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# ECON 7201 Applied Econometrics

**Assignment 2** 

#### **Due Date**

**Sunday October 5, 2025** at 11:59 PM

#### **Directions**

Answer all questions. Submit both a PDF and Quarto file to the nexus assignment portal.

#### Git and GitHub

- 1. (a) Create a new R project in your **econ\_3201** directory called **assignment\_2**.
  - (b) Download the assignment PDF and Quarto file the **assignment 2** folder.
  - (c) Commit and push the changes to your **econ 3201** repository on GitHub.com.

#### LaTeX

Matrices are created in LaTeX using the \begin{bmatrix}...\end{bmatrix} command. To separate entries along the same row, use &. To end a line, use \\. To make vertical elipses (:), use \vdots. Practice writing the following matrices and vectors in LaTeX. Write the following matrices in LaTeX.

2. (a) 
$$X'X = \begin{bmatrix} n & \sum_{i=1}^{n} x_{1i} & \sum_{i=1}^{n} x_{2i} \\ \sum_{i=1}^{n} x_{1i} & \sum_{i=1}^{n} x_{1i}^{2i} & \sum_{i=1}^{n} x_{1i} x_{2i} \\ \sum_{i=1}^{n} x_{2i} & \sum_{i=1}^{n} x_{1i} x_{2i} & \sum_{i=1}^{n} x_{2i}^{2i} \end{bmatrix}$$
 (b)

(b) 
$$\Omega = \begin{bmatrix} \sigma_1^2 & 0 & 0 & 0 \\ 0 & \sigma_2^2 & 0 & 0 \\ 0 & 0 & \sigma_3^2 & 0 \\ 0 & 0 & 0 & \sigma_4^2 \end{bmatrix}$$

R

3. In this question we compare standard errors based on (incorrect) asymptotic assumptions with those based on alternate (appropriate) estimator (White). Consider one sample drawn from the following data generating process (DGP) which we will simulate in R:

```
set.seed(123)
n <- 25
x <- rnorm(n,mean=0.0,sd=1.0)
beta0 <- 1
beta1 <- 0
## x is irrelevant in this model, the data generating process is as follows:
dgp <- beta0 + beta1*x
## The residual is heteroskedastic by construction
e <- x^2*rnorm(n,mean=0.0,sd=1.0)
y <- dgp + e</pre>
```

(a) Compute the OLS estimator of  $\beta_2$  and its standard error using the lm() command in R for the model  $y_i = \beta_1 + \beta_2 x_i + \epsilon_i$  based on the DGP given above.

```
OLS_Estimator <-lm(y~x)
summary(OLS_Estimator)</pre>
```

#### Call:

 $lm(formula = y \sim x)$ 

Residuals:

```
Min
              1Q Median
                               3Q
                                      Max
-3.8764 -0.1604 0.0928 0.5748 2.1268
Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept) 0.8890
                          0.2379
                                  3.737 0.00108 **
               0.4610
                          0.2563
                                    1.799 0.08520 .
___
Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
Residual standard error: 1.189 on 23 degrees of freedom
Multiple R-squared: 0.1233,
                                 Adjusted R-squared:
F-statistic: 3.235 on 1 and 23 DF, p-value: 0.0852
      (b) Next, compute the standard error of \hat{\beta}_2 by computing \hat{\sigma}^2(X'X)^{-1} in R using
         matrix commands, and verify that the two standard error estimates are
         identical.
X \leftarrow cbind(1, x)
residuals <-resid(OLS_Estimator)</pre>
sigmasqr_hat <-sum(residuals^2/(n-2))</pre>
trasposeX.Xinv <-solve(t(X) %*% X)</pre>
VarianceMatrix <-sigmasqr_hat * trasposeX.Xinv</pre>
beta2_se <-sqrt(VarianceMatrix[2,2])</pre>
#Comparing the result of beta_se with the standard error of x from the OLS_Estimates regress
beta2_se
[1] 0.2563086
coef(summary(OLS_Estimator))["x", "Std. Error"]
[1] 0.2563086
cbind(matrix= as.numeric(beta2_se), coef(summary(OLS_Estimator))["x", "Std. Error"])
        matrix
```

[1,] 0.2563086 0.2563086

(c) Compute White's heteroskedasticity consistent covariance matrix estimator using matrices in R and report the White estimator of the standard error of  $\hat{\beta}_2$ . Compare this with that from 3 (a) above.

```
White_X <- cbind(1,x)
transposeX.Xinv <-solve(t(X)%*%X)</pre>
#Extracting the residuals from the original regression output
resid_extract <-resid(OLS_Estimator)</pre>
# Transforming the residuals
resid_extract_hat <-diag(resid_extract^2)</pre>
White_resid <- t(White_X)%*% resid_extract_hat %*% White_X
CovWhite <- transposeX.Xinv %*% White_resid %*% transposeX.Xinv
# Extracting the standard corrected standard errors from
White_SE_Beta2 <- sqrt(CovWhite[2,2])</pre>
# Comparing results
White_SE_Beta2
[1] 0.4529172
coef(summary(OLS_Estimator))["x", "Std. Error"]
[1] 0.2563086
cbind(White_SE_Beta2,coef(summary(OLS_Estimator))["x", "Std. Error"])
     White_SE_Beta2
[1,]
          0.4529172 0.2563086
```

4. Let  $\hat{\theta}$  be an estimator for the population parameter  $\theta$ .  $\hat{\theta}$  is said to be unbiased if  $E(\hat{\theta}) = \theta$ . That is, if the mean of the sampling distribution of  $\hat{\theta}$  is equal to the true population value.

Consider the model

$$y_i = \beta_0 + \beta_1 x_{1,i} - \beta_2 x_{2,i} + \epsilon_i.$$

Lets provide empirical evidence that the ordinary least squares estimators  $\hat{\beta}_0$ ,  $\hat{\beta}_1$ , and  $\hat{\beta}_2$  are unbiased estimators of  $\beta_0$ ,  $\beta_1$ ,  $\beta_2$ , respectively, using R.

(a) Set the seed to 1, i.e., set.seed(1).

```
#nrm(list=ls())
set.seed(1)
```

(b) Set the number of observations \$n=100\$

```
n=(100)
```

(c) Generate the following model  $\$y_i=2+3.5x_{1,i}-9.2x_{2,i}+\epsilon_i$ , where  $\$x_1\leq 1$ 

```
x_1 <- rnorm(n,mean=3,sd=sqrt(6))
x_2 <-rnorm(n,mean=2,sd=sqrt(4))
epsilon <- rnorm(n,mean=0,sd=sqrt(100))
y <-2+3.5*x_1-9.2*x_2+epsilon
model4c <- lm(y ~x_1+x_2)
summary(model4c)</pre>
```

```
Call:
```

```
lm(formula = y \sim x_1 + x_2)
```

#### Residuals:

```
Min 1Q Median 3Q Max -29.4359 -4.3645 0.0202 6.3692 26.3941
```

#### Coefficients:

```
Estimate Std. Error t value Pr(>|t|)
(Intercept) 2.5297 2.1508 1.176 0.242
x_1 3.5862 0.4766 7.524 2.72e-11 ***
x_2 -9.4673 0.5474 -17.295 < 2e-16 ***
```

```
Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1

Residual standard error: 10.43 on 97 degrees of freedom

Multiple R-squared: 0.7859, Adjusted R-squared: 0.7815

F-statistic: 178 on 2 and 97 DF, p-value: < 2.2e-16
```

(d) Estimate the model coefficients using the `lm()` command. (Search `?lm()` in the console

```
lm(y~x_1+x_2)$coefficients
```

```
(Intercept) x_1 x_2
2.529657 3.586182 -9.467334
```

(e) Using a `for()` loop, replicate the model above \$M=1000\$ times and save the coefficient

```
set.seed(1)
n=100
M=1000

estimates_hat <- matrix( , nrow=M, ncol=3)

for (i in 1:M){
    x_1hat <- rnorm(n,mean=3,sd=sqrt(6))
    x_2hat <-rnorm(n,mean=2,sd=sqrt(4))
    epsilonhat <- rnorm(n,mean=0,sd=sqrt(100))
    yhat <-2+3.5*x_1hat-9.2*x_2hat+epsilonhat
    model4c_hat <- lm(yhat ~x_1hat+x_2hat)
    summary(model4c_hat)

estimates_hat[i, ]<-coef(model4c_hat)
}

summary(model4c_hat)</pre>
```

```
Call:
lm(formula = yhat ~ x_1hat + x_2hat)

Residuals:
    Min    1Q    Median    3Q    Max
```

```
-23.736 -6.136 1.829 5.846 20.767
Coefficients:
           Estimate Std. Error t value Pr(>|t|)
(Intercept) -0.1526 1.7297 -0.088 0.93
x_1hat
            3.3185
                       0.3333 9.958 <2e-16 ***
x_2hat
            -8.6074
                       0.4425 -19.452 <2e-16 ***
Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
Residual standard error: 8.771 on 97 degrees of freedom
Multiple R-squared: 0.8461, Adjusted R-squared: 0.8429
F-statistic: 266.6 on 2 and 97 DF, p-value: < 2.2e-16
colnames(estimates_hat) <-c("Intercept", "Beta1_hat", "Beta2_hat" )</pre>
lm(yhat~x_1hat+x_2hat)$coefficients
(Intercept)
             x_1hat
                           x_2hat
 -0.152614
             3.318500 -8.607369
(f) Using `hist()`, plot the sampling distributions of the coefficient estimates, $\beta_1$;
```

## Sampling Distribution from for Beta-1 from 1000 iterations

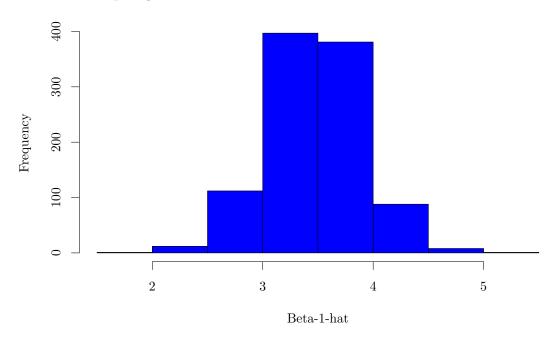


Figure 1: Sampling distribution from 1000 iterations for Beta-1-hat

```
hist(estimates_hat[, "Beta2_hat"],
    main="Sampling Distribution from for Beta-2-hat from 1000 iterations",
    xlab = "Beta-2-hat",
    col= "green")
```

## Sampling Distribution from for Beta-2-hat from 1000 iterations

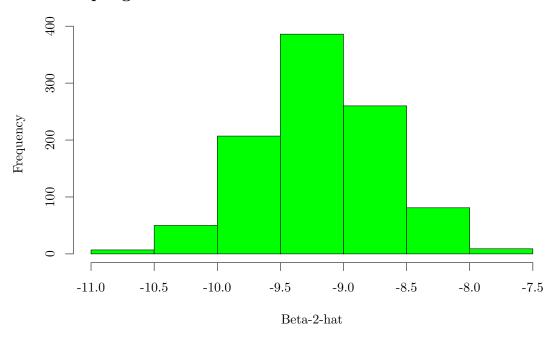


Figure 2: Sampling distribution from 1000 iterations for Beta-2-hat

(g) Add a vertical line to each figure at the mean of the respective variable. Search `?abli:

```
hist(estimates_hat[ , "Beta1_hat"],
    main= "Sampling Distribution from for Beta-1-hat from 1000 iterations",
    xlab = "Beta-1-hat",
    col= "blue")

abline(v= mean(estimates_hat[, "Beta1_hat"]),
    col= "red",
    lwd=2,
    lty="dashed")
```

## Sampling Distribution from for Beta-1-hat from 1000 iterations

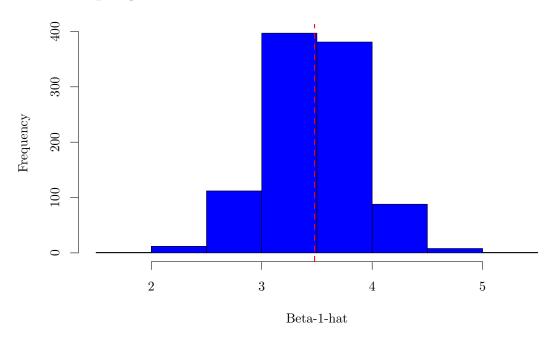


Figure 3: Sampling distribution from 1000 iterations for Beta-1-hat

```
hist(estimates_hat[ , "Beta2_hat"],
    main="Sampling Distribution from for Beta-2-hat from 1000 iterations",
    xlab = "Beta-2-hat",
    col= "green")

abline(v= mean(estimates_hat[, "Beta2_hat"]),
    col= "red",
    lwd=2,
    lty= "dashed")
```

## Sampling Distribution from for Beta-2-hat from 1000 iterations

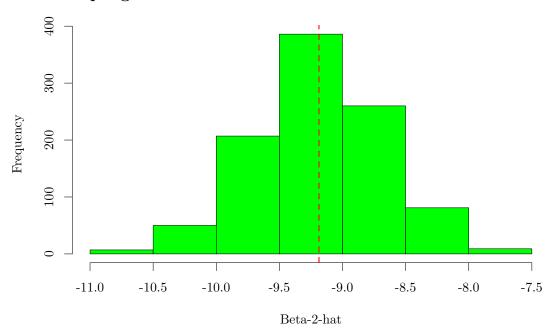


Figure 4: Sampling distribution from 1000 iterations for Beta-2-hat