

# Stochastic Processes: Homework 9

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**Problem 1.** Durrett, Exercise 2.30

- (a) The total number of calls in an hour is Poisson with mean 4. Hence, by Theorem 2.11, the number of men calling in an hour is Poisson with mean  $4 \cdot (3/4) = 3$  and the number of women calling in an hour is Poisson with mean  $4 \cdot (1/4) = 1$ . These two Poisson processes are independent as well. Thus, the probability of seeing exactly two men and three women is given by,

$$e^{-3} \cdot \frac{3^2}{2!} \cdot e^{-1} \cdot \frac{1^3}{3!} \approx 0.014$$

- (b) The sex of the caller is independent of the time of the call, so we can consider this binomial distribution with probability  $3/4$  for male and  $1/4$  for female. The probability for 3 males in the first 3 calls is thus given by,

$$\begin{aligned} (3/4)^3 &= 27/64 \\ &\approx 0.422 \end{aligned}$$

**Problem 2.** Durrett, Exercise 2.33

- a By Theorem 2.15, the set of arrival times  $\{T_1, T_2\}$  has the same distribution as  $\{U_1, U_2\}$ , which are independent and uniformly distributed on  $[0, t]$ . Given that the customers arrived in the first 5 minutes, the arrival times are uniformly distributed on  $[0, 5]$ , and so the probability that each of them arrived in the first 2 minutes is  $2/5$ . The arrival times for customers are independent, hence the probability that both customers arrived in the first 2 minutes is given by,

$$\begin{aligned} (2/5) \cdot (2/5) &= 4/25 \\ &= 0.16 \end{aligned}$$

- b The probability that at least 1 customer arrived in the first 2 minutes is given by 1 minus the probability that both arrived in the last 3 minutes. Each customer has probability  $3/5$  of arriving in the last 3 minutes and, as before, their arrivals are independent. Hence, the probability that at least 1 arrived in the first 2 minutes is given by,

$$\begin{aligned} 1 - (3/5)^2 &= 16/25 \\ &= 0.64 \end{aligned}$$

**Problem 3.** Durrett, Exercise 2.54

We have,

$$\begin{aligned} ES(t) &= E \left[ S_0 \prod_{i=1}^{N(t)} X_i \right] \\ &= S_0 E \left[ \prod_{i=1}^{N(t)} X_i \right] \end{aligned}$$

By the independence of the  $X_i$  variables and law of total expectation, we get,

$$\begin{aligned} ES(t) &= S_0 \prod_{i=1}^{\lambda t} EX_i \\ &= S_0 \prod_{i=1}^{\lambda t} \mu \\ &= S_0 \cdot \mu^{\lambda t} \end{aligned}$$

Now note that since  $\text{Var}(X_i) = \sigma^2$  for each  $i$ , we have,

$$\begin{aligned} \text{Var}(X_i) &= E[X_i^2] - E[X_i]^2 \\ \implies \sigma^2 &= E[X_i^2] - \mu^2 \\ \implies E[X_i^2] &= \sigma^2 + \mu^2 \end{aligned}$$

Then the variance is given by,

$$\begin{aligned} \text{Var}S(t) &= ES(t)^2 - (ES(t))^2 \\ &= E \left[ \left( S_0 \prod_{i=1}^{N(t)} X_i \right)^2 \right] - (S_0 \cdot \mu^{\lambda t})^2 \\ &= E \left[ S_0^2 \prod_{i=1}^{N(t)} X_i^2 \right] - S_0^2 \cdot \mu^{2\lambda t} \\ &= S_0^2 \prod_{i=1}^{\lambda t} E[X_i^2] - S_0^2 \cdot \mu^{2\lambda t} \\ &= S_0^2 \prod_{i=1}^{\lambda t} [\sigma^2 + \mu^2] - S_0^2 \cdot \mu^{2\lambda t} \\ &= S_0^2 [(\sigma^2 + \mu^2)^{\lambda t} - \mu^{2\lambda t}] \end{aligned}$$

**Problem 4.** Durrett, Exercise 2.58

We have that the sum of independent exponential random variables is a gamma distribution. Hence,

$$T = t_1 + \cdots + t_N \sim \text{gamma}(N, \lambda)$$

Notice that  $N$  is a random variable. We can circumvent this by conditioning on  $N = n$ , which yields,

$$T = t_1 + \cdots + t_n \sim \text{gamma}(n, \lambda)$$

Now note that if we let  $A = "T = t"$  and  $B = "N = n"$ , we get,

$$P(A \mid B) \cdot P(B) = P(A \cap B)$$

We already know that  $P(A \mid B) = \text{gamma}(n, \lambda)$ , and from the problem we know that  $P(N = n) = (1 - p)^{n-1}$ . Hence, we have that  $P(B) = (1 - p)^{n-1}$  and so,

$$\begin{aligned} P(A \cap B) &= \text{gamma}(n, \lambda) \cdot (1 - p)^{n-1} \\ &= \lambda e^{-\lambda t} \cdot \frac{(\lambda \cdot t)^{n-1}}{(n-1)!} \cdot (1 - p)^{n-1} \end{aligned}$$