# Bayesian Learning Lecture 5 - Regression and Regularization

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#### Lecture overview

- The linear regression model
- Non-linear regression
- Regularization priors

## Linear regression

The linear regression model in matrix form

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \boldsymbol{\varepsilon}_{(\mathbf{n} \times \mathbf{1})} + \boldsymbol{\varepsilon}_{(\mathbf{n} \times \mathbf{1})}$$

$$y = \begin{pmatrix} y_1 \\ \vdots \\ y_n \end{pmatrix}, \beta = \begin{pmatrix} \beta_1 \\ \vdots \\ \beta_k \end{pmatrix}, \varepsilon = \begin{pmatrix} \varepsilon_1 \\ \vdots \\ \varepsilon_n \end{pmatrix}$$

$$X = \begin{pmatrix} x'_1 \\ \vdots \\ x'_n \end{pmatrix} = \begin{pmatrix} x_{11} & \cdots & x_{1k} \\ \vdots & \ddots & \vdots \\ x_{n1} & \cdots & x_{nk} \end{pmatrix}$$

- Usually  $x_{i1} = 1$ , for all i.  $\beta_1$  is the intercept.
- Likelihood

$$y|\beta, \sigma^2, X \sim N(X\beta, \sigma^2 I_n)$$

# Linear regression - uniform prior

Standard non-informative prior: uniform on  $(β, log σ^2)$ 

$$p(\beta, \sigma^2) \propto \sigma^{-2}$$

**Joint posterior** of  $\beta$  and  $\sigma^2$ :

$$\beta | \sigma^2, y \sim N [\hat{\beta}, \sigma^2(X'X)^{-1}]$$
  
 $\sigma^2 | y \sim Inv \cdot \chi^2(n - k, s^2)$ 

where  $\hat{\beta}=(\mathsf{X}'\mathsf{X})^{-1}\mathsf{X}'\mathsf{y}$  and  $s^2=\frac{1}{n-k}(\mathsf{y}-\mathsf{X}\hat{\beta})'(\mathsf{y}-\mathsf{X}\hat{\beta}).$ 

- Simulate from the joint posterior by simulating from
  - $ightharpoonup p(\sigma^2|y)$
  - $\triangleright p(\beta|\sigma^2,y)$
- **Marginal posterior of**  $\beta$  :

$$\beta | \mathbf{y} \sim t_{n-k} \left[ \hat{\beta}, s^2 (X'X)^{-1} \right]$$

# Linear regression - conjugate prior

**Joint prior** for  $\beta$  and  $\sigma^2$ 

$$\begin{split} \beta | \sigma^2 &\sim \textit{N}\left(\mu_0, \sigma^2 \Omega_0^{-1}\right) \\ \sigma^2 &\sim \textit{Inv} - \chi^2\left(\nu_0, \sigma_0^2\right) \end{split}$$

Posterior

$$\begin{split} \beta | \sigma^2, \mathbf{y} &\sim \textit{N}\left[\mu_{\textit{n}}, \sigma^2 \Omega_{\textit{n}}^{-1}\right] \\ \sigma^2 | \mathbf{y} &\sim \textit{Inv} - \chi^2\left(\nu_{\textit{n}}, \sigma_{\textit{n}}^2\right) \end{split}$$

$$\mu_{n} = \left(\mathsf{X}'\mathsf{X} + \Omega_{0}\right)^{-1} \left(\mathsf{X}'\mathsf{X}\hat{\beta} + \Omega_{0}\mu_{0}\right)$$

$$\Omega_{n} = \mathsf{X}'\mathsf{X} + \Omega_{0}$$

$$\nu_{n} = \nu_{0} + n$$

$$\nu_{n}\sigma_{n}^{2} = \nu_{0}\sigma_{0}^{2} + \left(\mathsf{y}'\mathsf{y} + \mu_{0}'\Omega_{0}\mu_{0} - \mu_{n}'\Omega_{n}\mu_{n}\right)$$

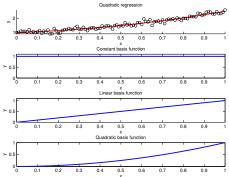
#### Polynomial regression

#### ■ Polynomial regression

$$f(x_i) = \beta_0 + \beta_1 x_i + \beta_2 x_i^2 + \dots + \beta_k x_i^k.$$
  
$$y = X_P \beta + \varepsilon,$$

where

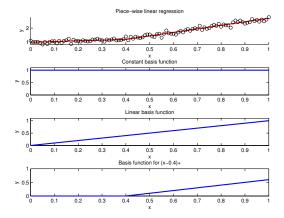
$$X_P = (1, x, x^2, ..., x^k).$$



#### Spline regression

- Polynomials are too global. Need more local basis functions.
- Truncated power splines given knot locations  $k_1, ..., k_m$

$$b_{ij} = \begin{cases} (x_i - k_j)^p & \text{if } x_i > k_j \\ 0 & \text{otherwise} \end{cases}$$



# **Splines**

**Spline regression is linear** in the m 'knot variables'  $b_j$ 

$$y = X_b \beta + \varepsilon$$
,

where  $X_b$  is the basis matrix

$$X_b = (b_1, ..., b_m).$$

Adding intercept and linear term

$$X_b = (1, x, b_1, ..., b_m).$$

# Smoothness prior for splines

- Problem: too many knots leads to over-fitting.
- Smoothness/shrinkage/regularization prior

$$\beta_j | \sigma^2 \stackrel{iid}{\sim} N\left(0, \frac{\sigma^2}{\lambda}\right)$$

- Larger  $\lambda$  gives smoother fit. More shrinkage. Note:  $\mu_0 = 0, \Omega_0 = \lambda I$ .
- Equivalent to penalized likelihood:

$$-2 \cdot \log p(\beta | \sigma^2, y, X) \propto (y - X\beta)'(y - X\beta) + \lambda \beta' \beta$$

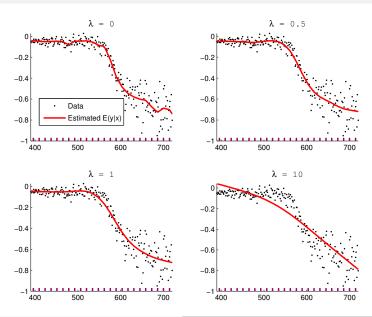
Posterior mean/mode gives ridge regression estimator

$$\tilde{\beta} = \left( \mathsf{X}'\mathsf{X} + \lambda \mathsf{I} \right)^{-1} \mathsf{X}'\mathsf{y}$$

■ When X'X = I (orthogonal, "uncorrelated" features)

$$\tilde{\beta} = \frac{1}{1+\lambda}\hat{\beta}$$

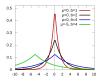
# Bayesian spline with smoothness prior



## Smoothness prior for splines

Lasso is equivalent to posterior mode under Laplace prior

$$\beta_i | \sigma^2 \stackrel{iid}{\sim} \text{Laplace} \left( 0, \frac{\sigma^2}{\lambda} \right)$$



- The Bayesian shrinkage prior is interpretable. Not ad hoc.
- Laplace prior:
  - tails in distribution die off slowly
  - ▶ many  $\beta_i$  close to zero, but some  $\beta_i$  very large.
- Normal prior:
  - tails in distribution die off rapidly
  - $\triangleright$  all  $\beta_i$ 's are similar in magnitude.

# Estimating the shrinkage

- lacksquare Cross-validation: determine  $\lambda$  by performance on test data.
- Bayesian:  $\lambda$  is unknown  $\Rightarrow$  use a prior for  $\lambda$ .
- $\lambda \sim Inv \chi^2(\eta_0, \lambda_0).$
- Hierarchical setup:

$$\begin{aligned} \textbf{y}|\beta,\sigma^2,\textbf{X} &\sim \textit{N}(\textbf{X}\beta,\sigma^2\textit{I}_n) \\ \beta|\sigma^2,\lambda &\sim \textit{N}\left(0,\sigma^2\lambda^{-1}\textit{I}_m\right) \\ \sigma^2 &\sim \textit{Inv} - \chi^2(\nu_0,\sigma_0^2) \\ \lambda &\sim \textit{Inv} - \chi^2(\eta_0,\lambda_0), \end{aligned}$$

so 
$$\mu_0 = 0$$
,  $\Omega_0 = \lambda I_m$ .

# Regression with estimated shrinkage

lacksquare The joint posterior of eta,  $\sigma^2$  and  $\lambda$  is

$$\begin{split} \beta|\sigma^2, \lambda, \mathbf{y} &\sim \textit{N}\left(\mu_n, \sigma^2\Omega_n^{-1}\right) \\ \sigma^2|\lambda, \mathbf{y} &\sim \textit{Inv} - \chi^2\left(\nu_n, \sigma_n^2\right) \\ \rho(\lambda|\mathbf{y}) &\propto \sqrt{\frac{|\Omega_0|}{|\mathbf{X}^T\mathbf{X} + \Omega_0|}} \left(\frac{\nu_n \sigma_n^2}{2}\right)^{-\nu_n/2} \cdot \rho(\lambda) \end{split}$$

where  $\Omega_0 = \lambda I_m$ , and  $p(\lambda)$  is the prior for  $\lambda$ , and

$$\mu_n = \left(X^T X + \Omega_0\right)^{-1} X^T y$$

$$\Omega_n = X^T X + \Omega_0$$

$$\nu_n = \nu_0 + n$$

$$\nu_n \sigma_n^2 = \nu_0 \sigma_0^2 + y^T y - \mu_n^T \Omega_n \mu_n$$

### More complexity

■ The location of the knots can be unknown. Joint posterior:

$$p(\beta, \sigma^2, \lambda, k_1, ..., k_m | y, X)$$

- The basic spline model can be extended with:
  - ► Heteroscedastic errors (also modelled with a spline)
  - ► Non-normal errors (student-t or mixture distributions)
  - Autocorrelated/dependent errors (AR process for the errors)