

Maximum Likelihood Estimation: Theoretical Properties



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- MLE estimator is random variable (as a function of the random data):

$$\hat{\theta} = \hat{\theta}(x_1, \dots, x_n), \quad \{x_i\} \stackrel{iid}{\sim} p(\cdot | \theta^*)$$

- Evaluation metrics: Bias, variance, mean square error (MSE).
- Unbiased estimators vs. consistent estimators.

$$\text{Bias}(\hat{\theta}) = E[\hat{\theta}(x_1, \dots, x_n)] - \theta^*$$

If $\text{Bias}(\hat{\theta}) = 0$, we call $\hat{\theta}$ "unbiased".

$\text{MSE}(\hat{\theta}) \rightarrow 0$, as $n \rightarrow \infty$, $\hat{\theta}$ "consistent".

$\text{Bias}(\hat{\theta}) \rightarrow 0$, as $n \rightarrow \infty$, "Asymptotic Unbiased"

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$$\text{MSE}(\hat{\theta}) = (\text{Bias}(\hat{\theta}))^2 + \text{Var}(\hat{\theta})$$

Proof: $\text{MSE}(\hat{\theta}) = E[(\hat{\theta} - \theta^*)^2]$

$$= E[(\hat{\theta} - E(\hat{\theta}) + E(\hat{\theta}) - \theta^*)^2]$$

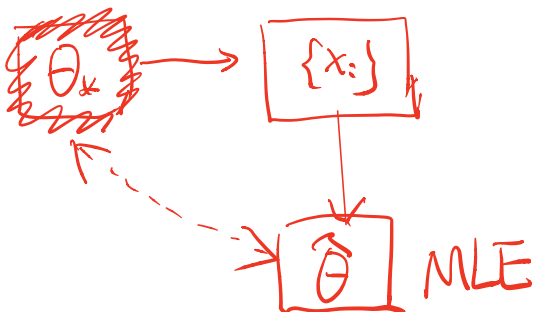
$$= E[(\hat{\theta} - E(\hat{\theta}))^2 + (E(\hat{\theta}) - \theta^*)^2 + 2(\hat{\theta} - E(\hat{\theta}))(E(\hat{\theta}) - \theta^*)]$$

$$= \text{Var}(\hat{\theta}) + (\text{bias}(\hat{\theta}))^2 = 0$$

$$E[(\hat{\theta} - E(\hat{\theta}))(E(\hat{\theta}) - \theta^*)] = E[\hat{\theta} - E(\hat{\theta})] (E(\hat{\theta}) - \theta^*)$$

$$= (E(\hat{\theta}) - E(\hat{\theta})) (E(\hat{\theta}) - \theta^*)$$

$$= 0$$



• Example: For $\{x_i\}_{i=1}^n \sim \mathcal{N}(\mu, \sigma^2)$, MLE is

"Unbiased"

$$\hat{\mu} = \frac{1}{n} \sum_{i=1}^n x_i, \quad \hat{\sigma}^2 = \frac{1}{n} \sum_{i=1}^n (x_i - \hat{\mu})^2.$$

$$\text{Bias}(\hat{\mu}) = E[\hat{\mu}] - \mu = E\left[\frac{1}{n} \sum_{i=1}^n x_i\right] - \mu = \frac{1}{n} \sum_{i=1}^n E[x_i] - \mu$$

$$\text{var}(\hat{\mu}) = E[(\hat{\mu} - \mu)^2] = \frac{1}{n} \sum_{i=1}^n \mu - \mu = 0$$

$$= E\left[\left(\frac{1}{n} \sum_{i=1}^n x_i - \mu\right)^2\right] = E\left[\frac{1}{n^2} \left(\sum_{i=1}^n (x_i - \mu)^2 + \sum_{i \neq j} (x_i - \mu)(x_j - \mu) \right)\right]$$

$$= \frac{1}{n^2} \sum_{i=1}^n E[(x_i - \mu)^2] + \sum_{i \neq j} E[(x_i - \mu)(x_j - \mu)]$$

$$= \frac{1}{n} \sigma^2 + \sum_{i \neq j} E[(x_i - \mu)E[(x_j - \mu)]] = 0$$

$$\text{MSE}(\hat{\mu}) = (\text{Bias}(\hat{\mu}))^2 + \text{var}(\hat{\mu}) = \frac{\sigma^2}{n}$$

If $n \rightarrow \infty$, $\text{MSE}(\hat{\mu}) \rightarrow 0 \Rightarrow \text{Consistent}$.

$$E[\hat{\sigma}^2] \neq \sigma^2, \quad \hat{\sigma}^2 = \frac{1}{n-1} \sum_{i=1}^n (x_i - \hat{\mu})^2$$

"Biased"

$$E[\hat{\sigma}^2] = \sigma^2$$

$\text{Bias}(\hat{\sigma}^2) \rightarrow 0$ as $n \rightarrow \infty$ (Asymp. Unbiased)

$\text{var}(\hat{\sigma}^2) \rightarrow 0$ as $n \rightarrow \infty$

$\text{MSE}(\hat{\sigma}^2) \rightarrow 0$ as $n \rightarrow \infty$ (Consistent)

Why MLE?

- MLE is equivalent to minimizing Kullback-Leibler (KL) Divergence.

$$KL(q \parallel p) = \mathbb{E}_q[\log q(x) - \log p(x)] = \begin{cases} \sum_x q(x) \left(\log \left(\frac{q(x)}{p(x)} \right) \right) \\ \int q(x) \log \left(\frac{q(x)}{p(x)} \right) dx \end{cases}$$

- $KL(q \parallel p) \geq 0$ for any q and p .
- $KL(q \parallel p) = 0$ if and only if $q = p$.

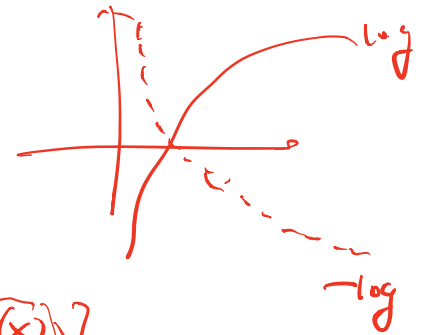
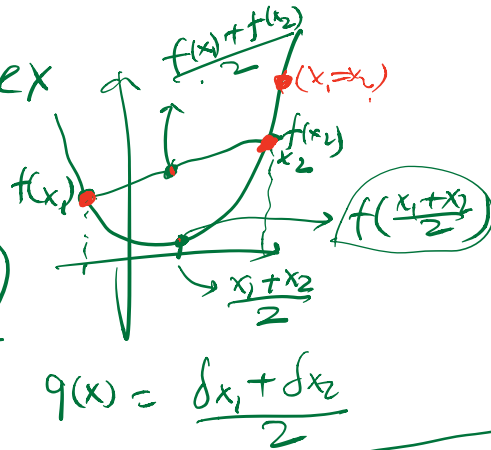
$$KL(q \parallel p) \neq KL(p \parallel q)$$

Jensen's Inequality:

If $f(x)$ is convex

then

$$\mathbb{E}_q[f(x)] \geq f(\mathbb{E}_q(x))$$



$$KL(q \parallel p) = \mathbb{E}_q[\log \left(\frac{q}{p} \right)] = \mathbb{E}_q[-\log \left(\frac{p(x)}{q(x)} \right)]$$

$$\geq -\log(\mathbb{E}_q[\frac{p(x)}{q(x)}])$$

$$\Rightarrow \frac{p(x)}{q(x)} = \text{const.}$$

$$= -\log \left(\sum_x q(x) \frac{p(x)}{q(x)} \right)$$

$$\Rightarrow p(x) = \text{const.} \cdot q(x)$$

$$= -\log \left(\sum_x p(x) \right) = -\log(1) = 0$$

$$\Rightarrow \sum_x p(x) = \text{const.} \cdot \sum_x q(x)$$

$$\Rightarrow 1 = \text{const.} \cdot 1 \Rightarrow \text{const.} = 1$$

$$\Rightarrow p = q$$

$$MLE \Leftrightarrow KL$$

$$KL(q \parallel p) = \mathbb{E}_q[\log q(x) - \log p(x)]$$

Assume q is data distribution (Given)

p_θ is "model" $\theta \in \Theta$

$$\min_{\theta} (KL(q \parallel p_\theta)) \Leftrightarrow \min_{\theta} \mathbb{E}_q[\log q(x) - \log p_\theta(x)]$$

$$\Leftrightarrow \min_{\theta} -E_q[\log P_{\theta}(x)]$$

$$\Leftrightarrow \boxed{\max_{\theta} E_q[\log P_{\theta}(x)]}$$

$$\approx \max_{\theta} \frac{1}{n} \sum_{i=1}^n \log P_{\theta}(x_i) \quad \{x_i\} \sim q$$

Avg log-likelihood.