## Module 1: Introduction to Bayesian Statistics

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### Agenda

- Motivations
- Traditional inference
- Bayesian inference
- Bernoulli, Beta
- Posterior of Bernoulli-Beta
- Conjugacy
- Example with 2012 election data
- Marginal likelihood
- Posterior Prediction
- ► Additional problems at the end of lecture (derivation + applied)

## What should you learn?

- You should learn the main principles of Bayesian inference/prediction and how to apply these to real data analysis.
- You will continue with this in lab/homework to make sure that you understand these key principles.

#### Traditional inference

You are given data X and there is an **unknown parameter** you wish to estimate  $\theta$ 

How would you estimate  $\theta$ ?

- $\triangleright$  Find an unbiased estimator of  $\theta$ .
- Find the maximum likelihood estimate (MLE) of  $\theta$  by looking at the likelihood of the data.
- Please review unbiased estimation and finding an MLE.
- Please also review other background material such as likelihoods, sufficient statistics, basic probability concepts, etc. Most of this material can be reviewed in Chapters 1-3 in Hoff.

### Bayesian inference

Bayesian methods trace its origin to the 18th century and English Reverend Thomas Bayes, who along with Pierre-Simon Laplace discovered what we now call **Bayes' Theorem** 

- $\triangleright$   $p(x \mid \theta)$  likelihood
- $ightharpoonup p(\theta)$  prior
- $\triangleright p(\theta \mid x)$  posterior
- $\triangleright$  p(x) marginal distribution

How can we derive  $p(\theta \mid x)$ ?

# Derivation of $p(\theta \mid x)$

#### Bernoulli distribution

The Bernoulli distribution is very common due to binary outcomes.

- Consider flipping a coin (heads or tails).
- We can represent this a binary random variable where the probability of heads is  $\theta$  and the probability of tails is  $1-\theta$ .

Consider  $X \sim \text{Bernoulli}(\theta)\mathbb{1}(0 < \theta < 1)$ 

The likelihood is

$$p(x \mid \theta) = \theta^{x} (1 - \theta)^{(1-x)} \mathbb{1}(0 < \theta < 1).$$

Exercise: what is the mean and the variance of X?

#### Bernoulli distribution

▶ Suppose that  $X_1, \ldots, X_n \stackrel{iid}{\sim} \mathsf{Bernoulli}(\theta)$ . Then for  $x_1, \ldots, x_n \in \{0, 1\}$  what is the likelihood?

### **Notation**

- ightharpoonup  $\propto$ : means "proportional to"
- $\triangleright$   $x_{1:n}$  denotes  $x_1, \ldots, x_n$

### Likelihood

$$p(x_{1:n}|\theta) = \mathbb{P}(X_1 = x_1, \dots, X_n = x_n \mid \theta)$$

$$= \prod_{i=1}^n \mathbb{P}(X_i = x_i \mid \theta)$$

$$= \prod_{i=1}^n p(x_i|\theta)$$

$$= \prod_{i=1}^n \theta^{x_i} (1-\theta)^{1-x_i}$$

$$= \theta^{\sum x_i} (1-\theta)^{n-\sum x_i}.$$

#### Beta distribution

Given a,b>0, we write  $\theta \sim \mathrm{Beta}(a,b)$  to mean that  $\theta$  has pdf

$$p(\theta) = \mathsf{Beta}(\theta|\mathsf{a},b) = \frac{1}{B(\mathsf{a},b)} \theta^{\mathsf{a}-1} (1-\theta)^{b-1} \mathbb{1}(0 < \theta < 1),$$

i.e.,  $p(\theta) \propto \theta^{a-1} (1-\theta)^{b-1}$  on the interval from 0 to 1.

► Here,

$$B(a,b) = \frac{\Gamma(a)\Gamma(b)}{\Gamma(a+b)}$$

.

- Parameters a, b control the shape of the distribution.
- This distribution models random behavior of percentages/proportions.

### Posterior of Bernoulli-Beta

Let's derive the posterior of  $\theta \mid x_{1:n}$ 

## Conjugacy

What do you notice about the prior and the posterior from the Bernoulli-Beta example that we just considered?

### Conjugacy

A class P of prior distributions for  $\theta$  is called **conjugate** for the likelihood  $p(x \mid \theta)$  if

$$p(\theta) \in P \implies p(\theta \mid x) \in P.$$

Tip: In practice, we check to see if the posterior has an updated form of the prior.

## Conjugacy

#### **Benefits**

- We do minimal or often no math. In fact, https://en.wikipedia.org/wiki/Conjugate\_prior provides many conjugate families.
- We have an exact posterior distribution. No approximations are needed.
- Computation is fast and simple!

#### **Downside**

Sometimes an unrealistic assumption, however, might provide guidance to us.

### Approval ratings of Obama

What is the proportion of people that approve of President Obama in PA?

- ▶ We take a random sample of 10 people in PA and find that 6 approve of President Obama. Likelihood
- ► The national approval rating (Zogby poll) of President Obama in mid-September 2015 was 50%. We'll assume that in PA his approval rating is also 50%. Prior
- ▶ Based on this prior information, we'll use a Beta prior for  $\theta$  and we'll choose a and b.

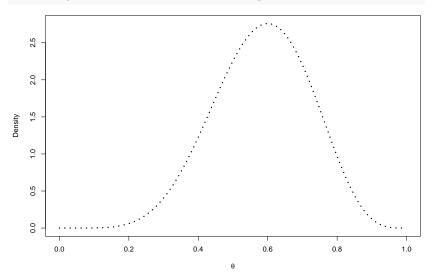
Let's consider choosing a prior distribution that is skewed and see what happens.

### Obama Example

```
n < -10
# Fixing values of a,b.
# I've chosen the prior on Beta to be skewed
\#a = 21/8
#b = 0.04
a < -0.25
b < -0.25
th \leftarrow seq(0,1, length=500)
x < -6
# we set the likelihood, prior, and posteriors with
# THETA as the sequence that we plot on the x-axis.
# Beta(c,d) refers to shape parameter
like \leftarrow dbeta(th, x+1, n-x+1)
prior <- dbeta(th, a, b)</pre>
# print(a/(a + b))
post \leftarrow dbeta(th, x+a, n-x+b)
```

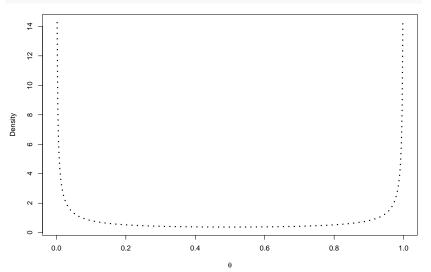
### Likelihood

```
plot(th, like, type='l', ylab = "Density",
    lty = 3, lwd = 3, xlab = expression(theta))
```



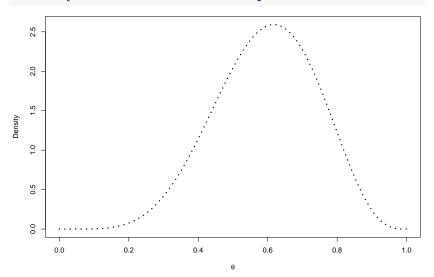
### **Prior**

```
plot(th, prior, type='l', ylab = "Density",
    lty = 3, lwd = 3, xlab = expression(theta))
```

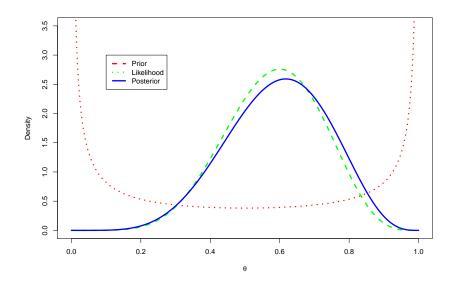


### Posterior

```
plot(th, post, type='1', ylab = "Density",
    lty = 3, lwd = 3, xlab = expression(theta))
```



## Likelihood, Prior, and Posterior



#### Back to the Prior

- ► We choose the prior here very randomly or naively. What happened? Can you think of a simple fix?
- ▶ In the supplemental material (end of lecture), find an example where we have more information and can set *a*, *b* from in a more subjective and principled manner.

#### Cast of characters

- Observed data: x
- This often involves many data points, e.g.,

$$x = x_{1:n} = (x_1, \ldots, x_n).$$

```
likelihood p(x_{1:n}|\theta) prior p(\theta) posterior p(\theta|x_{1:n}) marginal likelihood p(x_{1:n}) posterior predictive p(x_{n+1}|x_{1:n})
```

## Marginal likelihood

The marginal likelihood is defined as

$$p(x) = \int p(x|\theta)p(\theta) d\theta$$

### Example: Back to the Bernoulli-Beta

$$X_1,\dots,X_n\mid heta\stackrel{\mathit{iid}}{\sim} \mathit{Bernoulli}( heta)$$
 and  $heta\sim \mathit{Beta}(a,b).$ 

What is the marginal likelihood for the Bernoulli-Beta?

## Marginal Likelihood: Bernoulli-Beta

Then the marginal likelihood is

$$p(x_{1:n})$$

$$= \int p(x_{1:n}|\theta)p(\theta) d\theta$$

$$= \int_{0}^{1} \theta^{\sum x_{i}} (1-\theta)^{n-\sum x_{i}} \frac{1}{B(a,b)} \theta^{a-1} (1-\theta)^{b-1} d\theta$$

$$= \frac{1}{B(a,b)} \int_{0}^{1} \theta^{\sum x_{i}+a-1} (1-\theta)^{n-\sum x_{i}+b-1} d\theta$$

$$= \frac{\frac{B(a+\sum x_{i}, b+n-\sum x_{i})}{B(a,b)} \int_{0}^{1} \frac{\theta^{\sum x_{i}+a-1} (1-\theta)^{n-\sum x_{i}+b-1}}{B(a+\sum x_{i}, b+n-\sum x_{i})} d\theta$$

$$= \frac{B(a+\sum x_{i}, b+n-\sum x_{i})}{B(a,b)},$$

by the integral definition of the Beta function.

### Posterior predictive distribution

- ▶ We may wish to predict a new data point  $x_{n+1}$
- Assumption 1: Assume that  $x_{1:(n+1)}$  are independent given  $\theta$

$$\begin{split} \rho(x_{n+1}|x_{1:n}) &= \int p(x_{n+1},\theta|x_{1:n}) \, d\theta \\ &\int \frac{p(x_{n+1},\theta,x_{1:n})}{p(x_{1:n})} \, d\theta \quad \text{(Conditional probability)} \\ &\int \frac{p(x_{n+1}|\theta,x_{1:n})p(\theta|x_{1:n})p(x_{1:n})}{p(x_{1:n})} \, d\theta \quad \text{(Product rule)} \\ &= \int p(x_{n+1}|\theta,x_{1:n})p(\theta|x_{1:n}) \, d\theta \\ &= \int p(x_{n+1}|\theta)p(\theta|x_{1:n}) \, d\theta \quad \text{By Assumption 1.} \end{split}$$

### Posterior predictive distribution: Bernoulli-Beta

$$X_1,\ldots,X_n\mid\theta\stackrel{iid}{\sim} Bernoulli(\theta)$$

and

$$\theta \sim Beta(a, b)$$
.

The posterior distribution can be shown to be  $p(\theta|x_{1:n}) = \text{Beta}(\theta|a_n, b_n)$ , where  $a_n = a + \sum x_i$  and  $b_n = b + n - \sum x_i$ .

### Posterior predictive distribution: Bernoulli-Beta

The posterior predictive can be derived to be

$$\mathbb{P}(X_{n+1} = 1 \mid x_{1:n}) = \int \mathbb{P}(X_{n+1} = 1 \mid \theta) p(\theta \mid x_{1:n}) d\theta$$

$$= \int \theta \ \mathsf{Beta}(\theta \mid a_n, b_n) d\theta$$

$$= \frac{a_n}{a_n + b_n} \ \ (\mathsf{Mean of Beta distribution}).$$

Similarly,

$$\mathbb{P}(X_{n+1}=0\mid x_{1:n})=1-\mathbb{P}(X_{n+1}=1\mid x_{1:n})=\frac{b_n}{a_n+b_n}.$$

## Posterior predictive distribution (continued)

This implies that

$$p(x_{n+1}|x_{1:n}) = \begin{cases} \frac{a_n}{a_n + b_n} & \text{if } x_{n+1} = 1\\ \frac{b_n}{a_n + b_n} & \text{if } x_{n+1} = 0 \end{cases}$$

More formally,

$$p(x_{n+1}|x_{1:n}) = \frac{a_n^{x_{n+1}}b_n^{1-x_{n+1}}}{a_n + b_n}\mathbb{1}(x_{n+1} \in \{0,1\}).$$

Either solution above is correct.

## **Overall Summary**

- We covered the "cast of characters" needed to work with Bayesian models
- ► These include the likelihood, prior, posterior, marginal likelihood, and posterior predictive distribution
- We derived Bayes' Theorem
- ► Bernoulli-Beta
- Conjugacy

## Background Knowledge

- ► Familiar with Discrete and Continuous Distributions
- Can calculate expectations and variances
- Change of variables
- Mean squared error
- Sufficiency
- Confident calculating the likelihood and log-likelihood
- Confident in working with partial derivatives
- Familiar maximizing or minimizing functions (and proving they are global max/min)

### Detailed Summary for Exam

- Bayes Theorem
- Likelihood
- Prior
- Posterior derivation
- Marginal likelihood
- Posterior predictive distribution
- Conjugacy
- Proportionality
- Understanding when models are appropriate for data given to you (Ex: Approval ratings for Obama)
- What is an informative prior
- What is a non-informative prior
- Proper posterior
- How do you incorporate a pilot study into your posterior analysis (Ex: See sleep study)

## Supplemental Material

Below you will find supplemental material, such as exercises to help you for the exam with solutions provided.

#### Exercise

We write  $X \sim \mathsf{Poisson}(\theta)$  if X has the Poisson distribution with rate  $\theta > 0$ , that is, its p.m.f. is

$$p(x|\theta) = Poisson(x|\theta) = e^{-\theta}\theta^x/x!$$

for  $x \in \{0, 1, 2, \dots\}$  (and is 0 otherwise). Suppose  $X_1, \dots, X_n \stackrel{iid}{\sim} \mathsf{Poisson}(\theta)$  given  $\theta$ , and your prior is

$$p(\theta) = \mathsf{Gamma}(\theta|a,b) = \frac{b^a}{\Gamma(a)} \theta^{a-1} e^{-b\theta} \mathbb{1}(\theta > 0).$$

What is the posterior distribution on  $\theta$ ?

#### Solution

Since the data is independent given  $\theta$ , the likelihood factors and we get

$$p(x_{1:n}|\theta) = \prod_{i=1}^{n} p(x_i|\theta)$$
$$= \prod_{i=1}^{n} e^{-\theta} \theta^{x_i} / x_i!$$
$$\propto e^{-n\theta} \theta^{\sum x_i}.$$

#### Solution

Thus, using Bayes' theorem,

$$\begin{aligned} \rho(\theta|x_{1:n}) &\propto \rho(x_{1:n}|\theta)\rho(\theta) \\ &\propto e^{-n\theta}\theta^{\sum x_i}\theta^{a-1}e^{-b\theta}\mathbb{1}(\theta>0) \\ &\propto e^{-(b+n)\theta}\theta^{a+\sum x_i-1}\mathbb{1}(\theta>0) \\ &\propto \mathsf{Gamma}\;(\theta\mid a+\sum x_i,\;b+n). \end{aligned}$$

Therefore, since the posterior density must integrate to 1, we have

$$p(\theta|x_{1:n}) = \text{Gamma}(\theta \mid a + \sum x_i, b + n).$$

#### Module 1 Derivations

Class notes from Module 1 can be found below:

 $https://github.com/resteorts/modern-bayes/blob/master/lectures\\ ModernBayes20/lecture-1/notes-module1.pdf$ 

 $https://github.com/resteorts/modern-bayes/blob/master/lectures\\ ModernBayes20/lecture-1/notes-module1.pdf$ 

## Additional Applied Example

Below, there is an additional applied example that you may find useful regarding this material.

## How Much Do You Sleep Example

We are interested in a population of American college students and the proportion of the population that sleep at least eight hours a night, which we denote by  $\theta$ .

### How Much Do You Sleep Example

- ► The Gamecock, at the USC printed an internet article "College Students Don't Get Enough Sleep" (2004).
  - Most students spend six hours sleeping each night.
- 2003: University of Notre Dame's paper, Fresh Writing.
  - ► The article reported took random sample of 100 students:
  - "approximately 70% reported to receiving only five to six hours of sleep on the weekdays,
  - ▶ 28% receiving seven to eight,
  - ▶ and only 2% receiving the healthy nine hours for teenagers."

- Have a random sample of 27 students is taken from UF.
- ▶ 11 students record that they sleep at least eight hours each night.
- **\triangleright** Based on this information, we are interested in estimating  $\theta$ .

- ► From USC and UND, believe it's probably true that most college students get less than eight hours of sleep.
- Want our prior to assign most of the probability to values of  $\theta < 0.5$ .
- From the information given, we decide that our best guess for  $\theta$  is 0.3, although we think it is very possible that  $\theta$  could be any value in [0,0.5].

#### Our Model

Our model can be summarized by the Binomial-Beta distribution

$$X|\theta \sim \mathsf{Binomial}(n,\theta)$$
 (1)

$$\theta \sim \text{Beta}(a, b)$$
 (2)

You can show that the posterior of

$$\theta \mid X \sim \text{Beta}(x + a, n - x + b)$$

### Choice of a,b for Beta Prior

- ▶ Given this information, we believe that the median of  $\theta$  is 0.3 and the 90th percentile is 0.5.
- Knowing this allows us to estimate the unknown values of a and b.
- ► How do we actually calculate *a* and *b*?

### Choice of a,b for Beta Prior

We would need to solve the following equations:

$$\int_0^{0.3} \frac{\Gamma(a+b)}{\Gamma(a)\Gamma(b)} \theta^{a-1} (1-\theta)^{b-1} d\theta = 0.5$$
$$\int_0^{0.5} \frac{\Gamma(a+b)}{\Gamma(a)\Gamma(b)} \theta^{a-1} (1-\theta)^{b-1} d\theta = 0.9$$

In non-calculus language, this means the 0.5 quantile (50th percentile)=0.3. The 0.9 quantile (90th percentile) = 0.5.

The equations are written as percentiles above!

- ▶ We can easily solve this numerically in R using a numerical solver BBsolve using the BB package. .
- The documentation for this package is not great, so beware.

```
#load the BB package
library(BB)
## using percentiles
myfn <- function(shape){</pre>
    test \leftarrow pbeta(q = c(0.3, 0.5), shape1 = shape[1],
     shape2 = shape[2]) - c(0.5, 0.9)
    return(test)
BBsolve(c(1,1), myfn)
##
     Successful convergence.
## $par
## [1] 3.263743 7.185121
##
## $residual
## [1] 5.905161e-08
##
```

Using our calculations from the Beta-Binomial our model is

$$X \mid \theta \sim \text{Binomial}(27, \theta)$$
  
 $\theta \sim \text{Beta}(3.3, 7.2)$   
 $\theta \mid x \sim \text{Beta}(x + 3.3, 27 - x + 7.2)$   
 $\theta \mid 11 \sim \text{Beta}(14.3, 23.2)$ 

```
th = seq(0,1,length=500)
a = estimated *par[1]
b = estimated$par[2]
n = 27
x = 11
prior = dbeta(th,a,b)
like = dbeta(th,x+1,n-x+1)
post = dbeta(th,x+a,n-x+b)
plot(th,post,type="l",ylab="Density",lty=2,lwd=3,
xlab = expression(theta))
lines(th,like,lty=1,lwd=3)
lines(th,prior,lty=3,lwd=3)
legend(0.7,4,c("Prior","Likelihood","Posterior"),
lty=c(3,1,2), lwd=c(3,3,3))
```



