

Reminder about Confidence Intervals

Marie Delacre<sup>1</sup>

<sup>1</sup> ULB

Author Note

Correspondence concerning this article should be addressed to Marie Delacre, Postal  
address. E-mail: marie.delacre@ulb.ac.be

7

Abstract

8

9       *Keywords:* keywords

10       Word count: X

## Reminder about Confidence Intervals

**Reference**

Cumming, G., & Finch, S. (2001). A primer on the understanding, use, and calculation of confidence intervals that are based on central and noncentral distributions. *Educational and Psychological Measurement*, 61(532).

**Method based on the use of a pivotal quantity**

When computing a (supposed normal) centered variable, divided by the standard error (i.e. an independent variable closely related with the  $\chi^2$  distribution), then computed quantity will follow a central  $t$ -distribution. This quantity is called a pivotal quantity (PQ), i.e. a quantity that is very interesting because its sampling distribution is not a function of the parameter we want to estimate (Cox & Hinkley, 1974 cited by Cumming and Finch, 2001). We can therefore use it, in order to define confidence limits for any parameter.

The method consists in four steps:

- 1) Compute a pivotal quantity (PQ) of the general form: (Estimator - parameter)/SE;
- 2) Determining the distribution of PQ;
- 3) Computing the confidence limits of PQ: determine a range of values, centered around 0, such as  $(1-\alpha)\%$  of the area under the distribution of PQ falls in this range;
- 4) Pivot in order to obtain the confidence interval around the parameter of interest.

As a first example, consider the case of 2 means difference, assuming normality and homoscedasticity. The pivotal quantity is defined as follows:

$$PQ = \frac{(\bar{X}_1 - \bar{X}_2) - (\mu_1 - \mu_2)}{SE} \quad (1)$$

$$\text{With } SE = \sigma_{pooled} \times \sqrt{\frac{1}{n_1} + \frac{1}{n_2}} \text{ and } \sigma_{pooled} = \sqrt{\frac{(n_1-1)*S_1^2 + (n_2-1)*S_2^2}{n_1+n_2-2}}$$

32 This quantity follows a  $t$ -distribution with  $n_1 + n_2 - 2$  degrees of freedom (therefore, it  
 33 depends only on  $n_1$  and  $n_2$ , it does NOT depend on the parameter of interest, i.e.  $\mu_1 - \mu_2$ ).

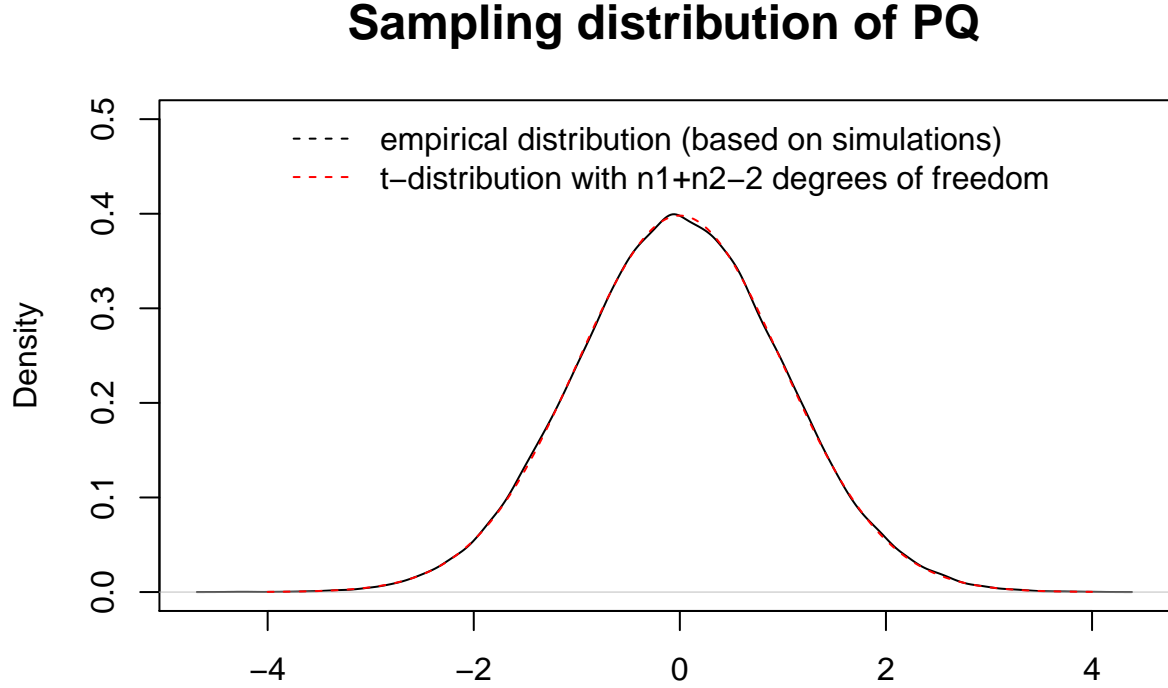


Figure 1. Sampling distribution of the pivotal quantity under the assumptions of normality and homoscedasticity

34 Because the theoretical distribution of PQ is known, one can compute the confidence  
 35 limits, for any confidence level:

$$Pr[t_{n_1+n_2-2}(\frac{\alpha}{2}) < \frac{(\bar{X}_1 - \bar{X}_2) - (\mu_1 - \mu_2)}{SE} < t_{n_1+n_2-2}(1 - \frac{\alpha}{2})] = 1 - \alpha \quad (2)$$

36 Because the  $t$ -distribution is symmetrically centered around 0, one can deduce that  
 37  $t_{n_1+n_2-2}(\frac{\alpha}{2}) = -t_{n_1+n_2-2}(1 - \frac{\alpha}{2})$ , and therefore:

$$Pr[-t_{n_1+n_2-2}(1 - \frac{\alpha}{2}) < \frac{(\bar{X}_1 - \bar{X}_2) - (\mu_1 - \mu_2)}{SE} < t_{n_1+n_2-2}(1 - \frac{\alpha}{2})] = 1 - \alpha \quad (3)$$

38

In pivoting the inequation, one can deduce that:

$$Pr[-t_{n_1+n_2-2}(1 - \frac{\alpha}{2}) \times SE < (\bar{X}_1 - \bar{X}_2) - (\mu_1 - \mu_2) < t_{n_1+n_2-2}(1 - \frac{\alpha}{2}) \times SE] = 1 - \alpha \quad (4)$$

$$\Leftrightarrow Pr[-(\bar{X}_1 - \bar{X}_2) - t_{n_1+n_2-2}(1 - \frac{\alpha}{2}) \times SE < -(\mu_1 - \mu_2) < -(\bar{X}_1 - \bar{X}_2) + t_{n_1+n_2-2}(1 - \frac{\alpha}{2}) \times SE] = 1 - \alpha \quad (5)$$

$$\Leftrightarrow Pr[(\bar{X}_1 - \bar{X}_2) + t_{n_1+n_2-2}(1 - \frac{\alpha}{2}) \times SE > \mu_1 - \mu_2 > (\bar{X}_1 - \bar{X}_2) - t_{n_1+n_2-2}(1 - \frac{\alpha}{2}) \times SE] = 1 - \alpha \quad (6)$$

$$\Leftrightarrow Pr[(\bar{X}_1 - \bar{X}_2) - t_{n_1+n_2-2}(1 - \frac{\alpha}{2}) \times SE < \mu_1 - \mu_2 < (\bar{X}_1 - \bar{X}_2) + t_{n_1+n_2-2}(1 - \frac{\alpha}{2}) \times SE] = 1 - \alpha \quad (7)$$

39

As a second example, consider the case of 2 means difference, assuming normality and

40

heteroscedasticity. The pivotal quantity is defined as follows:

$$PQ = \frac{(\bar{X}_1 - \bar{X}_2) - (\mu_1 - \mu_2)}{SE} \quad (8)$$

41

With  $SE = \sqrt{\frac{S_1^2}{n_1} + \frac{S_2^2}{n_2}}$

42 This quantity follows a  $t$ - distribution with  $\frac{(\frac{S_1^2}{n_1} + \frac{S_2^2}{n_2})^2}{\frac{S_1^2}{n_1-1} + \frac{S_2^2}{n_2-1}}$  degrees of freedom (therefore, it  
 43 depends on  $n_1$  and  $n_2$ ,  $S_1$  and  $S_2$ , and does NOT depend on the parameter of interest,  
 44 i.e.  $\mu_1 - \mu_2$ ).

## Sampling distribution of PQ

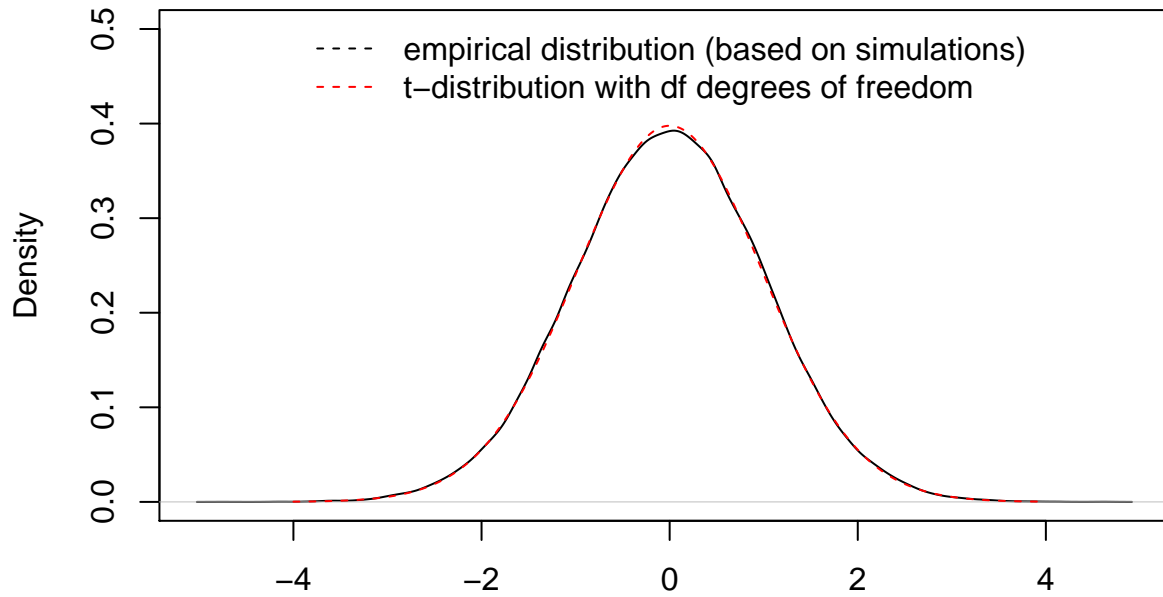


Figure 2. Sampling distribution of the pivotal quantity under the assumptions of normality and heteroscedasticity

45 Because the theoretical distribution of PQ is known, one can compute the confidence  
 46 limits, for any confidence level (see the first example for more details):

$$Pr[(\bar{X}_1 - \bar{X}_2) - t_{n_1+n_2-2}(1-\frac{\alpha}{2}) \times SE < \mu_1 - \mu_2 < (\bar{X}_1 - \bar{X}_2) + t_{n_1+n_2-2}(1-\frac{\alpha}{2}) \times SE] = 1 - \alpha \quad (9)$$

47 With  $SE = \sqrt{\frac{S_1^2}{n_1} + \frac{S_2^2}{n_2}}$