Stochastic Processes II: Homework 9

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Problem I. LPW 12.3

Let $P_L = (P+I)/2$ be the transition matrix of the lazy version of the chain with transition matrix P. Suppose P_L has eigenvalue $\lambda < 0$ with corresponding eigenfunction f. Then,

$$P_L f = (P/2 + I/2)f$$
$$= (Pf + f)/2$$
$$= \lambda f$$

From the above equalities, we have,

$$Pf/2 + f/2 = \lambda f$$

$$\iff Pf/2 = \lambda f - f/2$$

$$\iff Pf = 2\lambda f - f$$

$$\iff Pf = (2\lambda - 1)f$$

Hence, $2\lambda - 1$ is an eigenvalue of P and so, by Lemma 12.1(i), we have that,

$$2\lambda - 1 \ge -1$$
$$\iff \lambda \ge 0$$

We assumed that $\lambda < 0$, and hence we have a contradiction. Thus, there are no eigenvalues λ of P_L such that $\lambda < 0$, as required.

Problem II. LPW 12.4

We have that,

$$E_{\pi}(P^{t}f) = \pi P^{t}f$$
$$= \pi f$$
$$= E_{\pi}(f)$$

We have that the first eigenfunction $f_1 \equiv 1$, and so we have,

$$P^{t}f - E_{\pi}(P^{t}f) = \sum_{j=1}^{|\mathcal{X}|} \langle f, f_{j} \rangle_{\pi} f_{j} \lambda_{j}^{t}$$
$$= \sum_{j=2}^{|\mathcal{X}|} \langle f, f_{j} \rangle_{\pi} f_{j} \lambda_{j}^{t}$$

Because the f_j 's are an orthonormal basis, this gives us,

$$\operatorname{Var}_{\pi}(f) = ||P^{t}f - E_{\pi}(P^{t}f)||_{\ell^{2}(\pi)}^{2}$$
$$= \sum_{j=2}^{|\mathcal{X}|} \langle f, f_{j} \rangle_{\pi}^{2} f_{j} \lambda_{j}^{2t}$$

Now note that for all λ_j , we have that $1 - \gamma_* \ge \lambda_j$. Thus, we have,

$$\operatorname{Var}_{\pi}(f) = \sum_{j=2}^{|\mathcal{X}|} \langle f, f_j \rangle_{\pi}^2 f_j \lambda_j^{2t}$$

$$\leq \sum_{j=2}^{|\mathcal{X}|} \langle f, f_j \rangle_{\pi}^2 f_j (1 - \gamma_*)^{2t}$$

$$= (1 - \gamma_*)^{2t} \sum_{j=2}^{|\mathcal{X}|} \langle f, f_j \rangle_{\pi}^2 f_j$$

Moreover, we have that,

$$\sum_{j=2}^{|\mathcal{X}|} \langle f, f_j \rangle_{\pi}^2 = E_{\pi}(f^2) - E_{\pi}^2(f)$$
$$= \operatorname{Var}_{\pi}(f)$$

And so,

$$\operatorname{Var}_{\pi}(f) \le (1 - \gamma_*)^{2t} \operatorname{Var}_{\pi}(f)$$

as required.

Problem III. LPW 18.1

Suppose that the n-th chain in a sequence of Markov chains satisfies,

$$\lim_{n \to \infty} d_n(ct_{\text{mix}}^n) - \begin{cases} 1 & \text{if } c < 1, \\ 0 & \text{if } c > 1 \end{cases}$$

Then for any $\gamma > 0$ and for n large enough, we have,

$$t_{\text{mix}}(\epsilon) \le (1+\gamma)t_{\text{mix}}^n$$
$$t_{\text{mix}}(1-\epsilon) \ge (1-\gamma)t_{\text{mix}}^n$$

These equations together yield,

$$\frac{t_{\min}(\epsilon)}{t_{\min}(1-\epsilon)} \le \frac{1+\gamma}{1-\gamma}$$

If we let γ approach 0, then n must approach ∞ and we get,

$$\lim_{n \to \infty} \frac{t_{\text{mix}}(\epsilon)}{t_{\text{mix}}(1 - \epsilon)} = 1$$

as required.

Now suppose that we have,

$$\lim_{n \to \infty} \frac{t_{\text{mix}}(\epsilon)}{t_{\text{mix}}(1 - \epsilon)} = 1$$

Fix $\gamma > 0$. Then for any $\epsilon > 0$ and for n large enough, we have $t_{\text{mix}}(\epsilon) \leq (1 + \gamma)t_{\text{mix}}^n$. That is, $\lim_{n\to\infty} d_n((1+\gamma)t_{\text{mix}}^n) \leq \epsilon$. Since this holds for all ϵ ,

$$\lim_{n \to \infty} d_n((1+\gamma)t_{\text{mix}}^n) = 0$$

Moreover, $\lim_{n\to\infty} d_n((1-\gamma)t_{\text{mix}}^n) \geq 1-\epsilon$ since $t_{\text{mix}}(1-\epsilon) \geq (1-\gamma)t_{\text{mix}}^n$ for n sufficiently large. Thus, we have that,

$$\lim_{n \to \infty} d_n((1 - \gamma)t_{\text{mix}}^n) = 1$$