

June-2021-Exam

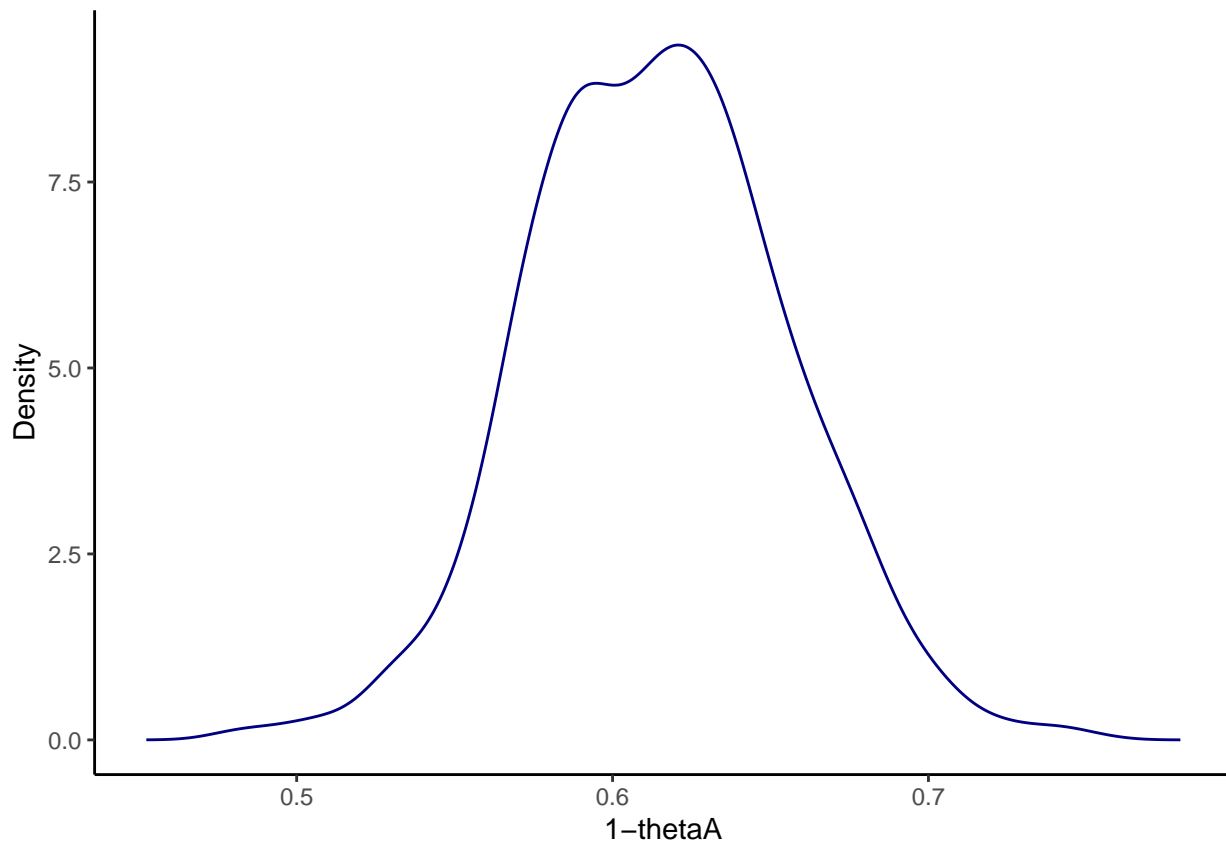
Christophoros Spyretos

Exercise 1 *Customers' choice of brands*

```
n <- 100  
s <- 38  
f <- 62  
a <- 16  
b <- 24
```

Task 1

```
set.seed(12345)  
  
#generates 1,000 random deviates.  
theta_A <- rbeta(n = 1000, shape1 = a+s, shape2 = b+f)  
  
#exact posterior prob  
#pbeta the distribution function  
thetaA_prob <- pbeta(q = 0.4, shape1 = a+s, shape2 = b+f, lower.tail = FALSE)  
  
dens_one_minus_thetaA <- density(1 - theta_A)  
  
df_plot_density <- data.frame("x" = dens_one_minus_thetaA$x, "y" = dens_one_minus_thetaA$y)  
  
ggplot(df_plot_density) +  
  geom_line(aes(x=x, y=y), color = "navy") +  
  xlab("1-thetaA") +  
  ylab("Density") +  
  theme_classic()
```



The posterior probability of $\theta_A > 0.4$ equals approximately 0.36.

Task 2

```
ratio <- (1-theta_A)/theta_A
interval <- quantile(ratio,probs = c(0.025,0.975))

#table for the interval
df_intervals <- data.frame("lower_bound" = interval[1], "upper_bound" = interval[2])
colnames(df_intervals) <- c("lower bound", "upper bound")
rownames(df_intervals) <- c("95% Equal Tail Credible Interval")
knitr::kable(df_intervals)
```

	lower bound	upper bound
95% Equal Tail Credible Interval	1.14581	2.272752

The ratio is the odds of not choosing brand A, i.e. it describes how many more times likely it is to not choose brand A compared to choosing brand A. The credible interval shows the values of the ratio with 95 % probability.

Task 3

```
beta(a+s,b+f)/beta(a,b)
```

```
## [1] 7.556771e-30
```

Task 4

```
set.seed(12345)
counts <- c(38,27,35)
c <- 2
a <- c*c(1,1,1)

N <- 1000
xDraws <- matrix(0,N,length(a))
thetaDraws <- matrix(0,N,length(a))

for (i in 1:length(a)){
  xDraws[,i] <- rgamma(N,a[i]+counts[i], rate = 1)
}

for (i in 1:N){
  thetaDraws[i,] <- xDraws[i,]/sum(xDraws[i,])
}

mean_val <- mean(thetaDraws[,1] > thetaDraws[,3])
```

The posterior probability is 0.656.

Exercise 2

Task a,b and c are hand written solutions.

Task d

```
LogPost <- function(theta,n, sum_x2){

  logLik <- n*log(theta) - theta*sum_x2
  logPrior <- -0.5 * theta

  return(logLik + logPrior)
}

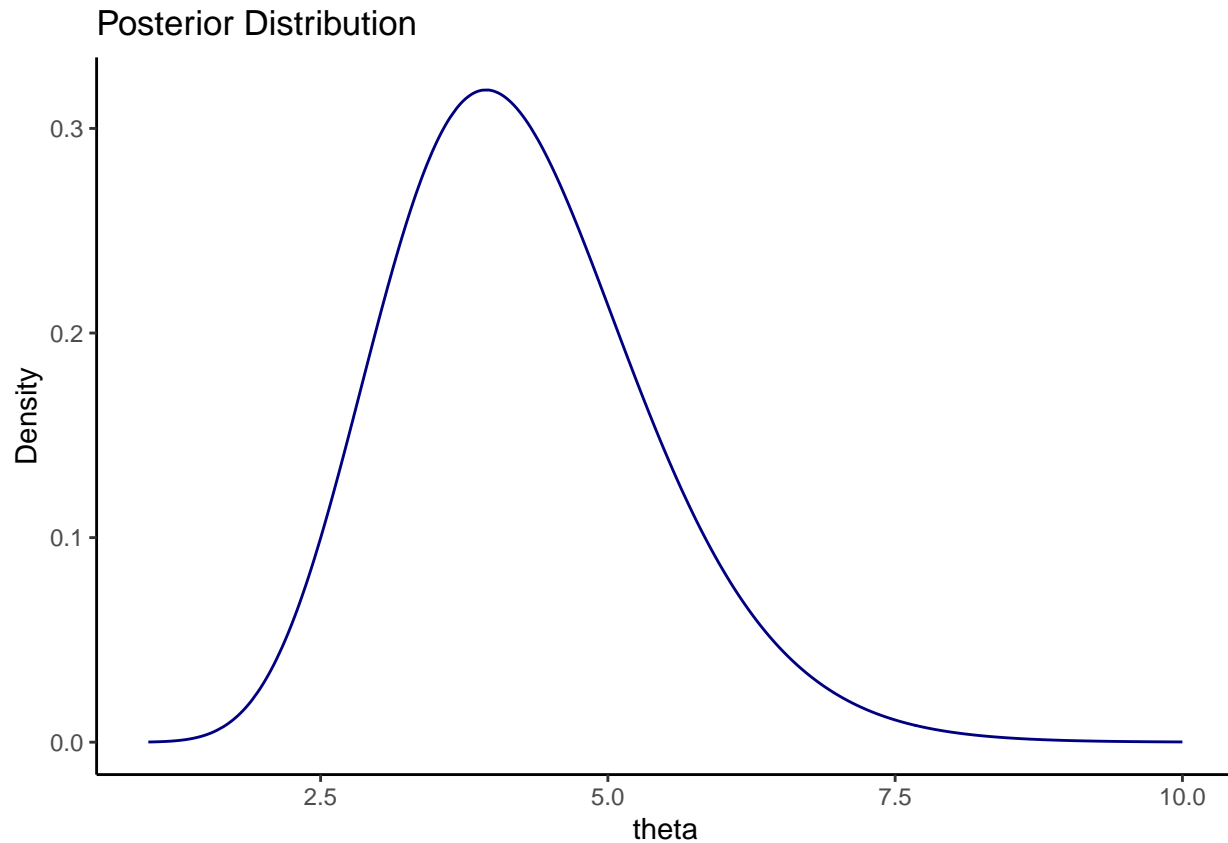
theta <- runif(1000,1,10)
n <- 13
sum_x2 <- 2.8

post_dens <- exp(LogPost(theta,n, sum_x2))

#normalise posterior density
post_dens <- post_dens/(0.01 * sum(post_dens))

df_plot <- data.frame("theta" = theta, "posterior" = post_dens)

ggplot(df_plot) +
  geom_line(aes(x=theta, y=posterior), color="navy") +
  ggtitle("Posterior Distribution") +
  ylab("Density") +
  theme_classic()
```



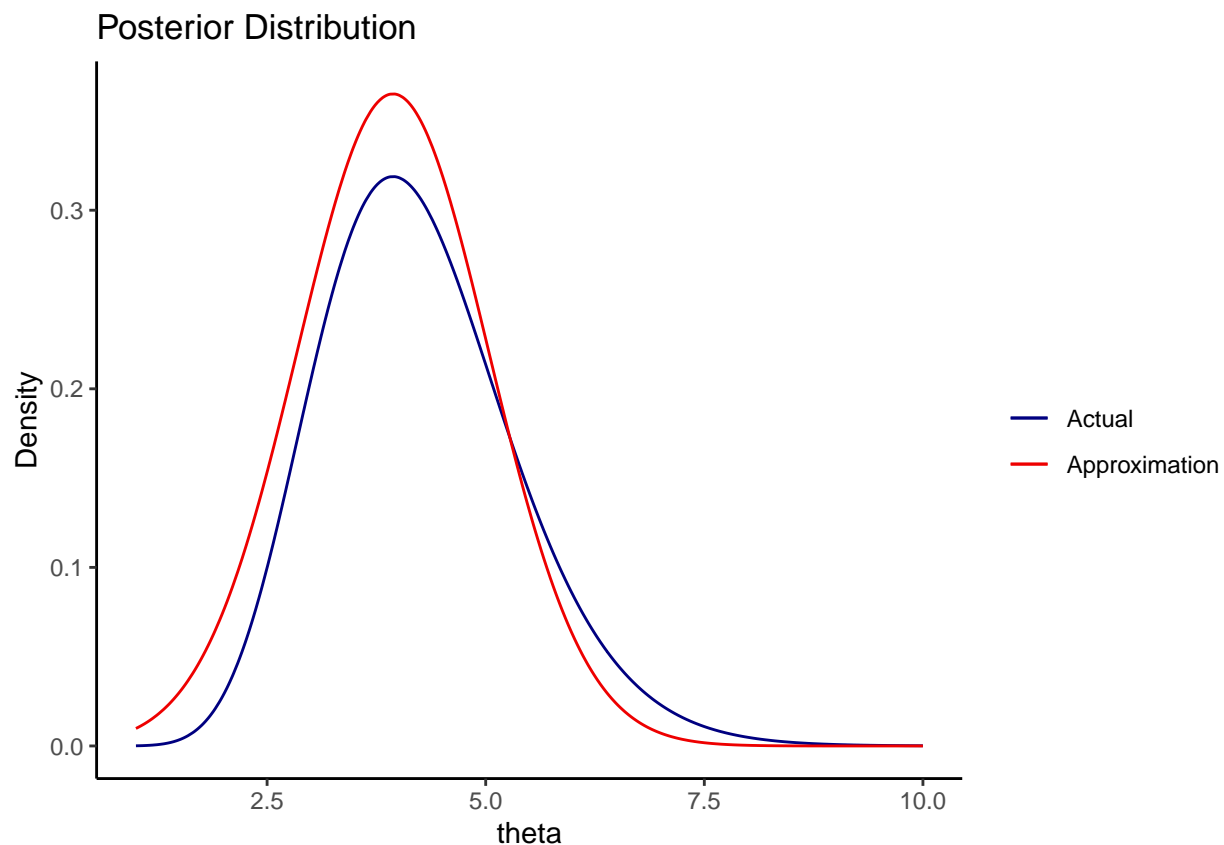
Task e

```
OptimRes <- optim(3, LogPost, gr=NULL, n, sum_x2, method=c("BFGS"), control=list(fnscale=-1), hessian=TRUE)

normal_approx <- dnorm(theta, OptimRes$par, sqrt(diag(-solve(OptimRes$hessian))))

df_plot$approximation <- normal_approx

ggplot(df_plot) +
  geom_line(aes(x=theta, y=posterior, color="navy")) +
  geom_line(aes(x=theta, y=approximation, color="red2")) +
  theme(legend.position="right") +
  scale_color_manual(values=c('navy', 'red2'),
                     name = "",
                     labels = c("Actual", "Approximation")) +
  ggtitle("Posterior Distribution") +
  ylab("Density") +
  theme_classic()
```



The posterior approximation is quite accurate, but the actual posterior distribution is skewed to the right.

Exercise 3

```
source("ExamData.R")
```

Task a

```
#BayesLinReg <- function(y, X, mu_0, Omega_0, v_0, sigma2_0, nIter)
library(mvtnorm)

mu_0 <- rep(0,7)
Omega_0 <- 1/25 * diag(7)
v_0 <- 1
sigma2_0 <- 4
nIter <- 1000

PostDraws <- BayesLinReg(y, X, mu_0, Omega_0, v_0, sigma2_0, nIter)

Betas <- PostDraws$betaSample

BetasMean <- colMeans(Betas)

intervals <- matrix(NA, nrow = 7, ncol = 2)

for (i in 1:ncol(Betas)){
```

```

  intervals[i,] <- quantile(Betas[,i], probs = c(0.025, 0.975))
}

colnames(intervals) <- c("2.5%", "97.5")
rownames(intervals) <- c("b0", "b1", "b2", "b3", "b4", "b5", "b6")
knitr::kable(intervals)

```

	2.5%	97.5
b0	1.1529187	1.4674344
b1	0.5377850	0.8711292
b2	0.0499087	0.2702009
b3	0.0225842	0.8239922
b4	-0.3683783	0.0361629
b5	-0.2882946	0.4179264
b6	-0.4500895	-0.0432380

Task b

```

Sigma2 <- PostDraws$sigma2Sample
Sigma2Median <- median(sqrt(Sigma2))

```

The posterior median of the standard deviation σ is approximately 0.64.

Task c

```

Effect_B <- Betas[,2] + Betas[,6]
Effect_C <- Betas[,2] + Betas[,7]

Diff <- Effect_B - Effect_C

DensDiff <- density(Diff)

df_DensDiff <- data.frame("x" = DensDiff$x, "y" = DensDiff$y)

DiffInterval <- quantile(Diff, probs=c(0.025, 0.975))

df_DiffInterval <- data.frame(lower_bound = DiffInterval[1], upper_bound = DiffInterval[2])
colnames(df_DiffInterval) <- c("lower bound", "upper bound")
rownames(df_DiffInterval) <- c("95% Equal Tail Credible Interval")
knitr::kable(df_DiffInterval)

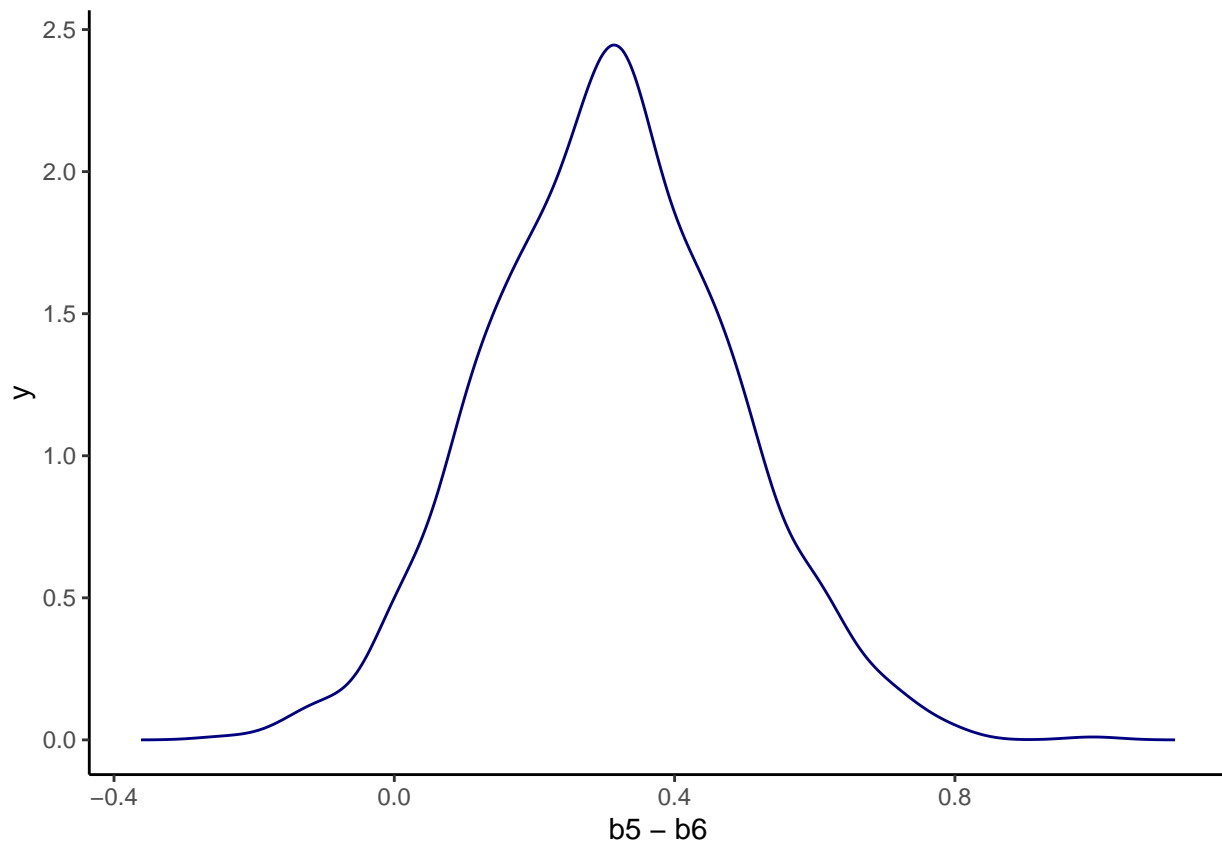
```

	lower bound	upper bound
95% Equal Tail Credible Interval	-0.0108499	0.6558339

```

ggplot(df_DensDiff) +
  geom_line(aes(x=x, y=y), color = "navy") +
  xlab("b5 - b6") +
  theme_classic()

```



There is substantial probability mass that the effect on y from x_1 is larger in high school B compared to high school C. However, the 95 % equal tail credible interval for the difference of the slopes of x_1 between the high schools reveals that the difference can be either negative or positive. Hence, the probability is not that high that this effect in high school B is larger than in high school C.

Task d

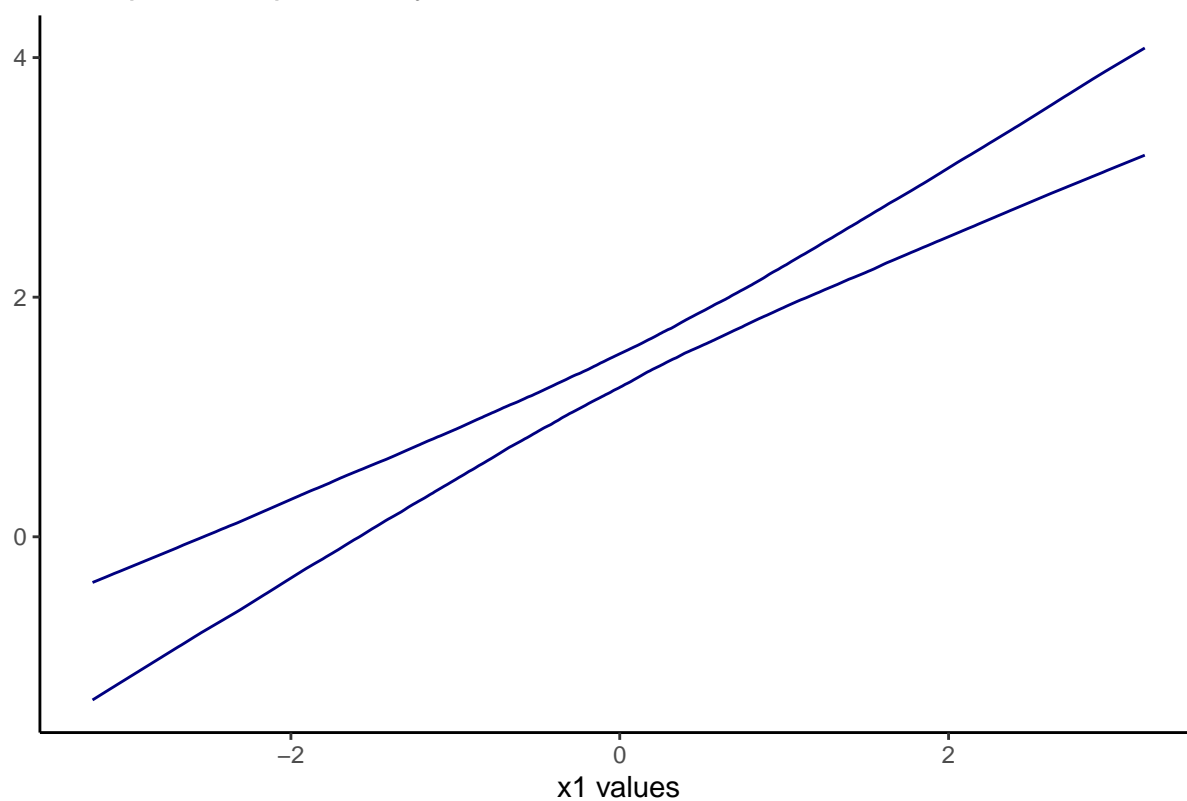
```
x1_grid <- seq(min(X[,2]), max(X[,2]), 0.01)
ExpectedMu <- matrix(NA, nrow = length(x1_grid), ncol = 2)

for (i in 1:length(x1_grid)){
  Mu <- Betas[,1] + Betas[,2]*x1_grid[i] + Betas[,3]*0.5
  ExpectedMu[i,] <- quantile(Mu, probs = c(0.05,0.95))
}

df_ExpectedMu <- data.frame("x1"= x1_grid,
                           "lower" = ExpectedMu[,1],
                           "upper" = ExpectedMu[,2])

ggplot(df_ExpectedMu) +
  geom_line(aes(x=x1, y=lower),color = "navy") +
  geom_line(aes(x=x1, y=upper),color = "navy") +
  ggtitle("90 % posterior probability intervals as a function of x1") +
  xlab("x1 values") +
  ylab("") +
  theme_classic()
```

90 % posterior probability intervals as a function of x1



Task e

```
Mu <- Betas[,1] + Betas[,2]*0.4 + Betas[,3] + Betas[,4] + Betas[,6]
N <- 1000

y <- rnorm(N, mean = Mu, sd = sqrt(Sigma2))

yDens <- density(y)

df_y <- data.frame("x"=yDens$x, "y"=yDens$y)

ggplot(df_y) +
  geom_line(aes(x=x,y=y), color = "navy") +
  ggtitle("Posterior Predictive Distribution Of y") +
  xlab("y") +
  ylab("Density") +
  theme_classic()
```