# Nonlinear Regression (Part 3)

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Stanford University, Spring 2016, STATS 205

Last time:

► Linear Smoothers

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  - Local Averages

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  - ► Local Regression

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- ► Bootstrap Confidence Bands

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▶ For now, assume that variance  $Var(\epsilon_i) = \sigma^2$  is independent of x



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- Choice of bandwidth matters which controls the amount of smoothing
- Small bandwidths give very rough estimates while larger bandwidths give smoother estimates

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  - ▶ The bias will be higher, because we are now using observations  $x_i$  further from  $x_0$ , and there is no guarantee that  $r(x_i)$  will be close to  $r(x_0)$

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- ► We use the data twice: to estimate the function and to estimate the risk

▶ A better idea is to use leave-one-out cross-validation

$$cv = \widehat{R}(h) = \frac{1}{n} \sum_{i=1}^{n} (Y_i - \widehat{r}_{(-i)}(x_i))^2$$

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$$I_{j,(-i)}(x) = \begin{cases} 0 & \text{if } j = i \\ \frac{I_j(x)}{\sum_{k \neq i} I_k(x)} & \text{if } j \neq i \end{cases}$$

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► Cross-validation is approximately the predictive risk (predicting the left-one-out observation)



▶ We can compute leave-one-out cross-validation without leaving one observation out

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- ▶ After some algebra, we can see that

$$\widehat{r}(x_i) = (1 - L_{ii})\widehat{r}_{(-i)}(x_i) + L_{ii}Y_i$$

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$$\mathsf{E}(\widehat{\sigma}^2) = \frac{\mathsf{E}(Y^T \Lambda Y)}{\mathsf{tr}(\Lambda)} = \sigma^2 + \frac{r^T \Lambda r}{n - 2\nu + \widetilde{\nu}}$$

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$$\Lambda = (I - L)^T (I - L)$$

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- Assuming that  $\nu$  and  $\widehat{\nu}$  do not grow too quickly, and that r is smooth, the second term is small for large n
- ▶ So E( $\hat{\sigma}^2$ )  $\approx \sigma^2$
- and one can show that  $Var(\widehat{\sigma^2}) \to 0$



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Therefore

$$\mathsf{E}(Y_{i+1}-Y_i)\approx \mathsf{E}(\epsilon_{i+1})+\mathsf{E}(\epsilon_i)=2\sigma^2$$



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with  $\bar{r}_n(x)$  being the mean

- First term converges to a normal
- ▶ If we do a good job trading off bias and variance, the second term doesn't vanish with large n

$$\frac{\overline{r}_n(x) - r(x)}{\widehat{\sigma}(x)} = \frac{\mathsf{Bias}(\widehat{r}_n(x))}{\sqrt{\mathsf{Variance}(\widehat{r}_n(x))}}$$

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- ▶ We will go with the first approach

▶ For linear smoother  $\hat{r}_n(x)$  with

$$\bar{r}(x) = \mathsf{E}(\hat{r}_n(x)) = \sum_{i=1}^n l_i(x) r(x_i)$$

and assuming constant variance

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and assuming constant variance

$$Var(\widehat{r}_n(x)) = \sigma^2 ||I(x)||^2$$

Consider confidence bands

$$\mathcal{I}(x) = (\hat{r}_n(x) - c\hat{\sigma} || I(x) ||, \hat{r}_n(x) + c\hat{\sigma} || I(x) ||)$$

for some c and  $a \le x \le b$ 

For now, suppose that  $\sigma$  is known, then probability of estimate not in confidence band in at least one position x

$$\mathsf{P}(\overline{r}(x) \notin \mathcal{I}(x) \text{ for some } x \in [a,b]) = \mathsf{P}\left(\max_{x \in [a,b]} \frac{|\widehat{r}(x) - \overline{r}|}{\sigma \|I(x)\|} > c\right)$$

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We are left just with the error term

$$= \mathsf{P}\left(\max_{x \in [a,b]} \frac{|\sum_{i} \epsilon_{i} l_{i}(x)|}{\sigma \|I(x)\|} > c\right) = \mathsf{P}\left(\max_{x \in [a,b]} |W(x)| > c\right)$$

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▶ This is a Gaussian process: a random function such that the vector  $(W(x_1), \ldots, W(x_k))$  has a multivariate normal distribution, for any finite set of points  $x_1, \ldots, x_k$ 

$$W(x) = \sum_{i=1}^{n} Z_i T_i(x), \quad Z_i = \epsilon_i / \sigma \sim N(0, 1), \quad T_i(x) = I_i(x) \|I(x)\|$$

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  - In our neuroimaging example, we used permutation test to find maximum voxel clusters

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$$P\left(\max_{x}\left|\sum_{i=1}^{n}Z_{i}T_{i}(x)\right|>c\right)\approx 2(1-\Phi(c))+\frac{\kappa_{0}}{\pi}e^{-c^{2}/2}$$

for large 
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- Intuition: The task is to calculate the volume of this tube

 $\triangleright$  One can show that (cdf of the standard normal  $\Phi$ )

$$P\left(\max_{x}\left|\sum_{i=1}^{n}Z_{i}T_{i}(x)\right|>c\right)\approx 2(1-\Phi(c))+\frac{\kappa_{0}}{\pi}e^{-c^{2}/2}$$

for large c,  $\kappa_0 = \int_a^b \|T'(x)\| dx$ , and  $T'(x) = \partial T_i(x)/\partial x$ 

- ▶ Think of T(x) as a curve in  $\mathbb{R}^n$ , and c as defining a tube around it with radius c
- Intuition: The task is to calculate the volume of this tube
- We choose *c* by solving for  $\alpha$  (e.g.  $\alpha = 0.05$ )

$$2(1 - \Phi(c)) + \frac{\kappa_0}{\pi} e^{-c^2/2} = \alpha$$

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Then this confidence is used

$$\mathcal{I}(x) = \hat{r}_n(x) \pm c \sqrt{\sum_{i=1}^n \hat{\sigma}^2(x_i) I_i^2(x)}$$

with c computed the same way

## Average Coverage

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- We can relax this requirement a bit
- ▶ Suppose we are estimating r(x) over an interval [0,1], then average coverage is defined as

$$C = \int_0^1 \mathsf{P}(r(x) \in [d(x), u(x)]) dx$$

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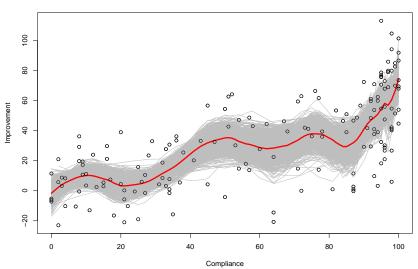
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  - Errors need to be iid

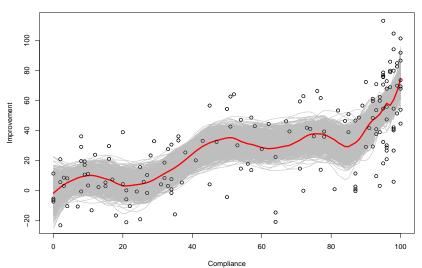
Experiment with n = 164 men to see if the drug cholostyramine lowered blood cholesterol levels

- ► Experiment with *n* = 164 men to see if the drug cholostyramine lowered blood cholesterol levels
- ► They were supposed to take six packets of cholostyramine per day, but many actually took much less

#### Resample Rows Bootstrap



#### Resample Residuals Bootstrap



#### References

▶ Wasserman (2006). All of Nonparametric Statistics

#### References

- ▶ Wasserman (2006). All of Nonparametric Statistics
- ▶ Efron and Tibshirani (1994). An Introduction to the Bootstrap