

Cox Model Building and Diagnostics

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Model building

Load the data

```
library(tidyverse)

## -- Attaching packages ----- tidyverse 1.3.2 --
## v ggplot2 3.4.0      v purrr   0.3.5
## v tibble  3.1.8      v dplyr  1.0.10
## v tidyr   1.2.1      v stringr 1.4.1
## v readr   2.1.3      v forcats 0.5.2
## -- Conflicts ----- tidyverse_conflicts() --
## x dplyr::filter() masks stats::filter()
## x dplyr::lag()    masks stats::lag()

library(survival)
library(asauro)
library(broom)

dat <- pharmacoSmoking
```

The 4 candidate models

```
M0 <- coxph(Surv(ttr, relapse) ~ 1, data = dat)
MA  <- coxph(Surv(ttr, relapse) ~ ageGroup4, data = dat)
MB  <- coxph(Surv(ttr, relapse) ~ employment, data = dat)
MC  <- coxph(Surv(ttr, relapse) ~ ageGroup4 + employment, data = dat)
```

```
summary(MA)
```

```
## Call:
## coxph(formula = Surv(ttr, relapse) ~ ageGroup4, data = dat)
##
##    n= 125, number of events= 89
##
##              coef exp(coef) se(coef)      z Pr(>|z|)
## ageGroup435-49  0.0293    1.0297  0.3093  0.095  0.9245
## ageGroup450-64 -0.7914    0.4532  0.3361 -2.355  0.0185 *
## ageGroup465+   -0.3173    0.7281  0.4435 -0.715  0.4744
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
##              exp(coef) exp(-coef) lower .95 upper .95
## ageGroup435-49    1.0297    0.9711    0.5616    1.8880
## ageGroup450-64    0.4532    2.2066    0.2345    0.8757
## ageGroup465+     0.7281    1.3734    0.3053    1.7367
##
## Concordance= 0.593 (se = 0.032 )
```

```
## Likelihood ratio test= 12.22 on 3 df, p=0.007
## Wald test = 11.36 on 3 df, p=0.01
## Score (logrank) test = 11.93 on 3 df, p=0.008
```

```
summary(MB)
```

```
## Call:
## coxph(formula = Surv(ttr, relapse) ~ employment, data = dat)
##
## n= 125, number of events= 89
##
##               coef exp(coef) se(coef)      z Pr(>|z|)
## employmentother 0.1982    1.2192  0.2371 0.836   0.403
## employmentpt    0.4500    1.5683  0.3229 1.394   0.163
##
##               exp(coef) exp(-coef) lower .95 upper .95
## employmentother    1.219    0.8202  0.7661    1.940
## employmentpt      1.568    0.6376  0.8328    2.953
##
## Concordance= 0.541 (se = 0.028 )
## Likelihood ratio test= 2.06 on 2 df, p=0.4
## Wald test = 2.17 on 2 df, p=0.3
## Score (logrank) test = 2.2 on 2 df, p=0.3
```

```
summary(MC)
```

```
## Call:
## coxph(formula = Surv(ttr, relapse) ~ ageGroup4 + employment,
##       data = dat)
##
## n= 125, number of events= 89
##
##               coef exp(coef) se(coef)      z Pr(>|z|)
## ageGroup435-49 -0.1299    0.8782  0.3213 -0.404  0.68594
## ageGroup450-64 -1.0239    0.3592  0.3585 -2.856  0.00429 **
## ageGroup465+   -0.7825    0.4573  0.5046 -1.551  0.12102
## employmentother 0.5257    1.6917  0.2748  1.913  0.05577 .
## employmentpt    0.5001    1.6489  0.3315  1.508  0.13143
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
##               exp(coef) exp(-coef) lower .95 upper .95
## ageGroup435-49    0.8782    1.1387  0.4678    1.6484
## ageGroup450-64    0.3592    2.7839  0.1779    0.7253
## ageGroup465+      0.4573    2.1868  0.1701    1.2295
## employmentother    1.6917    0.5911  0.9871    2.8992
## employmentpt      1.6489    0.6065  0.8610    3.1578
##
## Concordance= 0.617 (se = 0.033 )
## Likelihood ratio test= 16.79 on 5 df, p=0.005
## Wald test = 15.77 on 5 df, p=0.008
## Score (logrank) test = 16.35 on 5 df, p=0.006
```

Comparing nested models: LRT

```
anova(MA, MC)
```

```
## Analysis of Deviance Table
## Cox model: response is Surv(ttr, relapse)
## Model 1: ~ ageGroup4
## Model 2: ~ ageGroup4 + employment
##      loglik  Chisq Df Pr(>|Chi|)
## 1 -380.04
## 2 -377.76 4.5666 2      0.1019
```

Comparing non-nested models: AIC

```
fits <- list(MA = MA, MB = MB, MC = MC)
sapply(fits, AIC)
```

```
##      MA      MB      MC
## 766.0860 774.2464 765.5194
```

Model 'MC' is the best.

Automatic model selection based on AIC

```
Mfull <- coxph(Surv(ttr, relapse) ~ grp + gender + race +
               employment + yearsSmoking + levelSmoking +
               age + I(age^2) + priorAttempts + longestNoSmoke,
               data = dat)
```

```
MAIC <- step(Mfull)
```

```
## Start: AIC=769.66
## Surv(ttr, relapse) ~ grp + gender + race + employment + yearsSmoking +
##      levelSmoking + age + I(age^2) + priorAttempts + longestNoSmoke
##
##              Df      AIC
## - race        3 767.42
## - gender       1 767.68
## - priorAttempts 1 767.87
## - levelSmoking  1 767.92
## - longestNoSmoke 1 767.99
## - yearsSmoking  1 768.31
## - I(age^2)      1 769.17
## <none>          769.66
## - age          1 771.40
## - employment   2 773.89
## - grp          1 775.27
```

```

##
## Step: AIC=767.42
## Surv(ttr, relapse) ~ grp + gender + employment + yearsSmoking +
##   levelSmoking + age + I(age^2) + priorAttempts + longestNoSmoke
##
##           Df    AIC
## - gender      1 765.42
## - levelSmoking 1 765.43
## - priorAttempts 1 765.59
## - yearsSmoking 1 765.76
## - longestNoSmoke 1 766.28
## - I(age^2)      1 766.63
## <none>          767.42
## - age          1 768.38
## - employment   2 770.69
## - grp          1 773.24
##
## Step: AIC=765.42
## Surv(ttr, relapse) ~ grp + employment + yearsSmoking + levelSmoking +
##   age + I(age^2) + priorAttempts + longestNoSmoke
##
##           Df    AIC
## - levelSmoking 1 763.43
## - priorAttempts 1 763.59
## - yearsSmoking 1 763.76
## - longestNoSmoke 1 764.29
## - I(age^2)      1 764.63
## <none>          765.42
## - age          1 766.42
## - employment   2 768.88
## - grp          1 771.29
##
## Step: AIC=763.43
## Surv(ttr, relapse) ~ grp + employment + yearsSmoking + age +
##   I(age^2) + priorAttempts + longestNoSmoke
##
##           Df    AIC
## - priorAttempts 1 761.60
## - yearsSmoking 1 761.78
## - longestNoSmoke 1 762.29
## - I(age^2)      1 762.75
## <none>          763.43
## - age          1 764.57
## - employment   2 766.88
## - grp          1 769.31
##
## Step: AIC=761.6
## Surv(ttr, relapse) ~ grp + employment + yearsSmoking + age +
##   I(age^2) + longestNoSmoke
##
##           Df    AIC
## - yearsSmoking 1 759.98
## - longestNoSmoke 1 760.46
## - I(age^2)      1 760.86

```

```

## <none>                761.60
## - age                 1 762.67
## - employment         2 764.95
## - grp                 1 767.34
##
## Step: AIC=759.98
## Surv(ttr, relapse) ~ grp + employment + age + I(age^2) + longestNoSmoke
##
##           Df      AIC
## - longestNoSmoke 1 758.95
## - I(age^2)        1 759.25
## <none>            759.98
## - age             1 760.72
## - employment      2 763.30
## - grp              1 766.07
##
## Step: AIC=758.95
## Surv(ttr, relapse) ~ grp + employment + age + I(age^2)
##
##           Df      AIC
## - I(age^2)       1 758.28
## <none>           758.95
## - age            1 759.89
## - employment     2 762.73
## - grp            1 765.15
##
## Step: AIC=758.28
## Surv(ttr, relapse) ~ grp + employment + age
##
##           Df      AIC
## <none>         758.28
## - employment  2 762.48
## - grp         1 764.18
## - age         1 767.24

```

```
summary(MAIC)
```

```

## Call:
## coxph(formula = Surv(ttr, relapse) ~ grp + employment + age,
##       data = dat)
##
## n= 125, number of events= 89
##
##           coef exp(coef) se(coef)      z Pr(>|z|)
## grppatchOnly  0.60788   1.83654  0.21837  2.784  0.00537 **
## employmentother 0.70348   2.02077  0.26929  2.612  0.00899 **
## employmentpt   0.65369   1.92262  0.32732  1.997  0.04581 *
## age           -0.03529   0.96533  0.01075 -3.282  0.00103 **
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
##           exp(coef) exp(-coef) lower .95 upper .95
## grppatchOnly    1.8365    0.5445    1.1971    2.8176
## employmentother  2.0208    0.4949    1.1920    3.4256

```

```
## employmentpt      1.9226      0.5201      1.0122      3.6518
## age                0.9653      1.0359      0.9452      0.9859
##
## Concordance= 0.638 (se = 0.03 )
## Likelihood ratio test= 22.03 on 4 df, p=2e-04
## Wald test           = 21.91 on 4 df, p=2e-04
## Score (logrank) test = 22.48 on 4 df, p=2e-04
```

There's only so much we can automate: transformed variables

```
M_ageLinear <- MAIC
M_ageCat <- coxph(Surv(ttr, relapse) ~ grp + employment + ageGroup4, data = dat)
AIC(M_ageLinear)
```

```
## [1] 758.2785
```

```
AIC(M_ageCat)
```

```
## [1] 758.4157
```

AIC prefers a linear term of Age. However, it might not necessarily be the best choice if the age effect is *not* actually linear, and we want to e.g. explicitly account for ethnicity:

```
M_new <- coxph(Surv(ttr, relapse) ~ grp + race +
               employment +
               ageGroup4,
               data = dat)
summary(M_new)
```

```
## Call:
## coxph(formula = Surv(ttr, relapse) ~ grp + race + employment +
##       ageGroup4, data = dat)
##
##      n= 125, number of events= 89
##
##              coef exp(coef) se(coef)      z Pr(>|z|)
## grppatchOnly    0.6497    1.9149   0.2190   2.967 0.00301 **
## racehispanic   -0.4010    0.6696   0.4984  -0.805 0.42099
## raceother      -1.1671    0.3113   1.0313  -1.132 0.25780
## racewhite      -0.2893    0.7488   0.2352  -1.230 0.21865
## employmentother 0.6914    1.9966   0.2777   2.490 0.01277 *
## employmentpt    0.5509    1.7348   0.3343   1.648 0.09943 .
## ageGroup435-49  -0.1579    0.8539   0.3273  -0.482 0.62951
## ageGroup450-64  -1.0792    0.3399   0.3623  -2.978 0.00290 **
## ageGroup465+    -0.8313    0.4355   0.5029  -1.653 0.09830 .
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
##              exp(coef) exp(-coef) lower .95 upper .95
## grppatchOnly    1.9149    0.5222    1.24664    2.9414
```

```
## racehispanic      0.6696      1.4934      0.25213      1.7785
## raceother         0.3113      3.2126      0.04124      2.3497
## racewhite         0.7488      1.3355      0.47220      1.1873
## employmentother   1.9966      0.5009      1.15857      3.4407
## employmentpt      1.7348      0.5764      0.90084      3.3407
## ageGroup435-49     0.8539      1.1711      0.44958      1.6219
## ageGroup450-64     0.3399      2.9423      0.16707      0.6914
## ageGroup465+       0.4355      2.2964      0.16252      1.1668
##
## Concordance= 0.667 (se = 0.032 )
## Likelihood ratio test= 28.78 on 9 df, p=7e-04
## Wald test           = 27.04 on 9 df, p=0.001
## Score (logrank) test = 27.93 on 9 df, p=0.001
```

```
AIC(MAIC)
```

```
## [1] 758.2785
```

```
AIC(M_new)
```

```
## [1] 761.5297
```

Measuring predictions performance - C statistic

Split data into a training and a testing set

```
set.seed(1234)
i.training <- sample.int(nrow(dat), size = ceiling(nrow(dat)/2), replace = FALSE)
i.testing <- setdiff(seq_len(nrow(dat)), i.training)
d_training <- dat[i.training, ]
d_testing <- dat[i.testing, ]
```

Train 3 candidate models

```
MA <- coxph(Surv(ttr, relapse) ~ gender + age, data = d_training)
MB <- coxph(Surv(ttr, relapse) ~ gender + employment, data = d_training)
MC <- coxph(Surv(ttr, relapse) ~ gender + age + employment, data = d_training)
```

Measure predictions performance - THE WRONG WAY

WROOOOOONG! THIS WOULD BE INCORRECT:

```
models <- list(A = MA, B = MB, C = MC)
map_dbl(models, ~ summary(.)$concordance[1])
```

```
##           A           B           C
## 0.5489433 0.6315350 0.6359844
```


WROOOOOOOONG!!!!

The same dataset is used for both training and performance evaluation. This will *always* favor the model with more parameters.

Make predictions in the testing dataset

```
d_testing$lp_A <- predict(MA, newdata = d_testing, type = "lp")
d_testing$lp_B <- predict(MB, newdata = d_testing, type = "lp")
d_testing$lp_C <- predict(MC, newdata = d_testing, type = "lp")
```

d_testing

##	id	ttr	relapse	grp	age	gender	race	employment	yearsSmoking
## 1	21	182	0	patchOnly	36	Male	white	ft	26
## 7	16	14	1	patchOnly	66	Male	black	pt	54
## 10	70	0	1	patchOnly	38	Male	black	ft	23
## 11	84	12	1	patchOnly	64	Female	black	other	30
## 12	85	182	0	combination	51	Male	black	ft	35
## 13	25	21	1	patchOnly	37	Female	white	pt	23
## 15	59	170	1	patchOnly	42	Female	white	ft	30
## 18	3	182	0	combination	52	Female	white	other	19
## 19	15	140	1	combination	43	Male	white	ft	27
## 20	32	63	1	combination	34	Female	white	ft	18
## 23	110	110	1	combination	49	Female	white	other	35
## 24	127	182	0	combination	58	Female	white	ft	38
## 25	119	0	1	patchOnly	48	Male	hispanic	other	33
## 27	62	15	1	patchOnly	49	Female	black	pt	35
## 29	112	4	1	patchOnly	33	Male	black	ft	12
## 30	60	56	1	combination	49	Female	black	ft	37
## 31	93	2	1	combination	46	Male	black	ft	25
## 33	130	182	0	combination	46	Female	white	pt	25
## 34	19	56	1	patchOnly	52	Male	black	other	25
## 36	4	14	1	patchOnly	48	Female	white	ft	34
## 37	20	14	1	patchOnly	48	Female	white	ft	39
## 39	26	182	0	patchOnly	58	Male	white	other	41
## 44	8	182	0	combination	70	Female	white	other	52
## 45	12	75	1	combination	62	Female	white	pt	46
## 46	13	30	1	combination	86	Male	white	other	40
## 50	36	182	0	combination	62	Male	white	ft	35
## 52	44	8	1	combination	40	Male	white	ft	27
## 53	61	140	1	combination	49	Male	white	other	14
## 55	69	63	1	combination	47	Female	white	pt	15
## 58	106	50	1	combination	29	Male	white	ft	12
## 59	114	14	1	combination	64	Female	white	pt	45
## 60	120	0	1	combination	52	Male	white	ft	38
## 61	40	84	1	patchOnly	38	Female	hispanic	ft	10
## 64	123	182	0	combination	63	Female	hispanic	other	45
## 65	7	182	0	patchOnly	58	Male	black	ft	40
## 68	52	182	0	patchOnly	34	Female	black	ft	20
## 69	86	0	1	patchOnly	49	Female	black	ft	36
## 73	75	12	1	combination	46	Female	black	other	30

##	75	1	49	1	patchOnly	53	Female	white	pt	18
##	76	6	182	0	patchOnly	58	Male	white	ft	46
##	77	11	182	0	patchOnly	40	Female	white	ft	25
##	81	56	182	0	patchOnly	64	Female	white	other	47
##	82	72	0	1	patchOnly	50	Female	white	ft	30
##	83	78	28	1	patchOnly	47	Male	white	ft	22
##	85	83	2	1	patchOnly	51	Male	white	other	36
##	88	95	1	1	patchOnly	22	Female	white	pt	9
##	89	99	140	1	patchOnly	34	Female	white	ft	18
##	91	115	28	1	patchOnly	30	Female	white	ft	12
##	92	116	1	1	patchOnly	31	Male	white	ft	19
##	94	126	77	1	patchOnly	56	Female	white	ft	44
##	96	2	182	0	combination	69	Male	white	other	50
##	97	5	182	0	combination	41	Male	white	ft	29
##	102	64	60	1	combination	70	Female	white	other	54
##	104	92	182	0	combination	58	Female	white	ft	36
##	106	103	182	0	combination	72	Male	white	other	55
##	109	96	182	0	combination	63	Female	other	other	52
##	110	46	2	1	combination	40	Female	hispanic	pt	23
##	116	109	30	1	patchOnly	51	Male	black	other	30
##	118	45	2	1	combination	47	Female	black	ft	33
##	121	74	10	1	combination	45	Female	black	ft	32
##	124	118	15	1	combination	56	Female	black	other	39
##	125	128	182	0	combination	50	Female	black	pt	30
##		levelSmoking	ageGroup2	ageGroup4	priorAttempts	longestNoSmoke			lp_A	
##	1	heavy	21-49	35-49	0	0	-0.109308465			
##	7	heavy	50+	65+	0	0	-0.510397624			
##	10	light	21-49	35-49	10	90	-0.136047742			
##	11	heavy	50+	50-64	12	365	-0.216673033			
##	12	heavy	50+	50-64	1	7	-0.309853044			
##	13	light	21-49	35-49	5	1095	0.144307211			
##	15	heavy	21-49	35-49	5	240	0.077459018			
##	18	light	50+	50-64	1	7	-0.056237369			
##	19	heavy	21-49	35-49	5	120	-0.202895935			
##	20	heavy	21-49	21-34	8	90	0.184416127			
##	23	heavy	21-49	35-49	10	60	-0.016128453			
##	24	heavy	50+	50-64	1	0	-0.136455201			
##	25	heavy	21-49	35-49	4	120	-0.269744128			
##	27	light	21-49	35-49	4	540	-0.016128453			
##	29	light	21-49	21-34	1	730	-0.069199549			
##	30	heavy	21-49	35-49	3	2920	-0.016128453			
##	31	light	21-49	35-49	2	1095	-0.243004851			
##	33	heavy	21-49	35-49	3	365	0.023980463			
##	34	light	50+	50-64	10	7	-0.323222683			
##	36	heavy	21-49	35-49	6	2555	-0.002758814			
##	37	heavy	21-49	35-49	1	2	-0.002758814			
##	39	light	50+	50-64	100	180	-0.403440515			
##	44	heavy	50+	65+	3	1	-0.296890864			
##	45	heavy	50+	50-64	8	1095	-0.189933755			
##	46	light	50+	65+	4	2190	-0.777790397			
##	50	heavy	50+	50-64	1	2555	-0.456919069			
##	52	heavy	21-49	35-49	2	90	-0.162787019			
##	53	heavy	21-49	35-49	0	0	-0.283113767			
##	55	heavy	21-49	35-49	1	90	0.010610824			

## 58	heavy	21-49	21-34	0	0	-0.015720994
## 59	heavy	50+	50-64	4	60	-0.216673033
## 60	heavy	50+	50-64	4	7	-0.323222683
## 61	heavy	21-49	35-49	2	3	0.130937572
## 64	heavy	50+	50-64	1	28	-0.203303394
## 65	heavy	50+	50-64	0	0	-0.403440515
## 68	heavy	21-49	21-34	5	90	0.184416127
## 69	heavy	21-49	35-49	1	6205	-0.016128453
## 73	light	21-49	35-49	2	55	0.023980463
## 75	heavy	50+	50-64	6	3650	-0.069607008
## 76	heavy	50+	50-64	10	14	-0.403440515
## 77	heavy	21-49	35-49	1	2920	0.104198295
## 81	heavy	50+	50-64	4	365	-0.216673033
## 82	heavy	50+	50-64	1	90	-0.029498092
## 83	heavy	21-49	35-49	3	4	-0.256374490
## 85	heavy	50+	50-64	1	5	-0.309853044
## 88	heavy	21-49	21-34	2	3	0.344851790
## 89	light	21-49	21-34	2	2190	0.184416127
## 91	light	21-49	21-34	0	0	0.237894681
## 92	heavy	21-49	21-34	10	120	-0.042460271
## 94	heavy	50+	50-64	4	1095	-0.109715923
## 96	heavy	50+	65+	6	5475	-0.550506540
## 97	heavy	21-49	35-49	20	180	-0.176156658
## 102	heavy	50+	65+	1	90	-0.296890864
## 104	light	50+	50-64	2	90	-0.136455201
## 106	light	50+	65+	30	30	-0.590615456
## 109	heavy	50+	50-64	2	180	-0.203303394
## 110	heavy	21-49	35-49	2	3	0.104198295
## 116	heavy	50+	50-64	2	30	-0.309853044
## 118	light	21-49	35-49	4	365	0.010610824
## 121	heavy	21-49	35-49	1	75	0.037350102
## 124	heavy	50+	50-64	3	7	-0.109715923
## 125	heavy	50+	50-64	0	0	-0.029498092
##	lp_B	lp_C				
## 1	-0.1527679	0.260670756				
## 7	0.6226736	0.351007796				
## 10	-0.1527679	0.208171444				
## 11	0.6993553	0.494356461				
## 12	-0.1527679	-0.133074083				
## 13	0.7754415	1.161156338				
## 15	0.0000000	0.152081340				
## 18	0.6993553	0.809352333				
## 19	-0.1527679	0.076923164				
## 20	0.0000000	0.362078587				
## 23	0.6993553	0.888101300				
## 24	0.0000000	-0.267913155				
## 25	0.5465874	0.865442437				
## 27	0.7754415	0.846160467				
## 29	-0.1527679	0.339419724				
## 30	0.0000000	-0.031666252				
## 31	-0.1527679	-0.001825803				
## 33	0.7754415	0.924909435				
## 34	0.5465874	0.760443813				
## 36	0.0000000	-0.005416596				

```
## 37  0.0000000 -0.005416596
## 39  0.5465874  0.602945877
## 44  0.6993553  0.336858526
## 45  0.7754415  0.504914940
## 46  0.5465874 -0.132044489
## 50 -0.1527679 -0.421820298
## 52 -0.1527679  0.155672132
## 53  0.5465874  0.839192781
## 55  0.7754415  0.898659779
## 58 -0.1527679  0.444418348
## 59  0.7754415  0.452415628
## 60 -0.1527679 -0.159323739
## 61  0.0000000  0.257079964
## 64  0.6993553  0.520606117
## 65 -0.1527679 -0.316821675
## 68  0.0000000  0.362078587
## 69  0.0000000 -0.031666252
## 73  0.6993553  0.966850268
## 75  0.7754415  0.741161843
## 76 -0.1527679 -0.316821675
## 77  0.0000000  0.204580652
## 81  0.6993553  0.494356461
## 82  0.0000000 -0.057915908
## 83 -0.1527679 -0.028075459
## 85  0.5465874  0.786693469
## 88  0.7754415  1.554901177
## 89  0.0000000  0.362078587
## 91  0.0000000  0.467077211
## 92 -0.1527679  0.391919036
## 94  0.0000000 -0.215413843
## 96  0.5465874  0.314199662
## 97 -0.1527679  0.129422476
## 102 0.6993553  0.336858526
## 104 0.0000000 -0.267913155
## 106 0.5465874  0.235450694
## 109 0.6993553  0.520606117
## 110 0.7754415  1.082407370
## 116 0.5465874  0.786693469
## 118 0.0000000  0.020833060
## 121 0.0000000  0.073332372
## 124 0.6993553  0.704353709
## 125 0.7754415  0.819910811
```

Assess predictive performance

```
models <- list(
  A = coxph(Surv(ttr, relapse) ~ lp_A, data = d_testing),
  B = coxph(Surv(ttr, relapse) ~ lp_B, data = d_testing),
  C = coxph(Surv(ttr, relapse) ~ lp_C, data = d_testing)
)

summary(models$A)
```

```
## Call:
## coxph(formula = Surv(ttr, relapse) ~ lp_A, data = d_testing)
##
## n= 62, number of events= 42
##
##      coef exp(coef) se(coef)      z Pr(>|z|)
## lp_A 1.3700    3.9354   0.7538 1.818   0.0691 .
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
##      exp(coef) exp(-coef) lower .95 upper .95
## lp_A      3.935      0.2541   0.8982   17.24
##
## Concordance= 0.571 (se = 0.046 )
## Likelihood ratio test= 3.44 on 1 df,  p=0.06
## Wald test              = 3.3 on 1 df,  p=0.07
## Score (logrank) test = 3.32 on 1 df,  p=0.07
```

```
summary(models$B)
```

```
## Call:
## coxph(formula = Surv(ttr, relapse) ~ lp_B, data = d_testing)
##
## n= 62, number of events= 42
##
##      coef exp(coef) se(coef)      z Pr(>|z|)
## lp_B -0.0540    0.9474   0.4062 -0.133   0.894
##
##      exp(coef) exp(-coef) lower .95 upper .95
## lp_B    0.9474      1.055   0.4273   2.101
##
## Concordance= 0.512 (se = 0.049 )
## Likelihood ratio test= 0.02 on 1 df,  p=0.9
## Wald test              = 0.02 on 1 df,  p=0.9
## Score (logrank) test = 0.02 on 1 df,  p=0.9
```

```
summary(models$C)
```

```
## Call:
## coxph(formula = Surv(ttr, relapse) ~ lp_C, data = d_testing)
##
## n= 62, number of events= 42
##
##      coef exp(coef) se(coef)      z Pr(>|z|)
## lp_C 0.6048    1.8308   0.3944 1.533   0.125
##
##      exp(coef) exp(-coef) lower .95 upper .95
## lp_C    1.831      0.5462   0.8451   3.966
##
## Concordance= 0.54 (se = 0.053 )
## Likelihood ratio test= 2.38 on 1 df,  p=0.1
## Wald test              = 2.35 on 1 df,  p=0.1
## Score (logrank) test = 2.37 on 1 df,  p=0.1
```

1. Check that the predictions go into the right direction

```
sapply(models, coef)
```

```
##      A.lp_A      B.lp_B      C.lp_C  
## 1.37000089 -0.05399827 0.60475515
```

2. Extract the C-statistic from the 'summary' output

```
map_dbl(models, ~ summary(.)$concordance[1])
```

```
##      A      B      C  
## 0.5710006 0.5119832 0.5401438
```

3. Flip directions as needed

```
benchmark <-  
  tibble(  
    model = names(models),  
    sign = sapply(models, coef) |> sign(),  
    C_summary = map_dbl(models, ~ summary(.)$concordance[1]),  
    C = ifelse(sign > 0, C_summary, 1 - C_summary)  
  )  
benchmark
```

```
## # A tibble: 3 x 4  
##   model  sign C_summary      C  
##   <chr> <dbl>   <dbl> <dbl>  
## 1 A      1      0.571 0.571  
## 2 B     -1      0.512 0.488  
## 3 C      1      0.540 0.540
```

```
benchmark |> select(model, C)
```

```
## # A tibble: 3 x 2  
##   model      C  
##   <chr> <dbl>  
## 1 A     0.571  
## 2 B     0.488  
## 3 C     0.540
```

Measuring predictions performance - effect size

```
d1 <- d_testing |> select(ttr, relapse, lp_A, lp_B, lp_C)
head(d1)
```

```
##      ttr relapse      lp_A      lp_B      lp_C
## 1  182        0 -0.1093085 -0.1527679  0.2606708
## 7   14        1 -0.5103976  0.6226736  0.3510078
## 10   0        1 -0.1360477 -0.1527679  0.2081714
## 11  12        1 -0.2166730  0.6993553  0.4943565
## 12 182        0 -0.3098530 -0.1527679 -0.1330741
## 13  21        1  0.1443072  0.7754415  1.1611563
```

Raw estimates

```
A <- coxph(Surv(ttr, relapse) ~ lp_A, data = d1) |> tidy()
B <- coxph(Surv(ttr, relapse) ~ lp_B, data = d1) |> tidy()
C <- coxph(Surv(ttr, relapse) ~ lp_C, data = d1) |> tidy()
bind_rows(A, B, C)
```

```
## # A tibble: 3 x 5
##   term estimate std.error statistic p.value
##   <chr>      <dbl>      <dbl>      <dbl>   <dbl>
## 1 lp_A      1.37         0.754        1.82    0.0691
## 2 lp_B     -0.0540        0.406       -0.133   0.894
## 3 lp_C      0.605         0.394        1.53    0.125
```

These log-HRs are not directly comparable.

```
d_testing |> select(lp_A, lp_B, lp_C) |> sapply(sd)
```

```
##      lp_A      lp_B      lp_C
## 0.2155693 0.3863618 0.4227361
```

```
d_testing |> select(lp_A, lp_B, lp_C) |> sapply(IQR)
```

```
##      lp_A      lp_B      lp_C
## 0.2770125 0.6993553 0.7281209
```

Standardized predictions

```
d2 <- mutate(d1,
              ZA = lp_A / sd(lp_A),
              ZB = lp_B / sd(lp_B),
              ZC = lp_C / sd(lp_C))
head(d2)
```

```
##      ttr relapse      lp_A      lp_B      lp_C      ZA      ZB
## 1  182         0 -0.1093085 -0.1527679  0.2606708 -0.5070688 -0.3954011
## 7   14         1 -0.5103976  0.6226736  0.3510078 -2.3676733  1.6116333
## 10  0         1 -0.1360477 -0.1527679  0.2081714 -0.6311091 -0.3954011
## 11 12         1 -0.2166730  0.6993553  0.4943565 -1.0051202  1.8101045
## 12 182        0 -0.3098530 -0.1527679 -0.1330741 -1.4373710 -0.3954011
## 13 21         1  0.1443072  0.7754415  1.1611563  0.6694238  2.0070344
##
##      ZC
## 1  0.6166277
## 7  0.8303238
## 10 0.4924384
## 11 1.1694211
## 12 -0.3147924
## 13 2.7467644
```

```
A <- coxph(Surv(ttr, relapse) ~ ZA, data = d2) |> tidy()
B <- coxph(Surv(ttr, relapse) ~ ZB, data = d2) |> tidy()
C <- coxph(Surv(ttr, relapse) ~ ZC, data = d2) |> tidy()
bind_rows(A, B, C) |>
  transmute(term, estimate, HR = exp(estimate), p.value)
```

```
## # A tibble: 3 x 4
##   term estimate   HR p.value
##   <chr>     <dbl> <dbl>   <dbl>
## 1 ZA       0.295  1.34  0.0691
## 2 ZB      -0.0209 0.979  0.894
## 3 ZC       0.256  1.29  0.125
```

Discretized predictions - based on quantiles

```
d3 <-
  mutate(d2,
    FA = factor((lp_A > median(lp_A)), levels = c(FALSE, TRUE), labels = c("low", "high")),
    FC = factor((lp_C > median(lp_C)), levels = c(FALSE, TRUE), labels = c("low", "high"))
  )
```

```
A <- coxph(Surv(ttr, relapse) ~ FA, data = d3) |> tidy()
C <- coxph(Surv(ttr, relapse) ~ FC, data = d3) |> tidy()
bind_rows(A, C) |>
  transmute(term, estimate, HR = exp(estimate), p.value)
```

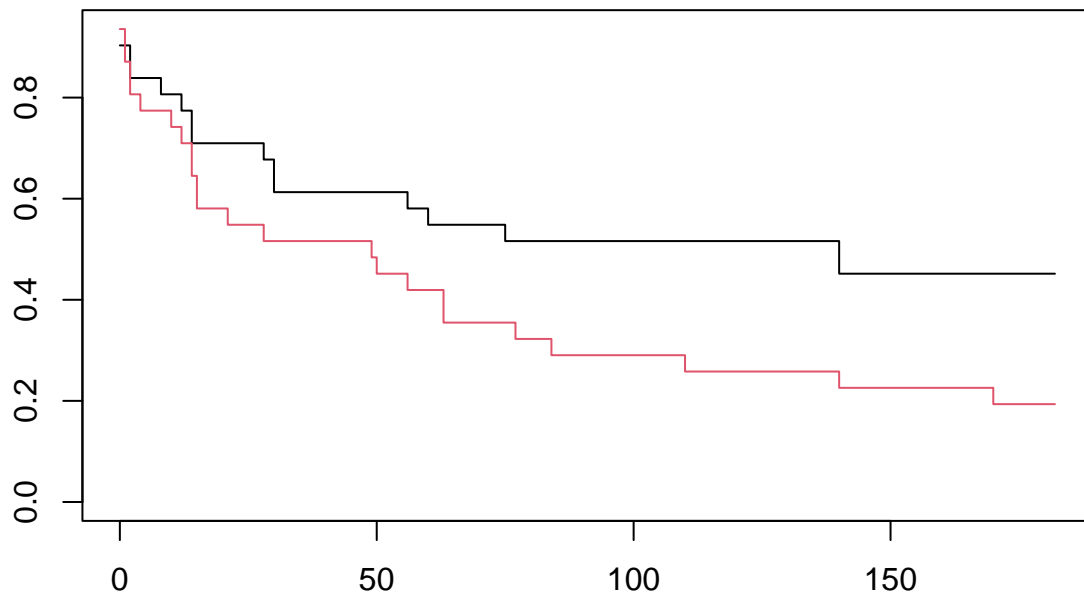
```
## # A tibble: 2 x 4
##   term estimate   HR p.value
##   <chr>     <dbl> <dbl>   <dbl>
## 1 FAhigh  0.599  1.82  0.0583
## 2 FChigh  0.253  1.29  0.416
```

```
fit.KM <- survfit(Surv(ttr, relapse) ~ FA, data = d3)
fit.KM
```



```
## Call: survfit(formula = Surv(ttr, relapse) ~ FA, data = d3)
##
##           n events median 0.95LCL 0.95UCL
## FA=low  31      17   140      30     NA
## FA=high 31      25    49      14    110
```

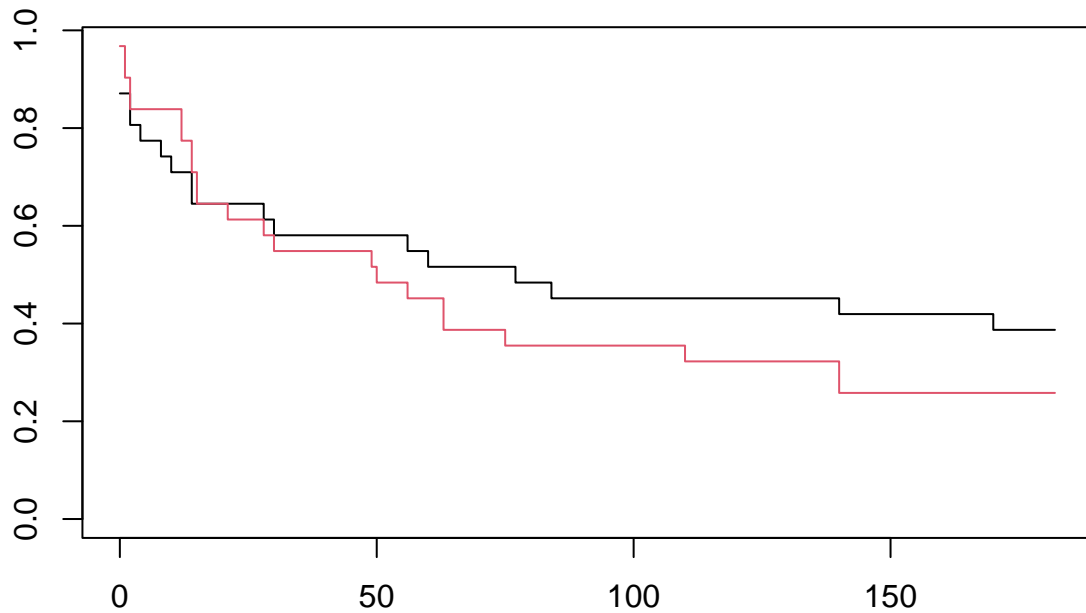
```
plot(fit.KM, col = 1:2)
```



```
fit.KM <- survfit(Surv(ttr, relapse) ~ FC, data = d3)
fit.KM
```

```
## Call: survfit(formula = Surv(ttr, relapse) ~ FC, data = d3)
##
##           n events median 0.95LCL 0.95UCL
## FC=low  31      19    77      14     NA
## FC=high 31      23    50      15    140
```

```
plot(fit.KM, col = 1:2)
```



Model diagnostics

Martingale residuals

```
library(survival)
library(asaur) ## dataset

data(pharmacoSmoking)
dat <- pharmacoSmoking
```

```
fit <- coxph(Surv(ttr, relapse) ~ grp + age + employment, data = dat)
dat$residual <- residuals(fit, type = "martingale")
```

```
par(mfrow = c(1, 3), mar = c(4.2, 2, 2, 2))
with(dat, {

  plot(age, residual)
  lines(lowess(age, residual), lwd = 2)
```

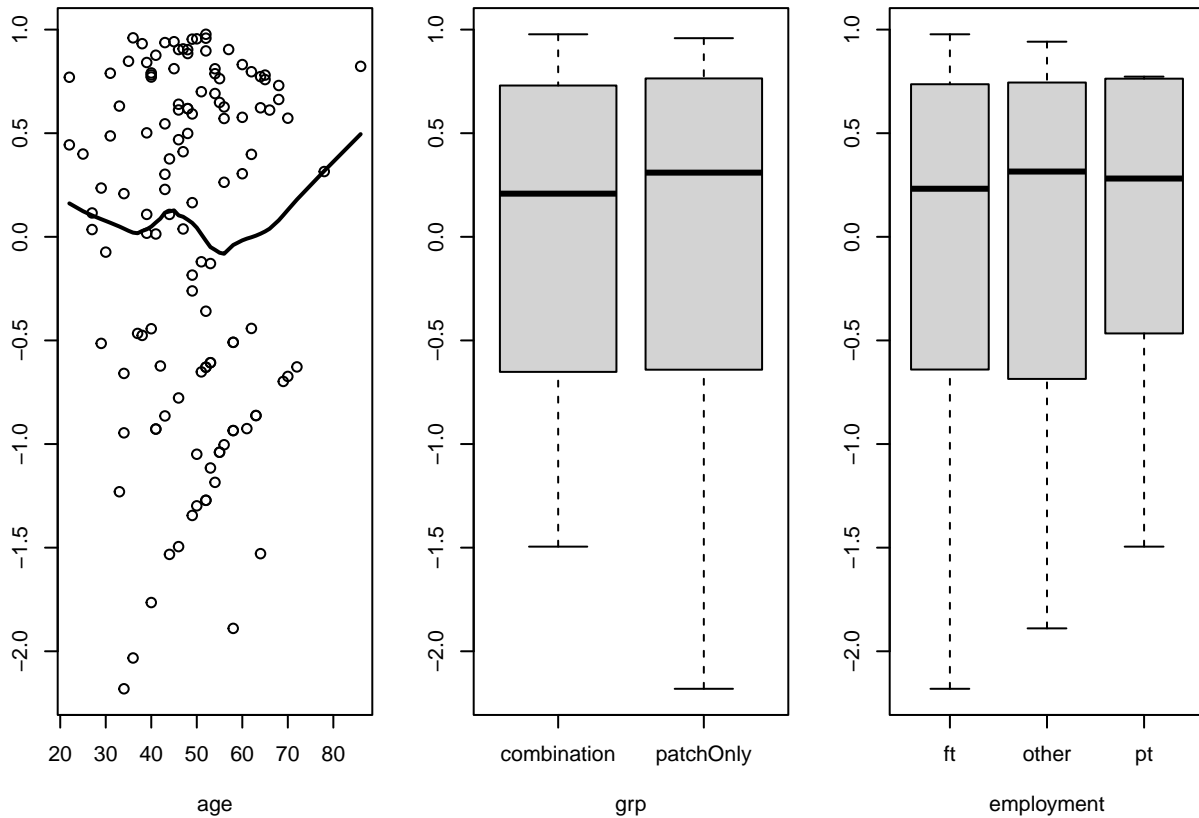
```

plot(residual ~ grp)

plot(residual ~ employment)

})

```



```

fit_better <- coxph(Surv(ttr, relapse) ~ grp + age + I(age^2) + employment, data = dat)
summary(fit_better)

```

```

## Call:
## coxph(formula = Surv(ttr, relapse) ~ grp + age + I(age^2) + employment,
##       data = dat)
##
##      n= 125, number of events= 89
##
##              coef exp(coef)    se(coef)      z Pr(>|z|)
## grppatchOnly   0.6206075  1.8600577  0.2188288  2.836  0.00457 **
## age            -0.1001902  0.9046654  0.0549849 -1.822  0.06843 .
## I(age^2)        0.0006729  1.0006732  0.0005572  1.208  0.22713
## employmentother 0.6800741  1.9740240  0.2754600  2.469  0.01355 *
## employmentpt   0.6757762  1.9655581  0.3278821  2.061  0.03930 *
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
##              exp(coef) exp(-coef) lower .95 upper .95

```

```
## grppatchOnly      1.8601      0.5376      1.2113      2.856
## age               0.9047      1.1054      0.8122      1.008
## I(age^2)          1.0007      0.9993      0.9996      1.002
## employmentother   1.9740      0.5066      1.1505      3.387
## employmentpt      1.9656      0.5088      1.0337      3.737
##
## Concordance= 0.633 (se = 0.031 )
## Likelihood ratio test= 23.36 on 5 df, p=3e-04
## Wald test           = 24.19 on 5 df, p=2e-04
## Score (logrank) test = 24.68 on 5 df, p=2e-04
```

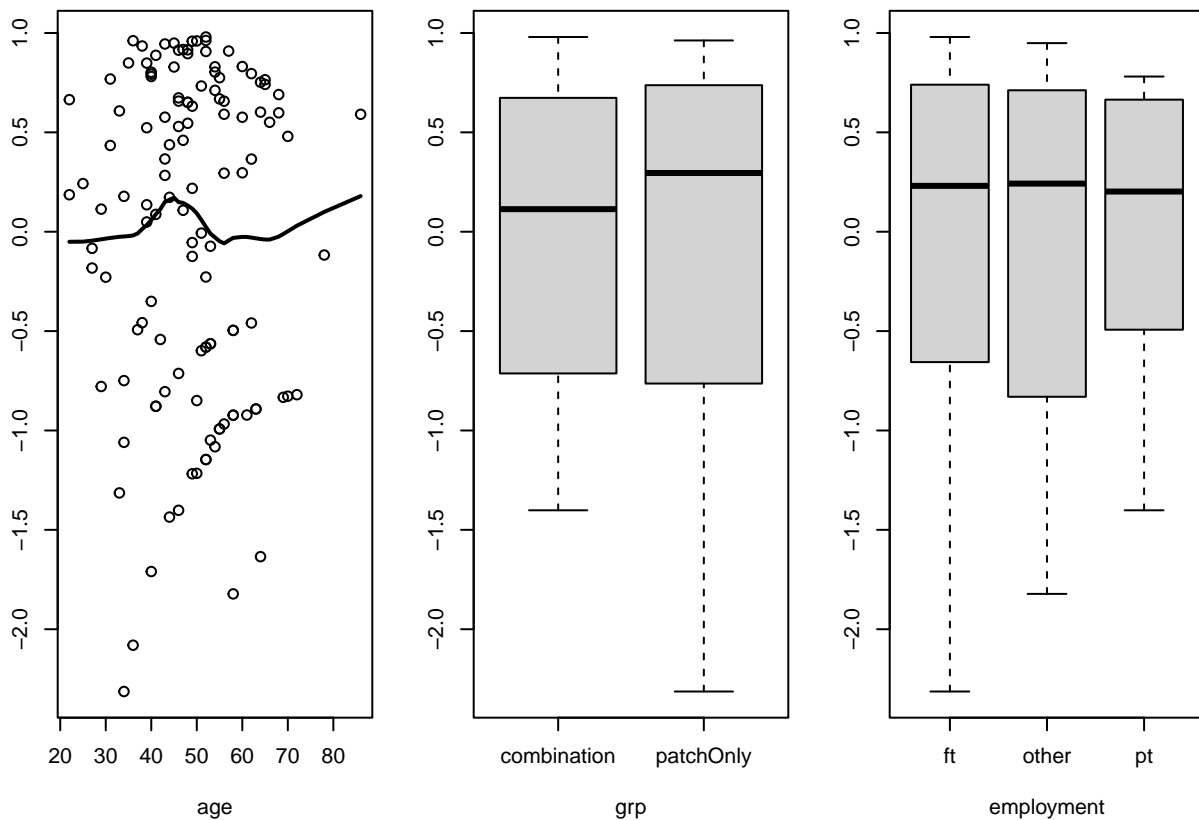
```
dat$residual <- residuals(fit_better, type = "martingale")
par(mfrow = c(1, 3), mar = c(4.2, 2, 2, 2))
with(dat, {

  plot(age, residual)
  lines(lowess(age, residual), lwd = 2)

  plot(residual ~ grp)

  plot(residual ~ employment)

})
```

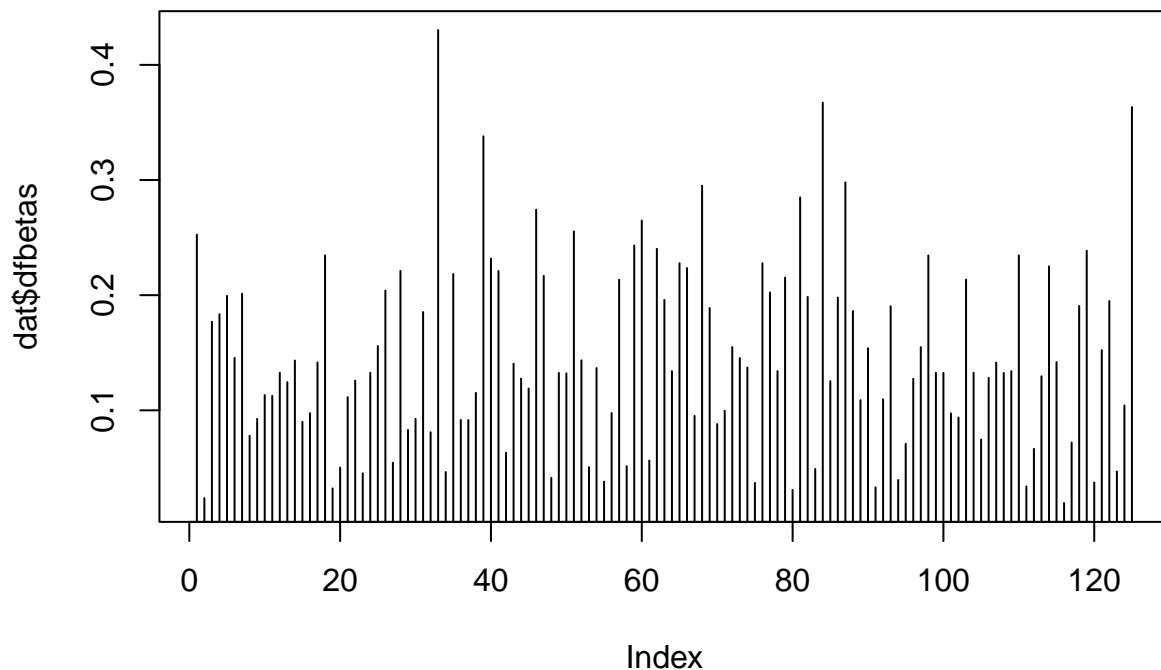


Case-deletion residuals

```
dfbetas <- residuals(fit, type = 'dfbetas')
head(dfbetas)
```

```
##           [,1]      [,2]      [,3]      [,4]
## 1 -0.164823287  0.15809557  0.05955969  0.089952431
## 2  0.006822493 -0.01266160  0.01817193  0.005336660
## 3 -0.050525661 -0.13050951  0.10683305  0.017798812
## 4 -0.101956037  0.08958552 -0.10516817 -0.064573127
## 5 -0.126362316 -0.07386963  0.13516905 -0.008288001
## 6  0.114896122  0.01309695  0.06960045  0.054304287
```

```
dat$dfbetas <- sqrt(rowSums(dfbetas^2))
plot(dat$dfbetas, type = 'h')
abline(h = 0)
```



Proportionality of hazards

Pancreatic cancer dataset - late vs early stage

```

library(survival)
library(asaur) ## dataset
library(plyr)

## -----

## You have loaded plyr after dplyr - this is likely to cause problems.
## If you need functions from both plyr and dplyr, please load plyr first, then dplyr:
## library(plyr); library(dplyr)

## -----

##
## Attaching package: 'plyr'

## The following objects are masked from 'package:dplyr':
##
##   arrange, count, desc, failwith, id, mutate, rename, summarise,
##   summarize

## The following object is masked from 'package:purrr':
##
##   compact

```

```

library(ggplot2)

fmt <- "%m/%d/%Y"
dat <- as.tibble(pancreatic) %>%
  mutate(
    onstudy = as.Date(as.character(onstudy), format = fmt),
    progression = as.Date(as.character(progression), format = fmt),
    death = as.Date(as.character(death), format = fmt),
    OS = death - onstudy,
    PFS = ifelse(is.na(progression), OS, pmin(progression - onstudy, OS))) %>%
  mutate(
    PFS = Surv(as.numeric(PFS / 30.5)),
    OS = Surv(as.numeric(OS / 30.5))
  )

```

```

## Warning: 'as.tibble()' was deprecated in tibble 2.0.0.
## i Please use 'as_tibble()' instead.
## i The signature and semantics have changed, see '?as_tibble'.

```

```
dat
```

```

## # A tibble: 41 x 6
##   stage onstudy   progression death          OS          PFS
##   <fct> <date>      <date>      <date>      <Surv>      <Surv>
## 1 M     2005-12-16 2006-02-02 2006-10-19 10.065574  1.573770
## 2 M     2006-01-06 2006-02-26 2006-04-19  3.377049  1.672131

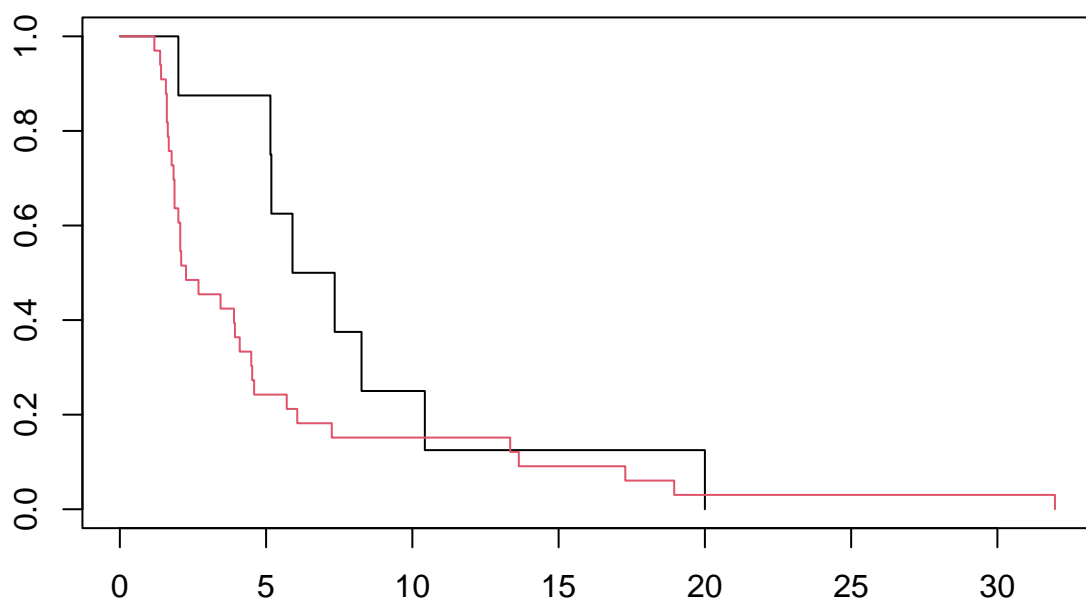
```

```
## 3 LA      2006-02-03 2006-08-02 2007-01-19 11.475410 5.901639
## 4 M      2006-03-30 NA      2006-05-11 1.377049 1.377049
## 5 LA      2006-04-27 2007-03-11 2007-05-29 13.016393 10.426230
## 6 M      2006-05-07 2006-06-25 2006-10-11 5.147541 1.606557
## 7 LA      2006-08-20 NA      2007-01-24 5.147541 5.147541
## 8 M      2007-01-22 2007-03-20 2007-04-14 2.688525 1.868852
## 9 LA      2007-03-02 NA      2008-11-01 20.000000 20.000000
## 10 M     2007-03-27 NA      2007-05-15 1.606557 1.606557
## # ... with 31 more rows
```

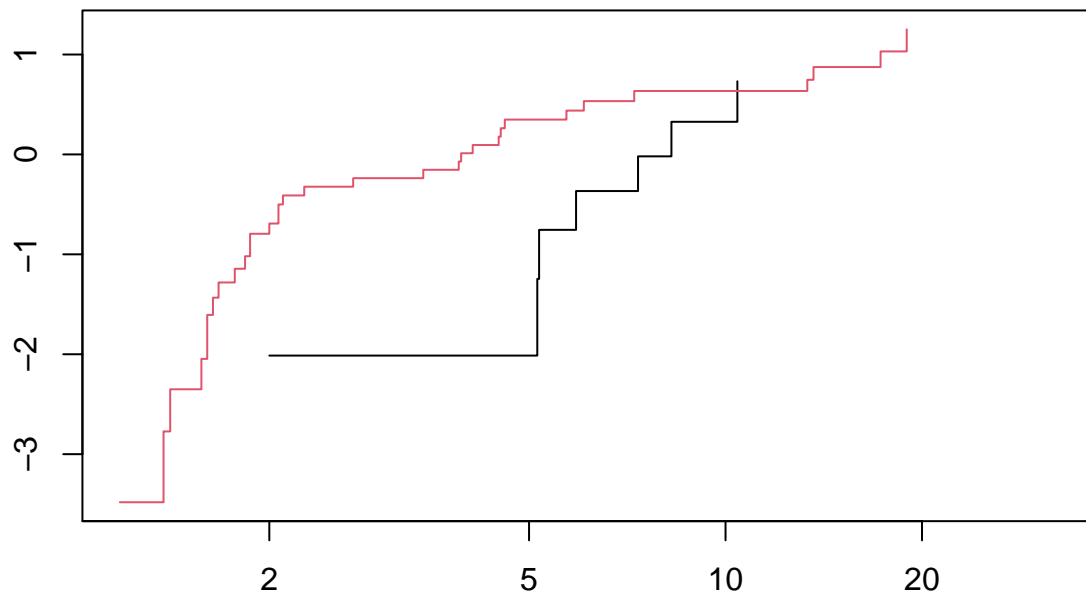
```
fit <- coxph(PFS ~ stage, data = dat)
summary(fit)
```

```
## Call:
## coxph(formula = PFS ~ stage, data = dat)
##
##      n= 41, number of events= 41
##
##              coef exp(coef) se(coef)      z Pr(>|z|)
## stageM 0.5931      1.8095   0.4007 1.48    0.139
##
##              exp(coef) exp(-coef) lower .95 upper .95
## stageM           1.81      0.5526   0.8251    3.969
##
## Concordance= 0.589 (se = 0.033 )
## Likelihood ratio test= 2.43 on 1 df,  p=0.1
## Wald test               = 2.19 on 1 df,  p=0.1
## Score (logrank) test = 2.25 on 1 df,  p=0.1
```

```
fit.KM <- survfit(PFS ~ stage, data = dat)
plot(fit.KM, col = 1:2)
```



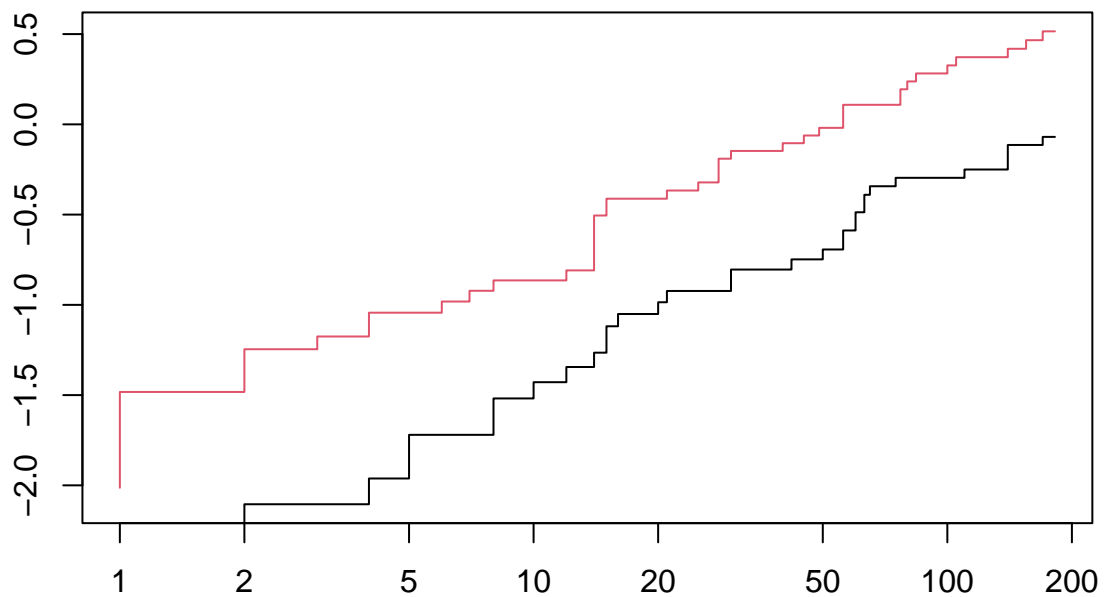
```
fit.KM <- survfit(PFS ~ stage, data = dat)
plot(fit.KM, fun= "cloglog", col = 1:2)
```

The figure indicates violation of proportionality of the hazards.

Pharmacosmoking dataset - treatment vs control

```
fit.KM <- survfit(Surv(ttr, relapse) ~ grp, data = pharmacoSmoking)
plot(fit.KM, fun = "cloglog", col = 1:2)
```



The figure indicates no violation of the assumption of proportionality of the hazards.

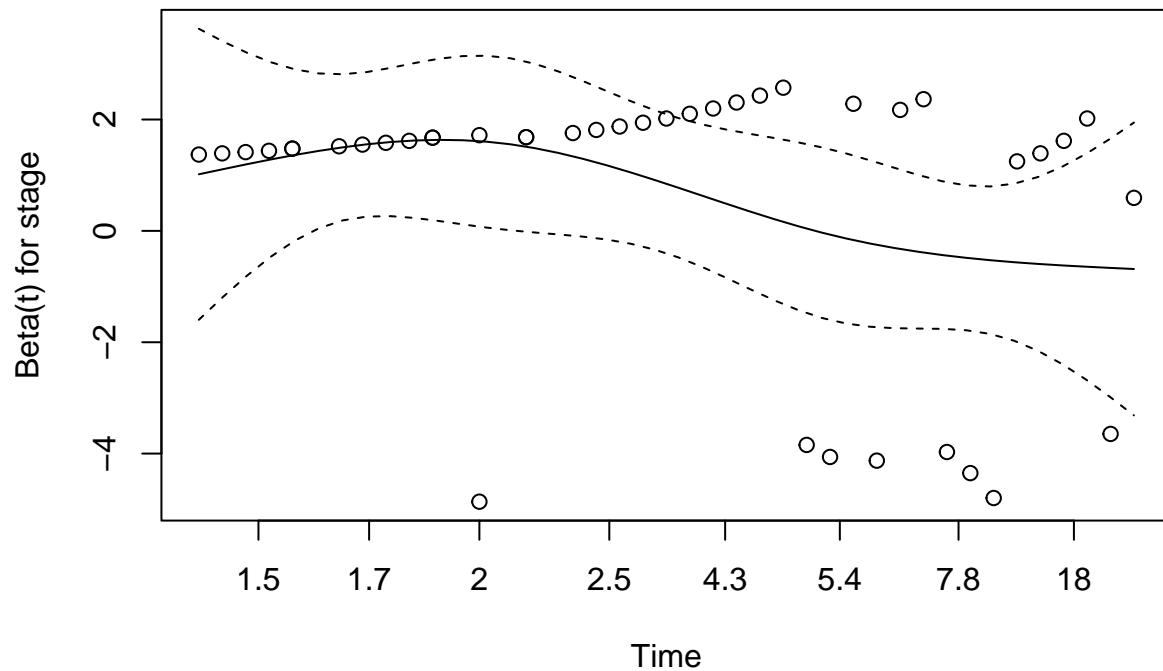
Schoenfeld residuals

```
fit <- coxph(PFS ~ stage, data = dat)
residual.sch <- cox.zph(fit)
residual.sch
```

```
##      chisq df    p
## stage  4.52  1 0.034
## GLOBAL  4.52  1 0.034
```

The null hypothesis is of constant covariates effects through time. Evidence against this hypothesis is also evidence against the assumption of proportionality of the hazards.

```
plot(residual.sch)
```



Dealing with assumptions violations

Stratification

```
library(asaur)
d <- pharmacoSmoking
d$employment <- ifelse(d$employment == "ft", "ft", "other")
table(d$employment)
```

```
##
##    ft other
##    72   53
```

Stratified Cox model:

```
fit <- coxph(Surv(ttr, relapse) ~ grp + strata(employment), data = d)
summary(fit)
```

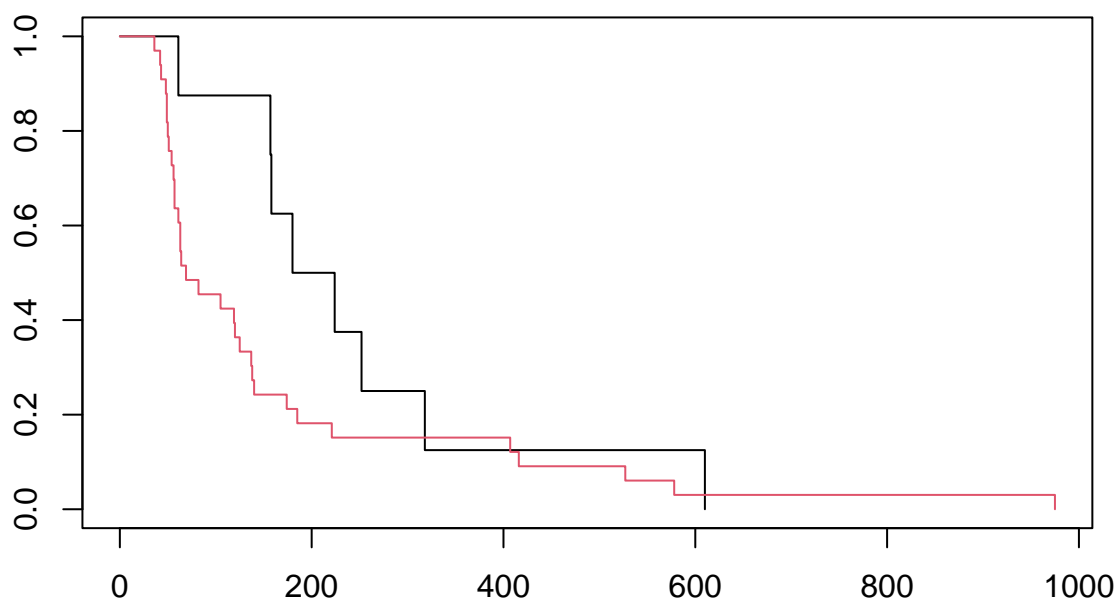
```
## Call:
## coxph(formula = Surv(ttr, relapse) ~ grp + strata(employment),
##       data = d)
##
## n= 125, number of events= 89
##
##               coef exp(coef) se(coef)      z Pr(>|z|)
## grppatchOnly 0.6391    1.8947   0.2187 2.922  0.00348 **
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
##               exp(coef) exp(-coef) lower .95 upper .95
## grppatchOnly    1.895    0.5278    1.234    2.909
##
## Concordance= 0.577 (se = 0.029 )
## Likelihood ratio test= 8.71 on 1 df,  p=0.003
## Wald test               = 8.54 on 1 df,  p=0.003
## Score (logrank) test = 8.81 on 1 df,  p=0.003
```

Note how there is no estimate associated with 'employment'.

Truncation

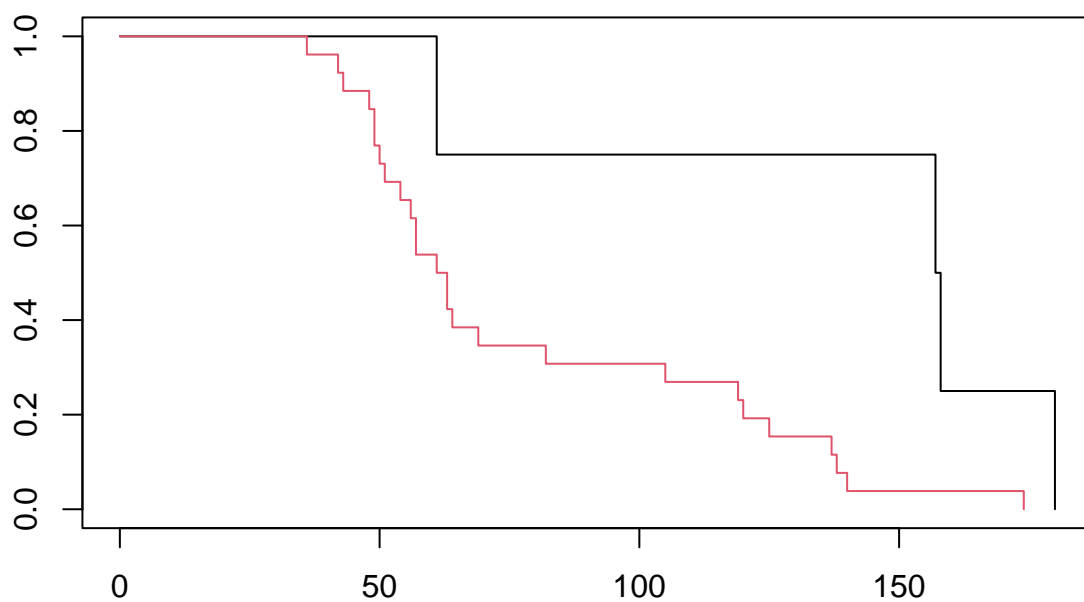
```
library(asaaur)
library(survival)
d <- pancreatic2

plot(survfit(Surv(pfs, status) ~ stage, data = d), col = 1:2)
```



THIS IS *NOT* HOW IT IS DONE:

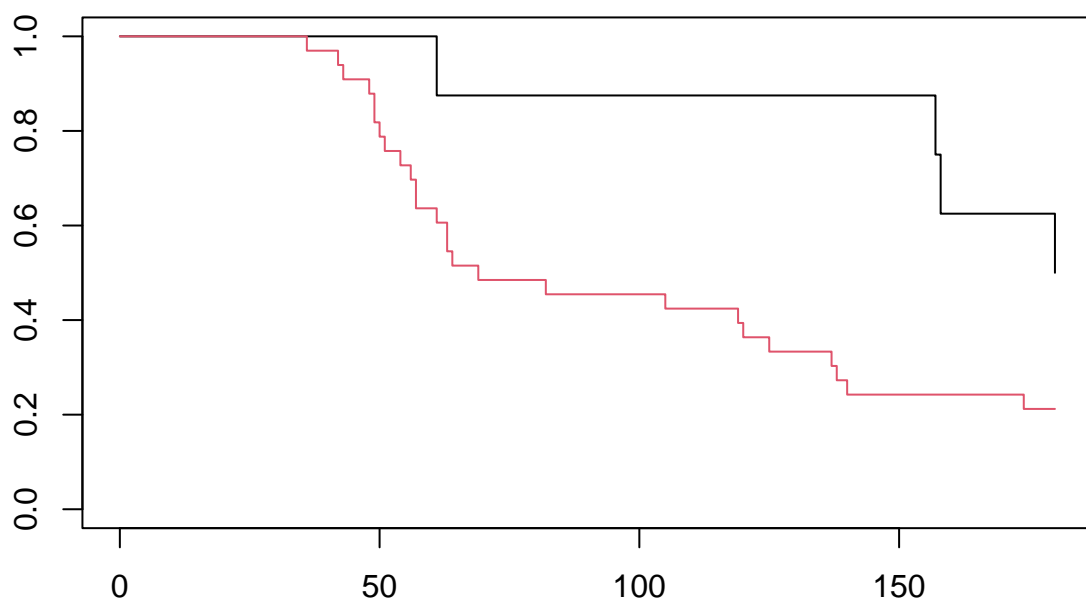
```
d_WRONG <- subset(d, pfs <= 180)
plot(survfit(Surv(pfs, status) ~ stage, data = d_WRONG), col = 1:2)
```



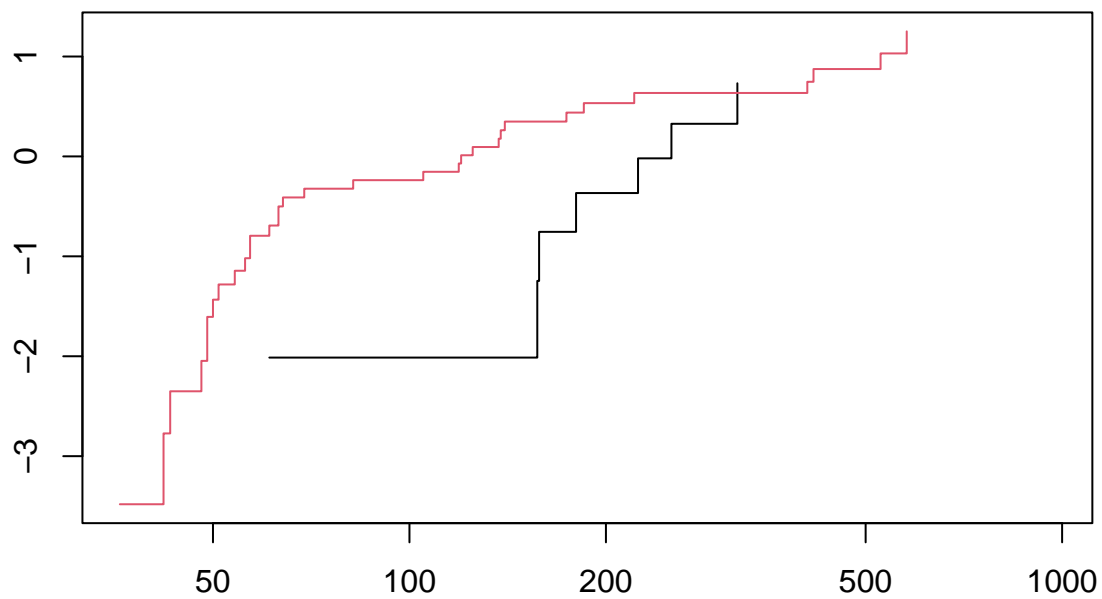
Here is how you do it:

```
d_RIGHT <- within(d, {
  status_truncated <- ifelse(pfs > 180, 0, status)
  pfs_truncated <- ifelse(pfs > 180, 180, pfs)
})
```

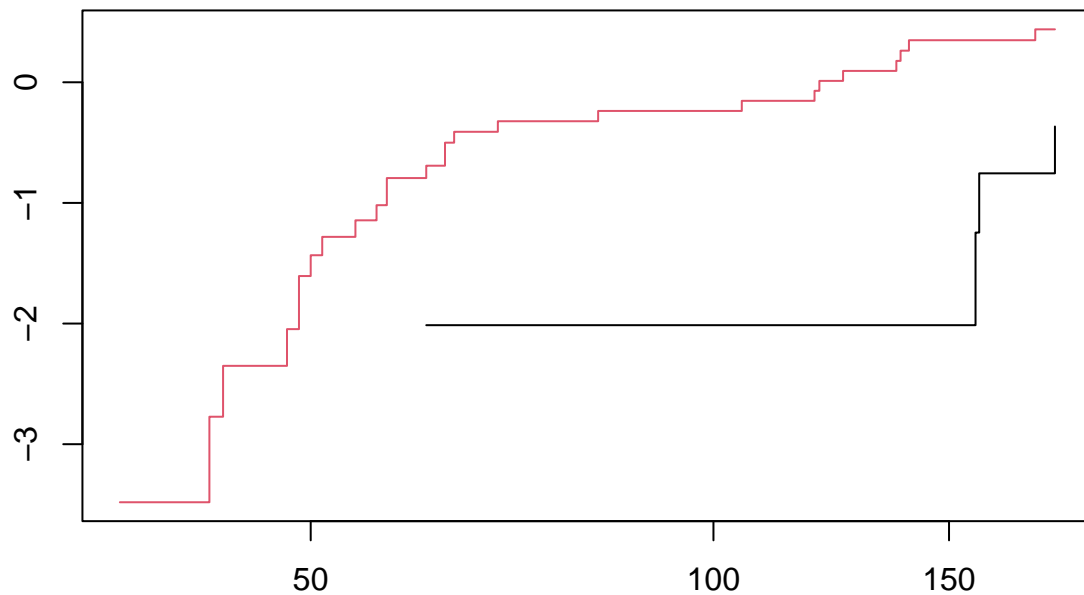
```
plot(survfit(Surv(pfs_truncated, status_truncated) ~ stage, data = d_RIGHT),
     col = 1:2)
```



```
plot(survfit(Surv(pfs, status) ~ stage, data = d_RIGHT),  
     fun = "cloglog",  
     col = 1:2)
```



```
plot(survfit(Surv(pfs_truncated, status_truncated) ~ stage, data = d_RIGHT),  
     fun = "cloglog",  
     col = 1:2)
```

```
summary(coxph(Surv(pfs_truncated, status_truncated) ~ stage, data = d_RIGHT))
```

```
## Call:
## coxph(formula = Surv(pfs_truncated, status_truncated) ~ stage,
##       data = d_RIGHT)
##
## n= 41, number of events= 30
##
##           coef exp(coef) se(coef)      z Pr(>|z|)
## stageM 1.0466   2.8479   0.5418 1.932   0.0534 .
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
##           exp(coef) exp(-coef) lower .95 upper .95
## stageM      2.848      0.3511   0.9848    8.236
##
## Concordance= 0.598 (se = 0.035 )
## Likelihood ratio test= 4.71 on 1 df,  p=0.03
## Wald test               = 3.73 on 1 df,  p=0.05
## Score (logrank) test = 4.07 on 1 df,  p=0.04
```