$$\pi_0 \stackrel{E}{\to} V_{\pi_0} \stackrel{I}{\to} \pi_1 \stackrel{E}{\to} V_{\pi_1} \stackrel{I}{\to} \pi_2 \stackrel{E}{\to} \cdots \stackrel{I}{\to} \pi^* \stackrel{E}{\to} V^*$$

Algorithm 3 Policy iteration algorithm

```
1: V(s) \in \mathbb{R} and \pi(s) \in \mathcal{A}(s) arbitrarily for all s \in \mathcal{S}
                                                                                                                           ▶ Initialization
 2: repeat
 3:
          repeat
                                                                                                                    ▶ Policy evaluation
                \Delta = 0
 4:
 5:
                for each s \in \mathcal{S} do
                     v = V(s)
 6:
                     V(s) = \sum_{a} \pi(a|s) \sum_{s',r} p(s',r|s,a) \left[ r + \gamma V(s') \right]
\Delta = \max \left( \Delta, |v - V(s)| \right)
 7:
 8:
                end for
 9:
          until \Delta < \theta
10:
          policy-stable = true
11:
                                                                                                                ▷ Policy improvement
          for each s \in \mathcal{S} do
12:
                old-action = \pi(s)
13:
                \pi(s) = \operatorname{argmax}_a \sum_{s',r} p(s',r|s,a)[r + \gamma V(s')]
if old-action \neq \pi(s) then
14:
15:
                     policy-stable = false
16:
                end if
17:
18:
          end for
     until policy-stable = true
20: return V \approx v^* and \pi \approx \pi^*
```

4.3.4 Value iteration

Policy iteration alternates complete policy evaluation and improvement up to the convergence Policy iteration framework allows also to find the optimal policy interleaving partial evaluation and improvement steps In particular, Value Iteration is one of the most popular GPI method In the policy evaluation step, only a single sweep of updates is performed:

$$\pi'(s) = \operatorname*{argmax}_{a} \left\{ r(s, a) + \gamma \sum_{s' \in \mathcal{S}} p(s'|s, a) V_{\pi}(s') \right\} \qquad \forall s \in \mathcal{S}$$
$$V_{k+1}(s) = \sum_{a} \in \mathcal{A}\pi'(a|s) \left(r(s, a) + \gamma \sum_{s' \in \mathcal{S}} p(s'|s, a) V_{k}(s') \right) \qquad \forall s \in \mathcal{S}$$

Combining them, we simply need to iterate the update of the value function using the Bellman optimality equation:

$$V_{k+1}(s) = \max_{a} \left[r(s, a) + \gamma \sum_{s' \in \mathcal{S}} p(s'|s, a) V_k(s') \right] \qquad \forall s \in \mathcal{S}$$

It can be proved that:

$$\lim_{k \to \infty} V_k = V^*$$

Algorithm 4 Iterative policy evaluation algorithm

```
1: Initialize V(s) for all s \in \mathcal{S}^+ arbitrarily
 2: V(\text{terminal}) = 0
 3: repeat
          \Delta = 0
 4:
 5:
          for each s \in \mathcal{S} do
               v = V(s)
 6:
               V(s) = \max_{a} \sum_{s',r} p(s',r|s,a) \left[r + \gamma V(s')\right]
 7:
               \Delta = \max(\Delta, |v - V(s)|)
 8:
          end for
 9:
10: until \Delta < \theta
```

It also has a small threshold $\theta > 0$, that is a parameter used to determine the accuracy of the estimation. The output is a deterministic policy $\pi \approx \pi^*$, such that:

$$\pi(s) = \operatorname*{argmax}_{a} \sum_{s',r} p(s',r|s,a) [r + \gamma V(s')]$$

4.3.5 Efficiency

All previously described DP methods mandate exhaustive sweeps across the complete state set. However, Asynchronous DP diverges from this approach by eschewing sweeps. Instead, it operates by selecting a state randomly, applying the relevant backup, and iterating until a convergence criterion is satisfied. We can chose states for backup in a more intelligent manner by noticing that an agent's experience can serve as a valuable guide in this regard.

The complexity of finding an optimal policy is polynomial in the number of states and actions:

- For value iteration: $O(|\mathcal{S}|^2 |\mathcal{A}|)$.
- For policy iteration:

- Iterative evaluation:
$$O\left(\frac{|\mathcal{S}|^2 \log\left(\frac{1}{\epsilon}\right)}{\log\left(\frac{1}{\gamma}\right)}\right)$$

- Improvement:
$$O\left(\frac{|\mathcal{A}|}{1-\gamma}\log\left(\frac{|\mathcal{S}|}{1-\gamma}\right)\right)$$

Unfortunately, the number of states can become extremely large, often growing exponentially with the number of state variables (known as the curse of dimensionality). Classical DP works well for problems with a few million states, but Asynchronous DP can handle larger ones and is suitable for parallel computing. However, there are MDPs where DP methods become impractical. Linear programming approaches are an alternative, but they don't scale well for larger problems.

4.4 Monte Carlo methods

Dynamic Programming enables us to determine the optimal value function and corresponding optimal policy. However, its major limitation lies in the assumption that we have full knowledge of the problem dynamics. To overcome this limitation, we seek methods that can learn the optimal policy directly from data.

Monte Carlo methods rely solely on experience (data) to learn value functions and policies. They can be utilized in two ways:

- Model-free: no model is necessary, yet it can still achieve optimality.
- Simulated: requires only a simulation, not a complete model.

Monte Carlo methods learn from complete sample returns and are exclusively defined for episodic tasks.

4.4.1 Policy evaluation

The goal of Monte Carlo policy evaluation is to learn $V_{\pi}(s)$ given some number of episodes under π which contain s. The idea is to average the returns observed after visits to s:

$$V_{\pi}(s) \doteq \mathbb{E}_{\pi}[G_t|S_t = s] \rightarrow V_{\pi}(s) \approx \operatorname{average}[G_t|S_t = s]$$

We can perform Monte Carlo policy evaluation in two ways:

- Every-Visit MC: average returns for every time s is visited in an episode
- First-visit MC: average returns only for first time s is visited in an episode

Note that Both converge asymptotically

Algorithm 5 Monte Carlo policy evaluation algorithm

```
1: Initialize V(s) \in \mathbb{R} arbitrarily, for all s \in \mathcal{S}
                                                                                                    ▶ Initialization
 2: Initialize Returns(s) as an empty list, for all s \in \mathcal{S}
 3: repeat
        for each episode do
 4:
             Generate an episode following \pi: S_0, A_0, R_1S_1, A_1, R_2, \cdots, S_{T-1}, A_{T-1}, R_T
 5:
 6:
             for each step of episode t = T - 1, T - 2, \dots, 0 do
 7:
                 G = \gamma G + R_{t+1}
 8:
                 if S_t \notin \{S_0, S_1, \dots, S_{t-1}\} then:
 9:
                      Append G to Returns(S_t)
10:
                      V(S_t) = \text{average}(\text{Returns}(S_t))
11:
                 end if
12:
             end for
13:
        end for
14:
15: until true
```

The input of this algorithm is a policy π to be evaluated. The incremental updates of lines ten and eleven can be done in the following way:

$$N(S_t) = N(S_t) + 1$$

$$V(S_t) = V(S_t) + \frac{1}{N(S_t)} (G - V(S_t)) \text{ or } V(S_t) = V(S_t) + \alpha (G_t - V(S_t))$$

4.4.2 Policy iteration

To improve the policy we need to find a policy that maximized the q value function:

$$\pi'(s) = \operatorname*{argmax}_{a} Q_{\pi}(s, a)$$

To do so, we average return starting from state s and action a following π :

$$Q_{\pi}(s, a) \doteq \mathbb{E}_{\pi}[G_t | S_t = s, A_t = a] \rightarrow Q_{\pi}(s, a) \approx \operatorname{average}[G_t | S_t = s, A_t = a]$$

This method Converges asymptotically if every state-action pair is visited.

To have this full exploration in a simple way we use exploring starts. We choose randomly the first state and the first action, and we perform the following algorithm.

Algorithm 6 Monte Carlo exploring starts

```
1: \pi(s) \in \mathcal{A}(s) arbitrarily, for all s \in \mathcal{S}
 2: Q(s, a) \in \mathbb{R} arbitrarily, for all s \in \mathcal{S}, a \in \mathcal{A}(s)
 3: Returns(s, a) = empty list, for all s2S, a \in \mathcal{A}(s)
 4: loop
         Choose S_0 \in \mathcal{S}, A_0 \in \mathcal{A}(S_0) randomly such that all pairs have probability greater than
 5:
    zero
         Generate an episode from S_0, A_0, following \pi: S_0, A_0, R_1, \cdots, S_{T-1}, A_{T-1}, R_T
 6:
 7:
         for each step of episode, t = T_1, T_2, \cdots, 0 do
 8:
              G = \gamma G + R_{t+1}
 9:
              if S_t, A_t \notin S_0, A_0, S_1, A_1, \dots, S_{t-1}, A_{t-1} then
10:
                   Append G to Returns(S_t, A_t)
11:
                   Q(S_t, A_t) = \text{average}(\text{Returns}(S_t, A_t))
12:
                   \pi(S_t) = \operatorname{argmax}_a Q(S_t, a)
13:
              end if
14:
         end for
15:
16: end loop
```

4.4.3 Epsilon-soft Monte Carlo policy iteration

Exploring starts is a simple idea but it is not always possible. But, we need to keep exploring during the learning process This leads to a key problem in RL: the Exploration-Exploitation Dilemma

 ε -Greedy Exploration is the simplest solution to the exploration-exploitation dilemma Instead of searching the optimal deterministic policy we search the optimal ε -soft policy, i.e., a policy that selects each action with a probability that is at least $\frac{\varepsilon}{|A|}$.

In particular we use ε -greedy policy:

$$\pi(a|s) = \begin{cases} \frac{\varepsilon}{|\mathcal{A}(s)|} + 1 - \varepsilon & \text{if } a^* = \operatorname{argmax}_{a \in \mathcal{A}} Q(s, a) \\ \frac{\varepsilon}{|\mathcal{A}(s)|} & \text{otherwise} \end{cases}$$

This algorithm takes as input a small $\varepsilon > 0$

Algorithm 7 ε -soft Monte Carlo policy iteration

```
1: \pi=an arbitrary \varepsilon-soft policy
 2: Q(s, a) \in \mathbb{R} arbitrarily, for all s \in \mathcal{S}, a \in \mathcal{A}(s)
 3: Returns(s, a) empty list, for all s \in \mathcal{S}, a \in \mathcal{A}(s)
 4: loop for each episode
           Generate an episode following \pi: S_0, A_0, R_1, \ldots, S_{T-1}, A_{T-1}, R_T
 5:
 6:
          for each step of episode, t = T - 1, T - 2, \dots, 0 do
 7:
                G = \gamma G + R_{t+1}
 8:
                if S_t, A_t \notin S_0, A_0, S_1, A_1, \dots, S_{t-1}, A_{t-1} then
 9:
                      Append G to Returns(S_t, A_t)
10:
                      Q(S_t, A_t) = \operatorname{average}(\operatorname{Returns}(S_t, A_t))
11:
                      A^* = \operatorname{argmax}_a Q(S_t, a)

    ▶ Ties broken arbitrarily

12:
                      for a \in \mathcal{A}(S_t) do
13:
                          \pi(a|S_t) = \begin{cases} 1 - \varepsilon + \frac{\varepsilon}{|A(S_t)|} & \text{if } a = A^* \\ \frac{\varepsilon}{|A(S_t)|} & \text{if } a \neq A^* \end{cases}
14:
                      end for
15:
                end if
16:
           end for
17:
18: end loop
```

Theorem 4.4.1. Any ε -greedy policy π' with respect to Q_{π} is an improvement over any ε -soft policy π .

Proof. We have that:

$$V_{\pi}(s) = Q_{\pi}(s, \pi'(s))$$

$$= \sum_{a \in \mathcal{A}} \pi'(a|s) Q_{\pi}(s, a)$$

$$= \varepsilon \sum_{a \in \mathcal{A}} \frac{1}{|A|} \sum_{a \in \mathcal{A}} Q_{\pi}(s, a) + (1 - \varepsilon) \max_{a \in \mathcal{A}} Q^{\pi}(s, a)$$

$$\geq \varepsilon \sum_{a \in \mathcal{A}} Q_{\pi}(s, a) + (1 - \varepsilon) \sum_{a \in \mathcal{A}} \frac{\pi(a|s) - \frac{\varepsilon}{|A|}}{1 - \varepsilon} \bar{Q}_{\pi}(s, a)$$

$$= \sum_{a \in \mathcal{A}} \pi(a|s) Q_{\pi}(s, a) = V_{\pi}(s)$$

4.4.4 Off-policy learning

On-policy learning On-policy learning involves the agent learning the value functions based on the same policy it uses to select actions. This method faces challenges in balancing exploration and exploitation, making it difficult to converge to an optimal deterministic policy.

Off-policy learning Off-policy learning, on the other hand, allows the agent to select actions using a behavior policy b(a|s), while learning the value functions of a different target policy

 $\pi(a|s)$. This flexibility enables the agent to use an explorative behavior policy, while still learning towards an optimal deterministic policy $\pi^*(a|s)$.

Regardless of our behavior policy, it's impossible to learn any policy $\pi(a|s)$ if there are actions in that state with zero probability according to the behavior policy b(a|s). This situation occurs when the behavior policy never transitions to a particular state from the current one.

Importance sampling Importance sampling enables the estimation of expectations of a distribution that differs from the one used to draw the samples:

$$\mathbb{E}_p[x] = \sum_{x \in X} x p(x) = \sum_{x \in X} x \frac{p(x)}{q(x)} q(x) = \sum_{x \in X} z \rho(x) q(x) = \mathbb{E}_q[x \rho(x)]$$

Consequently, for sample-based estimation:

$$\mathbb{E}_p[x] \approx \frac{1}{N} \sum_{i=1}^N x_i \text{ if } x_i \sim p(x) \to \mathbb{E}_p[x] \approx \frac{1}{N} \sum_{i=1}^N x_i \rho(x_i) \text{ if } x_i \sim q(x)$$

Importance sampling in policy evaluation When adhering to policy π , the computation of the state value function is expressed as:

$$V_{\pi}(s) \approx \text{average}(\text{Returns}[0], \text{Returns}[1], \text{Returns}[2], \cdots)$$

However, under policy b, the value function transforms into:

$$V_{\pi}(s) \approx \text{average}(\rho_0 \text{Returns}[0], \rho_1 \text{Returns}[1], \rho_2 \text{Returns}[2], \cdots)$$

Here, ρ_i denotes the probability of executing the trajectory observed in episode *i* while adhering to policy π , relative to the probability of observing the same trajectory while following policy *b*:

$$\rho = \frac{\Pr(\text{trajectory under } \pi)}{\Pr(\text{trajectory under } b)}$$

In practical terms, $\rho_{t:T-1}$ can be calculated as:

$$\rho_{t:T-1} = \prod_{k=t}^{T-1} \frac{\pi(A_k|S_k)}{b(A_k|S_k)}$$

Sampling methods include:

• Ordinary: unbiased with higher variance:

$$V_{\pi}(s) \approx \frac{\sum_{i} \rho[i] \operatorname{Return}[i]}{N(s)}$$

• Weighted: biased (bias converges to zero) with lower variance:

$$V_{\pi}(s) \approx \frac{\sum_{i} \rho[i] \operatorname{Return}[i]}{\sum_{i} \rho[i]}$$

Algorithm 8 Off-Policy every visit Monte Carlo prediction

```
1: V(s) \in \mathbb{R} arbitrarily, for all s \in S
 2: Returns(s) = an empty list, for all s \in S
 3: for each episode do
        Generate an episode following b: S_0, A_0, R_1, S_1, \ldots, S_{T-1}, A_{T-1}, R_T
 4:
 5:
        G=0
        W = 1
 6:
        for each step of episode, t = T - 1, T - 2, \dots, 0 do
 7:
             G = \gamma WG + R_{t+1}
 8:
             Append G to Returns(S_t)
 9:
             V(S_t) = \text{average}(\text{Returns}(S_t))
10:
             W = W \frac{\pi(A_t | \tilde{S}_t)}{b(A_t | S_t)}
11:
        end for
12:
13: end for
```

The input of this algorithm is a policy π to be evaluated.

4.5 Multi-armed bandits

In the k-armed bandit problem, an agent faces a decision-making scenario where it selects from k actions and receives a reward based on the chosen action. The objective is to identify the optimal action among the available options. Unlike many decision problems, this setting lacks contextual information; decisions are made in isolation, without considering a broader state.

Feedback in this problem manifests as evaluations (rewards) of decisions made under uncertainty, with learning occurring through trial and error and interaction with the environment.

The value associated with each action is represented by its expected reward:

$$q^* \doteq \mathbb{E}[R_t|A_t = a] = \sum p(r|a)r \quad \forall a \in \{1, \dots, k\}$$

Here, the agent's aim is to maximize the expected reward by selecting:

$$\operatorname*{argmax}_{a} q^{*}(a)$$

Since the exact distribution p(r|a) is typically unknown, the agent estimates $q^*(a)$ based on its experiences:

$$Q_t(a) \doteq \frac{\sum_{i=1}^{t-1} R_i \mathbb{1}_{A_t = a}}{\sum_{i=1}^{t-1} \mathbb{1}_{A_t = a}}$$

This expression represents the ratio of the cumulative rewards received when action a was chosen before time step t, divided by the number of times action a was selected up to time step t.

4.5.1 Incremental update of action-values

Let's examine the update for a single action:

$$Q_{n+1} = \frac{1}{n} \sum_{i=1}^{n} R_i$$

$$= \frac{1}{n} \left(R_n + (n-1) \frac{1}{n-1} \sum_{i=1}^{n-1} R_i \right)$$

$$= \frac{1}{n} \left(R_n + (n-1)Q_n \right)$$

$$= Q_n + \frac{1}{n} \left(R_n - Q_n \right)$$

In this equation, Q_{n+1} represents the new estimate, Q_n denotes the old estimate, $\frac{1}{n}$ stands for the step size, and $(R_n - Q_n)$ serves as the target for the old estimate.

Non-stationary bandit problem For non-stationary bandit problems, the update equation takes the form:

$$Q_{n+1} = Q_n + \alpha \left(R_n - Q_n \right)$$

Here, the parameter α varies over time.

4.5.2 Epsilon-greedy action selection

Selecting the action with the highest value isn't always optimal, as it may not lead to the best outcome. Thus, striking a balance between exploration and exploitation becomes crucial:

- Exploitation: The agent leverages its current knowledge to gain immediate rewards.
- Exploration: The agent seeks to enhance its knowledge for long-term gains.

To navigate this trade-off, we can employ epsilon-greedy action selection:

$$A_t = \begin{cases} \operatorname{argmax}_a Q_t(a) & \text{with probability } 1 - \varepsilon \\ \operatorname{Uniform}(\{a_1, \cdots, a_k\}) & \text{with probability } \varepsilon \end{cases}$$

4.5.3 Optimistic initial values

Traditionally, we've initialized action-values to 0.0. However, initializing them with values different from zero can yield varied outcomes.

Optimistic initial values encourage early exploration but may not be suitable for non-stationary problems, where the environment's dynamics change over time.

Determining the appropriate optimistic initial value can also pose a challenge, as it's often unclear what value would be most effective in driving exploration.

4.5.4 UCB action selection

In epsilon-greedy action selection, we had:

$$A_t = \begin{cases} \operatorname{argmax}_a Q_t(a) & \text{with probability } 1 - \varepsilon \\ \operatorname{Uniform}(\{a_1, \cdots, a_k\}) & \text{with probability } \varepsilon \end{cases}$$

However, we can improve upon the uniform function with the following approach:

$$A_t = \operatorname*{argmax}_{a} \left[Q_t(a) + c \sqrt{\frac{\ln(t)}{N_t(a)}} \right]$$

Here, $Q_t(a)$ represents exploitation, c is a user-defined coefficient, and $\frac{\ln(t)}{N_t(a)}$ accounts for exploration.

4.6 Temporal difference learning

Dynamic Programming (DP) necessitates knowledge of the Markov Decision Process (MDP) dynamics. On the other hand, Monte Carlo (MC) learning relies on experience but mandates complete episodes for updating. Consequently, it's solely applicable to episodic tasks. However, even within episodic tasks, MC might encounter challenges.

4.6.1 Temporal-Difference Policy Evaluation with TD(0)

Temporal-Difference combines MC (model-free) with DP (bootstrapping):

$$V(S_t) = V(S_t) + \alpha[G_t - V(S_t)] = V(S_t) = V(S_t) + \alpha[R_{t+1} + \gamma V(S_{t+1}) - V(S_t)]$$

Here, $R_{t+1} + \gamma V(S_{t+1}) - V(S_t)$ is the Temporal-Difference Error δt .

Algorithm 9 TD(0) Policy Evaluation

```
1: Initialize V(s) arbitrarily, for all s \in \mathcal{S}^+
 2: V(\text{terminal}) = 0
 3: for each episode do
       Initialize S
 4:
       repeat for each step of episode
 5:
            A = action given by \pi for S
 6:
            Take action A, observe R, S'
 7:
            V(S) = V(S) + \alpha [R + \gamma V(S') - V(S)]
 8:
            S = S'
 9:
        until S is terminal
10:
11: end for
```

The algorithm takes as input the policy π to be evaluated, and a parameter $\alpha \in (0, 1]$ that represent the step size. One main advantage is that the value function is updated during the episode and not after.

4.6.2 Comparison

Temporal Difference (TD) learning has several advantages over Monte Carlo (MC) learning:

- TD can learn before knowing the final outcome and can update its estimates after every step, whereas MC must wait until the end of an episode before the return is known.
- TD can learn from incomplete sequences, making it more flexible than MC, which can only learn from complete sequences.
- TD is suitable for both continuing (non-terminating) and episodic (terminating) environments, while MC is limited to episodic tasks.
- MC returns are unbiased estimates of the value function, whereas TD targets are biased estimates due to the use of bootstrapping.
- TD targets have lower variance compared to MC returns because they depend on fewer random actions, transitions, and rewards.
- MC works well with function approximation and is less sensitive to initial values, while TD may face challenges with function approximation and is more sensitive to initial values.

Bootstrapping, where updates involve an estimate, is a characteristic of TD and Dynamic Programming (DP), while MC does not bootstrap. Monte Carlo does not rely on Markov assumption. Sampling, where updates do not involve an expected value, is a feature of MC and TD, while DP does not sample.

4.6.3 SARSA

SARSA is an algorithm employed for policy evaluation in reinforcement learning. It operates as an on-policy optimization method, meaning it evaluates and improves the same policy that is used to make decisions. The update rule for SARSA is defined as follows:

$$Q(S_t, A_t) = Q(S_t, A_t) + \alpha(R_{t+1} + \gamma Q(S_{t+1}, A_{t+1}) - Q(S_t, A_t))$$

SARSA is typically paired with an ε -greedy policy for improvement. This means that most of the time, the policy selects the action with the highest estimated value, but occasionally explores other actions with probability ε .

The algorithm requires two parameters: a step size $\alpha \in (0,1]$ and a small $\varepsilon > 0$.

Algorithm 10 SARSA on-policy control algorithm

```
1: Initialize Q(s, a) arbitrarily, for all s \in \mathcal{S}^+, a \in \mathcal{A}(s)
 2: Q(\text{terminal}, \cdot) = 0
 3: loop
       Initialize S
 4:
 5:
        Choose A from S using policy derived from Q
       repeat for each step of episode
 6:
            Take action A, observe R, S'
 7:
            Choose A' from S' using policy derived from Q
 8:
            Q(S, A) = Q(S, A) + \alpha(R + \gamma Q(S', A') - Q(S, A))
 9:
            S = S'
10:
            A = A'
11:
        until S is terminal
12:
13: end loop
```

4.6.4 Q-learning

Q-learning is an algorithm utilized for policy evaluation in reinforcement learning. It operates as an off-policy optimization method, meaning it evaluates a policy while following a different policy for action selection.

As opposed to SARSA, which is a sampled version of the Bellman expectation equation, Q-learning is based on a sampled version of the Bellman optimality equation:

$$Q(S_t, A_t) = Q(S_t, A_t) + \alpha \left(R_{t+1} + \gamma \max_{a} Q(S_{t+1}, a) - Q(S_t, A_t) \right)$$

The algorithm requires two parameters: a step size $\alpha \in (0,1]$ and a small $\varepsilon > 0$.

Algorithm 11 Q-learning algorithm

```
1: Initialize Q(s, a) arbitrarily, for all s \in \mathcal{S}^+, a \in \mathcal{A}(s)
 2: Q(\text{terminal}, \cdot) = 0
 3: loop
 4:
        Initialize S
        repeat for each step of episode
 5:
 6:
            Choose A from S using policy derived from Q
            Take action A, observe R, S'
 7:
            Q(S, A) = Q(S, A) + \alpha \left(R + \gamma \max_{a} Q(S', a) - Q(S, A)\right)
 8:
             S = S'
 9:
        until S is terminal
10:
11: end loop
```

Q-learning updates the Q-values based on the observed rewards and transitions, aiming to find the optimal policy by maximizing the estimated action values over time. Since it's off-policy, it doesn't follow the policy it's evaluating, making it particularly useful in scenarios where exploration and exploitation need to be decoupled.

4.6.5 Eligibility traces

Eligibility traces are a concept in reinforcement learning that play a crucial role in updating the value estimates of states or actions. They are used in combination with temporal difference (TD) learning methods like SARSA (State-Action-Reward-State-Action) or Q-learning. The main features are:

- 1. Temporal credit assignment: eligibility traces help in assigning credit or blame to actions taken in the past for the rewards received in the future.
- 2. Memory mechanism: eligibility traces serve as a memory mechanism that tracks the eligibility of states or actions to be updated based on future rewards. They maintain a record of recent state-action pairs that are likely to contribute to future rewards.
- 3. *Decay factor*: determines how much past experiences influence the current update. It helps balance between short-term and long-term credit assignment.
- 4. *Updating value estimates*: when a reward is received, the eligibility traces are used to update the value estimates of relevant states or actions. This updating process is done more efficiently because the traces highlight which experiences are relevant.
- 5. Efficiency and learning speed: By allowing updates to propagate more efficiently through the learning process, eligibility traces can speed up learning and improve the convergence of RL algorithms.

Overall, eligibility traces enhance the efficiency and effectiveness of temporal difference learning methods by providing a mechanism for assigning credit over time, which is crucial for learning in complex environments with delayed rewards.