Zero adjusted distributions on the positive real line

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1 Introduction

The new R package gamlss.inf is designed to fit both inflated distributions on the interval [0,1], including 0 and 1, and also zero adjusted distributions defined on $[0,\infty)$ including 0. This vignette focuses on the zero adjusted distributions defined on $[0,\infty)$ including 0 (i.e. the non-negative real line). This is a mixed discrete-continuous distribution, comprising a value Y=0 with probability p_0 and $Y=Y_1$ with probability $(1-p_0)$ where Y_1 has a continuous distribution on the interval $(0,\infty)$, i.e. the positive real line.

The gamlss package, Rigby and Stasinopoulos [2005], Stasinopoulos and Rigby [2007], Stasinopoulos et al. [2017], already provides two zero adjusted distributions, the zero adjusted gamma, ZAGA and the zero adjusted inverse Gaussian, ZAIG. The probability at the value Y=0 may depend on explanatory variables. Note the gamma distribution has 2 parameters, so the zero adjusted gamma with probability at 0 has a total of 3 parameters. In practice, and for complicated data sets, the part of the response which lies on positive real line may need more than 2 distribution parameters to be captured correctly.

There are three methods within gamlss packages to obtain a more flexible continuous distribution with up to 4 parameters with range $(0, \infty)$, i.e. the positive real line, not including zero:

- 1. use a flexible continuous distribution with range $(0, \infty)$ within the **gamlss.dist** package, e.g. $BCCG(\mu, \sigma, \nu)$, $BCT(\mu, \sigma, \nu, \tau)$ or $BCPE(\mu, \sigma, \nu, \tau)$, see Section 2.1
- 2. Generate (i.e. create) a flexible continuous distribution on $(0, \infty)$ by transforming any continuous distribution with range $(-\infty, \infty)$ available in the **gamlss.dist** package to range $(0, \infty)$ using the inverse log (i.e. exponential) transformation through the function gen.Family() with argument type="log", see Section 2.2
- 3. Generate (i.e. create) a flexible continuous distribution with range $(0, \infty)$ by truncating any continuous distribution on $(-\infty, \infty)$ available in **gamlss.dist** package to range $(0, \infty)$ using the function **gen.trun()** of the package **gamlss.tr**, see Section 2.3

The new **R** package **gamlss.inf** allows any continuous distribution on $(0, \infty)$, obtained by any of the above three methods, to be zero adjusted to range $[0, \infty)$, by including a point probability at zero.

So the package **gamlss.inf** enhances the capability of the standard **gamlss()** function by allowing an extra parameter for modelling the probability at zero. The overall distribution can then have up to five parameters. Let μ, σ, ν, τ represent the four parameters of the distribution defined on $(0, \infty)$ and parameter $\xi_0 = p_0$, the probability at 0. Then the general zero adjusted model that the new package gamlss.inf can fit can be written as:

$$Y \stackrel{\text{ind}}{\sim} \mathcal{D}(\boldsymbol{\mu}, \boldsymbol{\sigma}, \boldsymbol{\nu}, \boldsymbol{\tau}, \boldsymbol{\xi}_{0})$$

$$\boldsymbol{\eta}_{1} = g_{1}(\boldsymbol{\mu}) = \mathbf{X}_{1}\boldsymbol{\beta}_{1} + s_{11}(\mathbf{x}_{11}) + \dots + s_{1J_{1}}(\mathbf{x}_{1J_{1}})$$

$$\boldsymbol{\eta}_{2} = g_{2}(\boldsymbol{\sigma}) = \mathbf{X}_{2}\boldsymbol{\beta}_{2} + s_{21}(\mathbf{x}_{21}) + \dots + s_{2J_{2}}(\mathbf{x}_{2J_{2}})$$

$$\boldsymbol{\eta}_{3} = g_{3}(\boldsymbol{\nu}) = \mathbf{X}_{3}\boldsymbol{\beta}_{3} + s_{31}(\mathbf{x}_{31}) + \dots + s_{3J_{3}}(\mathbf{x}_{3J_{3}})$$

$$\boldsymbol{\eta}_{4} = g_{4}(\boldsymbol{\tau}) = \mathbf{X}_{4}\boldsymbol{\beta}_{4} + s_{41}(\mathbf{x}_{41}) + \dots + s_{4J_{4}}(\mathbf{x}_{4J_{4}})$$

$$\boldsymbol{\eta}_{5} = g_{5}(\boldsymbol{\xi}_{0}) = \mathbf{X}_{5}\boldsymbol{\beta}_{5} + s_{51}(\mathbf{x}_{51}) + \dots + s_{5J_{5}}(\mathbf{x}_{5J_{5}})$$

$$(1)$$

where $\mathcal{D}(\boldsymbol{\mu}, \boldsymbol{\sigma}, \boldsymbol{\nu}, \boldsymbol{\tau}, \boldsymbol{\xi}_0)$ is a zero adjusted distribution for the response variable Y defined on $[0, \infty)$ including 0, where \mathbf{X}_k are the design matrices incorporating the linear additive terms in the model, $\boldsymbol{\beta}_k$ are the linear coefficient parameters and $s_{kj}(\mathbf{x}_{kj})$ represent smoothing functions for explanatory variables \mathbf{x}_{kj} , for k=1,2,3,4,5 and $j=1,\ldots,J_k$. Note that the quantitative explanatory variables in the \mathbf{X} 's can be the same or different for the ones defined in the smoothers. The vectors $\boldsymbol{\eta}_1$, $\boldsymbol{\eta}_2$, $\boldsymbol{\eta}_3$, $\boldsymbol{\eta}_4$ and $\boldsymbol{\eta}_5$ are called the *predictors* of the distribution parameters $\boldsymbol{\mu}$, $\boldsymbol{\sigma}$, $\boldsymbol{\nu}$, $\boldsymbol{\tau}$ and $\boldsymbol{\xi}_0$ respectively. Note that $\boldsymbol{\xi}_0 = p_0 = P(Y=0)$.

2 Distributions on the positive real line

2.1 Explicit continuous distributions on $(0, \infty)$

Within the gamlss.dist package there are currently several continuous distributions defined on $(0, \infty)$, including:

- 1. the exponential distribution, $EXP(\mu)$, with one parameter,
- 2. the gamma distribution, $GA(\mu, \sigma)$, with two parameters,
- 3. the inverse gamma distribution, IGAMMA(μ, σ), with two parameters,
- 4. the inverse Gaussian distribution, $IG(\mu, \sigma)$, with two parameters,
- 5. the log normal distribution, LOGNO(μ, σ), with two parameters
- 6. the Pareto type 2 distribution, PARETO2(μ, σ), with two parameters
- 7. the Weibull distribution, $WEI(\mu, \sigma)$, $WEI2(\mu, \sigma)$, $WEI3(\mu, \sigma)$, with two parameters
- 8. the Box-Cox Cole and Green distribution, $BCCG(\mu, \sigma, \nu)$ or $BCCGo(\mu, \sigma, \nu)$, with three parameters.
- 9. the generalized gamma distribution, $GG(\mu, \sigma, \nu)$, with three parameters.
- 10. the generalized inverse Gaussian distribution, $GIG(\mu, \sigma, \nu)$, with three parameters.
- 11. the Box-Cox t distribution, $BCT(\mu, \sigma, \nu, \tau)$ or $BCTo(\mu, \sigma, \nu, \tau)$, with four parameters.
- 12. the Box-Cox power exponential distribution, $\mathsf{BCPE}(\mu, \sigma, \nu, \tau)$ or $\mathsf{BCPEo}(\mu, \sigma, \nu, \tau)$, with four parameters.
- 13. the generalized beta type 2 distribution, $GB2(\mu, \sigma, \nu, \tau)$, with four parameters

2.2 Log transform distributions on $(0, \infty)$.

Any continuous random variable say Z defined on $(-\infty, \infty)$ can be transformed by the inverse log (i.e. exponential) transformation $Y = \exp(Z)$ to a random variable Y defined on $(0, \infty)$. The resulting distribution is called a log-transform distribution. For example, if Z is a t-family distributed variable i.e. $Z \sim \mathsf{TF}(\mu, \sigma, \nu)$, and the inverse log transformation is applied, then $Y \sim \mathsf{logTF}(\mu, \sigma, \nu)$, i.e. a log-t family distribution on $(0, \infty)$.

The following is an example on how to take a gamlss.family distribution on $(-\infty,\infty)$ and create a corresponding log-transform distribution defined on $(0,\infty)$. The gamlss function gen.Family() of the gamlss.dist package generates the d (pdf), p (cdf), q (inverse cdf) and r (random generation) functions of the log-transform distribution, together with the function which can be used for fitting within gamlss. Here first generate a log-t family distribution, i.e. $\log TF(\mu, \sigma, \nu)$, and plot the distribution for different values of μ , σ and ν . Note that μ , σ and ν are defined on the original t-distribution ranges, i.e. $(-\infty, \infty)$ for μ and $(0, \infty)$ for σ and ν . This implies that $\exp(\mu)$ is not the mean of the log-transform, $\log TF(\mu, \sigma, \nu)$, distribution but its median. Also σ and ν are related to the scale and shape of the distribution. We use gen.Family("TF", type="log") to generate a log-t family distribution and in Figure 1 we plot the distribution for different values of μ , σ and ν using the function curve().

```
# generate the distribution
library(gamlss)
gen.Family("TF", type="log")
## A log family of distributions from TF has been generated
## and saved under the names:
## dlogTF plogTF qlogTF rlogTF logTF
# different mu
curve(dlogTF(x, mu=-5, sigma=1, nu=10), 0,10, n = 1001, ylim=c(0,1),
      lwd=2, lty=2, col=2)
title("(a)")
curve(dlogTF(x, mu=-1, sigma=1, nu=10), 0,10, add=TRUE, lwd=2, lty=3, col=3)
curve(dlogTF(x, mu=0, sigma=1, nu=10), 0,10, add=TRUE, lwd=2, lty=4, col=4)
curve(dlogTF(x, mu=1, sigma=1, nu=10), 0,10, add=TRUE, lwd=2, lty=5, col=5)
curve(dlogTF(x, mu=2, sigma=1, nu=10), 0.10, add=TRUE, lwd=2, lty=6, col=6)
legend("topright",
       legend=\mathbf{c}("mu = -5", "mu = -1", "mu = 0", "mu = 1", "mu = 2"),
       lty=2:6, col=2:6, cex=1)
# different sigma
curve(dlogTF(x, mu=0, sigma=.5, nu=10), 0,2, ylim=c(0,3),
      lwd=2, lty=2, col=2)
title(("(b)"))
curve(dlogTF(x, mu=0, sigma=1, nu=10), 0,2, add=TRUE, lwd=2, lty=3, col=3)
curve(dlogTF(x, mu=0, sigma=2, nu=10), 0,2, add=TRUE, lwd=2, lty=4, col=4)
curve(dlogTF(x, mu=0, sigma=5, nu=10), 0,2, add=TRUE, lwd=2, lty=5, col=5)
legend("topright",
       legend=c("sigma = .5", "sigma = 1", "sigma = 2", "sigma = 5"),
       lty=2:5, col=2:5, cex=1)
# different nu
curve(dlogTF(x, mu=0, sigma=.5, nu=1000), 0,3, ylim=c(0,1.1),
       lwd=2, lty=2, col=2)
title("(c)")
curve(dlogTF(x, mu=0, sigma=.5, nu=10), 0,3, add=TRUE, lwd=2, lty=3, col=3)
curve(dlogTF(x, mu=0, sigma=.5, nu=2), 0,3, add=TRUE, lwd=2, lty=4, col=4)
curve(dlogTF(x, mu=0, sigma=.5, nu=1), 0,3, add=TRUE, lwd=2, lty=5, col=5)
legend("topright",
       legend=\mathbf{c}("nu = 1000", "nu = 10", "nu = 2", "nu = 1"),
       lty=2:5, col=2:5, cex=1)
```

Figure 1

Figure 1 shows the different shapes the $\log TF(\mu, \sigma, \nu)$ distribution can take. Panel (a) shows, for fixed $\sigma = 1$ and $\nu = 10$ how the distribution changes for different values of $\mu = (-5, -1, 0, 1, 2)$. Panel (b) fixes $\mu = 0$ and $\nu = 10$ and varies $\sigma = (0.5, 1, 2, 5)$. Finally panel (c) fixes $\mu = 0$ and $\sigma = 0.5$ and varies $\nu = (1000, 10, 2, 1)$.

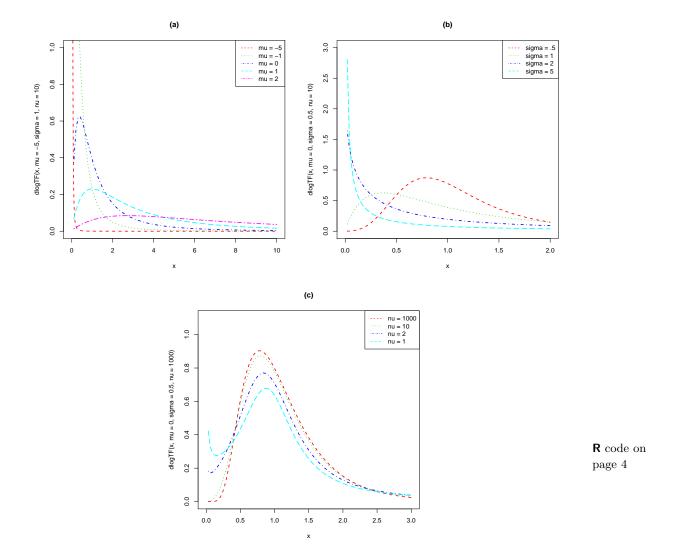


Figure 1: A log-t distribution: (a) with values $\mu = (-5, -1, 0, 1, 2)$, $\sigma = 1$ and $\nu = 10$, (b) with values $\mu = 0$, $\sigma = (0.5, 1, 2, 5)$ and $\nu = 10$ and (c) with values $\mu = 0$, $\sigma = 0.5$ and $\nu = (1000, 10, 2, 1)$.

2.3 Truncated distributions on $(0, \infty)$.

Any distribution defined on the real line $(-\infty, \infty)$ can be left truncated at 0 to give a truncated distribution on $(0, \infty)$ using the function gen.trun() from the R package gamls.tr.

Below we transform a skew student t, $SST(\mu, \sigma, \nu, \tau)$, distribution by left truncating it at zero to give a truncated SST distribution, $SSTtr(\mu, \sigma, \nu, \tau)$ defined on $(0, \infty)$. The range of each parameter of $SSTtr(\mu, \sigma, \nu, \tau)$ is the same as for $SST(\mu, \sigma, \nu, \tau)$, i.e. $-\infty < \mu < \infty, \sigma > 0, \nu > 0$ and $\tau > 2$.

Figure 2

```
# generate the distribution
library(gamlss.tr)
gen.trun(0, "SST", type="left")
## A truncated family of distributions from SST has been generated
## and saved under the names:
## dSSTtr pSSTtr qSSTtr rSSTtr SSTtr
## The type of truncation is left
## and the truncation parameter is 0
# different mu
curve(dSSTtr(x, mu=.1, sigma=.5, nu=1, tau=10), 0,4, lwd=2, lty=2, col=2)
curve(dSSTtr(x, mu= 1, sigma=.5, nu=1, tau=10), 0,4, lwd=2, lty=3, col=3,
      add=TRUE)
curve(dSSTtr(x, mu= 2, sigma=.5, nu=1, tau=10), 0,4, lwd=2, lty=4, col=4,
       add=TRUE)
title("(a)")
legend("topright",
       legend=c("mu = .1","mu = 1","mu = 2"),
       lty=2:4, col=2:4, cex=1)
# different sigma
curve(dSSTtr(x, mu=2, sigma=0.5, nu=1, tau=10), 0,4, lwd=2, lty=2, col=2 )
curve(dSSTtr(x, mu=2, sigma= 1, nu=1, tau=10), 0,4, lwd=2, lty=3, col=3,
      add=TRUE)
curve(dSSTtr(x, mu=2, sigma= 2, nu=1, tau=10), 0,4, lwd=2, lty=4, col=4,
      add=TRUE)
title("(b)")
legend("topright",
       legend=c("sigma = .5", "sigma = 1", "sigma = 2"),
       lty=2:4, col=2:4, cex=1)
# different nu
curve(dSSTtr(x, mu=2, sigma=.5, nu=0.1, tau=10), 0,4, lwd=2, lty=2, col=2)
curve(dSSTtr(x, mu=2, sigma=.5, nu= 1, tau=10), 0,4, lwd=2, lty=3, col=3,
      add=TRUE)
curve(dSSTtr(x, mu=2, sigma=.5, nu= 2, tau=10), 0,4, lwd=2, lty=4, col=4,
      add=TRUE)
title("(c)")
legend("topright",
       legend=c("nu = 0.1", "nu = 1", "nu = 2"),
      lty=2:4, col=2:4, cex=1)
```

Figure 2 shows the different shapes of the $SSTtr(\mu, \sigma, \nu, \tau)$ distribution can take. Panel (a) shows, for fixed $\sigma = 0.5$ $\nu = 1$ and $\tau = 10$, how the distribution changes for different values of $\mu = (0.1, 1, 2)$. Panel (b) fixes $\mu = 2$, $\nu = 1$ and $\tau = 10$ and varies $\sigma = (0.5, 1, 2)$. Panel (c) fixes $\mu = 2$, $\sigma = 0.5$ and $\tau = 10$ and varies $\nu = (0.1, 1, 2)$. Panel (d) fixes $\mu = 2$, $\sigma = 0.5$ and $\nu = 1$ and varies $\tau = (3, 5, 100)$.

3 Generating zero adjusted distributions

Next it is shown how any gamlss.family distribution defined on $(0, \infty)$ can be zero adjusted to $[0, \infty)$, by including a point probability at 0.

The function gen.Zadj() takes as an argument a gamlss.family distribution on $(0, \infty)$, obtained by any of the three methods of Section 2, and generates a zero adjusted version of the distribution having a point probability at 0. The function has only one argument, family, which specifies a distribution family on $(0, \infty)$.

The resulting mixed (continuous-discrete) probability (density) function (pdf) is given by:

$$f_Y(y|\boldsymbol{\theta}, \xi_0) = \begin{cases} \xi_0 & \text{if } y = 0\\ (1 - \xi_0) f_W(y|\boldsymbol{\theta}) & \text{if } 0 < y < \infty \end{cases}$$
 (2)

for $0 \le y \le \infty$, where $f_W(y|\boldsymbol{\theta})$ is any probability density function defined on $(0, \infty)$, i.e. for $0 < y < \infty$, with parameters $\boldsymbol{\theta}^{\top} = (\theta_1, \theta_2, \dots, \theta_p)$ and $0 < \xi_0 < 1$, where $\xi_0 = P(Y = 0)$. Note that, in the **gamlss.inf** implementation, $\boldsymbol{\theta}$ has a maximum of four parameters denoted by $\boldsymbol{\theta}^T = (\mu, \sigma, \nu, \tau)$.

In the example below first take the skew student t-family distribution, SST(μ, σ, ν, τ), defined on $(-\infty, \infty)$, and use the gen.Family() function in the gamlss.dist package to generate the distribution logSST(μ, σ, ν, τ), defined on $(0, \infty)$. By using the function gen.Zadj() on the new generated logSST(μ, σ, ν, τ) distribution, a zero adjusted logSST distribution, logSSTZadj(μ, σ, ν, τ) is created, defined on $[0, \infty)$.

```
library(gamlss.inf)
gen.Family(family="SST", type="log")
## A log family of distributions from SST has been generated
## and saved under the names:
## dlogSST plogSST qlogSST rlogSST
```

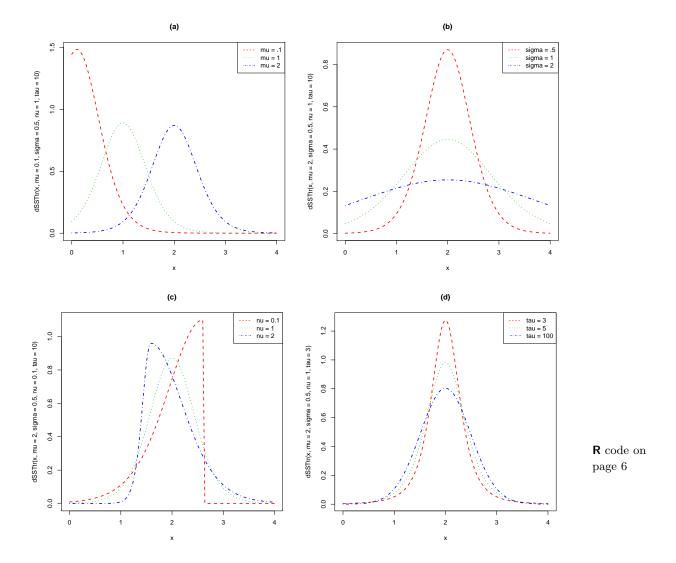


Figure 2: A truncated from zero SST distribution: (a) with values $\mu=(0.1,1,2),\,\sigma=0.5,\,\nu=1$ and $\tau=10$, (b) with values $\mu=2,\,\sigma=(0.5,1,2),\,\nu=1$ and $\tau=10$ (c) with values $\mu=2,\,\sigma=0.5,\,\nu=(0.1,1,2)$ and $\tau=10$. (d) with values $\mu=2,\,\sigma=0.5,\,\nu=1$ and $\tau=(3,5,100)$

```
gen.Zadj(family="logSST")

## A zero adjusted logSST distribution has been generated
## and saved under the names:
## dlogSSTZadj plogSSTZadj qlogSSTZadj rlogSSTZadj
## plotlogSSTZadj
```

There are five function generated here:

```
dlogSSTZadj() The pdf of the distribution, d function.
```

plogSSTZadj() The cdf of the distribution, p function.

qlogSSTZadj() The inverse cdf of the distribution, q function.

rlogSSTZadj() The random generating function of the distribution, r function.

plotlogSSTZadj() The function for plotting the pdf of the distribution.

To fit the distribution the function gamlssZadj() can be used, see Section 5.

4 Plotting zero adjusted distributions

The newly created plotlogSSTZadj() function can be used to plot the pdf of the adjusted distribution (which is a mixed continuous-discrete distribution). Figure 3 shows the use of the plotlogSSTZadj() function. The function plots the zero adjusted probability (density) function including the point probability at zero. Figure 3 shows eight different realisations of the distribution for different values of the parameters.

```
plotlogSSTZadj(mu= 1, sigma=1, nu=1, tau=10, xi0=.1); title("(a)")
plotlogSSTZadj(mu=-1, sigma=1, nu=1, tau=10, xi0=.1); title("(b)")
plotlogSSTZadj(mu=-1, sigma=2, nu=1, tau=10, xi0=.1); title("(c)")
plotlogSSTZadj(mu=0, sigma=2, nu=1, tau=10, xi0=.1); title("(d)")
plotlogSSTZadj(mu=0, sigma=1, nu=10, tau=10, xi0=.1); title("(e)")
plotlogSSTZadj(mu=0, sigma=1, nu=1, tau=3, xi0=.1); title("(f)")
plotlogSSTZadj(mu=0, sigma=1, nu=2, tau=3, xi0=.5); title("(g)")
plotlogSSTZadj(mu=0, sigma=1, nu=.3, tau=100, xi0=.1); title("(h)")
```

The standard plotting functions of R can also be used to plot the created mixed distribution as is shown below. Figure 4 shows how the pdf, cdf, inverse cdf and randomisation functions can be displayed for different values of the distribution parameters.

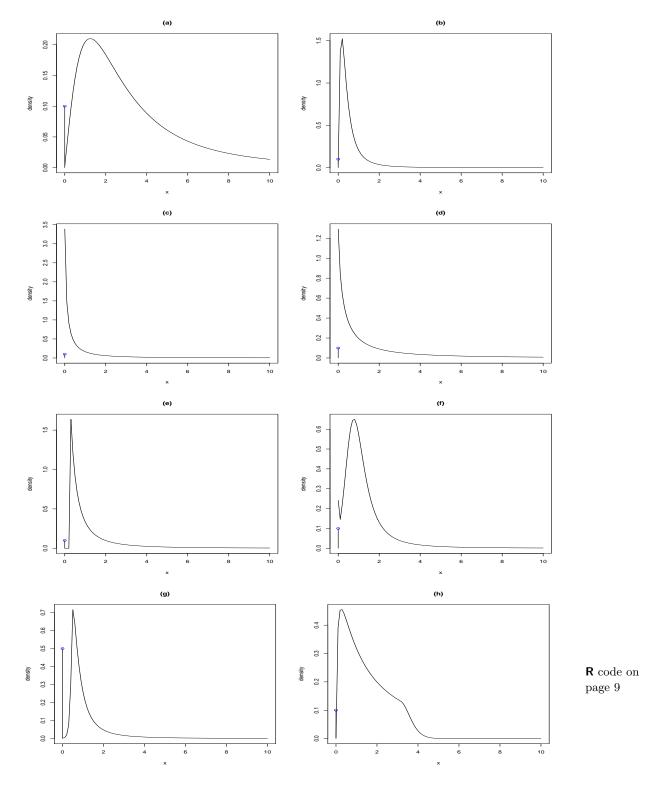


Figure 3: A logSST distribution: (a) with values $\mu=1,\,\sigma=1,\,\nu=1,\,\tau=10$ and $\xi_0=.1$ (b) with values $\mu=-1,\,\sigma=1,\,\nu=1,\,\tau=10$ and $\xi_0=.1$ (c) with values $\mu=-1,\,\sigma=2,\,\nu=1,\,\tau=10$ and $\xi_0=.1$ (d) with values $\mu=0,\,\sigma\stackrel{\underline{10}}{=}2,\,\nu=1,\,\tau=10$ and $\xi_0=.1$ (e) with values $\mu=0,\,\sigma=1,\,\nu=10,\,\tau=10$ and $\xi_0=.1$ (f) with values $\mu=0,\,\sigma=1,\,\nu=1,\,\tau=3$ and $\xi_0=.1$ (g) with values $\mu=0,\,\sigma=1,\,\nu=2,\,\tau=3$ and $\xi_0=.5$ (h) with values $\mu=0,\,\sigma=1,\,\nu=3,\,\tau=100$ and $\xi_0=.1$.

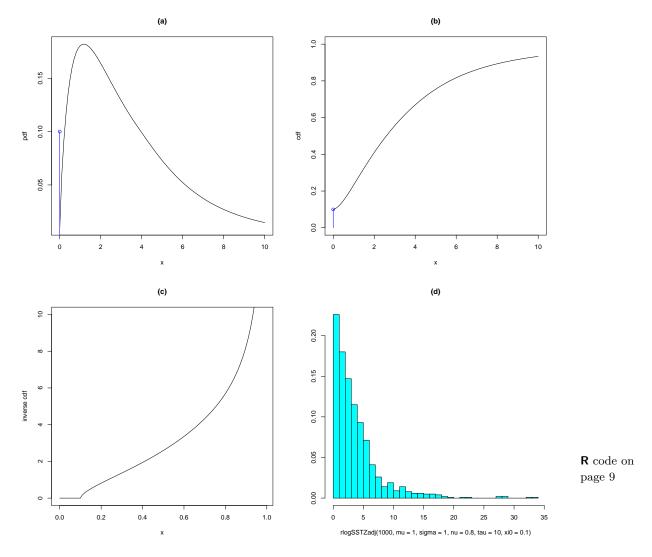


Figure 4: The (a) pdf (b) cdf (c) inverse pdf and (d) simulated data from a zero adjusted log-SST distribution with $\mu=1,\,\sigma=1,\,\nu=.8,\,\tau=10,\,\xi_0=.1.$

The next section demonstrates how to use the function gamlssZadj() to fit a model which has a response variable on the interval $[0, \infty)$, i.e. including 0.

5 Fitting a zero adjusted distribution

5.1 The gamlssZadj() function

The main function for fitting a zero adjusted distribution model to a response variable Y on $[0,\infty)$ i.e. including 0, is gamlssZadj(). In a zero adjusted distribution the parameters μ , σ , ν and τ (of the unadjusted distribution) are orthogonal to the parameter $\xi_0 = P(Y=0)$ in the sense that the log-likelihood function can be factorised in two components, one containing the μ , σ , ν and τ and another containing ξ_0 . This means that the two sets of parameters can be estimated separately. The function gamlssZadj() takes advantage of this separation and works as follows:

- It picks the argument family which defines a gamlss. family distribution defined on $(0,\infty)$
- It creates a binary response variable $Y_1 = 1$ (if Y = 0) + 0(if Y > 0) and it fits a binary logistic model (using the gamlss.family BI)
- Fits a GAMLSS model to the data cases with Y in interval $(0, \infty)$ using the distribution defined by family, by weighting out the observations with Y = 0.
- Creates the (normalised randomized) quantile residuals for the all cases (including cases with Y=0) using the fitted zero adjusted distribution model.
- Saves the output as an gamlssZadj object which is a subclass of a gamlss object.

The idea is that the object gamlssZadj should behave similar to a gamlss object. For this purpose the following S3 methods are created.

```
    fitted.gamlssZadj(),
    coef.gamlssZadj(),
```

- print.gamlssZadj(),
- deviance.gamlssZadj(),
- vcov.gamlssZadj(),
- 6. summary.gamlssZadj(),
- 7. predict.gamlssZadj(),
- 8. formula.gamlssZadj().

The methods are demonstrated in the next sections.

The function gamlssZadj() has the following arguments:

y the response variable on $[0, \infty)$ (including values at zero)

mu.formula a model formula for the μ parameter

sigma.formula a model formula for the σ parameter

nu.formula a model formula for the ν parameter

tau.formula a model formula for the τ parameter

xi0.formula a model formula for the ξ_0 parameter which is equals to the probability at zero **data** a data frame containing the variables occurring in the formula.

family any gamlss() distribution family defined on $(0, \infty)$

weights a vector of prior weights as in gamlss()

trace logical, if TRUE information on model estimation will be printed during the fitting

... for extra arguments which can be passed to gamlss().

Since the individual models fitted within the algorithm used in gamlssZadj() are GAMLSS models, the parameter formulae above can take any linear or additive GAMLSS terms including smoothers and random effects.

To demonstrate the use of the gamlssZadj() function a simulated example is used below.

5.2 Simulating data

To compare the results obtained by the function gamlssZadj() to the ones obtained from standard gamlss(), simulate data from the zero adjusted gamma distribution, $ZAGA(\mu, \sigma, \nu)$.

The mixed continuous-discrete probability (density) function of $Y \sim \mathsf{ZAGA}(\mu, \sigma, \nu)$ is given by

$$f_Y(y) = \begin{cases} p_0 & \text{if } y = 0\\ (1 - p_0) f_W(y) & \text{if } 0 < y < \infty \end{cases}$$
 (3)

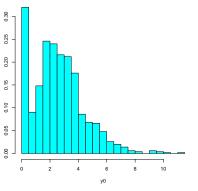
for $0 \le y < \infty$, where $W \sim \mathsf{GA}(\mu, \sigma)$ has a gamma distribution with $0 < \mu < \infty$ and $0 < \sigma < \infty$ and $0 < \nu < 1$, where $\nu = p_0$. For $\mathsf{ZAGA}(\mu, \sigma, \nu)$ the default link functions for μ and σ are the log link function and for ν the logit link function, giving predictors $\eta_1 = \log \mu$, $\eta_2 = \log \sigma$ and $\eta_3 = \log [\nu/(1-\nu)]$.

Here $f_W(y)$ is given by

$$f_W(y|\mu,\sigma) = \frac{1}{(\sigma^2 \mu)^{1/\sigma^2}} \frac{y^{\frac{1}{\sigma^2} - 1} e^{-y/(\sigma^2 \mu)}}{\Gamma(1/\sigma^2)}$$
(4)

for y > 0, where $\mu > 0$ and $\sigma > 0$ and $E(W) = \mu$ and $Var(W) = \sigma^2 \mu^2$.

The simulated example comes from a zero adjusted gamma distribution, $\mathsf{ZAGA}(\mu, \sigma, \nu)$, with $\mu = 3$ and $\sigma = 0.5$ and $\nu = 0.15$, so $p_0 = 0.15$.



R code on page 13

Figure 5: Generated data using the zero adjusted gamma distribution: with values $\mu = 3$, $\sigma = 0.5$, and $\nu = 0.15$.

5.3 Fitting a zero adjusted distribution

The zero adjusted distribution is fitted using both gamlss() and gamlssZadj() functions. Note that the family argument in gamlssZadj() takes a gamlss.family distribution defined on $(0, \infty)$. The trace=TRUE argument is used in gamlssZadj() to check the convergence of the two different models fitted, one using the BI(μ) family and the other using the GA(μ , σ).

```
g0 <- gamlss(y0~1, family=ZAGA)</pre>
## GAMLSS-RS iteration 1: Global Deviance = 3860.799
## GAMLSS-RS iteration 2: Global Deviance = 3860.799
 t0 <- gamlssZadj(y=y0, mu.formula=~1, family=GA, trace=TRUE)
## ****
               The binomial model
                                         ****
## GAMLSS-RS iteration 1: Global Deviance = 865.9541
## GAMLSS-RS iteration 2: Global Deviance = 865.9541
## **** The continuous distribution model ****
## GAMLSS-RS iteration 1: Global Deviance = 2994.845
  GAMLSS-RS iteration 2: Global Deviance = 2994.845
##
               The Final Global Deviance = 3860.799
AIC(g0,t0, k=0)
      df
              AIC
## g0 3 3860.799
## t0 3 3860.799
```

Note that the global deviance of the fitted t0 model, using gamlssZadj(), is obtained by adding the individual deviances from the binomial and the gamma model. The third fitted parameter in both models, is equal to the the probability at zero. The third parameter is called nu (i.e. ν) in gamlss but xi0 (i.e. ξ_0) in gamlssZadj(), where $\nu = p_0$ and $\xi_0 = p_0$. The default link functions are logit(ν) and logit(ξ_0) giving the same predictor $\eta_3 = \log [p_0/(1-p_0)]$. Hence the intercept coefficients (β_{30}) for the predictor η_3 for the third parameter are the same for both

models as shown below.

```
coef(g0, "nu")
## (Intercept)
## -1.688296

coef(t0, "xi0")
## (Intercept)
## -1.688296
```

The fitted values for ν and ξ_0 , which are the estimated probabilities at zero, are also identical.

```
fitted(t0, "xi0")[1]
## [1] 0.156
fitted(g0, "nu")[1]
## [1] 0.156
```

The summary() function makes it clear that the two models are identical.

```
summary(t0)
## *********************
## Family: "ZadjGA"
##
## Call: gamlssZadj(y = y0, mu.formula = ~1, family = GA, trace = TRUE)
##
## Fitting method: RS()
##
## Mu link function: log
## Mu Coefficients:
     Estimate Std. Error t value Pr(>|t|)
## (Intercept) 1.10322 0.01795 61.45 <2e-16 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## -----
## Sigma link function: log
## Sigma Coefficients:
## Estimate Std. Error t value Pr(>|t|)
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## -----
## xi0 link function: logit
## xi0 Coefficients:
          Estimate Std. Error t value Pr(>|t|)
```

```
## (Intercept) -1.68830 0.08715 -19.37 <2e-16 ***
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## -----
## No. of observations in the fit: 1000
## Degrees of Freedom for the fit: 3
## Residual Deg. of Freedom: 997
##
                at cycle:
##
## Global Deviance: 3860.799
## AIC:
               3866.799
         SBC:
                3881.522
## **********************
summary(g0)
## *********************
## Family: c("ZAGA", "Zero adjusted GA")
## Call: gamlss(formula = y0 ~ 1, family = ZAGA)
##
## Fitting method: RS()
##
## -----
## Mu link function: log
## Mu Coefficients:
## Estimate Std. Error t value Pr(>|t|)
## (Intercept) 1.10322 0.01795 61.45 <2e-16 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## -----
## Sigma link function: log
## Sigma Coefficients:
## Estimate Std. Error t value Pr(>|t|)
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
## -----
## Nu link function: logit
## Nu Coefficients:
## Estimate Std. Error t value Pr(>|t|)
## (Intercept) -1.68830 0.08715 -19.37 <2e-16 ***
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
## No. of observations in the fit:
## Degrees of Freedom for the fit:
                               3
##
       Residual Deg. of Freedom:
                               997
##
                     at cycle:
##
## Global Deviance:
                    3860.799
##
                     3866.799
            AIC:
##
             SBC:
                    3881.522
## **********************
```

The estimated variance covariance matrix of the estimates of the intercept coefficients (β_{0k} for k=1,2,3) in the predictors of (μ,σ,ν) and (μ,σ,ξ_0) for the fitted g0 and t0 models, respectively, can be obtained as follows:

```
vcov(t0)
##
                 (Intercept)
                               (Intercept) (Intercept)
## (Intercept) 3.222917e-04 -1.812589e-10
                                           0.00000000
## (Intercept) -1.812589e-10 5.438034e-04
                                           0.00000000
## (Intercept) 0.000000e+00 0.000000e+00
                                           0.00759509
vcov(g0)
##
                 (Intercept)
                               (Intercept) (Intercept)
## (Intercept) 3.222917e-04 -1.812639e-10
                                           0.00000000
                             5.438034e-04
## (Intercept) -1.812639e-10
                                            0.00000000
## (Intercept) 0.000000e+00 0.000000e+00
                                           0.00759509
```

Note that, because of the partition of the likelihood function, parameters μ and σ are orthogonal to ν or ξ_0 .

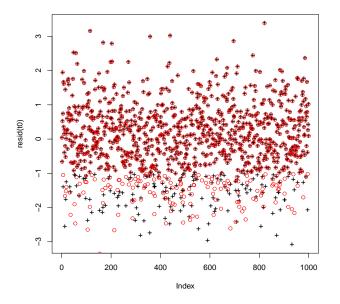
The residuals for the two models should be identical for the not zero response variable values, i.e. Y > 0. Due to the randomization of the residuals at discrete values of Y (i.e. Y = 0 here) we expect differences in the randomized residuals when the response Y is zero. This is demonstrated in the lower part of Figure 6 where the residuals are plotted against the observation index.

```
plot(resid(t0), pch="+")
points(resid(g0), col="red")
Figure 6
```

Next we will plot the fitted distribution in Figure 7. The standard $ZAGA(\mu, \sigma)$ distribution in **gamlss.dist** has its own plotting function called plotZAGA() which can be used here. For the model fitted with gamlssZadj() such a function plotGAZadj() is created using the gen.Zadj() function.

```
# generate the
gen.Zadj("GA")

## A zero adjusted GA distribution has been generated
## and saved under the names:
## dGAZadj pGAZadj rGAZadj
## plotGAZadj
plotZAGA(mu=fitted(g0, "mu")[1], sigma=fitted(g0, "sigma")[1],
```



 ${f R}$ code on page 17

Figure 6: Superimposed residuals from models t0, (+), and g0, (o). Because of the randomization in the residuals for the zero values of the response, the values of the residuals in the lower part of the plot are not identical.

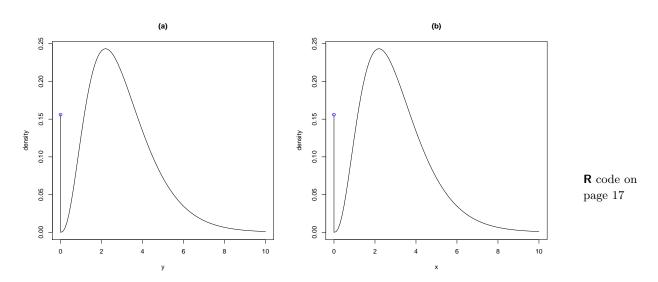


Figure 7: The fitted distribution using (a) plotZAGA() and (b) plotGAZadj().

The fitted distributions are identical.

6 Fitting a regression model

6.1 Simulation example.

We first generate the values for the explanatory variable x and then create three different functions of x, $f_{\mu}(x)$, $f_{\sigma}(x)$ and $f_{\nu}(x)$, for the parameters μ , σ , ν ; respectively. We will use for the simulations the functions fmu(), fsigma() and fnu() produced by the following code and shown in Figure 8. Note that the data frame sda contains values for a previous fitted model for μ , σ and ν given an explanatory variable x.

```
# generating x ------
set.seed(3210)
x <- (runif(1000)*4)-2
range(x)
## [1] -1.995186  1.999197

data(sda)
fmu <- splinefun(sda$x, sda$mu)
curve(fmu, -2,2)
fsigma <- splinefun(sda$x, sda$sigma)</pre>
```

```
curve(fsigma, -2,2)
fnu <- function(x)
  {f <- splinefun(sda$x, sda$nu)
  f(x)/6
  }
curve(fnu, -2,2)</pre>
```

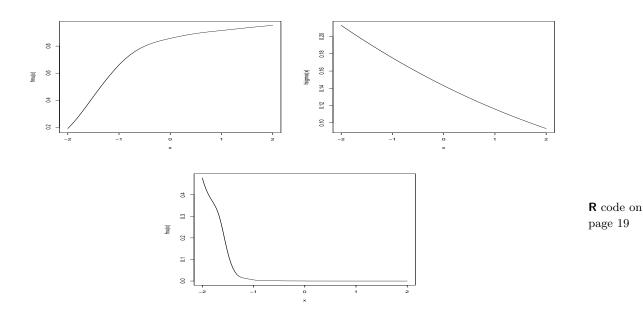


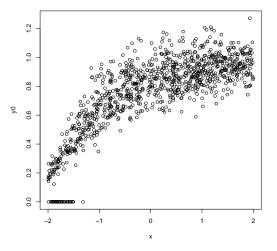
Figure 8: Showing the three functions $\mu = f_{\mu}(x)$, $\sigma = f_{\sigma}(x)$ and $\nu = f_{\nu}(x)$ used for the simulation of the data in this section. From top left to the bottom we have the functions for μ , σ and ν .

Next we randomly generate the response variable values, y0, from a zero adjusted gamma distribution model, ZAGA(μ , σ , ν) where $\mu = f_{\mu}(x)$, $\sigma = f_{\sigma}(x)$ and $\nu = f_{\nu}(x)$. Figure 9 shows the response variable y0 against x.

```
# generating y0 -----
set.seed(123)
y0 <- rZAGA(1000, mu=fmu(x), sigma=fsigma(x), nu=fnu(x))
plot(x,y0)</pre>
Figure 9
```

We will fit a zero adjusted distribution using the function gamlss(). The same model will be also fitted using the gamlssZadj() function. Since, in general, the type of relationship existing between the parameters and the explanatory variable x is unknown we will use smooth functions for x. In the following code we used the P-spline smother implemented in the additive term function pb().

```
g0p <- gamlss(y0~pb(x), sigma.fo=~pb(x), nu.fo=~pb(x), family=ZAGA)
## GAMLSS-RS iteration 1: Global Deviance = -1503.31</pre>
```



R code on page 19

Figure 9: Showing the response variable y_0 against the single explanatory variable x for the simulated from a zero adjusted gamma, $\mathsf{ZAGA}(\mu, \sigma, \nu)$ distributed response variable y_0 with values defined on $[0, \infty)$.

```
## GAMLSS-RS iteration 2: Global Deviance = -1502.221
## GAMLSS-RS iteration 3: Global Deviance = -1502.186
## GAMLSS-RS iteration 4: Global Deviance = -1502.184
## GAMLSS-RS iteration 5: Global Deviance = -1502.184
t0p \leftarrow gamlssZadj(y=y0, mu.fo=pb(x), sigma.fo=pb(x),
                 xi0.fo=~pb(x), family="GA", trace=TRUE)
              The binomial model
                                        ****
## GAMLSS-RS iteration 1: Global Deviance = 155.0575
## GAMLSS-RS iteration 2: Global Deviance = 155.0575
## **** The continuous distribution model ****
## GAMLSS-RS iteration 1: Global Deviance = -1658.367
## GAMLSS-RS iteration 2: Global Deviance = -1657.278
## GAMLSS-RS iteration 3: Global Deviance = -1657.243
## GAMLSS-RS iteration 4: Global Deviance = -1657.241
## GAMLSS-RS iteration 5: Global Deviance = -1657.241
              The Final Global Deviance = -1502.184
##
AIC(g0p, t0p)
##
            df
                     AIC
## g0p 16.72940 -1468.725
## t0p 16.72942 -1468.725
summary(t0p)
## **********************
```

```
## Family: "ZadjGA"
## Call: gamlssZadj(y = y0, mu.formula = pb(x), sigma.formula = pb(x),
   xi0.formula = pb(x), family = "GA", trace = TRUE
##
## Fitting method: RS()
##
## Mu link function: log
## Mu Coefficients:
          Estimate Std. Error t value Pr(>|t|)
## (Intercept) -0.317598   0.004905   -64.75   <2e-16 ***
## pb(x) 0.247489 0.004178 59.23 <2e-16 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## -----
## Sigma link function: log
## Sigma Coefficients:
## Estimate Std. Error t value Pr(>|t|)
## pb(x) -0.21021 0.02101 -10.01 <2e-16 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
## -----
## xi0 link function: logit
## xi0 Coefficients:
## Estimate Std. Error t value Pr(>|t|)
## (Intercept) -7.126 2.054 -3.470 0.000544 ***
        -2.781 1.176 -2.366 0.018174 *
## pb(x)
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## -----
## NOTE: Additive smoothing terms exist in the formulas:
## i) Std. Error for smoothers are for the linear effect only.
## ii) Std. Error for the linear terms maybe are not accurate.
## -----
## No. of observations in the fit: 1000
## Degrees of Freedom for the fit: 16.72942
      Residual Deg. of Freedom: 983.2706
##
                  at cycle:
##
## Global Deviance:
                 -1502.184
##
           AIC:
                 -1468.725
                -1386.621
##
           SBC:
## *********************
```

The two models have the same deviance as we would expect. Make sure not to try to interpret the coefficients and the standard errors of the smoothing functions in the summary table. They are here as a consequence of how the model is fitted within the gamlss() algorithm, (which uses backfitting to alternate between fitting the linear and smoothing component) and they are for the linear term rather than indicating the coefficients and the standard errors of the smoothing functions. To roughly check whether the smoothing function as a whole is significant use the function drop1(). In our case since we are dealing with simulated data we can actually plot the true functions $\mu = f_{\mu}(x)$, $\sigma = f_{\sigma}(x)$ and $\nu = f_{\nu}(x)$ with their fitted functions. Figure 10 shows the results. Note that the fitted values for the parameters μ , σ and ν (or ξ_0) are identical in both fitted models.

```
# mu
curve(fmu, -2,2, main="mu")
lines(fitted(g0p)[order(x)]~x[order(x)], col="red", lty=2, lwd=2)
lines(fitted(t0p, "mu")[order(x)]~x[order(x)], col="blue", lty=2, lwd=2)
# sigma
curve(fsigma, -2,2, main="sigma")
lines(fitted(g0p, "sigma")[order(x)]~x[order(x)], col="red", lty=2, lwd=2)
lines(fitted(t0p, "sigma")[order(x)]~x[order(x)], col="blue", lty=2, lwd=2)
# nu
curve(fnu, -2,2, main="nu")
lines(fitted(g0p, "nu")[order(x)]~x[order(x)], col="red", lty=2, lwd=2)
lines(fitted(t0p, "xi0")[order(x)]~x[order(x)], col="red", lty=2, lwd=2)
lines(fitted(t0p, "xi0")[order(x)]~x[order(x)], col="blue", lty=2, lwd=2)
```

Note that the term.plot() functions is not working for models fitted through the gamlssZadj() function so use the function term.plotZadj(). Figure 11 shows the fitted additive terms using the gamlss ZAGA fit on the left and the gamlssZadj fit on the right.

```
term.plot(g0p);title("gamlss")
term.plotZadj(t0p);title("gamlssZadj")
term.plot(g0p, "sigma")
term.plotZadj(t0p, "sigma")
term.plot(g0p, "nu")
term.plotZadj(t0p, "xi0")
```

Note that diagnostics plots like the residual plot, plot(), the worm plot, wp(), and the Q-statistics plot, Q.stats(), are all working with a gamlssZadj object. Also in the case in which only one explanatory variable exists, the function centile.Zadj() can be used to plot centile curves, see Figure 12.

```
plot(t0p)
                                                                Figure 12
## *********************
##
   Summary of the Randomised Quantile Residuals
##
                      mean = -0.00233167
##
                   variance
                          = 1.007856
             coef. of skewness = -0.02696412
##
##
             coef. of kurtosis = 2.950955
## Filliben correlation coefficient = 0.9996238
## **********************
```

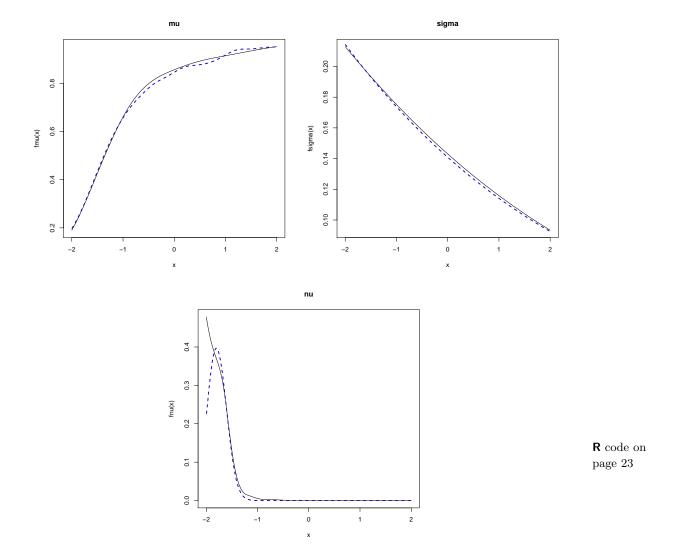


Figure 10: Showing the functions from which the data were simulated, together with the fitted smooth functions, for μ , σ and ν (or ξ_0). The solid black line is the true function. The dashed line is the fitted function from both models \mathfrak{gop} and \mathfrak{top} since they are identical.

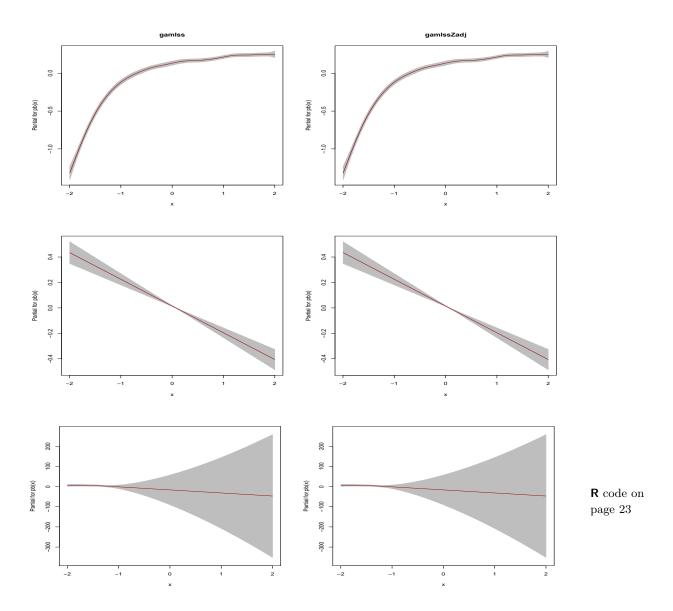


Figure 11: Showing the fitted additive predictors for μ , σ and ν and their approximate 95% confidence bands for models fitted using the function gamlss() on the left and using the function gamlssZadj() on the right.

```
wp(t0p, ylim.all=1)
Q.stats(t0p, xvar=x)
## Warning in Q.stats(t0p, xvar = x): the number of intervals have change to 13
##
##
                           Z1
                                       Ζ2
                                                 Z3
                                                            Z4
## -1.9951 to -1.6674 -0.5160602 -0.07459136 -0.8466418 -0.01777688
## -1.6674 to -1.3183 0.3294259 0.11634319 0.4385309 -0.68914681
## -1.3183 to -0.9880 -0.4294565 0.92735955 0.1518008 0.81464548
## -0.9880 to -0.6424 0.3398515 -0.43128824 1.3475734 -1.94348926
## -0.3589 to -0.0118 -0.5828591 -1.16560779 0.1170826 0.61381914
## -0.0118 to 0.3096 0.2503127 0.96743562 -0.7127370 -0.40662572
## 0.3096 to 0.5540 -0.6179411 -0.67474528 0.6117738 -0.68051577
## 0.5540 to 0.8607 1.0656147 0.52222204 -0.1508687 -0.46407877
## 0.8607 to 1.0907 -0.7681959 0.74471171 -0.5059944 1.45314330
## 1.0907 to 1.3660 0.4438157 0.77534908 -0.1793687 0.26486940
## 1.3660 to 1.6778 -0.3202311 -1.18703600 -1.2576778 0.64733508
## 1.6779 to 1.9992 0.1351939 -0.31670204 0.2662090 1.38642236
## TOTAL Q stats
                     3.6676560 6.90549334 6.1320730 11.73838077
## df for Q stats
                     2.5274836 11.49838086 13.0000000 13.00000000
                     ## p-val for Q stats
##
                    AgostinoK2
                                 Ν
                                77
## -1.9951 to -1.6674 0.7171184
## -1.6674 to -1.3183 0.6672327
                                77
## -1.3183 to -0.9880 0.6866907
                                77
## -0.9880 to -0.6424 5.5931046
                                77
## -0.6424 to -0.3589 1.6031721
                                77
## -0.3589 to -0.0118 0.3904823
                                77
## -0.0118 to 0.3096 0.6733385
                                76
   0.3096 to 0.5540 0.8373689
                                77
## 0.5540 to 0.8607 0.2381305
                                77
## 0.8607 to 1.0907 2.3676558
                                77
## 1.0907 to 1.3660 0.1023289
                                77
## 1.3660 to 1.6778 2.0007961
                                77
## 1.6779 to 1.9992 1.9930342
                                77
## TOTAL Q stats
                    17.8704537 1000
## df for Q stats
                    26.0000000
                                 0
## p-val for Q stats
                     0.8804356
                                 0
centiles.Zadj(t0p, xvar=x)
## % of cases below 0.4 centile is 4.1
## % of cases below 2 centile is 5.5
## % of cases below 10 centile is 11.8
## % of cases below 25 centile is
                                25.7
## % of cases below 50 centile is
                                 50.3
## % of cases below 75 centile is
                                75.2
## % of cases below 90 centile is
                                 90.5
```

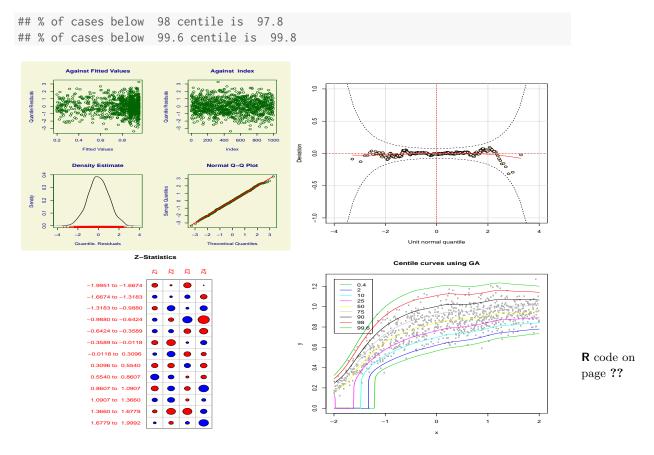


Figure 12: Showing the diagnostic plots created by plot(), wp(), and Q.stats(), and the centile curves produced by centile.Zadj().

Note that the % of cases below the 0.4, 2, 20 and 25 centile curves should be ignored because their centiles curves include value Y = 0, see the centiles curves in Figure 12.

6.2 Fitting alternative zero adjusted distributions on $[0, \infty)$.

6.2.1 Zero adjusted log-transform distributions on $[0, \infty)$.

The function gamlssZadj() can be used with a log-transform distribution on $(0, \infty)$ or with a truncated distribution on $(0, \infty)$. For zero adjusted log-transform distribution use the following code:

```
# generate a log-tranform distribution from 0 to infinity
gen.Family("SST", "log")

## A log family of distributions from SST has been generated
## and saved under the names:
## dlogSST plogSST qlogSST rlogSST
```

6.2.2 Zero adjusted truncated distributions on $[0, \infty)$.

For a zero adjusted truncated distribution use:

```
# generate a left truncated distribution SST from 0 to Infty
library(gamlss.tr)
gen.trun(c(0), "SST", type="left")
## A truncated family of distributions from SST has been generated
## and saved under the names:
## dSSTtr pSSTtr qSSTtr rSSTtr SSTtr
## The type of truncation is left
   and the truncation parameter is 0
# fit the model
t0ssttr <- gamlssZadj(y=y0, mu.fo=~pb(x), sigma.fo=~pb(x),
                       xi0.fo= pb(x), family="SSTtr")
GAIC(g0p, t0p, t0sst,t0ssttr )
##
                 df
                          AIC
           16.72940 -1468.725
## g0p
## t0p
           16.72942 -1468.725
## t0sst 18.67549 -1463.862
## t0ssttr 10.37618 -1060.759
```

Note that both $logSST(\mu, \sigma, \nu, \tau)$ and $SSTtr(\mu, \sigma, \nu, \tau)$ distributions have four parameters, so their corresponding zero adjusted distributions have five parameters [including the extra parameters $\xi_0 = P(Y=0)$]. Hence models tosst and tosstr can include models for parameters ν and τ , e.g. formula.nu=~pb(x), as well as models for μ , σ and ξ_0 .

6.2.3 Generalised Tobit model distributions on $[0, \infty)$.

In general for a restricted values response variable, that is, having a distribution on a restricted range, the Tobit model (which requires a survival analysis response variable), can be appropriate. Here we show how this model can be fitted within gamlss. Note though that for any Tobit model the probability at zero can **not** be modelled independently as a function of explanatory variables, but is equal to the probability of being left censored at zero.

Below we fit a Tobit normal model and a Tobit SST model, both left censored at 0 to provide point probabilities at 0.

```
library(survival)
y0surv<- Surv(y0, y0!=0, type="left")
# creating the distribution
library(gamlss.cens)
# Gaussian</pre>
```

```
gen.cens("NO", type="left")
## A censored family of distributions from NO has been generated
## and saved under the names:
## dNOlc pNOlc qNOlc NOlc
## The type of censoring is left
# SST distribution
gen.cens("SST", type="left")
## A censored family of distributions from SST has been generated
## and saved under the names:
## dSSTlc pSSTlc qSSTlc SSTlc
## The type of censoring is left
# fitting the model
# Tobit model
s0no <- gamlss( y0surv ~ pb(x), sigma.formula=~pb(x),</pre>
               family=NOlc)
## GAMLSS-RS iteration 1: Global Deviance = -1315.557
## GAMLSS-RS iteration 2: Global Deviance = -1314.02
## GAMLSS-RS iteration 3: Global Deviance = -1313.902
## GAMLSS-RS iteration 4: Global Deviance = -1313.863
## GAMLSS-RS iteration 5: Global Deviance = -1313.848
## GAMLSS-RS iteration 6: Global Deviance = -1313.841
## GAMLSS-RS iteration 7: Global Deviance = -1313.839
## GAMLSS-RS iteration 8: Global Deviance = -1313.838
## GAMLSS-RS iteration 9: Global Deviance = -1313.837
# generalised Tobit
s0sst <- gamlss( y0surv ~ pb(x), sigma.formula=~pb(x),</pre>
               family=SSTlc)
## GAMLSS-RS iteration 1: Global Deviance = -1290.533
## GAMLSS-RS iteration 2: Global Deviance = -1298.536
## GAMLSS-RS iteration 3: Global Deviance = -1302.704
## GAMLSS-RS iteration 4: Global Deviance = -1304.586
## GAMLSS-RS iteration 5: Global Deviance = -1305.147
## GAMLSS-RS iteration 6: Global Deviance = -1305.298
## GAMLSS-RS iteration 7: Global Deviance = -1305.348
## GAMLSS-RS iteration 8: Global Deviance = -1305.365
## GAMLSS-RS iteration 9: Global Deviance = -1305.373
## GAMLSS-RS iteration 10: Global Deviance = -1305.374
## GAMLSS-RS iteration 11: Global Deviance = -1305.375
GAIC(g0p, t0p, t0sst,t0ssttr, s0no, s0sst)
##
                 df
                          AIC
## g0p
           16.72940 -1468.725
## t0p
          16.72942 -1468.725
## t0sst 18.67549 -1463.862
```

```
## s0no 13.03170 -1287.774
## s0sst 11.98696 -1281.401
## t0ssttr 10.37618 -1060.759
```

The Tobit models do not perform as well as the zero adjusted models in this case. This is not surprising since the response variable y0 was simulated from a ZAGA distribution (see earlier Section 5.2) which is a zero adjusted distribution.

6.3 Real data example.

The data we will analyse here are motor vehicle insurance data, i.e. motor vehicle insurance policies. The data frame mvi is a sample of 2000 observations from the data frame mviBig which has 67143 observations.

```
R data file: mvi in package gamlss.data of dimensions 2000 \times 11
var retval: a numeric vector showing the value of the vehicle
     whetherclm: a numeric vector showing whether a claim is made, 0 no claim, 1 at
         least one claim
     numclaims : a numeric vector showing the number of claims
     claimcst0: a numeric vector showing the total amount of claim, so if for num-
         claims=0 it is zero.
     vehmake: a factor showing the make of the car with levels BMW DAEWOO FORD
         MITSUBISHI.
     vehbody: a factor showing the type of the car, with levels BUS CONT COUPE
         HACK HDTOP HRSE MCARA MIBUS PANVN RDSTR SEDAN STNWG
         TRUCK UTE.
     vehage: a numeric vector showing the age of the car
     gender: a factor showing the gender of the policy holder with levels F M
     area: factor showing the area of residence of the policy holder with levels A B C D
     agecat: a factor showing the age band of the policy holder with levels 1 2 3 4 5 6
         one is youngest
     exposure: a numeric vector showing the time of exposure with values from zero to
         one
```

In the following analysis we ignore the exposure time and we fit zero adjusted models, with models for μ and for the probability at zero (i.e. ν or ξ_0).

```
data(mvi)
# zero adjusted GA
m1 <- gamlss(claimcst0~vehmake+vehbody+vehage+gender+area,</pre>
```

```
nu.fo=~vehmake+vehbody+vehage+gender+area,
             family=ZAGA, data=mvi )
## GAMLSS-RS iteration 1: Global Deviance = 3229.012
## GAMLSS-RS iteration 2: Global Deviance = 3229.011
# zero adjusted IG
m2 <- gamlss(claimcst0~vehmake+vehbody+vehage+gender+area,</pre>
             nu.fo=~vehmake+vehbody+vehage+gender+area,
             family=ZAIG, data=mvi )
## GAMLSS-RS iteration 1: Global Deviance = 3227.486
## GAMLSS-RS iteration 2: Global Deviance = 3225.318
## GAMLSS-RS iteration 3: Global Deviance = 3225.317
# zero adjusted GG
m3 <- gamlssZadj(claimcst0, ~vehmake+vehbody+vehage+gender+area,
             xi0.fo=~vehmake+vehbody+vehage+gender+area,
             family=GG, data=mvi, trace=T, n.cyc=30 )
## ****
               The binomial model
                                         ****
## GAMLSS-RS iteration 1: Global Deviance = 942.9831
## GAMLSS-RS iteration 2: Global Deviance = 942.9828
## **** The continuous distribution model ****
## GAMLSS-RS iteration 1: Global Deviance = 2284.271
## GAMLSS-RS iteration 2: Global Deviance = 2280.202
## GAMLSS-RS iteration 3: Global Deviance = 2275.738
## GAMLSS-RS iteration 4: Global Deviance = 2271.351
## GAMLSS-RS iteration 5: Global Deviance = 2267.604
## GAMLSS-RS iteration 6: Global Deviance = 2264.868
## GAMLSS-RS iteration 7: Global Deviance = 2263.079
## GAMLSS-RS iteration 8: Global Deviance = 2261.965
## GAMLSS-RS iteration 9: Global Deviance = 2261.276
## GAMLSS-RS iteration 10: Global Deviance = 2260.844
## GAMLSS-RS iteration 11: Global Deviance = 2260.568
## GAMLSS-RS iteration 12: Global Deviance = 2260.389
## GAMLSS-RS iteration 13: Global Deviance = 2260.271
## GAMLSS-RS iteration 14: Global Deviance = 2260.192
## GAMLSS-RS iteration 15: Global Deviance = 2260.138
## GAMLSS-RS iteration 16: Global Deviance = 2260.101
## GAMLSS-RS iteration 17: Global Deviance = 2260.076
## GAMLSS-RS iteration 18: Global Deviance = 2260.059
## GAMLSS-RS iteration 19: Global Deviance = 2260.047
## GAMLSS-RS iteration 20: Global Deviance = 2260.039
## GAMLSS-RS iteration 21: Global Deviance = 2260.033
## GAMLSS-RS iteration 22: Global Deviance = 2260.029
## GAMLSS-RS iteration 23: Global Deviance = 2260.026
## GAMLSS-RS iteration 24: Global Deviance = 2260.024
## GAMLSS-RS iteration 25: Global Deviance = 2260.022
## GAMLSS-RS iteration 26: Global Deviance = 2260.021
```

```
The Final Global Deviance = 3203.004
##
# zero adjusted BCTo
m4 <- gamlssZadj(claimcst0, ~vehmake+vehbody+vehage+gender+area,
             xi0.fo=~vehmake+vehbody+vehage+gender+area,
             family=BCTo, data=mvi, trace=T )
               The binomial model
## ****
## GAMLSS-RS iteration 1: Global Deviance = 942.9831
## GAMLSS-RS iteration 2: Global Deviance = 942.9828
## **** The continuous distribution model ****
## GAMLSS-RS iteration 1: Global Deviance = 2297.874
## GAMLSS-RS iteration 2: Global Deviance = 2262.812
  GAMLSS-RS iteration 3: Global Deviance = 2261.301
## GAMLSS-RS iteration 4: Global Deviance = 2260.482
## GAMLSS-RS iteration 5: Global Deviance = 2260.253
## GAMLSS-RS iteration 6: Global Deviance = 2260.177
## GAMLSS-RS iteration 7: Global Deviance = 2260.155
## GAMLSS-RS iteration 8: Global Deviance = 2260.149
## GAMLSS-RS iteration 9: Global Deviance = 2260.147
  GAMLSS-RS iteration 10: Global Deviance = 2260.147
##
               The Final Global Deviance = 3203.129
AIC(m1, m2, m3, m4)
##
      df
              AIC
## m3 48 3299.004
## m4 49 3301.129
## m2 47 3319.317
## m1 47 3323.011
```

The best model according to AIC seems to be the model m3 with a zero adjusted generalised gamma distribution, but it is not very far from model m4, the zero adjusted BCTo distribution. Figure 13 shows the worm plot for both distributions with the zero adjusted BCTo looking a bit better.

```
wp(m3, ylim.all=1); title('(a)')
wp(m4, ylim.all=1); title('(b)')
```

Figure 13

7 Conclusions

GAMLSS is a framework where different models can be fitted and compared. In this vignette, we have shown how models with a response variable defined on $[0, \infty)$, including 0, (i.e. the nonnegative real line), can be fitted within the GALMSS framework. The gamlssZadj() function can be used to fit a variety of different zero adjusted models, in which the response variable lies on $[0, \infty)$. A zero adjusted distribution for Y comprises a value Y = 0 with probability p_0 and $Y = Y_1$ with probability $(1 - p_0)$ where Y_1 has a continuous distribution on the interval $(0, \infty)$, i.e. the positive real line. In gamlssZadj() the parameter xi0, (ξ_0) equals $p_0 = P(Y = 0)$. The continuous distribution component must either exist already within gamlss.dist or be created

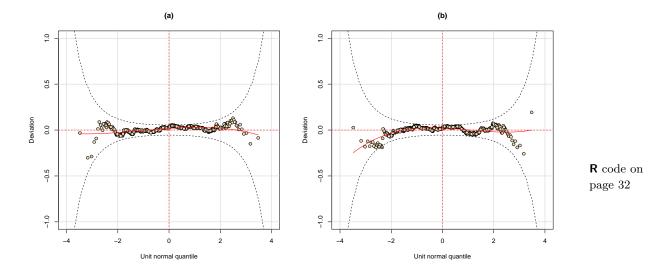


Figure 13: The worm plot of the residuals from model (a) m3, the zero adjusted generalised gamma distribution and (b) m4 the zero adjusted BCTo distribution.

using the gen.Family() or gen.trun() function for transformed or truncated distributions, respectively. In addition the function gen.Zadj() generates d, p, q, r and plot functions for a zero adjusted distribution.

More information about GAMLSS can be found in Stasinopoulos et al. [2017] or the GAMLSS website www.gamlss.org. We're hoping that the gamlss.inf package will be useful.

References

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