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Statistics C173/C273

### Ordinary kriging

Kriging (Matheron 1963) owes its name to D. G. Krige a South African mining engineer and it was first applied in mining data. Kriging assumes a random field expressed through a variogram or covariance function. It is a very popular method to solve the spatial prediction problem. Let  $\mathbf{Z} = (Z(s_1), Z(s_2), ..., Z(s_n))'$  be the vector of the observed data at known spatial locations  $s_1, s_2, ..., s_n$ . The objective is to estimate the unobserved value  $Z(s_0)$  at location  $s_0$ .

### The model:

The model assumption is:

$$Z(s) = \mu + \delta(s)$$

where  $\delta(s)$  is a zero mean stochastic term with variogram  $2\gamma(\cdot)$ . The variogram was discussed in previous handouts in detail.

### The Kriging System

The predictor assumption is

$$\hat{Z}(s_0) = \sum_{i=1}^n w_i Z(s_i)$$

i.e. it is a weighted average of the sample values, and  $\sum_{i=1}^{n} w_i = 1$  to ensure unbiasedness. The  $w_i$ 's are the weights that will be estimated.

Kriging minimizes the mean squared error of prediction

min 
$$\sigma_e^2 = E[Z(s_0) - \hat{Z}(s_0)]^2$$

or

min 
$$\sigma_e^2 = E \left[ Z(s_0) - \sum_{i=1}^n w_i Z(s_i) \right]^2$$

For intrinsically stationary process the last equation can be written as:

$$\sigma_e^2 = 2\sum_{i=1}^n w_i \gamma(s_0 - s_i) - \sum_{i=1}^n \sum_{j=1}^n w_i w_j \gamma(s_i - s_j)$$
 (1)

See next page for the proof:

Let's examine  $(Z(s_0) - \sum_{i=1}^n w_i Z(s_i))^2$ :

$$\left(z(s_0) - \sum_{i=1}^n w_i z(s_i)\right)^2 =$$

$$z^2(s_0) - 2z(s_0) \sum_{i=1}^n w_i z(s_i) + \sum_{i=1}^n \sum_{j=1}^n w_i w_j z(s_i) z(s_j) =$$

$$\sum_{i=1}^n w_i z^2(s_0) - 2 \sum_{i=1}^n w_i z(s_0) z(s_i) + \sum_{i=1}^n \sum_{j=1}^n w_i w_j z(s_i) z(s_j)$$

$$-\frac{1}{2} \sum_{i=1}^n w_i z^2(s_i) - \frac{1}{2} \sum_{j=1}^n w_j z^2(s_j) + \sum_{i=1}^n w_i z^2(s_i) =$$

$$-\frac{1}{2} \sum_{i=1}^n \sum_{j=1}^n w_i w_j \left[z(s_i) - z(s_j)\right]^2 + \sum_{i=1}^n w_i \left[z(s_0) - z(s_i)\right]^2$$

If we take expectations on the last expression we have

$$-\frac{1}{2}\sum_{i=1}^{n}\sum_{j=1}^{n}w_{i}w_{j}E\left[z(s_{i})-z(s_{j})\right]^{2} + \sum_{i=1}^{n}w_{i}E\left[z(s_{0})-z(s_{i})\right]^{2} =$$

$$-\frac{1}{2}\sum_{i=1}^{n}\sum_{j=1}^{n}w_{i}w_{j}\operatorname{var}\left[z(s_{i})-z(s_{j})\right] + \sum_{i=1}^{n}w_{i}\operatorname{var}\left[z(s_{0})-z(s_{i})\right]$$

But var  $[z(s_i) - z(s_j)] = 2\gamma(\cdot)$  is the definition of the variogram, and therefore the previous expression is written as:  $2\sum_{i=1}^{n} w_i \gamma(s_0 - s_i) - \sum_{i=1}^{n} \sum_{j=1}^{n} w_i w_j \gamma(s_i - s_j)$ Therefore kriging minimizes

$$\sigma_e^2 = E[(Z(s_0) - \sum_{i=1}^n w_i Z(s_i)]^2 = 2\sum_{i=1}^n w_i \gamma(s_0 - s_i) - \sum_{i=1}^n \sum_{j=1}^n w_i w_j \gamma(s_i - s_j)$$
subject to
$$\sum_{i=1}^n w_i = 1$$

The minimization is carried out over  $(w_1, w_2, ..., w_n)$ , subject to the constraint  $\sum_{i=1}^n w_i = 1$ . Therefore the minimization problem can be written as:

$$\min \ 2\sum_{i=1}^{n} w_i \gamma(s_0 - s_i) - \sum_{i=1}^{n} \sum_{j=1}^{n} w_i w_j \gamma(s_i - s_j) - 2\lambda(\sum_{i=1}^{n} w_i - 1)$$
 (2)

where  $\lambda$  is the Lagrange multiplier. After differentiating (2) with respect to  $w_1, w_2, ..., w_n$ , and  $\lambda$  and set the derivatives equal to zero we find that

$$-\sum_{i=1}^{n} w_j \gamma(s_i - s_j) + \gamma(s_0 - s_i) - \lambda = 0, \quad i = 1, ..., n$$

and

$$\sum_{i=1}^{n} w_i = 1$$

Using matrix notation the previous system of equations can be written as

$$\mathbf{\Gamma}\mathbf{W}=oldsymbol{\gamma}$$

Therefore the weights  $w_1, w_2, ..., w_n$  and the Lagrange multiplier  $\lambda$  can be obtained by

$$\mathbf{W} = \mathbf{\Gamma^{-1} \gamma}$$

where

$$\mathbf{W} = (w_1, w_2, ..., w_n, \lambda)$$

$$\mathbf{\gamma} = (\gamma(s_0 - s_1), \gamma(s_0 - s_2), ..., \gamma(s_0 - s_n), 1)'$$

$$\mathbf{\Gamma} = \begin{cases} \gamma(s_i - s_j), & i = 1, 2, ..., n, \quad j = 1, 2, ..., n, \\ 1, & i = n + 1, \quad j = 1, ..., n, \\ 1, & j = n + 1, \quad i = 1, ..., n, \\ 0, & i = n + 1, \quad i = n + 1 \end{cases}$$

### The variance of the estimator:

So far, we found the weights and therefore we can compute the estimator:  $\hat{Z}(s_0) = \sum_{i=1}^n w_i Z(s_i)$ . How about the variance of the estimator, namely  $\sigma_e^2$ ?

We multiply

$$-\sum_{j=1}^{n} w_j \gamma(s_i - s_j) + \gamma(s_0 - s_i) - \lambda = 0$$

by  $w_i$  and we sum over all  $i = 1, \dots, n$  to get:

$$-\sum_{i=1}^{n} w_i \sum_{j=1}^{n} w_j \gamma(s_i - s_j) + \sum_{i=1}^{n} w_i \gamma(s_0 - s_i) - \sum_{i=1}^{n} w_i \lambda = 0$$

Or

$$-\sum_{i=1}^{n} \sum_{j=1}^{n} w_i w_j \gamma(s_i - s_j) + \sum_{i=1}^{n} w_i \gamma(s_0 - s_i) - \sum_{i=1}^{n} w_i \lambda = 0$$

Therefore,

$$\sum_{i=1}^{n} \sum_{j=1}^{n} w_i w_j \gamma(s_i - s_j) = \sum_{i=1}^{n} w_i \gamma(s_0 - s_i) - \sum_{i=1}^{n} w_i \lambda$$

If we substitute this result into equation (1) we finally get:

$$\sigma_e^2 = \sum_{i=1}^n w_i \gamma(s_i - s_0) + \lambda \tag{3}$$

# The Kriging System

$(\gamma(s_0-s_1)^{\gamma}$	$\gamma(s_0 - s_2)$			$\gamma(s_0-s_n)$	П	
II						
$\begin{pmatrix} w_1 \end{pmatrix}$	$w_2$			$w_n$	<	
_						
$\vdash$	$\vdash$	$\vdash$	Т	П	0	
$\gamma(s_1-s_n)$ 1	$\gamma(s_2-s_n)$	÷	÷	$\gamma(s_n - s_n)$	1	
:	:	:	··	:	:	
$\gamma(s_1-s_3)$	$\gamma(s_2-s_3)$	÷		$\gamma(s_n-s_3)$	:	
$\gamma(s_1-s_2)$	$\gamma(s_2-s_2)$	:		$\gamma(s_n-s_2)$	1	
$\left( \gamma(s_1 - s_1) \right)$	$\gamma(s_2 - s_1)$	:		$\gamma(s_n - s_1)$	1	

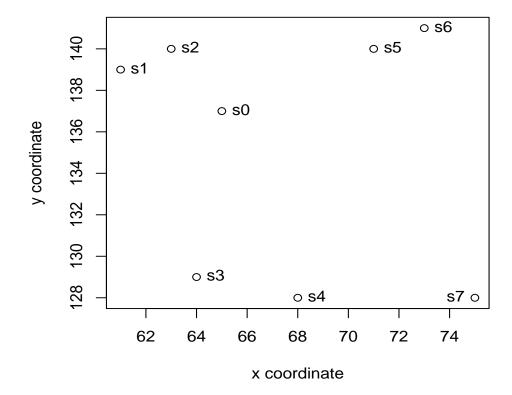
and uniqueness of the solution to the kriging system. In other words the sample variogram version of  $\Gamma$  might not be positive definite. To be guaranteed of having one and only one solution, we must ensure that our system has the property of positive We understand now why the sample variogram cannot be used directly in the kriging system, and instead a theoretical variogram is used. First, the  $\gamma$  vector may call for variogram values for distances that are not available from sample data. There are situations where the distance from the point being estimated to a particular sample is smaller than the distance between pairs of available samples. Since the sample data set cannot provide any pairs for these small distances, we must rely on a function that provides variogram values for all distances. Second, the use of the sample variogram does not guarantee the existence definiteness. This is ensured by the choice of one of the model variograms mentioned earlier.

## A simple example:

Consider the following data

$s_i$	$\boldsymbol{x}$	y	$z(s_i)$
$s_1$	61	139	477
$s_2$	63	140	696
$s_3$	64	129	227
$s_4$	68	128	646
$s_5$	71	140	606
$s_6$	73	141	791
$s_7$	75	128	783
$\mathbf{s_0}$	65	137	???

Our goal is to estimate the unknown value at location  $s_0$ . Here is the x-y plot:



For these data, let's assume that we use the exponential semivariogram model with parameters  $c_0 = 0, c_1 = 10, \alpha = 3.33$ .

$$\gamma(h) = 10(1 - e^{-\frac{h}{3.33}}).$$

We need to construct the matrix  $\Gamma$  and the vector  $\gamma$ . First we calculate the distance matrix as shown below:

$$\text{Distance matrix} = \begin{pmatrix} s_0 & s_1 & s_2 & s_3 & s_4 & s_5 & s_6 & s_7 \\ s_0 & 0.00 & 4.47 & 3.61 & 8.06 & 9.49 & 6.71 & 8.94 & 13.45 \\ s_1 & 4.47 & 0.00 & 2.24 & 10.44 & 13.04 & 10.05 & 12.17 & 17.80 \\ s_2 & 3.61 & 2.24 & 0.00 & 11.05 & 13.00 & 8.00 & 10.05 & 16.97 \\ s_3 & 8.06 & 10.44 & 11.05 & 0.00 & 4.12 & 13.04 & 15.00 & 11.05 \\ s_4 & 9.49 & 13.04 & 13.00 & 4.12 & 0.00 & 12.37 & 13.93 & 7.00 \\ s_5 & 6.71 & 10.05 & 8.00 & 13.04 & 12.37 & 0.00 & 2.24 & 12.65 \\ s_6 & 8.94 & 12.17 & 10.05 & 15.00 & 13.93 & 2.24 & 0.00 & 13.15 \\ s_7 & 13.45 & 17.80 & 16.90 & 11.05 & 7.00 & 2.65 & 13.15 & 0.00 \end{pmatrix}$$

The  $ij_{th}$  entry in the matrix above was computed as follows:

$$d_{ij} = \sqrt{(x_i - x_j)^2 + (y_i - y_j)^2}$$

Now, we can find the entries of the matrix  $\Gamma$  and the vector  $\gamma$ . Each entry will be computed using the exponential semivariogram  $\gamma(h) = 10(1 - e^{-\frac{h}{3.33}})$ . Here they are:

$$\Gamma = \begin{pmatrix} 0 & 4.893 & 9.564 & 9.800 & 9.510 & 9.740 & 9.952 & 1 \\ 4.893 & 0 & 9.637 & 9.798 & 9.093 & 9.510 & 9.938 & 1 \\ 9.564 & 9.637 & 0 & 7.095 & 9.800 & 9.889 & 9.637 & 1 \\ 9.800 & 9.798 & 7.095 & 0 & 9.755 & 9.847 & 8.775 & 1 \\ 9.510 & 9.093 & 9.800 & 9.755 & 0 & 4.893 & 9.775 & 1 \\ 9.740 & 9.510 & 9.889 & 9.847 & 4.893 & 0 & 9.806 & 1 \\ 9.952 & 9.938 & 9.637 & 8.775 & 9.775 & 9.806 & 0 & 1 \\ 1 & 1 & 1 & 1 & 1 & 1 & 1 & 0 \end{pmatrix}$$

$$\gamma = \begin{pmatrix}
7.384 \\
6.614 \\
9.109 \\
9.420 \\
8.664 \\
9.316 \\
9.823 \\
1
\end{pmatrix}$$

The weights and the Lagrange multiplier can be obtained as follows:

$$\boldsymbol{W} = \boldsymbol{\Gamma}^{-1} \boldsymbol{\gamma} = \begin{pmatrix} 0 & 4.893 & 9.564 & 9.800 & 9.510 & 9.740 & 9.952 & 1 \\ 4.893 & 0 & 9.637 & 9.798 & 9.093 & 9.510 & 9.938 & 1 \\ 9.564 & 9.637 & 0 & 7.095 & 9.800 & 9.889 & 9.637 & 1 \\ 9.800 & 9.798 & 7.095 & 0 & 9.755 & 9.847 & 8.775 & 1 \\ 9.510 & 9.093 & 9.800 & 9.755 & 0 & 4.893 & 9.775 & 1 \\ 9.740 & 9.510 & 9.889 & 9.847 & 4.893 & 0 & 9.806 & 1 \\ 9.952 & 9.938 & 9.637 & 8.775 & 9.775 & 9.806 & 0 & 1 \\ 1 & 1 & 1 & 1 & 1 & 1 & 1 & 0 \end{pmatrix}^{-1} \begin{pmatrix} 7.384 \\ 6.614 \\ 9.109 \\ 9.420 \\ 8.664 \\ 9.316 \\ 9.823 \\ 1 \end{pmatrix}$$

The answer is:

$$\mathbf{W} = \begin{pmatrix} 0.174 \\ 0.317 \\ 0.129 \\ 0.086 \\ 0.151 \\ 0.057 \\ 0.086 \\ 0.906 \end{pmatrix}.$$

The last element of the **W** vector is the Lagrange multiplier,  $\lambda = 0.906$ . We can verify that the sum of the elements 1 through 7 is equal to 1, as it should be.

The predicted value at location  $s_0$  is equal to:

$$\hat{z}(s_0) = \sum_{i=1}^n w_i z(s_i) = 0.174(477) + \dots + 0.086(783) = 592.59.$$

And the variance:

$$\sigma_e^2 = \sum_{i=1}^n w_i \gamma(s_i - s_0) + \lambda = 0.174(7.384) + \dots + 0.086(9.823) + 0.906 = 8.96.$$

Under the assumption that Z(s) is Gaussian a 95% confidence interval can be computed as follows:

$$592.59 \pm 1.96\sqrt{8.96}$$

Or

$$577.09 \le Z(s_0) \le 588.83$$

### Ordinary kriging using the covariance function:

The same example is solved using the covariance function:  $C(h) = 10e^{-\frac{h}{3.33}}$ . Reminder: The corresponding semivariogram function was  $\gamma(h) = 10(1 - e^{-\frac{h}{3.33}})$ .

$$\boldsymbol{W} = \boldsymbol{C^{-1}c} = \begin{pmatrix} 10 & 5.103 & 0.435 & 0.199 & 0.489 & 0.259 & 0.048 & 1 \\ 5.103 & 10 & 0.362 & 0.202 & 0.905 & 0.489 & 0.061 & 1 \\ 0.435 & 0.362 & 10 & 2.902 & 0.199 & 0.111 & 0.362 & 1 \\ 0.199 & 0.202 & 2.902 & 10 & 0.244 & 0.152 & 1.222 & 1 \\ 0.489 & 0.905 & 0.199 & 0.244 & 10 & 5.103 & 0.224 & 1 \\ 0.259 & 0.489 & 0.111 & 0.152 & 5.103 & 10 & 0.193 & 1 \\ 0.048 & 0.061 & 0.362 & 1.222 & 0.224 & 0.193 & 10 & 1 \\ 1 & 1 & 1 & 1 & 1 & 1 & 1 & 0 \end{pmatrix}^{-1} \begin{pmatrix} 2.612 \\ 3.382 \\ 0.889 \\ 0.579 \\ 1.333 \\ 0.682 \\ 0.176 \\ 1 \end{pmatrix}.$$

The answer is:

$$\mathbf{W} = \begin{pmatrix} 0.174 \\ 0.317 \\ 0.129 \\ 0.086 \\ 0.151 \\ 0.057 \\ 0.086 \\ -0.906 \end{pmatrix} = \begin{pmatrix} w_1 \\ w_2 \\ w_3 \\ w_4 \\ w_5 \\ w_6 \\ w_7 \\ -\lambda \end{pmatrix}.$$

We observe that the weights are same as before when we used the variogram, and therefore

$$\hat{z}(s_0) = \sum_{i=1}^n w_i z(s_i) = 0.174(477) + \dots + 0.086(783) = 592.59.$$

The Lagrange multiplier is now  $\lambda = -0.907$ . However the variance of the estimator is the same as before because:

$$\sigma_e^2 = C(0) - \sum_{i=1}^n w_i C(s_i - s_0) + \lambda = 8.96.$$

### Short code for ordinary kriging in terms of covariance:

```
a <- read.table("kriging_1.txt", header=TRUE) #See next page.
b <- read.table("kriging_11.txt", header=TRUE) #See next page.
x <- as.matrix(cbind(a$x, a$y))</pre>
x1 \leftarrow rep(rep(0,8),8) #Initialize
dist <- matrix(x1,nrow=8,ncol=8) #the distance matrix</pre>
for (i in 1:8){
for (j in 1:8){
dist[i,j]=((x[i,1]-x[j,1])^2+(x[i,2]-x[j,2])^2)^.5
}
c0 <- 0
c1 <- 10
alpha <- 3.33
x1 \leftarrow rep(rep(0,8),8)
                                   #Initialize
C <- matrix(x1,nrow=8,ncol=8) #the C matrix</pre>
for(i in 1:8){
for (j in 1:8){
C[i,j]=c1*exp(-dist[i,j]/alpha)
                if(i==j)\{C[i,j]=c0+c1\}
if(i==8){C[i,j]=1}
if(j==8){C[i,j]=1}
if(i==8 \& j==8) \{C[i,j]=0\}
}
}
c \leftarrow rep(0,8)
                                  #Initialize
                                   #the c vector
for(j in 1:8){
c[j]=c1*exp(-dist[8,j]/alpha)
if(j == 8) \{c[j]=1\}
}
w <- solve(C) %*% c #Obtain the weights and the Lagrange parameter
z_{hat} \leftarrow w[-8] \% \% b$z #Compute the estimate
var_z_hat \leftarrow c0+c1 - t(w) \%*\% c #Compute the variance of the estimate
```

File kriging\_1.txt: All the coordinates (observed data points plus the point to be predicted).

- x y 61 139 63 140
- 64 129
- 68 128
- 71 140
- 73 141
- 75 128
- 65 137

File kriging\_11.txt: Observed data points and their coordinates.

- x y z
- 61 139 477
- 63 140 696
- 64 129 227
- 68 128 646
- 71 140 606
- 73 141 791
- 75 128 783