

Stat Review

HUDM 6122: Multivariate Analysis

- **Expected value**
- **Linear Transformations**
- **Covariance**

Mean, Variance, Covariance, Correlation

$X_1, X_2 \equiv$ Random Variables $a, b, c \equiv$ constants

$$E\{X_1\} = \mu_1 \quad V\{X_1\} = E\{(X_1 - \mu_1)^2\} = \sigma_1^2 = \sigma_{11} \quad E\{X_2\} = \mu_2 \quad V\{X_2\} = \sigma_2^2 = \sigma_{22}$$

$$\text{COV}\{X_1, X_2\} = E\{(X_1 - \mu_1)(X_2 - \mu_2)\} = \sigma_{12}$$

$$E\{cX_1\} = cE\{X_1\} = c\mu_1 \quad V\{cX_1\} = E\{(cX_1 - c\mu_1)^2\} = c^2 E\{(X_1 - \mu_1)^2\} = c^2 \sigma_{11}$$

$$\text{COV}\{aX_1, bX_2\} = E\{(aX_1 - a\mu_1)(bX_2 - b\mu_2)\} = abE\{(X_1 - \mu_1)(X_2 - \mu_2)\} = ab\sigma_{12}$$

$$E\{aX_1 + bX_2\} = aE\{X_1\} + bE\{X_2\} = a\mu_1 + b\mu_2$$

$$\begin{aligned} V\{aX_1 + bX_2\} &= E\left\{\left[(aX_1 + bX_2) - (a\mu_1 + b\mu_2)\right]^2\right\} = E\left\{\left[a(X_1 - \mu_1) + b(X_2 - \mu_2)\right]^2\right\} \\ &= E\left\{\left[a^2(X_1 - \mu_1)^2 + b^2(X_2 - \mu_2)^2 + 2ab(X_1 - \mu_1)(X_2 - \mu_2)\right]\right\} = a^2\sigma_{11} + b^2\sigma_{22} + 2ab\sigma_{12} \end{aligned}$$

$$\text{CORR}\{aX_1, bX_2\} = \frac{\text{COV}\{aX_1, bX_2\}}{\sqrt{V\{aX_1\}}\sqrt{V\{bX_2\}}} = \frac{ab\sigma_{12}}{\sqrt{a^2\sigma_{11}}\sqrt{b^2\sigma_{22}}} = \frac{\sigma_{12}}{\sqrt{\sigma_{11}}\sqrt{\sigma_{22}}} = \text{CORR}\{X_1, X_2\} = \rho_{12}$$

Random Vectors and Matrices

$$\mathbf{X} = \begin{bmatrix} X_{11} & X_{12} & \cdots & X_{1p} \\ X_{21} & X_{22} & \cdots & X_{2p} \\ \vdots & \vdots & \ddots & \vdots \\ X_{n1} & X_{n2} & \cdots & X_{np} \end{bmatrix} \quad E\{\mathbf{X}\} = \begin{bmatrix} E\{X_{11}\} & E\{X_{12}\} & \cdots & E\{X_{1p}\} \\ E\{X_{21}\} & E\{X_{22}\} & \cdots & E\{X_{2p}\} \\ \vdots & \vdots & \ddots & \vdots \\ E\{X_{n1}\} & E\{X_{n2}\} & \cdots & E\{X_{np}\} \end{bmatrix} \quad \mathbf{A}, \mathbf{B} \text{ constants } \Rightarrow E\{\mathbf{AXB}\} = \mathbf{AE}\{\mathbf{X}\}\mathbf{B}$$

$r_A \times n \quad p \times c_B$

Random Vectors: Shown for case of $n=3$, generalizes to any n :

$$\text{Random variables: } X_1, X_2, X_3 \Rightarrow \mathbf{X} = \begin{bmatrix} X_1 \\ X_2 \\ X_3 \end{bmatrix} \quad \text{Expectation: } \mathbf{E}\{\mathbf{X}\} = \begin{bmatrix} E\{X_1\} \\ E\{X_2\} \\ E\{X_3\} \end{bmatrix} = \boldsymbol{\mu}_X$$

Variance-Covariance Matrix for a Random Vector:

$$\sigma^2\{\mathbf{X}\} = E\left\{\left[\mathbf{X} - \mathbf{E}\{\mathbf{X}\}\right]\left[\mathbf{X} - \mathbf{E}\{\mathbf{X}\}\right]'\right\} = \mathbf{E}\left\{\begin{bmatrix} X_1 - E\{X_1\} \\ X_2 - E\{X_2\} \\ X_3 - E\{X_3\} \end{bmatrix} \begin{bmatrix} X_1 - E\{X_1\} & X_2 - E\{X_2\} & X_3 - E\{X_3\} \end{bmatrix}\right\} =$$

$$= \mathbf{E}\left\{\begin{bmatrix} (X_1 - E\{X_1\})^2 & (X_1 - E\{X_1\})(X_2 - E\{X_2\}) & (X_1 - E\{X_1\})(X_3 - E\{X_3\}) \\ (X_2 - E\{X_2\})(X_1 - E\{X_1\}) & (X_2 - E\{X_2\})^2 & (X_2 - E\{X_2\})(X_3 - E\{X_3\}) \\ (X_3 - E\{X_3\})(X_1 - E\{X_1\}) & (X_3 - E\{X_3\})(X_2 - E\{X_2\}) & (X_3 - E\{X_3\})^2 \end{bmatrix}\right\} = \begin{bmatrix} \sigma_{11} & \sigma_{12} & \sigma_{13} \\ \sigma_{21} & \sigma_{22} & \sigma_{23} \\ \sigma_{31} & \sigma_{32} & \sigma_{33} \end{bmatrix} = \boldsymbol{\Sigma}_X$$

Independence and covariance

The p random variables X_1, X_2, \dots, X_p are (statistically) *independent* if their joint density can be factored as

$$f(x_1, x_2, \dots, x_p) = f_1(x_1)f_2(x_2) \dots f_p(x_p)$$

Theorem: If X_i and X_k are independent, then
 $\text{cov}(X_i, X_k) = 0$

Note: The converse is not true in general!

Mean and Variance of Linear Functions of \mathbf{X}

$\mathbf{A} \equiv$ matrix of fixed constants $\mathbf{X} \equiv$ random vector
 $k \times p$ $p \times 1$

$$\mathbf{W} = \mathbf{A}\mathbf{X} = \begin{bmatrix} a_{11} & \cdots & a_{1p} \\ \vdots & & \vdots \\ a_{k1} & \cdots & a_{kp} \end{bmatrix} \begin{bmatrix} X_1 \\ \vdots \\ X_p \end{bmatrix} \equiv \text{random vector: } \mathbf{W} = \begin{bmatrix} W_1 \\ \vdots \\ W_k \end{bmatrix} = \begin{bmatrix} a_{11}X_1 + \dots + a_{1p}X_p \\ \vdots \\ a_{k1}X_1 + \dots + a_{kp}X_p \end{bmatrix}$$

$$\begin{aligned} \mathbf{E}\{\mathbf{W}\} &= \begin{bmatrix} E\{W_1\} \\ \vdots \\ E\{W_k\} \end{bmatrix} = \begin{bmatrix} E\{a_{11}X_1 + \dots + a_{1p}X_p\} \\ \vdots \\ E\{a_{k1}X_1 + \dots + a_{kp}X_p\} \end{bmatrix} = \begin{bmatrix} a_{11}E\{X_1\} + \dots + a_{1p}E\{X_p\} \\ \vdots \\ a_{k1}E\{X_1\} + \dots + a_{kp}E\{X_p\} \end{bmatrix} = \\ &= \begin{bmatrix} a_{11} & \cdots & a_{1p} \\ \vdots & & \vdots \\ a_{k1} & \cdots & a_{kp} \end{bmatrix} \begin{bmatrix} E\{X_1\} \\ \vdots \\ E\{X_p\} \end{bmatrix} = \mathbf{A}\mathbf{E}\{\mathbf{X}\} = \mathbf{A}\boldsymbol{\mu}_{\mathbf{X}} \end{aligned}$$

$$\begin{aligned} \sigma^2\{\mathbf{W}\} &= \mathbf{E}\left\{\left[\mathbf{A}\mathbf{X} - \mathbf{A}\mathbf{E}\{\mathbf{X}\}\right]\left[\mathbf{A}\mathbf{X} - \mathbf{A}\mathbf{E}\{\mathbf{X}\}\right]'\right\} = \mathbf{E}\left\{\left[\mathbf{A}(\mathbf{X} - \mathbf{E}\{\mathbf{X}\})\right]\left[\mathbf{A}(\mathbf{X} - \mathbf{E}\{\mathbf{X}\})\right]'\right\} = \\ &= \mathbf{E}\left\{\left[\mathbf{A}(\mathbf{X} - \mathbf{E}\{\mathbf{X}\})\right]\left[(\mathbf{X} - \mathbf{E}\{\mathbf{X}\})'\mathbf{A}'\right]\right\} = \mathbf{A}\mathbf{E}\left\{(\mathbf{X} - \mathbf{E}\{\mathbf{X}\})(\mathbf{X} - \mathbf{E}\{\mathbf{X}\})'\right\}\mathbf{A}' = \mathbf{A}\boldsymbol{\Sigma}_{\mathbf{X}}\mathbf{A}' \end{aligned}$$

Standard Deviation and Correlation Matrices

Variance-Covariance Matrix: $V\{\mathbf{X}\} = \mathbf{\Sigma}_x = \begin{bmatrix} \sigma_{11} & \sigma_{12} & \cdots & \sigma_{1p} \\ \sigma_{12} & \sigma_{22} & \cdots & \sigma_{2p} \\ \vdots & \vdots & \ddots & \vdots \\ \sigma_{1p} & \sigma_{2p} & \cdots & \sigma_{pp} \end{bmatrix}$

Standard Deviation Matrix: $\mathbf{V}^{1/2} = \text{diag}\{\sqrt{\sigma_{ii}}\} = \begin{bmatrix} \sqrt{\sigma_{11}} & 0 & \cdots & 0 \\ 0 & \sqrt{\sigma_{22}} & \cdots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \cdots & \sqrt{\sigma_{pp}} \end{bmatrix}$ $\mathbf{V}^{-1/2} = \text{diag}\{1/\sqrt{\sigma_{ii}}\} = \begin{bmatrix} 1/\sqrt{\sigma_{11}} & 0 & \cdots & 0 \\ 0 & 1/\sqrt{\sigma_{22}} & \cdots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \cdots & 1/\sqrt{\sigma_{pp}} \end{bmatrix}$

Correlation: $\rho_{ij} = \frac{\sigma_{ij}}{\sqrt{\sigma_{ii}}\sqrt{\sigma_{jj}}}$ Correlation Matrix: $\mathbf{\rho} = \begin{bmatrix} 1 & \rho_{12} & \cdots & \rho_{1p} \\ \rho_{12} & 1 & \cdots & \rho_{2p} \\ \vdots & \vdots & \ddots & \vdots \\ \rho_{1p} & \rho_{2p} & \cdots & 1 \end{bmatrix} = \mathbf{V}^{-1/2}\mathbf{\Sigma}_x\mathbf{V}^{-1/2} \Rightarrow \mathbf{V}^{1/2}\mathbf{\rho}\mathbf{V}^{1/2} = \mathbf{V}^{1/2}\mathbf{V}^{-1/2}\mathbf{\Sigma}_x\mathbf{V}^{-1/2}\mathbf{V}^{1/2} = \mathbf{\Sigma}_x$

Partitioned Covariance Matrix

Suppose the p variables can be split into 2 groups: Group 1: X_1, \dots, X_q Group 2: X_{q+1}, \dots, X_p

$$\mathbf{X}_{p \times 1} = \begin{bmatrix} X_1 \\ \vdots \\ X_q \\ \text{---} \\ X_{q+1} \\ \vdots \\ X_p \end{bmatrix} = \begin{bmatrix} \mathbf{X}^{(1)} \\ \text{---} \\ \mathbf{X}^{(2)} \end{bmatrix} \quad \mu_{\mathbf{X}} = E\{\mathbf{X}\} = \begin{bmatrix} \mu_1 \\ \vdots \\ \mu_q \\ \text{---} \\ \mu_{q+1} \\ \vdots \\ \mu_p \end{bmatrix} = \begin{bmatrix} \mu^{(1)} \\ \text{---} \\ \mu^{(2)} \end{bmatrix}$$

$$E\{(\mathbf{X} - \mu_{\mathbf{X}})(\mathbf{X} - \mu_{\mathbf{X}})'\} = E\left\{ \begin{bmatrix} \mathbf{X}^{(1)} - \mu^{(1)} \\ \text{---} \\ \mathbf{X}^{(2)} - \mu^{(2)} \end{bmatrix} \begin{bmatrix} (\mathbf{X}^{(1)} - \mu^{(1)})' & | & (\mathbf{X}^{(2)} - \mu^{(2)})' \end{bmatrix} \right\} = E\left\{ \begin{bmatrix} (\mathbf{X}^{(1)} - \mu^{(1)})(\mathbf{X}^{(1)} - \mu^{(1)})' & (\mathbf{X}^{(1)} - \mu^{(1)})(\mathbf{X}^{(2)} - \mu^{(2)})' \\ (\mathbf{X}^{(2)} - \mu^{(2)})(\mathbf{X}^{(1)} - \mu^{(1)})' & (\mathbf{X}^{(2)} - \mu^{(2)})(\mathbf{X}^{(2)} - \mu^{(2)})' \end{bmatrix} \right\}$$

$$= \begin{bmatrix} \Sigma_{\mathbf{X}}^{(11)} & \Sigma_{\mathbf{X}}^{(12)} \\ \Sigma_{\mathbf{X}}^{(21)} & \Sigma_{\mathbf{X}}^{(22)} \end{bmatrix} = \Sigma_{\mathbf{X}} = \begin{bmatrix} \sigma_{11} & \cdots & \sigma_{1q} & \sigma_{1,q+1} & \cdots & \sigma_{1p} \\ \vdots & \ddots & \vdots & \vdots & \ddots & \vdots \\ \sigma_{1q} & \cdots & \sigma_{qq} & \sigma_{q,q+1} & \cdots & \sigma_{qp} \\ \sigma_{1,q+1} & \cdots & \sigma_{q,q+1} & \sigma_{q+1,q+1} & \cdots & \sigma_{q+1,p} \\ \vdots & \ddots & \vdots & \vdots & \ddots & \vdots \\ \sigma_{1p} & \cdots & \sigma_{qp} & \sigma_{q+1,p} & \cdots & \sigma_{pp} \end{bmatrix} \quad \text{Note: } \Sigma_{\mathbf{X}}^{(21)} = \Sigma_{\mathbf{X}}^{(12)'} ,$$

Similar Partitioning for sample mean vector and Variance-Covariance Matrix: $\bar{\mathbf{x}} = \begin{bmatrix} \bar{x}_1 \\ \vdots \\ \bar{x}_q \\ \text{---} \\ \bar{x}_{q+1} \\ \vdots \\ \bar{x}_p \end{bmatrix} = \begin{bmatrix} \bar{\mathbf{x}}^{(1)} \\ \text{---} \\ \bar{\mathbf{x}}^{(2)} \end{bmatrix}$

$$\mathbf{S}_n = \begin{bmatrix} \mathbf{S}_n^{(11)} & \mathbf{S}_n^{(12)} \\ \mathbf{S}_n^{(21)} & \mathbf{S}_n^{(22)} \end{bmatrix}$$

Matrix Inequalities and Maximization

Cauchy-Schwarz Inequality \mathbf{b}, \mathbf{d}
 $p \times 1 \quad p \times 1$

$$(\mathbf{b}'\mathbf{d})^2 \leq (\mathbf{b}'\mathbf{b})(\mathbf{d}'\mathbf{d}) \quad \text{with equality only if } \mathbf{b} = c\mathbf{d} \text{ for some constant } c$$

Extended Cauchy-Schwarz Inequality $\mathbf{b}, \mathbf{d}, \mathbf{B}$ $\mathbf{B} \equiv$ positive definite
 $p \times 1 \quad p \times 1 \quad p \times p$

$$(\mathbf{b}'\mathbf{d})^2 \leq (\mathbf{b}'\mathbf{B}\mathbf{b})(\mathbf{d}'\mathbf{B}^{-1}\mathbf{d}) \quad \text{with equality only if } \mathbf{b} = c\mathbf{B}^{-1}\mathbf{d} \text{ or } \mathbf{d} = c\mathbf{B}\mathbf{b} \text{ for some constant } c$$

Maximization Lemma \mathbf{B} $\mathbf{B} \equiv$ positive definite \mathbf{d} \equiv given vector \mathbf{x} \equiv arbitrary non-zero vector
 $p \times p \quad p \times 1 \quad p \times 1$

$$\max_{\mathbf{x} \neq \mathbf{0}} \frac{(\mathbf{x}'\mathbf{d})^2}{\mathbf{x}'\mathbf{B}\mathbf{x}} = \mathbf{d}'\mathbf{B}^{-1}\mathbf{d} \quad @ \quad \mathbf{x} = c\mathbf{B}^{-1}\mathbf{d} \quad \text{for any constant } c \neq 0$$

Maximization of Quadratic Forms for Points on Unit Sphere

\mathbf{B} $\mathbf{B} \equiv$ positive definite with eigenvalues and eigenvectors: $(\lambda_1, \mathbf{e}_1), \dots, (\lambda_p, \mathbf{e}_p)$ $\lambda_1 \geq \dots \geq \lambda_p$
 $p \times p$

$$\max_{\mathbf{x} \neq \mathbf{0}} \frac{\mathbf{x}'\mathbf{B}\mathbf{x}}{\mathbf{x}'\mathbf{x}} = \lambda_1 \quad @ \quad \mathbf{x} = \mathbf{e}_1 \quad \min_{\mathbf{x} \neq \mathbf{0}} \frac{\mathbf{x}'\mathbf{B}\mathbf{x}}{\mathbf{x}'\mathbf{x}} = \lambda_p \quad @ \quad \mathbf{x} = \mathbf{e}_p$$

$$\max_{\mathbf{x} \neq \mathbf{0} \perp \mathbf{e}_1, \dots, \mathbf{e}_k} \frac{\mathbf{x}'\mathbf{B}\mathbf{x}}{\mathbf{x}'\mathbf{x}} = \lambda_{k+1} \quad @ \quad \mathbf{x} = \mathbf{e}_{k+1} \quad k = 1, \dots, p-1$$

Matrix Inequalities and Maximization

Example:

Let $\mathbf{b}' = [2, -1, 4, 0]$ and $\mathbf{d}' = [-1, 3, -2, 1]$. Verify the Cauchy-Schwartz inequality $(\mathbf{b}' \mathbf{d})^2 \leq (\mathbf{b}' \mathbf{b})(\mathbf{d}' \mathbf{d})$

Multivariate Statistics

Chapter 3

X-Matrix, Mean Vector, Deviation Vectors

Data: p Variables/Characteristics Observed on n Experimental/Sampling Units:

$$\mathbf{X} = \begin{bmatrix} x_{11} & x_{12} & \cdots & x_{1p} \\ x_{21} & x_{22} & \cdots & x_{2p} \\ \vdots & \vdots & \ddots & \vdots \\ x_{n1} & x_{n2} & \cdots & x_{np} \end{bmatrix} = \begin{bmatrix} \mathbf{x}_1' \\ \mathbf{x}_2' \\ \vdots \\ \mathbf{x}_n' \end{bmatrix} = [\mathbf{y}_1 \quad \mathbf{y}_2 \quad \cdots \quad \mathbf{y}_p] \quad \mathbf{x}_j' = [x_{j1} \quad x_{j2} \quad \cdots \quad x_{jp}] \quad j = 1, \dots, n \quad \mathbf{y}_i = \begin{bmatrix} x_{1i} \\ x_{2i} \\ \vdots \\ x_{ni} \end{bmatrix} \quad i = 1, \dots, p$$

Mean Vector and Vector of a Single Mean:

$$\bar{\mathbf{x}} = \begin{bmatrix} \bar{x}_1 \\ \bar{x}_2 \\ \vdots \\ \bar{x}_p \end{bmatrix} \quad \bar{x}_i = \frac{\sum_{j=1}^n x_{ji}}{n} \quad i = 1, \dots, p \quad \bar{\mathbf{x}} = \frac{1}{n} \mathbf{X}' \mathbf{1}_n \quad \bar{x}_i \mathbf{1}_n = \begin{bmatrix} \bar{x}_i \\ \bar{x}_i \\ \vdots \\ \bar{x}_i \end{bmatrix} = \left[\mathbf{y}_i' \left(\frac{1}{\sqrt{n}} \mathbf{1}_n \right) \right] \left(\frac{1}{\sqrt{n}} \mathbf{1}_n \right) \quad i = 1, \dots, p$$

Projection of \mathbf{y}_i on the vector $\mathbf{1}_n$ is $\bar{x}_i \mathbf{1}_n$ which has Length: $\sqrt{\bar{x}_i \mathbf{1}_n' \bar{x}_i \mathbf{1}_n} = |\bar{x}_i| \sqrt{n}$

Vector of Deviations for a Single Variable, Sums of Squares and Cross-Products, and Correlation:

$$\mathbf{d}_i = \begin{bmatrix} x_{1i} - \bar{x}_i \\ x_{2i} - \bar{x}_i \\ \vdots \\ x_{ni} - \bar{x}_i \end{bmatrix} = \mathbf{y}_i - \bar{x}_i \mathbf{1}_n \quad i = 1, \dots, p \quad \mathbf{d}_i' \mathbf{d}_i = \sum_{j=1}^n (x_{ji} - \bar{x}_i)^2 \quad \mathbf{d}_i' \mathbf{d}_k = \sum_{j=1}^n (x_{ji} - \bar{x}_i)(x_{jk} - \bar{x}_k) = L_{\mathbf{d}_i} L_{\mathbf{d}_k} \cos(\theta_{ik}) \Rightarrow r_{ik} = \frac{\sum_{j=1}^n (x_{ji} - \bar{x}_i)(x_{jk} - \bar{x}_k)}{\sqrt{\sum_{j=1}^n (x_{ji} - \bar{x}_i)^2} \sqrt{\sum_{j=1}^n (x_{jk} - \bar{x}_k)^2}} = \cos(\theta_{ik}) \Rightarrow \theta_{ik} = \cos^{-1}(r_{ik})$$

$$L_{\mathbf{d}_i} = \sqrt{\mathbf{d}_i' \mathbf{d}_i} = \sqrt{\sum_{j=1}^n (x_{ji} - \bar{x}_i)^2} = \sqrt{ns_{ii}}$$

X-Matrix, Mean Vector, Deviation Vectors

Example: Given the data matrix

$$\mathbf{X} = \begin{bmatrix} 9 & 1 \\ 5 & 3 \\ 1 & 2 \end{bmatrix}$$

- a) Graph the data, calculate \bar{x} and locate it on the graph.
- b) Calculate the deviation vectors \mathbf{d}_1 and \mathbf{d}_2
- c) Calculate the lengths and angle between the two deviation vectors from part b). Relate to r , s_{11} and s_{22} .

Deviation, SSCP, and Variance-Covariance Matrices

Matrix of Deviations, Sum of Squares Cross-Products (SSCP) Matrix, and Sample Variance-Covariance Matrix (\mathbf{S}_n):

$$\mathbf{E} = \begin{bmatrix} \mathbf{d}_1 & \mathbf{d}_2 & \cdots & \mathbf{d}_p \end{bmatrix} = \begin{bmatrix} x_{11} - \bar{x}_1 & x_{12} - \bar{x}_2 & \cdots & x_{1p} - \bar{x}_p \\ x_{21} - \bar{x}_1 & x_{22} - \bar{x}_2 & \cdots & x_{2p} - \bar{x}_p \\ \vdots & \vdots & \ddots & \vdots \\ x_{n1} - \bar{x}_1 & x_{n2} - \bar{x}_2 & \cdots & x_{np} - \bar{x}_p \end{bmatrix}$$

$$SSCP = \mathbf{E}'\mathbf{E} = \begin{bmatrix} \sum_{j=1}^n (x_{j1} - \bar{x}_1)^2 & \sum_{j=1}^n (x_{j1} - \bar{x}_1)(x_{j2} - \bar{x}_2) & \cdots & \sum_{j=1}^n (x_{j1} - \bar{x}_1)(x_{jp} - \bar{x}_p) \\ \sum_{j=1}^n (x_{j1} - \bar{x}_1)(x_{j2} - \bar{x}_2) & \sum_{j=1}^n (x_{j2} - \bar{x}_2)^2 & \cdots & \sum_{j=1}^n (x_{j2} - \bar{x}_2)(x_{jp} - \bar{x}_p) \\ \vdots & \vdots & \ddots & \vdots \\ \sum_{j=1}^n (x_{j1} - \bar{x}_1)(x_{jp} - \bar{x}_p) & \sum_{j=1}^n (x_{j2} - \bar{x}_2)(x_{jp} - \bar{x}_p) & \cdots & \sum_{j=1}^n (x_{jp} - \bar{x}_p)^2 \end{bmatrix}$$

$$\mathbf{S}_n = \frac{1}{n} \mathbf{E}'\mathbf{E} = \frac{1}{n} \begin{bmatrix} \sum_{j=1}^n (x_{j1} - \bar{x}_1)^2 & \sum_{j=1}^n (x_{j1} - \bar{x}_1)(x_{j2} - \bar{x}_2) & \cdots & \sum_{j=1}^n (x_{j1} - \bar{x}_1)(x_{jp} - \bar{x}_p) \\ \sum_{j=1}^n (x_{j1} - \bar{x}_1)(x_{j2} - \bar{x}_2) & \sum_{j=1}^n (x_{j2} - \bar{x}_2)^2 & \cdots & \sum_{j=1}^n (x_{j2} - \bar{x}_2)(x_{jp} - \bar{x}_p) \\ \vdots & \vdots & \ddots & \vdots \\ \sum_{j=1}^n (x_{j1} - \bar{x}_1)(x_{jp} - \bar{x}_p) & \sum_{j=1}^n (x_{j2} - \bar{x}_2)(x_{jp} - \bar{x}_p) & \cdots & \sum_{j=1}^n (x_{jp} - \bar{x}_p)^2 \end{bmatrix}$$

Random Samples of Vectors of Units

$$\mathbf{X} = \begin{bmatrix} X_{11} & X_{12} & \cdots & X_{1p} \\ X_{21} & X_{22} & \cdots & X_{2p} \\ \vdots & \vdots & \ddots & \vdots \\ X_{n1} & X_{n2} & \cdots & X_{np} \end{bmatrix} = \begin{bmatrix} \mathbf{X}_1' \\ \mathbf{X}_2' \\ \vdots \\ \mathbf{X}_n' \end{bmatrix} \quad \mathbf{X}_j' = [X_{j1} \quad X_{j2} \quad \cdots \quad X_{jp}] \quad j=1, \dots, n$$

Row Vectors \equiv independent observations from common joint probability distribution: $f(\mathbf{x}) = f(x_1, \dots, x_p)$

$\Rightarrow \mathbf{X}_1, \dots, \mathbf{X}_n$ can be treated as a Random Sample from $f(\mathbf{x})$

\Rightarrow joint density is $f(\mathbf{x}_1) \cdots f(\mathbf{x}_n)$ where $f(\mathbf{x}_j) = f(x_{j1}, \dots, x_{jp})$

Treating Units (Rows) as a Random Sample from an Underlying Population of Units on the p Variables:

$$E\{X_{ji}\} = \mu_i \quad V\{X_{ji}\} = E\{(X_{ji} - \mu_i)^2\} = \sigma_{ii} \quad \text{COV}\{X_{ji}, X_{jk}\} = E\{(X_{ji} - \mu_i)(X_{jk} - \mu_k)\} = \sigma_{ik}$$

For Random Samples, Rows are Independent (Note: Not a Time-Series where rows are the same units being observed over Time):

$$\text{COV}\{X_{ji}, X_{lk}\} = 0 \quad j \neq l, \forall i, k$$

$$E\{\mathbf{X}_j\} = \boldsymbol{\mu}_x = \begin{bmatrix} \mu_1 \\ \mu_2 \\ \vdots \\ \mu_p \end{bmatrix} \quad V\{\mathbf{X}_j\} = E\{(\mathbf{X}_j - \boldsymbol{\mu}_x)(\mathbf{X}_j - \boldsymbol{\mu}_x)'\} = \boldsymbol{\Sigma}_x = \begin{bmatrix} \sigma_{11} & \sigma_{12} & \cdots & \sigma_{1p} \\ \sigma_{21} & \sigma_{22} & \cdots & \sigma_{2p} \\ \vdots & \vdots & \ddots & \vdots \\ \sigma_{p1} & \sigma_{p2} & \cdots & \sigma_{pp} \end{bmatrix} \quad \text{COV}\{\mathbf{X}_j, \mathbf{X}_l\} = E\{(\mathbf{X}_j - \boldsymbol{\mu}_x)(\mathbf{X}_l - \boldsymbol{\mu}_x)'\} = \mathbf{0}_{p \times p} \quad j \neq l$$

Expectation and Covariance Matrix of Sample Mean Vector

$$E\{\mathbf{X}_j\} = \boldsymbol{\mu}_x = \begin{bmatrix} \mu_1 \\ \mu_2 \\ \vdots \\ \mu_p \end{bmatrix} \quad V\{\mathbf{X}_j\} = E\{(\mathbf{X}_j - \boldsymbol{\mu}_x)(\mathbf{X}_j - \boldsymbol{\mu}_x)'\} = \boldsymbol{\Sigma}_x = \begin{bmatrix} \sigma_{11} & \sigma_{12} & \cdots & \sigma_{1p} \\ \sigma_{21} & \sigma_{22} & \cdots & \sigma_{2p} \\ \vdots & \vdots & \ddots & \vdots \\ \sigma_{p1} & \sigma_{p2} & \cdots & \sigma_{pp} \end{bmatrix}$$

$$\text{COV}\{\mathbf{X}_j, \mathbf{X}_l\} = E\{(\mathbf{X}_j - \boldsymbol{\mu}_x)(\mathbf{X}_l - \boldsymbol{\mu}_x)'\} = \mathbf{0}_{p \times p}$$

$$\text{Let } \bar{\mathbf{X}} = \frac{1}{n}(\mathbf{X}_1 + \dots + \mathbf{X}_n) = \frac{1}{n} \left(\begin{bmatrix} X_{11} \\ X_{12} \\ \vdots \\ X_{1p} \end{bmatrix} + \dots + \begin{bmatrix} X_{n1} \\ X_{n2} \\ \vdots \\ X_{np} \end{bmatrix} \right) = \begin{bmatrix} \bar{X}_1 \\ \bar{X}_2 \\ \vdots \\ \bar{X}_p \end{bmatrix}$$

$$E\{\bar{\mathbf{X}}\} = E\left\{\frac{1}{n}(\mathbf{X}_1 + \dots + \mathbf{X}_n)\right\} = \frac{1}{n}nE\{\mathbf{X}_j\} = \boldsymbol{\mu}_x$$

$$V\{\bar{\mathbf{X}}\} = E\{(\bar{\mathbf{X}} - \boldsymbol{\mu}_x)(\bar{\mathbf{X}} - \boldsymbol{\mu}_x)'\} = E\left\{\left[\frac{1}{n}\sum_{j=1}^n(\mathbf{X}_j - \boldsymbol{\mu}_x)\right]\left[\frac{1}{n}\sum_{l=1}^n(\mathbf{X}_l - \boldsymbol{\mu}_x)\right]'\right\} = E\left\{\frac{1}{n^2}\sum_{j=1}^n\sum_{l=1}^n(\mathbf{X}_j - \boldsymbol{\mu}_x)(\mathbf{X}_l - \boldsymbol{\mu}_x)'\right\} =$$

$$E\left\{\frac{1}{n^2}\sum_{j=1}^n(\mathbf{X}_j - \boldsymbol{\mu}_x)(\mathbf{X}_j - \boldsymbol{\mu}_x)'\right\} = \frac{1}{n^2}n\boldsymbol{\Sigma}_x = \frac{1}{n}\boldsymbol{\Sigma}_x$$

Expectation of Sample Variance-Covariance Matrix

$$\begin{aligned} \mathbf{S}_n &= \frac{1}{n} \mathbf{E}' \mathbf{E} = \frac{1}{n} \begin{bmatrix} \sum_{j=1}^n (x_{j1} - \bar{x}_1)^2 & \sum_{j=1}^n (x_{j1} - \bar{x}_1)(x_{j2} - \bar{x}_2) & \cdots & \sum_{j=1}^n (x_{j1} - \bar{x}_1)(x_{jp} - \bar{x}_p) \\ \sum_{j=1}^n (x_{j1} - \bar{x}_1)(x_{j2} - \bar{x}_2) & \sum_{j=1}^n (x_{j2} - \bar{x}_2)^2 & \cdots & \sum_{j=1}^n (x_{j2} - \bar{x}_2)(x_{jp} - \bar{x}_p) \\ \vdots & \vdots & \ddots & \vdots \\ \sum_{j=1}^n (x_{j1} - \bar{x}_1)(x_{jp} - \bar{x}_p) & \sum_{j=1}^n (x_{j2} - \bar{x}_2)(x_{jp} - \bar{x}_p) & \cdots & \sum_{j=1}^n (x_{jp} - \bar{x}_p)^2 \end{bmatrix} = \frac{1}{n} \sum_{j=1}^n [(\mathbf{X}_j - \bar{\mathbf{X}})(\mathbf{X}_j - \bar{\mathbf{X}})'] \\ &= \frac{1}{n} \sum_{j=1}^n [\mathbf{X}_j \mathbf{X}_j' - \mathbf{X}_j \bar{\mathbf{X}}' - \bar{\mathbf{X}} \mathbf{X}_j' + \bar{\mathbf{X}} \bar{\mathbf{X}}'] = \frac{1}{n} \sum_{j=1}^n \mathbf{X}_j \mathbf{X}_j' - \bar{\mathbf{X}} \bar{\mathbf{X}}' - \bar{\mathbf{X}} \bar{\mathbf{X}}' + \bar{\mathbf{X}} \bar{\mathbf{X}}' = \frac{1}{n} \sum_{j=1}^n \mathbf{X}_j \mathbf{X}_j' - \bar{\mathbf{X}} \bar{\mathbf{X}}' \end{aligned}$$

Aside: Let \mathbf{V} be a random vector with mean vector and Variance-Covariance matrix: $\boldsymbol{\mu}_V$ and $\boldsymbol{\Sigma}_V$:

$$\begin{aligned} \boldsymbol{\Sigma}_V &= E\{(\mathbf{V} - \boldsymbol{\mu}_V)(\mathbf{V} - \boldsymbol{\mu}_V)'\} = E\{\mathbf{V}\mathbf{V}' - \mathbf{V}\boldsymbol{\mu}_V' - \boldsymbol{\mu}_V\mathbf{V}' + \boldsymbol{\mu}_V\boldsymbol{\mu}_V'\} = E\{\mathbf{V}\mathbf{V}'\} - \boldsymbol{\mu}_V\boldsymbol{\mu}_V' - \boldsymbol{\mu}_V\boldsymbol{\mu}_V' + \boldsymbol{\mu}_V\boldsymbol{\mu}_V' = E\{\mathbf{V}\mathbf{V}'\} - \boldsymbol{\mu}_V\boldsymbol{\mu}_V' \\ \Rightarrow E\{\mathbf{V}\mathbf{V}'\} &= \boldsymbol{\Sigma}_V + \boldsymbol{\mu}_V\boldsymbol{\mu}_V' \end{aligned}$$

$$\Rightarrow E\{\mathbf{S}_n\} = E\left\{\frac{1}{n} \sum_{j=1}^n \mathbf{X}_j \mathbf{X}_j' - \bar{\mathbf{X}} \bar{\mathbf{X}}'\right\} = \left[\frac{1}{n} n(\boldsymbol{\Sigma}_X + \boldsymbol{\mu}_X \boldsymbol{\mu}_X')\right] - \left[\frac{1}{n} \boldsymbol{\Sigma}_X + \boldsymbol{\mu}_X \boldsymbol{\mu}_X'\right] = \frac{n-1}{n} \boldsymbol{\Sigma}_X$$

$$\text{Defining: } \mathbf{S} = \frac{n}{n-1} \mathbf{S}_n = \frac{1}{n-1} \sum_{j=1}^n [(\mathbf{X}_j - \bar{\mathbf{X}})(\mathbf{X}_j - \bar{\mathbf{X}})'] \Rightarrow E\{\mathbf{S}\} = \boldsymbol{\Sigma}_X$$

Generalized Variance

$$\mathbf{S} = \begin{bmatrix} s_{11} & s_{12} & \cdots & s_{1p} \\ s_{12} & s_{22} & \cdots & s_{2p} \\ \vdots & \vdots & \ddots & \vdots \\ s_{1p} & s_{2p} & \cdots & s_{pp} \end{bmatrix} \quad s_{ik} = \frac{1}{n-1} \sum_{j=1}^n (x_{ji} - \bar{x}_i)(x_{jk} - \bar{x}_k) \quad i, k = 1, \dots, p \quad \text{Generalized Sample Variance} = \det(\mathbf{S}) = |\mathbf{S}|$$

$$p = 2: \quad \mathbf{S} = \begin{bmatrix} s_{11} & s_{12} \\ s_{12} & s_{22} \end{bmatrix} = \begin{bmatrix} s_{11} & r_{12} \sqrt{s_{11}} \sqrt{s_{22}} \\ r_{12} \sqrt{s_{11}} \sqrt{s_{22}} & s_{22} \end{bmatrix} \Rightarrow |\mathbf{S}| = s_{11}s_{22} - r_{12}^2 s_{11}s_{22} = s_{11}s_{22} (1 - r_{12}^2) = s_{11}s_{22} (1 - \cos^2 \theta_{12}) = s_{11}s_{22} \sin^2 \theta_{12}$$

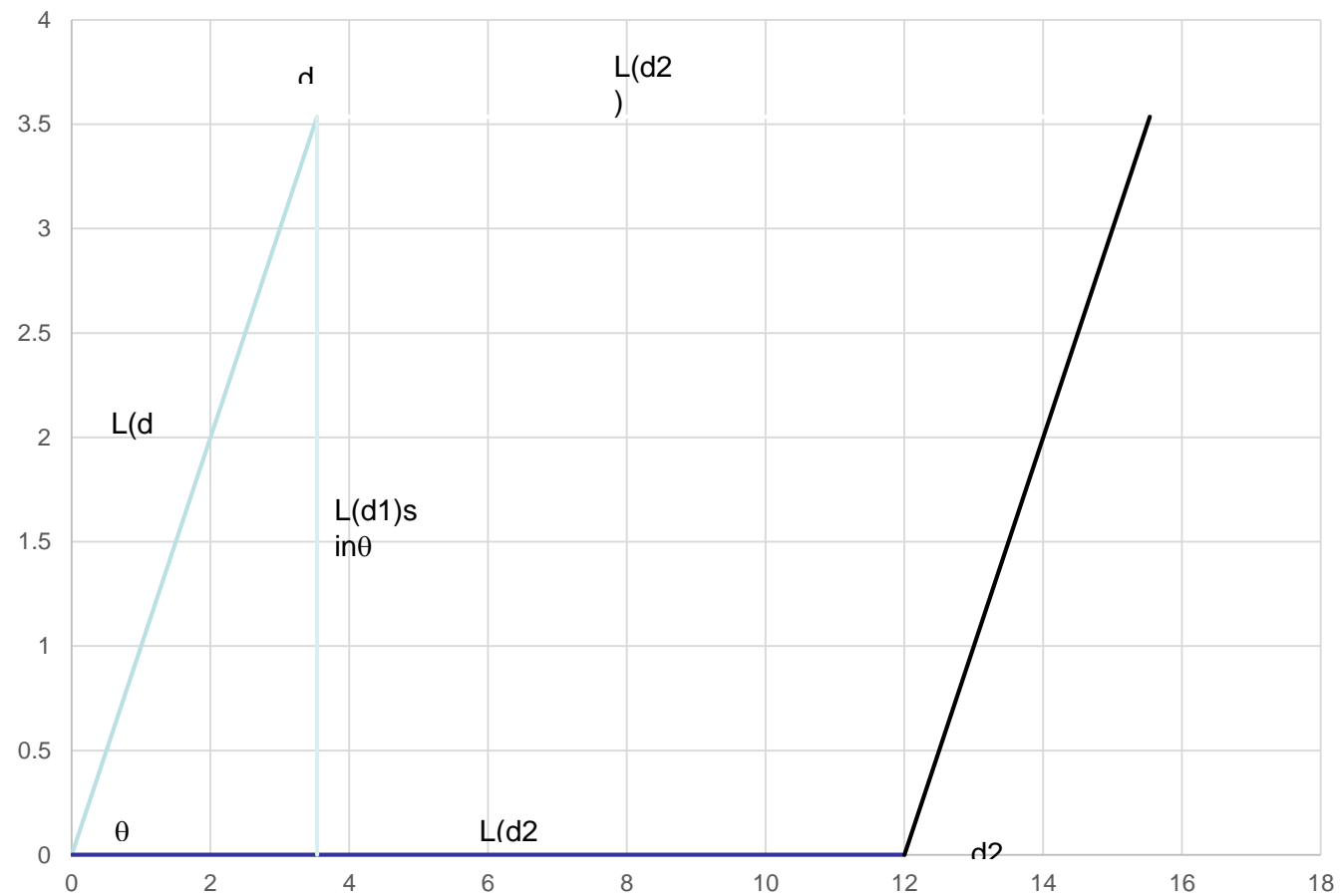
Area of Trapezoid created from vectors $\mathbf{d}_1, \mathbf{d}_2$ (plot on next slide): Area = $L_{\mathbf{d}_1} L_{\mathbf{d}_2} |\sin \theta_{12}|$ $\theta_{12} = \cos^{-1} r_{12}$

$$L_{\mathbf{d}_1} = \sqrt{\sum_{j=1}^n (x_{j1} - \bar{x}_1)^2} = \sqrt{(n-1)s_{11}} \quad L_{\mathbf{d}_2} = \sqrt{\sum_{j=1}^n (x_{j2} - \bar{x}_2)^2} = \sqrt{(n-1)s_{22}}$$

$$\Rightarrow \text{Area} = (n-1) \sqrt{s_{11}s_{22}} |\sin \theta_{12}| \Rightarrow |\mathbf{S}| = \frac{\text{Area}^2}{(n-1)^2} \quad \text{For general } p: |\mathbf{S}| = \frac{\text{Volume}^2}{(n-1)^p}$$

$$\text{Volume of a hyperellipsoid centered at } \bar{\mathbf{x}}: \left\{ \mathbf{x} : (\mathbf{x} - \bar{\mathbf{x}})' \mathbf{S}^{-1} (\mathbf{x} - \bar{\mathbf{x}}) \leq c^2 \right\} = k_p |\mathbf{S}|^{1/2} c^p \quad k_p = \frac{2\pi^{p/2}}{p\Gamma(p/2)}$$

Generalized Sample Variance as an Area



Matrix Form of Sample Mean, Covariance, Correlation

$$\bar{\mathbf{x}} = \begin{bmatrix} \bar{x}_1 \\ \bar{x}_2 \\ \vdots \\ \bar{x}_p \end{bmatrix} = \frac{1}{n} \mathbf{X}' \mathbf{1}_n \quad n \times p \text{ matrix of Column means: } \begin{bmatrix} \bar{x}_1 \mathbf{1}_n & \bar{x}_2 \mathbf{1}_n & \cdots & \bar{x}_p \mathbf{1}_n \end{bmatrix} = \mathbf{1}_n \bar{\mathbf{x}}' = \frac{1}{n} \mathbf{1}_n \mathbf{1}_n' \mathbf{X} = \frac{1}{n} \mathbf{J}_n \mathbf{X}$$

$$\text{Note: } \frac{1}{n} \mathbf{J}_n \frac{1}{n} \mathbf{J}_n = \frac{1}{n^2} \mathbf{J}_n \mathbf{J}_n = \frac{1}{n^2} n \mathbf{J}_n = \frac{1}{n} \mathbf{J}_n$$

$$n \times p \text{ matrix of Deviations from Column means: } \mathbf{E} = \mathbf{X} - \frac{1}{n} \mathbf{J}_n \mathbf{X} = \left(\mathbf{I}_n - \frac{1}{n} \mathbf{J}_n \right) \mathbf{X}$$

$$\text{Note: } \left(\mathbf{I}_n - \frac{1}{n} \mathbf{J}_n \right) \left(\mathbf{I}_n - \frac{1}{n} \mathbf{J}_n \right) = \mathbf{I}_n - \frac{1}{n} \mathbf{J}_n - \frac{1}{n} \mathbf{J}_n + \frac{1}{n^2} \mathbf{J}_n \mathbf{J}_n = \mathbf{I}_n - \frac{1}{n} \mathbf{J}_n$$

$p \times p$ matrix of Sums of Squares and Cross-Products and Estimated Variance-Covariance Matrix:

$$\mathbf{E}'\mathbf{E} = \mathbf{X}' \left(\mathbf{I}_n - \frac{1}{n} \mathbf{J}_n \right)' \left(\mathbf{I}_n - \frac{1}{n} \mathbf{J}_n \right) \mathbf{X} = \mathbf{X}' \left(\mathbf{I}_n - \frac{1}{n} \mathbf{J}_n \right) \mathbf{X} \Rightarrow \mathbf{S} = \frac{1}{n-1} \mathbf{E}'\mathbf{E} = \frac{1}{n-1} \mathbf{X}' \left(\mathbf{I}_n - \frac{1}{n} \mathbf{J}_n \right) \mathbf{X}$$

$p \times p$ Standard Deviation Matrix and its Inverse:

$$\mathbf{D}^{1/2} = \begin{bmatrix} \sqrt{s_{11}} & 0 & \cdots & 0 \\ 0 & \sqrt{s_{22}} & \cdots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \cdots & \sqrt{s_{pp}} \end{bmatrix} \quad \mathbf{D}^{-1/2} = \begin{bmatrix} 1/\sqrt{s_{11}} & 0 & \cdots & 0 \\ 0 & 1/\sqrt{s_{22}} & \cdots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \cdots & 1/\sqrt{s_{pp}} \end{bmatrix} \quad \mathbf{R} = \mathbf{D}^{-1/2} \mathbf{S} \mathbf{D}^{-1/2} \quad \mathbf{S} = \mathbf{D}^{1/2} \mathbf{R} \mathbf{D}^{1/2}$$

Linear Combinations of Sample Values

$$\mathbf{X} = \begin{bmatrix} X_1 \\ X_2 \\ \vdots \\ X_p \end{bmatrix} \quad \mathbf{c}'\mathbf{X} = c_1X_1 + \dots + c_pX_p \quad \text{observed value for Unit } j: \quad \mathbf{c}'\mathbf{x}_j = c_1x_{j1} + \dots + c_px_{jp} \quad j = 1, \dots, n$$

$$\text{Sample Mean: } \frac{\mathbf{c}'\mathbf{x}_1 + \dots + \mathbf{c}'\mathbf{x}_n}{n} = \frac{1}{n}\mathbf{c}'(\mathbf{x}_1 + \dots + \mathbf{x}_n) = \mathbf{c}'\bar{\mathbf{x}}$$

$$\text{Sample Variance: } \frac{1}{n-1} \sum_{j=1}^n (\mathbf{c}'\mathbf{x}_j - \mathbf{c}'\bar{\mathbf{x}})^2 = \frac{1}{n-1} \sum_{j=1}^n [\mathbf{c}'(\mathbf{x}_j - \bar{\mathbf{x}})]^2 = \frac{1}{n-1} \sum_{j=1}^n [\mathbf{c}'(\mathbf{x}_j - \bar{\mathbf{x}})(\mathbf{x}_j - \bar{\mathbf{x}})'\mathbf{c}] = \mathbf{c}' \frac{\sum_{j=1}^n (\mathbf{x}_j - \bar{\mathbf{x}})(\mathbf{x}_j - \bar{\mathbf{x}})'}{n-1} \mathbf{c} = \mathbf{c}'\mathbf{S}\mathbf{c}$$

Second Linear Combination: $\mathbf{b}'\mathbf{X}$:

Sample Mean: $\mathbf{b}'\bar{\mathbf{x}}$ Sample Variance: $\mathbf{b}'\mathbf{S}\mathbf{b}$

$$\text{Sample Variance of } \mathbf{b}'\mathbf{X} \text{ and } \mathbf{c}'\mathbf{X}: \quad \frac{1}{n-1} \sum_{j=1}^n [\mathbf{b}'(\mathbf{x}_j - \bar{\mathbf{x}})\mathbf{c}'(\mathbf{x}_j - \bar{\mathbf{x}})] = \frac{1}{n-1} \sum_{j=1}^n [\mathbf{b}'(\mathbf{x}_j - \bar{\mathbf{x}})(\mathbf{x}_j - \bar{\mathbf{x}})'\mathbf{c}] = \mathbf{b}' \frac{\sum_{j=1}^n (\mathbf{x}_j - \bar{\mathbf{x}})(\mathbf{x}_j - \bar{\mathbf{x}})'}{n-1} \mathbf{c} = \mathbf{b}'\mathbf{S}\mathbf{c}$$

$$\text{Generalized to } q \text{ Linear Combinations: } a_{i1}X_1 + \dots + a_{ip}X_p \quad i = 1, \dots, q \quad \mathbf{A}\mathbf{X} \text{ where } \mathbf{A} = \begin{bmatrix} a_{11} & a_{12} & \cdots & a_{1p} \\ a_{21} & a_{22} & \cdots & a_{2p} \\ \vdots & \vdots & \ddots & \vdots \\ a_{q1} & a_{q2} & \cdots & a_{qp} \end{bmatrix}$$

Sample Mean: $\mathbf{A}\bar{\mathbf{x}}$ Sample Variance: $\mathbf{A}\mathbf{S}\mathbf{A}'$