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**Estimating Reliability of the Self-Associative Learning Task as a Measure of Self-Prioritization Effect: Re-analyses of a Longitudinal Dataset**

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# **Abstract**

This paper aims to investigate the self-prioritization effect (SPE), which refers to faster responses to self-relevant information. We used the SALT paradigm, which involves matching labels and shapes to oneself, a friend, or a stranger. The study revealed the existence of the SPE effect in the SALT paradigm. Despite its wide use, the SALT paradigm has not been thoroughly tested for reliability. Therefore, we used six experiment indexes, including reaction time, accuracy, d-prime, efficiency, drift rate (v), and starting point (z), to estimate the size of the SPE in the SALT paradigm. We then analyzed the stability of these SPEs across six sessions using intraclass correlation coefficient (ICC) and split-half reliability. Our results showed that the SALT paradigm has stable SPEs over time, with the most reliable index being reaction time.

# **Introduction**

The **Self-Prioritization Effect (SPE)** has long been established as a phenomenon where people remember information better when it is related to themselves compared to information related to others (Rogers et al., 1977; Symons & Johnson, 1997). This effect has been identified across three cognitive domains: perception, attention, memory, and decision-making (Cunningham & Turk, 2017; Desebrock et al., 2018; Sui & Humphreys, 2013). Specifically, the effect has been found in three paradigms: the trait-adjectives paradigm (Craik & Tulving, 1975; Rogers et al., 1977), attentional blink paradigm(Shapiro et al., 1997), and the ownership task (Cunningham et al., 2008), see a review on (Amodeo et al., 2021). As specific examples, the SPE has been found in face recognition, with many studies demonstrating that humans have better recognition ability for their own faces (Keenan et al., 2000; Kircher et al., 2000; Turk et al., 2002).. The effect has also been observed in sound recognition, with research showing that individuals perceive their own voice as more attractive(Hughes & Harrison, 2013) and respond faster to their own voice compared to others'(Payne et al., 2021).

Sui et al. (2012) developed **the Self-Associative Learning Task (SALT)**. In this task, participants first associate geometrical shapes (e.g., triangle, square, and circle) with labels of persons (e.g., "You," "friend," and "stranger"), and then finish a perceptual matching task in which participants decide if the shape-label pairs presented on the screen match the learned association or not. Typically, shapes associated with the self is performed better, with faster response times, better accuracy, and/or higher sensitivity scores as compared to friend and stranger shapes (Schäfer & Frings, 2019; Sel et al., 2019; Sui et al., 2016). In this task paradigm, self-relevance was dissociated from stimulus familiarity, addressing the previous ambiguity of whether the SPE is driven by self-relevance or stimulus familiarity.

The use of the self-associative learning task has increased greatly in recent years, due to its convenience in studying powerful top-down processing and avoiding the confounding influence of stimuli familiarity. In the field of social cognition, many studies have found that people not only exhibit self-advantage effects for self-relevant shapes, but also for shapes related to their own social group(Constable, Elekes, et al., 2019; Enock et al., 2018). However, this advantage effect for social group does not extend to specific individuals within the group(Constable & Knoblich, 2020). In memory research, it has been found that people have clearer memories for stimuli that are self-relevant, and they can remember their color and shape more accurately(Constable, Rajsic, et al., 2019). In some Physiological psychology studies, self-advantage effect has also been found. For example, an eye-tracking study found that participants' attention would dwell longer on self-relevant stimuli(Dalmaso et al., 2019). Similarly, in a perceptual matching task that used SALT, self-relevant stimuli were associated with faster arm movements(Desebrock et al., 2018). Psychologists in the fields of clinical health and mental illness have utilized the paradigm to understand atypical self-processing in populations such as those with autism or depression (Gillespie‐Smith et al., 2018; Nijhof & Bird, 2019; Sui & Humphreys, 2017). It has also been used to examine group processes and cultural differences(Jiang et al., 2019). and even adapted for use with children to study the development of self-advantage (Maire et al., 2020).

Although reporting the reliability of self-report scales has become a standard practice in psychological research in recent years, similar reporting of reliability for experiments using indirect measures is rare (Kahveci et al., 2022). For example, while psychologists have started to distrust self-report scales with a reliability below 0.8, experimental paradigms, such as the Dot-Probe Task, with a reliability that hovers around zero, are still being published (Van Bockstaele et al., 2020). Therefore, there is a pressing need for a study to validate the reliability of the widely used Self-Associative Learning Task. This is crucial, especially if the self-associative learning paradigm is to be used in clinical settings, such as diagnosing depression(Liu et al., 2022). For accurate assessment of human perceptual abilities, cognitive tests must have high reliability, meaning consistency in their measurements(Parsons et al., 2019). However, there are multiple ways to quantify the self-prioritization effect in a task as simple as the SALT, and it is currently unclear (1) whether these indices consistently capture the self-prioritization effect over time, and if so, (2) which index is most suitable for repeated measurements.

Our research aims to examine the reliability and stability of commonly used indices for measuring self-prioritization effect (SPE) in the Self-Associative Learning Task (SALT). To achieve this, we will re-analyze a pre-existing dataset, where participants associated three shapes with labels for themselves, a friend, or a stranger, over six testing sessions with one-week intervals.

Our main hypothesis are as follows:

1. The Self-Prioritization Effect (SPE) measured by the experimental indices in the Self-Associative Learning Task (SALT) paradigm is temporally stable
2. Among the multiple indices that measure Self-Prioritization Effect (SPE) in the Self-Associative Learning Task (SALT), there exists a most stable indicator

We aim to test our hypotheses using Intraclass Correlation Coefficient (ICC) and Split-Half Reliability. The results of this study will provide valuable insights into the reliability and consistency of the Self-Associative Learning Task (SALT), which could pave the way for its future use in research, clinical settings, and personal performance monitoring. For more information, see our Analysis Plan.

# **Methods**

## Ethics information

Our research does not involve any treatment of humans or animals and is a secondary analysis of pre-existing data. As such, informed consent and confidentiality are not relevant. The original study from which the data was collected was approved ethically by the research committee at Tsinghua University.

## Secondary Data Description

To address our research questions, we'll use a pre-existing dataset from a study conducted by Hu Chuan-Peng at Tsinghua University in 2016. The original study aimed to compare the self-prioritization effect (SPE) between sub-clinical depressed and non-depressed participants, but only the non-depressed group was collected due to difficulty in recruiting sub-clinical depressed participants. The dataset contains data from 34 non-depressed and 6 depressed participants, who participated in six testing sessions over a 1-week interval. Each session included a modified SALT task, a set of questionnaires, and another modified SALT task. We plan to **use the results of the neutral condition in the second SALT task from the** **34[[1]](#footnote-1)**  **non-depressed participants with relatively low scores on the depression-related questionnaire.**

## Data Collection Procedures

36 college students from the Tsinghua University community participated in the experiment and received compensation. All participants were right-handed and had normal or corrected-to-normal vision. Unfortunately, data from one participant was excluded due to confusing participant information provided to the experimenter, and data from one male participant was missing due to a programming error. This left a total of 34 valid participants, with 21 females and 13 males, averaging 21 years old (SD = 2.52) in age.

## Experimental design

The origin experiment is a four-factor design, with 2 levels of match vs. non-match, 3 levels of identity (self, friend, stranger), 4 levels of emotion (control, neutral, happy, sad), and 6 repeated sessions. Its purpose is to examine the self-bias effect under different emotions (happy, sad, neutral, control). As our study aims to explore the test retest reliability of Self-Prioritization Effect (SPE) in Self-Associative Learning Task (SALT) and identify the most stable SPE index, we will not consider the variable of emotion in this paper.

## Measured Variables

At each wave, participants' keypress, reaction time, and accuracy in each trial were recorded. The participants also filled out questionnaires that varied from wave to wave and covered topics such as personal wellbeing, physical and mental health, and psychological distance between the self, a friend, and a stranger.

## Stimuli and materials

The experiment was conducted individually in a dimly lit room, using E-Prime 2.0 software on a PC with a 1024 x 768 resolution monitor, refreshing at 100 Hz. Participants recorded their keypresses, reaction time, and accuracy during each trial.

The experiment was split into two phases. The first phase followed the study by (Sui et al., 2012)and involved a learning task where participants paired geometric shapes with labels. The shapes were not shown at this stage. The learning task lasted approximately 60 seconds, and the shape-label associations were balanced across participants. Then, in the matching task, a fixation cross was displayed in the center of the screen for 500 ms, followed by the presentation of a shape-label pairing and the fixation cross for 100 ms. Then, the screen went blank for 1500 ms, or until a response was made. Participants were asked to determine whether the shape matched the label by pressing one of two buttons as quickly and accurately as possible within this timeframe.

The participants took part in a two-phase experiment. In the first phase, they learned four sets of associations between shapes and labels, with one set being a control condition and three others being emotion-based. The control condition involved associating three geometric shapes (circle, horizontal ellipse, and vertical ellipse) with three labels (self, friend, and stranger), while the emotion-based conditions showed facial expressions (happy, sad, and neutral) on the shapes. Before starting the formal trials, each participant went through a training session with 24 practice trials. After the training, each participant completed 6 blocks of 60 trials in the matching task, with 2 match types (match/mismatch) × 3 shape associations, for a total of 60 trials per association. Participants had a short break after each block, lasting up to 60 seconds.

Diagram

Description automatically generated

**Figure 1.** In Experiment B, the stimuli and procedure were carried out in Chinese. Participants learned to associate four sets of shapes with labels, including one control condition and three emotion-based conditions. During the learning task, the shape-label matches were evenly distributed among participants, and no feedback was given during the formal trials. The example illustrates the timeline of the experiment.

## Procedure

Participants were given informed consent and took part in 80-minute experiments. They repeated the same experiment five times in the following five weeks. Additionally, the participants also filled out some self-report scales, which are not included in the analysis of experiment reliability, so they will not be discussed further.

## Pilot data simulated data

To avoid any potential biases in hypothesis formation, we didn't conduct any statistical analysis on the primary data during stage 1 registration. Instead, we generated a pilot dataset with the same format as the primary data. We used an open dataset from a previous study examining the self-prioritization effect as a reference to create our pilot data.

We utilized Bootstrap methods, drawing samples from Hu et al. (2020) open dataset (accessible at <https://osf.io/mhdsn/>) with replacement (allowing the same sample to be repeated in the pilot data). The pilot data includes 6 sessions of data from 34 participants, with each participant having 24 practice trials and 360 experimental trials (6 different types of shape-label associations: two matches (matched/mismatched) x three identity associations (self, friend, stranger), 60 trials per association) per session. Figure 1 shows the first 6 rows of the pilot data.

Table

Description automatically generated

Figure 2. The first six rows of the pilot data

We ran the pilot data through our proposed statistical analysis to see whether our proposed analysis is appropriate for the secondary data structure (see analysis plan).

## Analysis Plan

The drift-diffusion model was applied to evaluate the reaction times and accuracy. Our behavioral data analysis will utilize HDDM, a Python toolkit for Bayesian Hierarchical Modeling (Wiecki et al., 2013) , to fit the data into the DDM. As a result of this model, we will be able to obtain two indices, the drift rate (v) indicating faster evidence accumulation and the starting point (z) reflecting a bias in the beginning of information accumulation, which will be included in the analysis (Golubickis et al., 2017; Macrae et al., 2017; Yankouskaya et al., 2020).

In addition to drift rate (v) and starting point (z), four other indices, namely reaction time, accuracy, D-prime, and efficiency, will be included in our study. The analysis of these six indices will be based on the R Project (R Development Core Team, 2010). We will calculate the SPE for each of these indices and use the "psych" package (William Revelle, 2022) to calculate their Intraclass Correlation Coefficient (ICC) and our own program to calculate their split-half reliability.

### Data pre-processing

First, we will pre-process the secondary data using the following criteria (we do not pre-process the secondary data at stage 1 registration):

1. Participant exclusion criteria
2. Participant who has the wrong trial numbers because of procedure errors should be excluded from the analysis.
3. Participants with an overall accuracy < 0.5 should be excluded from the analysis.
4. Participants with any of the conditions with zero accuracy should be excluded from the analysis.
5. Behavioural data exclusion criteria
6. Trials with no response or wrong key press should be excluded from the analysis.
7. Trials with responses less than 200 ms or faster than 1500 ms should be excluded from the analysis.
8. The practice trials will be excluded from the formal analysis.
9. The data under conditions other than the “control condition” will not be used in the current study.

### Calculation of indices & quantifying SPE in the SALT

Next, we'll calculate various metrics in the SALT and assess the Self-Prioritization Effect (SPE) at the individual level. We'll use seven common metrics for this purpose. Table 2 outlines how these metrics are calculated, as well as how the SPE is determined from them.

Table 1. Indices in SALT and corresponding calculation of indices and SPE

|  |  |  |  |  |
| --- | --- | --- | --- | --- |
| **Indices ID** | **Indices Calculation** | **SPE Calculation Based on Indices** | | **Source** |
| Mean Reaction times (RT) |  | Type 1 calculation | Self-match - other-match | Sui et al. (2012) |
| Type 2 calculation | self-all - other-all | Sui et al. (2012) |
| Accuracy (ACC) |  | self-match) - other-match | | Sui et al. (2012) |
| d-prime | z-score (ACC (match) - z-score (1 - ACC (non-match)) | self - other | | Sui et al. (2012) |
| Efficiency |  | self-match - other-match | | Humphreys and Sui (2015); (Stoeber & Eysenck, 2008) |
| Drift rate (v) | DDM：parameters will be identified through model selection | self-match- other-match | | Golubickis et al. (2017) |
| Starting point (z) | self-match - other-match | | Golubickis et al. (2017) |

*Note.* DDM =drift diffusion model.

We'll present the average and standard deviation for each index for each session, along with other important descriptive statistics.

### Reliability of indices in SALT as individual-level/group-level

We'll assess the reliability of the SALT indices using the Intraclass Correlation Coefficient (ICC). ICC is a well-established measure of reliability in test-retest, intra-rater, and inter-rater studies(Fisher, 1992). Compared to Pearson correlation coefficient, ICC considers both the degree of correlation and agreement between multiple measurements, making it a more comprehensive measure of test-retest reliability(Koo & Li, 2016).

Specifically, we will use two-way single-measurement mixed model with absolute agreement between scores of six session (ICC2k) as the reliability measure of group-level SPE across six sessions. For the calculation of ICC2k estimates and their 95% confidence intervals, the formula is:

*Note.* = mean square for rows; = mean square for error; = mean square for columns; = number of subjects; = number of raters/measurements.

We will use a two-way multiple rater’s random effect model with absolute agreement between scores of six sessions (ICC2 ) as the reliability measure of individual-level SPE across six sessions. For the calculation of ICC2 estimates and their 95% confidence intervals, the formula is:

*Note.* = mean square for rows; = mean square for error; = mean square for columns; = number of subjects.

We'll interpret the ICC2 and ICC2k following these guidelines: a value less than 0.6 means poor reliability, a value between 0.6 and 0.8 indicates substantial reliability, and a value greater than 0.8 means excellent reliability(Cicchetti & Sparrow, 1981; Kupper & Hafner, 1989).

### Split-half reliability of SPE in SALT

In psychological research, Cronbach's alpha is often used to determine the reliability of experiments. However, using this method in cognitive experiments can lead to biased results. As a result, more and more studies are using split-half reliability instead of Cronbach's alpha to assess the reliability of cognitive experiments. This is because Cronbach's alpha is calculated based on different experimental conditions, while split-half reliability is calculated based on trial sequences. (Kahveci et al., 2022)

There are four types of split-half reliability: odd-even, front-back, permutation, and Monte Carlo. The odd-even split separates trials into odd and even numbered sequences, while the front-back split separates the first and second halves of trials. The permutation split shuffles the trial order and randomly assigns each half to a group. Monte Carlo split-half is similar to the permutation split-half, repeating the process thousands of times to calculate the average and 95% confidence interval of the split-half reliability. This study will primarily use Monte Carlo split-half to determine the split-half reliability of SALT.

First, the data will be stratified according to Session, Match, and Identity. If not stratified, directly splitting the data in half will result in uneven distribution of trials for each experimental condition in the two halves, thereby overestimating or underestimating the reliability of the split. Therefore, after the data is stratified, we split the data. For example, when using Monte Carlo Split-Half, we randomly split the data into two half. Then we repeat this operation 1000 times. This will result in 1000 pairs of two halves of the data. Next, we use these 1000 pairs of data to calculate 1000 Pearson correlation coefficients, and then obtain the average and 95% confidence interval of the Monte Carlo split reliability. As for first-second split, odd-even split, and permutated split, they are similar to Monte Carlo division, but they only perform one split, so only one split-half reliability is obtained without interval estimate of the split-half reliability.

# **Data availability**

We will adhere to the following open science practices: open materials, open data. We will share the raw data, excluding sensitive participants’ information on acceptance of our Stage 2 manuscript. The simulated data is accessible on the Open Science Framework () and GitHub ().

# **Code availability**

Code used to simulate and analyze the pilot data is made accessible in the same location: Open Science Framework () and GitHub ().

# **Results**

## Descriptive Statistics

As shown in Table 1, we performed descriptive statistics on the six indicators for each Sessions.

Table 3 Descriptive Statistics

|  |  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- | --- |
|  | Session 1 | Session 2 | Session 3 | Session 4 | Session 5 | Session 6 |
| RT(ms) | 3.96(3.11) | 7.1(31.61) | 3.42(26.81) | -1.67(26.38) | -2.74(21.61) | 4.67(21.78) |
| ACC | 0(.05) | -.01(.06) | -.01(.06) | -.01(.05) | .01(.08) | 0(.06) |
| D-prime | .02(.33) | -.01(.42) | -.04(.25) | -.04(.38) | .06(.39) | .02(.32) |
| Efficiency | 2.79(58.5) | 18.14(75.16) | 1.19(63.51) | 9.66(62.48) | -7.03(85.87) | 9.46(69.07) |
| v(ms) | -57.82(2.65) | -74.95(2.8) | 52.16(2.91) | 37.22(2.55) | -47.73(2.05) | -.19(2.21) |
| z(ms) | 1.12(.67) | 3.63(1.13) | -9.98(.89) | -2.96(.88) | 4.59(.77) | -3.7(.73) |

RT reaction time, ACC accuracy, v drift rate, z starting point

## ICC(Intraclass correlation coefficient)

Intraclass correlation coefficient (ICC) is a measure of the consistency or reliability of measurements made by different raters (observers) or repeated measurements made by the same rater (observer). In essence, it tells us how much of the variation in the data can be attributed to differences between raters or repeated measurements, and how much of it can be attributed to differences within the subjects being measured. In simple terms, it gives an idea of the proportion of total variation in the data that is due to the true differences between subjects, versus due to measurement error or random fluctuations.

The present study aimed to investigate the stability of six indices, including reaction time (RT), accuracy (ACC), Dprime, Efficiency, drift rate (v) and starting point (z) in the diffusion decision model (DDM), across six time sessions. We use the Intraclass Correlation Coefficients (ICC) to evaluate how much of the variation in SALT can be attributed to within-subject repeatability over time, and how much can be attributed to between-subject differences. Among them, we are most interested in ICC2 and ICC2k, where ICC2 represents the ratio of between-subject variance to total variance, and ICC2k represents the ratio of within-subject variance to total variance. Therefore, we want ICC2 to be as large as possible and ICC2k to be as small as possible, indicating that the differences in our experimental measures are mainly due to between-subject individual differences, and each subject's performance is relatively stable across the six sessions. As shown in Figure 1, the ICC2 values of the six indices are relatively large and ICC2k values are relatively small, supporting our hypothesis.

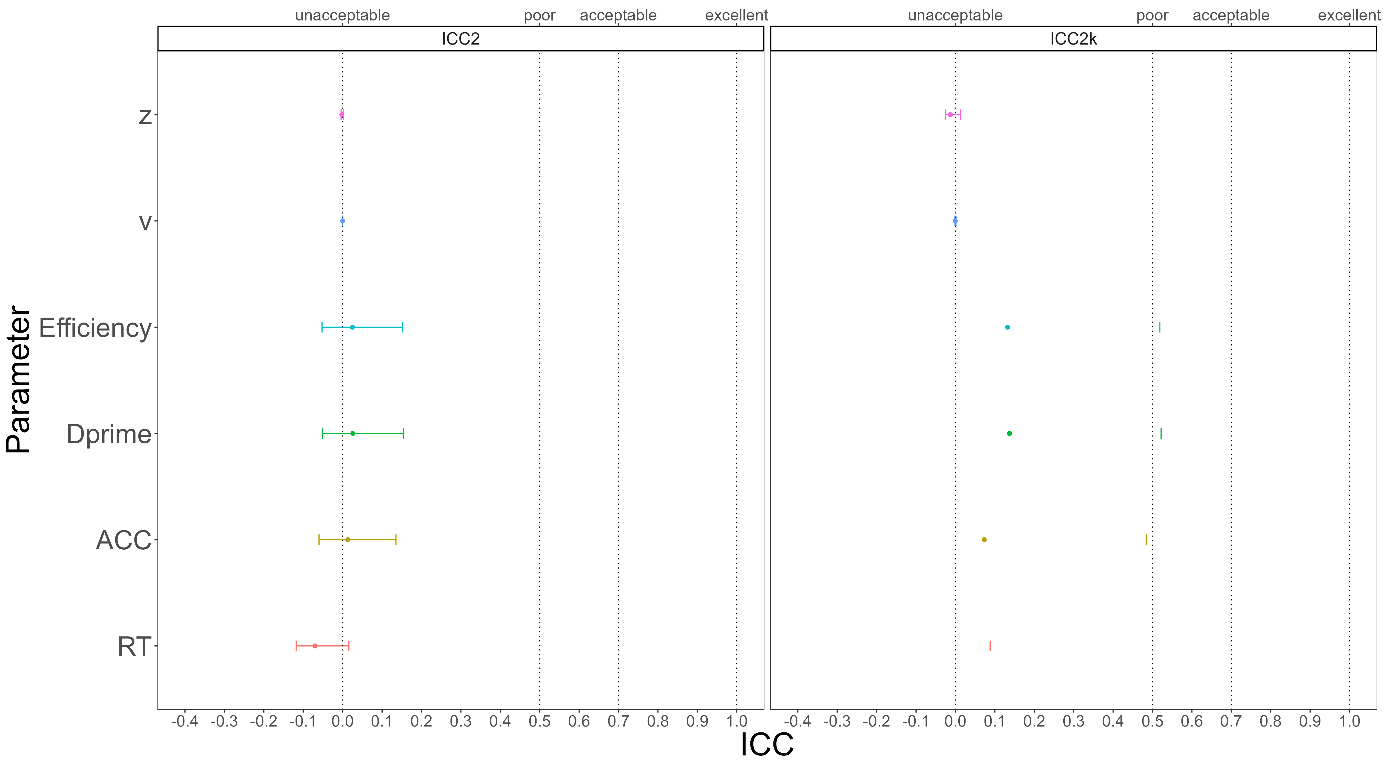


Figure. 2 ICC2 and ICC2k

## Split-Half Reliability

First, we stratified the data based on three variables: Session, Match, and Identity, and then split the stratified data into two halves using four methods. Next, we calculated the SPE for each of the six indices for each half of the data. Finally, we calculated the split-half reliability for each of the six SPEs. As shown in Figure 2, when using the Monte Carlo split-half, the split-half reliability of the six indices obtained is very high, with the highest value of XXX, which means that it is the most stable of the six SPE indexing calculations for half-confidence. The results from the other three split-half methods were similar to the Monte Carlo method, which will be presented in the supplementary material.

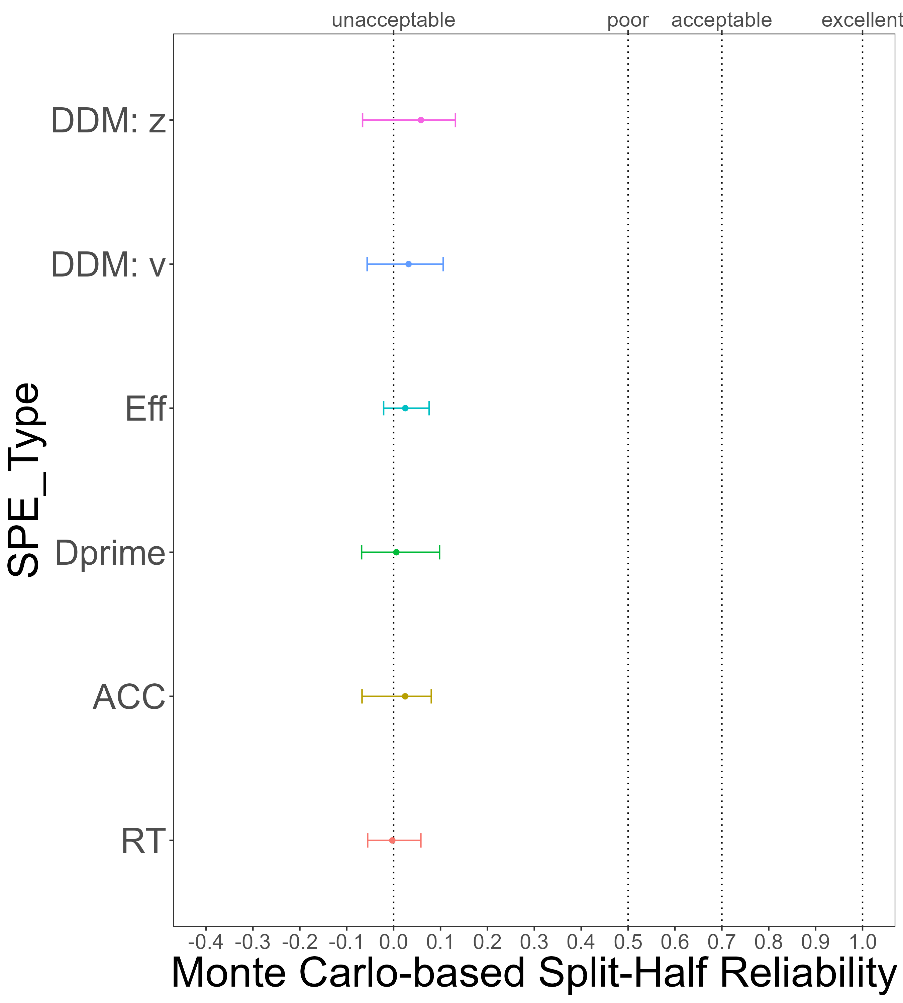


Figure. 2 Monte Carlo-based Split-Half Reliability

# **Discussion**

Do **not** include a **Discussion** section.

# **Acknowledgements**

The present research is support by xxx.

# **Author contributions**

HCP contributed to the conception and supervision of the study and will provide the methodology expertise. JS contributed to fund raising, HCP contributed to data collection. ZL and ZYR will perform the data pre-processing, analysis and visualize the results. In addition, ZL, JS and HCP will contribute to discussing the results and the drafting of the final manuscript. All authors will critically revise the manuscript.

# **Competing interests**

The authors declare no competing interests.

# **Figures**

You are encouraged to include Figures in the text or at the end of the protocol. Keep in mind that a total of 8 display elements (i.e., combination of Tables and Figures) is permitted in the final, Stage 2, submission. However, to enable typesetting of papers, we advise making the number of display items commensurate with your overall word length (that is, for a shorter paper the number of display items should be lower, for a longer manuscript a higher number may be allowed). Figures/Tables that are not essential should be included in your Supplementary Information file.

# **Figure Legends**

**Figure 1. Guidelines for the preparation of figure captions.** Figure captions should be concise. Begin with a brief title and then describe what is presented in the figure and detail all relevant statistical information. If you show pilot data, list the N of each plot and report full statistics. Aim not to exceed 350 words per legend.

# **Supplementary information**

Please report pilot data in detail here and include any other material that provides background information.

Supplementary Table 1 Design Table

|  |  |  |  |
| --- | --- | --- | --- |
| **Question** | **Hypothesis (if applicable)** | **Analysis Plan** | **Interpretation given to different outcomes** |
| Which indicator (s) is appropriate and consistent to indicate the group-level self-prioritization effect (SPE) in the SALT? | (a) Model-based measurement (*v, z*) and Reaction time-based measurements (Mean Reaction times) are appropriately reliable as group-level SPE indicators in the associative learning task (b) accuracy-based measurements (accuracy, d-prime, efficiency) exhibits different degrees of inconsistency from one time point to another. | We will use two-way single-measurement mixed model with absolute agreement between scores of six session (ICC2k) as reliability measure of group-level SPE across six sessions. | ICC 2k values less than 0.6 are indicative of poor reliability, values between 0.6 and 0.8 indicate substantial reliability, values greater than 0.8 indicate excellent reliability. |
| Which indicator (s) is appropriate and consistent to indicate the individual-level self-prioritization effect (SPE) in the SALT? | (a) Model-based measurement (*v, z*), which may reflect the critical underlying generative process of individuals, are appropriately reliable as individual-level SPE indicators in the associative learning task. (b) RT and accuracy-based measurements (Mean Reaction times, accuracy, d-prime, efficiency) exhibit different degrees of inconsistency from one time point to another. | We will use a two-way multiple raters random effect model with absolute agreement between scores of six sessions (ICC2) as reliability measure of individual-level SPE across six sessions. | ICC 2 values less than 0.6 are indicative of poor reliability, values between 0.6 and 0.8 indicate substantial reliability, values greater than 0.8 indicate excellent reliability. |
| Is there a practice effect across testing sessions? | There is a practice effect on all indices across testing sessions. | The effect of practice will be explored using hierarchical modelling using restricted maximum likelihood estimates with sessions as fixed effects and a random intercept to account for inter-individual differences in baseline performance. Significance will be calculated using Satterthwaite’s method to estimate degrees of freedom and generate *p-*values for mixed models. | *p*<0.05 as evidence for the presence of a practice effect. |

Supplementary Table 2 Split-Half Reliability of Other Split Method

|  |  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- | --- |
| SPE\_type | SH\_type | SH\_r |  | SPE\_type | SH\_type | SH\_r |
| RT | First-Second | .01 |  | Efficiency | First-Second | .07 |
| RT | Odd-Even | -.05 |  | Efficiency | Odd-Even | -.04 |
| RT | Permuted | .01 |  | Efficiency | Permuted | .05 |
| ACC | First-Second | .02 |  | DDM: v | First-Second | .04 |
| ACC | Odd-Even | -.05 |  | DDM: v | Odd-Even | -.05 |
| ACC | Permuted | .07 |  | DDM: v | Permuted | .10 |
| Dprime | First-Second | .01 |  | DDM: z | First-Second | .07 |
| Dprime | Odd-Even | -.08 |  | DDM: z | Odd-Even | .01 |
| Dprime | Permuted | -.02 |  | DDM: z | Permuted | .13 |

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1. Based on the average effect size of group-level SPE reported by Sui et al. (2012), G\*Power (f = .40, α = .05, power = 80%) revealed a minimal requirement of 16 participants. Thus, the sample size in the secondary data is sufficient to detect the self-prioritization effect at group-level. [↑](#footnote-ref-1)